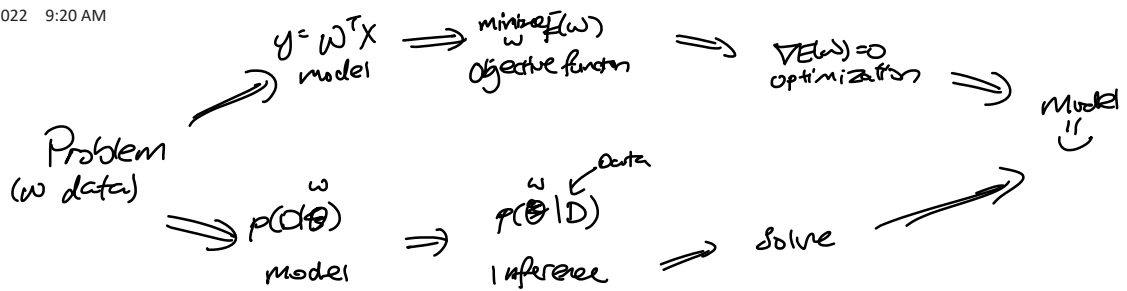


CS3244 Lecture 2

Wednesday, 11 May 2022 9:20 AM

Probability Space

$$P(\{\square, \square, \square\}) = \frac{1}{2} \quad P(\cdot) = 1$$

$\in E$

 (Ω, E, P)

$$P: E \mapsto [0, 1]$$

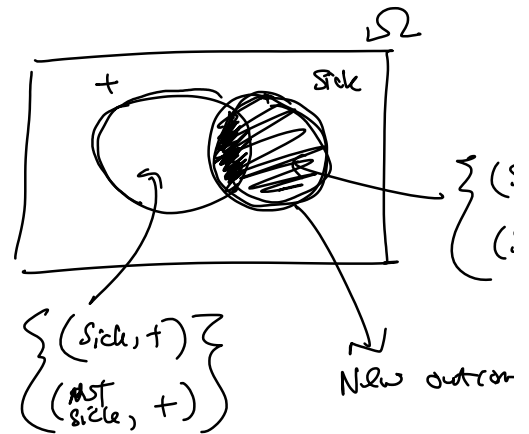
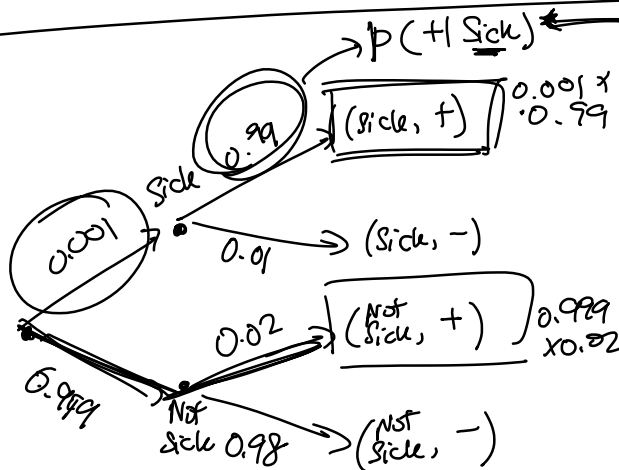
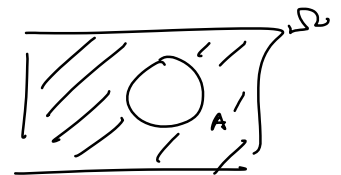
$\{\square, \square, \square, \dots, \square\}$

Set of all subsets of Ω
(technically wrong)

- ① $P(\Omega) = 1$
- ② $P(\emptyset) = 0$
- ③ $P(A \cup B) = P(A) + P(B)$

E satisfies:

- ① if $A \in E, B \in E$, then $A \cup B \in E$
- ② " " " " $A \cap B \in E$
- ③ if $A \in E$, then $A^c \in E$
- ④ $\emptyset \in E, \Omega \in E$



$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

$$P(+|Sick) = \frac{P(Sick, +)}{P(Sick)}$$

Product Rule

$$P(Sick|+) = \frac{P(Sick, +)}{P(+)} = \frac{0.001 \times 0.99}{0.001 \times 0.99 + 0.999 \times 0.02} \neq \frac{0.001 \times 0.99}{0.001 \times 0.99}$$

Sum rule.

