# 13. Properties of Least Squares

Spring 2023

Matthew Blackwell

Gov 2002 (Harvard)

## Where are we? Where are we going?

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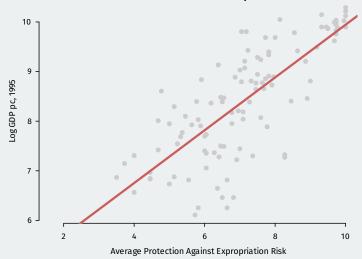
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#### Where are we? Where are we going?

- Before: learned about CEFs and linear projections in the population.
- · Last time: OLS estimator, its algebraic properties.
- Now: its statistical properties, both finite-sample and asymptotic.

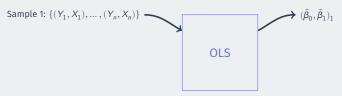
# Acemoglu, Johnson, and Robinson (2001)

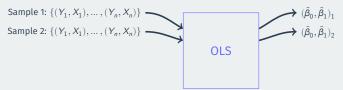


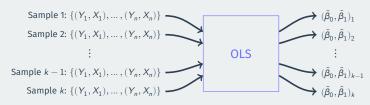


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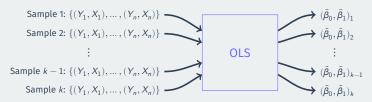
OLS



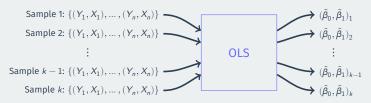




· OLS is an estimator—we plug data into and we get out estimates.



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- · Has a sampling distribution, with a sampling variance/standard error.

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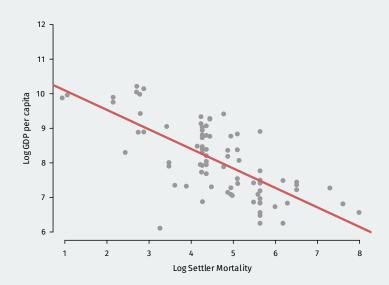
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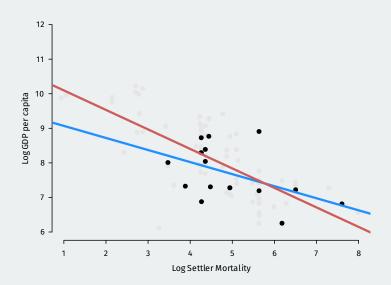
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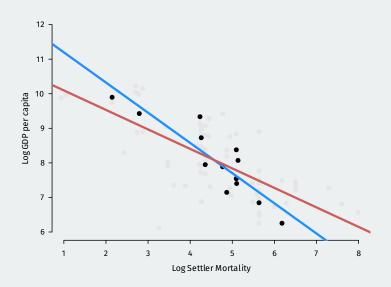
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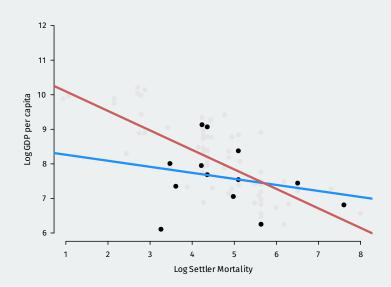
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- 3. Plot the estimated regression line

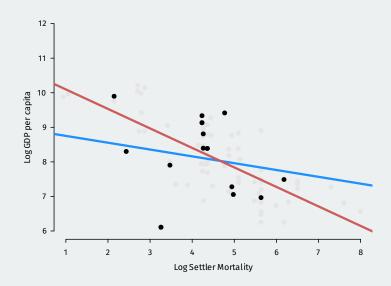
# **Population Regression**

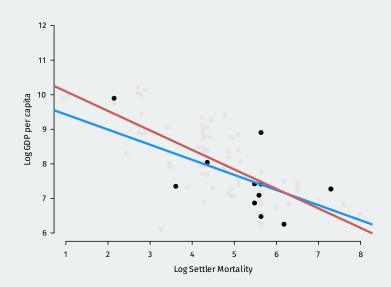


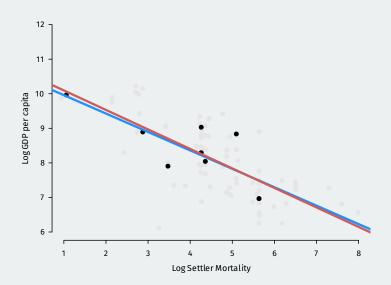


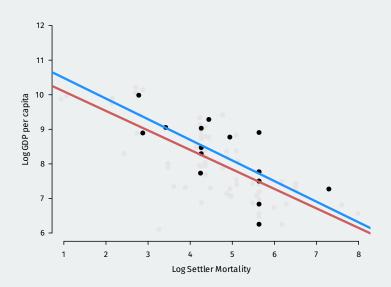












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### **Big picture**

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  - Linear projection model for asymptotic results.
  - Linear regression/CEF model for finite samples.

