6. Multivariate Distributions

Spring 2021

Matthew Blackwell

Gov 2002 (Harvard)

Where are we? Where are we going?

 Distributions of one variable: how to describe and summarize uncertainty about one variable.

Where are we? Where are we going?

- Distributions of one variable: how to describe and summarize uncertainty about one variable.
- Today: distributions of multiple variables to describe relationships between variables.

Where are we? Where are we going?

- Distributions of one variable: how to describe and summarize uncertainty about one variable.
- Today: distributions of multiple variables to describe relationships between variables.
- Later: use data to **learn** about probability distributions.

Why multiple random variables?

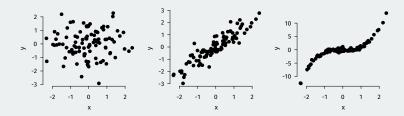
1. How to measure the relationship between two variables X and Y?

Why multiple random variables?

- 1. How to measure the relationship between two variables X and Y?
- 2. What if we have many observations of the same variable, X_1, X_2, \dots, X_n ?

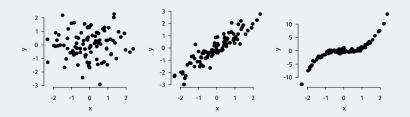
1/ Distributions of Multiple Random Variables

Joint distributions



 The joint distribution of two r.v.s, X and Y, describes what pairs of observations, (x, y) are more likely than others.

Joint distributions



- The joint distribution of two r.v.s, X and Y, describes what pairs of observations, (x, y) are more likely than others.
- Shape of the joint distribution \leadsto the relationship between X and Y

Definition

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

Definition

The **joint probability mass function (p.m.f.)** of a pair of discrete r.v.s, (X, Y) describes the probability of any pair of values:

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

• Properties of a joint p.m.f.:

Definition

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

- Properties of a joint p.m.f.:
 - $f_{X,Y}(x,y) \ge 0$

Definition

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

- · Properties of a joint p.m.f.:
 - $f_{X,Y}(x,y) \ge 0$ (probabilities can't be negative)

Definition

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

- · Properties of a joint p.m.f.:
 - $f_{X,Y}(x,y) \ge 0$ (probabilities can't be negative)
 - $\sum_{x} \sum_{y} f_{X,Y}(x,y) = 1$

Definition

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

- · Properties of a joint p.m.f.:
 - $f_{X,Y}(x,y) \ge 0$ (probabilities can't be negative)
 - $\sum_{x} \sum_{y} f_{X,Y}(x,y) = 1$ (something must happen)

Definition

$$f_{X,Y}(x,y) = \mathbb{P}(X=x,Y=y)$$

- · Properties of a joint p.m.f.:
 - $f_{X,Y}(x,y) \ge 0$ (probabilities can't be negative)
 - $\sum_{x} \sum_{y} f_{X,Y}(x,y) = 1$ (something must happen)
 - \sum_{x} is shorthand for sum over all possible values of X

	Support Gay	Oppose Gay
	Marriage	Marriage
	Y=1	Y = 0
Female $X = 1$	0.30	0.21
Male $X = 0$	0.22	0.27

• Joint p.m.f. can be summarized in a cross-tab:

	Support Gay	Oppose Gay
	Marriage	Marriage
	Y=1	Y = 0
Female $X = 1$	0.30	0.21
Male $X = 0$	0.22	0.27

- Joint p.m.f. can be summarized in a cross-tab:
 - Each is the probability of that combination, $p_{X,Y}(x,y)$

	Support Gay	Oppose Gay
	Marriage	Marriage
	Y=1	Y = 0
Female $X = 1$	0.30	0.21
Male $X = 0$	0.22	0.27

- Joint p.m.f. can be summarized in a cross-tab:
 - Each is the probability of that combination, $p_{X,Y}(x,y)$
- Probability that we randomly select a woman who supports gay marriage?

	Support Gay	Oppose Gay
	Marriage	Marriage
	Y=1	Y = 0
Female $X = 1$	0.30	0.21
Male $X = 0$	0.22	0.27

- Joint p.m.f. can be summarized in a cross-tab:
 - Each is the probability of that combination, $p_{X,Y}(x,y)$
- Probability that we randomly select a woman who supports gay marriage?

$$p_{X,Y}(1,1) = \mathbb{P}(X=1, Y=1) = 0.3$$

• Can we get the distribution of just one of the r.v.s alone?

- Can we get the distribution of just one of the r.v.s alone?
 - Called the marginal distribution in this context.

- Can we get the distribution of just one of the r.v.s alone?
 - · Called the marginal distribution in this context.
- Computing marginal p.m.f. from the joint p.m.f.:

$$\mathbb{P}(Y=y) = \sum_{x} \mathbb{P}(X=x, Y=y)$$

- Can we get the distribution of just one of the r.v.s alone?
 - Called the marginal distribution in this context.
- Computing marginal p.m.f. from the joint p.m.f.:

$$\mathbb{P}(Y = y) = \sum_{x} \mathbb{P}(X = x, Y = y)$$

 Intuition: sum over the probability that Y = y and X = x for all possible values of x

- Can we get the distribution of just one of the r.v.s alone?
 - · Called the marginal distribution in this context.
- · Computing marginal p.m.f. from the joint p.m.f.:

$$\mathbb{P}(Y = y) = \sum_{x} \mathbb{P}(X = x, Y = y)$$

- Intuition: sum over the probability that Y = y and X = x for all
 possible values of x
 - Called marginalizing out X.

- Can we get the distribution of just one of the r.v.s alone?
 - · Called the marginal distribution in this context.
- · Computing marginal p.m.f. from the joint p.m.f.:

$$\mathbb{P}(Y=y) = \sum_{x} \mathbb{P}(X=x, Y=y)$$

- Intuition: sum over the probability that Y = y and X = x for all
 possible values of x
 - Called marginalizing out X.
 - Works because values of X are disjoint.

	Support Gay Marriage	Oppose Gay Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	
Male $X = 0$	0.22	0.27	
Marginal			

• What's $\mathbb{P}(Y=1)$?

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	
Male $X = 0$	0.22	0.27	
Marginal			

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(\mathit{Y}=1)$$

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	
Male $X = 0$	0.22	0.27	
Marginal			

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(Y=1)=\mathbb{P}(X=1,Y=1)+\mathbb{P}(X=0,Y=1)$$

	Support Gay Marriage $Y = 1$	Oppose Gay Marriage $Y = 0$	Marginal
Female $X = 1$	0.30	0.21	
Male $X = 0$	0.22	0.27	
Marginal	0.30 + 0.22		

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(Y=1) = \mathbb{P}(X=1, Y=1) + \mathbb{P}(X=0, Y=1) = 0.3 + 0.22$$

	Support Gay Marriage $Y = 1$	Oppose Gay Marriage $Y = 0$	Marginal
Female $X = 1$	0.30	0.21	
Male $X = 0$	0.22	0.27	
Marginal	0.52		

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(Y=1) = \mathbb{P}(X=1, Y=1) + \mathbb{P}(X=0, Y=1) = 0.3 + 0.22 = 0.52$$

	Support Gay Marriage Y=1	Oppose Gay Marriage $Y = 0$	Marginal
Female $X = 1$	0.30	0.21	0.30 + 0.21
Male $X = 0$	0.22	0.27	0.22 + 0.27
Marginal	0.52	0.21 + 0.27	

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(Y=1) = \mathbb{P}(X=1, Y=1) + \mathbb{P}(X=0, Y=1) = 0.3 + 0.22 = 0.52$$

Works for all marginals.

