

# Heterogeneous Intermediaries and the Transmission(s) of Monetary Policy <sup>\*</sup>

Federic Holm-Hadulla  
European Central Bank

Matteo Leombroni  
Stanford University

November 24, 2022

*Click [here](#) for latest draft*

## Abstract

This paper studies the transmission of monetary policy to the corporate bond market. We show that corporate bond purchases by the central bank mainly give rise to credit spread shocks, whereas (safe) government bond purchases mainly cause term spread shocks. We also document that yields of bonds held by different intermediaries respond heterogeneously to the two shocks. This heterogeneity in responses reflects bond characteristics rather than the investor base composition for a given bond, implying that intermediary sectors systematically select into different types of assets. We explain these findings through the lens of a model of the fixed-income market with multiple risk factors. Levered intermediaries, such as insurance corporations and pension funds, select into assets with a high interest rate risk exposure to match their long-duration liabilities. For mutual funds, this liability-matching motive is absent, and they instead select into securities carrying credit risk. Different policy tools heterogeneously affect the market prices of those factors, thereby redistributing risks across intermediary sectors and ultimately across the households investing in them.

---

<sup>\*</sup>Matteo is grateful to Monika Piazzesi, Martin Schneider, Hanno Lustig and Luigi Bocola for invaluable guidance. Matteo also thanks Juliane Begenau, Matteo Benetton, Marianna Kudlyak and Cedomir Malgieri for their priceless support. We thank Adrien Auclert, Antonio Coppola, Chris Tonetti, Sebastian Di Tella as well as all participants of the MM reading group for helpful discussions, comments, and questions. The views expressed here are those of the authors and do not necessarily reflect those of the European Central Bank.

# 1. Introduction

Over the past two decades, central banks around the world have engaged in an unprecedented expansion of their monetary policy toolkits to lean against disinflationary pressures and address impairments in monetary policy transmission. A key element of this tool expansion consists of introducing central bank asset purchase programs, covering not only sovereign securities typically characterized by low credit risk but also assets issued by riskier private sector entities, e.g., in the form of corporate bonds.

A common view on the impact of quantitative easing (QE) supports the idea that QE has narrow effects (Krishnamurthy (2022)). This research argues that the impact of QE shocks is highest in the asset market targeted, with limited spillover to other assets. At the same time, there is a long tradition of thinking about fixed-income markets in terms of a small number of risk factors. Factor models of bond prices or returns go back to Litterman and Scheinkman (1991) for government bonds and Fama and French (1993) for corporate bonds. Two factors—interest rate risk and credit risk—explain a large share of the variation in fixed-income markets (Begenau et al. (2015)).

Against this backdrop, this paper provides novel empirical evidence on factor models’ relevance to understanding different monetary policy tools and their transmissions to asset prices. In our factor approach, QE transmission works through two channels: an interest rate channel (empirically measured via term-spreads) and a credit risk channel (measured through credit spreads). By trading different securities, the central bank can affect the market price of risk factors (e.g., Papoutsis et al. (2021)).<sup>1</sup>

Central banks’ interventions in financial markets further highlighted the relevance of financial intermediaries (He and Krishnamurthy (2013)). A flourishing literature, originated by Koijen and Yogo (2019), studies characteristics-based portfolio *selection* by heterogeneous intermediaries and the importance of diverse institutional demand shocks for equilibrium prices. Recently, Coppola (2021) emphasized this second channel for corporate bond price dynamics in response to credit shocks. By contrast, we (i) highlight the importance of bond characteristics to explain the cross-sectional response of corporate bonds to QE shocks and (ii) emphasize a selection mechanism that explains the empirical correlation we uncover between price reactions and intermediaries’ holdings.

Because of their characteristics, corporate bonds are heterogeneously exposed to the two risk factors (i.e., different  $\beta$  loadings on interest rate risk and credit risk) and hence to different monetary policy shocks. High credit  $\beta$  bonds are naturally more exposed to the QE credit risk channel, while bonds with high interest rate  $\beta$  exhibit higher sensitivity to the QE interest rate

---

<sup>1</sup>While in our paper we focus on the interest rate risk and credit risk factors, other factors are however important. Papoutsis et al. (2021) suggest that central banks can affect carbon emissions by trading a climate risk factor since the price of that factor moves polluting firms’ cost of capital. In the Euro area a sovereign risk factor is also relevant.

channel.

In turn, due to their institutional settings, heterogeneous intermediaries choose bond portfolios with different exposure to the risk factors. Levered intermediaries, such as insurance companies and pension funds (ICPFs), select into assets with a high interest rate risk exposure to match their long-duration liabilities. For mutual funds, this liability-matching motive is absent, and they instead select into securities carrying credit risk. This selection mechanism primarily explains the association between price response to QE and bond holdings.

The Euro area provides an interesting case for understanding monetary policy transmission of different instruments. Like many other central banks, the ECB has included various asset classes in the purchase programs adopted to counter disinflationary pressures since mid-2014. Besides securities issued by the government sector, these have also comprised corporate bonds, with the latter accounting for the second most significant component of the purchase envelope since their introduction in 2016. Using micro data corporate bond prices and ECB confidential dataset of security-level portfolio holdings by investor type (Kojien et al. (2021)), we study the transmission of ECB monetary policy to corporate bonds.

Based on our two-factor approach, we construct two monetary policy shocks: the first related to interest rate risk and the second related to credit risk. The term spread monetary policy shock is measured using the high-frequency change in the difference between 10-year and 2-year OIS rates in a narrow window around ECB monetary policy announcements. The credit spread monetary policy shock is measured as the change in the credit spread of Euro area corporate bond yields around ECB announcement dates.

In a second step, we show that the prevalence of these two shocks depends on the design of central bank asset purchase programs. Government bond purchases primarily give rise to term spread shocks (Altavilla et al. (2019), Eser et al. (2019), Swanson (2021)), whereas ECB announcements related to corporate sector purchases primarily give rise to credit spread shocks.

Credit spreads are an essential determinant of economic activity (Gilchrist and Zakrajšek (2012)) and a crucial channel for monetary policy transmission (Gertler and Karadi (2011)). Central banks can affect credit spreads either through their interest rate policies (Gertler and Karadi (2015), Gilchrist et al. (2015)) or directly through a new set of different policy interventions (e.g., corporate sector purchases). This second channel motivates our new credit spread monetary policy shocks. We show that while ECB corporate bond purchases substantially affect credit spreads, interest rate policies' effects on credit spreads are limited. Taken together, this evidence demonstrates that the term and credit spread shocks capture distinct mechanisms by which different central bank purchase programs transmit to bond markets.

Using these shocks, we can then estimate the cross-sectional effects of monetary policy on the corporate bond market. To this end, we group a large number of corporate bonds according to key characteristics, including maturity, rating, and sector. For each group, we estimate the sensitivity of the variation in their credit spread around ECB announcements to our credit spread monetary

policy shock. Moreover, we lever the granular information of the ECB Securities Holdings Statistics (SHS) database to compute, for the bonds in each group, their holding structure across financial intermediaries sectors.

This sorting, in turn, exposes a strong correlation between the holdings of different intermediaries and the sensitivity coefficients. Bonds held by mutual funds exhibit higher sensitivity to the credit spread shock, whereas bonds held by insurance companies and pension funds are less sensitive to this shock. Moreover, conducting an analogous exercise to estimate the sensitivity of bonds to our term spread shock, we obtain the opposite result: insurance companies and pension funds hold more of the more responsive bonds, whereas mutual funds hold more of the less responsive bonds.

This finding sheds new light on a recent literature studying the interaction between bond-holding structures among financial intermediary sectors and their responses to shocks. For instance, [Coppola \(2021\)](#) and [Faia et al. \(2022\)](#) highlight the importance of the investor base composition for the pricing of corporate bonds and their response to shocks. By contrast, we emphasize how the sensitivity to credit spread shocks primarily reflects bond characteristics rather than the investor base composition for a given bond. Vice versa, this implies that different euro area intermediary sectors systematically select into different types of financial assets.

To rationalize our main findings, we develop a dynamic risk-factor asset pricing model with heterogeneous intermediaries.

Our modeling framework for different intermediaries is motivated by the critical differences in their liability structure. Insurance companies' debt-type liabilities mainly consist of insurance policies, which are discounted by the risk-free curve (the €STR -OIS curve for the Euro area). These policies tend to have a long duration, and their valuation is therefore highly exposed to interest rate risk ([Domanski et al. \(2017\)](#)). Risk-management concerns, as well as regulatory rules, thus induce insurance companies to use their portfolio allocations on the asset side to hedge their interest rate exposure on the liability side of their balance sheet. Similar considerations apply to pension funds which, in the Euro area, are primarily defined benefits and hence characterized by long-duration liabilities. By contrast, mutual funds have limited debt-type liabilities and are mostly (or entirely) funded with mutual fund shares.

The model's set-up is an exchange economy with two risky claims, each exposed to one of the two risk factors (i.e., interest rate risk and credit risk). There are two types of households: the intermediary shareholder and the passive household. The intermediary shareholder decides how to allocate its wealth across intermediaries ([Kojen and Yogo \(2022\)](#)), an asset manager and a levered intermediary. The asset manager is not able to take on leverage, whereas the levered intermediary can borrow from passive households while being subject to a collateral constraint and a value-at-risk type of regulation cost. A vital feature of the model is that the debt of the levered intermediaries is also exposed to interest rate risk. Both intermediaries maximize shareholder value and allocate their portfolios across an interest rate risk factor claim and a credit risk factor

claim. The passive household cannot invest in the equity of intermediaries or the two risky claims.

The model provides a set of predictions regarding the allocation of risks, asset pricing, and the effects of central bank purchases. Levered intermediaries choose to allocate more their portfolio to the interest rate risk factor because: (i) the interest risk factor hedges the corresponding exposure of their debt, and (ii) the interest rate risk factor has superior collateral value compared to the credit risk factor. By contrast, mutual funds allocate more of their portfolio to credit risk as they do not face any regulatory cost in doing so.

Consequently, the intermediary shareholders consumption is exposed to credit risk mainly through the balance sheet of mutual funds and to interest rate risk through the balance sheet of levered intermediaries. Part of the interest rate risk, however, is offloaded to the passive investor through the debt of levered intermediaries: it is the net exposure to risk that determines the risk premia commanded on the two assets.

The central bank purchases risky securities by issuing a risk-free liability. The balance sheet of the central bank, therefore, reallocates risk in the economy. By redistributing the risk exposure from the intermediary shareholders to the passive households—who is not pricing the two risky claims—the central bank, in turn, can affect premia (Silva (2016)). Within our two-factors framework, the composition of asset purchases by the central bank heterogeneously affects premia on different factors (Papoutsis et al. (2021)).

We then provide further empirical evidence to corroborate the predictions of our model. We further sort corporate bonds based on their credit spread and maturity. We use the ECB granular holding database to show that levered intermediaries mainly sort into high interest rate risk and low credit risk assets while mutual funds sort into credit risky assets. Conditioning on credit ratings, we still observe a similar sorting mechanism (i.e., mutual funds hold more with higher credit spread).

Consistently with our factor approach, bonds with higher credit spreads (i.e., higher  $\beta$  loadings on credit risk) are more sensitive to credit spread monetary policy shocks. We also provide further evidence on the effects of two specific announcements: the corporate sector purchase programme announcement (CSPP) (Todorov (2020), De Santis and Zaghini (2021)) and the pandemic emergency purchase programme (PEPP) announcement. Sub-investment grade corporate bonds, which were ineligible to be purchased by the ECB, exhibited strong price responses to both announcements. This evidence contrasts with previous findings on the effects of Federal Reserve’s announcement of corporate bond purchases (Haddad et al. (2021)).

Our results have important implications for the transmission of monetary policy, financial stability, and redistribution.

First, the results suggest that the asset prices effects of monetary policy primarily depend on the type of monetary policy shock (i.e., the type of risk factor absorbed: credit risk or interest rate risk) and the exposure of the bond to the different risk factors—based on its characteristics. Once we account for the characteristic of the bond, the investor base composition of the single

bond has minor effects on the transmission of shocks to asset prices.

Second, different monetary policy shocks have significantly different effects on the balance sheet of different intermediaries. A quantitative easing or tightening operated via corporate bonds significantly affects the balance sheet of mutual funds and the funding costs of credit riskier non-financial corporations. A quantitative easing or tightening operated via buying or selling (safe) government bonds mainly affects the balance sheet of insurance companies and pension funds and the borrowing costs of safer non-financial corporations.

Third, different central banks purchases have different redistributive effects across households. The entire household sector ultimately bears the risks on the balance sheet of the central bank. A quantitative easing or tightening redistribute risks from intermediaries shareholders to the household sectors as a whole.

## 1.1 Literature Review

By providing empirical evidence and new theoretical insights on the transmission mechanism of different central bank policies, we contribute to the vast literature on large-scale asset purchases and the role of intermediaries.

From an empirical perspective, [Krishnamurthy and Vissing-Jorgensen \(2011\)](#); [Gagnon et al. \(2011\)](#) take an event study approach to study the effects of QE on asset prices. Several papers also investigate the transmission of QE on asset prices in the Euro area, including [Krishnamurthy et al. \(2017\)](#); [Altavilla et al. \(2015\)](#).

Our paper also focuses on the effects of ECB corporate bond purchases. We hence relate to a relevant set of papers that investigate the impact of the announcement of ECB CSPP (Corporate Sector Purchase Programme) on asset pricing ([Abidi and Miquel-Flores \(2018\)](#); [De Santis et al. \(2018\)](#); [Grosse-Rueschkamp et al. \(2019\)](#); [Zaghini \(2020\)](#); [Todorov \(2020\)](#); [De Santis and Zaghini \(2021\)](#)). In our paper, we take a systematic approach by constructing two monetary policy shocks, one affecting credit spreads and the other affecting term spreads. We then show that corporate bond purchases by the central bank mainly give rise to credit spread shocks. We also show that (safe) government bond purchases mainly cause term spread shocks ([Eser et al. \(2019\)](#)), by relating our shocks to the previous work of [Altavilla et al. \(2019\)](#) and [Swanson \(2021\)](#).<sup>2</sup>

We show that our credit spread monetary policy shocks reflect ECB announcements related to corporate bond purchases. [Nakamura and Steinsson \(2018\)](#); [Jarociński and Karadi \(2020\)](#), discuss how information effects may also affect the price of risky assets. [Jarociński \(2021\)](#) disentangle Fed’s unconventional policy shocks from information shocks.

We then study the effects of the two shocks to the cross-section of corporate bond yields.

---

<sup>2</sup>There is an extensive literature devoted to estimating monetary policy shocks using high-frequency changes in prices around monetary policy announcements, including ([Cochrane and Piazzesi \(2002\)](#); [Gürkaynak et al. \(2004\)](#); [Bernanke and Kuttner \(2005\)](#)). [Leombroni et al. \(2021\)](#) discusses the effects of different monetary policy shocks on the Euro area government bond market



Many papers have studied the cross-sectional effects of monetary policy (measured through the change in interest rate) on corporate bonds, including [Gilchrist et al. \(2015\)](#); [Gertler and Karadi \(2015\)](#); [Smolyansky and Suarez \(2021\)](#); [Anderson and Cesa-Bianchi \(2020\)](#). We extend this earlier work by (i) studying new monetary policy shocks and (ii) relating the cross-sectional effects to intermediary holdings. We show that yield on bonds held by different intermediaries responds heterogeneously to the two shocks.

However, this heterogeneity in responses primarily reflects bond characteristics rather than the investor base composition for a given bond. We emphasize the factor structure in fixed-income assets ([Fama and French \(1993\)](#)) and discuss how different  $\beta$  loadings on different risk factors explain the cross-sectional response of corporate bonds. [Kelly et al. \(2020\)](#), [Bai et al. \(2019\)](#) propose different factor models for corporate bond returns.

This result extends the work of [Coppola \(2021\)](#) to different types of shock, i.e., monetary policy shocks. [Faia et al. \(2022\)](#) also stresses how institutional investors' mandates determine bond demand, which, through granularity, affects bond prices.

In our work we then emphasized how heterogeneous intermediaries systematically select into different types of assets. [Koijen and Yogo \(2022\)](#) discusses the equilibrium allocation of insurance companies in a model with different intermediaries where insurance companies have relatively cheap access to leverage through their underwriting activity. Other papers have instead emphasized the importance of duration matching for the insurance sector (for instance, [Domanski et al. \(2017\)](#); [Ozdagli and Wang \(2022\)](#); [Carboni and Ellison \(2022\)](#)).

We extend the setting of [Koijen and Yogo \(2022\)](#) to a dynamic equilibrium asset pricing model with heterogeneous intermediaries where the debt of insurance companies and pension funds is risky. Intermediaries allocate across two risk factors: interest rate risk and credit risk (these risk factors have also been used to document banks' risk exposure by [Begenau et al. \(2015\)](#)). We hence emphasize the portfolio choice across different risk factors, the importance of liabilities hedging for asset allocation, and the equilibrium risk premia commanded on different risks.<sup>3</sup> Finally, we also use the framework to study the effects of central banks purchases.

A growing literature has emphasized the role of investor heterogeneity in asset pricing. [Koijen and Yogo \(2019\)](#) proposes a model with flexible heterogeneity in asset demand across investors, and [Bretscher et al. \(2021\)](#) applies this methodology to corporate bonds. [Koijen et al. \(2021\)](#) uses this approach to evaluate the impact of ECB QE on asset pricing using the ECB SHS database. We take a different approach by studying the systematic effects of different monetary policy shocks on the cross-section of corporate bonds around ECB announcements and then discuss how bonds held by different intermediaries are exposed to the different shocks and how different policy instruments have different cross-sectional implications.

These empirical findings motivate a modeling framework where central banks affect the market

---

<sup>3</sup>Our model shows that the size of the insurance sector affects the interest rate risk premium, consistently with the work of [Greenwood and Vissing-Jorgensen \(2018\)](#)

price of different risk factors; the cross-sectional effects are then pinned down by the different exposure of corporate bonds to the two risk factors (based on bond characteristics). [Papoutsis et al. \(2021\)](#) propose a model of unconventional monetary policy where central bank purchases change factor prices because they lower private intermediaries' marginal cost of holding risky bonds.

We contribute by providing a theoretical mechanism—based on the risk redistribution of different risk factors—that generates the central bank effects we observed in the data and by studying the exposure to risk factors of heterogeneous intermediaries. The risk redistribution channel (based on a single risk factor) was introduced by [Silva \(2016\)](#).

Our theoretical approach contributes to the extensive literature on quantitative easing ([Gertler and Karadi \(2011\)](#); [He and Krishnamurthy \(2013\)](#); [Brunnermeier and Sannikov \(2014\)](#); [Piazzesi and Schneider \(2021\)](#); [Vayanos and Vila \(2021\)](#)) as well other studies on the role of conventional monetary policy for risk premia ([Kekre and Lenel \(2021\)](#)).

## 2. The Euro Area Corporate Bond Market

The Euro area corporate bond market has been growing over the past two decades, and it is increasing its role in the debt financing of non-financial corporations.<sup>4</sup> Although bank loans have been historically the primary source of debt financing for firms in the euro area, the corporate bond market is gaining share. In 2021, corporate bonds accounted for 15% of non-financial corporations' long-term debt financing.<sup>5</sup> Figure 1a shows the rising share of market financing for non-financial corporations. Bond financing is even more relevant during recessions (the shaded area in Figure 1a), when access to bank loans may be impaired.

The leading investors in corporate bonds issued by Euro area non-financial corporations are mutual funds, insurance companies and pension funds (ICPF), banks, and the ECB.<sup>6</sup> At the end of 2021, the total amount of bonds held by Euro area investors totaled €1.36Tn. The largest Euro area investors in corporate bonds were mutual funds (34.7%), followed by insurance companies and pension funds (27.2%). Of the total amount held by insurance companies and pension funds (ICPF), insurance corporations account for 89% while pension funds for the remaining 11%. The third largest investor is the ECB, which at the of 2021 held 18%. Finally, banks held 11%. The rest of the world holds an additional €0.35Tn. Figure 1b splits the amount of long-term corporate bonds issued by Euro area non-financial corporations and held by Euro area investors.<sup>7</sup> Appendix

---

<sup>4</sup>See [Darmouni and Papoutsis \(2022\)](#) for a discussion on the rise of bond financing in Europe

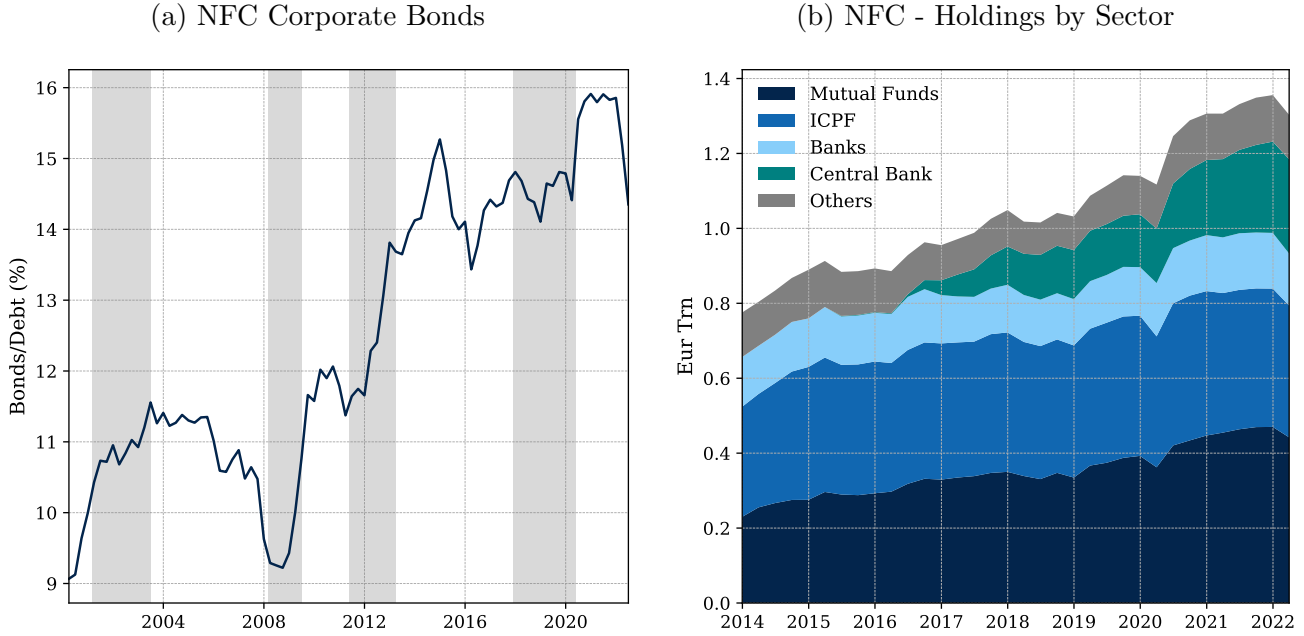
<sup>5</sup>We consider long-term, any loans or bonds with a maturity greater than one year, in accordance with the European System of Account (ESA) 2010 definition.

