6. Multivariate Distributions

Spring 2023

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Gov 2002 (Harvard)

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- Distributions of one variable: how to describe and summarize uncertainty about one variable.
- Today: distributions of multiple variables to describe relationships between variables.
- Later: use data to **learn** about probability distributions.

Why multiple random variables?

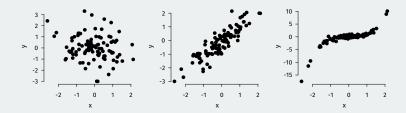
1. How to measure the relationship between two variables X and Y?

Why multiple random variables?

- 1. How to measure the relationship between two variables X and Y?
- 2. What if we have many observations of the same variable, X_1, X_2, \dots, X_n ?

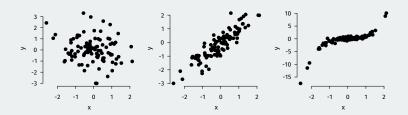
1/ Distributions of Multiple Random Variables

Joint distributions



 The joint distribution of two r.v.s, X and Y, describes what pairs of observations, (x, y) are more likely than others.

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- The joint distribution of two r.v.s, X and Y, describes what pairs of observations, (x, y) are more likely than others.
- Shape of the joint distribution \leadsto the relationship between X and Y

Definition

$$p_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

Definition

The **joint probability mass function (p.m.f.)** of a pair of discrete r.v.s, (X, Y) describes the probability of any pair of values:

$$p_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

• Properties of a joint p.m.f.:

Definition

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 - $p_{X,Y}(x,y) \geq 0$

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 - $\sum_{x} \sum_{y} p_{X,Y}(x,y) = 1$ (something must happen)
 - \sum_{x} is shorthand for sum over all possible values of X

	Support Gay	Oppose Gay
	Marriage	Marriage
	Y=1	Y = 0
Female $X = 1$	0.32	0.19
Male $X = 0$	0.29	0.20

• Joint p.m.f. can be summarized in a cross-tab:

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$$p_{X,Y}(1,1) = \mathbb{P}(X=1,Y=1) = 0.32$$

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 - Works because values of X are disjoint.

	Support Gay Marriage $Y = 1$	Oppose Gay Marriage $Y = 0$	Marginal
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Marginal			

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$\frac{\text{Male } X = 0}{\text{Marginal}}$	0.29	0.20	

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(Y=1)$$

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Marginal	0.32 + 0.29		

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Marginal	0.61	0.19 + 0.20	

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Conditional p.m.f.

Definition

The **conditional probability mass function** or conditional p.m.f. of Y conditional on X is

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(X = x)}$$

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$$P(Y = y \mid X = x) \ge 0$$
 and $\sum_{y} \mathbb{P}(Y = y \mid X = x) = 1$

• Can define the **conditional expectation** of this p.m.f.:

$$E[Y \mid X = x] = \sum_{y} y \mathbb{P}(Y = y \mid X = x)$$

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• Probability of favoring gay marriage conditional on **male**?

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$$\mathbb{P}(Y=1\mid X=0)$$

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Probability of favoring gay marriage conditional on male?

$$\mathbb{P}(Y = 1 \mid X = 0) = \frac{\mathbb{P}(X = 0, Y = 1)}{\mathbb{P}(X = 0)}$$

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	Marriage	Marriage	Marginal
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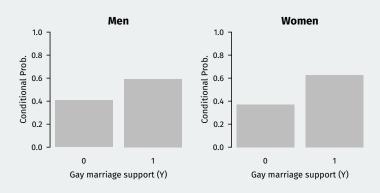
· Probability of favoring gay marriage conditional on male?

$$\mathbb{P}(Y=1 \mid X=0) = \frac{\mathbb{P}(X=0, Y=1)}{\mathbb{P}(X=0)} = \frac{0.29}{0.29 + 0.20}$$

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$$\mathbb{P}(Y=1 \mid X=0) = \frac{\mathbb{P}(X=0, Y=1)}{\mathbb{P}(X=0)} = \frac{0.29}{0.29 + 0.20} = 0.592$$



• Two values of $X \rightsquigarrow$ two **univariate** conditional distributions of Y

Bayes and LTP

· Bayes' rule for r.v.s:

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x \mid Y = y)\mathbb{P}(Y = y)}{\mathbb{P}(X = x)}$$

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· Bayes' rule for r.v.s:

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x \mid Y = y)\mathbb{P}(Y = y)}{\mathbb{P}(X = x)}$$

· Law of total probability for r.v.s:

$$\mathbb{P}(X = x) = \sum_{y} \mathbb{P}(X = x \mid Y = y) \mathbb{P}(Y = y)$$

Joint c.d.f.s

Definition

For two r.v.s X and Y, the **joint cumulative distribution function** or joint c.d.f. $F_{X,Y}(x,y)$ is a function such that for finite values x and y,

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- Well-defined for discrete and continuous X and Y.
- · For discrete we simply have:

$$F_{X,Y}(x,y) = \sum_{i \le x} \sum_{j \le y} \mathbb{P}(X = i, Y = j)$$

Continuous r.v.s

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 Two continuous r.v.s: probability of being in some subset of the 2-dimensional plane.



