

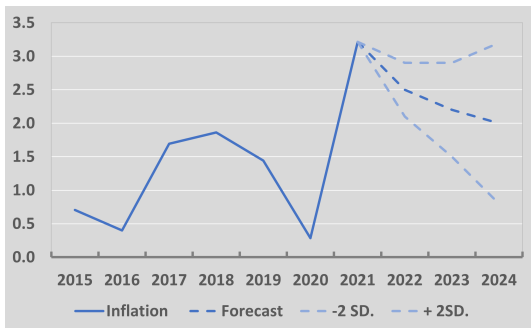
How accurate are survey-based measures of macroeconomic uncertainty?

Alexander Glas and Matthias Hartmann

Freier Fachvortrag
Berufungsverfahren W2-Professur in Volkswirtschaftslehre
ESB Business School, Hochschule Reutlingen

March 1, 2022

Macroeconomic uncertainty



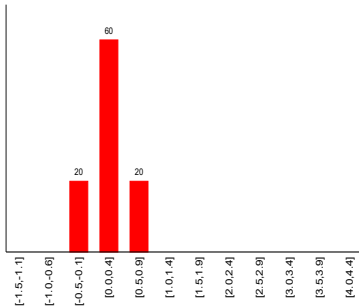
- **Forecasting applications:** Point forecasts without uncertainty bands are incomplete (Dawid, 1984; Gneiting and Raftery, 2007).
- **Monetary policy applications:** Heightened inflation uncertainty may be related to de-anchoring of inflation expectations (Kumar et al., 2015).

This study

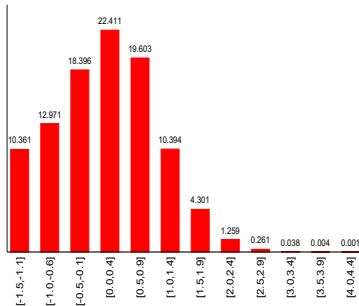
- **Aim:** Measurement of average uncertainty, from a panel of survey-based histogram forecasts.
- **Starting point (Kenny et al., 2014):**
Variances of histogram forecasts \ll mean squared forecast errors.
- **Research questions:**
Why do ex-ante uncertainty and ex-post uncertainty differ?
Can we correct for this?
- **Novel findings and contribution:**
 - A subgroup of forecasters reports strongly rounded numbers.
 - Histogram forecasts of this group are very narrow.
 - Discarding these observations helps to re-align average ex-ante uncertainty and ex-post uncertainty.

Example histogram forecasts for HICP inflation, 1Q ahead

Forecaster #102



Forecaster #95



Both histogram forecasts were reported as part of the ECB's *Survey of Professional Forecasters* in 2016Q4.

Survey of Professional Forecasters (SPF)

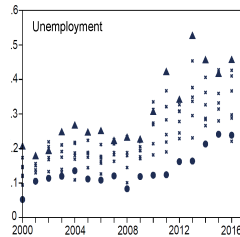
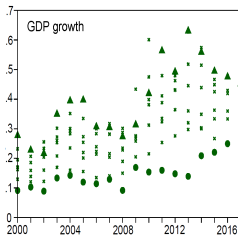
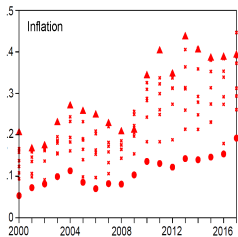
- Data set comprises forecasts from **ECB-SPF** and **FED-SPF**.
- Quarterly **fixed-event** histogram forecasts for inflation, real GDP growth and unemployment rates in the euro area and the U.S.
- **Common sample**: 1999Q1-2017Q4 (target years: 2000-2017).
- Unemployment rate in FED-SPF available since 2009Q2.

	Forecast horizon h	
	Current year	Next year
Q1	4	8
Q2	3	7
Q3	2	6
Q4	1	5

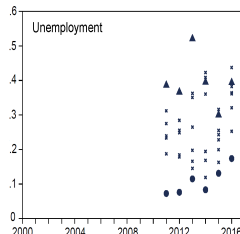
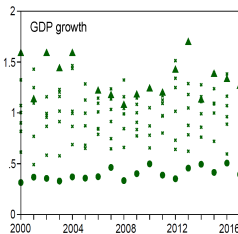
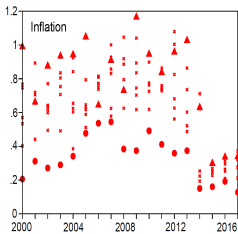
- Sequence of h -step-ahead histogram forecasts for target year $t = 1, \dots, T$ with forecast horizon $h \in \{8, 7, \dots, 1\}$.

Consensus ex-ante uncertainty

ECB-SPF



FED-SPF



Target period

Notes: Graphs depict the cross-sectional averages over the variances derived from the individual histograms, i.e. $\sigma_{i,t,h}^2$, at forecast horizons $h \in \{8, 7, \dots, 1\}$. Triangles " \triangle " depict the 8-step-ahead variance forecasts and bullets " \bullet " the 1-step-ahead variance forecast. The sample period is 1999Q1-2017Q4.

