# 15. Properties of Least Squares

Spring 2021

Matthew Blackwell

Gov 2002 (Harvard)

## Where are we? Where are we going?

• Before: learned about CEFs and linear projections in the population.

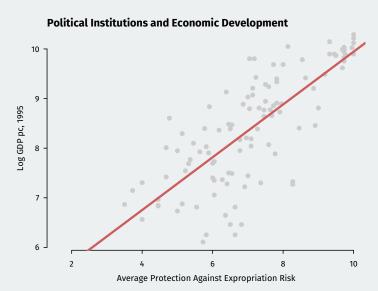
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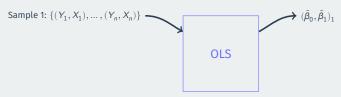
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- · Last time: OLS estimator, its algebraic properties.
- · Now: its statistical properties, both finite-sample and asymptotic.

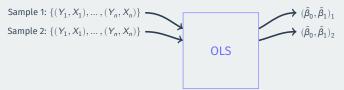
# Acemoglu, Johnson, and Robinson (2001)

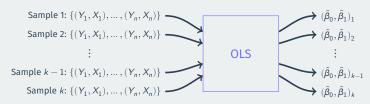


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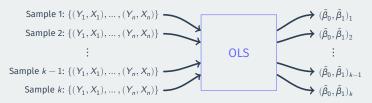
OLS



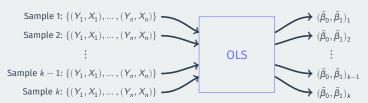




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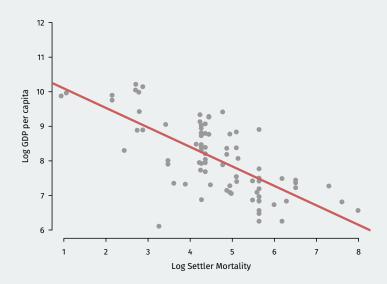
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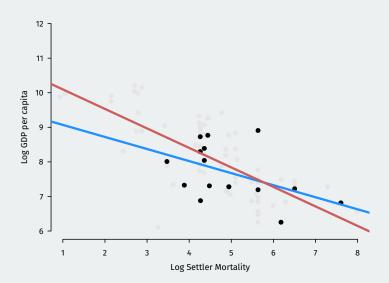
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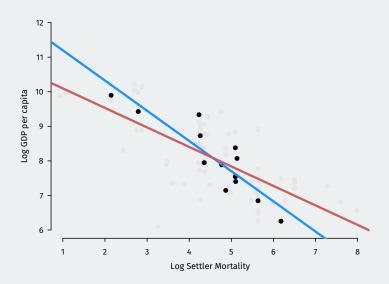
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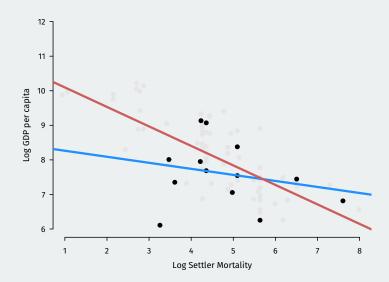
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- 3. Plot the estimated regression line