1/ Linear projection model and Large-sample Properties

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#### Linear projection model

$$Y = \mathbf{X}' \boldsymbol{\beta} + e$$
  
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1. For the variables  $(Y, \mathbf{X})$ , we assume the linear projection of Y on  $\mathbf{X}$  is defined as:

$$Y = \mathbf{X}' \boldsymbol{\beta} + e$$
  $\mathbb{E}[\mathbf{X}e] = 0.$ 

2. The design matrix is invertible, so  $\mathbb{E}[\mathbf{X}_i \mathbf{X}_i'] > 0$  (positive definite).

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    - Implies coefficients are  $\pmb{\beta} = (\mathbb{E}[\mathbf{X}\mathbf{X}'])^{-1}\mathbb{E}[\mathbf{X}Y]$
  - What properties can we derive under such weak assumptions?

$$\hat{\boldsymbol{\beta}} = \left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} Y_{i}\right) = \boldsymbol{\beta} + \underbrace{\left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} e_{i}\right)}_{\text{estimation error}}$$

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- Sample means in the estimation error follow the law of large numbers:

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•  $\mathbf{Q}_{\mathbf{X}\mathbf{X}}$  is invertible by assumption, so by the continuous mapping theorem:

$$\left(\frac{1}{n}\sum_{i=1}^{n}\mathbf{X}_{i}\mathbf{X}_{i}^{\prime}\right)^{-1}\overset{\rho}{\rightarrow}\mathbf{Q}_{\mathbf{XX}}^{-1}\quad\Longrightarrow\quad\hat{\boldsymbol{\beta}}\overset{\rho}{\rightarrow}\boldsymbol{\beta}+\mathbf{Q}_{\mathbf{XX}}^{-1}\cdot\mathbf{0}=\boldsymbol{\beta},$$

Theorem (Consistency of OLS)

Under the linear projection model and i.i.d. data,  $\hat{\beta}$  is consistent for  $\beta$ .

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  - If we have a linear CEF, then it's consistent for the CEF coefficients.
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- Not guaranteed to be unbiased (unless CEF is linear, as we'll see...)

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$$\mathbb{E}\left[\frac{1}{n}\sum_{i=1}^{n}g(\mathbf{X}_{i})\right] = \mathbb{E}[g(\mathbf{X}_{i})] \quad \text{var}\left[\frac{1}{n}\sum_{i=1}^{n}g(\mathbf{X}_{i})\right] = \frac{\text{var}[g(\mathbf{X}_{i})]}{n}$$

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· CLT implies:

$$\sqrt{n}\left(\frac{1}{n}\sum_{i=1}^n g(\mathbf{X}_i) - \mathbb{E}[g(\mathbf{X}_i)]\right) \overset{d}{\to} \mathcal{N}(0, \mathrm{var}[g(\mathbf{X}_i)])$$

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• If  $\mathbb{E}[g(\mathbf{X}_i)] = 0$ , then we have

$$\sqrt{n}\left(\frac{1}{n}\sum_{i=1}^{n}g(\mathbf{X}_{i})\right) = \frac{1}{\sqrt{n}}\sum_{i=1}^{n}g(\mathbf{X}_{i}) \stackrel{d}{\to} \mathcal{N}(0,\mathbb{E}[g(\mathbf{X}_{i})g(\mathbf{X}_{i})'])$$

$$\sqrt{n}\left(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}\right) = \left(\frac{1}{n}\sum_{i=1}^{n} \mathbf{X}_{i}\mathbf{X}_{i}'\right)^{-1}\left(\frac{1}{\sqrt{n}}\sum_{i=1}^{n} \mathbf{X}_{i}e_{i}\right)$$

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• Remember that  $(n^{-1}\sum_{i=1}^n \mathbf{X}_i\mathbf{X}_i')^{-1} \stackrel{p}{\to} \mathbf{Q}_{\mathbf{XX}}^{-1}$  so we have

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  - Rewrite as  $\sqrt{n}$  times an average of i.i.d. mean-zero random vectors.
- Let  $\mathbf{\Omega} = \mathbb{E}[e_i^2 \mathbf{X}_i \mathbf{X}_i']$  and apply the CLT:

$$\left(\frac{1}{\sqrt{n}}\sum_{i=1}^{n}\mathbf{X}_{i}e_{i}\right)\overset{d}{\rightarrow}\mathcal{N}(0,\mathbf{\Omega})$$