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \Rightarrow P(A \cap B) = P(B)P(A|B)$$

Product Rule.

$$\begin{cases} P(A_1 \cap A_2 \cap \dots, A_n) = P(A_1)P(A_2|A_1)P(A_3|A_2, A_1) \dots \\ P(A_n|A_1, \dots, A_{n-1}) \\ = P(A_1)P(A_2|A_1)P(A_3|A_2, A_1) \dots \end{cases}$$

Independence

def $P(A|B) = P(A)P(B)$

if A and B disjoint

$\Rightarrow A$ and B are dependent.

$$\frac{P(A \cap B)}{P(A)P(B)} = \frac{P(A)P(B|A)}{P(A)P(B)} = 1$$

Random Variable

if X is a random variable (r.v.), then:

$$X: \Omega \rightarrow \mathbb{R}$$

Example:

$$\Omega = \{ \boxed{1}, \boxed{2}, \dots, \boxed{6} \}$$

$$X(\omega) = \begin{cases} 1 & \text{if } \omega = \boxed{1} \\ 2 & \text{if } \omega = \boxed{2} \\ \vdots & \vdots \\ 6 & \text{if } \omega = \boxed{6} \end{cases}$$

$$(1 + 6) = X_1(\omega) + X_2(\omega) =$$

$$\boxed{1} + \boxed{6}$$

Indicator r.v.

$$\mathbb{1}_L(\omega) = \begin{cases} 0 & \text{if } L \text{ is false} \\ 1 & \text{if } L \text{ is true.} \end{cases}$$

is even?

$$\mathbb{1}_{\text{even}}(\omega) = \begin{cases} 0 & \text{if } \omega \in \{ \boxed{1} \} \\ 1 & \text{if } \omega \in \{ \boxed{2}, \boxed{4}, \boxed{6} \} \end{cases}$$

19 + :

$(\{0,1\}, E, P)$

\downarrow
 $\{0, \{0,1\}, \{0,1\}, \{0,1\}\}$

$P(X=1)$
 $P(X=0)$

Distributions

Discrete

(Ω, E, P)

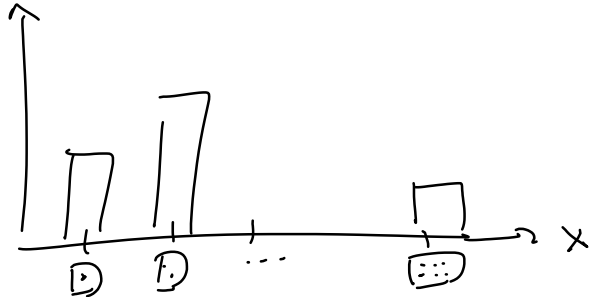
$\{0, \dots, \infty\}$

Discrete PMF

Uniform $P(X) = \frac{1}{|\Omega|}$

Poisson $P(X) = \frac{\lambda^x e^{-\lambda}}{x!}$

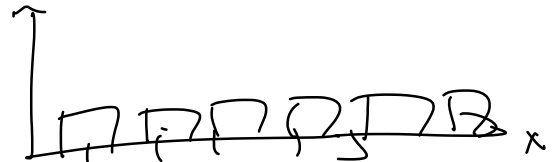
$P(\{x\})$



$P(\{0\})$

\vdots

$P(\{3\})$

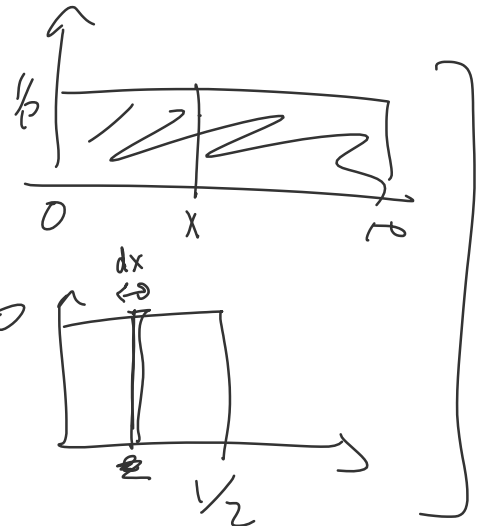


$\{0, 1, 2, 3, 4, 5\}$

$\Rightarrow \sum_{x \in \{0, 1, 2, 3, 4, 5\}}$

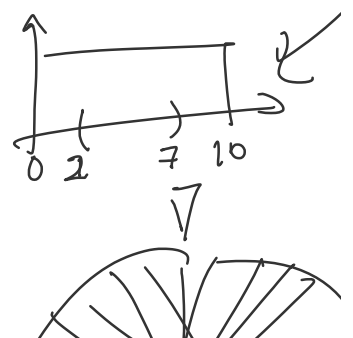
Continuous distribution.