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	0.51
Male $X = 0$	0.22	0.27	0.49
Marginal	0.52	0.48	

- What's $\mathbb{P}(Y=1)$?
 - Probability that a man supports gay marriage plus the probability that a woman supports gay marriage.

$$\mathbb{P}(Y=1) = \mathbb{P}(X=1, Y=1) + \mathbb{P}(X=0, Y=1) = 0.3 + 0.22 = 0.52$$

· Works for all marginals.

Conditional p.m.f.

Definition

The **conditional probability mass function** or conditional p.m.f. of Y conditional on X is

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(X = x)}$$

for all values x s.t. $\mathbb{P}(X = x) > 0$.

Conditional p.m.f.

Definition

The **conditional probability mass function** or conditional p.m.f. of Y conditional on X is

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(X = x)}$$

for all values x s.t. $\mathbb{P}(X = x) > 0$.

· This is a valid univariate probability distribution!

Conditional p.m.f.

Definition

The **conditional probability mass function** or conditional p.m.f. of *Y* conditional on *X* is

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(X = x)}$$

for all values x s.t. $\mathbb{P}(X = x) > 0$.

This is a valid univariate probability distribution!

•
$$P(Y = y \mid X = x) \ge 0$$
 and $\sum_{y} \mathbb{P}(Y = y \mid X = x) = 1$

Conditional p.m.f.

Definition

The **conditional probability mass function** or conditional p.m.f. of *Y* conditional on *X* is

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(X = x)}$$

for all values x s.t. $\mathbb{P}(X = x) > 0$.

· This is a valid univariate probability distribution!

•
$$P(Y = y \mid X = x) \ge 0$$
 and $\sum_{y} \mathbb{P}(Y = y \mid X = x) = 1$

• Can define the **conditional expectation** of this p.m.f.:

$$E[Y \mid X = x] = \sum_{y} y \mathbb{P}(Y = y \mid X = x)$$

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	0.51
Male $X = 0$	0.22	0.27	0.49
Marginal	0.52	0.48	

• Probability of favoring gay marriage conditional on **male**?

	Support Gay Marriage	Oppose Gay Marriage	Marginal
	Y=1	Y = 0	J
Female $X = 1$	0.30	0.21	0.51
Male $X = 0$	0.22	0.27	0.49
Marginal	0.52	0.48	

• Probability of favoring gay marriage conditional on **male**?

$$\mathbb{P}(Y=1\mid X=0)$$

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	0.51
Male $X = 0$	0.22	0.27	0.49
Marginal	0.52	0.48	

Probability of favoring gay marriage conditional on male?

$$\mathbb{P}(Y = 1 \mid X = 0) = \frac{\mathbb{P}(X = 0, Y = 1)}{\mathbb{P}(X = 0)}$$

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	0.51
Male $X = 0$	0.22	0.27	0.49
Marginal	0.52	0.48	

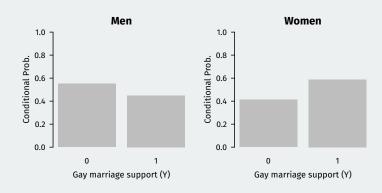
Probability of favoring gay marriage conditional on male?

$$\mathbb{P}(Y=1 \mid X=0) = \frac{\mathbb{P}(X=0, Y=1)}{\mathbb{P}(X=0)} = \frac{0.22}{0.22 + 0.27}$$

	Support Gay	Oppose Gay	
	Marriage	Marriage	Marginal
	Y = 1	Y = 0	
Female $X = 1$	0.30	0.21	0.51
Male $X = 0$	0.22	0.27	0.49
Marginal	0.52	0.48	

· Probability of favoring gay marriage conditional on male?

$$\mathbb{P}(Y=1 \mid X=0) = \frac{\mathbb{P}(X=0, Y=1)}{\mathbb{P}(X=0)} = \frac{0.22}{0.22 + 0.27} = 0.449$$



• Two values of $X \rightsquigarrow$ two **univariate** conditional distributions of Y

Bayes and LTP

· Bayes' rule for r.v.s:

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x \mid Y = y)\mathbb{P}(Y = y)}{\mathbb{P}(X = x)}$$

Bayes and LTP

· Bayes' rule for r.v.s:

$$\mathbb{P}(Y = y \mid X = x) = \frac{\mathbb{P}(X = x \mid Y = y)\mathbb{P}(Y = y)}{\mathbb{P}(X = x)}$$

· Law of total probability for r.v.s:

$$\mathbb{P}(X = x) = \sum_{y} \mathbb{P}(X = x \mid Y = y) \mathbb{P}(Y = y)$$

Joint c.d.f.s

Definition

For two r.v.s X and Y, the **joint cumulative distribution function** or joint c.d.f. $F_{X,Y}(x,y)$ is a function such that for finite values x and y,

$$F_{X,Y}(x,y) = \mathbb{P}(X \le x, Y \le y)$$

Joint c.d.f.s

Definition

For two r.v.s X and Y, the **joint cumulative distribution function** or joint c.d.f. $F_{X,Y}(x,y)$ is a function such that for finite values x and y,

$$F_{X,Y}(x,y) = \mathbb{P}(X \le x, Y \le y)$$

Well-defined for discrete and continuous X and Y.

Joint c.d.f.s

Definition

For two r.v.s X and Y, the **joint cumulative distribution function** or joint c.d.f. $F_{X,Y}(x,y)$ is a function such that for finite values x and y,

$$F_{X,Y}(x,y) = \mathbb{P}(X \le x, Y \le y)$$

- Well-defined for discrete and continuous X and Y.
- · For discrete we simply have:

$$F_{X,Y}(x,y) = \sum_{i \le x} \sum_{j \le y} \mathbb{P}(X = i, Y = j)$$

Continuous r.v.s

• One continuous r.v.: prob. of being in a subset of the real line.



Continuous r.v.s

• One continuous r.v.: prob. of being in a subset of the real line.



 Two continuous r.v.s: probability of being in some subset of the 2-dimensional plane.



Continuous joint p.d.f.

