Reading Group on Stochastic Modelling

A Structural Model of Dense Network Formation

Mele (2017)

Econometrica



Literature

Related works

- ▶ similar paper [6]
- extensions [11], [3], [9]

Model

- network formation models [1]
- **ERGM** theory [7], [15]

Asymptotics

- many networks asymptotics [14]
- large network asymptotics [8], [12], [10]
 - graph limits [2]
 - large deviations [5]
 - variational methods for the exponential family [16]

Estimation

ERGM estimation [15], [4]

Introduction

Strategic models of network formation provide a framework to interpret the observed network as the equilibrium of a game.

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- 1) multiple equilibria \Rightarrow links generate externalities not fully accounted for by agents
- 2) curse dimensionality \Rightarrow # network configs grows exponentially with # agents
- 3) data on single graph \Rightarrow only one network snapshot is observable

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Estimation and identification of strategic models is challenging

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Proposed model of network formation

- * combines features from the strategic and random network formation literature
- players' utilities depend on payoffs from direct links and link externalities (e.g., reciprocity, indirect friends, popularity, . . .)
- network formation is dynamic: each period, a player meets another one and decides whether to form a new link, keep an existing link, or do nothing
- * process generates a sequence of directed dense graphs

Model of Network Formation

Setup

- *n* agents, with characteristics $X_i \in \mathbb{R}^A \ \forall i \in \mathcal{I} := \{1, \dots, n\}$
- discrete time $t \in \mathbb{N}$
- ullet directed, binary network $G \in \mathcal{G}$, realisations each time g^t
- utility of network = sum utility from links, each having four components

$$U_{i}(g, X|\theta) = \underbrace{\sum_{j=1}^{n} g_{ij} u_{ij}^{\theta_{u}}}_{\text{direct links}} + \underbrace{\sum_{j=1}^{n} g_{ij} g_{ji} m_{ij}^{\theta_{m}}}_{\text{mutual links}} + \underbrace{\sum_{j=1}^{n} g_{ij} \sum_{\substack{k=1 \\ k \neq i,j}}^{n} g_{jk} v_{ik}^{\theta_{v}}}_{\text{indirect links}} + \underbrace{\sum_{j=1}^{n} g_{ij} \sum_{\substack{k=1 \\ k \neq i,j}}^{n} g_{ki} w_{kj}^{\theta_{w}}}_{\text{popularity}}$$

where
$$u_{ij}^{\theta_u} := u(X_i, X_j | \theta_u)$$
, $m_{ij}^{\theta_m} := m(X_i, X_j | \theta_m)$, $v_{ij}^{\theta_v} := v(X_i, X_j | \theta_v)$, $w_{ij}^{\theta_w} := w(X_i, X_j | \theta_w)$

Assumption 1. (Preferences)

$$m_{ij}^{\theta_m} = m(X_i, X_j | \theta_m) = m(X_j, X_i | \theta_m) = m_{ji}^{\theta_m} \quad \forall i, j \in \mathcal{I} \times \mathcal{I}$$

$$w_{ij}^{\theta_v} = w(X_k, X_j | \theta_v) = v(X_k, X_j | \theta_v) = v_{ji}^{\theta_v} \quad \forall k, j \in \mathcal{I} \times \mathcal{I}$$

- ▶ first is necessary for identification of the utility from indirect links and popularity;
- ▶ second makes another agent *i* internalise the externality she creates.

Proposition 1 (Existence Potential Function).

Under Assumption 1, the deterministic components of the incentives of any player in any state of the network are summarized by a **potential function** $\mathcal{Q}:\mathcal{G}\times\mathcal{X}\to\mathbb{R}$ and the network game is a Potential Game

$$Q(g, X|\theta) = \sum_{i=1}^{n} \sum_{j=1}^{n} g_{ij} u_{ij}^{\theta_{u}} + \sum_{i=1}^{n} \sum_{j>1}^{n} g_{ij} g_{ji} m_{ij}^{\theta_{m}} + \sum_{i=1}^{n} \sum_{\substack{j=1 \ j \neq 1}}^{n} \sum_{\substack{k=1 \ j \neq i}}^{n} g_{ij} g_{jk} v_{ik}^{\theta_{v}}$$

Q is an aggregate function summarising: (i) state of network; (ii) deterministic incentives of players in each state.

