11. (Linear) Regression

Spring 2023

Matthew Blackwell

Gov 2002 (Harvard)

Where are we? Where are we going?

- Learned about estimation and inference in general.

Where are we? Where are we going?

- Learned about estimation and inference in general.
- Now: building to a specific estimator, least squares regression.

Where are we? Where are we going?

- Learned about estimation and inference in general.
- Now: building to a specific estimator, least squares regression.
- First we need to understand what a “linear model” is and when/why we need it.

Where are we? Where are we going?

- Learned about estimation and inference in general.
- Now: building to a specific estimator, least squares regression.
- First we need to understand what a "linear model" is and when/why we need it.
    - No estimators quite yet. First, let's understand what we are estimating.

Where are we? Where are we going?

- Learned about estimation and inference in general.
- Now: building to a specific estimator, least squares regression.
- First we need to understand what a "linear model" is and when/why we need it.
    - No estimators quite yet. First, let's understand what we are estimating.
- Linear model is ubiquitous but poorly understood. Lots of subtlety here.

Regression derivatives and partial effects

- Goal of regression: how mean of Y changes with X.

$$\mu(\mathbf{x}) = \mathbb{E}[Y \mid \mathbf{X} = \mathbf{x}]$$

Regression derivatives and partial effects

- Goal of regression: how mean of Y changes with X.

$$\mu(\mathbf{x}) = \mathbb{E}[Y \mid \mathbf{X} = \mathbf{x}]$$

- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1, \dots, x_k)}{\partial x_1}$$

Regression derivatives and partial effects

- Goal of regression: how mean of Y changes with X.

$$\mu(\mathbf{x}) = \mathbb{E}[Y \mid \mathbf{X} = \mathbf{x}]$$

- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1, \ldots, x_k)}{\partial x_1}$$

- For binary X_1, we can use the difference in conditional expectations:

$$\mu(1, x_2, \ldots, x_k) - \mu(0, x_2, \ldots, x_k)$$

Regression derivatives and partial effects

- Goal of regression: how mean of Y changes with X.

$$\mu(\mathbf{x}) = \mathbb{E}[Y \mid \mathbf{X} = \mathbf{x}]$$

- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1, \ldots, x_k)}{\partial x_1}$$

- For binary X_1, we can use the difference in conditional expectations:

$$\mu(1, x_2, \ldots, x_k) - \mu(0, x_2, \ldots, x_k)$$

- "Partial effect" of X_1 holding other included variables constant

Regression derivatives and partial effects

- Goal of regression: how mean of Y changes with X.

$$\mu(\mathbf{x}) = \mathbb{E}[Y \mid \mathbf{X} = \mathbf{x}]$$

- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1, \dots, x_k)}{\partial x_1}$$

- For binary X_1, we can use the difference in conditional expectations:

$$\mu(1, x_2, \dots, x_k) - \mu(0, x_2, \dots, x_k)$$

- "Partial effect" of X_1 holding other included variables constant
- Exact form will depend on the functional form of $\mu(\mathbf{x})$.

Regression derivatives and partial effects

- Goal of regression: how mean of Y changes with X.

$$\mu(\mathbf{x}) = \mathbb{E}[Y \mid \mathbf{X} = \mathbf{x}]$$

- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1, \dots, x_k)}{\partial x_1}$$

- For binary X_1, we can use the difference in conditional expectations:

$$\mu(1, x_2, \dots, x_k) - \mu(0, x_2, \dots, x_k)$$

- "Partial effect" of X_1 holding other included variables constant
- Exact form will depend on the functional form of $\mu(\mathbf{x})$.
    - How do we decide what form $\mu(\mathbf{x})$ should take?

Estimating the CEF for discrete covariates

- To motivate function form, useful to think about estimation.

Estimating the CEF for discrete covariates

- To motivate function form, useful to think about estimation.
- How do we estimate $\mu(x) = \mathbb{E}[Y|X = x]$ for binary X?

Estimating the CEF for discrete covariates

- To motivate function form, useful to think about estimation.
- How do we estimate $\mu(x) = \mathbb{E}[Y|X = x]$ for binary X?
- **Subclassification**: calculate sample averages with levels of X_i:

$$\hat{\mu}(1) = \frac{1}{n_1} \sum_{i=1}^{n} Y_i X_i$$

Estimating the CEF for discrete covariates

- To motivate function form, useful to think about estimation.
- How do we estimate $\mu(x) = \mathbb{E}[Y|X = x]$ for binary X?
- **Subclassification**: calculate sample averages with levels of X_i:

$$\hat{\mu}(1) = \frac{1}{n_1} \sum_{i=1}^{n} Y_i X_i$$

  - $n_1 = \sum_{i=1}^{n} X_i$ is the number of units with $X_i = 1$ in the sample.

Estimating the CEF for discrete covariates

- To motivate function form, useful to think about estimation.
- How do we estimate $\mu(x) = \mathbb{E}[Y|X = x]$ for binary X?
- **Subclassification**: calculate sample averages with levels of X_i:

$$\hat{\mu}(1) = \frac{1}{n_1} \sum_{i=1}^{n} Y_i X_i$$

  - $n_1 = \sum_{i=1}^{n} X_i$ is the number of units with $X_i = 1$ in the sample.
- More generally for any discrete X_i:

$$\hat{\mu}(x) = \frac{\sum_{i=1}^{N} Y_i \mathbb{I}(X_i = x)}{\sum_{i=1}^{N} \mathbb{I}(X_i = x)}$$

Continuous covariates

- What if X is continuous? Subclassification fall apart.

Continuous covariates

- What if X is continuous? Subclassification fall apart.
    - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$

Continuous covariates

- What if X is continuous? Subclassification fall apart.
  - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates

Continuous covariates

- What if X is continuous? Subclassification fall apart.
  - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?

Continuous covariates

- What if X is continuous? Subclassification fall apart.
    - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
    - Very noisy estimates
    - What about any x not in the sample?
- **Stratification**: bin X_i into categories and treat like as discrete.

Continuous covariates

- What if X is continuous? Subclassification fall apart.
  - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin X_i into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.

Continuous covariates

- What if X is continuous? Subclassification fall apart.
  - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin X_i into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - Depends on arbitrary bin cutoffs/sizes.