Definition

If two continuous r.v.s X and Y with joint c.d.f. $F_{X,Y}$, their **joint p.d.f.** $f_{X,Y}(x,y)$ is the derivative of $F_{X,Y}$ with respect to x and y,

$$f_{X,Y}(x,y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x,y)$$

• Integrate over both dimensions to get the probability of a region:

$$\mathbb{P}((X,Y)\in A)=\iint_{(x,y)\in A}f_{X,Y}(x,y)dxdy.$$

Continuous joint p.d.f.

Definition

If two continuous r.v.s X and Y with joint c.d.f. $F_{X,Y}$, their **joint p.d.f.** $f_{X,Y}(x,y)$ is the derivative of $F_{X,Y}$ with respect to x and y,

$$f_{X,Y}(x,y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x,y)$$

• Integrate over both dimensions to get the probability of a region:

$$\mathbb{P}((X,Y) \in A) = \iint_{(x,y) \in A} f_{X,Y}(x,y) dx dy.$$

• $\{(x,y): f_{X,Y}(x,y) > 0\}$ is called the **support** of the distribution.

• Joint p.d.f. must meet the following conditions:

• Joint p.d.f. must meet the following conditions:

1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y),

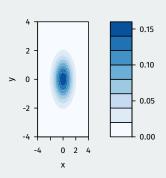
- Joint p.d.f. must meet the following conditions:
 - 1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y), (nonnegative)

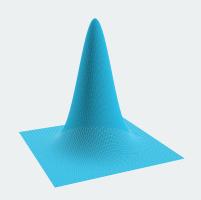
- Joint p.d.f. must meet the following conditions:
 - 1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y), (nonnegative)
 - 2. $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1,$

- Joint p.d.f. must meet the following conditions:
 - 1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y), (nonnegative)
 - 2. $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1$, (probabilities "sum" to 1)

- Joint p.d.f. must meet the following conditions:
 - 1. $f_{X,Y}(x,y) \ge 0$ for all values of (x,y), (nonnegative)
 - 2. $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1$, (probabilities "sum" to 1)
- $\mathbb{P}(X = x, Y = y) = 0$ for similar reasons as with single r.v.s.

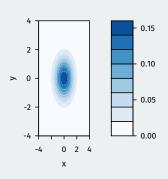
Joint densities are 3D

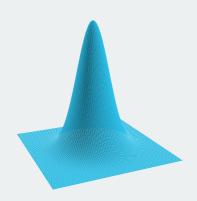




• X and Y axes are on the "floor," height is the value of $f_{X,Y}(x,y)$.

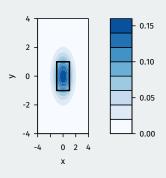
Joint densities are 3D

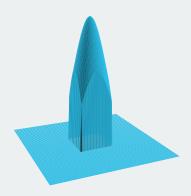




- X and Y axes are on the "floor," height is the value of $f_{X,Y}(x,y)$.
- Remember $f_{X,Y}(x,y) \neq \mathbb{P}(X=x,Y=y)$.

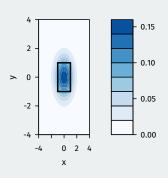
Probability = volume

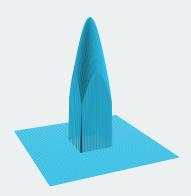




•
$$\mathbb{P}((X,Y) \in A) = \iint_{(x,y)\in A} f_{X,Y}(x,y) dx dy$$

Probability = volume





- $\mathbb{P}((X,Y) \in A) = \iint_{(X,Y) \in A} f_{X,Y}(x,y) dx dy$
- Probability = volume above a specific region.

Continuous marginal distributions

 We can recover the marginal PDF of one of the variables by integrating over the distribution of the other variable:

$$f_{Y}(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

Continuous marginal distributions

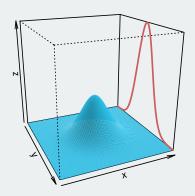
 We can recover the marginal PDF of one of the variables by integrating over the distribution of the other variable:

$$f_{Y}(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

· Works for either variable:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$

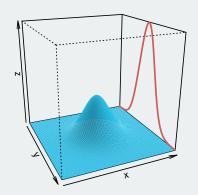
Visualizing continuous marginals



• Marginal integrates (sums, basically) over other r.v.:

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

Visualizing continuous marginals



Marginal integrates (sums, basically) over other r.v.:

$$f_{Y}(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

• Pile up/flatten all of the joint density onto a single dimension.

Continuous conditional distributions

Definition

The conditional p.d.f. of a continuous random variable is

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

for all values x s.t. $f_X(x) > 0$.

Implies

$$\mathbb{P}(a < Y < b | X = x) = \int_a^b f_{Y|X}(y|x) dy$$

Continuous conditional distributions

Definition

The conditional p.d.f. of a continuous random variable is

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

for all values x s.t. $f_X(x) > 0$.

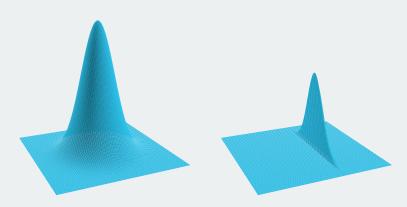
Implies

$$\mathbb{P}(a < Y < b | X = x) = \int_a^b f_{Y|X}(y|x) dy$$

 Based on the definition of the conditional p.m.f./p.d.f., we have the following factorization:

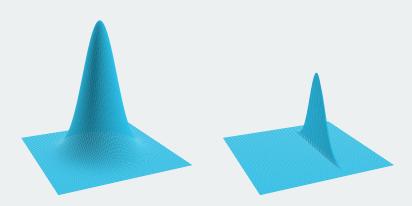
$$f_{X,Y}(x,y) = f_{Y|X}(y|x)f_X(x)$$

Conditional distributions as slices



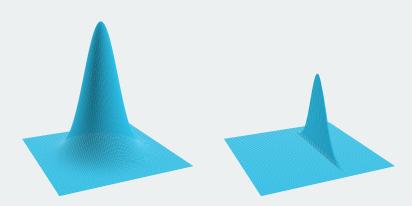
+ $f_{Y|X}(y|x_0)$ is the conditional p.d.f. of Y when $X=x_0$

Conditional distributions as slices



- $f_{Y|X}(y|x_0)$ is the conditional p.d.f. of Y when $X=x_0$
- + $f_{Y|X}(y|x_0)$ is proportional to joint p.d.f. along x_0 : $f_{X,Y}(y,x_0)$

Conditional distributions as slices



- $f_{Y|X}(y|x_0)$ is the conditional p.d.f. of Y when $X=x_0$
- $f_{Y|X}(y|x_0)$ is proportional to joint p.d.f. along x_0 : $f_{X,Y}(y,x_0)$
- Normalize by dividing by $f_X(x_0)$ to ensure proper p.d.f.

Independence

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp \!\!\! \perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

• Knowing the value of X gives us no information about the value of Y.