Continuous joint p.d.f.

Definition

If two continuous r.v.s X and Y with joint c.d.f. $F_{X,Y}$, their **joint p.d.f.** $f_{X,Y}(x,y)$ is the derivative of $F_{X,Y}$ with respect to x and y,

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• Integrate over both dimensions to get the probability of a region:

$$\mathbb{P}((X,Y)\in A)=\iint_{(x,y)\in A}f_{X,Y}(x,y)dxdy.$$

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$$\mathbb{P}((X,Y) \in A) = \iint_{(x,y) \in A} f_{X,Y}(x,y) dx dy.$$

• $\{(x,y): f_{X,Y}(x,y) > 0\}$ is called the **support** of the distribution.

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1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y),

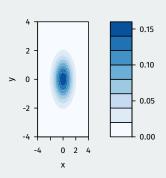
- Joint p.d.f. must meet the following conditions:
 - 1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y), (nonnegative)

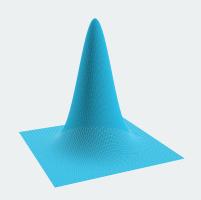
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- $\mathbb{P}(X = x, Y = y) = 0$ for similar reasons as with single r.v.s.

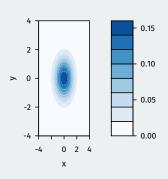
Joint densities are 3D

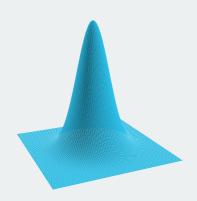




• X and Y axes are on the "floor," height is the value of $f_{X,Y}(x,y)$.

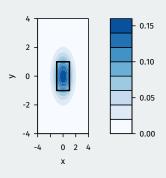
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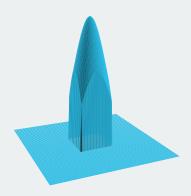




- X and Y axes are on the "floor," height is the value of $f_{X,Y}(x,y)$.
- Remember $f_{X,Y}(x,y) \neq \mathbb{P}(X=x,Y=y)$.

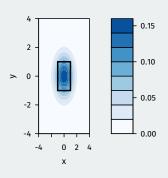
Probability = volume

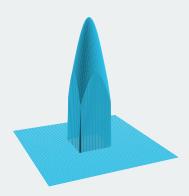




•
$$\mathbb{P}((X,Y) \in A) = \iint_{(x,y)\in A} f_{X,Y}(x,y) dx dy$$

Probability = volume





- $\mathbb{P}((X,Y) \in A) = \iint_{(X,Y) \in A} f_{X,Y}(x,y) dx dy$
- Probability = volume above a specific region.

Continuous marginal distributions

 We can recover the marginal PDF of one of the variables by integrating over the distribution of the other variable:

$$f_{Y}(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

Continuous marginal distributions

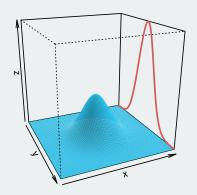
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· Works for either variable:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$

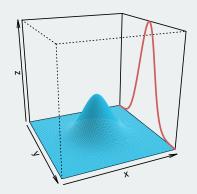
Visualizing continuous marginals



• Marginal integrates (sums, basically) over other r.v.:

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• Pile up/flatten all of the joint density onto a single dimension.

Continuous conditional distributions

Definition

The conditional p.d.f. of a continuous random variable is

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

for all values x s.t. $f_X(x) > 0$.

Implies

$$\mathbb{P}(a < Y < b | X = x) = \int_a^b f_{Y|X}(y|x) dy$$

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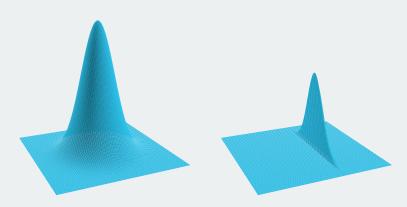
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 Based on the definition of the conditional p.m.f./p.d.f., we have the following factorization:

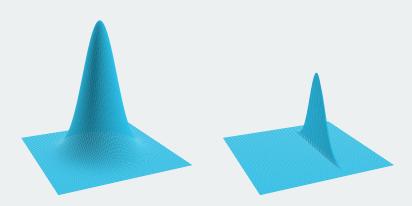
$$f_{X,Y}(x,y) = f_{Y|X}(y|x)f_X(x)$$

Conditional distributions as slices



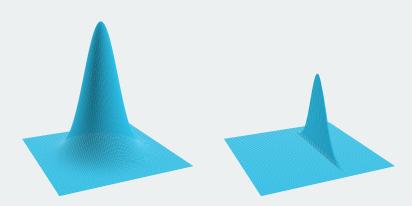
+ $f_{Y|X}(y|x_0)$ is the conditional p.d.f. of Y when $X=x_0$

