# 13. Linear Model

Spring 2021

Matthew Blackwell

Gov 2002 (Harvard)

# Where are we? Where are we going?

· Learned about the CEF in general, iterated expectation, etc.

# Where are we? Where are we going?

- · Learned about the CEF in general, iterated expectation, etc.
- Now: focusing on when the CEF is (and isn't) linear.

# Where are we? Where are we going?

- Learned about the CEF in general, iterated expectation, etc.
- Now: focusing on when the CEF is (and isn't) linear.
- Linear model is ubiquitous but poorly understood. Lots of subtlety here.

• Goal of regression: how mean of Y changes with X.

- Goal of regression: how mean of Y changes with X.
- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1,\ldots,x_k)}{\partial x_1}$$

- Goal of regression: how mean of Y changes with X.
- · For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1,\ldots,x_k)}{\partial x_1}$$

• For binary  $X_1$ , we can use the difference in conditional expectations:

$$\mu(1,x_2,\dots,x_k) - \mu(0,x_2,\dots,x_k)$$

- Goal of regression: how mean of Y changes with X.
- · For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1,\ldots,x_k)}{\partial x_1}$$

• For binary  $X_1$ , we can use the difference in conditional expectations:

$$\mu(1,x_2,\dots,x_k) - \mu(0,x_2,\dots,x_k)$$

• "Partial effect" of  $X_1$  holding other included variables constant

- Goal of regression: how mean of Y changes with X.
- · For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1,\ldots,x_k)}{\partial x_1}$$

• For binary  $X_1$ , we can use the difference in conditional expectations:

$$\mu(1,x_2,\dots,x_k) - \mu(0,x_2,\dots,x_k)$$

- "Partial effect" of  $X_1$  holding other included variables constant
- Exact form will depend on the functional form of  $\mu(x)$ .

- Goal of regression: how mean of Y changes with X.
- For continuous regressors, we can use the partial derivative:

$$\frac{\partial \mu(x_1,\ldots,x_k)}{\partial x_1}$$

• For binary  $X_1$ , we can use the difference in conditional expectations:

$$\mu(1,x_2,\dots,x_k) - \mu(0,x_2,\dots,x_k)$$

- "Partial effect" of  $X_1$  holding other included variables constant
- Exact form will depend on the functional form of  $\mu(x)$ .
  - How do we decide what form  $\mu(x)$  should take?

• To motivate function form, useful to think about estimation.

- To motivate function form, useful to think about estimation.
- How do we estimate  $\mu(x) = \mathbb{E}[Y|X=x]$  for binary X?

- To motivate function form, useful to think about estimation.
- How do we estimate  $\mu(x) = \mathbb{E}[Y|X=x]$  for binary X?
- **Subclassification**: calculate sample averages with levels of  $X_i$ :

$$\hat{\mu}(1) = \frac{1}{n_1} \sum_{i=1}^{n} Y_i X_i$$

- To motivate function form, useful to think about estimation.
- How do we estimate  $\mu(x) = \mathbb{E}[Y|X=x]$  for binary X?
- **Subclassification**: calculate sample averages with levels of  $X_i$ :

$$\hat{\mu}(1) = \frac{1}{n_1} \sum_{i=1}^{n} Y_i X_i$$

•  $n_1 = \sum_{i=1}^n X_i$  is the number of units with  $X_i = 1$  in the sample.

- To motivate function form, useful to think about estimation.
- How do we estimate  $\mu(x) = \mathbb{E}[Y|X=x]$  for binary X?
- **Subclassification**: calculate sample averages with levels of  $X_i$ :

$$\hat{\mu}(1) = \frac{1}{n_1} \sum_{i=1}^{n} Y_i X_i$$

- $n_1 = \sum_{i=1}^n X_i$  is the number of units with  $X_i = 1$  in the sample.
- More generally for any discrete X<sub>i</sub>:

$$\hat{\mu}(x) = \frac{\sum_{i=1}^{N} Y_i \mathbb{I}(X_i = x)}{\sum_{i=1}^{N} \mathbb{I}(X_i = x)}$$

• What if X is continuous? Subclassification fall apart.

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - · Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin  $X_i$  into categories and treat like as discrete.

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - · Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin  $X_i$  into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.

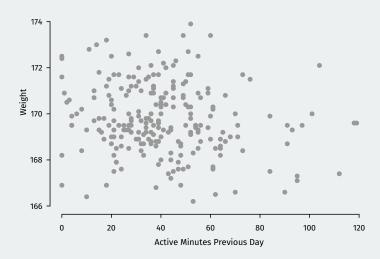
- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin  $X_i$  into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - Depends on arbitrary bin cutoffs/sizes.

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin  $X_i$  into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - Depends on arbitrary bin cutoffs/sizes.
- Example:

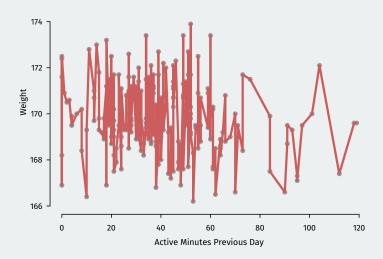
- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin  $X_i$  into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - · Depends on arbitrary bin cutoffs/sizes.
- · Example:
  - Personal data science: I wear an activity tracker and have a smart scale.