Continuous covariates

- What if X is continuous? Subclassification fall apart.
  - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin X_i into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - Depends on arbitrary bin cutoffs/sizes.
- Example:

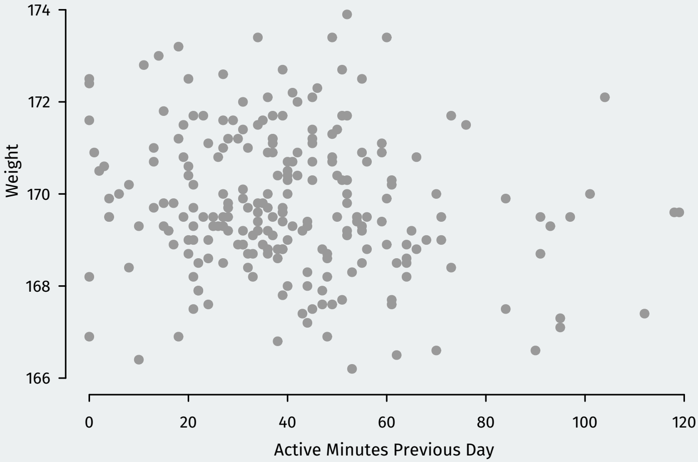
Continuous covariates

- What if X is continuous? Subclassification fall apart.
  - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin X_i into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - Depends on arbitrary bin cutoffs/sizes.
- Example:
  - Personal data science: I wear an activity tracker and have a smart scale.

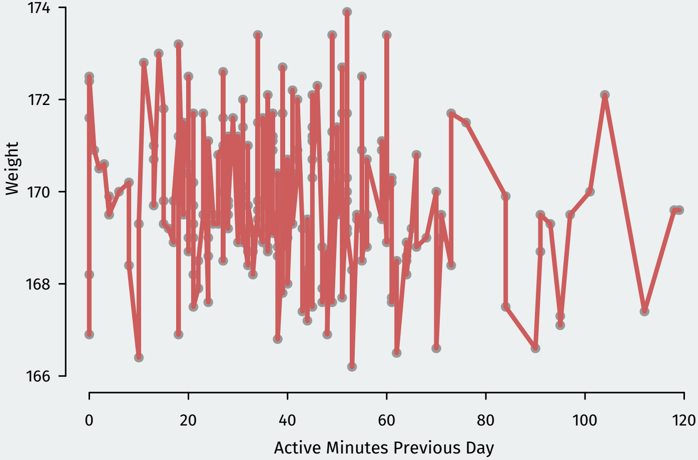
Continuous covariates

- What if X is continuous? Subclassification fall apart.
    - Each i has a unique value: $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
    - Very noisy estimates
    - What about any x not in the sample?
- **Stratification**: bin X_i into categories and treat like as discrete.
    - Every x in the same bin gets the same conditional expectation.
    - Depends on arbitrary bin cutoffs/sizes.
- Example:
    - Personal data science: I wear an activity tracker and have a smart scale.
    - Relationship between my weight and active minutes in the previous day.

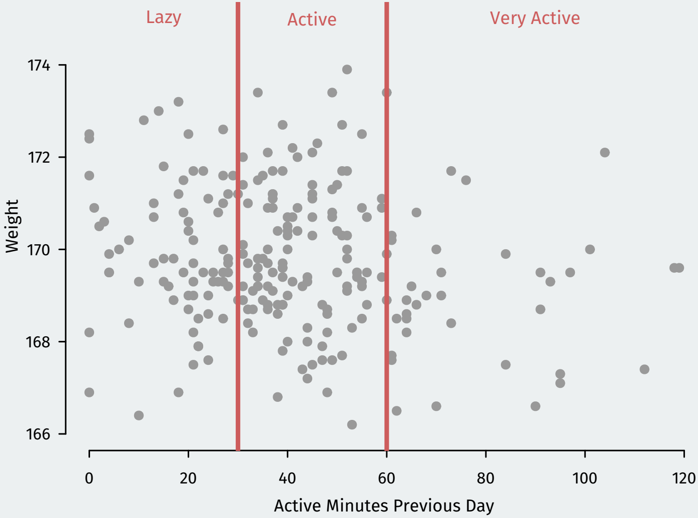
Continuous covariate example



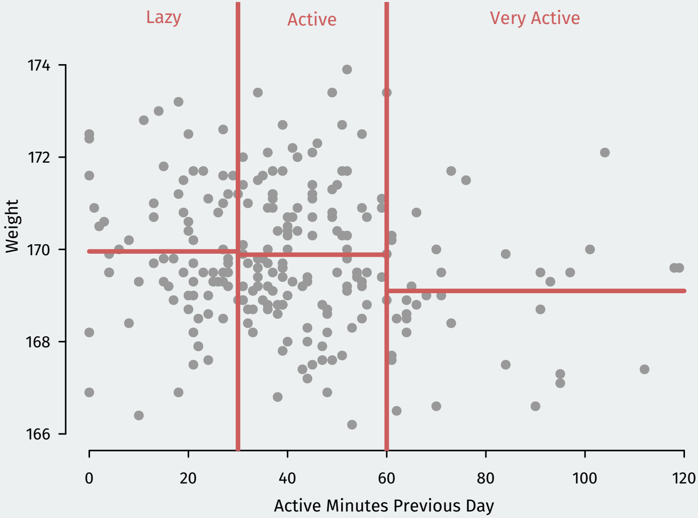
Continuous covariate CEF: interpolation



Continuous covariate CEF: stratification



Continuous covariate CEF: stratification



Linear CEFs

- Statification requires lots of choices/hidden assumptions.

Linear CEFs

- Statification requires lots of choices/hidden assumptions.
    - Number of categories, cutoffs for the categories, constant means within strata, etc.

Linear CEFs

- Statification requires lots of choices/hidden assumptions.
    - Number of categories, cutoffs for the categories, constant means within strata, etc.
- Alternative: assuming that the CEF is **linear**:

$$\mu(x) = \mathbb{E}[Y_i | X_i = x] = \beta_0 + \beta_1 x$$

Linear CEFs

- Statification requires lots of choices/hidden assumptions.
  - Number of categories, cutoffs for the categories, constant means within strata, etc.
- Alternative: assuming that the CEF is **linear**:

$$\mu(x) = \mathbb{E}[Y_i|X_i = x] = \beta_0 + \beta_1 x$$

- **Intercept**, β_0: the condition expectation of Y_i when $X_i = 0$

Linear CEFs

- Statification requires lots of choices/hidden assumptions.
  - Number of categories, cutoffs for the categories, constant means within strata, etc.
- Alternative: assuming that the CEF is **linear**:

$$\mu(x) = \mathbb{E}[Y_i | X_i = x] = \beta_0 + \beta_1 x$$

- **Intercept**, β_0: the condition expectation of Y_i when $X_i = 0$
- **Slope**, β_1: change in the CEF of Y_i given a one-unit change in X_i

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
    - β_0: average income among people with 0 years of education.

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
  - β_0: average income among people with 0 years of education.
  - β_1: expected difference in income between two adults that differ by 1 year of education.