Conditional distributions as slices



- $f_{Y|X}(y|x_0)$ is the conditional p.d.f. of Y when $X=x_0$
- + $f_{Y|X}(y|x_0)$ is proportional to joint p.d.f. along x_0 : $f_{X,Y}(y,x_0)$

Conditional distributions as slices



- $f_{Y|X}(y|x_0)$ is the conditional p.d.f. of Y when $X=x_0$
- $f_{Y|X}(y|x_0)$ is proportional to joint p.d.f. along x_0 : $f_{X,Y}(y,x_0)$
- Normalize by dividing by $f_X(x_0)$ to ensure proper p.d.f.

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp \!\!\! \perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

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- Knowing the value of X gives us no information about the value of Y.
- If X and Y are independent, then:
 - $f_{X,Y}(x,y) = f_X(x)f_Y(y)$ (joint is the product of marginals)

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp\!\!\!\perp Y$) if for all sets A and B:

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- Conditional independence implies similar to conditional distributions:

$$\mathbb{P}(X \in A, Y \in B \mid Z) = \mathbb{P}(X \in A \mid Z)\mathbb{P}(Y \in B \mid Z)$$

2/ Expectations of Joint Distributions

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$$\mathbb{E}[Y] = \sum_{x} \sum_{y} y \ f_{X,Y}(x,y)$$

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If X and Y are independent r.v.s, then

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3/ Covariance and Correlation

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 - In observational studies, treatment assignment is usually not independent of background characteristics.

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Covariance

The **covariance** between two r.v.s, *X* and *Y* is defined as:

$$\mathsf{Cov}[X,Y] = \mathbb{E}\Big[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\Big]$$

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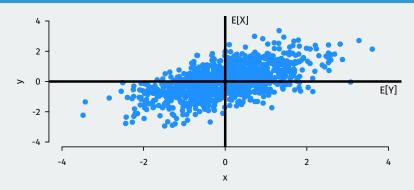
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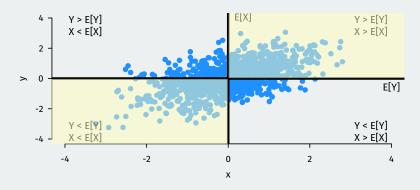
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 - $Cov[X, Y] = \mathbb{E}[XY] \mathbb{E}[X]\mathbb{E}[Y]$
 - If $X \perp \!\!\! \perp Y$, then Cov[X, Y] = 0

Covariance intuition

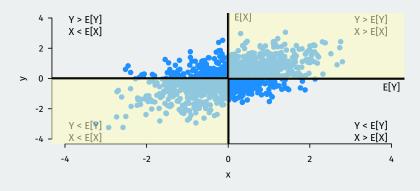


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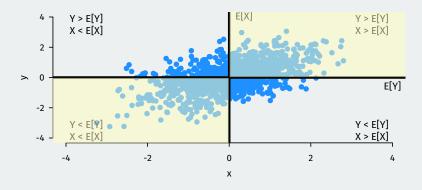


• Large values of *X* tend to occur with large values of *Y*:

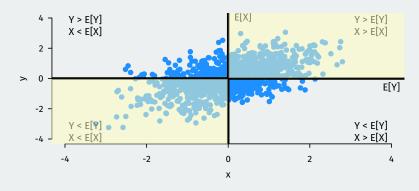
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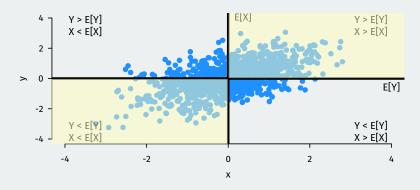
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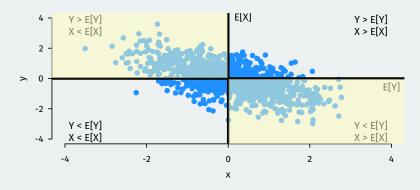
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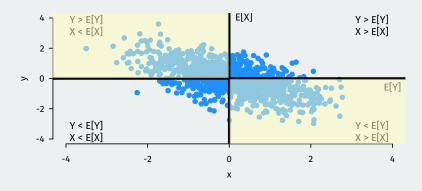
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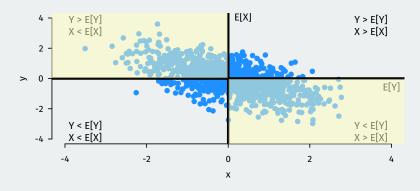
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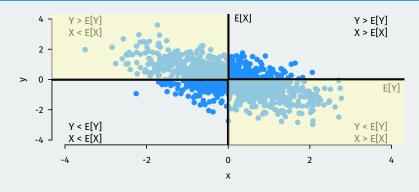
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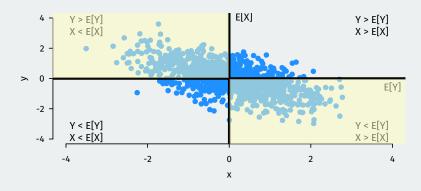
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 - 5. Cov[X + Y, Z] = Cov[X, Z] + Cov[Y, Z]
 - $\mathbf{6.} \ \, \mathsf{Cov}[X+Y,Z+W] = \mathsf{Cov}[X,Z] + \mathsf{Cov}[Y,Z] + \mathsf{Cov}[X,W] + \mathsf{Cov}[Y,W]$