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp \!\!\! \perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

- Knowing the value of X gives us no information about the value of Y.
- If X and Y are independent, then:

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp \!\!\! \perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

- Knowing the value of X gives us no information about the value of Y.
- If X and Y are independent, then:
 - $f_{X,Y}(x,y) = f_X(x)f_Y(y)$ (joint is the product of marginals)

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp\!\!\!\perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

- Knowing the value of X gives us no information about the value of Y.
- If X and Y are independent, then:
 - $f_{X,Y}(x,y) = f_X(x)f_Y(y)$ (joint is the product of marginals)
 - $F_{X,Y}(x,y) = F_X(x)F_Y(y)$

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp\!\!\!\perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

- Knowing the value of X gives us no information about the value of Y.
- If X and Y are independent, then:
 - $f_{X,Y}(x,y) = f_X(x)f_Y(y)$ (joint is the product of marginals)
 - $F_{X,Y}(x,y) = F_X(x)F_Y(y)$
 - $f_{Y|X}(y|x) = f_Y(y)$ (conditional is the marginal)

Independence

Two r.v.s Y and X are **independent** (which we write $X \perp\!\!\!\perp Y$) if for all sets A and B:

$$\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$$

- Knowing the value of X gives us no information about the value of Y.
- If X and Y are independent, then:
 - $f_{X,Y}(x,y) = f_X(x)f_Y(y)$ (joint is the product of marginals)
 - $F_{X,Y}(x,y) = F_X(x)F_Y(y)$
 - $f_{Y|X}(y|x) = f_Y(y)$ (conditional is the marginal)
- Conditional independence implies similar to conditional distributions:

$$\mathbb{P}(X \in A, Y \in B \mid Z) = \mathbb{P}(X \in A \mid Z)\mathbb{P}(Y \in B \mid Z)$$

2/ Expectations of Joint Distributions

Properties of joint distributions

• Single r.v.: summarized $f_X(x)$ with $\mathbb{E}[X]$ and $\mathbb{V}[X]$

Properties of joint distributions

- Single r.v.: summarized $f_X(x)$ with $\mathbb{E}[X]$ and $\mathbb{V}[X]$
- With 2 r.v.s: how strong is the dependence is between *X* and *Y*?

Properties of joint distributions

- Single r.v.: summarized $f_X(x)$ with $\mathbb{E}[X]$ and $\mathbb{V}[X]$
- With 2 r.v.s: how strong is the dependence is between X and Y?
- First: expectations over joint distributions.

• 2-d LOTUS: take expectations over the joint distribution.

- **2-d LOTUS**: take expectations over the joint distribution.
- With discrete X and Y:

$$\mathbb{E}[g(X,Y)] = \sum_{x} \sum_{y} g(x,y) \ f_{X,Y}(x,y)$$

- 2-d LOTUS: take expectations over the joint distribution.
- With discrete X and Y:

$$\mathbb{E}[g(X,Y)] = \sum_{x} \sum_{y} g(x,y) \ f_{X,Y}(x,y)$$

• With continuous X and Y:

$$\mathbb{E}[g(X,Y)] = \int_{X} \int_{Y} g(x,y) f_{X,Y}(x,y) dx dy$$

- 2-d LOTUS: take expectations over the joint distribution.
- With discrete X and Y:

$$\mathbb{E}[g(X,Y)] = \sum_{x} \sum_{y} g(x,y) \ f_{X,Y}(x,y)$$

With continuous X and Y:

$$\mathbb{E}[g(X,Y)] = \int_{X} \int_{Y} g(x,y) f_{X,Y}(x,y) dx dy$$

· Marginal expectations:

$$\mathbb{E}[Y] = \sum_{x} \sum_{y} y \ f_{X,Y}(x,y)$$

Theorem

If X and Y are independent r.v.s, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

Theorem

If X and Y are independent r.v.s, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

Theorem

If X and Y are independent r.v.s, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

$$\mathbb{E}[XY] = \sum_{x} \sum_{y} xy \ f_{X,Y}(x,y)$$

Theorem

If X and Y are independent r.v.s, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

$$\mathbb{E}[XY] = \sum_{x} \sum_{y} xy \ f_{X,Y}(x,y)$$
$$= \sum_{x} \sum_{y} xy \ f_{X}(x)f_{Y}(y)$$

Theorem

If X and Y are independent r.v.s, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

$$\mathbb{E}[XY] = \sum_{x} \sum_{y} xy \ f_{X,Y}(x,y)$$
$$= \sum_{x} \sum_{y} xy \ f_{X}(x) f_{Y}(y)$$
$$= \left(\sum_{x} x \ f_{X}(x)\right) \left(\sum_{y} y \ f_{Y}(y)\right)$$

Theorem

If X and Y are independent r.v.s, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

$$\begin{split} \mathbb{E}[XY] &= \sum_{x} \sum_{y} xy \ f_{X,Y}(x,y) \\ &= \sum_{x} \sum_{y} xy \ f_{X}(x) f_{Y}(y) \\ &= \left(\sum_{x} x \ f_{X}(x) \right) \left(\sum_{y} y \ f_{Y}(y) \right) \\ &= \mathbb{E}[X] \mathbb{E}[Y] \end{split}$$

3/ Covariance and Correlation

• Independence assumptions are **everywhere** in statistics.

- Independence assumptions are **everywhere** in statistics.
 - Each response in a poll is considered independent of all other responses.

- Independence assumptions are everywhere in statistics.
 - Each response in a poll is considered independent of all other responses.
 - In a randomized control trial, treatment assignment is independent of background characteristics.

- · Independence assumptions are everywhere in statistics.
 - Each response in a poll is considered independent of all other responses.
 - In a randomized control trial, treatment assignment is independent of background characteristics.
- · Lack of independence is a blessing or a curse:

- Independence assumptions are everywhere in statistics.
 - Each response in a poll is considered independent of all other responses.
 - In a randomized control trial, treatment assignment is independent of background characteristics.
- Lack of independence is a blessing or a curse:
 - Two variables not independent → potentially interesting relationship.

- Independence assumptions are everywhere in statistics.
 - Each response in a poll is considered independent of all other responses.
 - In a randomized control trial, treatment assignment is independent of background characteristics.
- Lack of independence is a blessing or a curse:
 - Two variables not independent → potentially interesting relationship.
 - In observational studies, treatment assignment is usually not independent of background characteristics.

• How do we measure the strength of the dependence between two r.v.?

· How do we measure the strength of the dependence between two r.v.?

Covariance

$$\mathsf{Cov}[X,Y] = \mathbb{E}\Big[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\Big]$$

· How do we measure the strength of the dependence between two r.v.?

Covariance

The **covariance** between two r.v.s, *X* and *Y* is defined as:

$$\mathsf{Cov}[X,Y] = \mathbb{E}\Big[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\Big]$$

• How often do high values of X occur with high values of Y?

· How do we measure the strength of the dependence between two r.v.?

Covariance

$$\mathrm{Cov}[X,Y] = \mathbb{E}\Big[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\Big]$$

- How often do high values of X occur with high values of Y?
- · Properties of covariances:

· How do we measure the strength of the dependence between two r.v.?