#### Theorem (Asymptotic Normality of OLS)

Under the linear projection model,

$$\sqrt{n}\left(\hat{\pmb{\beta}} - \pmb{\beta}\right) \overset{d}{\rightarrow} \mathcal{N}(\mathbf{0}, \mathbf{V}_{\pmb{\beta}}),$$

where,

$$\mathbf{V}_{\pmb{\beta}} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{\Omega}\mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1} = \left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}\mathbb{E}[e_i^2\mathbf{X}_i\mathbf{X}_i']\left(\mathbb{E}[\mathbf{X}_i\mathbf{X}_i']\right)^{-1}$$

+  $\hat{\pmb{\beta}}$  is approximately normal with mean  $\pmb{\beta}$  and variance  $\mathbf{Q}_{\mathbf{XX}}^{-1}\mathbf{\Omega}\mathbf{Q}_{\mathbf{XX}}^{-1}/n$ 

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- $\hat{m{eta}}$  is approximately normal with mean  $m{eta}$  and variance  $m{Q}_{m{X}m{X}}^{-1}m{\Omega}m{Q}_{m{X}m{X}}^{-1}/n$
- $\mathbf{V}_{\hat{m{eta}}} = \mathbf{V}_{m{eta}}/n$  is the asymptotic covariance matrix of  $\hat{m{eta}}$

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- +  $\hat{m{eta}}$  is approximately normal with mean  $m{eta}$  and variance  $m{Q}_{m{X}m{X}}^{-1}m{\Omega}m{Q}_{m{X}m{X}}^{-1}/n$
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#### Theorem (Asymptotic Normality of OLS)

Under the linear projection model,

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  - Square root of the diagonal of  $\mathbf{V}_{\hat{oldsymbol{eta}}}$  = standard errors for  $\hat{oldsymbol{eta}}_{j}$
- Allows us to formulate (approximate) confidence intervals, tests.

# **Estimating OLS variance**

$$\mathbf{V}_{\pmb{\beta}} = \mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}\mathbf{\Omega}\mathbf{Q}_{\mathbf{X}\mathbf{X}}^{-1}$$

- Estimation of  $\mathbf{V}_{\pmb{\beta}}$  uses plug-in estimators.

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  - Replace  $\mathbf{Q}_{\mathbf{XX}} = \mathbb{E}[\mathbf{X}_i \mathbf{X}_i']$  with  $n^{-1} \sum_{i=1}^n \mathbf{X}_i \mathbf{X}_i' = \mathbb{X}' \mathbb{X}/n$ .

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- · Putting these together:

$$\widehat{\mathbf{V}}_{\beta} = \left(\frac{1}{n} \mathbb{X}' \mathbb{X}\right)^{-1} \left(\frac{1}{n} \sum_{i=1}^{n} \hat{e}_{i}^{2} \mathbf{X}_{i} \mathbf{X}'_{i}\right) \left(\frac{1}{n} \mathbb{X}' \mathbb{X}\right)^{-1}$$
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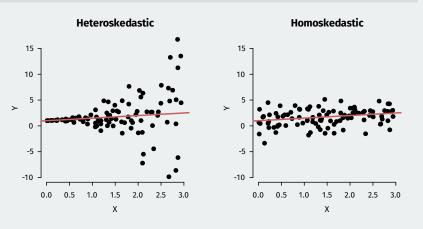
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  ightarrow} \mathbf{V}_{\pmb{\beta}}.$
- Square root of the diagonal of  $\widehat{\mathbf{V}}_{\hat{\beta}} = n^{-1}\widehat{\mathbf{V}}_{\hat{\beta}}$ : heteroskedasticity-consistent (HC) SEs

## Homoskedasticity

#### Assumption: Homoskedasticity

The variance of the error terms is constant in  $\mathbf{X}$ ,  $\mathbb{E}[e^2 \mid \mathbf{X}] = \sigma^2(\mathbf{X}) = \sigma^2$ .



• Homoskedasticity implies  $\mathbb{E}[e_i^2\mathbf{X}_i\mathbf{X}_i'] = \mathbb{E}[e_i^2]\mathbb{E}[\mathbf{X}_i\mathbf{X}_i'] = \sigma^2\mathbf{Q}_{\mathbf{XX}}$ 

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$$s^{2} = \frac{1}{n-k} \sum_{i=1}^{n} \hat{\mathbf{e}}_{i}^{2} \qquad \widehat{\mathbf{V}}_{\hat{\beta}}^{\text{lm}} = \frac{1}{n} s^{2} \left( \frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}' \right)^{-1} = s^{2} \left( \mathbb{X}' \mathbb{X} \right)^{-1}$$

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• LLN implies  $s^2\stackrel{p}{ o}\sigma^2$  and so  $n\widehat{m V}_{\hatm eta}^{lm}$  is consistent for  $m V_{m eta}^{lm}$ 

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  - Mostly small, ad hoc changes to improve finite-sample performance.

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• Software often uses t critical values instead of normal (we'll see why).