Network formation process

Process follows a **stochastic best-response dynamics**, that generates a Markov chain of networks

- for each t, randomly chosen player i meets j according to meeting technology
- meeting process is a stochastic sequence $\mathbf{m} = \{m^t\}_t$ supported on $\mathcal{I} \times \mathcal{I}$, with realisations $m^t = ij = \{i, j\}$

$$\mathbb{P}(m^t = ij|g^{t-1}, X) = \rho(g^{t-1}, X_i, X_j)$$

Assumption 2. (Meeting process)

The meeting probability between i, j does not depend on the existence of a link between them, and each meeting has a positive probability of occurring, that is

$$\rho(g^{t-1}, X_i, X_j) = \rho(g_{-ij}^{t-1}, X_i, X_j) > 0 \quad \forall ij$$

- guarantees any equilibrium network can be reached with positive probability
- \blacktriangleright identification: allow ρ to depend on <u>current</u> link g_{ij} prevents closed form likelihood

Utility

Players' rules

- conditional on meeting $m^t = ij$, player i updates link g_{ij} to maximise her utility
- existing network g_{-ii}^{t-1} is taken as given
- complete information: everybody known each others' attributes and whole network
- myopia: agents not account for effects of their linking strategy on future evolution of network
- idiosyncratic shock on individual preferences: $\varepsilon \sim EV_1(\varepsilon)$ Type I extreme value distribution, *iid* among links and across time
- link established if and only if

$$U_i(g_{ij}^t = 1, g_{-ij}^{t-1}, X|\theta) + \varepsilon_{1t} > U_i(g_{ij}^t = 0, g_{-ij}^{t-1}, X|\theta) + \varepsilon_{0t}$$

Process generates a Markov chain of networks:

- ✓ transition proba determined by: (i) meeting process, (ii) agents' linking choices
- ✓ irreducible, aperiodic

Equilibrium

Remark 1.

Any change in utility for any agent is equivalent to change in potential Q. So, any deviation from Nash (equilibrium) network must decrease the potential.

Thus, the Nash network is a local maximizer of the potential function over the set of networks that differ from the current network for at most one link.

Theorem 1 (Uniqueness and Characterisation of Stationary Equilibrium).

The network formation game, under Assumptions 1–3, converges to a unique stationary distribution

$$\pi(g, X|\theta) = \frac{\exp\{\mathcal{Q}(g, X|\theta)\}}{\sum_{\omega \in \mathcal{G}} \exp\{\mathcal{Q}(\omega, X|\theta)\}}$$
(1)

Comments

- existence and uniqueness come from irreducibility and aperiodicity of Markov chain
- uniqueness crucial for estimation (avoid multiple equilibria)
- closed form stationary $\pi(g, X|\theta)$ corresponds to the likelihood of observing a specific network configuration in the long run
 - \implies can estimate heta with only one network, assuming it is drawn from the stationary equilibrium
- $\pi(g, X|\theta)$ coincides with likelihood of ERGM (Exponential Random Graph Model), where probability observing a network is proportional to exponential of linear combination of network statistics

Corollary 1.1.

Let Assumptions 1–3 hold. If the <u>utility functions</u> are <u>linear</u> in parameters, the stationary distribution $\pi(g, X|\theta)$ describes an ERGM, with $\mathbf{t}(g, X)$ a vector of canonical statistics

$$\pi(g, X | \theta) = \frac{\exp\{\theta' \mathbf{t}(g, X)\}}{\sum_{\omega \in \mathcal{G}} \exp\{\theta' \mathbf{t}(\omega, X)\}}$$
(2)

Extensions

Utility functions

possible to include additional utility components, as long as possible to find restrictions on payoffs that guarantee the existence of a potential function

Undirected networks

possible to extend existence results, characterisation of equilibrium, relation with ERGM and asymptotic results to undirected networks

Sparsity

model with negative linking externalities is compatible with a certain degree of sparsity

Estimation and Identification

Likelihood function

$$L(g, X|\theta) = \pi(g, X|\theta) = \frac{\mathcal{Q}(g, X|\theta)}{\sum_{\omega \in \mathcal{G}} \mathcal{Q}(\omega, X|\theta)} = \frac{\mathcal{Q}(g, X|\theta)}{c(\mathcal{G}, X, \theta)}$$

whose normalizing constant $c(\mathcal{G}, X, \theta)$ is intractable since it sums $2^{n(n-1)}$ terms.