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
    - β_0: average income among people with 0 years of education.
    - β_1: expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
    - β_0: average income among people with 0 years of education.
    - β_1: expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11]$$

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
  - β_0: average income among people with 0 years of education.
  - β_1: expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] = \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15]$$

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
  - β_0: average income among people with 0 years of education.
  - β_1: expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

$$
\begin{aligned}
\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] &= \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15] \\
&= \beta_1
\end{aligned}
$$

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
  - β_0: average income among people with 0 years of education.
  - β_1: expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

$$\begin{aligned}\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] &= \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15] \\ &= \beta_1\end{aligned}$$

- Effect of HS degree is the same as the effect of college degree.

Why is linearity an assumption?

- Example: Y_i is income, X_i is years of education.
  - β_0: average income among people with 0 years of education.
  - β_1: expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

$$\begin{aligned}\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] &= \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15] \\ &= \beta_1\end{aligned}$$

- Effect of HS degree is the same as the effect of college degree.
- Put another way: average partial effects are constant $\frac{\partial \mu(x)}{\partial x} = \beta_1$

Linear CEF with nonlinear effects

- What if we think the effect is nonlinear?

Linear CEF with nonlinear effects

- What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

Linear CEF with nonlinear effects

- What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

  - Partial effect now varies: $\partial\mu(x)/\partial x = \beta_1 + 2x\beta_2$

Linear CEF with nonlinear effects

- What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

    - Partial effect now varies: $\partial\mu(x)/\partial x = \beta_1 + 2x\beta_2$
- **Linear** means linear in the parameters $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)$, not $\mathbf{X}$.

Linear CEF with nonlinear effects

- What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

  - Partial effect now varies: $\partial\mu(x)/\partial x = \beta_1 + 2x\beta_2$
- **Linear** means linear in the parameters $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)$, not $\mathbf{X}$.
- We can also include **interactions** between covariates:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

Linear CEF with nonlinear effects

- What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

  - Partial effect now varies: $\partial\mu(x)/\partial x = \beta_1 + 2x\beta_2$
- **Linear** means linear in the parameters $\boldsymbol{\beta} = (\beta_1, \ldots, \beta_k)$, not **X**.
- We can also include **interactions** between covariates:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

  - Average partial effect of X_1 depends on X_2: $\partial\mu(x_1, x_2)/\partial x_1 = \beta_1 + x_2\beta_3$

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)
  - Two possible values of the CEF: μ_1 for whites and μ_0 for POC.

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)
  - Two possible values of the CEF: μ_1 for whites and μ_0 for POC.
- Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1 - x)\mu_0 = \mu_0 + x\left(\mu_1 - \mu_0\right) = \beta_0 + x\beta_1$$

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)
    - Two possible values of the CEF: μ_1 for whites and μ_0 for POC.
- Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1 - x)\mu_0 = \mu_0 + x\left(\mu_1 - \mu_0\right) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)
    - Two possible values of the CEF: μ_1 for whites and μ_0 for POC.
- Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1 - x)\mu_0 = \mu_0 + x\left(\mu_1 - \mu_0\right) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:
    - $\beta_0 = \mu_0$: expected wait-time for POC

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)
  - Two possible values of the CEF: μ_1 for whites and μ_0 for POC.
- Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1 - x)\mu_0 = \mu_0 + x\left(\mu_1 - \mu_0\right) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:
  - $\beta_0 = \mu_0$: expected wait-time for POC
  - $\beta_1 = \mu_1 - \mu_0$: diff. in avg. wait times between whites and POC.

Linear CEF with a binary covariate

- Wait-times (Y_i) and race ($X_i = 1$ for white, $X_i = 0$ for POC)
  - Two possible values of the CEF: μ_1 for whites and μ_0 for POC.
- Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1 - x)\mu_0 = \mu_0 + x\left(\mu_1 - \mu_0\right) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:
  - $\beta_0 = \mu_0$: expected wait-time for POC
  - $\beta_1 = \mu_1 - \mu_0$: diff. in avg. wait times between whites and POC.
- >2 categories: dummies for all but category and everything is linear.

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:
  - $\beta_0 = \mu_{00}$: average wait times for rural POC.

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:
    - $\beta_0 = \mu_{00}$: average wait times for rural POC.
    - $\beta_1 = \mu_{10} - \mu_{00}$: diff. in means for rural whites vs rural POC.

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:
    - $\beta_0 = \mu_{00}$: average wait times for rural POC.
    - $\beta_1 = \mu_{10} - \mu_{00}$: diff. in means for rural whites vs rural POC.
    - $\beta_2 = \mu_{01} - \mu_{00}$: diff. in means for urban POC vs rural POC.

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:
    - $\beta_0 = \mu_{00}$: average wait times for rural POC.
    - $\beta_1 = \mu_{10} - \mu_{00}$: diff. in means for rural whites vs rural POC.
    - $\beta_2 = \mu_{01} - \mu_{00}$: diff. in means for urban POC vs rural POC.
    - $\beta_3 = (\mu_{11} - \mu_{01}) - (\mu_{10} - \mu_{00})$: diff. in urban racial diff. vs rural racial diff.

Linear CEF with multiple binary covariates

- What if we have two binary covariates, X_1 (race) and X_2 (1 urban/0 rural):

$$\mu(x_1, x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

- Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:
    - $\beta_0 = \mu_{00}$: average wait times for rural POC.
    - $\beta_1 = \mu_{10} - \mu_{00}$: diff. in means for rural whites vs rural POC.
    - $\beta_2 = \mu_{01} - \mu_{00}$: diff. in means for urban POC vs rural POC.
    - $\beta_3 = (\mu_{11} - \mu_{01}) - (\mu_{10} - \mu_{00})$: diff. in urban racial diff. vs rural racial diff.
- Generalizes to p binary variables if **all interactions included (saturated)**

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $m(x) = \beta_0 + \beta_1 X$ is called the **linear projection** of Y onto X.

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $m(x) = \beta_0 + \beta_1 X$ is called the **linear projection** of Y onto X.
    - $\beta_1 = \text{Cov}(X, Y)/\mathbb{V}[X]$

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $m(x) = \beta_0 + \beta_1 X$ is called the **linear projection** of Y onto X.
    - $\beta_1 = \text{Cov}(X, Y)/\mathbb{V}[X]$
    - $\beta_0 = \mu_Y - \mu_X \beta_1$, where $\mu_Y = \mathbb{E}[Y]$ and $\mu_X = \mathbb{E}[X]$

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $m(x) = \beta_0 + \beta_1 X$ is called the **linear projection** of Y onto X.
  - $\beta_1 = \text{Cov}(X, Y)/\mathbb{V}[X]$
  - $\beta_0 = \mu_Y - \mu_X \beta_1$, where $\mu_Y = \mathbb{E}[Y]$ and $\mu_X = \mathbb{E}[X]$
- In general, $m(x)$ distinct from the CEF:

Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $m(x) = \beta_0 + \beta_1 X$ is called the **linear projection** of Y onto X.
  - $\beta_1 = \text{Cov}(X, Y)/\mathbb{V}[X]$
  - $\beta_0 = \mu_Y - \mu_X\beta_1$, where $\mu_Y = \mathbb{E}[Y]$ and $\mu_X = \mathbb{E}[X]$
- In general, $m(x)$ distinct from the CEF:
  - CEF, $\mu(x)$ is the best predictor of Y_i among all functions.