Covariance

$$\mathsf{Cov}[X,Y] = \mathbb{E}\Big[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\Big]$$

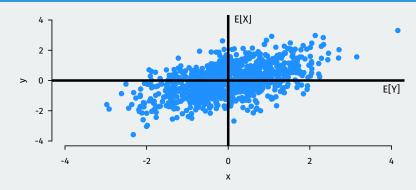
- How often do high values of X occur with high values of Y?
- · Properties of covariances:
 - $Cov[X, Y] = \mathbb{E}[XY] \mathbb{E}[X]\mathbb{E}[Y]$

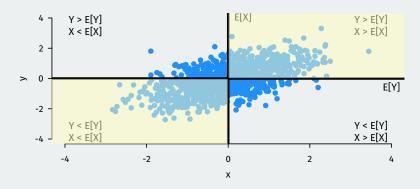
· How do we measure the strength of the dependence between two r.v.?

Covariance

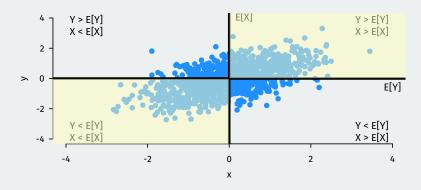
$$\mathrm{Cov}[X,Y] = \mathbb{E}\Big[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\Big]$$

- How often do high values of X occur with high values of Y?
- · Properties of covariances:
 - $Cov[X, Y] = \mathbb{E}[XY] \mathbb{E}[X]\mathbb{E}[Y]$
 - If $X \perp \!\!\! \perp Y$, then Cov[X, Y] = 0

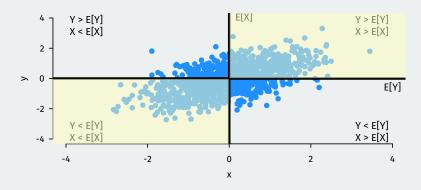




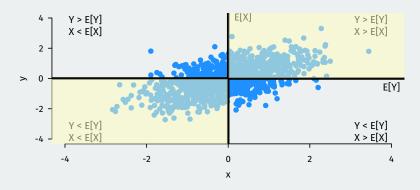
• Large values of *X* tend to occur with large values of *Y*:



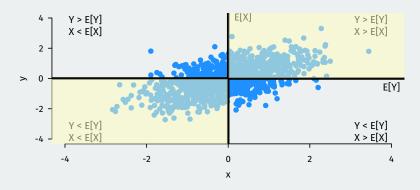
- Large values of *X* tend to occur with large values of *Y*:
 - $\bullet \ (X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{pos. num}) = +$



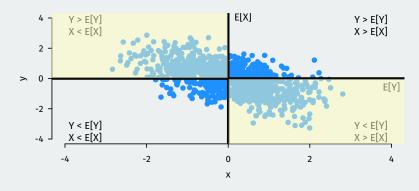
- Large values of *X* tend to occur with large values of *Y*:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{pos. num}) = +$
- Small values of X tend to occur with small values of Y:



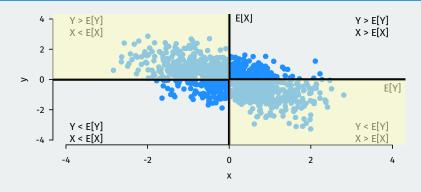
- Large values of *X* tend to occur with large values of *Y*:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{pos. num}) = +$
- Small values of X tend to occur with small values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{neg. num.}) \times (\text{neg. num}) = +$



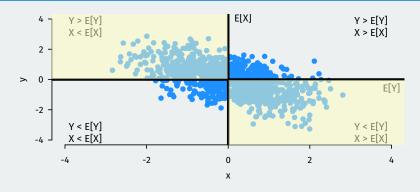
- Large values of *X* tend to occur with large values of *Y*:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{pos. num}) = +$
- Small values of X tend to occur with small values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{neg. num.}) \times (\text{neg. num}) = +$
- If these dominate → positive covariance.



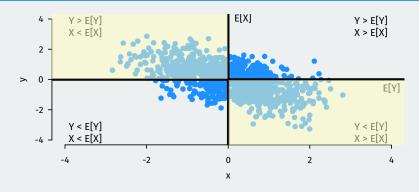
• Large values of *X* tend to occur with small values of *Y*:



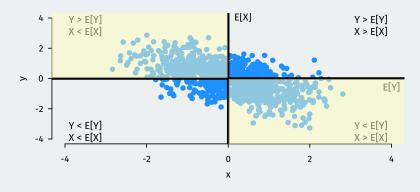
- Large values of X tend to occur with small values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{neg. num}) = -$



- Large values of X tend to occur with small values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{neg. num}) = -$
- Small values of X tend to occur with large values of Y:



- Large values of X tend to occur with small values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{neg. num}) = -$
- Small values of X tend to occur with large values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{neg. num.}) \times (\text{pos. num}) = -$



- Large values of X tend to occur with small values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{pos. num.}) \times (\text{neg. num}) = -$
- Small values of X tend to occur with large values of Y:
 - $(X \mathbb{E}[X])(Y \mathbb{E}[Y]) = (\text{neg. num.}) \times (\text{pos. num}) = -$

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

• Properties of covariances:

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

· Properties of covariances:

1.
$$Cov[X, X] = V[X]$$

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

- Properties of covariances:
 - 1. Cov[X, X] = V[X]
 - 2. Cov[X, Y] = Cov[Y, X]

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

- Properties of covariances:
 - 1. Cov[X, X] = V[X]
 - 2. Cov[X, Y] = Cov[Y, X]
 - 3. Cov[X, c] = 0 for any constant c

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

- · Properties of covariances:
 - 1. Cov[X, X] = V[X]
 - 2. Cov[X, Y] = Cov[Y, X]
 - 3. Cov[X, c] = 0 for any constant c
 - 4. Cov[aX, Y] = aCov[X, Y].

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

- · Properties of covariances:
 - 1. Cov[X, X] = V[X]
 - 2. Cov[X, Y] = Cov[Y, X]
 - 3. Cov[X, c] = 0 for any constant c
 - 4. Cov[aX, Y] = aCov[X, Y].
 - 5. Cov[X + Y, Z] = Cov[X, Z] + Cov[Y, Z]

$$\mathsf{Cov}[X,Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

- Properties of covariances:
 - 1. Cov[X, X] = V[X]
 - 2. Cov[X, Y] = Cov[Y, X]
 - 3. Cov[X, c] = 0 for any constant c
 - 4. Cov[aX, Y] = aCov[X, Y].
 - 5. Cov[X + Y, Z] = Cov[X, Z] + Cov[Y, Z]
 - $\mathbf{6.} \ \, \mathsf{Cov}[X+Y,Z+W] = \mathsf{Cov}[X,Z] + \mathsf{Cov}[Y,Z] + \mathsf{Cov}[X,W] + \mathsf{Cov}[Y,W]$

• Can now state a few more properties of variances.