Integer-based categorization

- Probabilistic survey forecasts often tend to contain numbers that are **multiples of five or ten** (Engelberg et al., 2009; Clements, 2011; Boero et al., 2015).
- Consider a histogram forecast by forecaster $i = 1, \dots, N$, consisting of probabilities, $p_{i,k}$, for outcome intervals ('bins') $k = 1, \dots, K$.
- Check whether $p_{i,k}$ is a **multiple of five**:

$$\tilde{D}_{i,k}^{\text{m5}} = \begin{cases} 1 & \text{if } 5 \cdot \lfloor \frac{p_{i,k}}{5} \rfloor = p_{i,k} \text{ and} \\ 0 & \text{else.} \end{cases}$$

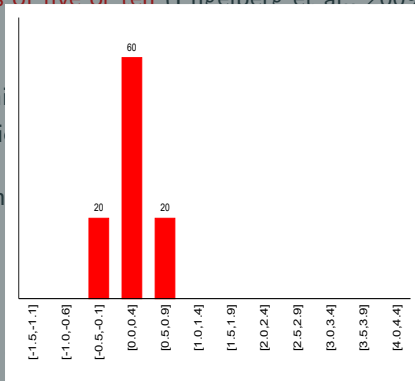
- Classify i as a **rounder** according to the following rule:

$$\tilde{D}_i^{\text{m5}} = \begin{cases} 1 & \text{if } \text{mode}(\tilde{D}_{i,1}^{\text{m5}}, \dots, \tilde{D}_{i,K}^{\text{m5}}) = 1 \text{ and} \\ 0 & \text{else.} \end{cases}$$

- Consider $D_i^{\text{m5}} = 1 - \tilde{D}_i^{\text{m5}}$ instead (focus on **non-rounders**).

Integer-based categorization

- Probabilistic survey forecasts often tend to contain numbers that are multiples of five or ten (Engelberg et al., 2009; Clements, 2011; Boero et al., 2012).



- Consider a histogram \tilde{D}_i^{m5} of probabilistic forecasts $\tilde{D}_i^{\text{m5}} = \{\tilde{D}_{i,1}^{\text{m5}}, \dots, \tilde{D}_{i,N}^{\text{m5}}\}$, consisting of N forecasts $\tilde{D}_{i,k}^{\text{m5}}$, $k = 1, \dots, N$.

- Check whether

- Classify i as

$$\tilde{D}_i^{\text{m5}} = \begin{cases} 1 & \text{if } \text{mode}(\tilde{D}_{i,1}^{\text{m5}}, \dots, \tilde{D}_{i,K}^{\text{m5}}) = 1 \text{ and} \\ 0 & \text{else.} \end{cases}$$

- Consider $D_i^{\text{m5}} = 1 - \tilde{D}_i^{\text{m5}}$ instead (focus on non-rounders).

Integer-based categorization

- Probabilistic survey forecasts often tend to contain numbers that are **multiples of five or ten** (Engelberg et al., 2009; Clements, 2011; Boero et al., 2015).
- Consider a histogram forecast by forecaster $i = 1, \dots, N$, consisting of probabilities, $p_{i,k}$, for outcome intervals ('bins') $k = 1, \dots, K$.
- Check whether $p_{i,k}$ is a **multiple of five**:

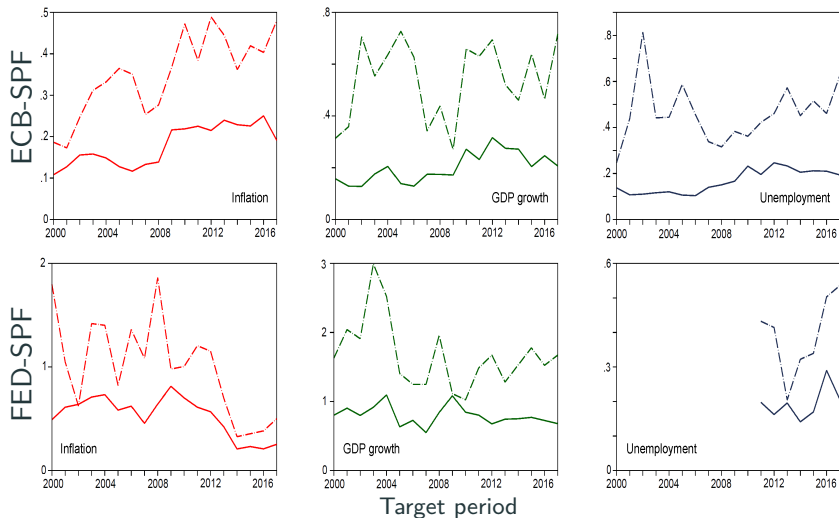
$$\tilde{D}_{i,k}^{m5} = \begin{cases} 1 & \text{if } 5 \cdot \lfloor \frac{p_{i,k}}{5} \rfloor = p_{i,k} \text{ and} \\ 0 & \text{else.} \end{cases}$$

- Classify i as a **rounder** according to the following rule:

$$\tilde{D}_i^{m5} = \begin{cases} 1 & \text{if } \text{mode}(\tilde{D}_{i,1}^{m5}, \dots, \tilde{D}_{i,K}^{m5}) = 1 \text{ and} \\ 0 & \text{else.} \end{cases}$$

- Consider $D_i^{m5} = 1 - \tilde{D}_i^{m5}$ instead (focus on **non-rounders**).