# 2/ Inference for Multiple Parameters

$$m(x,z) = \beta_0 + X\beta_1 + Z\beta_2 + XZ\beta_3$$

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- What if we want the variance of this effect for any value of Z?

$$\mathbb{V}\left(\frac{\partial \widehat{m}(x,z)}{\partial x}\right) = \mathbb{V}\left[\widehat{\beta}_1 + z\widehat{\beta}_3\right] = \mathbb{V}[\widehat{\beta}_1] + z^2 \mathbb{V}[\widehat{\beta}_3] + 2z \mathsf{cov}[\widehat{\beta}_1,\widehat{\beta}_3]$$

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· Use the estimated covariance matrix:

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$$m(X,Z) = \beta_0 + X\beta_1 + Z\beta_2 + XZ\beta_3$$

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  - · Distribution depends on the variance/covariance of the coefficients.
  - · Need to normalize like the t-statistic.

• Usually t-test of  $H_0: \beta_i = b_0$  based on the t-statistic:

$$t = rac{\hat{eta}_j - b_0}{\widehat{\mathsf{se}}(\hat{eta}_j)},$$

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  - Reminder:  $\chi_k^2$  is the sum of k squared standard normals.
  - Could get the critical value for  $t^2$  directly from  $\chi_1^2$ .

• We can rewrite the null hypothesis as  $H_0: \mathbf{L} \boldsymbol{\beta} = \mathbf{c}$  where,

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• If this covariance matrix where identity, then these would be standard normal and  $\hat{\beta}_1^2 + \hat{\beta}_3^2$  would be  $\chi_2^2$  under the null

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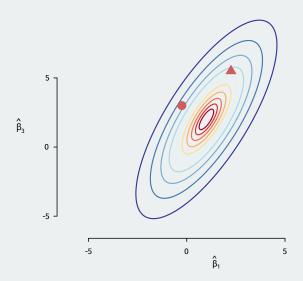
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# **Weighting by the distribution**



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  - "Usual" F-test reports test of all coef = 0 except intercept (pointless?)

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#### **Wald test steps**

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  - When applied to a single coefficient, equivalent to a t-test.
  - Use packages like {aod} or {clubSandwich} in R.

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- · Illustration:
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- By design, no effect of any variable on any other.

#### Multiple test example

noise <- data.frame(matrix(rnorm(2100), nrow = 100, ncol = 21))
summary(lm(noise))</pre>

```
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.028039
                          0.113820
                                      -0.25
                                              0.8061
## X2
              -0.150390
                          0.112181
                                      -1.34
                                              0.1839
## X3
                0.079158
                          0.095028
                                     0.83
                                              0.4074
## X4
              -0.071742
                          0.104579
                                      -0.69
                                              0.4947
## X5
                0.172078
                          0.114002
                                      1.51
                                              0.1352
## X6
                0.080852
                           0.108341
                                      0.75
                                              0.4577
## X7
                0.102913
                          0.114156
                                      0.90
                                              0.3701
## X8
              -0.321053
                          0.120673
                                      -2.66
                                              0.0094 **
## X9
              -0.053122
                          0.107983
                                      -0.49
                                              0.6241
## X10
                0.180105
                          0.126443
                                      1.42
                                              0.1583
## X11
                0.166386
                           0.110947
                                      1.50
                                              0.1377
## X12
               0.008011
                          0.103766
                                      0.08
                                              0.9387
## X13
               0.000212
                          0.103785
                                      0.00
                                              0.9984
## X14
              -0.065969
                           0.112214
                                      -0.59
                                              0.5583
## X15
              -0.129654
                           0.111575
                                      -1.16
                                              0.2487
                                              0.6647
## X16
              -0.054446
                           0.125140
                                      -0.44
## X17
                0.004335
                           0.112012
                                      0.04
                                              0.9692
## X18
              -0.080796
                           0.109853
                                      -0.74
                                              0.4642
## X19
              -0.085806
                           0.118553
                                      -0.72
                                              0.4713
## X20
              -0.186006
                          0.104560
                                      -1.78
                                              0.0791 .
## X21
                0.002111
                          0.108118
                                     0.02
                                              0.9845
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.999 on 79 degrees of freedom
## Multiple R-squared: 0.201. Adjusted R-squared: -0.00142
## F-statistic: 0.993 on 20 and 79 DF. p-value: 0.48
```

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- Bonferroni correction: use p-value cutoff  $\alpha/m$  where m is the number of hypotheses.
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  - Ensures that the family-wise error rate (probability of making at least 1 Type I error) is less than  $\alpha$ .

3/ Linear Regression Model and Finite-sample Properties

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- We continue to maintain  $\{(Y_i, \mathbf{X}_i)\}$  are i.i.d.

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• Useful when linearity holds by default (discrete X in experiments, etc)

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- $\bullet\,$  Upshot: OLS will have the smaller SEs than any other linear estimator.

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- With reasonable *n*, asymptotic normality has the same effect.