- · Can now state a few more properties of variances.
- · Variance of a sum:

$$\mathbb{V}[X+Y] = \mathbb{V}[X] + \mathbb{V}[Y] + 2\mathsf{Cov}[X,Y]$$

- · Can now state a few more properties of variances.
- · Variance of a sum:

$$\mathbb{V}[X+Y] = \mathbb{V}[X] + \mathbb{V}[Y] + 2\mathsf{Cov}[X,Y]$$

• More generally for n r.v.s X_1, \dots, X_n :

$$\mathbb{V}[X_1+\cdots+X_n]=\mathbb{V}[X_1]+\cdots+\mathbb{V}[X_n]+2\sum_{i< j}\operatorname{Cov}(X_i,X_j)$$

- · Can now state a few more properties of variances.
- · Variance of a sum:

$$\mathbb{V}[X+Y] = \mathbb{V}[X] + \mathbb{V}[Y] + 2\mathsf{Cov}[X,Y]$$

• More generally for n r.v.s X_1, \dots, X_n :

$$\mathbb{V}[X_1+\cdots+X_n]=\mathbb{V}[X_1]+\cdots+\mathbb{V}[X_n]+2\sum_{i< j}\operatorname{Cov}(X_i,X_j)$$

• If X and Y independent, V[X + Y] = V[X] + V[Y].

- · Can now state a few more properties of variances.
- · Variance of a sum:

$$\mathbb{V}[X+Y] = \mathbb{V}[X] + \mathbb{V}[Y] + 2\mathsf{Cov}[X,Y]$$

• More generally for n r.v.s X_1, \dots, X_n :

$$\mathbb{V}[X_1+\cdots+X_n]=\mathbb{V}[X_1]+\cdots+\mathbb{V}[X_n]+2\sum_{i< j}\operatorname{Cov}(X_i,X_j)$$

- If X and Y independent, V[X + Y] = V[X] + V[Y].
 - Beware: V[X Y] = V[X] + V[Y] as well.

• We saw that $X \perp \!\!\! \perp Y \leadsto \text{Cov}[X,Y] = 0$.

- We saw that $X \perp \!\!\! \perp Y \leadsto \text{Cov}[X,Y] = 0$.
- Does Cov[X, Y] = 0 imply that $X \perp \!\!\! \perp Y$?

- We saw that $X \perp \!\!\!\perp Y \leadsto \text{Cov}[X,Y] = 0$.
- Does Cov[X, Y] = 0 imply that $X \perp \!\!\! \perp Y$? **No!**

- We saw that $X \perp \!\!\!\perp Y \rightsquigarrow Cov[X, Y] = 0$.
- Does Cov[X, Y] = 0 imply that $X \perp \!\!\! \perp Y$? **No!**
- Counterexample: $X \in \{-1,0,1\}$ with equal probability and $Y = X^2$.

- We saw that $X \perp \!\!\!\perp Y \rightsquigarrow Cov[X, Y] = 0$.
- Does Cov[X, Y] = 0 imply that $X \perp \!\!\! \perp Y$? **No!**
- Counterexample: $X \in \{-1, 0, 1\}$ with equal probability and $Y = X^2$.
- Covariance is a measure of linear dependence, so it can miss non-linear dependence.

• Correlation is a scale-free measure of linear dependence.

· Correlation is a scale-free measure of linear dependence.

Definition

$$\rho = \rho(X, Y) = \frac{\mathsf{Cov}[X, Y]}{\sqrt{\mathbb{V}[X]\mathbb{V}[Y]}} = \mathsf{Cov}\left(\frac{X - \mathbb{E}[X]}{SD[X]}, \frac{Y - \mathbb{E}[Y]}{SD[Y]}\right)$$

· Correlation is a scale-free measure of linear dependence.

Definition

The **correlation** between two r.v.s X and Y is defined as:

$$\rho = \rho(X, Y) = \frac{\mathsf{Cov}[X, Y]}{\sqrt{\mathbb{V}[X]\mathbb{V}[Y]}} = \mathsf{Cov}\left(\frac{X - \mathbb{E}[X]}{SD[X]}, \frac{Y - \mathbb{E}[Y]}{SD[Y]}\right)$$

· Covariance after dividing out the scales of the respective variables.

· Correlation is a scale-free measure of linear dependence.

Definition

$$\rho = \rho(X,Y) = \frac{\mathsf{Cov}[X,Y]}{\sqrt{\mathbb{V}[X]\mathbb{V}[Y]}} = \mathsf{Cov}\left(\frac{X - \mathbb{E}[X]}{SD[X]}, \frac{Y - \mathbb{E}[Y]}{SD[Y]}\right)$$

- · Covariance after dividing out the scales of the respective variables.
- · Correlation properties:

· Correlation is a scale-free measure of linear dependence.

Definition

$$\rho = \rho(X,Y) = \frac{\mathsf{Cov}[X,Y]}{\sqrt{\mathbb{V}[X]\mathbb{V}[Y]}} = \mathsf{Cov}\left(\frac{X - \mathbb{E}[X]}{SD[X]}, \frac{Y - \mathbb{E}[Y]}{SD[Y]}\right)$$

- · Covariance after dividing out the scales of the respective variables.
- Correlation properties:

•
$$-1 \le \rho \le 1$$

· Correlation is a scale-free measure of linear dependence.

Definition

$$\rho = \rho(X,Y) = \frac{\mathsf{Cov}[X,Y]}{\sqrt{\mathbb{V}[X]\mathbb{V}[Y]}} = \mathsf{Cov}\left(\frac{X - \mathbb{E}[X]}{SD[X]}, \frac{Y - \mathbb{E}[Y]}{SD[Y]}\right)$$

- · Covariance after dividing out the scales of the respective variables.
- · Correlation properties:
 - $-1 \le \rho \le 1$
 - $|\rho(X,Y)| = 1$ if and only if X and Y are perfectly correlated with a deterministic linear relationship: Y = a + bX.

4/ Random vectors

• When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m
 - x is now a length-m vector and potential value of X

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m
 - x is now a length-m vector and potential value of X
 - Generalizes all ideas from 2 variables to $\it m$

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m
 - x is now a length-m vector and potential value of X
 - Generalizes all ideas from 2 variables to m
- Joint distribution function: $F(x) = \mathbb{P}(X \le x) = \mathbb{P}(X_1 \le x_1, \dots, X_m \le x_m)$.

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m
 - x is now a length-m vector and potential value of X
 - Generalizes all ideas from 2 variables to m
- Joint distribution function: $F(x) = \mathbb{P}(X \le x) = \mathbb{P}(X_1 \le x_1, \dots, X_m \le x_m)$.
 - Discrete: joint p.m.f. $\mathbb{P}(X = x)$.