- What if X is continuous? Subclassification fall apart.
  - Each *i* has a unique value:  $\sum_{i=1}^{N} \mathbb{I}(X_i = x) = 1$
  - Very noisy estimates
  - What about any x not in the sample?
- **Stratification**: bin  $X_i$  into categories and treat like as discrete.
  - Every x in the same bin gets the same conditional expectation.
  - · Depends on arbitrary bin cutoffs/sizes.
- · Example:
  - Personal data science: I wear an activity tracker and have a smart scale.
  - Relationship between my weight and active minutes in the previous day.

# **Continuous covariate example**



# **Continuous covariate CEF: interpolation**



# **Continuous covariate CEF: stratification**



# **Continuous covariate CEF: stratification**



• Statification requires lots of choices/hidden assumptions.

- · Statification requires lots of choices/hidden assumptions.
  - Number of categories, cutoffs for the categories, constant means within strata, etc.

- · Statification requires lots of choices/hidden assumptions.
  - Number of categories, cutoffs for the categories, constant means within strata, etc.
- Alternative: assuming that the CEF is linear:

$$\mu(x) = \mathbb{E}[Y_i | X_i = x] = \beta_0 + \beta_1 x$$

- · Statification requires lots of choices/hidden assumptions.
  - Number of categories, cutoffs for the categories, constant means within strata, etc.
- Alternative: assuming that the CEF is linear:

$$\mu(x) = \mathbb{E}[Y_i | X_i = x] = \beta_0 + \beta_1 x$$

• **Intercept,**  $\beta_0$ : the condition expectation of  $Y_i$  when  $X_i = 0$ 

- · Statification requires lots of choices/hidden assumptions.
  - Number of categories, cutoffs for the categories, constant means within strata, etc.
- Alternative: assuming that the CEF is **linear**:

$$\mu(x) = \mathbb{E}[Y_i|X_i = x] = \beta_0 + \beta_1 x$$

- **Intercept**,  $\beta_0$ : the condition expectation of  $Y_i$  when  $X_i = 0$
- **Slope**,  $\beta_1$ : change in the CEF of  $Y_i$  given a one-unit change in  $X_i$

# Why is linearity an assumption?

• Example:  $Y_i$  is income,  $X_i$  is years of education.

# Why is linearity an assumption?

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.
- Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i=12]-\mathbb{E}[Y_i|X_i=11]$$

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.
- · Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] = \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15]$$

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.
- · Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] = \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15]$$
$$= \beta_1$$

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.
- · Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] = \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15]$$
$$= \beta_1$$

• Effect of HS degree is the same as the effect of college degree.

- Example:  $Y_i$  is income,  $X_i$  is years of education.
  - $\beta_0$ : average income among people with 0 years of education.
  - $\beta_1$ : expected difference in income between two adults that differ by 1 year of education.
- · Why is linearity an assumption?

$$\mathbb{E}[Y_i|X_i = 12] - \mathbb{E}[Y_i|X_i = 11] = \mathbb{E}[Y_i|X_i = 16] - \mathbb{E}[Y_i|X_i = 15]$$
$$= \beta_1$$

- Effect of HS degree is the same as the effect of college degree.
- Put another way: average partial effects are constant  $rac{\partial \mu(x)}{\partial x} = oldsymbol{eta}_1$

· What if we think the effect is nonlinear?

- · What if we think the effect is nonlinear?
- · We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

- · What if we think the effect is nonlinear?
- · We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

• Partial effect now varies:  $\partial \mu(x)/\partial x = \beta_1 + 2x\beta_2$ 

- · What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

- Partial effect now varies:  $\partial \mu(x)/\partial x = \beta_1 + 2x\beta_2$
- **Linear** means linear in the parameters  $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)$ , not **X**.

- What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

- Partial effect now varies:  $\partial \mu(x)/\partial x = \beta_1 + 2x\beta_2$
- **Linear** means linear in the parameters  $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)$ , not **X**.
- · We can also include **interactions** between covariates:

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

- · What if we think the effect is nonlinear?
- We can include nonlinear transformations:

$$\mu(x) = \beta_0 + x\beta_1 + x^2\beta_2$$

- Partial effect now varies:  $\partial \mu(x)/\partial x = \beta_1 + 2x\beta_2$
- **Linear** means linear in the parameters  $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)$ , not **X**.
- · We can also include **interactions** between covariates:

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

• Average partial effect of  $X_1$  depends on  $X_2$ :  $\partial \mu(x_1,x_2)/\partial x_1=\beta_1+x_2\beta_3$ 

• Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 

- Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 
  - Two possible values of the CEF:  $\mu_1$  for whites and  $\mu_0$  for POC.

- Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 
  - Two possible values of the CEF:  $\mu_1$  for whites and  $\mu_0$  for POC.
- · Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1-x)\mu_0 = \mu_0 + x(\mu_1 - \mu_0) = \beta_0 + x\beta_1$$

- Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 
  - Two possible values of the CEF:  $\mu_1$  for whites and  $\mu_0$  for POC.
- · Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1-x)\mu_0 = \mu_0 + x(\mu_1 - \mu_0) = \beta_0 + x\beta_1$$

• No assumptions, just rewriting! Interpretations:

- Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 
  - Two possible values of the CEF:  $\mu_1$  for whites and  $\mu_0$  for POC.
- · Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1-x)\mu_0 = \mu_0 + x(\mu_1 - \mu_0) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:
  - $oldsymbol{eta}_0=\mu_0$ : expected wait-time for POC

- Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 
  - Two possible values of the CEF:  $\mu_1$  for whites and  $\mu_0$  for POC.
- · Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1-x)\mu_0 = \mu_0 + x(\mu_1 - \mu_0) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:
  - $\beta_0 = \mu_0$ : expected wait-time for POC
  - $oldsymbol{eta}_1=\mu_1-\mu_0$ : diff. in avg. wait times between whites and POC.

