

SIMULATION FROM GLM

12/03/25

Two objectives:

- ① learn the idea behind simulation
- ② Recall the theory behind GLMs

We are going to consider 2 models :

- ① Poisson regression (count response)
- ② Logistic regression (classification model, binary response)

EX 1 POISSON REGRESSION

Y response (count) $X = X_1, \dots, X_p$

$$Y_i = Y | X = x_i, i = 1, \dots, n$$

y_1, \dots, y_n observations of Y_i
(independent but not
identically distributed)

	1	...	p
1	x_{11}	...	x_{1p}
...
i	x_{i1}	...	x_{ip}
...
n	x_{n1}	...	x_{np}

x_i y_i

$$\mu_i = E[Y_i] \in [0, \infty)$$

$$\eta_i = x_i^T \beta \in (-\infty, \infty)$$

$$\log(\mu_i) = \eta_i \text{ canonical link } (g)$$

$\Rightarrow \mu_i = e^{\eta_i}$

$$Y_i \sim \text{Poisson}(\mu_i) = \text{Poisson}(e^{\eta_i}) = \text{Poisson}(e^{x_i^T \beta})$$

$$Y_i \sim \text{Poisson}(e^{x_i^T \beta})$$

(8) e^{-1}

SIMULATION

$$p = 1, n = 100$$

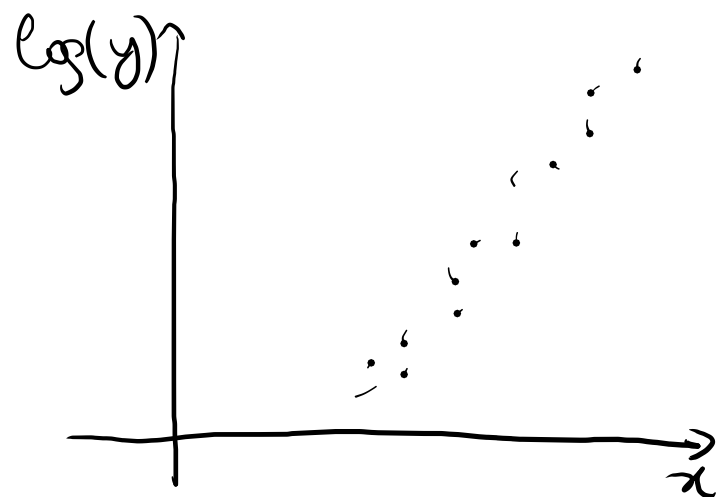
x_i
...
...
...
...