<sup>6</sup>Note that throughout the paper, we use the term ECB to represent the whole Eurosystem. Large parts of the central bank asset portfolio in the Euro area are held by the National Central Banks that form the Eurosystem along with the ECB.

<sup>7</sup>Note that this was not considered in the calculation of the shares.



Figure 1: Euro area non-financial corporate bonds



*Note:* Figure 1a shows the total amount of bond issued by Euro area non-financial corporations as a share of total non-financial corporations debt. Figure 1b splits the total amount of bond issued by Euro area non-financial corporations by holding sector: mutual funds, insurance corporations and pension funds (ICPF), banks, central bank and others. Source: Quarterly Sector Account.

A provides further information on the corporate bond market and the regulation of different intermediaries.

## 2.1 Data

In order to study the effects of monetary policy shocks on corporate bonds and on intermediaries balance sheets, we combine a number of different data sets.

### 2.1.1 Holdings

We compute the shares of bonds held by different intermediaries by leveraging granular information from the ECB Securities Holding Statistics (SHS). SHS is a security-level portfolio holdings database of all euro-area investors. The securities are uniquely identified by an International Securities Identification Number. The security types include government and corporate bonds, equities, and mutual fund shares. Data are reported at a quarterly frequency, and our sample covers 2013Q4 to 2021Q4. Securities Holding Statistics reports portfolio holdings by country of domicile and investor sector. We aggregate data for all Euro area countries and only split by investor sector. We use data for monetary financial institutions (which we refer to as banks), insurance companies and pension funds (ICPF), and mutual funds.

### 2.1.2 Asset Prices

We collect data from Markit iBoxx. The dataset provides detailed information on the universe of bonds used in their index. We only include bonds denominated in euros. The data are available at a daily frequency and include bond bid price, ask price, accrued interest, yield to maturity, duration, and ratings. We also construct a set of additional measures. We compute a measure of illiquidity defined as the autocorrelation in returns in a given quarter, following the methodology of [Bai et al. \(2019\)](#). We compute the corresponding credit risk-free interest rate for each bond using data on swap curves. From 2002 to 2006, we used Euribor with a maturity of up to 1 year and Swap rates written on 6m Euribor for longer maturities. We use OIS EONIA rates from 2007 to 2021. Finally, from 2021 onward, we use the OIS-€STR rates. We use the 1m to 12m maturity, the 1y to 10y maturity, 12y, 15y, 20y, 25y, and 30y. We interpolate swap yield using cubic spline.

### 2.1.3 Other data sources

We use data from the Euro Area Monetary Policy Event-Study Database (EA-MPD), developed by [Altavilla et al. \(2019\)](#), to measure the high-frequency intraday movement in OIS rates around ECB announcements.

## 3. Monetary Policy Shocks

The rapid growth in bond financing in the Euro area increases the relevance of the corporate bond market for the transmission of monetary policy. In order to systematically investigate monetary policy transmission to the cross-section of corporate bonds, we construct two monetary policy shocks series: i) a new credit spread monetary policy shock and ii) a term spread monetary policy shock. We measure our monetary policy shocks in a short window around all ECB schedule Governing Council announcements.

The term spread monetary policy shock is the change in term spread (defined as the 10-year OIS minus the 2-year OIS) in the intraday window around the ECB announcements.

The credit spread monetary policy shock is the change in corporate bond credit spread in a two-day window around the ECB announcement: from the day before the announcement to the day after. It is standard practice to measure the effects of monetary policy on corporate bonds in a larger window to reduce the effects of stale pricing or low liquidity. We construct the credit spread component by decomposing the market corporate bond yield into an interest rate component and a credit spread component:

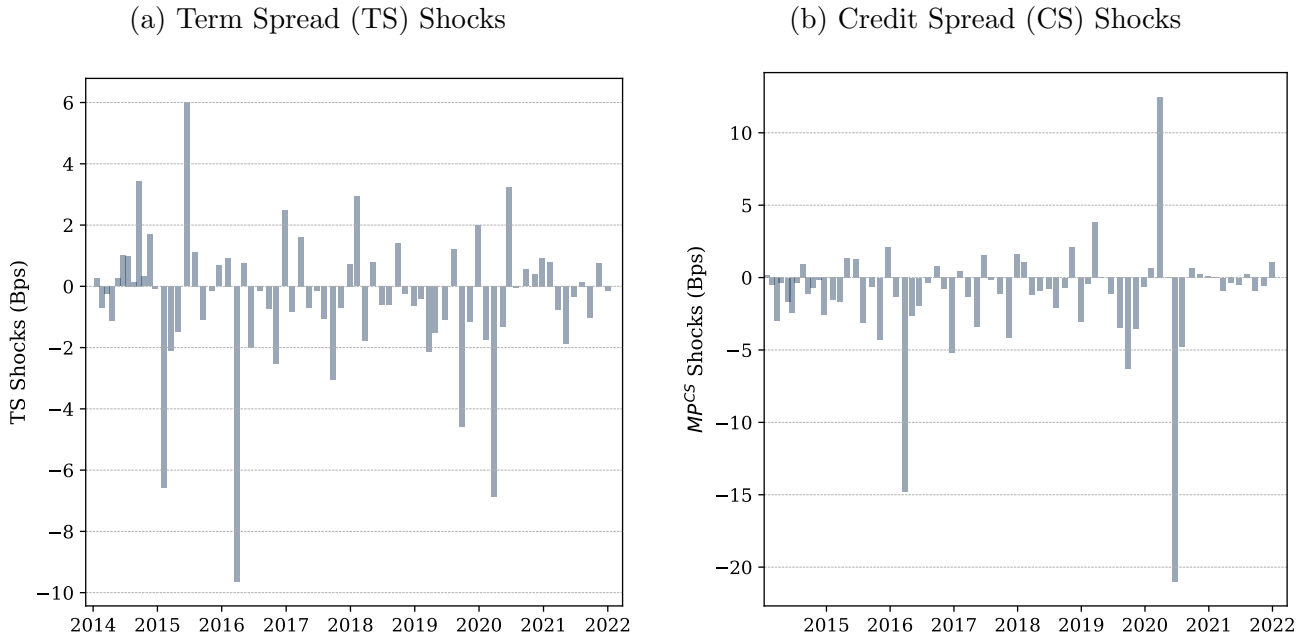
$$\underbrace{Y_t}_{\text{Market Corporate Yield}} = \sum_i w_{i,t} \times y_{i,t} = \underbrace{\sum_i w_{i,t} \times l_{i,t}}_{\text{Interest Rate Component}} + \underbrace{\sum_i w_{i,t} \times ys_{i,t}}_{\text{Credit Spread Component}} \quad (1)$$

where  $y_{i,t}$  is the yield-to-maturity of bond  $i$  at time  $t$ , and  $w_{i,t}$  is the notional outstanding amount of bond  $i$  as a share of the total outstanding amount of all bonds in the sample,  $\iota_{i,t}$  is the OIS swap yield with maturity equal to the maturity of bond  $i$  while  $ys_{i,t}$  is the spread of corporate bond  $i$ . The interest rate component (IR) measures the variation of credit yield due to variation in credit-risk-free rates: a combination of future expected short-term interest rate and term-premium. The second component, the credit spread component (CS), combines the compensation for default risk as well as the credit risk premium. We use the variation in the second component, the credit spread component, as our measure of credit spread monetary policy shocks:

$$MP_t^{CS} = \sum_i w_{i,t} \times \Delta ys_{i,t}, \quad (2)$$

where we define  $\Delta ys_{i,t} = (ys_{i,t+1} - ys_{i,t-1})$ , as the change around monetary policy announcement  $t$ .<sup>8</sup> Figure 2 plot the time series of the two shocks.

Figure 2: Monetary Policy Shocks



*Note:* The figure shows our two monetary policy shocks. Figure 2a plots the term-spread monetary policy shock, defined as the change in term spread (10-year OIS minus the 2-year OIS), in the intraday window around the ECB announcements. Figure 2b plots the credit spread monetary policy shocks, the change in Euro-denominated investment grade non-financial corporate bond credit spread in a two-day window around the ECB announcement (from the day before the announcement to the day after). The sample runs from January 2014 through December 2021.

<sup>8</sup>We drop observations where  $ys_{i,t}$  is smaller than 300bps or higher than 3000bps. We also winsorize  $\Delta ys_{i,t}$  at the bottom 5% and the top 5%; as we are using a large cross-section of corporate bond yields, we want to capture the overall variation in credit spreads, and therefore we want to limit the effects of idiosyncratic corporate bond variation on our monetary policy shock.

### 3.1 Credit Spread Monetary Policy Shocks and ECB Corporate Purchases

In this section, we discuss the relationship between our credit spread monetary policy shocks and ECB policy. In March 2016, the ECB announced its decision to extend its asset purchase programme to euro-denominated investment-grade corporate bonds issued by non-financial corporations established in the Euro area. On the day of the CSPP announcement, we observe a sharp contraction of market corporate bond yield, with yields continuing to fall in the period after the announcement (see Figure 3a). The figure clearly shows that the CSPP announcement successfully lowered bond yields.

In 2020, the Covid crisis prompted the ECB to increase the size of its interventions by extending its asset purchase programme. The spike in the number of Covid cases in Eurozone countries and the uncertainty related to the economic outlook led to a sudden adverse repricing of financial assets. Amid the turbulence in the financial market, after the scheduled governing council of March 12, 2020, ECB President Christine Lagarde announced a series of measures to support the economy. The ECB announced to "add a temporary envelope of additional net asset purchases of €120 billion until the end of the year, ensuring a strong contribution from the private sector purchase programmes".<sup>9</sup> After the announcement, bond market prices plunged as financial markets considered the ECB inadequate to face the magnitude of the shocks.

On March 18, 2020, outside a regularly scheduled governing council meeting, the ECB announced its decision to massively increase the size of its interventions through pandemic emergency purchase programme (PEPP), with an initial envelope of €750bn. The two vertical lines in Figure 3b denote the dates of these two events. The figure exhibits the rise in yields in March 2020, which was exacerbated by the ECB meeting. The PEPP announcement stopped the fierce rise in bond yields and led to a sudden reversal and a reduction in funding costs. Overall, Figure 3 shows the powerful effects of ECB corporate purchases on corporate bond yields.<sup>10</sup>

The transmission of ECB monetary policy can work either through a reduction in the interest rate component of corporate bond yields or via compression of the credit risk component. In order to shed light on the transmission mechanism, we use the decomposition described in Equation 1, to disentangle the variation in the two components around ECB announcements. Following the CSPP announcement, the net interest rate contribution is null. We instead observe a sizeable reduction in credit spread, which is driving most of the variation.

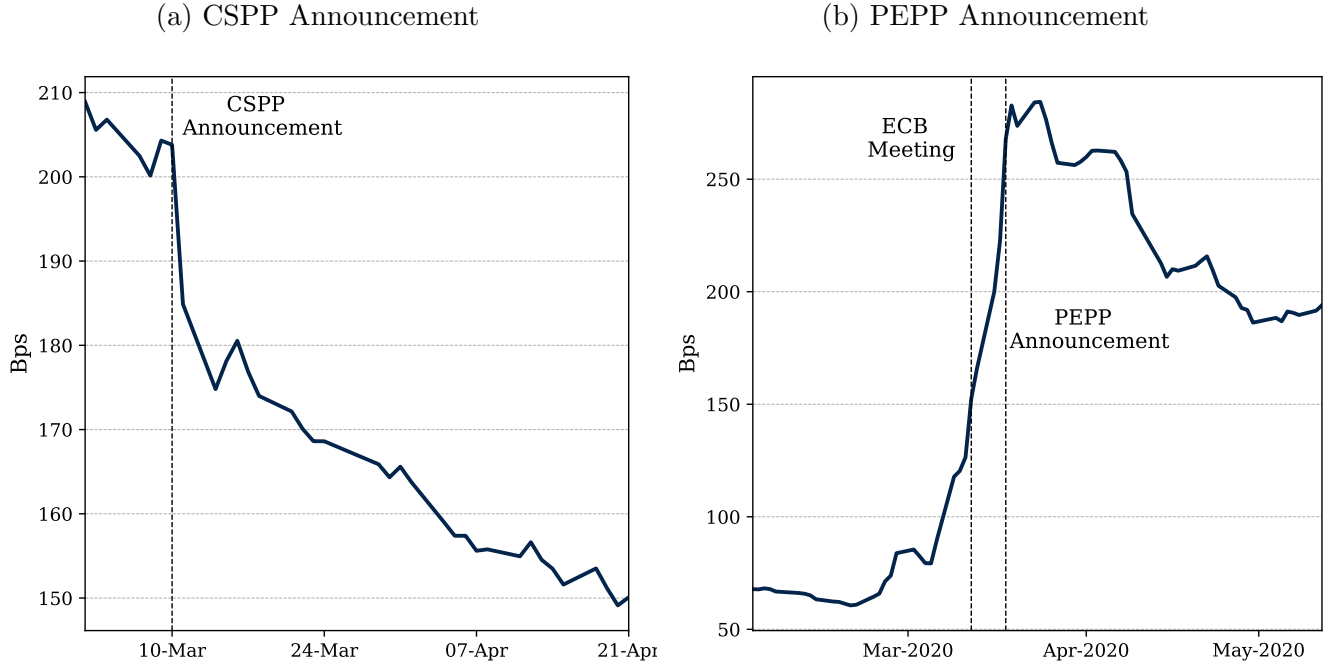
Figure 4b also shows that the relative contribution of the credit risk component dwarfed the

---

<sup>9</sup>ECB Press Conference, March 12, 2020

<sup>10</sup>The two policy packages did not only affect prices; the period following the announcements also saw an increase in the volume of bond issuance by euro area non-financial corporations, as shown in Figure 20, in Appendix B. Several papers in the literature (see Abidi and Miquel-Flores (2018); Grosse-Rueschkamp et al. (2019); Zaghini (2020); De Santis et al. (2018); Todorov (2020); De Santis and Zaghini (2021)) further demonstrated the causal link between CSPP announcement and the effects on bond yields and bond volumes.

Figure 3: Announcements of CSPP and PEPP



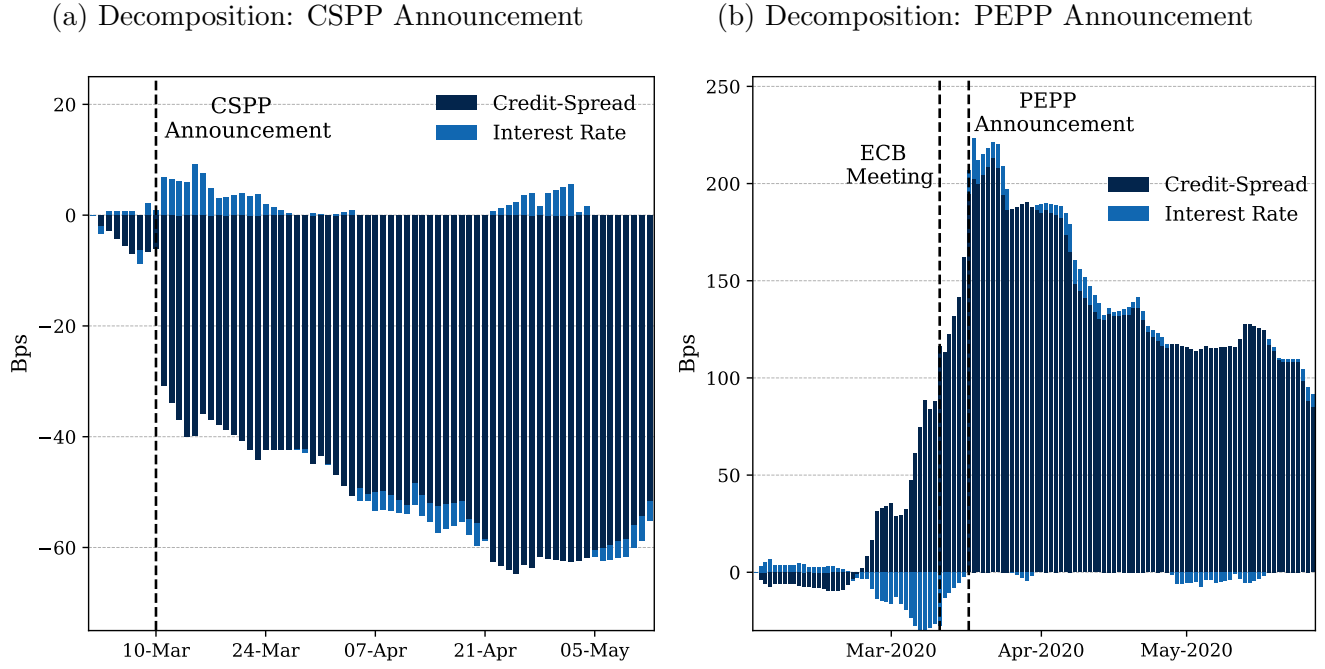
*Note:* The figure shows the evolution of the value-weighted Euro area non-financial corporations bond yields. Figure 3a shows the evolution around the Corporate Sector Purchase Programme (CSPP) announcement on March 10, 2016. Figure 3b shows the evolution around the Pandemic Emergency Purchase Programme (PEPP) announcement on March 20, 2020. Source: Markit iBoxx.

interest rate component at the launch of the PEPP programme. As expected, the rise in bond yields at the onset of the Covid crisis was entirely driven by a surge in credit spread. The timely interventions of the ECB halted the spike in credit spread and led to a gradual reduction in yields.

The two examples show the importance of ECB corporate sector purchases for credit spreads. Figure 5 plots the time series of credit spread shocks from 2014 through 2021. Each bar in the figure corresponds to a scheduled ECB governing council meeting: a negative number means credit spreads fall on that meeting. In the figure, we show that the most significant shocks (highlighted in dark blue) all correspond to announcements regarding corporate sector purchases. In March 2016, the ECB launched the CSPP programme, and in December 2016 the ECB extended the duration of the APP programme. In December 2019, the ECB re-started the APP programme after its termination in December 2018. In March 2020, the ECB announced its measure to fight the turmoil in the market due to the Covid Crisis. Finally, in June 2020, the ECB extended the PEPP in terms of the amount of purchases (increasing the envelope for PEPP by €600 billion to a total of €1,350 billion) as well in terms of the duration of the programme.

We also show that the credit spread monetary policy shocks are poorly explained by monetary policy shocks estimated using OIS rates. The total variance explained by all shocks is roughly zero (see Table I). This analysis is further confirmed in Section E.1 where we conduct a variance decomposition of the variation in corporate bond yields around ECB announcement and demon-

Figure 4: Decomposition of CSPP and PEPP announcement



*Note:* The figure shows the decomposition of bond yields around ECB CSPP and PEPP announcement. We decompose the market corporate bond yield into an interest rate component and a credit spread component. The figure plots the individual contribution of the two components to the observed variation the overall corporate bond market yields. Source: Markit iBoxx.

strate the relevance of the credit spread component. Our analysis shows the importance of using our new credit spread monetary policy shock to study the effects of ECB policies on the corporate bond market.

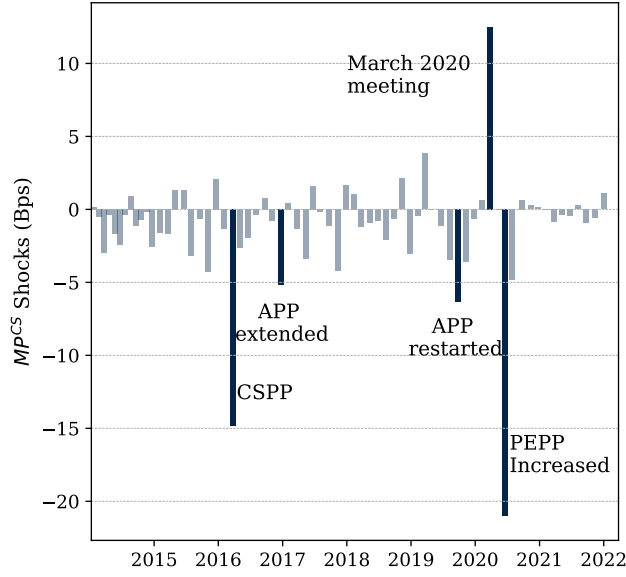
### 3.2 Term Spread Monetary Policy Shocks and ECB Government Bond Purchases

This section discusses the main drivers of our term spread monetary policy shock. The term spread monetary policy shock is highly correlated with the QE-OIS<sup>11</sup> shocks developed by [Altavilla et al. \(2019\)](#) based on the Euro OIS curve based on the methodology of [Swanson \(2021\)](#). The two shocks have a correlation of 0.7 in our sample (see Table II). The OIS-QE shock was developed to capture the effects of government bond purchases. Consistently, We find that the two most significant QE-OIS shocks correspond to January 2015 and December 2015. In January 2015, the ECB announced its government sector purchase programme (PSPP). In December 2015, markets expected a significant increase in the size of the PSPP programme. The package announced was

<sup>11</sup>We call the shocks QE-OIS as it captures the effect of quantitative easing through changes in the OIS rates. Our credit spread monetary policy shock also captures changes in QE (as corporate sector purchases are also a type of QE), using corporate sector credit spreads.



Figure 5: Time Series of Credit Spread Monetary Policy Shocks



*Note:* The figure shows the measured credit spread shocks  $MP^{CS}$  on ECB announcements from January 2014 through December 2021. We identify the largest shocks. The first coincides with the announcement of the Corporate Sector Purchase Programme (CSPP), the second and third with two announcements related to the Asset Purchase Programme (APP), the fourth with the March 2020 where the ECB discussed the policy package after the Covid outbreak, the fifth coincides with the ECB extension of Pandemic Emergency Purchase Programme (PEPP). Source: Markit iBoxx.