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m
 - x is now a length-m vector and potential value of X
 - Generalizes all ideas from 2 variables to m
- Joint distribution function: $F(x) = \mathbb{P}(X \le x) = \mathbb{P}(X_1 \le x_1, \dots, X_m \le x_m)$.
 - Discrete: joint p.m.f. $\mathbb{P}(X = x)$.
 - · Continuous: joint p.d.f.

$$f(x) = \frac{\partial^m}{\partial x_1 \cdots \partial x_m} F(x)$$

- When we have many r.v.s, we sometimes group them into random vectors $X = (X_1, \dots, X_m)^T$
 - X is a function from the sample space to \mathbb{R}^m
 - x is now a length-m vector and potential value of X
 - Generalizes all ideas from 2 variables to m
- Joint distribution function: $F(x) = \mathbb{P}(X \le x) = \mathbb{P}(X_1 \le x_1, \dots, X_m \le x_m)$.
 - Discrete: joint p.m.f. $\mathbb{P}(X = x)$.
 - · Continuous: joint p.d.f.

$$f(x) = \frac{\partial^m}{\partial x_1 \cdots \partial x_m} F(x)$$

• Expectation of a random vector is just the vector of expectations:

$$\mathbb{E}[X] = \left(\mathbb{E}[X_1], \mathbb{E}[X_2], \dots, \mathbb{E}[X_m]\right)^T$$

· Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[X] = \mathbb{E}\left[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T \right]$$

· Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[X] = \mathbb{E}\left[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T\right]$$

• We usually write $V[X] = \Sigma$ and it is a $m \times m$ symmetric matrix:

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1m} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{m1} & \sigma_{m2} & \cdots & \sigma_m^2 \end{pmatrix}$$

where,
$$\sigma_j^2 = \mathbb{V}[X_j]$$
 and $\sigma_{ij} = \text{Cov}(X_i, X_j)$.

Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[X] = \mathbb{E}\left[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T\right]$$

• We usually write $V[X] = \Sigma$ and it is a $m \times m$ symmetric matrix:

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1m} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{m1} & \sigma_{m2} & \cdots & \sigma_m^2 \end{pmatrix}$$

where,
$$\sigma_j^2 = \mathbb{V}[X_j]$$
 and $\sigma_{ij} = \text{Cov}(X_i, X_j)$.

• Symmetric ($\Sigma = \Sigma^T$) because $Cov(X_i, X_i) = Cov(X_i, X_i)$.

Theorem

If $X \in \mathbb{R}^m$ with $m \times 1$ expectation μ and $m \times m$ covariance matrix Σ , and \mathbf{A} is a $q \times m$ matrix, then $\mathbf{A}X$ is a random vector with mean $\mathbf{A}\mu$ and covariance matrix $\mathbf{A}\Sigma\mathbf{A}^T$.

- Can group r.v.s into random vectors $\mathbf{X} = (X_1, \dots, X_k)'$

- Can group r.v.s into random vectors $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k

- Can group r.v.s into random vectors $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k
 - \mathbf{x} is now a length-k vector and potential value of \mathbf{X}

- Can group r.v.s into random vectors $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k
 - x is now a length-k vector and potential value of X
 - Generalizes all ideas from 2 variables to k

- Can group r.v.s into random vectors $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k
 - x is now a length-k vector and potential value of X
 - Generalizes all ideas from 2 variables to k
- Joint distribution function: $F(\mathbf{x}) = \mathbb{P}(\mathbf{X} \leq \mathbf{x}) = \mathbb{P}(X_1 \leq x_1, \dots, X_k \leq x_k)$.

- Can group r.v.s into random vectors $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k
 - x is now a length-k vector and potential value of X
 - Generalizes all ideas from 2 variables to k
- Joint distribution function: $F(\mathbf{x}) = \mathbb{P}(\mathbf{X} \leq \mathbf{x}) = \mathbb{P}(X_1 \leq x_1, \dots, X_k \leq x_k)$.
 - Discrete: joint p.m.f. $\mathbb{P}(\mathbf{X} = \mathbf{x})$.

- Can group r.v.s into **random vectors** $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k
 - x is now a length-k vector and potential value of X
 - Generalizes all ideas from 2 variables to k
- Joint distribution function: $F(\mathbf{x}) = \mathbb{P}(\mathbf{X} \leq \mathbf{x}) = \mathbb{P}(X_1 \leq x_1, \dots, X_k \leq x_k)$.
 - Discrete: joint p.m.f. $\mathbb{P}(\mathbf{X} = \mathbf{x})$.
 - · Continuous: joint p.d.f.

$$f(\mathbf{x}) = \frac{\partial^m}{\partial x_1 \cdots \partial x_k} F(\mathbf{x})$$

- Can group r.v.s into **random vectors** $\mathbf{X} = (X_1, \dots, X_k)'$
 - **X** is a function from the sample space to \mathbb{R}^k
 - x is now a length-k vector and potential value of X
 - Generalizes all ideas from 2 variables to k
- Joint distribution function: $F(\mathbf{x}) = \mathbb{P}(\mathbf{X} \leq \mathbf{x}) = \mathbb{P}(X_1 \leq x_1, \dots, X_k \leq x_k)$.
 - Discrete: joint p.m.f. $\mathbb{P}(\mathbf{X} = \mathbf{x})$.
 - · Continuous: joint p.d.f.

$$f(\mathbf{x}) = \frac{\partial^m}{\partial x_1 \cdots \partial x_k} F(\mathbf{x})$$

• Expectation of a random vector is just the vector of expectations:

$$\mathbb{E}[\mathbf{X}] = \left(\mathbb{E}[X_1], \mathbb{E}[X_2], \dots, \mathbb{E}[X_k]\right)'$$

· Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[\mathbf{X}] = \mathbb{E}\left[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])' \right]$$

Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[\mathbf{X}] = \mathbb{E}\left[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])' \right]$$

• We usually write $\mathbb{V}[\mathbf{X}] = \mathbf{\Sigma}$ and it is a $k \times k$ symmetric matrix:

$$oldsymbol{\Sigma} = egin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1k} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2k} \\ dots & dots & \ddots & dots \\ \sigma_{k1} & \sigma_{k2} & \cdots & \sigma_k^2 \end{pmatrix}$$

where, $\sigma_j^2 = \mathbb{V}[X_j]$ and $\sigma_{ij} = \operatorname{Cov}(X_i, X_j)$.

Covariance matrix generalizes (co)variance to this setting:

$$\mathbb{V}[\mathbf{X}] = \mathbb{E}\left[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])' \right]$$

• We usually write $V[X] = \Sigma$ and it is a $k \times k$ symmetric matrix:

$$oldsymbol{\Sigma} = egin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1k} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2k} \\ dots & dots & \ddots & dots \\ \sigma_{k1} & \sigma_{k2} & \cdots & \sigma_k^2 \end{pmatrix}$$

where, $\sigma_j^2 = \mathbb{V}[X_j]$ and $\sigma_{ij} = \mathsf{Cov}(X_i, X_j)$.