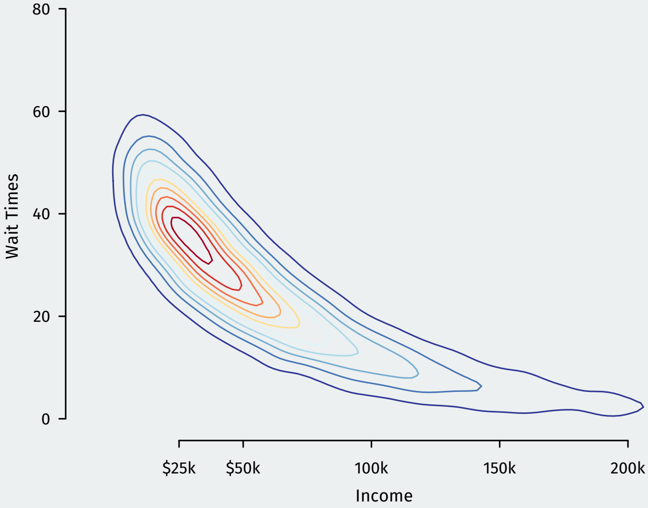
Linear approximation

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

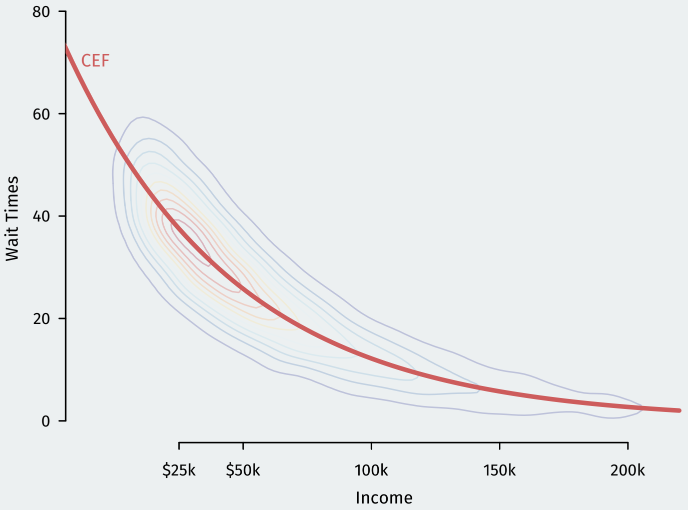
$$(\beta_0, \beta_1) = \underset{(b_0, b_1)}{\arg\min}\, \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $m(x) = \beta_0 + \beta_1 X$ is called the **linear projection** of Y onto X.
  - $\beta_1 = \text{Cov}(X, Y)/\mathbb{V}[X]$
  - $\beta_0 = \mu_Y - \mu_X \beta_1$, where $\mu_Y = \mathbb{E}[Y]$ and $\mu_X = \mathbb{E}[X]$
- In general, $m(x)$ distinct from the CEF:
  - CEF, $\mu(x)$ is the best predictor of Y_i among all functions.
  - Linear projection is best predictor among linear functions.

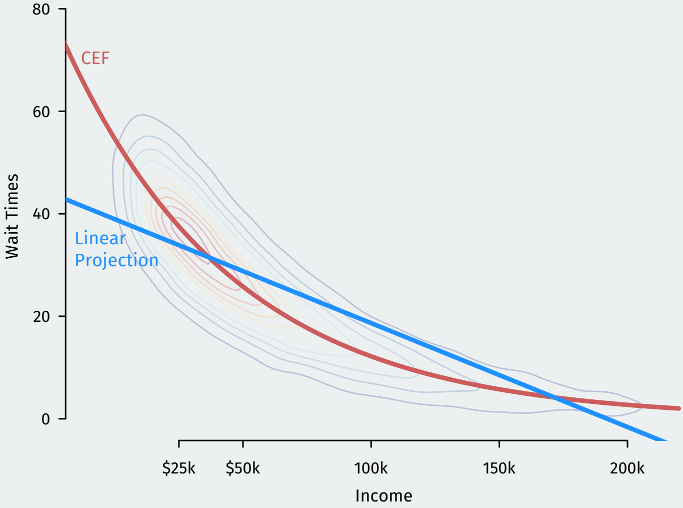
Linear approximation



Linear approximation



Linear approximation



Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \ldots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \ldots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$
- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$
- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:
  - May contain nonlinear transformations/interactions of "real" variables.

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

    - Linear predictor when $\mathbf{X} = \mathbf{x}$
- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:
    - May contain nonlinear transformations/interactions of "real" variables.
    - Typically, $X_1 = 1$ and is the intercept/constant.

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$
- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically, $X_1 = 1$ and is the intercept/constant.
- Assumptions ("Regularity conditions"):

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$
- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically, $X_1 = 1$ and is the intercept/constant.
- Assumptions ("Regularity conditions"):
  1. $\mathbb{E}[Y^2] < \infty$ (outcome has finite mean/variance)

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$

- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically, $X_1 = 1$ and is the intercept/constant.

- Assumptions ("Regularity conditions"):

  1. $\mathbb{E}[Y^2] < \infty$ (outcome has finite mean/variance)
  2. $\mathbb{E}\|\mathbf{X}\|^2 < \infty$ ($\mathbf{X}$ has finite means/variances/covariances)

Best linear predictor

- We'll almost always condition on a vector $\mathbf{X} = (X_1, \dots, X_k)'$:

$$m(\mathbf{x}) = m(x_1, \dots, x_k) = x_1\beta_1 + \cdots + x_k\beta_k = \mathbf{x}'\boldsymbol{\beta}$$

  - Linear predictor when $\mathbf{X} = \mathbf{x}$
- $\mathbf{X}$ is now a $k \times 1$ random vector of covariates:
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically, $X_1 = 1$ and is the intercept/constant.
- Assumptions ("Regularity conditions"):
  1. $\mathbb{E}[Y^2] < \infty$ (outcome has finite mean/variance)
  2. $\mathbb{E}\|\mathbf{X}\|^2 < \infty$ ($\mathbf{X}$ has finite means/variances/covariances)
  3. $\mathbf{Q}_{\mathbf{XX}} = \mathbb{E}[\mathbf{XX}']$ is positive definite (columns of $\mathbf{X}$ are linearly independent)

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \arg\min_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[(Y - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \underset{\mathbf{b} \in \mathbb{R}^k}{\arg\min}\, \mathbb{E}\left[\left(Y - \mathbf{X}'\boldsymbol{\beta}\right)^2\right]$$

- After some calculus:

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{XX}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \arg\min_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[(Y - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- After some calculus:

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{XX}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

  - $\mathbb{E}[\mathbf{XX}']$ is $k \times k$ and $\mathbb{E}[\mathbf{X}Y]$ is $k \times 1$

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \arg\min_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[(Y - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- After some calculus:

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{XX}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

  - $\mathbb{E}[\mathbf{XX}']$ is $k \times k$ and $\mathbb{E}[\mathbf{X}Y]$ is $k \times 1$

- Notes about the $m(\mathbf{x}) = \mathbf{x}'\boldsymbol{\beta}$:

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \underset{\mathbf{b}\in\mathbb{R}^k}{\arg\min}\, \mathbb{E}\left[(Y - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- After some calculus:

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{XX}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

  - $\mathbb{E}[\mathbf{XX}']$ is $k \times k$ and $\mathbb{E}[\mathbf{X}Y]$ is $k \times 1$
- Notes about the $m(\mathbf{x}) = \mathbf{x}'\boldsymbol{\beta}$:
  - $\boldsymbol{\beta}$ is a population quantity and possible quantity of interest.

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \arg\min_{\mathbf{b}\in\mathbb{R}^k} \mathbb{E}\left[(Y - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- After some calculus:

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{XX}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

  - $\mathbb{E}[\mathbf{XX}']$ is $k \times k$ and $\mathbb{E}[\mathbf{X}Y]$ is $k \times 1$

- Notes about the $m(\mathbf{x}) = \mathbf{x}'\boldsymbol{\beta}$:
  - $\boldsymbol{\beta}$ is a population quantity and possible quantity of interest.
  - Well-defined under very mild assumptions!

Linear Projection

- How to find $\boldsymbol{\beta}$? Minimize squared prediction error!

$$\boldsymbol{\beta} = \underset{\mathbf{b} \in \mathbb{R}^k}{\arg\min}\, \mathbb{E}\left[(Y - \mathbf{X}'\boldsymbol{\beta})^2 \right]$$

- After some calculus:

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{XX}}^{-1} \mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{XX}'] \right)^{-1} \mathbb{E}[\mathbf{X}Y]$$

  - $\mathbb{E}[\mathbf{XX}']$ is $k \times k$ and $\mathbb{E}[\mathbf{X}Y]$ is $k \times 1$
- Notes about the $m(\mathbf{x}) = \mathbf{x}'\boldsymbol{\beta}$:
  - $\boldsymbol{\beta}$ is a population quantity and possible quantity of interest.
  - Well-defined under very mild assumptions!
  - Not necessarily a conditional mean nor a causal effect!

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:
    - $\mathbb{E}[\mathbf{X}e] = 0$

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$ when $\mathbf{X}$ contains a constant.

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:
    - $\mathbb{E}[\mathbf{X}e] = 0$
    - $\mathbb{E}[e] = 0$ when $\mathbf{X}$ contains a constant.
    - Together, implies $\text{Cov}(X_j, e) = 0$ for all $j = 1, \dots, k$

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$ when $\mathbf{X}$ contains a constant.
  - Together, implies $\text{Cov}(X_j, e) = 0$ for all $j = 1, \dots, k$
- Distinct from CEF errors: $u = Y - \mu(\mathbf{X})$ which had the additional property: $\mathbb{E}[u \mid \mathbf{X}] = 0$

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:
    - $\mathbb{E}[\mathbf{X}e] = 0$
    - $\mathbb{E}[e] = 0$ when $\mathbf{X}$ contains a constant.
    - Together, implies $\text{Cov}(X_j, e) = 0$ for all $j = 1, \dots, k$
- Distinct from CEF errors: $u = Y - \mu(\mathbf{X})$ which had the additional property: $\mathbb{E}[u \mid \mathbf{X}] = 0$
    - Zero conditional mean is stronger: CEF errors are 0 at every value of $\mathbf{X}$

Projection errors

- Projection error: $e = Y - \mathbf{X}'\boldsymbol{\beta}$
- Decomposition of Y into the linear projection and error: $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$ when $\mathbf{X}$ contains a constant.
  - Together, implies $\text{Cov}(X_j, e) = 0$ for all $j = 1, \ldots, k$
- Distinct from CEF errors: $u = Y - \mu(\mathbf{X})$ which had the additional property: $\mathbb{E}[u \mid \mathbf{X}] = 0$
  - Zero conditional mean is stronger: CEF errors are 0 at every value of $\mathbf{X}$
  - $\mathbb{E}[\mathbf{X}e] = 0$ just says they are uncorrelated.

Regression coefficients

- Sometimes useful to separate the constant:

$$Y = \beta_0 + \mathbf{X}'\boldsymbol{\beta} + e$$

where $\mathbf{X}$ doesn't have a constant.

Regression coefficients

- Sometimes useful to separate the constant:

$$Y = \beta_0 + \mathbf{X}'\boldsymbol{\beta} + e$$

  where $\mathbf{X}$ doesn't have a constant.

- Solution for $\boldsymbol{\beta}$ more interpretable here:

$$\boldsymbol{\beta} = \mathbb{V}[\mathbf{X}]^{-1}\text{Cov}(\mathbf{X}, Y), \qquad \beta_0 = \mu_Y - \boldsymbol{\mu}_{\mathbf{X}}'\boldsymbol{\beta}$$

Interpretation of the coefficients

- Interpretation of β_j depends on what nonlinearities are included.

Interpretation of the coefficients

- Interpretation of β_j depends on what nonlinearities are included.
- Simplest case: no polynomials or interactions.

Interpretation of the coefficients

- Interpretation of β_j depends on what nonlinearities are included.
- Simplest case: no polynomials or interactions.
- β_j is the average change in predicted outcome for a one-unit change in X_j holding other variables fixed.

Interpretation of the coefficients

- Interpretation of β_j depends on what nonlinearities are included.
- Simplest case: no polynomials or interactions.
- β_j is the average change in predicted outcome for a one-unit change in X_j holding other variables fixed.
- Let's compare:

$$m(x_1 + 1, x_2) = \beta_0 + \beta_1(x_1 + 1) + \beta_2 x_2$$
$$m(x_1, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2,$$

Interpretation of the coefficients

- Interpretation of β_j depends on what nonlinearities are included.
- Simplest case: no polynomials or interactions.
- β_j is the average change in predicted outcome for a one-unit change in X_j holding other variables fixed.
- Let's compare:

$$
\begin{aligned}
m(x_1 + 1, x_2) &= \beta_0 + \beta_1(x_1 + 1) + \beta_2 x_2 \\
m(x_1, x_2) &= \beta_0 + \beta_1 x_1 + \beta_2 x_2,
\end{aligned}
$$

- Then:

$$
m(x_1 + 1, x_2) - m(x_1, x_2) = \beta_1
$$

Interpretation of the coefficients

- Interpretation of β_j depends on what nonlinearities are included.
- Simplest case: no polynomials or interactions.
- β_j is the average change in predicted outcome for a one-unit change in X_j holding other variables fixed.
- Let's compare:

$$\begin{aligned} m(x_1 + 1, x_2) &= \beta_0 + \beta_1(x_1 + 1) + \beta_2 x_2 \\ m(x_1, x_2) &= \beta_0 + \beta_1 x_1 + \beta_2 x_2, \end{aligned}$$

- Then:

$$m(x_1 + 1, x_2) - m(x_1, x_2) = \beta_1$$

- Holds for all values of x_2 and even if we add more variables.