Average ex-ante uncertainty for $h = 4$



Dashed lines: uncertainty from non-rounded forecasts,
solid lines: uncertainty from rounded forecasts.

Horizon-specific variance misalignment

- Individuals' ex-ante uncertainty at horizon h is computed as:

$$\overline{\sigma_{i,h}^2} = \frac{1}{T_{i,h}} \sum_{t=1}^{T_{i,h}} \sigma_{i,t,h}^2.$$

- Ex-post uncertainty is defined as the mean squared error:

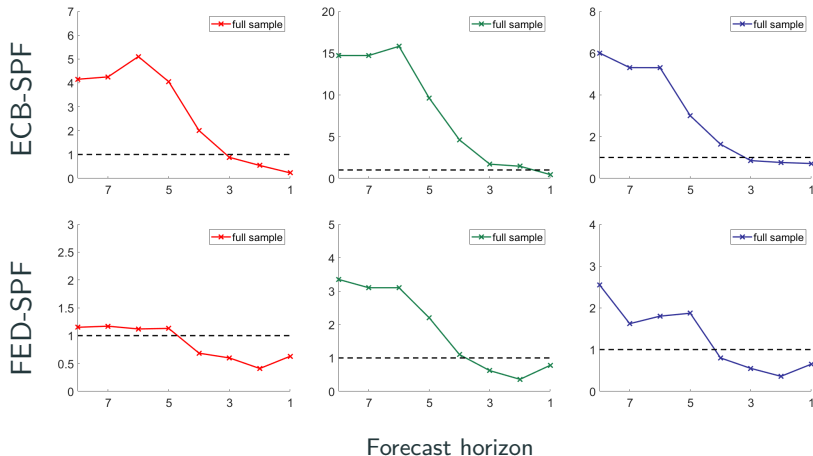
$$\text{MSE}_{i,h} = \frac{1}{T_{i,h}} \sum_{t=1}^{T_{i,h}} e_{i,t,h}^2,$$

where $e_{i,t,h} = x_t - \mu_{i,t,h}$ denotes forecast errors.

- **Average misalignment ratio** (Clements, 2014):

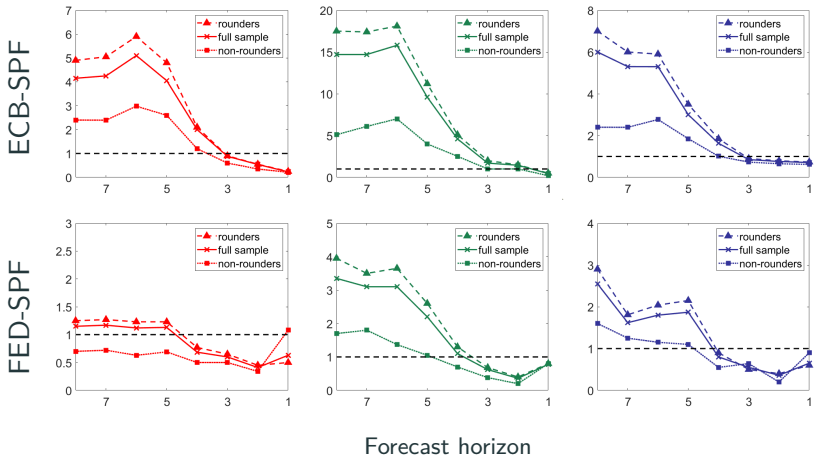
$$m_h = \frac{1}{N_h} \sum_{i=1}^{N_h} \frac{\text{MSE}_{i,h}}{\overline{\sigma_{i,h}^2}}.$$

Variance misalignment



Notes: For each forecast horizon, the plots display the average misalignment ratios m_h for the ECB- (first row) and FED-SPF (second row). The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Variance misalignment



Notes: For each forecast horizon, the plots display the average misalignment ratios m_h for the ECB- (first row) and FED-SPF (second row). Non-rounders are classified by means of $D_{i,t,h}^{m5}$. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Conclusion

- Rounding as a reliable indicator of misalignment between ex-ante and ex-post variances.
- Empirically, we confirm that rounders in the SPF...
 - ...use fewer of the available bins and report smaller ex-ante variances than the non-rounders
 - ...do not differ in terms of forecast errors from the non-rounders.
- Consensus ex-ante uncertainty based solely on non-rounders is about twice as high as overall consensus, also distinct dynamics.
- Next steps: Application of corrected macroeconomic uncertainty indexes in SVARs for Euro area and U.S.
- Extension to uncertainty measures from household and firm surveys of the Deutsche Bundesbank.