- X standard ML infeasible
- MCMC with standard MH step infeasible (ratio of normalizing constants)

Estimation Algorithm

ERGM literature \Rightarrow approximate $c(\mathcal{G}, X, \theta)$ via MCMC (for fixed θ_0)

Algorithm 1 Metropolis-Hastings for Network Simulations

- 1: **procedure** MH_NETSIM(θ_0, g_0, R)
- 2: **for** r = 1, ..., R **do**
- 3: 1) propose network $g' \sim q_g(g'|g^{(r)})$
- 4: 2) accept network g' with probability

$$\alpha(\mathbf{g}^{(r)}, \mathbf{g}') = \min \left\{ 1, \frac{\exp\{\mathcal{Q}(\mathbf{g}', X|\boldsymbol{\theta}_0)\}}{\exp\{\mathcal{Q}(\mathbf{g}^{(r)}, X|\boldsymbol{\theta}_0)\}} \frac{q_{\mathbf{g}}(\mathbf{g}^{(r)}|\mathbf{g}')}{q_{\mathbf{g}}(\mathbf{g}'|\mathbf{g}^{(r)})} \right\}$$

- 5: end for
- 6: **return** sequence of *R* networks $\{g^{(r)}\}_r$
- 7: end procedure
- ✓ not requires $c(G, X, \theta)$
- X slow convergence
- **X** local sampler at each iteration, update link g_{ii} according to $\alpha(\cdot, \cdot)$
- ✗ degeneracy problem: large probability mass on few networks

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- 5: end for
- 6: **return** sequence of *R* networks $\{g^{(r)}\}_r$
- 7: end procedure

▶ how to choose $q_g(\cdot|g^{(r)})$?

Classes of asymptotics for networks

- 1) many networks ⇒ same players, growing number of networks
- 2) *large networks* ⇒ growing number of players, same network
- ▶ Hp: homogeneous players (i.e. $X_i = X_j$, $\forall i, j$)
- \blacktriangleright potential function re-scaled by $n^{\nu(H)}$, with $\nu(H)$ # players in each utility term
- ▶ example **re-scaled likelihood**, with $\mathcal{T}(g) = \alpha t(H_1, g) + \beta t(H_2, g)$ re-scaled potential and $\psi_n = n^{-2} \log(c(\alpha, \beta, \mathcal{G}_n))$ log-normalising constant

$$\pi_{n}(g|\alpha,\beta) = \frac{\exp\left\{n^{2}\left[\alpha \frac{\sum_{i=1}^{n} \sum_{j=1}^{n} g_{ij}}{n^{2}} + \beta \frac{\sum_{i=1}^{n} \sum_{j=1}^{n} \sum_{k\neq i}^{n} g_{ij}g_{jk}}{n^{3}}\right]\right\}}{c(\alpha,\beta,\mathcal{G}_{n})}$$

$$= \exp\left\{n^{2}\left[\mathcal{T}(g) - \psi_{n}\right]\right\}$$
(3)

Theorem 2 (Nonnegative Link Externalities).