- Wait-times  $(Y_i)$  and race  $(X_i = 1 \text{ for white, } X_i = 0 \text{ for POC})$ 
  - Two possible values of the CEF:  $\mu_1$  for whites and  $\mu_0$  for POC.
- · Can write the CEF as follows:

$$\mu(x) = x\mu_1 + (1-x)\mu_0 = \mu_0 + x(\mu_1 - \mu_0) = \beta_0 + x\beta_1$$

- No assumptions, just rewriting! Interpretations:
  - $\beta_0 = \mu_0$ : expected wait-time for POC
  - $\beta_1 = \mu_1 \mu_0$ : diff. in avg. wait times between whites and POC.
- ullet > 2 categories: dummies for all but category and everything is linear.

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

Can rewrite this without assumptions as a linear CEF with interaction:

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

Interpretations:

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

- · Interpretations:
  - $\beta_0 = \mu_{00}$ : average wait times for rural POC.

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

- Interpretations:
  - $\beta_0 = \mu_{00}$ : average wait times for rural POC.
  - $eta_1 = \mu_{10} \mu_{00}$ : diff. in means for rural whites vs rural POC.

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

- Interpretations:
  - $\beta_0 = \mu_{00}$ : average wait times for rural POC.
  - $\beta_1 = \mu_{10} \mu_{00}$ : diff. in means for rural whites vs rural POC.
  - $\beta_2 = \mu_{01} \mu_{00}$ : diff. in means for urban POC vs rural POC.

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

$$\mu(x_1, x_2) = \beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_1 x_2 \beta_3$$

- · Interpretations:
  - $\beta_0 = \mu_{00}$ : average wait times for rural POC.
  - $\beta_1 = \mu_{10} \mu_{00}$ : diff. in means for rural whites vs rural POC.
  - $\beta_2 = \mu_{01} \mu_{00}$ : diff. in means for urban POC vs rural POC.
  - $eta_3=(\mu_{11}-\mu_{01})-(\mu_{10}-\mu_{00})$ : diff. in urban racial diff. vs rural racial diff.

• What if we have two binary covariates,  $X_1$  (race) and  $X_2$  (1 urban/0 rural):

$$\mu(x_1,x_2) = \begin{cases} \mu_{00} & \text{if } x_1 = 0 \text{ and } x_2 = 0 \text{ (POC, rural)} \\ \mu_{10} & \text{if } x_1 = 1 \text{ and } x_2 = 0 \text{ (white, rural)} \\ \mu_{01} & \text{if } x_1 = 0 \text{ and } x_2 = 1 \text{ (POC, urban)} \\ \mu_{11} & \text{if } x_1 = 1 \text{ and } x_2 = 1 \text{ (white, urban)} \end{cases}$$

$$\mu(x_1, x_2) = \beta_0 + x_1\beta_1 + x_2\beta_2 + x_1x_2\beta_3$$

- Interpretations:
  - $\beta_0 = \mu_{00}$ : average wait times for rural POC.
  - $\beta_1 = \mu_{10} \mu_{00}$ : diff. in means for rural whites vs rural POC.
  - $\beta_2 = \mu_{01} \mu_{00}$ : diff. in means for urban POC vs rural POC.
  - $m{\beta}_3=(\mu_{11}-\mu_{01})-(\mu_{10}-\mu_{00})$ : diff. in urban racial diff. vs rural racial diff.
- Generalizes to p binary variables if all interactions included (saturated)

• Outside of saturated discrete settings, CEF almost never truly linear.

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of *Y* given *X*.

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find best linear predictor of Y given X.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0,\beta_1) = \mathop{\arg\min}_{(b_0,b_1)} \mathbb{E}[(Y-(b_0+b_1X))^2]$$

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find best linear predictor of Y given X.
- Formally, linear function of *X* that **minimizes squared prediction errors**:

$$(\beta_0,\beta_1) = \operatorname*{arg\,min}_{(b_0,b_1)} \mathbb{E}[(Y-(b_0+b_1X))^2]$$

•  $\mathbb{L}[Y \mid X] = \beta_0 + \beta_1 X$  is called the **linear projection** of Y onto X.

- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find best linear predictor of Y given X.
- Formally, linear function of *X* that **minimizes squared prediction errors**:

$$(\beta_0,\beta_1) = \operatorname*{arg\,min}_{(b_0,b_1)} \mathbb{E}[(Y-(b_0+b_1X))^2]$$

- $\mathbb{L}[Y \mid X] = \beta_0 + \beta_1 X$  is called the **linear projection** of Y onto X.
  - $\beta_1 = \operatorname{Cov}(X, Y) / \mathbb{V}[X]$

- · Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of *Y* given *X*.
- Formally, linear function of X that **minimizes squared prediction errors**:

$$(\beta_0,\beta_1) = \mathop{\arg\min}_{(b_0,b_1)} \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $\mathbb{L}[Y \mid X] = \beta_0 + \beta_1 X$  is called the **linear projection** of Y onto X.
  - $\beta_1 = \text{Cov}(X, Y)/\mathbb{V}[X]$
  - $eta_0 = \mu_Y \mu_X eta_1$ , where  $\mu_Y = \mathbb{E}[Y]$  and  $\mu_X = \mathbb{E}[X]$

- · Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of Y given X.
- Formally, linear function of *X* that **minimizes squared prediction errors**:

$$(\beta_0,\beta_1) = \mathop{\arg\min}_{(b_0,b_1)} \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

- $\mathbb{L}[Y \mid X] = \beta_0 + \beta_1 X$  is called the **linear projection** of Y onto X.
  - $\beta_1 = \operatorname{Cov}(X, Y) / \mathbb{V}[X]$
  - $eta_0 = \mu_Y \mu_X eta_1$ , where  $\mu_Y = \mathbb{E}[Y]$  and  $\mu_X = \mathbb{E}[X]$
- In general,  $\mathbb{L}[Y \mid X]$  distinct from the CEF:

- · Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find **best linear predictor** of *Y* given *X*.
- Formally, linear function of X that **minimizes squared prediction errors**:

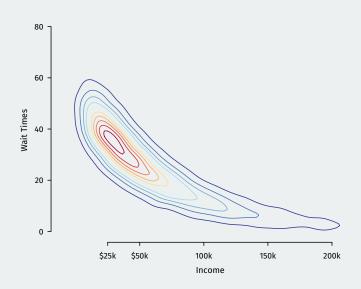
$$(\pmb{\beta}_0, \pmb{\beta}_1) = \mathop{\arg\min}_{(b_0,b_1)} \mathbb{E}[(Y - (b_0 + b_1 X))^2]$$

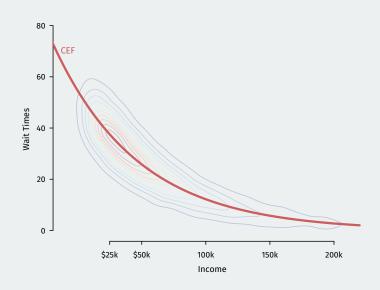
- $\mathbb{L}[Y \mid X] = \beta_0 + \beta_1 X$  is called the **linear projection** of Y onto X.
  - $\beta_1 = \operatorname{Cov}(X, Y) / \mathbb{V}[X]$
  - $\beta_0 = \mu_Y \mu_X \beta_1$ , where  $\mu_Y = \mathbb{E}[Y]$  and  $\mu_X = \mathbb{E}[X]$
- In general,  $\mathbb{L}[Y \mid X]$  distinct from the CEF:
  - CEF,  $\mu(x)$  is the best predictor of  $Y_i$  among all functions.

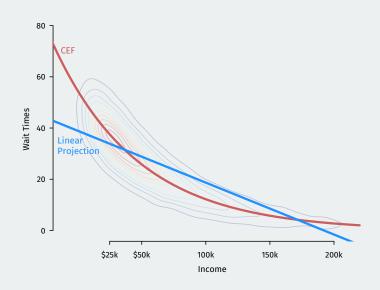
- Outside of saturated discrete settings, CEF almost never truly linear.
- Alternative goal: find best linear predictor of Y given X.
- Formally, linear function of *X* that **minimizes squared prediction errors**:

$$(\beta_0,\beta_1) = \operatorname*{arg\,min}_{(b_0,b_1)} \mathbb{E}[(Y-(b_0+b_1X))^2]$$

- $\mathbb{L}[Y \mid X] = \beta_0 + \beta_1 X$  is called the **linear projection** of Y onto X.
  - $\beta_1 = \operatorname{Cov}(X, Y) / \mathbb{V}[X]$
  - $\beta_0 = \mu_Y \mu_X \beta_1$ , where  $\mu_Y = \mathbb{E}[Y]$  and  $\mu_X = \mathbb{E}[X]$
- In general,  $\mathbb{L}[Y \mid X]$  distinct from the CEF:
  - CEF,  $\mu(x)$  is the best predictor of  $Y_i$  among all functions.
  - Linear projection is best predictor among linear functions.







$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}[Y \mid X_1, \dots, X_k] = X_1 \beta_1 + \dots + X_k \beta_k = \mathbf{X}' \boldsymbol{\beta}$$

• We'll almost always condition on a vector X:

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}[Y \mid X_1, \dots, X_k] = X_1 \beta_1 + \dots + X_k \beta_k = \mathbf{X}' \boldsymbol{\beta}$$

• Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}[Y \mid X_1, \dots, X_k] = X_1 \beta_1 + \dots + X_k \beta_k = \mathbf{X}' \boldsymbol{\beta}$$

- Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 
  - May contain nonlinear transformations/interactions of "real" variables.

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}[Y \mid X_1, \dots, X_k] = X_1 \beta_1 + \dots + X_k \beta_k = \mathbf{X}' \boldsymbol{\beta}$$

- Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically,  $X_1 = 1$  and is the intercept/constant.

$$\mathbb{L}[Y\mid \mathbf{X}] = \mathbb{L}[Y\mid X_1,\dots,X_k] = X_1\boldsymbol{\beta}_1 + \dots + X_k\boldsymbol{\beta}_k = \mathbf{X}'\boldsymbol{\beta}$$

- Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically,  $X_1 = 1$  and is the intercept/constant.
- Assumptions ("Regularity conditions"):

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}[Y \mid X_1, \dots, X_k] = X_1 \beta_1 + \dots + X_k \beta_k = \mathbf{X}' \boldsymbol{\beta}$$

- Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically,  $X_1 = 1$  and is the intercept/constant.
- Assumptions ("Regularity conditions"):
  - 1.  $\mathbb{E}[Y^2] < \infty$  (outcome has finite mean/variance)

$$\mathbb{L}[Y\mid \mathbf{X}] = \mathbb{L}[Y\mid X_1,\dots,X_k] = X_1\boldsymbol{\beta}_1 + \dots + X_k\boldsymbol{\beta}_k = \mathbf{X}'\boldsymbol{\beta}$$

- Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically,  $X_1 = 1$  and is the intercept/constant.
- Assumptions ("Regularity conditions"):
  - 1.  $\mathbb{E}[Y^2] < \infty$  (outcome has finite mean/variance)
  - 2.  $\mathbb{E}\|\mathbf{X}\|^2 < \infty$  (**X** has finite means/variances/covariances)

$$\mathbb{L}[Y\mid \mathbf{X}] = \mathbb{L}[Y\mid X_1,\dots,X_k] = X_1\boldsymbol{\beta}_1 + \dots + X_k\boldsymbol{\beta}_k = \mathbf{X}'\boldsymbol{\beta}$$

- Random vector  $(k \times 1)$  of covariates:  $\mathbf{X} = (X_1, \dots, X_k)'$ 
  - May contain nonlinear transformations/interactions of "real" variables.
  - Typically,  $X_1 = 1$  and is the intercept/constant.
- Assumptions ("Regularity conditions"):
  - 1.  $\mathbb{E}[Y^2] < \infty$  (outcome has finite mean/variance)
  - 2.  $\mathbb{E}\|\mathbf{X}\|^2 < \infty$  (**X** has finite means/variances/covariances)
  - 3.  $\mathbf{Q}_{\mathbf{XX}} = \mathbb{E}[\mathbf{XX}']$  is positive definite (columns of  $\mathbf{X}$  are linearly independent)

• How to find  $\beta$ ? Minimize squared prediction error!

$$\pmb{\beta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[ \left( Y - \mathbf{X}' \pmb{\beta} \right)^2 \right]$$

How to find β? Minimize squared prediction error!

$$\pmb{\beta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[ \left( Y - \mathbf{X}' \pmb{\beta} \right)^2 \right]$$

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{X}\mathbf{X}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

• How to find  $\beta$ ? Minimize squared prediction error!

$${m eta} = \mathop{\mathrm{arg\,min}}_{{m b} \in \mathbb{R}^k} \mathbb{E}\left[ \left( Y - {f X}' {m eta} 
ight)^2 
ight]$$

· After some calculus:

$$\pmb{\beta} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{X}\mathbf{X}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

•  $\mathbb{E}[\mathbf{X}\mathbf{X}']$  is  $k \times k$  and  $\mathbb{E}[\mathbf{X}Y]$  is  $k \times 1$ 

• How to find  $\beta$ ? Minimize squared prediction error!

$$oldsymbol{eta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[\left(Y - \mathbf{X}' oldsymbol{eta}\right)^2
ight]$$

$$\pmb{\beta} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{X}\mathbf{X}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

- $\mathbb{E}[\mathbf{X}\mathbf{X}']$  is  $k \times k$  and  $\mathbb{E}[\mathbf{X}Y]$  is  $k \times 1$
- Notes about the  $\mathbb{L}[Y \mid X]$ :

• How to find  $\beta$ ? Minimize squared prediction error!

$$oldsymbol{eta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[\left(Y - \mathbf{X}' oldsymbol{eta}\right)^2
ight]$$

$$\boldsymbol{\beta} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{X}\mathbf{X}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

- $\mathbb{E}[\mathbf{X}\mathbf{X}']$  is  $k \times k$  and  $\mathbb{E}[\mathbf{X}Y]$  is  $k \times 1$
- Notes about the  $\mathbb{L}[Y \mid X]$ :
  - $\beta$  is a population quantity and possible quantity of interest.

• How to find  $\beta$ ? Minimize squared prediction error!

$$oldsymbol{eta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^k} \mathbb{E}\left[\left(Y - \mathbf{X}' oldsymbol{eta}\right)^2
ight]$$

$$oldsymbol{eta} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{X}\mathbf{X}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

- $\mathbb{E}[\mathbf{X}\mathbf{X}']$  is  $k \times k$  and  $\mathbb{E}[\mathbf{X}Y]$  is  $k \times 1$
- Notes about the  $\mathbb{L}[Y \mid X]$ :
  - $\beta$  is a population quantity and possible quantity of interest.
  - · Well-defined under very mild assumptions!

• How to find  $\beta$ ? Minimize squared prediction error!

$${m eta} = \mathop{\mathrm{arg\,min}}_{{m b} \in \mathbb{R}^k} \mathbb{E}\left[ \left( Y - {f X}' {m eta} 
ight)^2 
ight]$$

$$oldsymbol{eta} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{Q}_{\mathbf{X}Y} = \left(\mathbb{E}[\mathbf{X}\mathbf{X}']\right)^{-1}\mathbb{E}[\mathbf{X}Y]$$

- $\mathbb{E}[\mathbf{X}\mathbf{X}']$  is  $k \times k$  and  $\mathbb{E}[\mathbf{X}Y]$  is  $k \times 1$
- Notes about the  $\mathbb{L}[Y \mid X]$ :
  - $\beta$  is a population quantity and possible quantity of interest.
  - · Well-defined under very mild assumptions!
  - Not necessarily a conditional mean nor a causal effect!