Link to data set and slides



Previous work on rounding and uncertainty

■ Rounding as a source of information

- Rounding behavior of surveyed forecasters sometimes interpreted as indicator of uncertainty – “Round numbers suggest round interpretations” (RNRI), Fischhoff and Bruine de Bruin (1999).
- Binder (2017) and Goldfayn-Frank and Wohlfart (2020) use rounding as a means to quantify uncertainty.

■ Rounding as a loss of information

- Heitjan and Rubin (1991) regard rounded survey responses as *coarse data*, with
missing data \prec coarse data \prec complete data
- Informative content of higher moments from survey-based histogram forecasts called in question e.g. by D’Amico and Orphanides (2008) or Clements (2014, 2016).

This study

- traces misalignment between ex-ante uncertainty and ex-post uncertainty to characteristics of 2 subsamples of the cross section,
- uses rounding behavior as a means to distinguish skillful variance forecasters from the rest,
- shows that computing ex-ante uncertainty from subsample of skillful forecasts alone provides better indication of realized forecast uncertainty.

Section 1 – Data

Survey of Professional Forecasters (SPF)

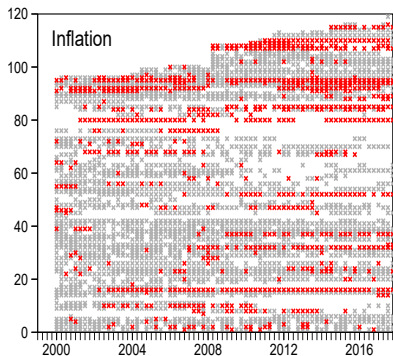
- Data set comprises forecasts from **ECB-SPF** and **FED-SPF**.
- Quarterly **fixed-event** histogram forecasts for inflation, real GDP growth and unemployment rates in the euro area and the U.S.
- **Common sample**: 1999Q1-2017Q4 (target years: 2000-2017)
- Unemployment rate in FED-SPF available since 2009Q2

	Forecast horizon h	
	Current year	Next year
Q1	4	8
Q2	3	7
Q3	2	6
Q4	1	5

- Sequence of h -step-ahead histogram forecasts for target year $t = 1, \dots, T$ with forecast horizon $h \in \{8, 7, \dots, 1\}$

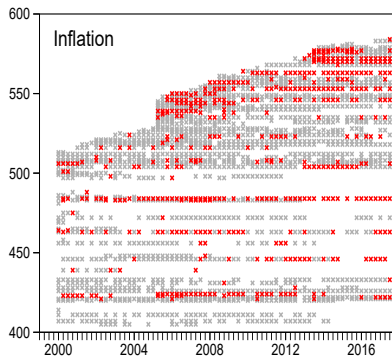
Forecaster panel for inflation

ECB-SPF



Survey period

FED-SPF



Survey period

Notes: Graphs depict forecaster participation for the (current year) inflation rate in the ECB- (left) and FED-SPF (right). The sample period is 1999Q1-2017Q4.

Target variables

- For inflation and real GDP growth, outcome variable x_t for year $t = 1, \dots, T$ refers to

$$x_t = 100 \times \left(\frac{X_t}{X_{t-1}} - 1 \right).$$

- X_t denotes annual average of either the respective price index or real GDP.
- For unemployment rate, x_t denotes annual average over monthly unemployment rate, i.e. $x_t = X_t$.
- First-vintage realizations drawn from
 - *Statistical Data Warehouse* of the ECB for euro area,
 - *Real-Time Data Set for Macroeconomists* of the FED Philadelphia for US.

Participants

- Majority of cross section from financial industry or research institutes.
- Most participants' names are published yet not associated with ID numbers to guarantee anonymity.
- ECB-SPF: on average, >50 participants each survey round between 1999 and 2017.
- Examples ECB-SPF: *Allianz SE, ifo Institute for Economic Research, Société Générale, ...*
- FED-SPF: on average, >30 participants each survey round between 1999 and 2017.
- Examples FED-SPF: *Oxford Economics USA, Inc., IHS Markit, Johns Hopkins University Center for Financial Economics, ...*

Moments of histogram forecasts – mean

- Participants $i = 1, \dots, N$ assign probabilities $p_{i,k,t,h}$ to $k = 1, \dots, K$ pre-specified outcome intervals, the ‘bins’.
- We assume for each bin that all probability mass is concentrated at its center (‘Mass-at-midpoint’, Lahiri et al., 1988; Kenny et al., 2015).
- Mean of forecaster i ’s histogram is given by

$$\mu_{i,t,h} = \frac{1}{100} \sum_{k=1}^K p_{i,k,t,h} \times \mathbf{m}_k,$$

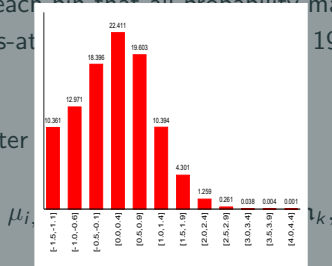
with \mathbf{m}_k denoting the midpoint of the k -th bin.