Model (3) with nonnegative link externalities $\beta \geq 0$ exhibits the following behaviour

1) asymptotic normalizing constant ψ solves

$$\psi = \lim_{n \to \infty} \psi_n = \max_{\mu \in [0,1]} \left\{ \alpha \mu + \beta \mu^2 - \mu \log(\mu) - (1-\mu) \log(1-\mu) \right\}$$
(4)

- 2) networks generated by the model are indistinguishable from directed Erdös–Rényi graph with linking probability μ^* , defined as follows:
 - (a) if the maximization (4) has a unique solution, then μ^* satisfies $2\beta\mu(1-mu)<1$ for almost all $\alpha\in\mathbb{R}$ and $\beta\geq0$, and solves

$$\mu = \frac{\exp\left\{\alpha + 2\beta\mu\right\}}{1 + \exp\left\{\alpha + 2\beta\mu\right\}} \tag{5}$$

(b) if the maximization (4) has two solutions, then μ^* picked randomly from same proba distribution over μ_1^* and μ_2^* , such that $\mu_1^* < 0.5 < \mu_2^*$, and both solve (5) and satisfy $2\beta\mu(1-\mu) < 1$.

Comments on Theorem 2

- consistent estimator of log-normalising constant analogue of variational representation of the discrete exponential family
- ▶ $\beta \ge 0$ ⇒ realisations using (α, β) indistinguishable from those using $(\alpha', 0) = (\log(\mu^*/(1 \mu^*)), 0)$, that is from Erdös-Rényi model

Corollary 2.1.

When $\beta \geq 0$, the externality cannot be identified.

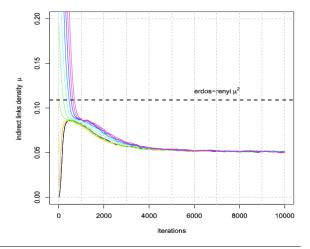
Corollary 2.2.

When $\beta \geq 0$, Algorithm 1 is not necessary since Erdös-Rényi graphs can be simulated using Bernoulli draws.

Theorem 3 (Negative Link Externalities).

If $\beta < 0$ and sufficiently large in magnitude, model (3) is asymptotically different from a directed Erdös-Rényi model.

✓ sparser graphs than Erdös-Rényi



Theorem 3 (Negative Link Externalities).

If $\beta < 0$ and sufficiently large in magnitude, model (3) is asymptotically different from a directed Erdös-Rényi model.

- ▶ how much "sufficiently large" magnitude?
- ▶ how to know it, since we must estimate β ?

Consider an additional utility component (cyclic triangles):

$$T(g) = \alpha t(H_1, g) + \beta t(H_2, g) + \gamma t(H_3, g), \qquad t(H_3, g) = \sum_{i=1}^n \sum_{j=1}^n \sum_{k \neq i}^n g_{ij} g_{jk} g_{ki} \quad (6)$$

Theorem 4.

Consider model (6) as $n \to \infty$

1) If $\beta \geq 0$ and $\gamma \geq 0$, then the asymptotic normalising constant ψ solves

$$\psi = \lim_{n \to \infty} \psi_n = \max_{\mu \in [0,1]} \left\{ \alpha \mu + \beta \mu^2 + \gamma \mu^3 - \mu \log(\mu) - (1-\mu) \log(1-\mu) \right\}$$
 (7)

and model is asymptotically indistinguishable from directed Erdös-Rényi graph, with μ^* maximising (7). If the maximisation problem has multiple solutions, then μ^* picked randomly from some distribution on maximisers.

2) If at least **one** externality is negative (i.e. $\beta < 0$ or $\gamma < 0$) and sufficiently large, then model (6) not converge asymptotically to directed Erdös-Rényi graph and externalities can be identified.

Summary of Asymptotics

Remark 2.

Homogeneous players $(X_i = X_i, \forall i, j)$:

- (a) positive externalities
 - asymptotically indistinguishable from Erdös-Rényi graph
 - externalities not identified
 - can approximate likelihood of model via likelihood of Erdös-Rényi graph
- (b) at least one externality negative and large
 - asymptotically sparser than Erdös-Rényi graph
 - externalities identified

Heterogeneous players $(\exists i, j \text{ such that } X_i \neq X_i)$:

- no results
- preliminary study in Mele & Zhu (2017) working paper

Sampler Convergence

Theorem 5 (Convergence of Local Sampler with Nonnegative Externalities).