Table I: Shocks and Various Interest Rate Shocks

	(1)	(2)	(3)	(4)
	$MP^{CS}$	$MP^{CS}$	$MP^{CS}$	$MP^{CS}$
	b/se	b/se	b/se	b/se
$MP^{TS}$	-0.010 (0.49)			0.052 (1.34)
QE-OIS		-0.127 (0.43)		-0.183 (2.50)
OIS 2Y			-0.060 (0.25)	-0.006 (0.88)
Observations	80	80	80	80
R-squared	0.000	0.002	0.002	0.002

*Note:* The table shows the results from three regression estimates. We regress the credit spread monetary policy shocks on (i) the term spread shocks, (ii) the QE shocks developed by [Altavilla et al. \(2019\)](#) based on OIS rates and (iii) the change in the two year OIS around ECB announcement, a standard measure of monetary policy shock and (iv) all shocks together. The sample runs from 2014 through 2021.

below market expectations and led to a sharp increase in interest rates (a positive QE-OIS shock).

The term spread monetary policy shock and credit spread monetary policy shocks have a low correlation. In the following sections, we separately study the effect of the two shocks on the corporate bond market and the balance sheet of different intermediaries.

Table II: Shocks Correlation

	$MP^{CS}$	$MP^{TS}$	IR	Stoxx 50	QE-OIS	OIS 2Y
$MP^{CS}$	1.00	-0.01	0.03	-0.39	-0.01	-0.01
$MP^{TS}$	-0.01	1.00	-0.03	0.13	0.70	-0.29
IR	0.03	-0.03	1.00	-0.58	0.53	0.72
Stoxx 50	-0.39	0.13	-0.58	1.00	-0.27	-0.54
QE-OIS	-0.01	0.70	0.53	-0.27	1.00	0.47
OIS 2Y	-0.01	-0.29	0.72	-0.54	0.47	1.00

*Note:* The table shows the correlation for different shocks: our credit spread monetary policy shocks, our term spread shocks, the total variation in the interest rate component of corporate bonds around ECB announcements, the variation in the Euro Stoxx 50 around ECB announcement, the QE shocks developed by [Altavilla et al. \(2019\)](#) based on OIS rates and the change in the two year OIS around ECB announcement. The sample runs from 2014 through 2021.

## 4. Heterogeneous Response to Monetary Policy Shocks

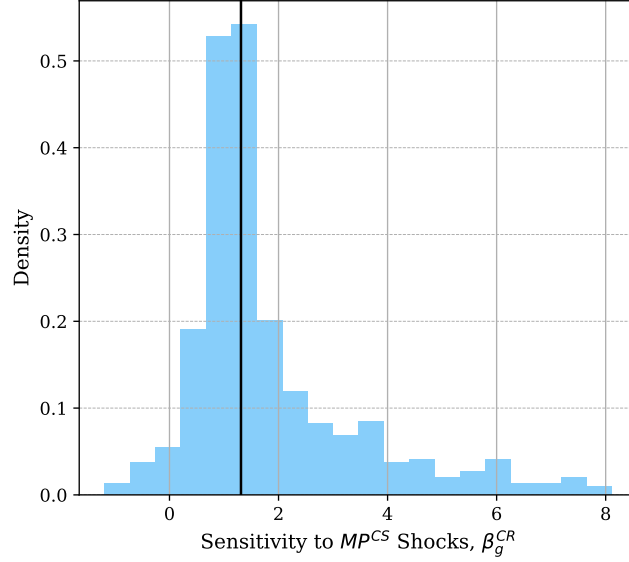
### 4.1 Corporate Bonds and Credit Spread Shocks

In this section, we estimate the cross-section of the sensitivity of corporate bonds to our credit spread monetary policy shock. We show how different bonds are affected by monetary policy shocks and discuss the relationship between the investor base composition of the bonds and the sensitivity of different bonds to the shocks.

There are several challenges in evaluating the response of corporate bonds to monetary policy shocks. First, bonds have limited maturity, and this limits their sample. Second, bond characteristics change over time: a bond mechanically changes maturity (or duration) as time goes by; the riskiness of the firm also varies over time. Third, corporate bonds exhibit a significant degree of idiosyncratic risk, which can confound our estimates. To circumvent these issues, we exploit the panel dimension of our corporate bond price database. We use a set of bond characteristics (maturity, ratings, sector) to cluster bonds into different groups.<sup>12</sup> Our final sample includes a

<sup>12</sup>We only use Euro-denominated corporate bonds. To have a larger sample, we use corporate bonds issued by non-financial corporations and by captive institutions. We include bonds with ratings ranging from AAA to B. We sample bonds according to maturity by dividing them into ten equally sized groups (maturity deciles). We use the average rating average of Fitch Ratings, Moody's, and Standard & Poor's (S&P) as well as the worst rating of the three rating agencies. Finally, we use sector according to the definition of Markit-iBoxx. We only consider groups with more than 20 observations

Figure 6: Distribution of Sensitivities to  $MP^{CS}$



*Note:* The figure shows the distribution estimated bond sensitivity to  $MP^{CS}$  for different groups of corporate bonds, as estimated in Equation 3. We use a set of bond characteristics (maturity, ratings, sector) to cluster bonds into different groups. The vertical line is the median estimate. The sample runs from 2013Q4 to 2021Q4. Source: Markit iBoxx.

total of 539 groups: each group is a combination of maturity-rating-sector.

We then evaluate the sensitivity of the yield spread of each of these groups to our  $MP^{CS}$  shocks. We run a battery of panel regressions:

$$\Delta ys_{i,t}^g = \gamma_i + \beta_g^{CS} MP_t^{CS} + \varepsilon_{i,t}^g, \quad (3)$$

where  $\Delta ys_{i,t}^g$  is the change in yield of corporate bond  $i$  of the group  $g$  at time  $t$ , measured using the 2-day change around each ECB announcement.<sup>13</sup>  $\gamma_i$  is a bond fixed effect.<sup>14</sup> A time  $t$  corresponds to each of the ECB announcements in our sample. We hence run a total of 539 panel regression (one for each group), where each individual observation is a combination of bond-time. This results in a total of 539 estimated  $\beta_{CS}^g$ . Figure 6 plots the distribution of sensitivity to our credit spread monetary policy shock  $MP^{CS}$ , illustrating the dispersion in  $\beta_g^{CS}$  estimates. The median sensitivity is around 1.2 (the black vertical line).

We now combine our corporate bond sensitivity results with asset holdings' data. For each group, exploiting the granular information of the ECB SHS database, we compute the holdings by two groups of investors: insurance corporations and pension funds (ICPF) and mutual funds. We pool data across the whole sample: for each group, we compute the share held by the different

<sup>13</sup>As we are comparing different corporate bonds, we allow for a slightly longer window to reduce the heterogeneous differences due to liquidity.

<sup>14</sup>We use a bond fixed effect to reduce the noise due to idiosyncratic factors.

intermediaries at each  $t$  and average across the entire sample. We then match the holdings for each group  $g$  with the estimated  $\beta_{CS}^g$ .

Figure 7 plots the results: each dot corresponds to a group. Figure 7a reports on the x-axis the estimated  $\beta_{CS}^g$  and on the y-axis the holding share (in percentage) by insurance corporations and pension funds (ICPF) for each group. The plot clearly shows a negative correlation between insurance and pension fund holdings and the sensitivity of bonds to  $MP^{CS}$  shocks. ICPFs hold more of the least responsive bonds (low  $\beta_{CS}^g$ ) and less of the most responsive bonds (high  $\beta_{CS}^g$ ).

By contrast, Figure 7b plots the equivalent estimates for mutual funds. The figure shows a positive correlation between mutual fund holdings and the bond sensitivity to CR shocks. In this case, mutual funds hold less of the least responsive bonds and more of the most responsive bonds.

To sum up, Figure 7 shows a striking correlation between holdings by different intermediaries and the sensitivity of corporate bonds to our credit spread monetary policy shocks. Understanding the nature of this correlation is vital to learn how monetary policy shocks transmit to the cross-section of corporate bonds and how the transmission interacts with heterogeneous intermediaries' balance sheets. In section 4.1.1 we document that this heterogeneity in responses primarily reflects bond characteristics, rather than the investor base composition for a given bond, implying that intermediary sectors systematically select into different types of assets.

#### 4.1.1 Holding Structure and Sensitivity to Credit Spread Shocks

Recent literature proposed that the investor base composition is an essential determinant of bond price dynamics. Within our context, bonds held by mutual funds may react differently to shocks because they are held by mutual funds. Coppola (2021) shows that the holding structure of the bond has significant effects on the transmission of credit crunches during a crisis to corporate bond yields.<sup>15</sup>

We use a similar analysis within our framework. We use data on holdings by ICPF (insurance and pension funds) and mutual funds and test the relevance of the holding structure on the transmission of  $MP^{CS}$  shocks to bond yield changes.

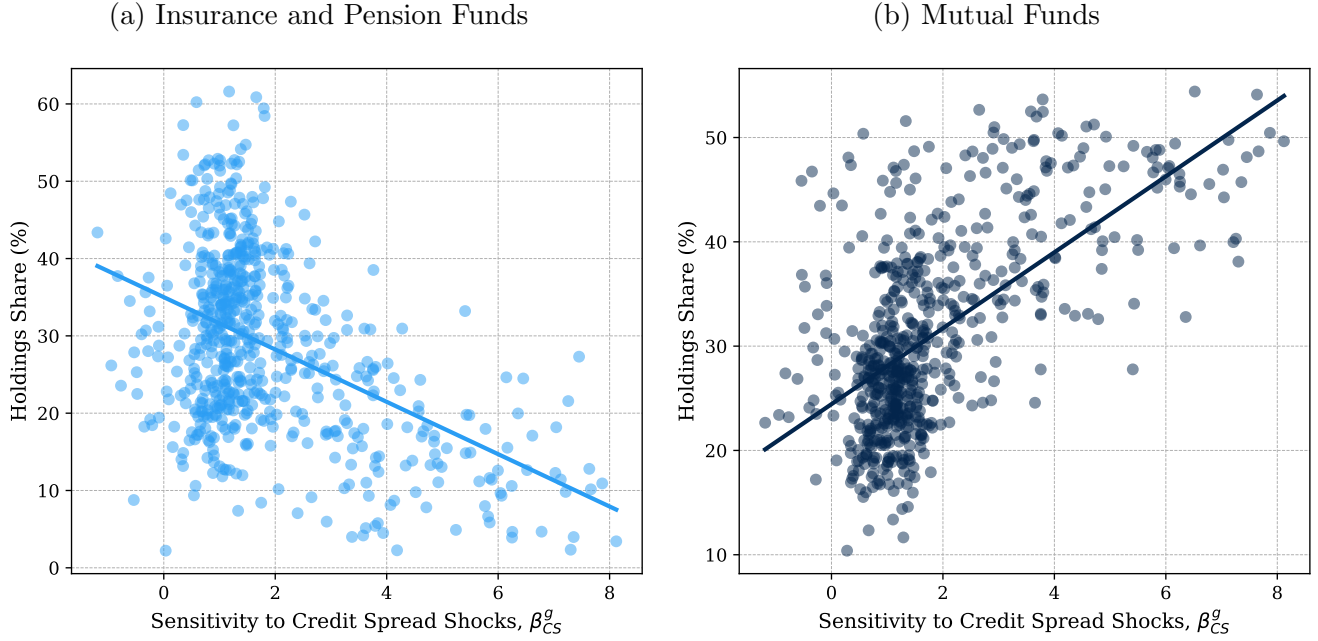
We use the information on bond characteristics to dissect the driver of heterogeneity. Controlling for bond characteristics, we can exploit heterogeneity in the holding structure of bonds (the share held by different intermediaries) to understand if intermediaries play a role in the transmission of monetary policy shocks to bond yields—beyond the characteristics of the bond. We run the panel regression:

$$\begin{aligned} \Delta y_{s,i,t} = & \gamma_i + \gamma^{CS} MP_t^{CS} + \gamma^M \theta_{M,i,t} \times MP_t^{CS} + \beta^L \theta_{L,i,t} \times MP_t^{CS} + \text{Interacted Controls} \times MP_t^{CS} \\ & + \theta_{M,i,t} + \theta_{L,i,t} + \text{Interacted Controls} + \varepsilon_{i,t}, \end{aligned} \quad (4)$$

---

<sup>15</sup>In Appendix D we show that those results are confirmed in our data and with our specifications.

Figure 7: Holding shares and Bond Sensitivity to  $MP^{CS}$  Shocks



*Note:* The figure plots on the x-axis the estimated  $\beta_{CS}^g$  from Equation 3 (i.e., the sensitivity of each bond group to our credit spread monetary policy shock,  $MP^{CS}$ ) and on the y-axis the holding share (in percentage) for each group by insurance corporations and pension funds (Figure 7a) and mutual funds (Figure 7b). The shares are computed as the average shares across the full sample for each group. The sample runs from 2014 through 2021. Source: Markit iBoxx and ECB SHS.

where  $\Delta ys_{i,t}$  is the two-day change in spread of bond  $i$  around announcement  $t$ ,  $\theta_{M,i,t}$  are the shares of bond  $i$  held by mutual funds at time  $t$ ,  $\theta_{L,i,t}$  are the shares of bond  $i$  held by levered intermediaries (insurance companies and pension funds) at time  $t$ .

We run the regression for the sample 2014-2021 and restrict the sample to investment-grade non-financial corporate bonds. We use a different set of controls: duration<sup>16</sup>, issuer country, average rating, worst rating, bid-ask spread quintile, liquidity quintile, issuer sector, issuer company.

Table III reports the results. The coefficient of interest is  $\beta^L$  and  $\beta^M$ : the marginal effects of bond holdings on the response of bond yields to monetary policy. Going from column (1) to column (11), we are increasing the number of controls.

We then find that once we account for characteristics, there are minor differences in the sensitivity of bonds to monetary policy shocks. Figure 8 displays how including fixed effects on characteristics reduce the difference between the two coefficients for each of the model detailed in Table III, suggesting the characteristics of the bond rather than the holders are the main drivers of the different sensitivity to monetary policy shocks.

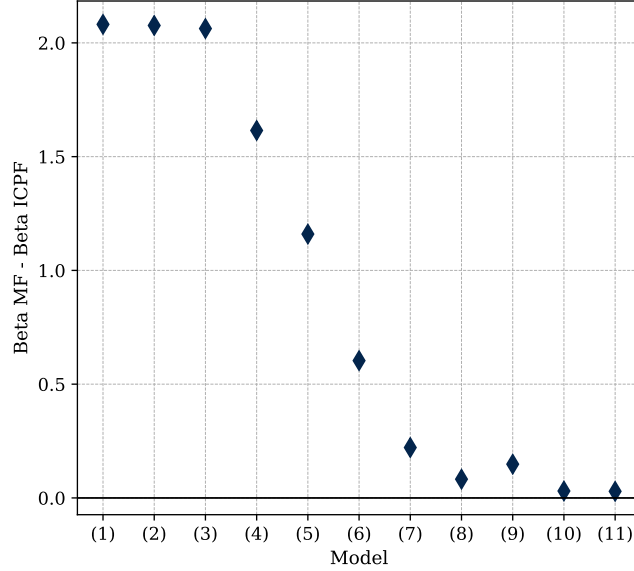
<sup>16</sup>We use a dummy for each duration group: 0y-1y, 1y-3y, 3y-5y, 5y-8y, 8y+.

Table III: Holdings and Sensitivity to  $MP^{CS}$  shocks

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys	$\Delta$ ys
	b/se	b/se	b/se	b/se	b/se	b/se	b/se	b/se	b/se	b/se	b/se
$MP^{CS} \times \theta_L$	0.223	0.144	0.191	0.314	0.314	0.144	0.110	0.106	-0.021	0.092	0.144
	(0.20)	(0.12)	(0.15)	(0.13)	(0.12)	(0.10)	(0.07)	(0.08)	(0.05)	(0.06)	(0.08)
$MP^{CS} \times \theta_M$	2.304	2.220	2.254	1.929	1.473	0.748	0.332	0.188	0.128	0.123	0.173
	(0.75)	(0.59)	(0.38)	(0.31)	(0.27)	(0.23)	(0.17)	(0.10)	(0.07)	(0.08)	(0.09)
Levels	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	40085	40070	39611	38666	37131	31184	18493	10340	25476	13199	7132
Adj. R-squared	0.561	0.710	0.719	0.758	0.798	0.878	0.923	0.962	0.963	0.971	0.974
Fixed Effects	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN
Duration	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Country	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Rating	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Rating Worst	No	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Liquidity	No	No	No	No	No	No	Yes	Yes	No	Yes	Yes
Bid-ask	No	No	No	No	No	No	No	Yes	No	No	Yes
Sector	No	No	No	No	No	Yes	No	Yes	Yes	Yes	Yes
Issuer	No	No	No	No	No	No	No	No	Yes	Yes	Yes

*Note:* The table shows the results of Equation 4. We run the regression for the sample 2014-2021 and restrict the sample to investment-grade non-financial corporate bonds. We use a different set of controls: duration (a dummy for each duration group: 0y-1y, 1y-3y, 3y-5y, 5y-8y, 8y+), issuer country, average rating, worst rating, bid-ask spread quintile, liquidity quintile, issuer sector, issuer company. Source: Markit iBoxx and ECB SHSs.

Figure 8: Levered Institutions vs Mutual Funds: Difference between sensitivity



*Note:* The figure shows the difference between the estimated  $\beta^M$  and  $\beta^L$  in Equation 4 for a difference set of controls (i.e., the marginal effects of intermediary holdings on the sensitivity of bonds to our credit spread monetary policy shock,  $MP^{CS}$ ). The x-axis numbering corresponds to the model estimated in each column of Table III. Each model use a different set of bond characteristics fixed-effects.



## 4.2 Corporate Bonds and Term Spread Shocks

In Section 4.1 we discussed the heterogeneous effects of corporate bonds on our credit spread monetary policy shock. A natural question is whether the correlation between intermediary holdings and monetary policy holds for a different type of shock. To this end, in this section, we study the relationship between the cross-sectional effects of the term spread monetary policy shocks and intermediary holdings.

As the effects of the term spread shock on a credit spread are close to zero, we estimate the sensitivity of the interest rate component  $\iota_{i,t}$  of the bond to the shocks. We run the following regression:

$$\Delta\iota_{i,t}^g = \gamma_i + \beta_g^{TS} MP_t^{TS} + \varepsilon_{i,t}^g, \quad (5)$$

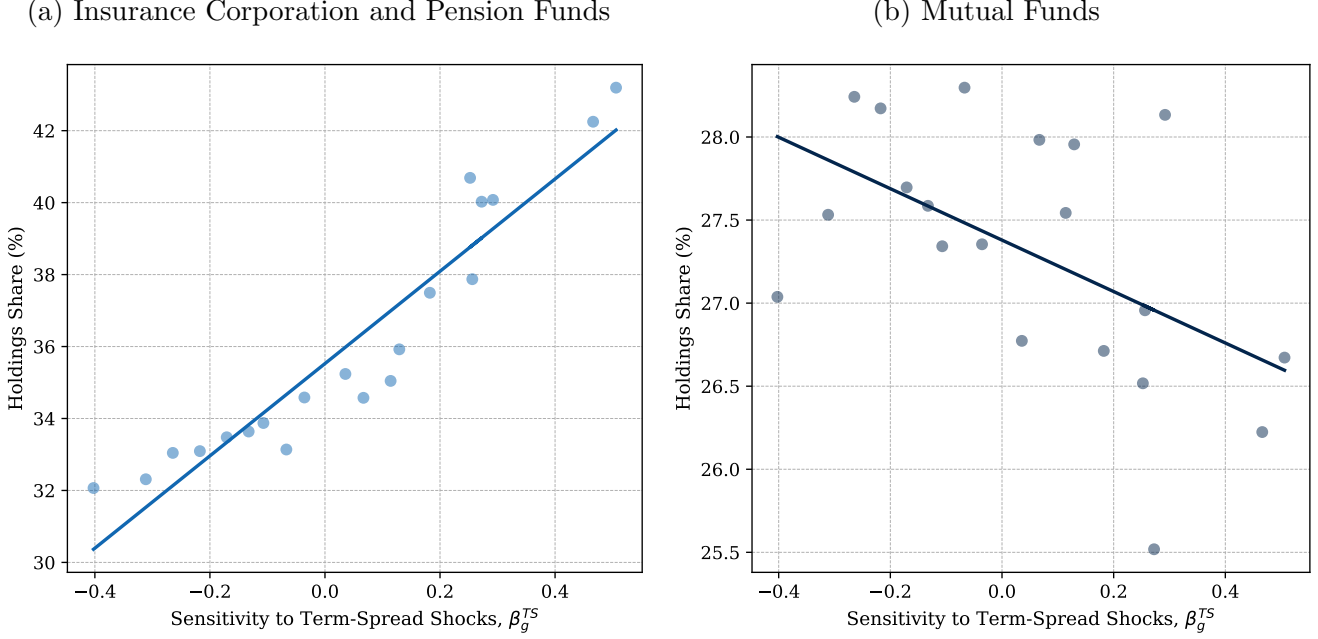
where  $\Delta\iota_{i,t}^g$  is the change in the credit risk-free component of bond  $i$  of group  $g$  at time  $t$ , measured using the change around each the ECB announcement. In this case, we can use the shorter one-day window as we are measuring the sensitivity of the OIS component of different bonds to the shocks. Moreover, bonds only differ in their maturity. We, therefore, split our sample into twenty maturity groups and estimate twenty  $\beta_g^{TS}$ . The results are reported in Figure 9.

Our results show a striking correlation between intermediary holdings and bond sensitivity to our term spread shocks. Insurance companies and pension funds hold larger shares of more responsive bonds (see Figure 9a). This result is antithetical to what we have found for the credit spread monetary policy shock. In section 4.1.1 we exploit that significant cross-sectional variation in corporate bond spreads to illustrate how bond sensitivity primarily depends on the characteristics of the bonds rather than the individual investor base composition. It is impossible to conduct an analogous exercise for the term spread shocks as the OIS component  $\iota_{i,t}$  is common to all bonds with the same maturity.

## 5. A Model of the Bond Market

In section 4 we showed a stark correlation between intermediaries' bond holdings and the sensitivity to different monetary policy shocks. Furthermore, we have shown that this correlation is mainly driven by how different intermediaries select into different bonds based on their characteristics. In order to rationalize these findings, we develop an intermediary asset pricing model with multiple risky assets. In section 5.1 we describe the environment, while in section 5.2 we show the main predictions of the model.