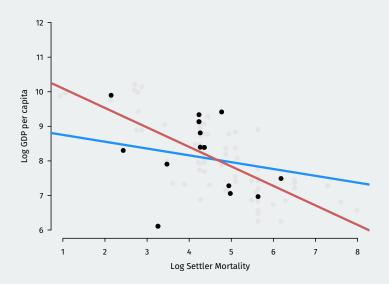
# **Population Regression**

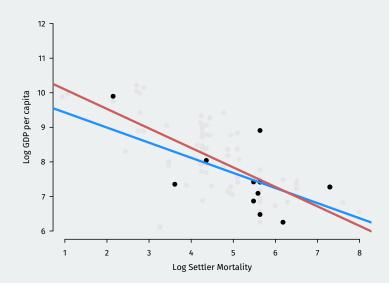


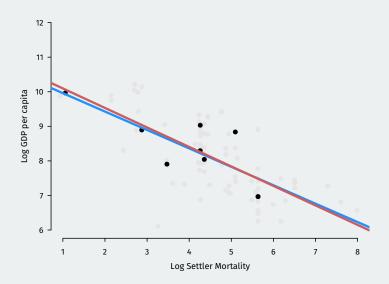


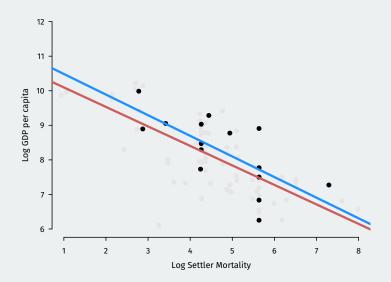












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### **Big picture**

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  - Linear projection model for asymptotic results.
  - Linear regression/CEF model for finite samples.

1/ Linear projection model and Large-sample Properties

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#### Linear projection model

$$Y = \mathbf{X}' \boldsymbol{\beta} + e$$
  
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#### Linear projection model

1. For the variables  $(Y, \mathbf{X})$ , we assume the linear projection of Y on  $\mathbf{X}$  is defined as:

$$Y = \mathbf{X}' \boldsymbol{\beta} + e$$
  $\mathbb{E}[\mathbf{X}e] = 0.$ 

2. The design matrix is invertible, so  $\mathbb{E}[XX'] > 0$  (positive definite).

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- What properties can we derive under such weak assumptions?

$$\hat{\boldsymbol{\beta}} = \left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} Y_{i}\right) = \boldsymbol{\beta} + \underbrace{\left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} e_{i}\right)}_{\text{estimation error}}$$

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- Sample means in the estimation error follow the law of large numbers:

$$\frac{1}{n}\sum_{i=1}^{n}\mathbf{X}_{i}\mathbf{X}_{i}'\overset{p}{\to}\mathbb{E}[\mathbf{X}_{i}\mathbf{X}_{i}']\equiv\mathbf{Q}_{\mathbf{X}\mathbf{X}}\qquad\frac{1}{n}\sum_{i=1}^{n}\mathbf{X}_{i}\mathbf{e}_{i}\overset{p}{\to}\mathbb{E}[\mathbf{X}\mathbf{e}]=\mathbf{0}$$

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 Q<sub>XX</sub> is invertible by assumption, so by the continuous mapping theorem:

$$\left(\frac{1}{n}\sum_{i=1}^{n}\mathbf{X}_{i}\mathbf{X}_{i}^{\prime}\right)^{-1}\overset{\rho}{\to}\mathbf{Q}_{\mathbf{XX}}^{-1}\quad\Longrightarrow\quad\widehat{\boldsymbol{\beta}}\overset{\rho}{\to}\boldsymbol{\beta}+\mathbf{Q}_{\mathbf{XX}}^{-1}\cdot\mathbf{0}=\boldsymbol{\beta}$$

Theorem (Consistency of OLS)

Under the linear projection model and i.i.d. data,  $\hat{\beta}$  is consistent for  $\beta$ .

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- Not guaranteed to be unbiased (unless CEF is linear, as we'll see...)

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$$\mathbb{E}\left[\frac{1}{n}\sum_{i=1}^{n}g(\mathbf{X}_{i})\right] = \mathbb{E}[g(\mathbf{X}_{i})] \quad \text{var}\left[\frac{1}{n}\sum_{i=1}^{n}g(\mathbf{X}_{i})\right] = \frac{\text{var}[g(\mathbf{X}_{i})]}{n}$$

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· CLT implies:

$$\sqrt{n}\left(\frac{1}{n}\sum_{i=1}^n g(\mathbf{X}_i) - \mathbb{E}[g(\mathbf{X}_i)]\right) \overset{d}{\to} \mathcal{N}(0, \mathrm{var}[g(\mathbf{X}_i)])$$

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• If  $\mathbb{E}[g(\mathbf{X}_i)] = 0$ , then we have

$$\sqrt{n}\left(\frac{1}{n}\sum_{i=1}^{n}g(\mathbf{X}_{i})\right) = \frac{1}{\sqrt{n}}\sum_{i=1}^{n}g(\mathbf{X}_{i}) \stackrel{d}{\to} \mathcal{N}(0,\mathbb{E}[g(\mathbf{X}_{i})g(\mathbf{X}_{i})'])$$