Model (6), with probability of meeting $\rho_{ij}=1/(n(n-1))$. Fix $\gamma \geq 0$. Then, in the case of nonnegative externalities $\beta \geq 0$, there exists a V-shaped region of the parameter space delimited by functions $S_{\gamma}(\phi_1(\alpha)), S_{\gamma}(\phi_2(\alpha))$ such that

- 1) if (α, β) belongs to the V-shaped region, then model converges to stationarity in e^{Cn^2} steps, C > 0. This results holds for any local sampler.
- 2) otherwise, model converges in $Cn^2 \log(n)$ steps, C > 0.

Intuition:

- (1a) in the V-shaped region problem (7) has 2 *local maxima*, the sampler spend exponential time at one of them (i.e. probability e^{-Cn^2} to escape from local max)
- (1b) increasing $\gamma \implies$ increase area of exponentially slow convergence
- (2a) when convergence is quadratic \implies sampler feasible for n < 500
- (2b) this happens when model is indistinguishable from directed Erdös-Rényi graph

Simulation and Estimation in Finite Networks

Posterior inference via approx version of exchange algorithm of MGM06 [13]

- ▶ double Metropolis-Hastings step to avoid computing $c(\mathcal{G}, X, \theta)$
- ▶ data augmentation via auxiliary network g'
- \blacktriangleright higher $R \implies$ better approximation of posterior, but higher rejection rate

Algorithm 2 Approximate Exchange Algorithm

```
1: procedure AEA(\theta, g, M, R)
```

- 2: **for** m = 1, ..., M **do**
- 3: 1) propose parameter $heta' \sim q_{ heta}(\cdot|m{ heta})$
- 4: 2) run Algorithm 1 for R iterations using θ' . Keep last simulated network g'
- 5: 3) accept parameter θ' with probability

$$\alpha(\boldsymbol{\theta}, \boldsymbol{\theta}', \boldsymbol{g}', \boldsymbol{g}) = \min \left\{ 1, \frac{\exp\{\mathcal{Q}(\boldsymbol{g}', \boldsymbol{X}|\boldsymbol{\theta})\}}{\exp\{\mathcal{Q}(\boldsymbol{g}, \boldsymbol{X}|\boldsymbol{\theta})\}} \frac{p(\boldsymbol{\theta}')}{p(\boldsymbol{\theta})} \frac{q_{\boldsymbol{\theta}}(\boldsymbol{\theta}|\boldsymbol{\theta}')}{q_{\boldsymbol{\theta}}(\boldsymbol{\theta}'|\boldsymbol{\theta})} \frac{\exp\left\{\mathcal{Q}(\boldsymbol{g}, \boldsymbol{X}|\boldsymbol{\theta}')\right\}}{\exp\left\{\mathcal{Q}(\boldsymbol{g}', \boldsymbol{X}|\boldsymbol{\theta}')\right\}} \right\}$$

- 6: end for
- 7: **return** sequence of M parameters $\{\boldsymbol{\theta}^{(m)}\}_m$
- 8: end procedure

Simulation and Estimation in Finite Networks

Algorithm 2 Approximate Exchange Algorithm

- 1: procedure $AEA(\theta, g, M, R)$
- 2: **for** m = 1, ..., M **do**
- 3: 1) propose parameter $\theta' \sim q_{\theta}(\cdot|\theta)$
- 4: 2) run Algorithm 1 for R iterations using θ' . Keep last simulated network g'
- 5: 3) accept parameter θ' with probability

$$\alpha(\boldsymbol{\theta}, \boldsymbol{\theta}', g', g) = \min \left\{ 1, \frac{\exp\{\mathcal{Q}(g', X|\boldsymbol{\theta})\}}{\exp\{\mathcal{Q}(g, X|\boldsymbol{\theta})\}} \frac{p(\boldsymbol{\theta}')}{p(\boldsymbol{\theta})} \frac{q_{\boldsymbol{\theta}}(\boldsymbol{\theta}|\boldsymbol{\theta}')}{q_{\boldsymbol{\theta}}(\boldsymbol{\theta}'|\boldsymbol{\theta})} \frac{\exp\left\{\mathcal{Q}(g, X|\boldsymbol{\theta}')\right\}}{\exp\left\{\mathcal{Q}(g', X|\boldsymbol{\theta}')\right\}} \right\}$$

- 6: end for
- 7: **return** sequence of M parameters $\{\theta^{(m)}\}_m$
- 8: end procedure

- ▶ what prior distribution $p(\theta)$?
- ▶ what proposal distribution $q_{\theta}(\cdot|\theta)$?