Figure 9: Holding shares and Bond Sensitivity to TS Shocks



*Note:* The figure plots on the x-axis the estimated  $\beta_{TS}^g$  of Equation 5 and on the y-axis the holding share (in percentage) for each group by insurance corporations and pension funds (Figure 9a) and mutual funds (Figure 9b). The shares are computed as the average shares across the full sample for each group. Source: Markit iBoxx and ECB SHS.

## 5.1 Environment and Equilibrium

### 5.1.1 Assets

**Factors** We exploit the factor structure in fixed-income securities to represent different agent portfolios in terms of simple factor portfolios. We focus on two orthogonal risk factors: interest rate risk (IR) and credit risk (CR). The two risk factors follow the dynamic:

$$\begin{aligned} f_{t+1}^{IR} &= \rho f_t^{IR} + \varepsilon_{t+1}^{IR}, \\ f_{t+1}^{CR} &= \rho f_t^{CR} + \varepsilon_{t+1}^{CR}, \\ \varepsilon^{IR} &\sim N(0, \sigma_{IR}^2) \quad \varepsilon^C \sim N(0, \sigma_{CR}^2), \end{aligned}$$

where the two shocks are uncorrelated. Define  $F_t = (f_t^{IR}, f_t^{CR})'$ , the vector of the two factors.

There are two claims, whose payments depend on the realization of the two factors. The IR claim depends on the realization on  $f^{IR}$  while the CR claim depends on the realization of  $f^{CR}$ . The payments are:

$$\log d_{t+1}^{IR} = \alpha^{IR} + f_{t+1}^{IR}$$

$$\log d_{t+1}^{CR} = \alpha^{CR} + f_{t+1}^C.$$

The IR claim and CR claim, are in positive supply:  $\Theta = (\theta^{IR}, \theta^{CR})'$ .

We also have an additional instrument, which is the debt of the intermediary. The gross supply is also fixed to  $\bar{\ell}$  and the payment is:

$$\log d_{t+1}^\ell = \alpha_\ell + \beta_\ell^{IR} f_{t+1}^{IR}$$

### 5.1.2 Central bank

The central bank chooses the amount of risk-free debt  $b_t^{CB}$ , its price  $q_t^{CB}$  and the allocation to risky assets  $X_{CB,t} = (x_{CB,t}^{IR}, x_{CB,t}^{CR})'$ . The central bank budget constraint is:

$$(P'_t + D'_t)X_{CB,t-1} - b_{t-1}^G = P'_t X_{CB,t} - q_t b_t^G + \Pi_t^{CB}. \quad (6)$$

where  $P_t = (p_t^{IR}, p_t^{CR})'$  is the vector of prices,  $D_t = (d_t^{IR}, d_t^{CR})'$  is the vector of payments from the two risky claims.  $\Pi_t^{CB}$  are the proceeds from the central bank portfolio. The proceeds are rebated to households: a share  $1 - \varphi$  is redistributed to the bond investor, while a share  $\varphi$  is redistributed to the passive household. This implies that the taxes/rebates to the two household types are:

$$T^a = (1 - \varphi)\Pi_t,$$

$$T^p = \varphi\Pi_t,$$

where  $T^a$  are total taxes/rebates to the active households while  $T^p$  are total taxes/rebates to the passive households.

### 5.1.3 Asset Managers

The asset managers choose the allocation across risky claims  $X_{M,t} = (x_{M,t}^{IR}, x_{M,t}^{CR})'$ . They do not have the technology to borrow. They further choose the net payoff to its shareholder,  $\pi_t^M$ .  $\pi_t^M$  can be either positive (net-dividend payment or buyback) or negative (equity injection by the household). The asset manager budget constraint is defined as:

$$\pi_t^M + P'_t X_{M,t} \leq W_t^M, \quad (7)$$

where  $W_t^M$  is the initial wealth. The wealth evolution follows the equation:

$$W_{t+1}^M = (P'_{t+1} + D'_{t+1})X_{M,t}. \quad (8)$$

The asset managers maximizes the value of the firm,

$$V^M(W_t^M, F_t) = \max_{\pi_t^M, X_{M,t+1}} \pi_t^M + \mathbb{E} [M_{t+1}^a V^M(W_{t+1}^M, F_{t+1})]$$

subject to Equation 7-8, where  $M_{t+1}^a$  is the stochastic discount factor of the active households, the shareholder of the firm.

#### 5.1.4 Levered Intermediaries

The levered intermediaries have initial net-wealth  $W_t^L$ . They choose their asset allocation across risky claims,  $X_{L,t} = (x_{L,t}^{IR}, x_{L,t}^{CR})'$  and the net payoff to its shareholder,  $\pi_t^L$ . The intermediary takes the amount of central bank debt  $b_{t+1}^{CB}$ , their price  $q_t^{CB}$  as well as the amount of borrowing  $\bar{\ell}$  as given.

**Regulatory Framework** The intermediary net-proceeds from the investment is:

$$\tilde{\pi}_{t+1}^L = D'_{t+1} X_t - \bar{\ell} d_{t+1}^{CR}. \quad (9)$$

The intermediary faces a regulatory cost proportional to the variance of its net-proceeds. The total cost is:

$$\frac{\psi}{2} \text{Var}(\tilde{\pi}_{t+1}^L), \quad (10)$$

where  $\psi$  is the parameter that governs the severity of the regulatory cost. The intermediary also faces a collateral constraint:

$$P'_t \Omega X_t^L \leq \bar{\ell} p_t^\ell \quad (11)$$

where  $\Omega$  is a diagonal matrix containing the collateral coefficients  $\omega^{IR}, \omega^{CR}$  of the two risky claims.

**Intermediary Problem** The intermediary budget constraint is defined as:

$$\pi_t^L + \frac{\psi}{2} \text{Var}(\tilde{\pi}_{t+1}^L) + P'_t X_{L,t} + q_t^{CB} b_t^{CB} - p_t^\ell \ell \leq W_t^L. \quad (12)$$

The wealth evolution follows the equation:

$$W_{t+1}^L = (P'_{t+1} + D'_{t+1}) X_{L,t} - \ell (p_{t+1}^\ell + d_{t+1}^\ell) + b_t^{CB}. \quad (13)$$

The intermediary maximizes the value of the firm,

$$V^L(W_t^L, F_t) = \max_{\pi_t^L, X_{L,t}} \pi_t^L + \mathbb{E} [M_{t+1}^a V^L(W_{t+1}^L, F_{t+1})] \quad (14)$$

subject to Equation 9-13, where  $M_{t+1}^a$  is the stochastic discount factor of the entrepreneur, the shareholder of the firm.

### 5.1.5 Active Investors

The active investors have Epstein-Zin Utility over consumption. They are the shareholders of the intermediaries. We also assume the active households receive a non-stochastic stream of income  $y$ . Their consumption is:

$$C_t^a + \leq y^a + \pi_t^L + \pi_t^M + T_t^a + \frac{\psi}{2} Var(\tilde{\pi}_{t+1}^L) \quad (15)$$

where  $C_t^a$  is consumption at time  $t$ ,  $\pi_t^M$  is the net payout from the asset manager,  $\pi_t^L$  is the net payout from the levered intermediary, and  $T_t^a$  are taxes/rebate. We also assume that the regulatory cost of the levered intermediary,  $\frac{\psi}{2} Var(\tilde{\pi}_{t+1}^L)$ , is rebated to the intermediary shareholders.

The active household utility is:

$$V^a(F_t) = \left\{ (1 - \beta) C_t^{a1 - \frac{1}{\sigma_a}} + \beta_e E_t [V(F_{t+1})^{1 - \gamma_a}]^{\frac{1 - \frac{1}{\sigma_a}}{1 - \gamma_a}} \right\}^{\frac{1}{1 - \frac{1}{\sigma_a}}}, \quad (16)$$

subject to Equation 15. Further define the SDF of the active investors as:

$$M_{t+1}^a = \beta^a \left( \frac{V_{t+1}^a}{C E_t^a} \right)^{1/\sigma_e - \gamma_e} \left( \frac{C_{t+1}^a}{C_t^a} \right)^{-1/\sigma_e} \quad (17)$$

### 5.1.6 Passive Households

The passive has no access to the  $IR$  and  $CR$  assets. However, it can saves in the liabilities of the intermediary. We also assume the passive households receive a non-stochastic stream of income  $y^p$ . The budget constraint is therefore defined as:

$$C_t^p + p_t^\ell \ell_{p,t+1} \leq y + W_t^p + T_t^\omega, \quad (18)$$

where  $W_t^p$  is the wealth of the passive at time  $t$ , and  $T_t^\omega$  are the total taxes and rebate. The wealth evolution of the passive is:

$$W_{t+1}^p = y^p + \ell_{p,t+1} (d_{t+1}^\ell + p_{t+1}^\ell) \quad (19)$$

The passive household solves the problem:

$$V^p(W_t^p, F_t) = \max_{C_t^p, \ell_{t+1}^p} \left\{ (1 - \beta) C_t^{p1 - \frac{1}{\sigma_p}} + \beta_p E_t [V(W_{t+1}^p, F_{t+1})^{1 - \gamma_p}]^{\frac{1 - \frac{1}{\sigma_p}}{1 - \gamma_p}} \right\}^{\frac{1}{1 - \frac{1}{\sigma_p}}} \quad (20)$$

subject to Equation 18 - 19. Further define the SDF of the passive household as:

$$M_{t+1}^p = \beta^p \left( \frac{V_{t+1}^p}{C E_t^p} \right)^{1/\sigma_p - \gamma_p} \left( \frac{C_{t+1}^p}{C_t^p} \right)^{-1/\sigma_p} \quad (21)$$

### 5.1.7 Equilibrium

After defining the asset space and the individual problems of each agent, we describe the equilibrium of the model.

Given a process for  $\{F_t\}_0^\infty$  and a Central bank policy  $\{X_{CB,t}, b_t^{CB}, q_t^{CB}\}_0^\infty$ , the equilibrium is:

1. a path of prices  $\{p_t^{IR}, p_t^{CR}, p_t^\ell\}_0^\infty$
2. consumption  $\{C_t^p, C_t^a\}_0^\infty$
3. and allocations  $\{X_{M,t}, X_{L,t}\}_0^\infty$

such that

1. the active households maximize their utility 16 subject to the budget constraint 15
2. the passive households maximize their utility 20 subject to the budget constraint 18
3. the asset managers maximize their utility subject to Equation 7-8
4. the levered intermediary maximize their utility subject to Equation 9-13
5. all markets clear:

$$x_{M,t}^{IR} + x_{L,t}^{IR} + x_{CB,t}^{IR} = \theta^{IR}$$

$$x_{M,t}^{CR} + x_{L,t}^{CR} + x_{CB,t}^{CR} = \theta^{CR}$$

$$\ell_t = \bar{\ell}$$

6. the good market clears:

$$C_t^a + C_t^p = \theta^{IR} d_t^{IR} + \theta^{CR} d_t^{CR} + y^p + y^a$$

### 5.1.8 Model Solution and Policy Experiment

We use the model to provide qualitative predictions. However, in order to better illustrate the mechanisms, we based our results on a calibrated set of parameters. We first solve the model for an economy where  $\{X_{CB,t}, b_t^{CB}\}_{t=0}^\infty = 0$ . We solve for a set of equilibrium prices and allocations  $(\tilde{P}_t, \tilde{p}_t^\ell, \tilde{X}_t^L, \tilde{X}_t^M)$ , for the value of shocks equal to 0:  $f_t^{IR} = f_t^{CR} = 0$ . We use these set of equilibrium prices and allocations to calibrate the parameters, as detailed in Section 5.1.9. We also discussed a set of predictions on allocations in Section 5.2.1 and on asset prices in Section 5.2.2. We then show



the effects of central bank interventions by means of a counterfactual economy with time-invariant central bank's policy  $\bar{X}^{CB}, \bar{b}^{CB}, \bar{q}^{CB}$ .

### 5.1.9 Calibration

We describe the set of parameters used to illustrate the model's mechanism. We normalize the  $\beta^\ell$ , the risk exposure of the levered intermediary debt to IR factor to 1. We set  $\beta^a = 0.96$ , targeting a real rate in the non-stochastic steady state of 4%.

The passive households are only pricing the debt of the levered intermediary. We set the  $\beta^p$  to 0.97, meaning that the passive households are slightly more patient than the active households. We assume that the passive households have log-utility;  $\sigma_p = 1$ ,  $\gamma_p = 1$ . We set the coefficient that controls the tightness of the constraint to  $\psi = 1$ . The quadratic form of the regulatory cost implies that the multiplier on the collateral constraint scales with  $\psi$ . Due to this convenient scaling, we show that allocations and prices do not depend on  $\psi$ . We define:  $\Omega = \text{diag}(\omega + \Delta, \omega)$ , where  $\Delta$  is the collateral advantage of interest rate risk holdings versus credit risk holdings. We set the value to 0.5; in a similar setting, [Gertler and Karadi \(2018\)](#) uses a value of 0.5 for  $\Delta$  as the collateral advantage of government bonds on corporate bonds.

We then jointly calibrate:  $\Theta, \gamma^a, \sigma^a, \bar{\ell}, \omega, \alpha, \sigma$  to match: (i) Risk-premium and volatility of IR and CR to 3% and 8%, (ii), market share of levered intermediaries = 50%, (iii) asset to liability ratio of levered intermediaries = 1.15. To simplify the exposition of results, we set the risk premia and volatility of IR claim and CR claim to be the same. Based on our estimates, the risk-premium on interest rate risk returns and credit risk returns ranges from 2% to 4%; we then calibrate the model to a value of 3%. Our estimates of the volatility are also between 6% and 12%. We calibrate the model to a value of 8%. In the corporate bond market, mutual funds and levered intermediaries (ICPFs) hold roughly similar shares. We, therefore, target bond shares of 50% for both asset managers and levered intermediaries. Based on a regulatory report from EIOPA, the asset-to-liability ratio of insurance companies in Europe is 115%, motivating our value of 1.15.

We calibrate the parameters jointly within pre-defined bounds.<sup>17</sup> We find the set of parameters to be:  $\Theta = (1.7, 1.3)'$ ,  $\gamma_a = 27.5$ ,  $\sigma_a = 1.4$ ,  $\bar{\ell} = 0.5$ ,  $\omega = 0.16$ ,  $\alpha = 0.4$ ,  $\sigma = 0.2$ .

## 5.2 Model Results

### 5.2.1 Asset Allocation

In this section we show how different intermediaries choose to be exposed to different risk factors. Levered intermediaries choose to be more exposed to the interest rate risk factor while asset managers choose to be exposed to the credit risk factors.

---

<sup>17</sup>We set the bounds for  $\theta^{IR}$  and  $\theta^{CR}$  to  $[0.5, 2]$ , for  $\gamma^a$  to  $[5, 30]$ , for  $\sigma_a$  to  $[0.1, 2]$ , for  $\bar{\ell}$  we set the bounds as shares of  $\theta^{IR}$  to  $[0.1, 0.95]$ , for  $\alpha$  and  $\sigma$  to  $[0.01, 0.5]$ .

We consider the allocations without central bank interventions and describe the allocation of the levered intermediary. Combining Equation 9 and 10, the total regulatory cost, at time  $t$ , as a function of  $X_{L,t}$  is:

$$\frac{\psi}{2} (X'_{L,t} \Sigma_t X_{L,t} + \bar{\ell} \sigma_{\ell,t}^2 - \bar{\ell} X'_{L,t} \Sigma_t^\ell).$$

$\Sigma_t$  is the conditional variance-covariance matrix for the payments on the two risky claims at time  $t + 1$ :  $\Sigma_t = \text{VarCov}(D_{t+1})$ .  $\sigma_{\ell,t}^2$  is the conditional variance of the payments of the debt of levered intermediaries:  $\sigma_{\ell,t} = \text{Var}_t(d_{t+1}^\ell)$ .  $\Sigma_t^\ell = (\sigma_{IR,\ell,t}, \sigma_{CR,\ell,t})'$  is the vector of conditional covariance between the payments on the two risky claims and the payments of the debt of levered intermediaries:  $\sigma_{IR,\ell,t} = \text{Cov}_t(d_{t+1}^{IR}, d_{t+1}^\ell)$  and  $\sigma_{CR,\ell,t} = \text{Cov}_t(d_{t+1}^{CR}, d_{t+1}^\ell)$ .

The optimality conditions for the levered intermediary is:

$$P_t = E [M_{t+1}^a (P_{t+1} + D_{t+1})] - \underbrace{\psi \Sigma_t X_{L,t}}_{\text{Regulatory Cost}} + \underbrace{\psi \bar{\ell} \Sigma_t^\ell}_{\text{Hedging Benefit}} + \underbrace{\mu_t \Omega P_t}_{\text{Collateral Benefit}} \quad (22)$$

The equation shows that the levered intermediary trade-off the marginal increase in regulatory cost by increasing the holdings of risky assets with the hedging and collateral benefits. The optimality condition of the asset managers are instead:

$$P_t = E [M_{t+1}^a (P_{t+1} + D_{t+1})]. \quad (23)$$

The asset managers does not have any regulatory concern and therefore only invest in the assets because of their future payoff. Using Equation 22 and 23, we find the optimal allocation of the levered intermediary:

$$X_{L,t+1} = \underbrace{\bar{\ell} \Sigma^{-1} \Sigma^\ell}_{\text{Hedging Term}} + \underbrace{\frac{\mu^B}{\psi} \Sigma^{-1} \Omega P}_{\text{Collateral Term}}. \quad (24)$$

**Allocation to IR** Using the orthogonality of the two factors, we could write the allocations to interest rate risk:

$$X_{I,t+1}^{IR} = \underbrace{\bar{\ell} \frac{\sigma_{IR,\ell,t}}{\sigma_{IR,t}^2}}_{\text{Hedging Term}} + \underbrace{\frac{\mu^B}{\psi} \frac{\omega^{IR}}{\sigma_{IR,t}^2} P}_{\text{Collateral Term}}$$

The first term, on the right hand side, is the hedging term. It states that the intermediaries tilts their portfolio to assets with higher covariance with their liabilities, per unit of squared risk. The first term is particularly relevant for insurance companies and pension funds. These intermediaries have long-term liability, which are marked-to-market using the risk-free swap curve. That negatively expose their balance sheet to interest rate risk. The hedging motive of insurance companies and pension funds explains the allocation to safe long-term bonds.

The second term, is the collateral term. Levered intermediaries are able to take leverage using assets as collateral. However, it is more costly for them to hold risk (due to the regulatory cost). They then allocate more of their portfolio to assets which provide better collateral per unit of squared risk. The higher the  $\omega$ , the higher the allocation to the risk factors.

**Allocation to CR** The allocation to credit risk is instead defined by:

$$X_{I,t+1}^{CR} = \underbrace{\frac{\mu^B}{\psi} \frac{\omega^{CR}}{\sigma_{CR}^2}}_{\text{Collateral Term}} P.$$

Levered intermediaries allocate their portfolio to credit risk insofar as it provides collateral value. The credit risk, instead does not provide any hedging value.

Figure 10 plots the allocation across the two types of risks of asset managers and levered intermediaries.<sup>18</sup> The figures shows that levered intermediary holds a larger share of interest rate risk while mutual funds hold larger share of credit risk.

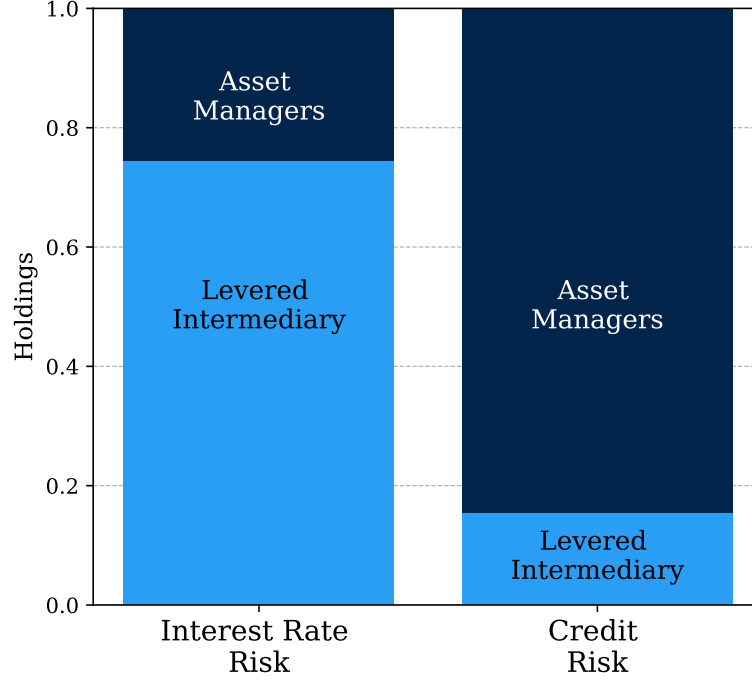
These results would suggest that—because of the different structure of their liabilities—asset managers and levered intermediaries choose to be exposed to different risk. Levered intermediaries, such as ICPFs, select into assets with a high interest rate risk exposure to match their long-duration liabilities. For asset managers, this liability-matching motive is absent, and they instead select into securities carrying credit risk as they do not face any regulatory cost in doing so. The wealth allocation decision of the household investing through the the different intermediaries is motivated by the trade-off between the ability to take leverage and the cost of regulation. They, therefore, allocate part of their wealth to levered intermediaries because of their ability to issue debt and the remainder to mutual funds because of their relatively lighter regulation on risk-taking.

A key measure for our allocation is the hedging term:  $\frac{\sigma_{IR,\ell,t}}{\sigma_{IR,t}^2}$ . In our model, the IR claim and the CR claim are exposed to only the IR risk factor and CR risk factor, respectively. However, this measure can be empirically estimate for all bonds. In Appendix E.2, we discuss how to empirically estimate this measure and show how bonds with low credit risk provide inferior hedging for the liability of levered intermediaries.

---

<sup>18</sup>The figure shows the allocations  $(\tilde{X}_t^L, \tilde{X}_t^M)$ , for an economy where  $\{X_{CB,t}, b_t^{CB}\}_{t=0}^\infty = 0$  and for the value of shocks equal to 0:  $f_t^{IR} = f_t^{CR} = 0$

Figure 10: Allocation Across Risky Claims



*Note:* The figure shows the allocation across risky claims of the two type of intermediary. The left bar plots the shares of the IR claim while the right bar plots the shares of the CR claim. The figure is based on the parametrization describe in Section 5.1.9.