① Generate x_1, \dots, x_{100} from any distribution

$$\eta_i = \beta_0 + \beta_1 x_i$$

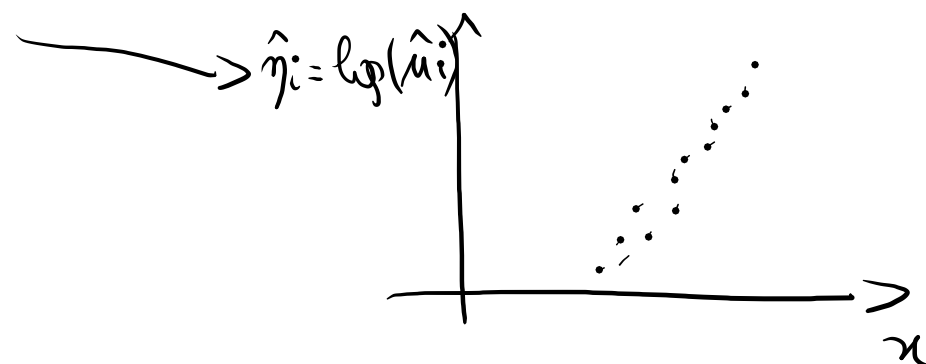
Fix $\beta_0 = 1, \beta_1 = 2$

③ Sample y_i from $\text{Poisson}(e^{\eta_i})$, $i = 1, \dots, 100$



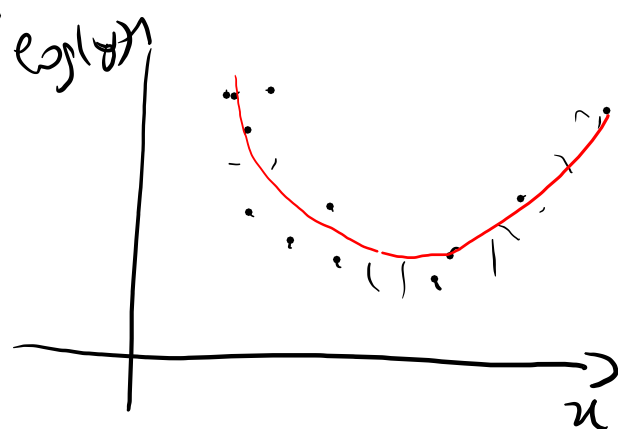
Non-parametric estimate of μ_i is

$$\hat{\mu}_i = y_i$$



$\log(\mu_i) = \beta_0 + \beta_1 x_i$
is a good idea!

If instead:



← This would suggest the inclusion of x^2 in the linear predictor

RESIDUALS

In linear models:

$$y_i = \eta_i + \varepsilon_i$$

(signal + noise)

$$\begin{aligned} e_i &= y_i - \hat{\eta}_i \\ &= y_i - \mathbf{x}_i^t \hat{\beta} \end{aligned}$$

\Leftrightarrow

$$y_i \sim \text{Normal}(\mathbf{x}_i^t \beta, \sigma^2)$$

(modern definition)



GLMs

For generalized linear models, the two main definitions of residuals are:

① Pearson residuals

② deviance residuals

PEARSON RESIDUALS

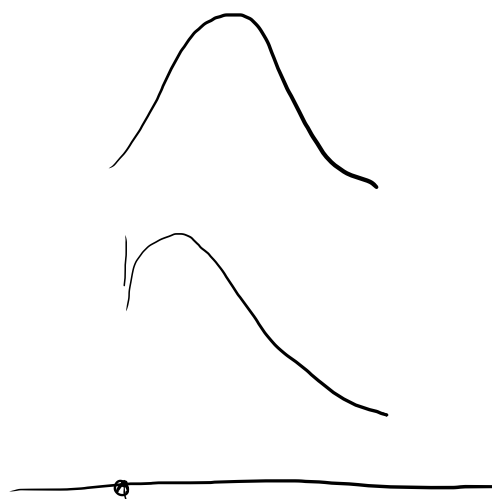
$$r_i^p = \frac{y_i - \hat{\mu}_i}{\sqrt{\widehat{\text{Var}}(y_i) / \hat{\phi}}}$$

For Poisson:

$$r_i^p = \frac{y_i - \hat{\mu}_i}{\sqrt{\hat{\mu}_i}}, i=1, \dots, n$$

$$\text{Var}(Y_i) = E[Y_i] = \mu_i$$

$$\phi = 1$$



→ Gaussian distribution for r_i^p is an approximation

→ The larger μ_i is, the better the approximation

DEVIANCE

The deviance of observation i is defined by:

$$d_i = 2 \left(\underset{\substack{\uparrow \\ \hat{\mu}_i = y_i}}{\ell_i(y_i; y_i)} - \underset{\substack{\uparrow \\ \text{log-likelihood}}}{\ell_i(\hat{\mu}_i; y_i)} \right), \quad i = 1, \dots, n$$

$\hat{\mu}_i = e^{x_i^T \hat{\beta}}$

$$d_i \geq 0$$

$$\Rightarrow D = \sum_{i=1}^n d_i \quad \text{deviance}$$

$$\Rightarrow r_i^D = \sqrt{d_i} \operatorname{sign}(y_i - \hat{\mu}_i) = \begin{cases} \sqrt{d_i} & y_i \geq \hat{\mu}_i \\ -\sqrt{d_i} & y_i < \hat{\mu}_i \end{cases}$$