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  - Rewrite as  $\sqrt{n}$  times an average of i.i.d. mean-zero random vectors.
- Let  $\mathbf{\Omega} = \mathbb{E}[e_i^2 \mathbf{X}_i \mathbf{X}_i']$  and apply the CLT:

$$\left(\frac{1}{\sqrt{n}}\sum_{i=1}^{n}\mathbf{X}_{i}e_{i}\right)\overset{d}{\rightarrow}\mathcal{N}(0,\mathbf{\Omega})$$

#### Theorem (Asymptotic Normality of OLS)

Under the linear projection model,

$$\sqrt{n}\left(\hat{\pmb{\beta}} - \pmb{\beta}\right) \overset{d}{\to} \mathcal{N}(0, \mathbf{V}_{\hat{\pmb{\beta}}}),$$

where,

$$\mathbf{V}_{\hat{\boldsymbol{\beta}}} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1} \mathbf{\Omega} \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1} = \left( \mathbb{E}[\mathbf{X}_i \mathbf{X}_i'] \right)^{-1} \mathbb{E}[e_i^2 \mathbf{X}_i \mathbf{X}_i'] \left( \mathbb{E}[\mathbf{X}_i \mathbf{X}_i'] \right)^{-1}$$

•  $\hat{m{\beta}}$  is approximately normal with mean  $m{\beta}$  and variance  $m{Q}_{m{X}m{X}}^{-1}m{\Omega}m{Q}_{m{X}m{X}}^{-1}/n$ 

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- $\hat{m{\beta}}$  is approximately normal with mean  $m{\beta}$  and variance  $\mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{\Omega}\mathbf{Q}_{\mathbf{XX}}^{-1}/n$
- $\mathbf{V}_{\hat{\boldsymbol{\beta}}}/n$  is the **asymptotic covariance matrix** of  $\hat{\boldsymbol{\beta}}$

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- $\hat{m{\beta}}$  is approximately normal with mean  $m{\beta}$  and variance  $\mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{\Omega}\mathbf{Q}_{\mathbf{XX}}^{-1}/n$
- $V_{\hat{\beta}}/n$  is the asymptotic covariance matrix of  $\hat{\beta}$ 
  - Square root of the diagonal of  $\mathbf{V}_{\hat{\boldsymbol{\beta}}}/n$  = standard errors for  $\hat{\boldsymbol{\beta}}_j$

#### Theorem (Asymptotic Normality of OLS)

Under the linear projection model,

$$\sqrt{n}\left(\hat{\pmb{\beta}} - \pmb{\beta}\right) \overset{d}{\to} \mathcal{N}(0, \mathbf{V}_{\hat{\pmb{\beta}}}),$$

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  - Square root of the diagonal of  $\mathbf{V}_{\hat{\boldsymbol{\beta}}}/n$  = standard errors for  $\hat{\beta}_j$
- Allows us to formulate (approximate) confidence intervals, tests.

# **Estimating OLS variance**

$$\mathbb{V}[\hat{\boldsymbol{\beta}}] = \frac{1}{n} \mathbf{V}_{\hat{\boldsymbol{\beta}}} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1} \mathbf{\Omega} \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}$$

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  - Replace  $\mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1} = \mathbb{E}[\mathbf{X}_i\mathbf{X}_i']$  with  $n^{-1}\sum_{i=1}^n\mathbf{X}_i\mathbf{X}_i' = \mathbb{X}'\mathbb{X}/n$ .