### 5.2.2 Asset Pricing

We now discuss the asset pricing implications of the model for the risk premia on the two claims. We define the risk premia as:

$$E_t \left[ R_{t+1}^i - R_t^f \right],$$

for  $i = IR, CR$ , where

$$R_{t+1}^i = \frac{P_{t+1}^i + d_{t+1}^i}{P_t^i},$$

and

$$R_t^f = E_t \left[ M_{t+1}^e \right].$$

In the model, as both intermediaries are held by the active investors and use their SDF to price the assets, the risk premium is determined by the covariance between  $C_t^a$  and  $V_{t+1}^a$  with the returns on the two risky claims. Furthermore, combining the balance sheets of the entrepreneur with the

balance sheets of intermediaries, we have:

$$\begin{aligned} C_t^a &= \pi_t^M + \pi_t^I + T_t^a + \frac{\psi}{2} \text{Var}(\tilde{\pi}_{t+1}^L) \\ &= \underbrace{\theta^{CR} d_t^{CR}}_{\text{CR Risk}} + \underbrace{\theta^{IR} d_t^{IR} - \ell_t d_t^\ell}_{\text{IR Risk}}. \end{aligned}$$

The consumption path of the entrepreneur depends on the total supply of credit risk and the *net* supply of interest rate risk. The larger the  $\beta^\ell$  of the debt of the levered intermediaries, the smaller the risk-premium on interest rate risk. The model highlights the relevance of the liability of the intermediary for the differential pricing of interest rate risk vis-a-vis credit risk. A large levered intermediary sector (i.e., a large insurance corporations and pension fund sector) reduces the exposure of the intermediary shareholders to interest rate risk and lower risk premium. While the intermediary sector (and hence the intermediary shareholders) has to absorb the supply of interest rate risk, is able to offload some of the risk to the passive households. This result also explains the findings of [Greenwood and Vissing-Jorgensen \(2018\)](#), which documents the effect of pension and insurance company assets on the long end of the yield curve.

### 5.2.3 Transmission of Monetary Policy Through Different Balance Sheets

The central bank portfolio  $X_{CB,t}, b_t^{CB}$  alters the equilibrium allocation and asset prices. The model predicts that—as heterogeneous intermediaries have different allocations—the monetary policy transmission depends on the policy package.

The total net-supply of assets available to intermediaries is now:

$$X_{L,t} + X_{M,t} = \Theta - X_{CB,t}.$$

The sorting into different markets will heterogeneously expose intermediaries to different policy packages. If the Central bank mainly absorbs interest rate risk, this will i) imply a reduction of interest rate risk exposure of levered intermediaries. If the Central bank absorbs credit risk, this mainly affects the balance sheets of mutual funds.

To provide further insight into the mechanism of the model, we study the equilibrium allocation and asset pricing in an economy with a time-invariant central bank allocation  $\bar{X}_{CB}, \bar{b}^{CB}$  and an invariant price for  $q^{CB} = \beta^a$ . Given the benchmark economy price:  $\tilde{P}$ , we define central banks policy such that:

$$\tilde{P}' \bar{X}_{CB} = \bar{q}^{CB} \bar{b}^{CB}.$$

One we consolidate balance sheets, the consumption path of the active households is:

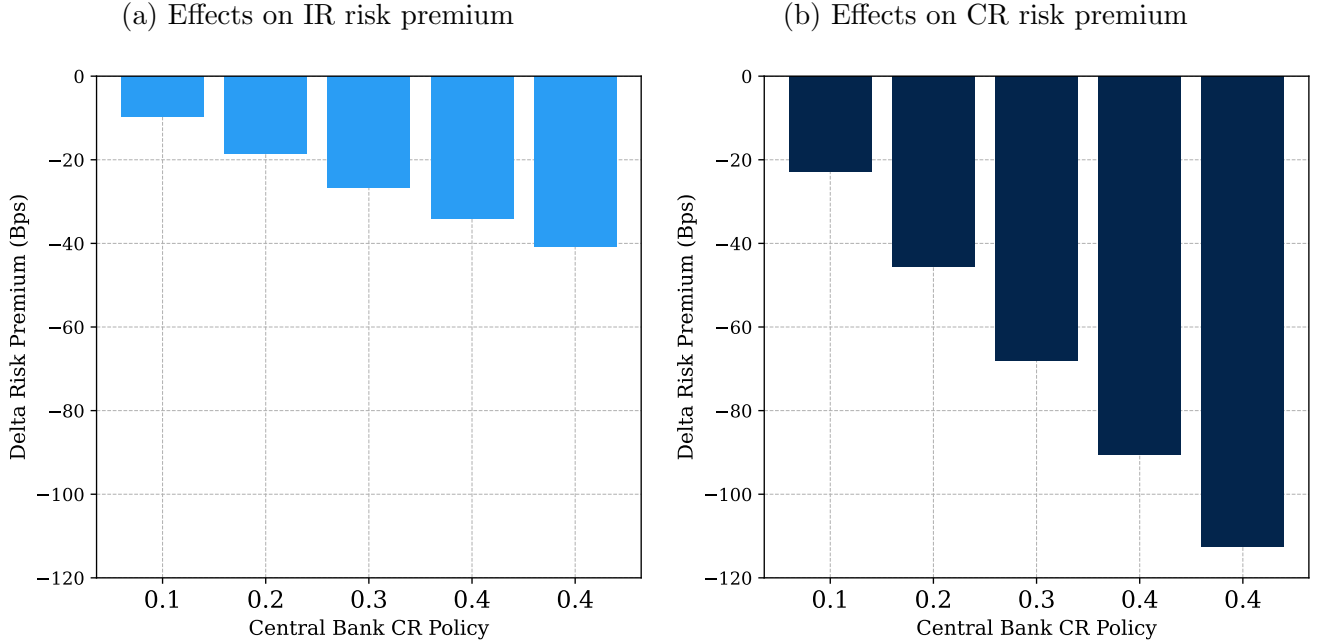
$$C_t^a = \underbrace{(\theta^{CR} - \varphi \bar{x}_{CR}^{CB}) d_t^{CR}}_{\text{CR Risk}} + \underbrace{(\theta^{IR} - \varphi \bar{x}_{CR}^{CB}) d_t^{IR} - \bar{\ell} d_t^\ell}_{\text{IR Risk}} + \underbrace{\varphi \bar{b}^G (1 - \bar{q}^{CB})}_{\text{Risk-free}}.$$

Central bank policy affects the exposure to the risk of the active households by affecting the riskiness of intermediaries' balance sheets. If the central bank buys credit risk, it reduces the net credit risk exposure and, in turn, the risk-premium commanded on credit risk. By contrast, if the central bank buys interest rate risk, it reduces the net interest rate risk exposure, and in turn, the risk-premium commanded on interest rate risk.

The different effects of different policies are a crucial element in our model. Depending on their balance sheet policy, the central bank can heterogeneously affect different risk premia. We let intermediaries endogenously sort into different risks in the model without any exogenous segmentation.

Figure 11 shows the effects for different central bank CR holdings. The effects on the credit risk premium (Figure 11a) is stronger than the effects on interest rate risk premium (Figure 11b).

Figure 11: Holding shares and Bond Sensitivity to TS Shocks



*Note:* The figure plots the effects on risk premia due to different central bank policies. We estimate the counterfactual difference in risk premia for different amount of CR claim absorbed by the central bank: 0.1, 0.2, 0.3, 0.4. Figure 11a plots the change in IR risk premium while Figure 11b plots the change in the CR risk premium. The figure is based on the parametrization describe in Section 5.1.9.

To sum up, the model shows that levered intermediaries choose to be exposed to interest rate risk. Central bank policies that absorb interest rate risk, such as (safe) government bond



purchases, mainly affect the market price of interest rate risk and, therefore, the balance sheet of levered intermediaries. On the other hand, asset managers choose to be exposed to credit risk; a central bank policy that absorbs credit risk, such as corporate bond purchases, mainly affects the balance sheet of asset managers.

#### 5.2.4 Redistributive Implications

The model also shows that central banks' policies have implications for passive households, who do not actively invest in risky claims. The consumption of the passive household is:

$$\begin{aligned}
C_t^p &= y^p + \bar{\ell}d_t^\ell + \underbrace{T_t^p}_{\text{Taxes/Rebate}} \\
&= y^p + \underbrace{\varphi \bar{X}_{CR,t}^G d_t^{CR}}_{\text{CR Risk}} + \underbrace{\varphi \bar{X}_{IR}^G + \bar{\ell}d_t^\ell}_{\text{IR Risk}} - \underbrace{(1 - \varphi)\bar{b}^G(1 - \bar{b}^G)}_{\text{Risk-free}}.
\end{aligned}$$

Central bank policies alter the consumption path of households by exposing them to credit risk and additional interest rate risk. The balance sheet of the central bank is risky, and the risk has to be borne by the household sector ultimately.

## 6. Risk Allocation and the Cross-Sectional effects of Monetary Policy Shocks

The fixed-income model, described in Section 5, explain why (i) different intermediaries choose different risk exposure and (ii) why different central bank policies (i.e., different risk absorption) have heterogeneous effects on risk premia. Individual bond returns have a factor structure:

$$R_{i,t+1} - R_t^f = \alpha_i + \beta_i^{IR}(R_{t+1}^{IR} - R_t^f) + \beta_i^{CR}(R_{t+1}^{CR} - R_t^f) + \varepsilon_{i,t+1}, \quad (25)$$

where  $R_{t+1}^{IR}$  are the returns on the interest rate portfolio and  $R_{t+1}^{CR}$  are the returns the credit risk portfolio. In this case, the risk premium on a each individual bond is:

$$E \left[ R_{i,t} - R_t^f \right] = \alpha_i + \beta_i^{IR} \underbrace{\lambda^{IR}}_{\text{Interest Rate Risk-Premium}} + \beta_i^{CR} \underbrace{\lambda^{CR}}_{\text{Credit Risk Risk Premium}}.$$

Central bank policies affect the risk premia on the two factors,  $\lambda^{IR}$  and  $\lambda^{CR}$ , while the cross-sectional effects are pinned down by  $\beta_i^{IR}$  and  $\beta_i^{CR}$ .

The correlation observed between intermediary holdings and the sensitivity of bonds to shocks

is due to (i) different monetary policy instruments have different transmissions: a credit risk absorption (e.g., through corporate sector purchases) mainly affects credit risk premium while an interest rate risk absorption (e.g., through government sector purchases) mainly affects interest rate risk premium (ii) a selection mechanism: levered intermediaries select into corporate bonds with high interest rate risk exposure (i.e., high  $\beta_i^{IR}$ ) while mutual funds select into corporate bonds with high credit risk exposure (i.e., high  $\beta_i^{CR}$ ) and, (iii) bonds with higher  $\beta_i^{CR}$  are more responsive to credit risk shocks while bonds with higher  $\beta_i^{IR}$  are more responsive to interest rate risk shocks.

In Section 3 we provided evidence on how different policies give rise to different shocks, in Section 6.1 we show how intermediaries select into the corporate bond market and in Section 6.2, we provide evidence on the cross-sectional effects of monetary policy shocks.

## 6.1 Risk Allocation

We rank bonds according to maturity and credit risk to provide empirical evidence on the sorting in the corporate bond market. We sample bonds into ten maturity deciles, where the 1<sup>st</sup> decile includes the shortest maturity bonds, and the 10<sup>th</sup> decile includes the longest maturity bonds.<sup>19</sup>

To measure credit risk, we use the yield-spread decile of the bond. The yield spreads should reflect a large set of observable characteristics that may contribute to the perceived riskiness of a bond (e.g., the rating, the sector, the country, and the liquidity). Moreover, a company may receive different credit ratings from rating agencies (Fitch Ratings, Moody's, and Standard & Poor's (S&P)). We hence sample bonds according to yield-spread decile (where the 1<sup>st</sup> decile are credit-safest bonds while the 10<sup>th</sup> are the credit-riskiest bonds).

For each group of bonds, we calculate the average share held by the different intermediary types by pooling data across the whole sample.<sup>20</sup>

Figure 12a reports the holding share in each risk-decile. The Figure clearly shows that mutual funds select into credit-riskier bonds. While mutual funds hold only 20% of the safest bonds, they hold almost 50% of the riskiest group. Insurance companies hold instead higher shares (around 35%) of the safest bonds, while their shares drop dramatically for the riskier deciles (they hold only 10% of the 10<sup>th</sup> risk decile).

Figure 19a in Appendix A also includes holdings by banks. Banks holds safe short-term maturity bonds. This explains the non-monotonic decrease in shares by insurance companies and pension funds in Figure 12a. In Appendix A, Figure 17 shows the sorting by different credit ratings.

By contrast, insurance companies and pension funds (ICPF) have larger share of long maturity

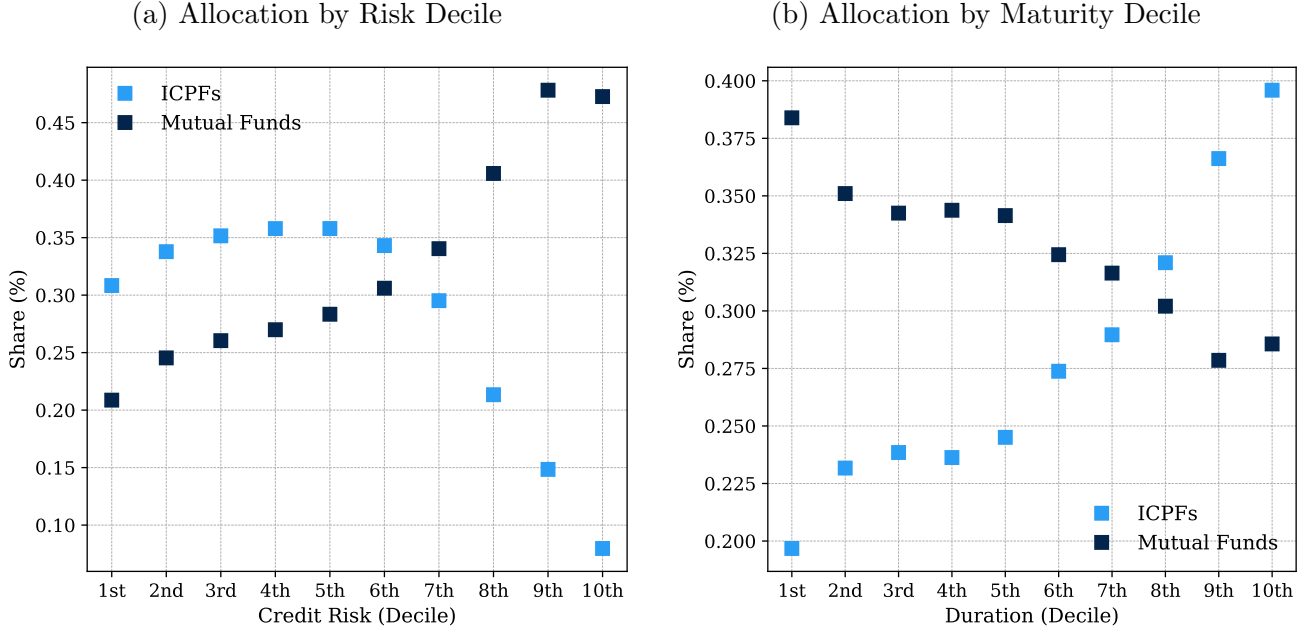
---

<sup>19</sup>The average maturity for each decile is: 1.5, 2.3, 3.0, 3.8, 4.5, 5.2, 6.1, 7.1, 8.6, 12.5. This exercise only considers bonds with a maturity greater than one year.

<sup>20</sup>We only include the sample of bonds available in iBoxx. The sample is, therefore the same used in our estimates in Section 4.

bonds, as displayed in Figure 12b. They hold roughly 20% of the shorter maturity bonds and 40% of the longer maturity bonds.

Figure 12: Allocation by Risk and Maturity



*Note:* The figure plots the average shares of insurance corporation and pension funds (ICPF) and mutual funds (MF) for different bond group. We divided bond based on their credit risk (Figure 12a) and maturity (Figure 12b). We only include bonds available both on iBoxx and on SHS and we estimate shares by pooling data across the full sample. The sample runs from 2014 to 2021. Source: Markit iBoxx and SHS. Source: Markit iBoxx and ECB SHSS.

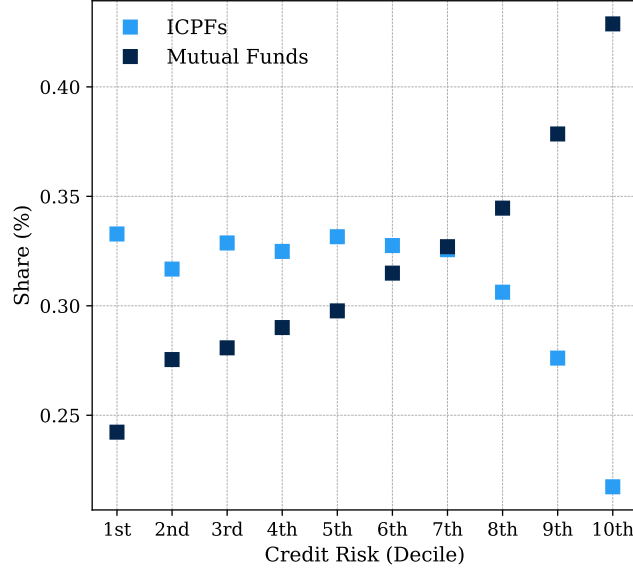
To further convey the importance of using the yield spread (and not only ratings, in Figure 13, we restrict our sample to corporate bonds with BBB ratings. We then rank to bonds according to credit risk decile (within the BBB rating sample), and compute the shares held by ICPFs and mutual funds. We find a striking sorting within the BBB-rated bonds. Mutual funds hold more of the riskier corporate bonds while insurance companies and pension funds hold more of the safer bonds.

### 6.1.1 Factor Allocation

To further illustrate the selection mechanism, in this section, we summarize the risk-exposure using two key risk factors of the corporate bond market: interest rate risk and credit risk. We use an empirical factor model to capture the risk exposure of each corporate bonds and, using the data on asset holdings, we match each corporate bond position to its risk exposure.

**Risk Factors** We measure two risk-factors for the corporate bond market. One risk-factor that measure interest rate risk (IR) and one risk-factor that measure credit risk (CR). We construct two market portfolios to replicate the returns of the two factors. The first portfolio is constructed

Figure 13: BBB Corporate Bonds: Allocation by Risk Decile



*Note:* In this figure, we restrict our sample to corporate bonds with BBB ratings. We then rank to bonds according to credit risk decile (within the BBB rating sample), where the 1<sup>st</sup> decile are credit-safest bonds while the 10<sup>th</sup> are the credit-riskiest bonds. Finally, we compute the shares held by ICPFs and mutual funds. We only include bonds available both on iBoxx and on SHS and we estimate shares by pooling data across the full sample. The sample runs from 2014 to 2021. Source: Markit iBoxx and SHS.

as the value-weighted return on the credit-risk-free component of our corporate bond index: the interest rate risk factor ( $R^{IR}$ ). The second risk-factor is constructed using the value-weighted excess return (vis-a-vis a credit-risk-free bonds with equal maturity) on all BBB corporate bonds: the credit risk factor ( $R^{CR}$ ). The return of any corporate bond can then be defined as in Equation 25.

**Risk Exposure** We use corporate bond characteristics (rating and maturity), to estimate the risk-exposure of a single corporate bonds. We run the following panel regression:

$$R_{i,t+1} = \beta_{Rat,j}^{IR} R_{t+1}^{IR} \times \mathbb{1}_j^{Rat} + \beta_{Mat,j}^{IR} R_{t+1}^{IR} \times \mathbb{1}_j^{Mat} + \beta_{Rat,j}^{CR} R_{t+1}^{CR} \times \mathbb{1}_j^{Rat} + \beta_{Mat,j}^{CR} R_{t+1}^{CR} \times \mathbb{1}_j^{Mat} \quad (26)$$

$$+ \text{Level Dummies} + \varepsilon_{i,t}, \quad (27)$$

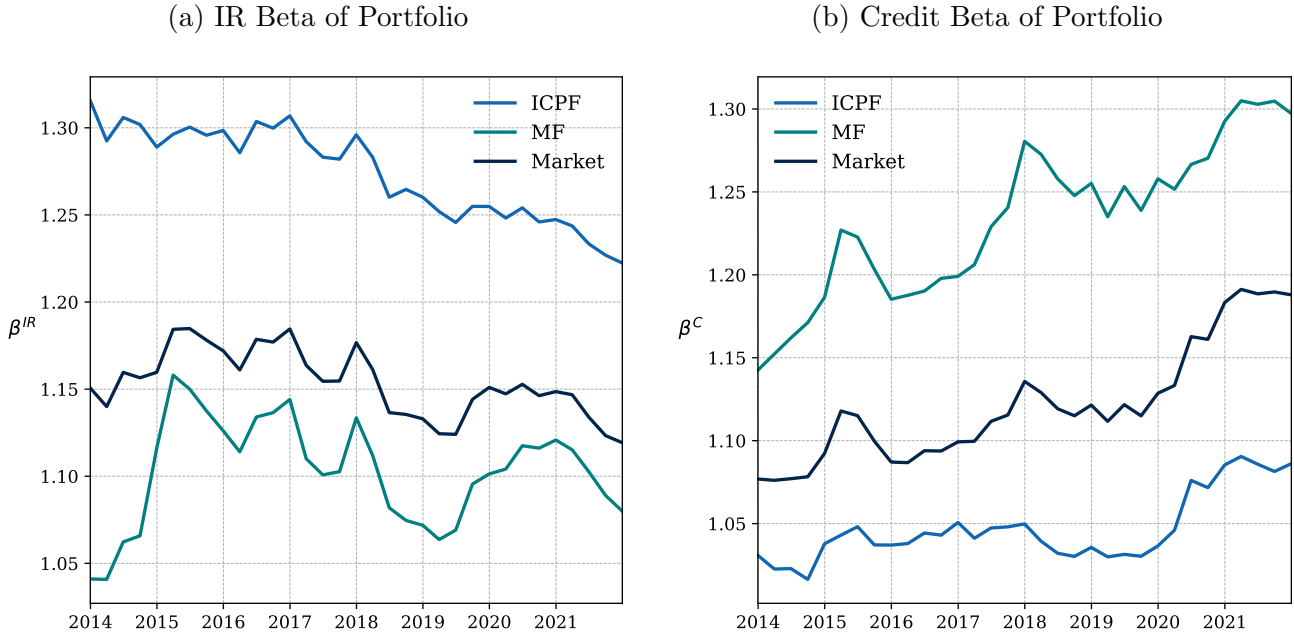
where  $\mathbb{1}_j^{Rat}$  is a dummy equal to 1 if the bond belongs to the  $j^{\text{th}}$  rating group and  $\mathbb{1}_j^{Mat}$  is a dummy equal to 1 if the bond belongs to the  $j^{\text{th}}$  maturity group. Level Dummies is the set of non-interacted dummies. The exposure to interest rate risk of a bond with rating of  $j$  and maturity  $j$  is given by the two estimated  $\beta_{Rat,j}^{IR}$  and  $\beta_{Mat,j}^{IR}$ . The exposure of the bond to credit risk is measured through the estimated  $\beta_{Rat,j}^{CR}$  and  $\beta_{Mat,j}^{CR}$ .