- The h -step-ahead ‘consensus’ forecast is given by

$$\bar{\mu}_{t,h} = \frac{1}{N} \sum_{i=1}^N \mu_{i,t,h}.$$

Moments of histogram forecasts – mean

- Participants $i = 1, \dots, N$ assign probabilities $p_{i,k,t,h}$ to $k = 1, \dots, K$ pre-specified outcome intervals, the ‘bins’.
- We assume for each bin that all probability mass is concentrated at its center (‘Mass-at-center’ assumption; see, e.g., Gneiting 1988; Kenny et al., 2015).
- Mean of forecaster i ’s h -step-ahead forecast



with m_k denoting the midpoint of the k -th bin.

- The h -step-ahead ‘consensus’ forecast is given by

$$\bar{\mu}_{t,h} = \frac{1}{N} \sum_{i=1}^N \mu_{i,t,h}.$$

Moments of histogram forecasts – mean

- Participants $i = 1, \dots, N$ assign probabilities $p_{i,k,t,h}$ to $k = 1, \dots, K$ pre-specified outcome intervals, the ‘bins’.
- We assume for each bin that all probability mass is concentrated at its center (‘Mass-at-midpoint’, Lahiri et al., 1988; Kenny et al., 2015).
- Mean of forecaster i ’s histogram is given by

$$\mu_{i,t,h} = \frac{1}{100} \sum_{k=1}^K p_{i,k,t,h} \times m_k,$$

with m_k denoting the midpoint of the k -th bin.

- The h -step-ahead ‘consensus’ forecast is given by

$$\bar{\mu}_{t,h} = \frac{1}{N} \sum_{i=1}^N \mu_{i,t,h}.$$

Moments of histogram forecasts – variance

- Based on the means from Eqn. (1), we calculate the individual variance according to

$$\sigma_{i,t,h}^2 = \frac{1}{100} \sum_{k=1}^K p_{i,k,t,h} \times (m_k - \mu_{i,t,h})^2.$$

as an ex-ante measure of individual uncertainty.

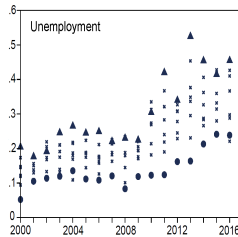
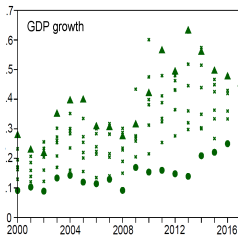
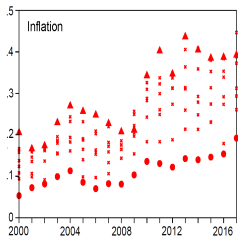
- Consensus ex-ante uncertainty obtains as

$$\overline{\sigma_{t,h}^2} = \frac{1}{N} \sum_{i=1}^N \sigma_{i,t,h}^2,$$

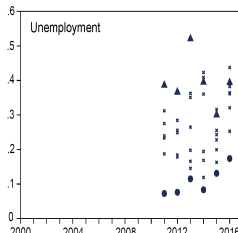
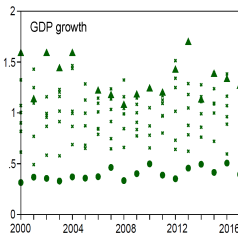
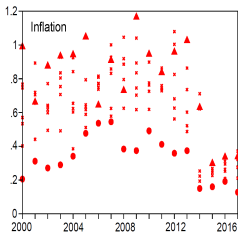
the ex-ante uncertainty of the consensus forecast $\bar{\mu}_{t,h}$ as proposed by Lahiri and Sheng (2010).

Consensus ex-ante uncertainty

ECB-SPF



FED-SPF



Target period

Notes: Graphs depict the cross-sectional averages over the variances derived from the individual histograms, i.e. $\sigma_{i,t,h}^2$, at forecast horizons $h \in \{8, 7, \dots, 1\}$. Triangles " \triangle " depict the 8-step-ahead variance forecasts and bullets " \bullet " the 1-step-ahead variance forecast. The sample period is 1999Q1-2017Q4.

Rounders vs. Non-rounders – Classification schemes

Integer-based categorization

- Probabilistic survey forecasts often tend to contain numbers that are **multiples of five or ten** (Engelberg et al., 2009; Clements, 2011; Boero et al., 2015)
- Consider a histogram forecast by forecaster $i = 1, \dots, N$, consisting of probabilities, $p_{i,k}$, for outcome intervals ('bins') $k = 1, \dots, K$.
- Check whether $p_{i,k}$ is a **multiple of five**:

$$\tilde{D}_{i,k}^{m5} = \begin{cases} 1 & \text{if } 5 \cdot \lfloor \frac{p_{i,k}}{5} \rfloor = p_{i,k} \text{ and} \\ 0 & \text{else.} \end{cases}$$