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- One-unit change in x_1 is more complicated:

$$m(x_1 + 1, (x_1 + 1)^2, x_2) = \beta_0 + \beta_1 (x_1 + 1) + \beta_2 (x_1 + 1)^2 + \beta_3 x_2$$
$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- One-unit change in x_1 is more complicated:

$$m(x_1+1, (x_1+1)^2, x_2) = \beta_0 + \beta_1 (x_1+1) + \beta_2 (x_1+1)^2 + \beta_3 x_2$$
$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- Better to think of the **marginal effect** of X_{i1}:

$$\frac{\partial m(x_1, x_1^2, x_2)}{\partial x_1} = \beta_1 + 2\beta_2 x_1$$

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- One-unit change in x_1 is more complicated:

$$m(x_1 + 1, (x_1 + 1)^2, x_2) = \beta_0 + \beta_1 (x_1 + 1) + \beta_2 (x_1 + 1)^2 + \beta_3 x_2$$
$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- Better to think of the **marginal effect** of X_{i1}:

$$\frac{\partial m(x_1, x_1^2, x_2)}{\partial x_1} = \beta_1 + 2\beta_2 x_1$$

- Interpretations:

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- One-unit change in x_1 is more complicated:

$$m(x_1 + 1, (x_1 + 1)^2, x_2) = \beta_0 + \beta_1(x_1 + 1) + \beta_2(x_1 + 1)^2 + \beta_3 x_2$$
$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- Better to think of the **marginal effect** of X_{i1}:

$$\frac{\partial m(x_1, x_1^2, x_2)}{\partial x_1} = \beta_1 + 2\beta_2 x_1$$

- Interpretations:
    - β_1: “effect” of X_{i1} on predicted Y_i when $X_{i1} = 0$ (holding X_{i2} fixed)

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- One-unit change in x_1 is more complicated:

$$m(x_1 + 1, (x_1 + 1)^2, x_2) = \beta_0 + \beta_1(x_1 + 1) + \beta_2(x_1 + 1)^2 + \beta_3 x_2$$
$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- Better to think of the **marginal effect** of X_{i1}:

$$\frac{\partial m(x_1, x_1^2, x_2)}{\partial x_1} = \beta_1 + 2\beta_2 x_1$$

- Interpretations:
  - β_1: "effect" of X_{i1} on predicted Y_i when $X_{i1} = 0$ (holding X_{i2} fixed)
  - $\beta_2/2$: how that "effect" changes as X_{i1} changes

Interpretation with nonlinear terms

- What if we include a nonlinear function of one covariate?

$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- One-unit change in x_1 is more complicated:

$$m(x_1 + 1, (x_1 + 1)^2, x_2) = \beta_0 + \beta_1 (x_1 + 1) + \beta_2 (x_1 + 1)^2 + \beta_3 x_2$$
$$m(x_1, x_1^2, x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \beta_3 x_2,$$

- Better to think of the **marginal effect** of X_{i1}:

$$\frac{\partial m(x_1, x_1^2, x_2)}{\partial x_1} = \beta_1 + 2\beta_2 x_1$$

- Interpretations:
    - β_1: "effect" of X_{i1} on predicted Y_i when $X_{i1} = 0$ (holding X_{i2} fixed)
    - $\beta_2/2$: how that "effect" changes as X_{i1} changes
    - Maybe better to visualize than to interpret

Interpretation with interactions

- What if we include an interaction between two covariates?

$$m(x_1, x_2, x_1x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

Interpretation with interactions

- What if we include an interaction between two covariates?

$$m(x_1, x_2, x_1x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

- Two different **marginal effects** of interest:

$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_1} = \beta_1 + \beta_3 x_2,$$
$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_2} = \beta_2 + \beta_3 x_1$$

Interpretation with interactions

- What if we include an interaction between two covariates?

$$m(x_1, x_2, x_1x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

- Two different **marginal effects** of interest:

$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_1} = \beta_1 + \beta_3 x_2,$$
$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_2} = \beta_2 + \beta_3 x_1$$

- Interpretations:

Interpretation with interactions

- What if we include an interaction between two covariates?

$$m(x_1, x_2, x_1x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

- Two different **marginal effects** of interest:

$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_1} = \beta_1 + \beta_3 x_2,$$
$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_2} = \beta_2 + \beta_3 x_1$$

- Interpretations:
    - β_1: the marginal effect of X_{i1} on predicted Y_i when $X_{i2} = 0$.

Interpretation with interactions

- What if we include an interaction between two covariates?

$$m(x_1, x_2, x_1x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

- Two different **marginal effects** of interest:

$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_1} = \beta_1 + \beta_3 x_2,$$
$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_2} = \beta_2 + \beta_3 x_1$$

- Interpretations:
    - β_1: the marginal effect of X_{i1} on predicted Y_i when $X_{i2} = 0$.
    - β_2: the marginal effect of X_{i2} on predicted Y_i when $X_{i1} = 0$.

Interpretation with interactions

- What if we include an interaction between two covariates?

$$m(x_1, x_2, x_1x_2) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

- Two different **marginal effects** of interest:

$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_1} = \beta_1 + \beta_3 x_2,$$
$$\frac{\partial m(x_1, x_2, x_1x_2)}{\partial x_2} = \beta_2 + \beta_3 x_1$$

- Interpretations:
    - β_1: the marginal effect of X_{i1} on predicted Y_i when $X_{i2} = 0$.
    - β_2: the marginal effect of X_{i2} on predicted Y_i when $X_{i1} = 0$.
    - β_3: the change in the marginal effect of X_{i1} due to a one-unit change in X_{i2} **OR** the change in the marginal effect of X_{i2} due to a one-unit change in X_{i2}.

Partitioned Regression

$$(\alpha, \beta, \gamma) = \underset{(a,b,c)\in\mathbb{R}^3}{\arg\min}\ \mathbb{E}[(Y_i - (a + bX_i + cZ_i))^2]$$

- Can we get an expression for just β? With some tricks, yes!