• Symmetric ($\Sigma = \Sigma'$) because $Cov(X_i, X_j) = Cov(X_j, X_i)$.

• Let $\mathbf{Z} = Z_1, Z_2, \dots, Z_k$ be i.i.d. $\mathcal{N}(0,1)$. What is their joint distribution?

- Let $\mathbf{Z} = Z_1, Z_2, \dots, Z_k$ be i.i.d. $\mathcal{N}(0,1)$. What is their joint distribution?
- For vector of values $\mathbf{z} = (z_1, z_2, \dots, z_k)^T$

$$f(\mathbf{z}) = \frac{1}{(2\pi)^{k/2}} \exp\left(-\frac{\mathbf{z}'\mathbf{z}}{2}\right)$$

- Let $\mathbf{Z} = Z_1, Z_2, \dots, Z_k$ be i.i.d. $\mathcal{N}(0,1)$. What is their joint distribution?
- For vector of values $\mathbf{z} = (z_1, z_2, \dots, z_k)^T$

$$f(\mathbf{z}) = \frac{1}{(2\pi)^{k/2}} \exp\left(-\frac{\mathbf{z}'\mathbf{z}}{2}\right)$$

• Easy to see the mean/variance: $\mathbb{E}[\mathbf{Z}] = 0$ and $\mathbb{V}[\mathbf{Z}] = \mathbf{I}_k$.

- Let $\mathbf{Z} = Z_1, Z_2, \dots, Z_k$ be i.i.d. $\mathcal{N}(0,1)$. What is their joint distribution?
- For vector of values $\mathbf{z} = (z_1, z_2, \dots, z_k)^T$

$$f(\mathbf{z}) = \frac{1}{(2\pi)^{k/2}} \exp\left(-\frac{\mathbf{z}'\mathbf{z}}{2}\right)$$

- Easy to see the mean/variance: $\mathbb{E}[\mathbf{Z}] = \mathbf{0}$ and $\mathbb{V}[\mathbf{Z}] = \mathbf{I}_k$.
 - \mathbf{I}_k is the k by k identity matrix because $\mathbb{V}[Z_j]=1$ and $\mathrm{Cov}(Z_i,Z_j)=0$.

Theorem

If $\mathbf{X} \in \mathbb{R}^k$ with $k \times 1$ expectation $\boldsymbol{\mu}$ and $k \times k$ covariance matrix $\boldsymbol{\Sigma}$, and $\boldsymbol{\Lambda}$ is a $q \times k$ matrix, then $\boldsymbol{\Lambda} \mathbf{X}$ is a random vector with mean $\boldsymbol{\Lambda} \boldsymbol{\mu}$ and covariance matrix $\boldsymbol{\Lambda} \boldsymbol{\Sigma} \boldsymbol{\Lambda}'$.

Theorem

If $\mathbf{X} \in \mathbb{R}^k$ with $k \times 1$ expectation $\boldsymbol{\mu}$ and $k \times k$ covariance matrix $\boldsymbol{\Sigma}$, and \boldsymbol{A} is a $q \times k$ matrix, then $\mathbf{A} \boldsymbol{X}$ is a random vector with mean $\mathbf{A} \boldsymbol{\mu}$ and covariance matrix $\mathbf{A} \boldsymbol{\Sigma} \mathbf{A}'$.

• Let $Z \sim \mathcal{N}(0, I_k)$ and $X = \mu + BZ$, where B is $q \times k$ then $X \sim \mathcal{N}(\mu, \Sigma)$

Theorem

If $\mathbf{X} \in \mathbb{R}^k$ with $k \times 1$ expectation $\boldsymbol{\mu}$ and $k \times k$ covariance matrix $\boldsymbol{\Sigma}$, and \boldsymbol{A} is a $q \times k$ matrix, then $\mathbf{A} \boldsymbol{X}$ is a random vector with mean $\mathbf{A} \boldsymbol{\mu}$ and covariance matrix $\mathbf{A} \boldsymbol{\Sigma} \mathbf{A}'$.

- Let $\mathbf{Z} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_k)$ and $\mathbf{X} = \boldsymbol{\mu} + \mathbf{B}\mathbf{Z}$, where \mathbf{B} is $q \times k$ then $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$
 - μ : $q \times 1$ mean vector $\mathbb{E}[X] = \mu$

Theorem

If $\mathbf{X} \in \mathbb{R}^k$ with $k \times 1$ expectation $\boldsymbol{\mu}$ and $k \times k$ covariance matrix $\boldsymbol{\Sigma}$, and \boldsymbol{A} is a $q \times k$ matrix, then $\mathbf{A} \boldsymbol{X}$ is a random vector with mean $\mathbf{A} \boldsymbol{\mu}$ and covariance matrix $\mathbf{A} \boldsymbol{\Sigma} \mathbf{A}'$.

- Let $\mathbf{Z} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_k)$ and $\mathbf{X} = \boldsymbol{\mu} + \mathbf{B}\mathbf{Z}$, where \mathbf{B} is $q \times k$ then $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$
 - μ : $q \times 1$ mean vector $\mathbb{E}[X] = \mu$
 - $\Sigma = BB'$: $q \times q$ covariance matrix $V[X] = \Sigma$.

Theorem

If $\mathbf{X} \in \mathbb{R}^k$ with $k \times 1$ expectation $\boldsymbol{\mu}$ and $k \times k$ covariance matrix $\boldsymbol{\Sigma}$, and \boldsymbol{A} is a $q \times k$ matrix, then $\mathbf{A} \boldsymbol{X}$ is a random vector with mean $\mathbf{A} \boldsymbol{\mu}$ and covariance matrix $\mathbf{A} \boldsymbol{\Sigma} \mathbf{A}'$.

- Let $\mathbf{Z} \sim \mathcal{N}(0, \mathbf{I}_k)$ and $\mathbf{X} = \boldsymbol{\mu} + \mathbf{B}\mathbf{Z}$, where \mathbf{B} is $q \times k$ then $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$
 - μ : $q \times 1$ mean vector $\mathbb{E}[X] = \mu$
 - $\Sigma = BB'$: $q \times q$ covariance matrix $V[X] = \Sigma$.
- More generally, if $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ then $\mathbf{Y} = \mathbf{a} + \mathbf{B}\mathbf{X} \sim \mathcal{N}(\mathbf{a} + \mathbf{B}\boldsymbol{\mu}, \mathbf{B}\boldsymbol{\Sigma}\mathbf{B}')$

Properties of the multivariate normal

• If (X_1, X_2, X_3) are MVN, then (X_1, X_2) is also MVN.

Properties of the multivariate normal

- If (X_1, X_2, X_3) are MVN, then (X_1, X_2) is also MVN.
- If (X, Y) are multivariate normal with Cov(X, Y) = 0, then X and Y are independent.