With our estimated  $\beta$  exposure for each corporate bond, we calculate the overall exposure

of the corporate bond portfolio of ICPF and mutual funds, as well the overall exposure of all private investors in the Euro area (the Market exposure). We compute the overall exposure for each quarter from 2013Q4 to 2021Q4.<sup>21</sup> We compute the risk exposure for the *market*, which consists of the whole Euro area (combining all sectors). We then also compute the exposure for the insurance corporation and pension fund sector (ICPFs) and the mutual fund sector (MF).

Figure 14 collect our results. The figure clearly shows that ICPF have high interest rate risk (IR) exposure and low credit risk (CR) exposure, compared to the market. By contrast mutual funds exhibit high exposure to credit risk and lower exposure to IR risk than the market.

Figure 14: Risk Exposure



*Note:* The figure plots the risk exposure to interest rate risk (Figure 14a) and credit risk (Figure 14b) for different sectors: the mutual fund sector (MF), the insurance corporations and pension funds sector (ICPFs) and the combined Euro area sectors (the market). We estimate the risk exposure for every single security by way of Equation 27 and use holdings data from SHS to aggregate exposures to the sector level. Sample: 2013-Q4 to 2021-Q4. Source: Markit iBoxx and SHS.

## 6.2 Bond Response to ECB Policies

We now further discuss the main source of heterogeneity in bonds yields sensitivity to credit shocks. We sort bonds according to yield spread and maturity decile.<sup>22</sup> However, we also report results by using the rating as a single measure of credit risk.

We then regress:

<sup>21</sup>We use  $\beta$  estimated over the entire sample. The time variation is given by the change in portfolio weights over time.

<sup>22</sup>The category are the same as discussed in Section 6.1

$$\Delta y_{s_{i,t}} = \gamma_i + \gamma_t + \beta_j^{Risk} MP_t^{CS} \times \mathbb{1}_j^{Risk} + \beta_j^{Mat} MP_t^{CS} \times \mathbb{1}_j^{Mat} + \gamma_j^{Risk} \mathbb{1}_j^{Risk} + \gamma_j^{Mat} \mathbb{1}_j^{Mat} + \varepsilon_{i,t}, \quad (28)$$

where  $\mathbb{1}_j^{Risk}$  is the dummy that indicates whether the bonds belong to the  $j^{\text{th}}$  risk bucket and  $\mathbb{1}_j^{Mat}$  is the dummy that indicates whether the bonds belong to the  $j^{\text{th}}$  maturity bucket.  $\gamma_i$  and  $\gamma_t$  are bond and time fixed-effects.

The coefficients of interests are:  $\beta_j^{Risk}$  and  $\beta_j^{Mat}$ , which measure how the sensitivity of corporate bonds vary with credit-risk and maturity, respectively.

Figure 15 plots the estimated coefficients from Equation 28 together with the confidence bands.<sup>23</sup> The coefficients are estimated relative to the first riskiness and maturity group. Figure 15a displayed the estimated  $\beta_j^{Risk}$ . The figure clearly shows that the sensitivity to credit spread shocks increase with risk. Figure 15b display the estimated  $\beta_j^{Mat}$  together with their confidence bands. The figure shows that for short to medium maturity the response of bonds is increasing. However, for longer-maturity bonds (with maturity above 4.5) it seems the sensitivity is lower.

Figure 24 shows the results for Equation 28 where we used ratings instead of credit-risk deciles. The results in Figure 24a are broadly in line with our previous results.

The combination of the results in Figure 15a and Figure 24b suggest that the sensitivity of corporate bonds to credit spread shocks is increasing with the risk of the bonds. Within a credit rating, it seems that the term-structure of credit risk is increasing: longer-maturity bonds have higher credit spread than shorter maturity bonds. Therefore, if we capture credit risk through the priced yield-spread—which should take into account all characteristics contributing to credit risk—the results suggest the maturity of the bonds is less important in determining the sensitivity of bonds to ECB credit spread shocks.

### 6.2.1 The Effects on Non-Investment Grade Bonds

Figure 4 clearly shows that credit spreads have been reduced by the action of the central bank. However, as discussed, the ECB only purchased investment-grade corporate bonds. Their purchases have not been extended to sub-investment grade bonds. Haddad et al. (2021) showed that in the US, they did not observe any effects on sub-investment grade bonds, consistently with the *narrow* channel of QE.

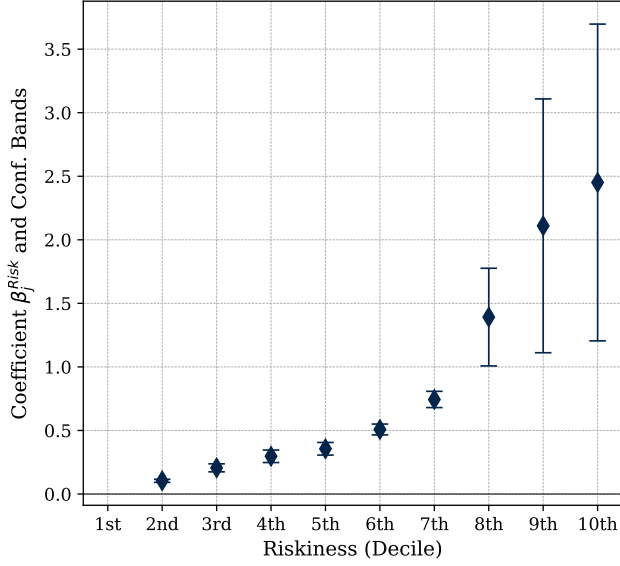
The ECB only bought corporate bonds whose best rating (S&P, Fitch, Moody's, and DBRS) is at least BBB-. Using our micro data, we compute the change in yields on the week of the CSPP announcement<sup>24</sup> and estimate the average change by rating group. We also further split

<sup>23</sup>Confidence bands are based on standard errors clustered at the ISIN-Time level.

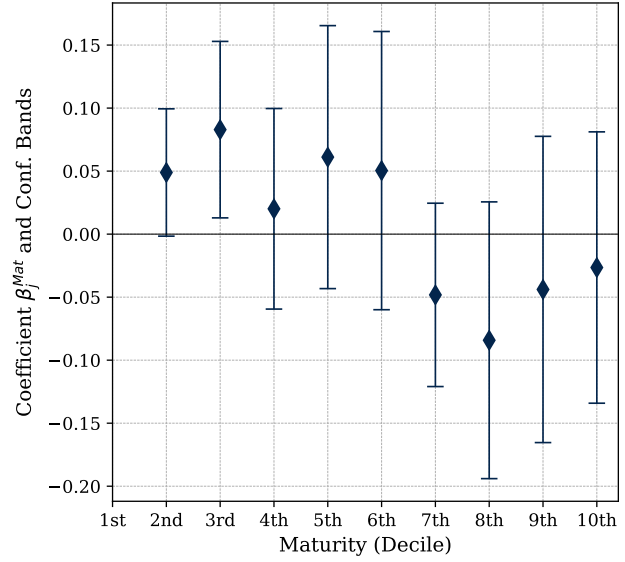
<sup>24</sup>To make sure our results are not driven by noise, we take a number of step. We compute the average yield for each bond one week after the announcement of CSPP and measure the difference with respect two days before the announcement. We drop outliers and compute the average change for each rating bucket.

Figure 15: Sensitivity to  $MP^{CS}$  by Credit Risk and Maturity

(a) Sensitivity to  $MP^{CS}$  by Credit Risk Decile



(b) Sensitivity to  $MP^{CS}$  Shocks by Maturity Decile



*Note:* The figure shows the estimated coefficient from Equation 28. Figure 15a displayed the estimated  $\beta_j^{Risk}$  together with their confidence bands. Figure 15b display the estimated  $\beta_j^{Mat}$  together with their confidence bands. The sample runs from 2014 through 2021. The coefficients are estimated relative to the first riskiness and maturity group (for this reason, the coefficient for the first group are missing). Source: Markit iBoxx.

the sample in short maturity bonds (maturity less or equal to four years) and long-maturity bonds. The results are displayed in Figure 16a: every point in the figure is the estimated change in yield for the corresponding rating-maturity group. Figure 16a suggests that bonds beyond the ECB cutoff (BB+ or worse) were notably affected by the ECB policy. We also note that, among the investment grade bonds, long-maturity bonds saw a larger compression in yields than short-maturity (as also discussed by Todorov (2020)). However, according to Figure 16a seems that the riskiness of the bonds explains most of the cross-section variation. In fact, in contrast to the US experience, our results suggest that high-yields bonds, seemed to be those that benefited the most by the ECB credit policy.

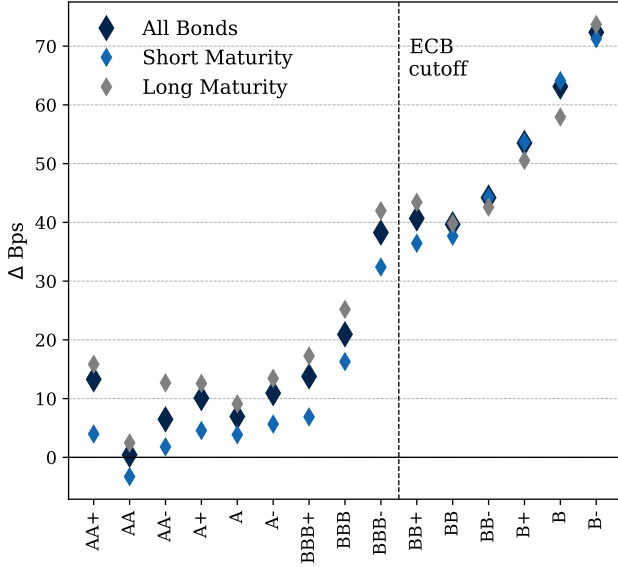
The bond yields behavior around the PEPP announcement also corroborates this conclusion. In Figure 16b we plot the evolution of bond yields in the first semester of 2020. The high-yields bonds, in orange colors, spiked in March 2020 to then retract after the announcement of the massive bond intervention of the ECB. The size and speed of the fall in bond yields by high-yield bonds is unprecedented. Although the effects on investment-grade bonds is also notable—as the PEPP managed to stop the rising cost of bond funding all across the board—the pandemic experience also confirmed the potent effects of ECB actions on high-yields bonds.

Figure 16 is based on micro-data on bond prices from Markit-iBoxx. We also replicate the results using a different datasource: aggregate option-adjusted-yield spread from the Bloomberg-

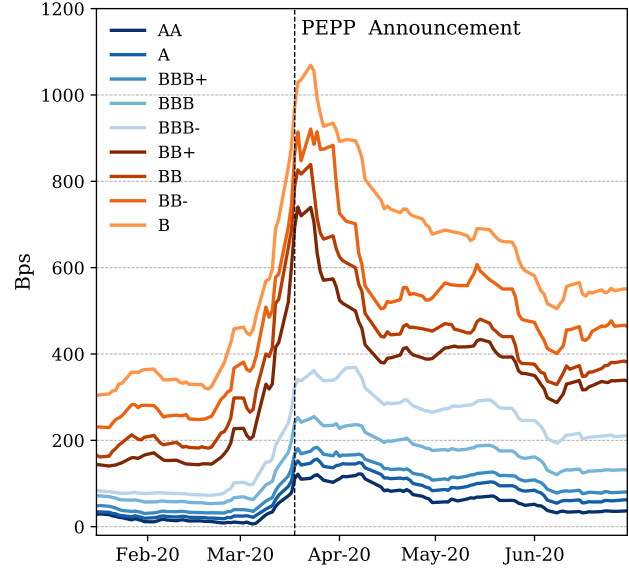
Barclays index. The results, reported in Figure 21 in Appendix B, are virtually identical, both for the CSPP as well as the PEPP announcements.

Figure 16: CSPP and PEPP announcement - by Rating

(a) Response to CSPP Announcement - by Rating



(b) PEPP Announcement - by Ratings



*Note:* The figure shows the change in bond yields around ECB CSPP and PEPP announcement. We compute the average yield for each rating group (and maturity group). Figure 16a shows the change in yields in the week of the CSPP announcement. Figure 16b shows the evolution of bond yields around the PEPP announcement. Source: Markit iBoxx.

## 7. Conclusion

This paper studies the transmission of ECB monetary policy to the Euro area corporate bond market. We construct two monetary policy shock series and show that their transmission depends on the type of policy instrument used by the central bank. The first shock, the credit spread monetary policy shock is measured using the change in corporate bond yield spread around ECB announcement. The second shock, the term spread monetary policy shock, is measured as the change in term spread (10 year OIS minus 2 year OIS) around ECB announcements.

We show that different policy instrument transmit through different shocks. A corporate bond purchases mainly affect the credit spread shock. A (safe) government bond purchases mainly affect the term spread shock.

With the two shocks at hand, we measure the cross-sectional effects of monetary policy on the corporate bond market. Based on granular information from the ECB Security Holdings Statistics, we document that the yield on bonds held by different intermediaries respond heterogeneously to the two shocks: bonds held by mutual funds exhibit a higher sensitivity to the credit spread



shocks, while bonds held by insurance companies and pension funds (ICPFs) are more sensitive to the term spread shocks.

We also show that these correlation mainly arise because of different intermediaries choose to be exposed to different risks: ICPFs select into assets with a high interest rate risk exposure (credit safe, long maturity bonds) while mutual funds select instead into securities carrying credit risk (credit risky bonds).

We explain these findings through the lens of a model of the fixed income market with multiple risk factors and heterogeneous intermediaries. Levered intermediaries, such as ICPFs, select into assets with a high interest rate risk exposure to match their long-duration liabilities. For mutual funds, this liability-matching motive is absent, and they instead select into securities carrying credit risk.

Different policy tools heterogeneously affect the market prices of those factors. In our model, the transmission channel of central bank purchases is through risk-shifting: the central bank, by purchasing financial assets, extracts risks from the market, previously held by private intermediaries. In turn, this affects the risk exposure of intermediaries shareholders—a subset of the household sector—and hence alter the market price of risk. Different monetary policy instruments absorb different types of risks and therefore have different effects on risk premia. Our results have important implications for the transmission of monetary policy, financial stability and redistribution.

## References

- ABIDI, N. AND I. MIQUEL-FLORES (2018): “Who benefits from the corporate QE? A regression discontinuity design approach,” Working Paper Series 2145, European Central Bank.
- ALTAVILLA, C., L. BRUGNOLINI, R. S. GÜRKAYNAK, R. MOTTO, AND G. RAGUSA (2019): “Measuring euro area monetary policy,” *Journal of Monetary Economics*, 108, 162–179.
- ALTAVILLA, C., G. CARBONI, AND R. MOTTO (2015): “Asset purchase programmes and financial markets: lessons from the euro area,” .
- ANDERSON, G. AND A. CESA-BIANCHI (2020): “Crossing the credit channel: credit spreads and firm heterogeneity,” Bank of England working papers 854, Bank of England.
- BAI, J., T. G. BALI, AND Q. WEN (2019): “Common risk factors in the cross-section of corporate bond returns,” *Journal of Financial Economics*, 131, 619–642.
- BEGENAU, J., M. PIAZZESI, AND M. SCHNEIDER (2015): “Banks’ Risk Exposures,” NBER Working Papers 21334, National Bureau of Economic Research, Inc.
- BERNANKE, B. S. AND K. N. KUTTNER (2005): “What explains the stock market’s reaction to Federal Reserve policy?” *The Journal of finance*, 60, 1221–1257.
- BRETSCHER, L., L. SCHMID, I. SEN, AND V. SHARMA (2021): “Institutional Corporate Bond Pricing,” *Swiss Finance Institute Research Paper*.
- BRUNNERMEIER, M. K. AND Y. SANNIKOV (2014): “A macroeconomic model with a financial sector,” *American Economic Review*, 104, 379–421.
- CARBONI, G. AND M. ELLISON (2022): “Preferred habitat and monetary policy through the looking-glass,” .
- COCHRANE, J. H. AND M. PIAZZESI (2002): “The Fed and Interest Rates - A High-Frequency Identification,” *American Economic Review*, 92, 90–95.
- COPPOLA, A. (2021): “In Safe Hands: The Financial and Real Impact of Investor Composition Over the Credit Cycle,” *Working Paper*.
- DARMOUNI, O. AND M. PAPOUTSI (2022): “The rise of bond financing in europe,” .
- DE SANTIS, R. A., A. GEIS, A. JUSKAITE, AND L. V. CRUZ (2018): “The impact of the corporate sector purchase programme on corporate bond markets and the financing of euro area non-financial corporations,” *Economic Bulletin Articles*, 3.

- DE SANTIS, R. A. AND A. ZAGHINI (2021): “Unconventional monetary policy and corporate bond issuance,” *European Economic Review*, 135, 103727.
- DOMANSKI, D., H. S. SHIN, AND V. SUSHKO (2017): “The hunt for duration: not waving but drowning?” *IMF Economic Review*, 65, 113–153.
- DOYLE, N., L. HERMANS, P. A. MOLITOR, AND C. WEISTROFFER (2016): “Shadow banking in the euro area: risks and vulnerabilities in the investment fund sector,” *ECB Occasional Paper*.
- ESER, F., W. LEMKE, K. NYHOLM, S. RADDE, AND A. VLADU (2019): “Tracing the impact of the ECB’s asset purchase programme on the yield curve,” .
- FAIA, E., J. SALOMAO, AND A. V. VEGHAZY (2022): “Granular investors and international bond prices: Scarcity-induced safety,” *Available at SSRN*.
- FAMA, E. F. AND K. R. FRENCH (1993): “Common risk factors in the returns on stocks and bonds,” *Journal of financial economics*, 33, 3–56.
- GAGNON, J., M. RASKIN, J. REMACHE, AND B. SACK (2011): “The Financial Market Effects of the Federal Reserve’s Large-Scale Asset Purchases,” *International Journal of Central Banking*, 7, 3–43.
- GERTLER, M. AND P. KARADI (2011): “A model of unconventional monetary policy,” *Journal of Monetary Economics*, 58, 17–34, *carnegie-Rochester Conference Series on Public Policy: The Future of Central Banking* April 16-17, 2010.
- (2015): “Monetary Policy Surprises, Credit Costs, and Economic Activity,” *American Economic Journal: Macroeconomics*, 7, 44–76.
- (2018): “Qe 1 vs. 2 vs. 3...: A framework for analyzing large-scale asset purchases as a monetary policy tool,” *29th issue (January 2013) of the International Journal of Central Banking*.
- GILCHRIST, S., D. LÓPEZ-SALIDO, AND E. ZAKRAJŠEK (2015): “Monetary Policy and Real Borrowing Costs at the Zero Lower Bound,” *American Economic Journal: Macroeconomics*, 7, 77–109.
- GILCHRIST, S. AND E. ZAKRAJŠEK (2012): “Credit Spreads and Business Cycle Fluctuations,” *American Economic Review*, 102, 1692–1720.
- GREENWOOD, R. M. AND A. VISSING-JORGENSEN (2018): “The impact of pensions and insurance on global yield curves,” *Harvard Business School Finance Working Paper*, 19–59.

- GROSSE-RUESCHKAMP, B., S. STEFFEN, AND D. STREITZ (2019): “A capital structure channel of monetary policy,” *Journal of Financial Economics*, 133, 357–378.
- GÜRKAYNAK, R. S., B. P. SACK, AND E. T. SWANSON (2004): “Do actions speak louder than words? The response of asset prices to monetary policy actions and statements,” *The Response of Asset Prices to Monetary Policy Actions and Statements (November 2004)*.
- HADDAD, V., A. MOREIRA, AND T. MUIR (2021): “When Selling Becomes Viral: Disruptions in Debt Markets in the COVID-19 Crisis and the Fed’s Response,” *The Review of Financial Studies*, 34, 5309–5351.
- HE, Z. AND A. KRISHNAMURTHY (2013): “Intermediary Asset Pricing,” *American Economic Review*, 103, 732–70.
- JAROCIŃSKI, M. (2021): “Estimating the Fed’s Unconventional Policy Shocks,” Working Paper Series 20210, European Central Bank.
- JAROCIŃSKI, M. AND P. KARADI (2020): “Deconstructing Monetary Policy Surprises—The Role of Information Shocks,” *American Economic Journal: Macroeconomics*, 12, 1–43.
- KEKRE, R. AND M. LENEL (2021): “Monetary policy, redistribution, and risk premia,” Tech. rep., National Bureau of Economic Research.
- KELLY, B. T., D. PALHARES, AND S. PRUITT (2020): “Modeling corporate bond returns,” *Available at SSRN 3720789*.
- KOIJEN, R. AND M. YOGO (2022): “Understanding the Ownership Structure of Corporate Bonds,” *Working Paper*.
- KOIJEN, R. S., F. KOULISCHER, B. NGUYEN, AND M. YOGO (2021): “Inspecting the mechanism of quantitative easing in the euro area,” *Journal of Financial Economics*, 140, 1–20.
- KOIJEN, R. S. J. AND M. YOGO (2019): “A Demand System Approach to Asset Pricing,” *Journal of Political Economy*, 127, 1475–1515.
- KRISHNAMURTHY, A. (2022): “Lessons for QE Policy from Research,” Panel: Shrinking the federal reserve balance sheet.
- KRISHNAMURTHY, A., S. NAGEL, AND A. VISSING-JORGENSEN (2017): “ECB Policies Involving Government Bond Purchases: Impact and Channels\*,” *Review of Finance*, 22, 1–44.
- KRISHNAMURTHY, A. AND A. VISSING-JORGENSEN (2011): “The Effects of Quantitative Easing on Interest Rates: Channels and Implications for Policy,” Working Paper 17555, National Bureau of Economic Research.