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 - Beware: V[X Y] = V[X] + V[Y] as well.

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- Does Cov[X, Y] = 0 imply that $X \perp \!\!\! \perp Y$? **No!**
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- Covariance is a measure of linear dependence, so it can miss non-linear dependence.

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Definition

$$\rho = \rho(X, Y) = \frac{\mathsf{Cov}[X, Y]}{\sqrt{\mathbb{V}[X]\mathbb{V}[Y]}} = \mathsf{Cov}\left(\frac{X - \mathbb{E}[X]}{SD[X]}, \frac{Y - \mathbb{E}[Y]}{SD[Y]}\right)$$

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- · Covariance after dividing out the scales of the respective variables.
- · Correlation properties:
 - $-1 \le \rho \le 1$
 - $|\rho(X, Y)| = 1$ if and only if X and Y are perfectly correlated with a deterministic linear relationship: Y = a + bX.

4/ Random vectors

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• Expectation of a random vector is just the vector of expectations:

$$\mathbb{E}[X] = \left(\mathbb{E}[X_1], \mathbb{E}[X_2], \dots, \mathbb{E}[X_m]\right)^T$$

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• We usually write $V[X] = \Sigma$ and it is a $m \times m$ symmetric matrix:

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1m} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{m1} & \sigma_{m2} & \cdots & \sigma_m^2 \end{pmatrix}$$

where,
$$\sigma_j^2 = \mathbb{V}[X_j]$$
 and $\sigma_{ij} = \text{Cov}(X_i, X_j)$.

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$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1m} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{m1} & \sigma_{m2} & \cdots & \sigma_m^2 \end{pmatrix}$$

where,
$$\sigma_j^2 = \mathbb{V}[X_j]$$
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• Symmetric ($\Sigma = \Sigma^T$) because $Cov(X_i, X_i) = Cov(X_i, X_i)$.

Theorem

If $X \in \mathbb{R}^m$ with $m \times 1$ expectation μ and $m \times m$ covariance matrix Σ , and \mathbf{A} is a $q \times m$ matrix, then $\mathbf{A}X$ is a random vector with mean $\mathbf{A}\mu$ and covariance matrix $\mathbf{A}\Sigma\mathbf{A}^T$.

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• Expectation of a random vector is just the vector of expectations:

$$\mathbb{E}[\mathbf{X}] = (\mathbb{E}[X_1], \mathbb{E}[X_2], \dots, \mathbb{E}[X_k])'$$

· Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[\mathbf{X}] = \mathbb{E}\left[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])' \right]$$

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 - I_k is the k by k identity matrix because $\mathbb{V}[Z_j]=1$ and $\mathrm{Cov}(Z_i,Z_j)=0$.

Theorem

If $X \in \mathbb{R}^k$ with $k \times 1$ expectation μ and $k \times k$ covariance matrix Σ , and Δ is a $a \times k$ matrix, then ΔX is a random vector with mean $\Delta \mu$ and covariance matrix $\Delta \Sigma \Delta'$.

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• Let $\mathbf{Z} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_k)$ and $\mathbf{X} = \boldsymbol{\mu} + \mathbf{BZ}$, where \mathbf{B} is $q \times k$ then $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \mathbf{BB}')$

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 - μ : $q \times 1$ mean vector $\mathbb{E}[X] = \mu$
 - V[X] = BB': $q \times q$ covariance matrix.
- More generally, if $\mathbf{X} \sim \mathcal{N}(\pmb{\mu}, \pmb{\Sigma})$ then $\mathbf{Y} = \mathbf{a} + \mathbf{B} \mathbf{X} \sim \mathcal{N}(\mathbf{a} + \mathbf{B} \pmb{\mu}, \mathbf{B} \pmb{\Sigma} \mathbf{B}')$

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