Partitioned Regression

$$(\alpha, \beta, \gamma) = \underset{(a,b,c)\in\mathbb{R}^3}{\arg\min} \; \mathbb{E}[(Y_i - (a + bX_i + cZ_i))^2]$$

- Can we get an expression for just β? With some tricks, yes!
- Population residuals from projection of X_i on Z_i:

$$\widetilde{X}_i = X_i - (\delta_0 + \delta_1 Z_i) \quad \text{where} \quad (\delta_0, \delta_1) = \underset{(d_0,d_1)\in\mathbb{R}^2}{\arg\min} \; \mathbb{E}[(X_i - (d_0 + d_1 Z_i))^2]$$

Partitioned Regression

$$(\alpha, \beta, \gamma) = \underset{(a,b,c)\in\mathbb{R}^3}{\arg\min}\ \mathbb{E}[(Y_i - (a + bX_i + cZ_i))^2]$$

- Can we get an expression for just β? With some tricks, yes!
- Population residuals from projection of X_i on Z_i:

$$\widetilde{X}_i = X_i - (\delta_0 + \delta_1 Z_i) \quad \text{where} \quad (\delta_0, \delta_1) = \underset{(d_0,d_1)\in\mathbb{R}^2}{\arg\min}\ \mathbb{E}[(X_i - (d_0 + d_1 Z_i))^2]$$

  - $\widetilde{X}_i$ is now **orthogonal** to Z_i.

Partitioned Regression

$$(\alpha, \beta, \gamma) = \underset{(a,b,c)\in\mathbb{R}^3}{\arg\min}\; \mathbb{E}[(Y_i - (a + bX_i + cZ_i))^2]$$

- Can we get an expression for just β? With some tricks, yes!
- Population residuals from projection of X_i on Z_i:

$$\widetilde{X}_i = X_i - (\delta_0 + \delta_1 Z_i) \quad \text{where} \quad (\delta_0, \delta_1) = \underset{(d_0,d_1)\in\mathbb{R}^2}{\arg\min}\; \mathbb{E}[(X_i - (d_0 + d_1 Z_i))^2]$$

  - $\widetilde{X}_i$ is now **orthogonal** to Z_i.
- Project Y onto these residuals gives β as coefficient:

$$\beta = \frac{\text{cov}(Y_i, \widetilde{X}_i)}{\mathbb{V}[\widetilde{X}_i]}$$

Partitioned Regression

$$(\alpha, \beta, \gamma) = \underset{(a,b,c)\in\mathbb{R}^3}{\arg\min}\ \mathbb{E}[(Y_i - (a + bX_i + cZ_i))^2]$$

- Can we get an expression for just β? With some tricks, yes!
- Population residuals from projection of X_i on Z_i:

$$\widetilde{X}_i = X_i - (\delta_0 + \delta_1 Z_i) \quad \text{where} \quad (\delta_0, \delta_1) = \underset{(d_0,d_1)\in\mathbb{R}^2}{\arg\min}\ \mathbb{E}[(X_i - (d_0 + d_1 Z_i))^2]$$

  - $\widetilde{X}_i$ is now **orthogonal** to Z_i.
- Project Y onto these residuals gives β as coefficient:

$$\beta = \frac{\text{cov}(Y_i, \widetilde{X}_i)}{\mathbb{V}[\widetilde{X}_i]}$$

- Works if $\mathbf{Z}_i$ is a vector and $\widetilde{X}_i = X_i - m_X(\mathbf{Z}_i)$.

Partitioned Regression

$$(\alpha, \beta, \gamma) = \underset{(a,b,c)\in\mathbb{R}^3}{\arg\min}\ \mathbb{E}[(Y_i - (a + bX_i + cZ_i))^2]$$

- Can we get an expression for just β? With some tricks, yes!
- Population residuals from projection of X_i on Z_i:

$$\widetilde{X}_i = X_i - (\delta_0 + \delta_1 Z_i) \quad \text{where} \quad (\delta_0, \delta_1) = \underset{(d_0,d_1)\in\mathbb{R}^2}{\arg\min}\ \mathbb{E}[(X_i - (d_0 + d_1 Z_i))^2]$$

  - $\widetilde{X}_i$ is now **orthogonal** to Z_i.
- Project Y onto these residuals gives β as coefficient:

$$\beta = \frac{\text{cov}(Y_i, \widetilde{X}_i)}{\mathbb{V}[\widetilde{X}_i]}$$

- Works if $\mathbf{Z}_i$ is a vector and $\widetilde{X}_i = X_i - m_X(\mathbf{Z}_i)$.
  - $m_X(\mathbf{Z}_i)$ is the BLP of X_i on $\mathbf{Z}_i$

Omitted variable bias

- Consider two projections/regressions with and without some Z:

$$m(\mathbf{X}_i, Z_i) = \mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma, \qquad m_{-z}(\mathbf{X}_i) = \mathbf{X}_i'\boldsymbol{\delta},$$

Omitted variable bias

- Consider two projections/regressions with and without some Z:

$$m(\mathbf{X}_i, Z_i) = \mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma, \qquad m_{-z}(\mathbf{X}_i) = \mathbf{X}_i'\boldsymbol{\delta},$$

- How do $\boldsymbol{\beta}$ and $\boldsymbol{\delta}$ relate? Use law of iterated projections:

$$\begin{aligned}
\boldsymbol{\delta} &= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iY_i] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_i(\mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma + e_i)] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\boldsymbol{\beta} + \mathbb{E}[\mathbf{X}_iZ_i]\gamma + \mathbb{E}[\mathbf{X}_ie_i]\right) \\
&= \boldsymbol{\beta} + \underbrace{\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]}_{\text{coefs from} Z\sim\mathbf{X}}\gamma
\end{aligned}$$

Omitted variable bias

- Consider two projections/regressions with and without some Z:

$$m(\mathbf{X}_i, Z_i) = \mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma, \qquad m_{-z}(\mathbf{X}_i) = \mathbf{X}_i'\boldsymbol{\delta},$$

- How do $\boldsymbol{\beta}$ and $\boldsymbol{\delta}$ relate? Use law of iterated projections:

$$\begin{aligned}
\boldsymbol{\delta} &= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iY_i] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_i(\mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma + e_i)] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\boldsymbol{\beta} + \mathbb{E}[\mathbf{X}_iZ_i]\gamma + \mathbb{E}[\mathbf{X}_ie_i]\right) \\
&= \boldsymbol{\beta} + \underbrace{\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]}_{\text{coefs from} Z\sim\mathbf{X}}\gamma
\end{aligned}$$

- Leads to the **"omitted variable bias" formula**:

$$\boldsymbol{\delta} = \boldsymbol{\beta} + \boldsymbol{\pi}\gamma, \qquad \boldsymbol{\pi} = \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]$$

Omitted variable bias

- Consider two projections/regressions with and without some Z:

$$m(\mathbf{X}_i, Z_i) = \mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma, \qquad m_{-z}(\mathbf{X}_i) = \mathbf{X}_i'\boldsymbol{\delta},$$

- How do $\boldsymbol{\beta}$ and $\boldsymbol{\delta}$ relate? Use law of iterated projections:

$$\begin{aligned}
\boldsymbol{\delta} &= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iY_i] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_i(\mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma + e_i)] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\boldsymbol{\beta} + \mathbb{E}[\mathbf{X}_iZ_i]\gamma + \mathbb{E}[\mathbf{X}_ie_i]\right) \\
&= \boldsymbol{\beta} + \underbrace{\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]}_{\text{coefs from} Z\sim\mathbf{X}}\gamma
\end{aligned}$$

- Leads to the **"omitted variable bias" formula**:

$$\boldsymbol{\delta} = \boldsymbol{\beta} + \boldsymbol{\pi}\gamma, \qquad \boldsymbol{\pi} = \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]$$

- $\boldsymbol{\delta} - \boldsymbol{\beta} = \boldsymbol{\pi}\gamma$ is the "bias" but this is misleading.