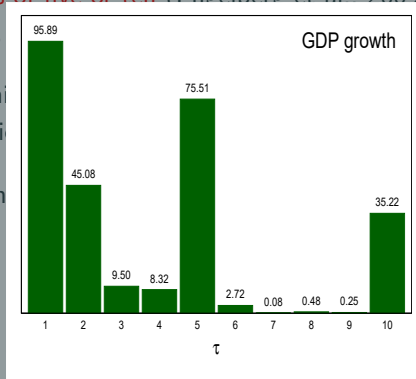
- Classify i as a **rounder** according to the following rule:

$$\tilde{D}_i^{m5} = \begin{cases} 1 & \text{if } \text{mode}(\tilde{D}_{i,1}^{m5}, \dots, \tilde{D}_{i,K}^{m5}) = 1 \text{ and} \\ 0 & \text{else.} \end{cases}$$

- Consider $D_i^{m5} = 1 - \tilde{D}_i^{m5}$ instead (focus on **non-rounders**)

Integer-based categorization

- Probabilistic survey forecasts often tend to contain numbers that are multiples of five or ten (Engelberg et al., 2009; Clements, 2011; Boero et al., 2012)



- Consider a histogram \tilde{D}_i^{m5} of probability mass function \tilde{D}_i^{m5} , consisting of N bins, indexed by $k = 1, \dots, K$.

- Check whether

- Classify i as a rounder if $\text{mode}(\tilde{D}_{i,1}^{\text{m5}}, \dots, \tilde{D}_{i,K}^{\text{m5}}) = 1$ and $\tilde{D}_{i,K}^{\text{m5}} > 0$. Rule:

$$\tilde{D}_i^{\text{m5}} = \begin{cases} 1 & \text{if } \text{mode}(\tilde{D}_{i,1}^{\text{m5}}, \dots, \tilde{D}_{i,K}^{\text{m5}}) = 1 \text{ and } \tilde{D}_{i,K}^{\text{m5}} > 0 \\ 0 & \text{else.} \end{cases}$$

- Consider $D_i^{\text{m5}} = 1 - \tilde{D}_i^{\text{m5}}$ instead (focus on non-rounders)

Integer-based categorization

- Probabilistic survey forecasts often tend to contain numbers that are **multiples of five or ten** (Engelberg et al., 2009; Clements, 2011; Boero et al., 2015)
- Consider a histogram forecast by forecaster $i = 1, \dots, N$, consisting of probabilities, $p_{i,k}$, for outcome intervals ('bins') $k = 1, \dots, K$.
- Check whether $p_{i,k}$ is a **multiple of five**:

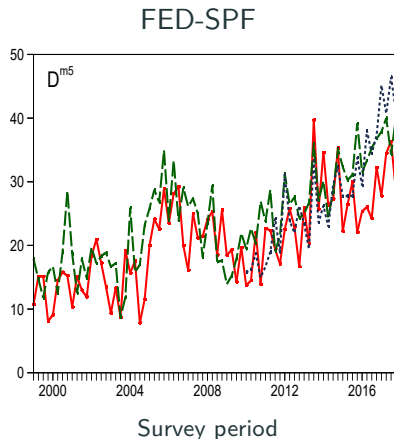
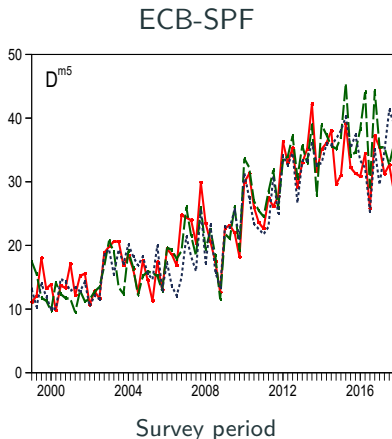
$$\tilde{D}_{i,k}^{m5} = \begin{cases} 1 & \text{if } 5 \cdot \lfloor \frac{p_{i,k}}{5} \rfloor = p_{i,k} \text{ and} \\ 0 & \text{else.} \end{cases}$$

- Classify i as a **rounder** according to the following rule:

$$\tilde{D}_i^{m5} = \begin{cases} 1 & \text{if } \text{mode}(\tilde{D}_{i,1}^{m5}, \dots, \tilde{D}_{i,K}^{m5}) = 1 \text{ and} \\ 0 & \text{else.} \end{cases}$$

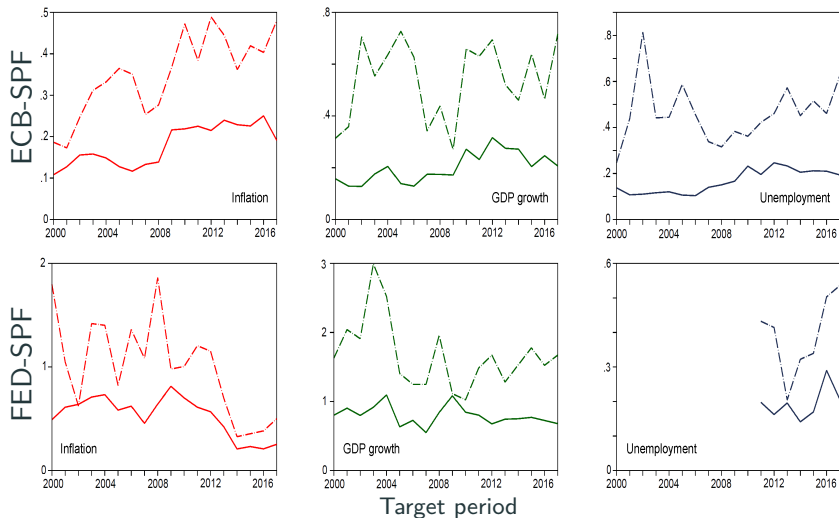
- Consider $D_i^{m5} = 1 - \tilde{D}_i^{m5}$ instead (focus on **non-rounders**)