- LEOMBRONI, M., A. VEDOLIN, G. VENTER, AND P. WHELAN (2021): “Central bank communication and the yield curve,” *Journal of Financial Economics*, 141, 860–880.
- LITTERMAN, R. AND J. SCHEINKMAN (1991): “Common factors affecting bond returns,” *Journal of fixed income*, 1, 54–61.
- NAKAMURA, E. AND J. STEINSSON (2018): “High-frequency identification of monetary non-neutrality: the information effect,” *The Quarterly Journal of Economics*, 133, 1283–1330.
- OECD (2011): “Pensions at a Glance,” Tech. rep.
- OZDAGLI, A. K. AND Z. K. WANG (2022): “Interest Rates and Insurance Company Investment Behavior,” *Working Paper*.
- PAPOUTSI, M., M. PIAZZESI, AND M. SCHNEIDER (2021): “How unconventional is green monetary policy,” *JEEA-FBBVA Lecture at the ASSA (January)*.
- PIAZZESI, M. AND M. SCHNEIDER (2021): “Payments, credit and asset prices,” Working paper.
- SILVA, D. H. (2016): “The risk channel of unconventional monetary policy,” Tech. rep., working paper.
- SMOLYANSKY, M. AND G. A. SUAREZ (2021): “Monetary policy and the corporate bond market: How important is the Fed information effect?” Finance and Economics Discussion Series 2021-010, Board of Governors of the Federal Reserve System (U.S.).
- SWANSON, E. T. (2021): “Measuring the effects of federal reserve forward guidance and asset purchases on financial markets,” *Journal of Monetary Economics*, 118, 32–53.
- TODOROV, K. (2020): “Quantify the quantitative easing: Impact on bonds and corporate debt issuance,” *Journal of Financial Economics*, 135, 340–358.
- VAYANOS, D. AND J.-L. VILA (2021): “A Preferred-Habitat Model of the Term Structure of Interest Rates,” *Econometrica*, 89, 77–112.
- ZAGHINI, A. (2020): “How ECB purchases of corporate bonds helped reduce firms’ borrowing costs,” *Research Bulletin*, 66.

## A. The Euro Area Corporate Bond Market

We analyze the Euro area corporate bond market using the total amount of holdings by private investors in the Euro area. These data are based on information from the SHS database.

**Issuer Sector** Table IV reports data for the Euro area corporate bond market by issuer sector. The first column reports the total amount in Eur bilions of holdings by Euro area private sector investors. The second to fifth columns reports the holding shares by different types of investors. The same columns are also reported for 2013. Note that the NFC total amount does not include holdings by the central bank, which are large—as Figure 1b has shown. The MFI sector is still the largest issuer of corporate bonds but the total amount has been roughly stable since 2013. This contrasts with the notable increase in NFC corporate bonds which almost doubled in the past decade from Eur 559bns to Eur 941bns. We also observe a notable increase in bonds issued by captive institutions: it is often the case that captive institutions issue bonds to NFC, showing the increase relevance of corporate bonds for NFC.

Table IV: Euro Area Corporate Bond Market - Issuer

	2021					2013				
	Tot (bn)	ICPF	MF	MFI	Others	Tot (bn)	ICPF	MF	MFI	Others
MFI	1896	21.4	20.6	51.9	6.2	1903	25.1	17.0	46.6	11.4
NFC	941	42.0	43.4	8.2	6.3	559	46.5	32.9	10.1	10.6
Captive	407	33.4	48.0	9.4	9.3	4	46.0	44.5	6.9	2.6
OFI	575	25.2	28.8	40.8	5.3	939	26.3	24.2	39.1	10.4
ICPF	31	47.8	39.9	3.7	8.6	10	58.0	19.0	12.7	10.4
F. Auxiliaries	116	39.8	45.2	7.2	7.8	6	29.3	41.8	18.3	10.6

*Note:* Table IV reports data for the Euro area corporate bond market by issuer sector. The first column reports the total amount in Eur bilions of holdings by Euro area private sector investors. The second to fifth columns reports the holding shares by different types of investors. The same columns are also reported for 2013. Source: ECB SHSSs.

**Ratings** We show the allocation for bonds issued by NFC according to the ratings. We show data for NFC and for all issuer sector. The first column reports the shares for each rating within the aggregate Euro area portfolio. The second to fifth columns exhibit the shares held by the different rating agencies.

Table V: Euro Area Corporate Bond Market - Ratings

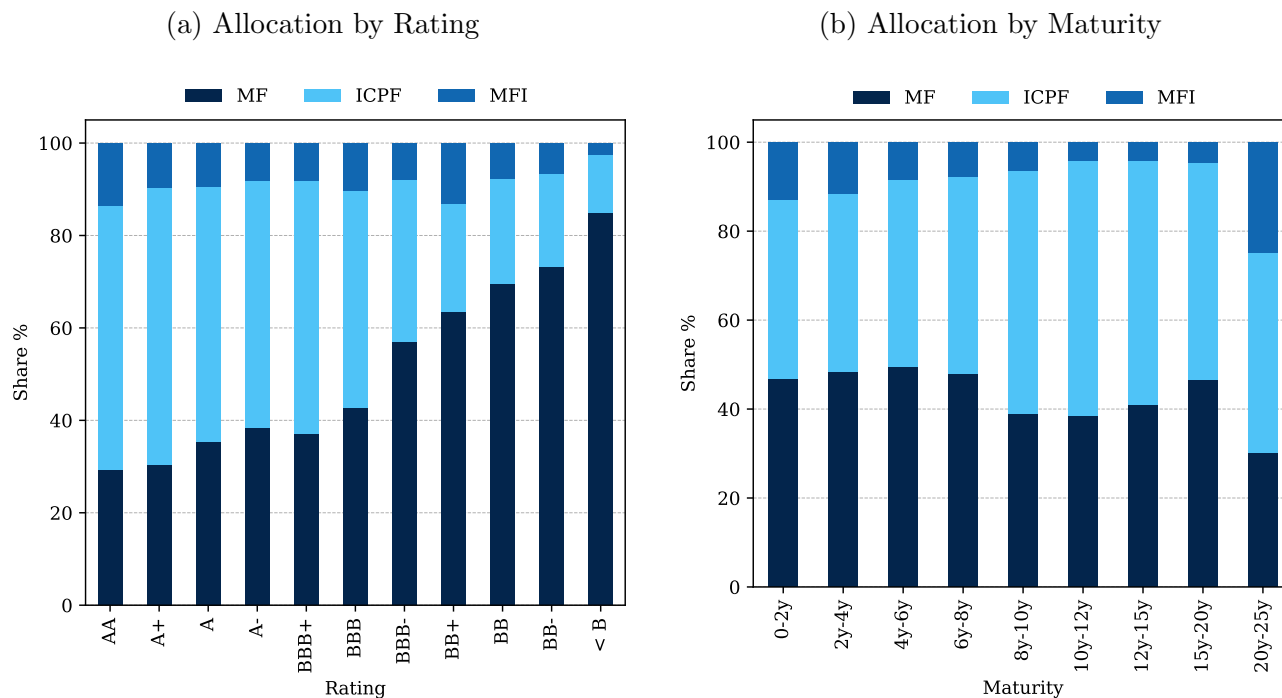
	NFC					All				
	Tot (Share)	ICPF	MF	MFI	Others	Tot (Share)	ICPF	MF	MFI	Others
AAA	1.1	59.2	25.2	10.5	5.0	26.1	16.8	15.2	62.1	5.9
AA+	0.5	51.1	18.7	27.9	2.3	5.0	13.1	10.8	73.2	3.0
AA	2.3	55.7	26.7	11.8	5.8	5.4	27.9	17.9	50.3	3.8
AA-	3.1	57.7	27.4	10.6	4.3	7.3	31.8	20.5	43.4	4.3
A+	5.5	54.2	34.4	6.7	4.7	8.0	40.1	32.1	21.6	6.1
A	7.3	49.9	34.2	10.2	5.8	8.2	37.8	35.3	20.0	6.9
A-	19.0	51.0	36.2	7.8	5.1	9.7	46.2	37.4	10.4	6.0
BBB+	22.4	49.5	38.4	8.1	4.0	11.2	39.6	38.6	16.2	5.5
BBB	14.6	40.2	44.8	9.7	5.3	7.8	34.8	46.2	12.2	6.9
BBB-	8.1	27.4	58.9	7.5	6.1	3.9	24.9	57.8	10.4	6.9
BB+	5.6	20.4	61.1	10.2	8.3	2.4	19.1	58.4	11.3	11.2
BB	3.7	17.4	67.9	5.1	9.6	1.3	16.2	66.0	7.9	10.0
BB-	1.7	12.3	65.4	5.8	16.5	0.8	11.0	62.8	8.9	17.3
B+	1.9	10.4	70.2	1.6	17.8	1.0	8.9	63.9	6.0	21.1
B	2.0	9.3	66.7	1.3	22.7	1.1	8.5	62.8	4.5	24.2
B-	0.7	5.9	61.8	0.9	31.5	0.4	5.5	67.6	2.0	24.9
< B	0.4	5.6	76.2	1.7	16.5	0.4	6.1	70.0	4.8	19.1

*Note:* The table shows the allocation for bonds issued by NFC according to the ratings. We show data for NFC and for all issuer sector. The first column reports the shares for each rating within the aggregate Euro area portfolio. The second to fifth columns exhibit the shares held by the different rating agencies. Source: ECB SHSS.

Of the total amount of NFC corporate bonds held by Euro area private investors, 30% are issued by countries outside the Euro area, while 70% are issued by Euro area countries.

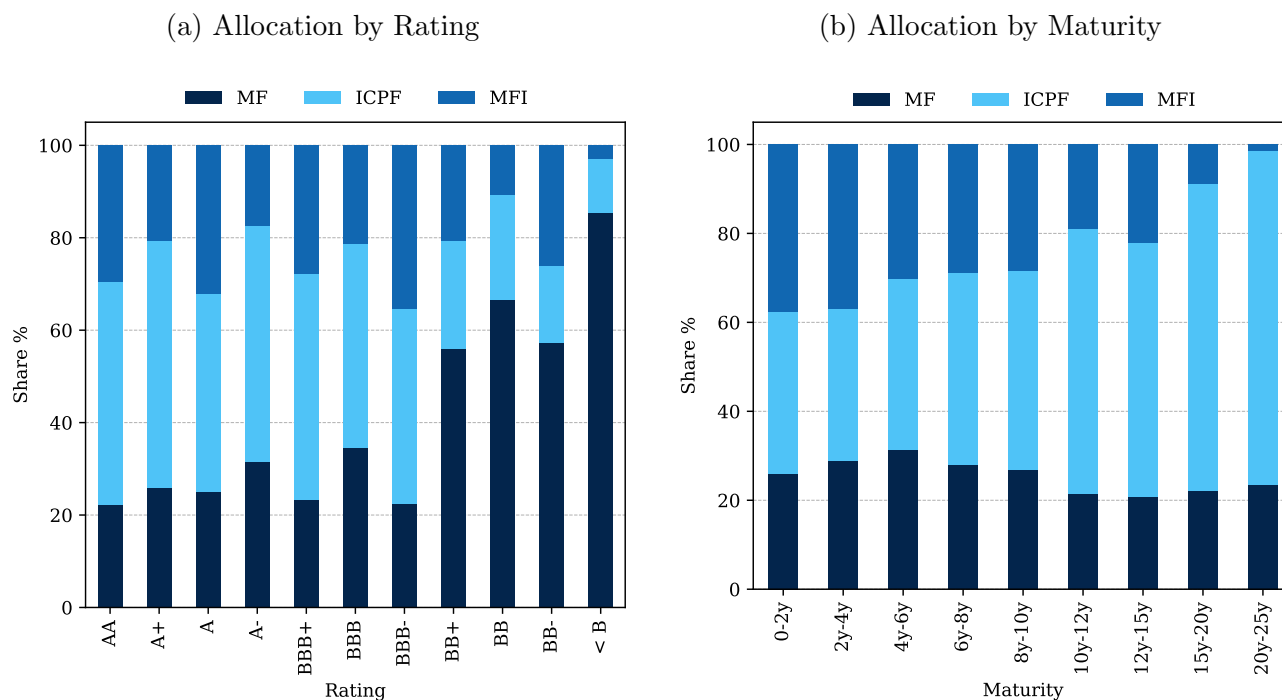
We then only select bonds issued by Government and NFC and show the allocation of mutual funds (MF), insurance corporations and pension funds and banks (MFI). Figure 17 plots the allocation of bonds issued by non-financial corporations. Figure 18 plots the allocation of total government and corporate bonds.

Figure 17: Allocation Corporate Bonds



Note: Figure 17 plots the allocation of bonds issued by non-financial corporations. Source: ECB SHSSs.

Figure 18: Allocation of All Bonds (Government and Corporate Bonds)

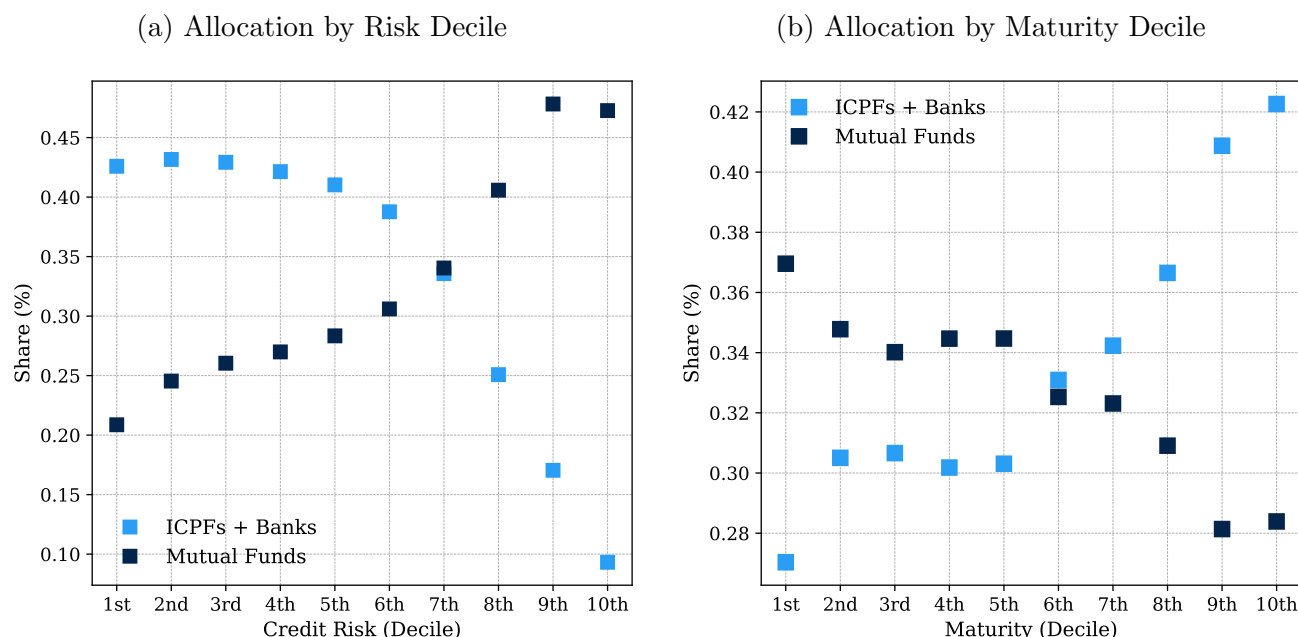


Note: Figure 18 plots the allocation of total government and corporate bonds. Source: ECB SHSSs.



## A.1 Allocation with MFI

Figure 19: Allocation by Risk and Maturity



*Note:* The figure plots the share held by different intermediaries. Source: Markit iBoxx, ECB SHSS.

## A.2 Financial Intermediaries

### A.2.1 Insurance Corporations

The total asset of insurance corporation in the Euro area accounts for 9.1tn at the end of 2021.

The insurance sector in the EU is supervised by the European Insurance and Occupational Pensions Authority (EIOPA) and is based on Solvency II, the prudential regime for insurance and reinsurance undertakings in the EU. Solvency II has a risk-based approach that enables to assess the “overall solvency” of insurance and reinsurance undertakings through quantitative and qualitative measures. The Solvency II framework set out qualitative and quantitative requirements for calculation of technical provisions and Solvency Capital Requirement (SCR).

The value of technical provision should be equal to the sum of best estimate of the liabilities and risk margin. The best estimate corresponds to the probability-weighted average of future cash-flows, discounted using the risk-free curve published by EIOPA. For Euro denominated liabilities, the risk-free curve was originally based on EONIA-OIS curve; it then transitioned to ESTR-OIS curve.<sup>25</sup> Technical provisions represent the current amount the (re)insurance company would have to pay for an immediate transfer of its obligations to a third party.

<sup>25</sup>See Appendix Data for further information on the risk-free curves.

The SCR is the capital required to ensure that the (re)insurance company will be able to meet its obligations over the next 12 months with a probability of at least 99.5%. In addition to the SCR capital a Minimum capital requirement (MCR) must be calculated which represents the threshold below which the national supervisor (regulator) would intervene. The MCR is intended to correspond to an 85% probability of adequacy over a one-year period and is bounded between 25% and 45% of the SCR.

For supervisory purposes, the SCR and MCR can be regarded as "soft" and "hard" floors respectively. That is, a regulatory ladder of intervention applies once the capital holding of the (re)insurance undertaking falls below the SCR, with the intervention becoming progressively more intense as the capital holding approaches the MCR.

The computation of SCR is based on the following risk modules: (a) non-life underwriting risk; (b) life underwriting risk; (c) health underwriting risk; (d) market risk; (e) counterparty default risk. Each of these risk modules is calibrated using a Value-at-Risk.

The capital requirement for market risk is by far the most important risk module.<sup>26</sup> The regulation stresses in particular that the market risk module "shall properly reflect the structural mismatch between assets and liabilities, in particular with respect to the duration".<sup>27</sup> Appendix A.2.2 includes detailed information on the market risk module.

The stress on the duration miss-match between asset and liabilities owes to the nature of the insurance business. Insurance companies tend to have long duration liabilities. According to EIOPA, in 2020, the median duration for life insurance companies was 13 (22 for the 90<sup>th</sup> percentile) and 4 for non-life insurance companies (8 for the 90<sup>th</sup> percentile).<sup>28</sup>

Insurance companies also provide so-called unit-linked and index-linked policies. The main feature of these policies is that the policyholder bears the risk and are therefore not included in the computation of the capital ratios. If we exclude such policies, the lion's share of insurance liabilities are life-insurance policies (89%) rather than non-life insurance policies (11%).<sup>29</sup>

## A.2.2 The market risk module

The market risk module<sup>30</sup> shall reflect the risk arising from the level or volatility of market prices of financial instruments which have an impact upon the value of the assets and liabilities of the undertaking. It shall properly reflect the structural mismatch between assets and liabilities, in particular with respect to the duration thereof.

---

<sup>26</sup>According to EIOPA 2022 Stress Test, the capital requirement for market risk accounts for 85.4% of the gross SCR before diversification benefits in the baseline.

<sup>27</sup>Source: Directive 2009/138/EC of the European Parliament and of the Council of 25 November 2009 on the taking-up and pursuit of the business of Insurance and Reinsurance (Solvency II).

<sup>28</sup>Source: [EIOPA Insurance Statistics](#)

<sup>29</sup>We compute this ratio using the value of technical provision of the aggregate EU balance sheet provided by EIOPA for 2021.

<sup>30</sup>Source: Directive 2009/138/EC of the European Parliament and of the Council of 25 November 2009 on the taking-up and pursuit of the business of Insurance and Reinsurance (Solvency II)

It shall be calculated, in accordance with point as a combination of the capital requirements for at least the following sub-modules:

- (a) the sensitivity of the values of assets, liabilities and financial instruments to changes in the term structure of interest rates, or in the volatility of interest rates (interest rate risk);
- (b) the sensitivity of the values of assets, liabilities and financial instruments to changes in the level or in the volatility of market prices of equities (equity risk);
- (c) the sensitivity of the values of assets, liabilities and financial instruments to changes in the level or in the volatility of market prices of real estate (property risk);
- (d) the sensitivity of the values of assets, liabilities and financial instruments to changes in the level or in the volatility of credit spreads over the risk-free interest rate term structure (spread risk);
- (e) the sensitivity of the values of assets, liabilities and financial instruments to changes in the level or in the volatility of currency exchange rates (currency risk);
- (f) additional risks to an insurance or reinsurance undertaking stemming either from lack of diversification in the asset portfolio or from large exposure to default risk by a single issuer of securities or a group of related issuers (market risk concentrations).

### **A.3 Pension Funds**

The pension fund sector accounts for a total of 3.1tn at the end of 2021. The pension fund system in the Euro area is mostly pay-as-you-go (or notional defined contribution)<sup>31</sup>, meaning that most of current pensions are funded by contributions from current workers. As a consequence, the total assets held by pensions are small. The only exception is Netherlands, whose pension fund sector's assets total 2tn (and hence more than two thirds of the total Euro area sector). Netherlands combine a pay-as-you-go system with a defined-benefit individual investment system, managed by pension funds. EIOPA is also in charged of the supervision of institutions for occupational retirement provision. The IORP II Directive sets common standards to protect pension scheme members and their beneficiaries. The directive also set rules on IORPs' own risk assessment. Differently from the insurance companies there are not a specific set of quantitative requirements.

### **A.4 Mutual Funds**

The mutual fund sector (non-MMF investment funds) accounted for 17tn at the end of 2021. Mutual funds are regulated by the Undertakings for Collective Investments in Transferable Securities

---

<sup>31</sup>See [OECD \(2011\)](#)

(UCITS) Directive. Under the UCITS Directive funds have to comply with limits on balance sheet leverage, and borrowing should not exceed 10% of assets on a temporary basis. The UCITS Directive also imposes limit for synthetic leverage (which stems from derivative instruments or securities financing transactions), either through limits to the ratio of overall exposure with respect to NAV (Net Asset Value) or through limits to VaR (Value at Risk).<sup>32</sup>

Mutual funds in the Euro area are mostly held directly by households or by insurance companies and pension funds. However, some of the insurance investment fund holdings are connected to the unit-linked or index-linked policies aforementioned. In this case, as the policyholder bears the risk these can be considered as indirect household holdings. Using data from EIOPA, we calculate that approximately unit-linked and index-linked account for 47% of the total insurance companies mutual fund holdings.