• Projection error:  $e = Y - \mathbf{X}'\boldsymbol{\beta}$ 

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}' \boldsymbol{\beta} + e$

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}' \boldsymbol{\beta} + e$
- Properties of the projection error:

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- · Properties of the projection error:

• 
$$\mathbb{E}[\mathbf{X}e] = 0$$

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- · Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$  when **X** contains a constant.

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- · Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$  when **X** contains a constant.
  - Together, implies  $\operatorname{Cov}(X_j,e)=0$  for all  $j=1,\ldots,k$

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}' \boldsymbol{\beta} + e$
- · Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$  when **X** contains a constant.
  - Together, implies  $Cov(X_j,e)=0$  for all  $j=1,\ldots,k$
- Distinct from CEF errors:  $u=Y-\mu(\mathbf{X})$  which had the additional property:  $\mathbb{E}[u\mid\mathbf{X}]=0$

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}' \boldsymbol{\beta} + e$
- · Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$  when **X** contains a constant.
  - Together, implies  $Cov(X_j,e)=0$  for all  $j=1,\ldots,k$
- Distinct from CEF errors:  $u=Y-\mu(\mathbf{X})$  which had the additional property:  $\mathbb{E}[u\mid\mathbf{X}]=0$ 
  - Zero conditional mean is stronger: CEF errors are 0 at every value of X

- Projection error:  $e = Y X'\beta$
- Decomposition of Y into the linear projection and error:  $Y = \mathbf{X}'\boldsymbol{\beta} + e$
- · Properties of the projection error:
  - $\mathbb{E}[\mathbf{X}e] = 0$
  - $\mathbb{E}[e] = 0$  when **X** contains a constant.
  - Together, implies  $Cov(X_j,e)=0$  for all  $j=1,\ldots,k$
- Distinct from CEF errors:  $u=Y-\mu(\mathbf{X})$  which had the additional property:  $\mathbb{E}[u\mid\mathbf{X}]=0$ 
  - Zero conditional mean is stronger: CEF errors are 0 at every value of X
  - $\mathbb{E}[\mathbf{X}e] = 0$  just says they are uncorrelated.

# **Regression coefficients**

· Sometimes useful to separate the constant:

$$Y = \beta_0 + \mathbf{X}'\boldsymbol{\beta} + e$$

where X doesn't have a constant.

### **Regression coefficients**

· Sometimes useful to separate the constant:

$$Y = \beta_0 + \mathbf{X}'\boldsymbol{\beta} + e$$

where X doesn't have a constant.

• Solution for  $\beta$  more interpretable here:

$$\pmb{\beta} = \mathbb{V}[\mathbf{X}]^{-1} \mathrm{Cov}(\mathbf{X}, Y), \qquad \pmb{\beta}_0 = \mu_Y - \pmb{\mu}_{\mathbf{X}}' \pmb{\beta}$$

$$\mathbb{L}(Y\mid \mathbf{X},\mathbf{Z}) = \mathbf{X}'\boldsymbol{\beta} + \mathbf{Z}'\boldsymbol{\gamma}$$

• Can we get an expression for just  $\beta$ ? With some tricks, yes!

$$\mathbb{L}(Y \mid \mathbf{X}, \mathbf{Z}) = \mathbf{X}'\boldsymbol{\beta} + \mathbf{Z}'\boldsymbol{\gamma}$$

- Can we get an expression for just  $\beta$ ? With some tricks, yes!
- Population residuals from projection of **X** on **Z**:  $\mathbf{R} = \mathbf{X} \mathbb{L}(\mathbf{X} \mid \mathbf{Z})$ .

$$\mathbb{L}(Y \mid \mathbf{X}, \mathbf{Z}) = \mathbf{X}'\boldsymbol{\beta} + \mathbf{Z}'\boldsymbol{\gamma}$$

- Can we get an expression for just  $\beta$ ? With some tricks, yes!
- Population residuals from projection of **X** on **Z**:  $\mathbf{R} = \mathbf{X} \mathbb{L}(\mathbf{X} \mid \mathbf{Z})$ .
  - · R is now orthogonal to Z.

$$\mathbb{L}(Y \mid \mathbf{X}, \mathbf{Z}) = \mathbf{X}' \boldsymbol{\beta} + \mathbf{Z}' \boldsymbol{\gamma}$$

- Can we get an expression for just  $\beta$ ? With some tricks, yes!
- Population residuals from projection of **X** on **Z**:  $\mathbf{R} = \mathbf{X} \mathbb{L}(\mathbf{X} \mid \mathbf{Z})$ .
  - · R is now **orthogonal** to **Z**.
- Project Y onto these residuals gives  $\pmb{\beta}$  as coefficient:  $\mathbb{L}(Y \mid \mathbf{R}) = \mathbf{R}' \pmb{\beta}$

$$\boldsymbol{\beta} = \left(\mathbb{E}[\mathsf{RR}']\right)^{-1}\mathbb{E}[\mathsf{R}Y]$$

$$\mathbb{L}(Y \mid \mathbf{X}, \mathbf{Z}) = \mathbf{X}'\boldsymbol{\beta} + \mathbf{Z}'\boldsymbol{\gamma}$$

- Can we get an expression for just  $\beta$ ? With some tricks, yes!
- Population residuals from projection of **X** on **Z**:  $\mathbf{R} = \mathbf{X} \mathbb{L}(\mathbf{X} \mid \mathbf{Z})$ .
  - · R is now orthogonal to Z.
- Project Y onto these residuals gives  $\pmb{\beta}$  as coefficient:  $\mathbb{L}(Y \mid \mathbf{R}) = \mathbf{R}' \pmb{\beta}$

$$\boldsymbol{\beta} = \left(\mathbb{E}[\mathsf{RR}']\right)^{-1}\mathbb{E}[\mathsf{R}Y]$$

Also holds if we get residuals from projection of Y on Z:

$$V = Y - \mathbb{L}(Y \mid \mathbf{Z}).$$

$$\mathbb{L}(V \mid \mathbf{R}) = \mathbf{R}' \boldsymbol{\beta}$$

#### **Omitted variable bias**

• Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

• Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

•  $\mathbb{L}[Y \mid \mathbf{X}, Z]$  is the long regression,  $\mathbb{L}[Y \mid \mathbf{X}]$  is the short regression.

Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

- $\mathbb{L}[Y \mid \mathbf{X}, Z]$  is the long regression,  $\mathbb{L}[Y \mid \mathbf{X}]$  is the short regression.
- How do  $\beta$  and  $\delta$  relate? Use law of iterated projections:

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}\left\{\mathbb{L}[Y \mid \mathbf{X}, Z] \mid \mathbf{X}\right\} = \mathbb{L}[\mathbf{X} \mid \mathbf{X}]'\boldsymbol{\beta} + \mathbb{L}[Z \mid \mathbf{X}]\boldsymbol{\gamma}$$

Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

- $\mathbb{L}[Y \mid \mathbf{X}, Z]$  is the long regression,  $\mathbb{L}[Y \mid \mathbf{X}]$  is the short regression.
- How do  $\beta$  and  $\delta$  relate? Use law of iterated projections:

$$\mathbb{L}[Y\mid \mathbf{X}] = \mathbb{L}\left\{\mathbb{L}[Y\mid \mathbf{X},Z]\mid \mathbf{X}\right\} = \mathbb{L}[\mathbf{X}\mid \mathbf{X}]'\boldsymbol{\beta} + \mathbb{L}[Z\mid \mathbf{X}]\boldsymbol{\gamma}$$

• First regress/project Z on X:  $\mathbb{L}[Z \mid X] = X'\pi$  and so:

$$\mathbb{L}[Y \mid X] = X'(\beta + \pi \gamma), \qquad \delta = \beta + \pi \gamma$$

• Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

- $\mathbb{L}[Y \mid \mathbf{X}, Z]$  is the long regression,  $\mathbb{L}[Y \mid \mathbf{X}]$  is the short regression.
- How do  $\beta$  and  $\delta$  relate? Use law of iterated projections:

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}\left\{\mathbb{L}[Y \mid \mathbf{X}, Z] \mid \mathbf{X}\right\} = \mathbb{L}[\mathbf{X} \mid \mathbf{X}]'\boldsymbol{\beta} + \mathbb{L}[Z \mid \mathbf{X}]\boldsymbol{\gamma}$$

• First regress/project Z on X:  $\mathbb{L}[Z \mid X] = X' \pi$  and so:

$$\mathbb{L}[Y \mid X] = X'(\beta + \pi \gamma), \qquad \delta = \beta + \pi \gamma$$

•  $\delta - \beta = \pi \gamma$  is the "bias" but this is misleading.

Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

- $\mathbb{L}[Y \mid \mathbf{X}, Z]$  is the long regression,  $\mathbb{L}[Y \mid \mathbf{X}]$  is the short regression.
- How do  $\beta$  and  $\delta$  relate? Use law of iterated projections:

$$\mathbb{L}[Y \mid \mathbf{X}] = \mathbb{L}\left\{\mathbb{L}[Y \mid \mathbf{X}, Z] \mid \mathbf{X}\right\} = \mathbb{L}[\mathbf{X} \mid \mathbf{X}]'\boldsymbol{\beta} + \mathbb{L}[Z \mid \mathbf{X}]\boldsymbol{\gamma}$$

• First regress/project Z on X:  $\mathbb{L}[Z \mid X] = X' \pi$  and so:

$$\mathbb{L}[Y \mid X] = X'(\beta + \pi \gamma), \qquad \delta = \beta + \pi \gamma$$

- $\delta \beta = \pi \gamma$  is the "bias" but this is misleading.
  - $oldsymbol{eta}$  not necessarily "correct", we're just relating two projections

Consider two projections/regressions with and without some Z:

$$\mathbb{L}[Y \mid \mathbf{X}, \mathbf{Z}] = \mathbf{X}' \boldsymbol{\beta} + Z \gamma, \qquad \mathbb{L}[Y \mid \mathbf{X}] = \mathbf{X}' \boldsymbol{\delta}$$

- $\mathbb{L}[Y \mid \mathbf{X}, Z]$  is the long regression,  $\mathbb{L}[Y \mid \mathbf{X}]$  is the short regression.
- How do  $\beta$  and  $\delta$  relate? Use law of iterated projections:

$$\mathbb{L}[Y\mid \mathbf{X}] = \mathbb{L}\left\{\mathbb{L}[Y\mid \mathbf{X},Z]\mid \mathbf{X}\right\} = \mathbb{L}[\mathbf{X}\mid \mathbf{X}]'\boldsymbol{\beta} + \mathbb{L}[Z\mid \mathbf{X}]\boldsymbol{\gamma}$$

• First regress/project Z on X:  $\mathbb{L}[Z \mid X] = X'\pi$  and so:

$$\mathbb{L}[Y \mid X] = X'(\beta + \pi \gamma), \qquad \delta = \beta + \pi \gamma$$

- $\delta \beta = \pi \gamma$  is the "bias" but this is misleading.
  - $\cdot$   $oldsymbol{eta}$  not necessarily "correct", we're just relating two projections
  - Difference is (coef of excluded) × (effect of included on excluded)

• What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .
  - But  $\mu(\mathbf{X})$  could be nonlinear, what then?