Omitted variable bias

- Consider two projections/regressions with and without some Z:

$$m(\mathbf{X}_i, Z_i) = \mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma, \qquad m_{-z}(\mathbf{X}_i) = \mathbf{X}_i'\boldsymbol{\delta},$$

- How do $\boldsymbol{\beta}$ and $\boldsymbol{\delta}$ relate? Use law of iterated projections:

$$\begin{aligned}
\boldsymbol{\delta} &= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iY_i] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_i(\mathbf{X}_i'\boldsymbol{\beta} + Z_i\gamma + e_i)] \\
&= \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\boldsymbol{\beta} + \mathbb{E}[\mathbf{X}_iZ_i]\gamma + \mathbb{E}[\mathbf{X}_ie_i]\right) \\
&= \boldsymbol{\beta} + \underbrace{\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]}_{\text{coefs from} Z\sim\mathbf{X}}\gamma
\end{aligned}$$

- Leads to the **"omitted variable bias" formula**:

$$\boldsymbol{\delta} = \boldsymbol{\beta} + \boldsymbol{\pi}\gamma, \qquad \boldsymbol{\pi} = \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[\mathbf{X}_iZ_i]$$

- $\boldsymbol{\delta} - \boldsymbol{\beta} = \boldsymbol{\pi}\gamma$ is the "bias" but this is misleading.
  - $\boldsymbol{\beta}$ not necessarily "correct", we're just relating two projections

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
    - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
  - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.
  - But $\mu(\mathbf{X})$ could be nonlinear, what then?

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
  - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.
  - But $\mu(\mathbf{X})$ could be nonlinear, what then?
- Linear projection justification: best linear approximation to $\mu(\mathbf{X})$:

$$\boldsymbol{\beta} = \underset{\mathbf{b}\in\mathbb{R}^K}{\arg\min}\, \mathbb{E}\left[(\mu(\mathbf{X}) - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
    - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.
    - But $\mu(\mathbf{X})$ could be nonlinear, what then?
- Linear projection justification: best linear approximation to $\mu(\mathbf{X})$:

$$\boldsymbol{\beta} = \underset{\mathbf{b}\in\mathbb{R}^K}{\arg\min}\, \mathbb{E}\left[\left(\mu(\mathbf{X}) - \mathbf{X}'\boldsymbol{\beta}\right)^2\right]$$

- Linear projection is best linear approximation to Y and $\mathbb{E}[Y \mid X]$.

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
  - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.
  - But $\mu(\mathbf{X})$ could be nonlinear, what then?
- Linear projection justification: best linear approximation to $\mu(\mathbf{X})$:

$$\boldsymbol{\beta} = \underset{\mathbf{b} \in \mathbb{R}^K}{\arg\min}\, \mathbb{E}\left[(\mu(\mathbf{X}) - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- Linear projection is best linear approximation to Y and $\mathbb{E}[Y \mid X]$.
- Limitations:

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
  - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.
  - But $\mu(\mathbf{X})$ could be nonlinear, what then?
- Linear projection justification: best linear approximation to $\mu(\mathbf{X})$:

$$\boldsymbol{\beta} = \underset{\mathbf{b}\in\mathbb{R}^K}{\arg\min}\, \mathbb{E}\left[(\mu(\mathbf{X}) - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- Linear projection is best linear approximation to Y and $\mathbb{E}[Y \mid X]$.
- Limitations:
  - If nonlinearity of $\mu(\mathbf{X})$ is severe, $m(\mathbf{X})$ can only be so good.

Best linear approximation

- What is the relationship between $m(\mathbf{X})$ and $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$?
    - If $\mu(\mathbf{X})$ is linear, then $\mu(\mathbf{X}) = m(\mathbf{X}) = \mathbf{X}'\boldsymbol{\beta}$.
    - But $\mu(\mathbf{X})$ could be nonlinear, what then?
- Linear projection justification: best linear approximation to $\mu(\mathbf{X})$:

$$\boldsymbol{\beta} = \underset{\mathbf{b} \in \mathbb{R}^K}{\arg\min}\, \mathbb{E}\left[(\mu(\mathbf{X}) - \mathbf{X}'\boldsymbol{\beta})^2\right]$$

- Linear projection is best linear approximation to Y and $\mathbb{E}[Y \mid X]$.
- Limitations:
    - If nonlinearity of $\mu(\mathbf{X})$ is severe, $m(\mathbf{X})$ can only be so good.
    - $m(\mathbf{X})$ can be sensitive to the marginal distribution of $\mathbf{X}$.

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- “The Linear Model”: is this an assumption?

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
    - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
    - If just $\mathbb{E}[\mathbf{X}e] = 0$, then this is just a linear projection.

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
    - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
    - If just $\mathbb{E}[\mathbf{X}e] = 0$, then this is just a linear projection.
    - First is very strong, second is very mild.

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- “The Linear Model”: is this an assumption?
- Depends on what we assume about the error, e
    - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
    - If just $\mathbb{E}[\mathbf{X}e] = 0$, then this is just a linear projection.
    - First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
  - If just $\mathbb{E}[\mathbf{X}e] = 0$, then this is just a linear projection.
  - First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.
  - Some finite-sample properties of OLS (unbiasedness) require linear CEF

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
    - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
    - If just $\mathbb{E}[\mathbf{X}e] = 0$, then this is just a linear projection.
    - First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.
    - Some finite-sample properties of OLS (unbiasedness) require linear CEF
    - Asymptotic results (consistency, asymptotic normality) apply to both.

Recap

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
    - If $\mathbb{E}[e \mid \mathbf{X}] = 0$, then we are assuming the CEF is linear, $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
    - If just $\mathbb{E}[\mathbf{X}e] = 0$, then this is just a linear projection.
    - First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.
    - Some finite-sample properties of OLS (unbiasedness) require linear CEF
    - Asymptotic results (consistency, asymptotic normality) apply to both.
    - OLS will consitently estimate something, but maybe not what you want.