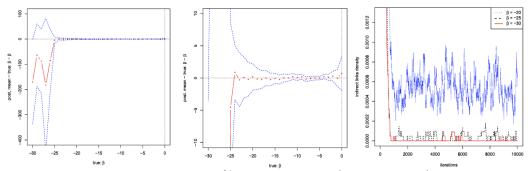


Figure: Left: Estimates of β < 0, with 95% credibility intervals (middle: zoom-in). Right: indirect links density.

- $\beta \ge 0 \implies$ Erdös-Rényi case, not identified
- $\beta < 0 \implies$ identified
- $\beta \ll 0 \implies$ estimation impossible: # indirect links close to 0

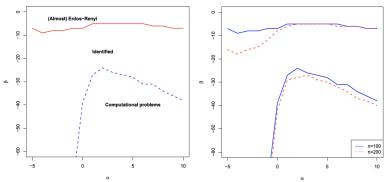


Figure: Left: approximate regions of identified parameters, for n = 100.

Right: comparison of regions for n = 100, n = 200.

Remark 3.

Regions of identified parameters (α, β) vary with n, the number of players.

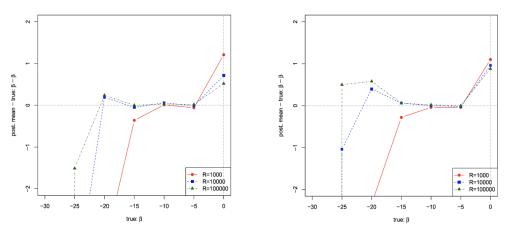


Figure: Difference between posterior estimates and true, for varying number of network simulations R: n = 100 (left) and n = 200 (right).

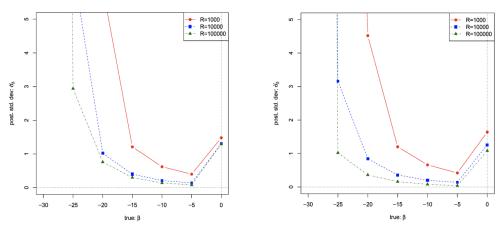


Figure: Posterior standard deviation for varying number of network simulations R: n = 100 (left) and n = 200 (right).

- $ightharpoonup R = 1000 \implies \text{imprecise estimates}$
- ▶ no significant difference between R = 10,000 and R = 100,000 \implies suggest rule-of-thumb R = 10,000
- ightharpoonup cost of increasing network simulations \implies almost linear $\mathcal{O}(R)$
- results suggest convergence is almost quadratic $\mathcal{O}(n^2)$ in this area of parameter space

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- results suggest convergence is almost quadratic $\mathcal{O}(n^2)$ in this area of parameter space

▶ what was the computing time?

▶ what about real data applications?

Conclusions

The paper in a nutshell:

- ❖ network formation model, combining strategic and random networks features
- payoffs depend on links: direct + indirect (externalities)
- ♦ homogeneous players meet sequentially at random, myopically updating links
- network formation process is a potential game and converges to ERGM, generating directed dense networks
- ❖ identification: only if at least 1 externality negative and sufficiently large
- ❖ standard estimation for ERGMs exponentially slow ⇒ Bayesian MCMC (almost quadratic time)

Conclusions

Unclear points and questions:

- theoretical quantification of "sufficiently large" (negative) magnitude of β ?
- choice of prior for parameters $p(\theta)$?
- choice of proposal for network $q_g(\cdot|g)$?
- choice of proposal for parameters $q_{\theta}(\cdot|\theta)$?
- duration of computing time in simulations?
- real data applications?

Thanks for your attention!

Any question?

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