Time-variation in the share of non-rounded histograms



Notes: These graphs depict the time series of the share of non-rounded observations in the ECB- (left) and FED-SPF (right) for **inflation**, **output growth** and **unemployment**. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Average ex-ante uncertainty for $h = 4$



Dashed lines: uncertainty from non-rounded forecasts,
solid lines: uncertainty from rounded forecasts.

Rounding and variance misalignment

Horizon-specific variance misalignment

- Individuals' ex-ante uncertainty at horizon h is computed as:

$$\overline{\sigma_{i,h}^2} = \frac{1}{T_{i,h}} \sum_{t=1}^{T_{i,h}} \sigma_{i,t,h}^2$$

- Ex-post uncertainty is defined as the mean squared error:

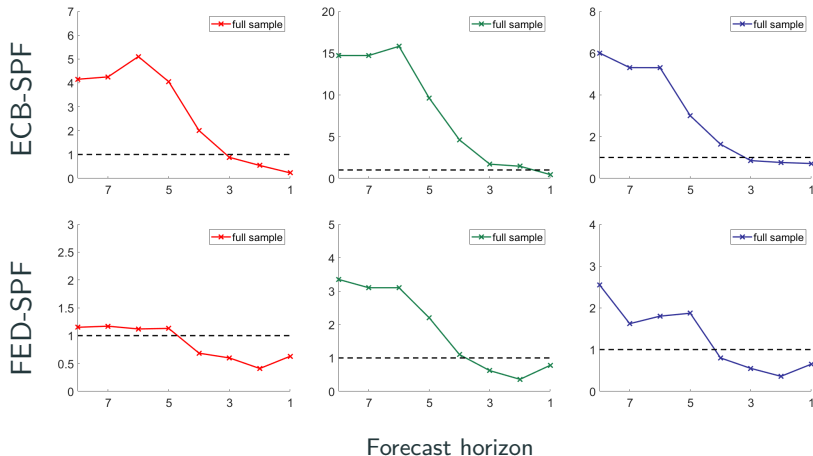
$$\text{MSE}_{i,h} = \frac{1}{T_{i,h}} \sum_{t=1}^{T_{i,h}} e_{i,t,h}^2,$$

where $e_{i,t,h} = x_t - \mu_{i,t,h}$ denotes forecast errors.

- **Average misalignment ratio** (Clements, 2014):

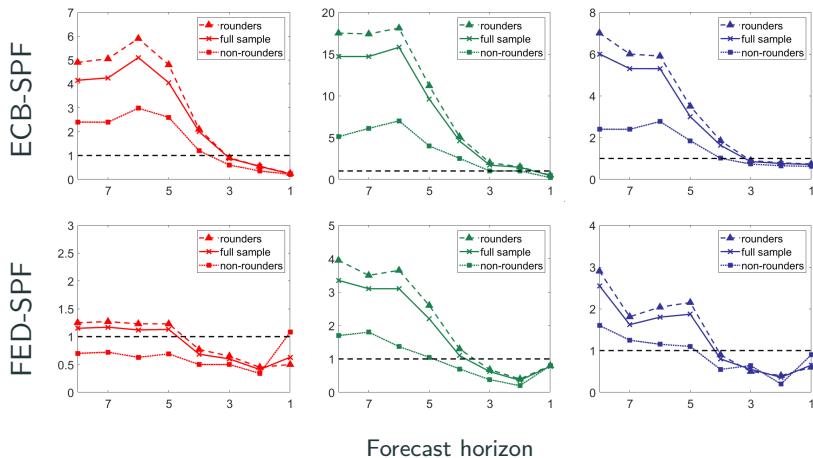
$$m_h = \frac{1}{N_h} \sum_{i=1}^{N_h} \frac{\text{MSE}_{i,h}}{\overline{\sigma_{i,h}^2}}$$

Variance misalignment



Notes: For each forecast horizon, the plots display the average misalignment ratios m_h for the ECB- (first row) and FED-SPF (second row). The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Variance misalignment



Notes: For each forecast horizon, the plots display the average misalignment ratios m_h for the ECB- (first row) and FED-SPF (second row). Non-rounders are classified by means of $D_{i,t,h}^{m5}$. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Sources of misalignment

For $h \in \{8, \dots, 1\}$, we estimate

$$y_{i,t,h} = \alpha_h + \beta_h D_{i,t,h}^{m5} + \gamma_{2,h} D2_{t,h} + \dots + \gamma_{T,h} DT_{t,h} + \varepsilon_{i,t,h}.$$