$\left[\begin{array}{l} \text{pdf}(x) > 1 ? \\ \text{pdf}(\frac{1}{4}) = 2 \\ \text{YES} \end{array} \right]$

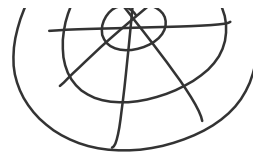


$P(C) = \int_C \text{pdf}(x) dx$

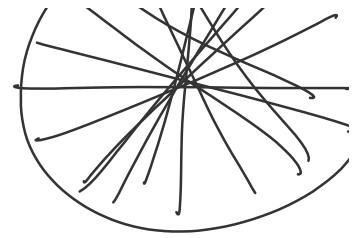
$P([2, 7]) = \int_{[2, 7]} \frac{1}{10} dx$



$$|P(\{SS\}) = 0|$$

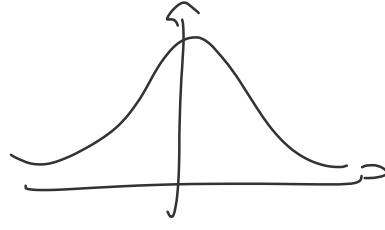


w.p zero.



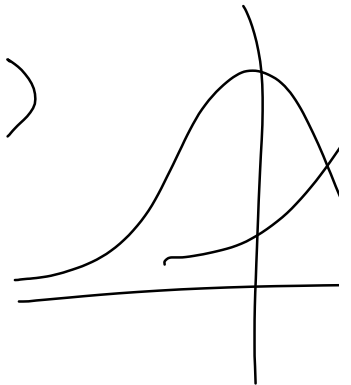
Gaussian distribution

$$p_{\theta_0}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$



if X, Y follows a gaussian distribution:

- then
- ① $X + Y$ is gaussian $N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$
 - ② aX is gaussian $N(a\mu, a^2\sigma_1^2)$
 - ③ $X + a$ is gaussian $N(\mu + a, \sigma_1^2)$



Learning parameter for gaussian distribution

Given $N(\mu^*, 1)$ $\mathcal{D} = \{x_1, x_2, \dots, x_N\}$ sampled from a
 don't know. Guess: $\frac{1}{N} \sum_{n=1}^N x_n \approx \mu^* = \hat{\mu}$

we need justification...

$$\boxed{\text{maximize}_{\mu} p(\mathcal{D}|\mu)}$$

Maximum Likelihood Estimation (MLE)

$$p(\mathcal{D}|\mu) = p(x_1|\mu) p(x_2|\mu) \dots p(x_N|\mu)$$

$$= \prod_{n=1}^N p(x_n|\mu)$$

$$= \prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(x_n - \mu)^2}{2}\right)$$

$$\forall n: p(x_n|\mu) = \frac{1}{\sqrt{2\pi}} \exp$$

Detour



log.

$$\begin{aligned} \log p(D|\mu) &= \log \prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(x_n - \mu)^2}{2}\right) \\ &= \sum_{n=1}^N \log \frac{1}{\sqrt{2\pi}} + \log \exp\left(-\frac{(x_n - \mu)^2}{2}\right) \\ &\Rightarrow \sum_{n=1}^N \log \exp\left(-\frac{(x_n - \mu)^2}{2}\right) \\ &= \sum_{n=1}^N -\frac{(x_n - \mu)^2}{2} \end{aligned}$$

min

$$\sum_{n=1}^N \frac{(x_n - \mu)^2}{2}$$

max $-()$
min $(-)$

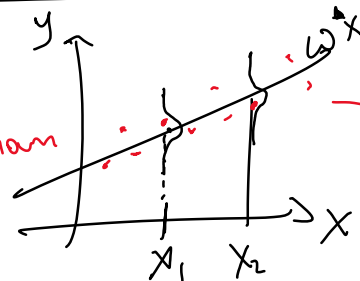
$$\frac{d}{d\mu} \sum_{n=1}^N -\frac{2(x_n - \mu)}{2} = \sum_{n=1}^N -(x_n - \mu) \stackrel{\text{set}}{=} 0$$

$$\sum_{n=1}^N \mu = \sum_{n=1}^N x_n \Rightarrow \mu_{ML} = \frac{1}{N} \sum_{n=1}^N x_n$$