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .
  - But  $\mu(\mathbf{X})$  could be nonlinear, what then?
- Linear projection justification: best linear approximation to  $\mu(\mathbf{X})$ :

$$\pmb{\beta} = \mathop{\arg\min}_{\mathbf{b} \in \mathbb{R}^K} \mathbb{E}\left[\left(\mu(\mathbf{X}) - \mathbf{X}' \pmb{\beta}\right)^2\right]$$

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .
  - But  $\mu(\mathbf{X})$  could be nonlinear, what then?
- Linear projection justification: best linear approximation to  $\mu(\mathbf{X})$ :

$$oldsymbol{eta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^K} \mathbb{E}\left[\left(\mu(\mathbf{X}) - \mathbf{X}' oldsymbol{eta}
ight)^2
ight]$$

• Linear projection is best linear approximation to Y and  $\mathbb{E}[Y \mid X]$ .

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .
  - But  $\mu(\mathbf{X})$  could be nonlinear, what then?
- Linear projection justification: best linear approximation to  $\mu(\mathbf{X})$ :

$$\pmb{\beta} = \mathop{\arg\min}_{\mathbf{b} \in \mathbb{R}^K} \mathbb{E}\left[\left(\mu(\mathbf{X}) - \mathbf{X}' \pmb{\beta}\right)^2\right]$$

- Linear projection is best linear approximation to Y and  $\mathbb{E}[Y \mid X]$ .
- · Limitations:

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .
  - But  $\mu(\mathbf{X})$  could be nonlinear, what then?
- Linear projection justification: best linear approximation to  $\mu(\mathbf{X})$ :

$$\pmb{\beta} = \mathop{\arg\min}_{\mathbf{b} \in \mathbb{R}^K} \mathbb{E}\left[\left(\mu(\mathbf{X}) - \mathbf{X}' \pmb{\beta}\right)^2\right]$$

- Linear projection is best linear approximation to Y and  $\mathbb{E}[Y \mid X]$ .
- · Limitations:
  - If nonlinearity of  $\mu(\mathbf{X})$  is severe,  $\mathbb{L}[Y \mid X]$  can only be so good.

- What is the relationship between  $\mathbb{L}[Y \mid \mathbf{X}]$  and  $\mu(\mathbf{X}) = \mathbb{E}[Y \mid \mathbf{X}]$ ?
  - If  $\mu(\mathbf{X})$  is linear, then  $\mu(\mathbf{X}) = \mathbb{L}[Y \mid \mathbf{X}]$ .
  - But  $\mu(\mathbf{X})$  could be nonlinear, what then?
- Linear projection justification: best linear approximation to  $\mu(\mathbf{X})$ :

$$oldsymbol{eta} = \mathop{\mathrm{arg\,min}}_{\mathbf{b} \in \mathbb{R}^K} \mathbb{E}\left[\left(\mu(\mathbf{X}) - \mathbf{X}' oldsymbol{eta}\right)^2\right]$$

- Linear projection is best linear approximation to Y and  $\mathbb{E}[Y \mid X]$ .
- · Limitations:
  - If nonlinearity of  $\mu(\mathbf{X})$  is severe,  $\mathbb{L}[Y \mid X]$  can only be so good.
  - $\mathbb{L}[Y \mid X]$  can be sensitive to the marginal distribution of X.

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

• "The Linear Model": is this an assumption?

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- ullet Depends on what we assume about the error, e

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}' \boldsymbol{\beta}$

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- ullet Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
  - If just  $\mathbb{E}[\mathbf{X}e] = 0$ , then this is just a linear projection.

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}' \boldsymbol{\beta}$
  - If just  $\mathbb{E}[\mathbf{X}e] = 0$ , then this is just a linear projection.
  - First is very strong, second is very mild.

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- ullet Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
  - If just  $\mathbb{E}[\mathbf{X}e] = 0$ , then this is just a linear projection.
  - · First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
  - If just  $\mathbb{E}[\mathbf{X}e] = 0$ , then this is just a linear projection.
  - · First is very strong, second is very mild.
- · Why do we care? Affects the properties of OLS.
  - · Some finite-sample properties of OLS (unbiasedness) require linear CEF

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
  - If just  $\mathbb{E}[\mathbf{X}e] = 0$ , then this is just a linear projection.
  - · First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.
  - · Some finite-sample properties of OLS (unbiasedness) require linear CEF
  - Asymptotic results (consistency, asymptotic normality) apply to both.

$$Y = \mathbf{X}'\boldsymbol{\beta} + e$$

- "The Linear Model": is this an assumption?
- Depends on what we assume about the error, e
  - If  $\mathbb{E}[e \mid \mathbf{X}] = 0$ , then we are assuming the CEF is linear,  $\mathbb{E}[Y \mid X] = \mathbf{X}'\boldsymbol{\beta}$
  - If just  $\mathbb{E}[\mathbf{X}e] = 0$ , then this is just a linear projection.
  - · First is very strong, second is very mild.
- Why do we care? Affects the properties of OLS.
  - · Some finite-sample properties of OLS (unbiasedness) require linear CEF
  - Asymptotic results (consistency, asymptotic normality) apply to both.
  - OLS will consitently estimate something, but maybe not what you want.