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• Replace  $\mathbf{\Omega} = \mathbb{E}[e_i^2 \mathbf{X}_i \mathbf{X}_i']$  with  $n^{-1} \sum_{i=1}^n \hat{e}_i^2 \mathbf{X}_i \mathbf{X}_i'$ 

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- · Putting these together:

$$\begin{split} \widehat{\mathbb{V}}[\widehat{\boldsymbol{\beta}}] &= \frac{1}{n} \left( \frac{1}{n} \mathbb{X}' \mathbb{X} \right)^{-1} \left( \frac{1}{n} \sum_{i=1}^{n} \widehat{e}_{i}^{2} \mathbf{X}_{i} \mathbf{X}'_{i} \right) \left( \frac{1}{n} \mathbb{X}' \mathbb{X} \right)^{-1} \\ &= \left( \mathbb{X}' \mathbb{X} \right)^{-1} \left( \sum_{i=1}^{n} \widehat{e}_{i}^{2} \mathbf{X}_{i} \mathbf{X}'_{i} \right) \left( \mathbb{X}' \mathbb{X} \right)^{-1} \end{split}$$

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• Straightforward to show this is consistent:  $n\hat{\mathbb{V}}[\hat{\boldsymbol{\beta}}] \stackrel{p}{\to} \mathbf{V}_{\hat{\boldsymbol{\beta}}}$ .

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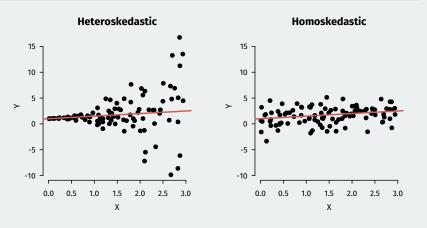
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- Square root of the diagonal of  $\hat{\mathbb{V}}[\hat{\boldsymbol{\beta}}]$ : heteroskedasticity-consistent (HC) **SEs**

## Homoskedasticity

#### Assumption: Homoskedasticity

The variance of the error terms is constant in  $\mathbf{X}$ ,  $\mathbb{E}[e^2 \mid \mathbf{X}] = \sigma^2(\mathbf{X}) = \sigma^2$ .



• Homoskedasticity implies  $\mathbb{E}[e_i^2 \mathbf{X}_i \mathbf{X}_i'] = \mathbb{E}[e_i^2] \mathbb{E}[\mathbf{X}_i \mathbf{X}_i'] = \sigma^2 \mathbf{Q}_{\mathbf{XX}}$ 

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- Simplifies the expression for the variance of  $\sqrt{n}(\hat{\beta} \beta)$ :

$$\mathbf{V}_{\hat{\boldsymbol{\beta}}}^{0} = \mathbf{Q}_{\mathbf{XX}}^{-1} \mathbb{E}[e_{i}^{2}] \mathbf{Q}_{\mathbf{XX}} \mathbf{Q}_{\mathbf{XX}}^{-1} = \sigma^{2} \mathbf{Q}_{\mathbf{XX}}^{-1}$$

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- Estimated variance of  $\hat{oldsymbol{eta}}$  under homoskedasticity

$$s^{2} = \frac{1}{n-k} \sum_{i=1}^{n} \hat{e}_{i}^{2} \qquad \hat{\mathbb{V}}^{0}[\hat{\boldsymbol{\beta}}] = \frac{1}{n} s^{2} \left( \frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}' \right)^{-1} = s^{2} \left( \mathbb{X}' \mathbb{X} \right)^{-1}$$

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• LLN implies  $s^2\stackrel{
ho}{ o}\sigma^2$  and so  $n\widehat{\mathbb{V}}^0[\widehat{\pmb{\beta}}]$  is consistent for  ${f V}^0_{\widehat{\pmb{\beta}}}$ 

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  - Mostly small, ad hoc changes to improve finite-sample performance.

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Software often uses t critical values instead of normal (we'll see why).

# 2/ Inference for Multiple Parameters

$$\mathbb{L}[Y|\mathbf{X}] = \boldsymbol{\beta}_0 + X\boldsymbol{\beta}_1 + Z\boldsymbol{\beta}_2 + XZ\boldsymbol{\beta}_3$$

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· Use the estimated covariance matrix:

$$\widehat{\mathbb{V}}\left(\frac{\partial \mathbb{L}\widehat{[Y|\mathbf{X}]}}{\partial X}\right) = \widehat{V}_{\beta_1} + Z\widehat{V}_{\beta_3} + 2Z\widehat{V}_{\beta_1\beta_3}$$

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• What about a test of no effect of X ever? Involves 2 coeffcients:

$$H_0: \beta_1 = \beta_3 = 0$$

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  - Compare model fit when we do and do not impose the null hypothesis.