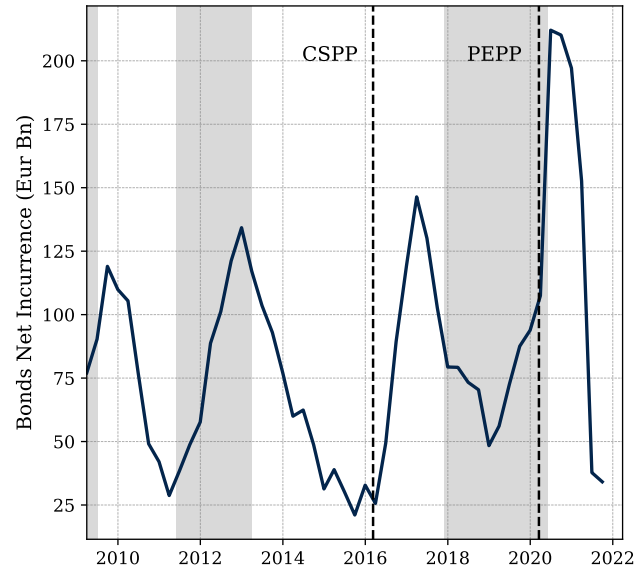
Table VI: Holding Shares Mutual Funds

	2014	2019	2021
Households	0.34	0.31	0.33
Insurance Corp.	0.28	0.30	0.30
Pension Funds	0.17	0.19	0.16
Non-Financial Corp.	0.08	0.07	0.07
OFIs	0.03	0.04	0.06
Government	0.04	0.05	0.05
Banks and MMF	0.06	0.04	0.04

<sup>32</sup>See [Doyle et al. \(2016\)](#) for further details).

## B. ECB and Corporate Bonds

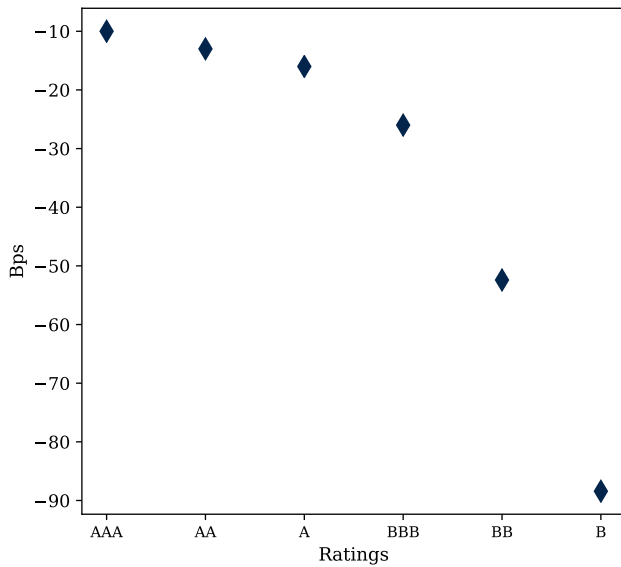
Figure 20: Corporate Bond Volumes, CSPP and PSPP



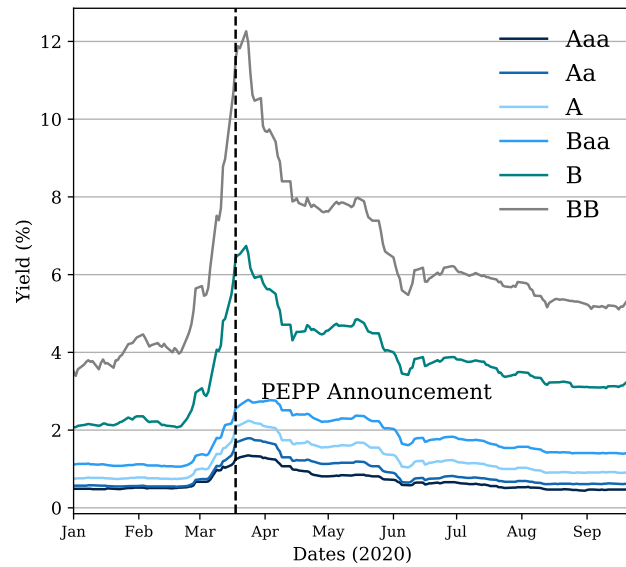
*Note:* The figure plots issuance of bonds by non-financial corporations. Source: Quarterly Sector Account.

Figure 21: Bloomberg-Barclays index, CSPP and PSPP, by Rating

(a) Response to CSPP, by Rating



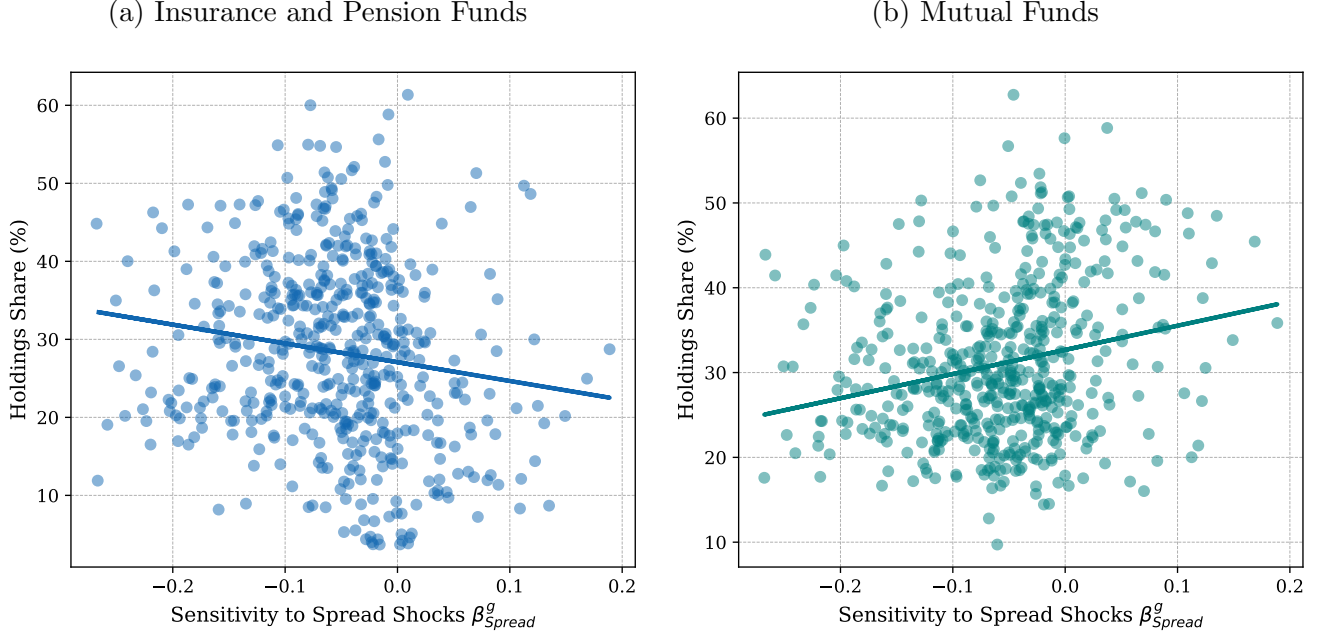
(b) PEPP Announcement, by Rating



*Note:* The figure plots the evolution of bonds around CSPP and PEPP announcements. Source: Bloomberg-Barclays.

## C. Heterogeneous Sensitivity

Figure 22: Holding shares and Bond Sensitivity to Spread Shocks



*Note:* The figure plots the share held by different intermediaries and the sensitivity to spread shocks. Source: Markit iBoxx, ECB SHSS.

## D. Drawdowns

Recent literature demonstrated that bonds held predominantly by domestic insurance companies rather than mutual funds suffer milder losses in crises. [Coppola \(2021\)](#) showed that during a downturn, mutual funds face withdrawals that push them to liquidate their positions and exacerbate the rise in bond yields. [Coppola \(2021\)](#) analysis mainly focused on the US but also present tentative evidence that similar mechanism holds in other countries, including the Euro area. We extend their analysis to our data to the Covid crisis. Compared to their analysis, we include detailed holdings by insurance companies, pension funds and banks. We analyze data in the first quarter of 2020, which saw a sharp rise in bond yields. We define drawdowns as the maximum variation experienced in bond yield spread in the quarter (maximum yield spread minus minimum yield spread). We then regress the drawdowns on the shares held by levered institutions or mutual funds, as in Equation 29. Table VII shows the results. We also find a sharper increase in spread for bonds held by mutual funds. Going from model 1 to model 7 we also find that the difference decreases, as part of the difference owes to selection rather than treatment.

$$Draw(y_{s_i}) = \theta^{Lev} + \theta^{MF} + \text{Interacted Fixed Effects} + \varepsilon_i \quad (29)$$

Table VII confirms the finding of Coppola (2021), showing that

Table VII: Holdings and Covid Drawdown

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Draw ys b/se	Draw ys b/se	Draw ys b/se	Draw ys b/se	Draw ys b/se	Draw ys b/se	Draw ys b/se
Levered	21.924 (22.16)	29.445 (22.96)	23.648 (32.47)	28.930 (18.46)	6.999 (17.29)	27.188* (13.96)	0.175 (5.70)
MF	178.570*** (26.87)	184.841*** (33.82)	148.305*** (32.29)	78.574*** (27.01)	44.905 (30.92)	84.086*** (27.79)	16.878** (7.04)
Observations	556	556	551	544	474	472	387
Adj. R-squared	0.086	0.082	0.182	0.503	0.635	0.725	0.964
Clustering		group	group	group	group	group	group
Fixed Effects	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN	ISIN
Duration	No	Yes	Yes	Yes	Yes	Yes	Yes
Country	No	No	Yes	Yes	Yes	Yes	Yes
Rating	No	No	Yes	Yes	Yes	Yes	Yes
Rating Worst	No	No	No	Yes	Yes	Yes	Yes
Liquidity	No	No	No	No	Yes	No	No
Sector	No	No	No	No	No	Yes	Yes
Issuer	No	No	No	No	No	No	Yes

## E. Additional Results

### E.1 Variance Decomposition

In the previous sections, we provided evidence that ECB corporate bond purchases exert notable effects on bond yields and especially on the credit component. In Equation 1 we decompose the bond market yields into two components:

$$Y_t = \underbrace{Y_t^{IR}}_{\text{Interest Rate Component}} + \underbrace{Y_t^{CR}}_{\text{Credit Risk Component}}.$$

We then collect the list of regularly scheduled ECB governing council meeting from 2006 till 2021 and measure the variation in the two components around the ECB announcements. Define as  $t = 1, ..T$  the day of each announcement. Further define  $\Delta Y_{t-1,t+j}$  the variation from the day before to  $j$ -days after the announcement. There is a trade-off between taking a window which is too short (as the corporate bond prices may have not fully incorporated information) or windows that are too long (as many other events take places in the meanwhile). For this reason, we provide information using different lengths for the monetary policy window. We then conduct a variance decomposition of the changes in corporate bond yield index around ECB announcements:

$$Var(\Delta Y_{t-1,t+j}) = \underbrace{Var(\Delta Y_{t-1,t+j}^{IR})}_{\text{Interest Rate Component}} + \underbrace{Var(\Delta Y_{t-1,t+j}^{CR})}_{\text{Credit Risk Component}} + 2Cov(\Delta Y_{t-1,t+j}^{IR}, \Delta Y_{t-1,t+j}^{CR}).$$

We also select the list of announcements which conveyed information about the corporate sector purchases. Table VIII lists the main announcements related to the corporate sector purchases (CSP) from March 2016 onward. We call the announcements listed in the table corporate sector purchases (CSP) announcement, to distinguish for the rest of other ECB announcements (Non-CSP announcements). Information regarding the purchase programme – as well as other monetary policy measure – was also released in non-scheduled meetings. As discussed, the most notable of this example is the announcement related to the deployment of the Pandemic Emergency Purchase Programme (PEPP), announced on March 18, 2020 with a press release on the ECB website. In this exercise we only focus on the announcements held during the scheduled governing council (GC) meetings.

Table VIII: Announcements Related to Corporate Sector Purchases (CSP)

Date	Announcement
2016-03-10	Launched CSPP, extend APP to 80bn
2016-04-21	Press release with information on CSPP
2016-12-08	Extension of APP from April 2017 to Dec 2017 at 60
2017-10-26	Extend APP from January 2018 to September 2018 at 30 billion
2018-07-26	Reduce APP to 15bn until end of December 2018 and then end net purchases
2019-09-12	Restart APP at 20 billion
2020-03-12	Additional net asset purchases of 120 billion until the end of the year
2020-06-04	Increase the envelope for PEPP by 600bn to a total of 1,350bn
2020-12-10	Increase the envelope of PEPP by 500 bn to a total of 1,850bn
2021-12-16	Reduced the pace APP

*Note:* The table shows the list of the main ECB announcements with information regarding corporate bond purchases.

We then compute the variance of the interest rate component and the credit risk component on both the CSP announcements and the Non-CSP announcements. Table IX reports the i) variance of IR changes ( $Var(\Delta Y_{t-1,t+j}^{IR})$ ) and CR changes ( $Var(\Delta Y_{t-1,t+j}^{CR})$ ) on CSP announcements as well as Non-CSP announcements, ii) the ratio between the variance on CSP vis-à-vis Non-CSP announcements, iii) the ratio between the CS variance and IR variance on the two types of the announcements. We report results for different windows, from one day lag to three days lag ( $j = 1, 2, 3$ ), in Panel a), b) and c), respectively. We also report results for different time periods: 2006-2021 and 2014-2021. Using the ratio between CSP and Non-CSP announcement days allow us to better compare standard announcement versus announcements on corporate bond asset



purchases. It also lessens the concern related to the overall volatility of the two components: it may be that the different in variance on announcements reflect the fact that the CR component is in general more volatile than the IR component, or vice versa. different variance of IR and CR components. Taking the ratio of the variance allows us to isolate whether announcement on corporate purchases lead to different effects if compared to standard announcements. The results in Table IX confirms the suggestive evidence of Figure 4. If we take the full sample 2006-2021, for the two days lag, the variance ratio for IR is 0.39: that means that the variance on CSP announcement is roughly half of the variance on CSP announcement. For CS, the ratio is approximately 14. This means that CSP announcement seems to have strong effects on credit spread. The table also suggest that Non-CSP announcement has small effect on CR if compared to IR: the relative variance is 0.5. Instead on CSP announcement, the ratio of variance is close to 20. This means that overall we are observing larger effects on CR than IR even in absolute terms.

Table IX: CPS announcements: IR vs CS decomposition

	Panel a) 1 day lag			2014-2021		
	2006-2021					
	IR	CR	Ratio CR/IR	IR	CR	Ratio CR/IR
Non-CSP	51.67	18.77	0.36	15.84	5.13	0.32
CSP	32.16	175.08	5.44	32.16	175.08	5.44
Ratio CSP/Non-CSP	0.62	9.33		2.03	34.11	
	Panel b) 2 days lag			2014-2021		
	2006-2021					
	IR	CR	Ratio CR/IR	IR	CR	Ratio CR/IR
Non-CSP	75.17	41.87	0.56	14.87	10.07	0.68
CSP	29.38	580.54	19.76	29.38	580.54	19.76
Ratio CSP/Non-CSP	0.39	13.86		1.98	57.65	
	Panel c) 3 days lag			2014-2021		
	2006-2021					
	IR	CR	Ratio CR/IR	IR	CR	Ratio CR/IR
Non-CSP	78.40	73.12	0.93	17.73	12.22	0.69
CSP	33.24	846.29	25.46	33.24	846.29	25.46
Ratio CSP/Non-CSP	0.42	11.57		1.87	69.27	

*Note:* The table shows the variance decomposition of the variation in bond yields around ECB announcement. Source: Markit iBoxx.

## E.2 Empirical Measure of Hedgeability

We estimate empirically the hedgeability coefficient, described in Equation 24. We use the full set of investment grade corporate bonds. We sort bonds based on their risk. We then compute monthly returns on each bond portfolio. We also use data on the Euribor Swap curve to measure the risk-free returns. As insurance companies measure their liabilities using the Swap curve, the returns on swap are a good measure of the variation in valuation of insurance companies assets.

The hedgeability coefficient can then be measured as a linear regression:

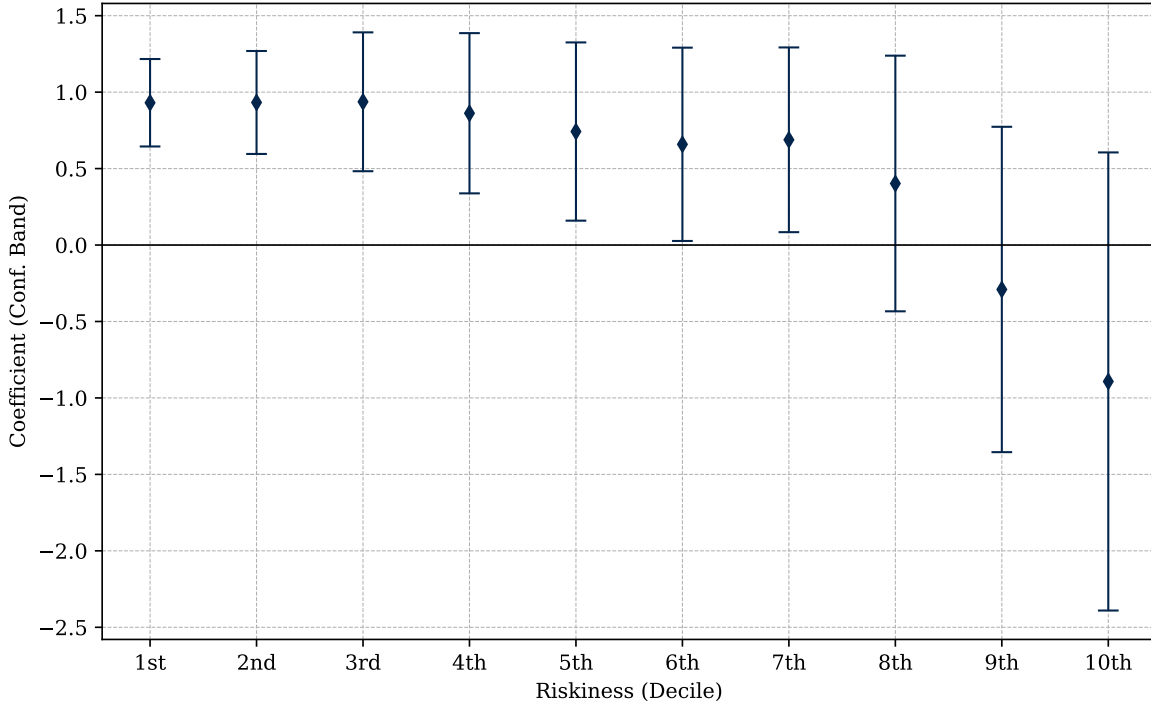
$$R_t^L = \alpha + \beta^{Hed} R_{i,t} + \varepsilon_t \quad (30)$$

where  $R_t^L$  are the returns on swap.  $R_{i,t}$  are the returns on our sorted bond portfolios. In fact:

$$\beta^{Hed} = \frac{Cov(R_t^L, R_{i,t})}{Var(R_{i,t})}$$

We evaluate the regression 30 on our data. The results are displayed in Figure 23.

Figure 23: Empirical Hedgeability Measure

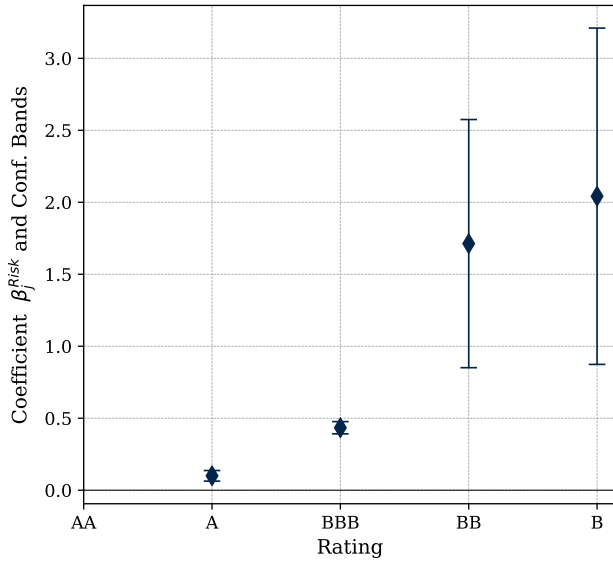


## E.3 Sensitivity to credit spread Shocks by Ratings and Maturity

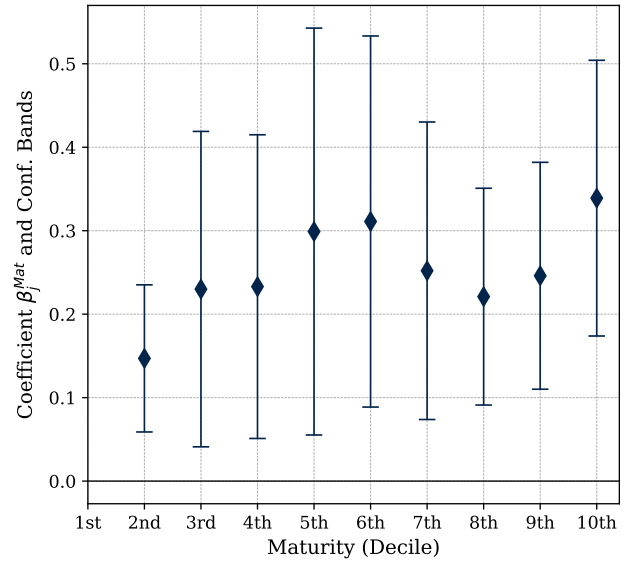
Figure 24 shows the results for Equation 28 where we used ratings instead of credit-risk deciles. The results in Figure 24a are broadly in line with our previous results.

Figure 24: Sensitivity to CR Shocks by Rating and Maturity

(a) Sensitivity to CR Shocks by Ratings



(b) Sensitivity to CR Shocks by Maturity Decile



*Note:* The figure shows the results of Equation 28. We divided bond based on their credit risk and maturity. We used ratings instead of credit-risk deciles. Figure 24a displayed the estimated  $\beta_j^{Risk}$  together with their confidence bands. Figure 24b display the estimated  $\beta_j^{Mat}$  together with their confidence bands. The sample runs from 2014 through 2021. The coefficients are estimated relative to the first riskiness and maturity group (for this reason, the coefficient for the first group are missing). Source: Markit iBoxx.