Nμ

Linear regression REVISITED

Given x true ω ← unknown
 $\Rightarrow \omega^T x + \epsilon$
 $\epsilon \sim N(0, 1)$



log p(D|ω)

$$= \log \prod_{n=1}^N p(t_n | x_n, \omega)$$

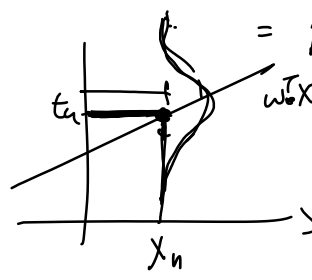
$$= \log \prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(t_n - \omega^T x_n)^2}{2}\right)$$

$$= \sum_{n=1}^N \log \frac{1}{\sqrt{2\pi}} + \frac{-(t_n - \omega^T x_n)^2}{2}$$

maximize.

 $p(t_n | x_n, \omega)$

$$= N(\omega^T x_n, 1)$$



$$t_n = \omega^T x_n + \epsilon$$

$$t_n \sim N(\omega^T x_n + \epsilon)$$

N(0, 1)

$$\Rightarrow - \sum_{n=1}^N \frac{(t_n - w^T x_n)^2}{2} \leftarrow \text{minimize}$$

$$\Rightarrow \sum_{n=1}^N \frac{(t_n - w^T x_n)^2}{2} \leftarrow E(w)$$

$$t_n = \frac{w^T x_n + \epsilon}{\sigma^2}$$

$$\leftarrow N(w^T x_n, 1)$$

Maximum a posteriori estimator

$$p(D|w)$$

$$p(w|D) = \frac{p(D|w)p(w)}{p(D)}$$

Linear Regression.

$$p(D)$$

Don't care

$$\log p(w|D) = \log \frac{p(D|w)p(w)}{p(D)}$$

$$= \log p(D|w) + \log p(w) - \underbrace{\log p(D)}_{\text{ind. } w}$$

$$\Rightarrow \log p(D|w) + \log p(w)$$

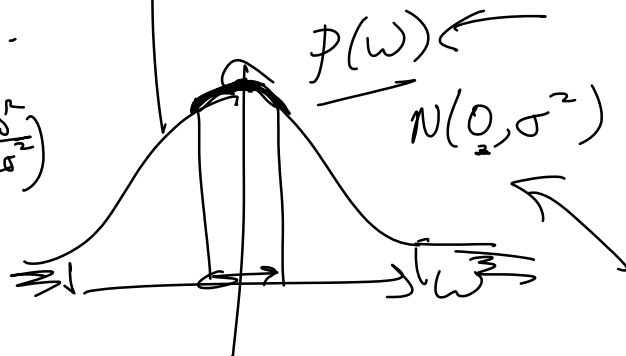
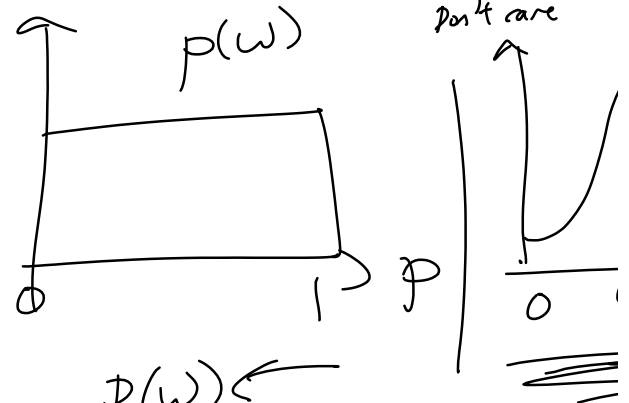
$$= \sum_{n=1}^N -\frac{(t_n - w^T x_n)^2}{2} + \log p(w)$$

$$= \sum_{n=1}^N -\frac{(t_n - w^T x_n)^2}{2} + \log \left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{w^2}{2\sigma^2}\right) \right)$$

$$= \sum_{n=1}^N \dots + \underbrace{\log \frac{1}{\sqrt{2\pi\sigma^2}}}_{\text{ind. } w} + \frac{-w^2}{2\sigma^2}$$

$$\Rightarrow - \sum_{n=1}^N \frac{(t_n - w^T x_n)^2}{2} - \frac{w^2}{2\sigma^2} \leftarrow \text{max}$$

$$\Rightarrow \sum_{n=1}^N \frac{(t_n - w^T x_n)^2}{2} + \frac{1}{2} \frac{w^2}{\sigma^2} \leftarrow \text{min}$$



$$(\Omega, \mathcal{E}, \mathbb{P})$$

$$P(\Omega)$$

has to have Ω, ϕ