#### **Restricted and unrestricted models**

• Estimates and SSR from unrestricted model (alternative is true):

$$\widehat{Y}_i = \widehat{\beta}_0 + X_i \widehat{\beta}_1 + Z_i \widehat{\beta}_2 + X_i Z_i \widehat{\beta}_3 \qquad SSR_u = \sum_{i=1}^n (Y_i - \widehat{Y}_i)^2$$

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  - $\bullet\,$  SSR mechanically increases even if you add noise, but not that much.

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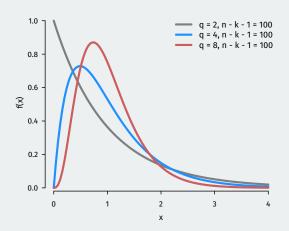
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- Asymptotic distribution of *F*:

$$\frac{(SSR_r - SSR_u)/q}{SSR_u/(n-k-1)} \stackrel{d}{\to} F_{q,n-(k+1)}$$

## **F** distribution



- Ratio of two  $\chi^2$  (Chi-squared) distributions

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  - When applied to a single coefficient, equivalent to a t-test.

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· Often reported with regression output.

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- · Illustration:
  - · Randomly draw 21 variables independently.
  - Run a regression of the first variable on the rest.
- By design, no effect of any variable on any other.

## Multiple test example

noise <- data.frame(matrix(rnorm(2100), nrow = 100, ncol = 21))
summary(lm(noise))</pre>

```
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.028039
                          0.113820
                                      -0.25
                                              0.8061
## X2
              -0.150390
                          0.112181
                                      -1.34
                                              0.1839
## X3
                0.079158
                          0.095028
                                     0.83
                                              0.4074
## X4
              -0.071742
                          0.104579
                                      -0.69
                                              0.4947
## X5
                0.172078
                          0.114002
                                      1.51
                                              0.1352
## X6
                0.080852
                           0.108341
                                      0.75
                                              0.4577
## X7
                0.102913
                          0.114156
                                      0.90
                                              0.3701
## X8
              -0.321053
                          0.120673
                                      -2.66
                                              0.0094 **
## X9
              -0.053122
                          0.107983
                                      -0.49
                                              0.6241
## X10
                0.180105
                          0.126443
                                      1.42
                                              0.1583
## X11
                0.166386
                           0.110947
                                      1.50
                                              0.1377
## X12
               0.008011
                          0.103766
                                      0.08
                                              0.9387
## X13
               0.000212
                          0.103785
                                      0.00
                                              0.9984
## X14
              -0.065969
                           0.112214
                                      -0.59
                                              0.5583
## X15
              -0.129654
                           0.111575
                                      -1.16
                                              0.2487
                                              0.6647
## X16
              -0.054446
                           0.125140
                                      -0.44
## X17
                0.004335
                           0.112012
                                     0.04
                                              0.9692
## X18
              -0.080796
                           0.109853
                                      -0.74
                                              0.4642
## X19
              -0.085806
                           0.118553
                                      -0.72
                                              0.4713
## X20
              -0.186006
                          0.104560
                                      -1.78
                                              0.0791 .
## X21
                0.002111
                          0.108118
                                     0.02
                                              0.9845
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.999 on 79 degrees of freedom
## Multiple R-squared: 0.201. Adjusted R-squared: -0.00142
## F-statistic: 0.993 on 20 and 79 DF. p-value: 0.48
```

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- 2 out of 20 variables significant at  $\alpha=0.1$
- Exactly the number of false positives we would expect.
- But notice the F-statistic: the variables are not **jointly** significant

3/ Linear Regression Model and Finite-sample Properties

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- We continue to maintain  $\{(Y_i, \mathbf{X}_i)\}$  are i.i.d.

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• Useful when linearity holds by default (discrete X in experiments, etc)

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- A matrix **C** is p.s.d. if  $\mathbf{x}'\mathbf{C}\mathbf{x} \geq 0$ .
- Upshot: OLS will have the smaller SEs than any other linear estimator.

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- · Software often implicitly assumes this for p-values.
- With reasonable *n*, asymptotic normality has the same effect.