- Histogram characteristics $y_{i,t,h}$ include
 - variables related to ex-ante uncertainty ($K_{i,t,h}^*$, $\sigma_{i,t,h}^2$) and
 - variables related to ex-post uncertainty ($|e_{i,t,h}|$, $e_{i,t,h}^2$)
- $K_{i,t,h}^* \in \{0, \dots, K\}$: Number of bins containing nonzero probability
- Time-fixed effects $D2_{t,h}, \dots, DT_{t,h}$
- Variance-covariance estimator by Newey and West (1987)

Sources of misalignment

For $h \in \{8, \dots, 1\}$, we estimate

$$y_{i,t,h} = \alpha_h + \beta_h D_{i,t,h}^{\text{m5}} + \gamma_{2,h} D_{2,t,h} + \dots + \gamma_{T,h} D_{T,t,h} + \varepsilon_{i,t,h}.$$

- Histogram characteristics $y_{i,t,h}$ include

- variables related to $(K_{i,t,h}^*, \sigma_{i,t,h}^2)$ and
- variables related to $(|e_{i,t,h}|, e_{i,t,h}^2)$

Misalignment ratio

$$m_h = \frac{1}{N_h} \sum_{i=1}^{N_h} \frac{\text{MSE}_{i,h}}{\sigma_{i,h}^2}$$

- $K_{i,t,h}^* \in \{0, \dots, K\}$ number of bins with nonzero probability
- Time-fixed effects $D_{2,t,h}, \dots, D_{T,t,h}$
- Variance-covariance estimator by Newey and West (1987)

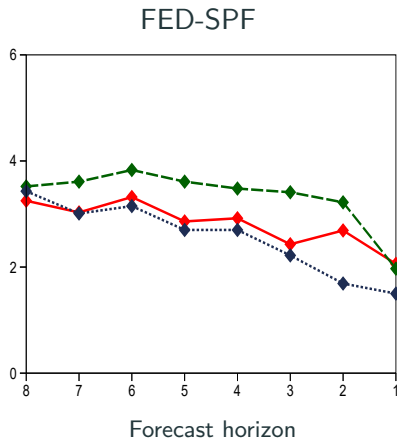
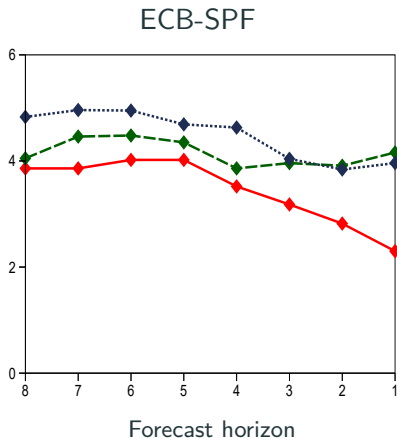
Sources of misalignment

For $h \in \{8, \dots, 1\}$, we estimate

$$y_{i,t,h} = \alpha_h + \beta_h D_{i,t,h}^{\text{m5}} + \gamma_{2,h} D2_{t,h} + \dots + \gamma_{T,h} DT_{t,h} + \varepsilon_{i,t,h}.$$

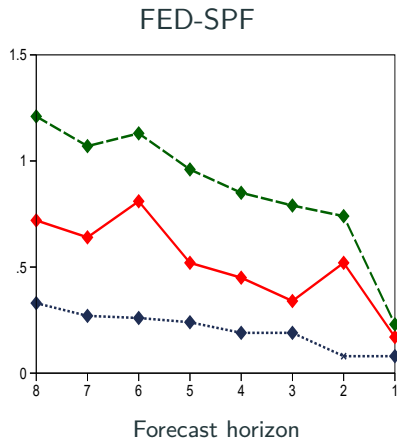
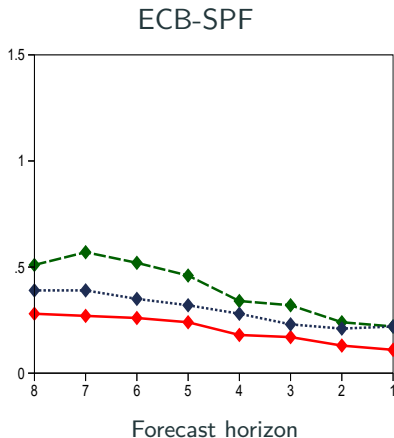
- Histogram characteristics $y_{i,t,h}$ include
 - variables related to ex-ante uncertainty ($K_{i,t,h}^*$, $\sigma_{i,t,h}^2$) and
 - variables related to ex-post uncertainty ($|e_{i,t,h}|$, $e_{i,t,h}^2$)
- $K_{i,t,h}^* \in \{0, \dots, K\}$: Number of bins containing nonzero probability
- Time-fixed effects $D2_{t,h}, \dots, DT_{t,h}$
- Variance-covariance estimator by Newey and West (1987)

Deviations in $K_{i,t,h}^*$



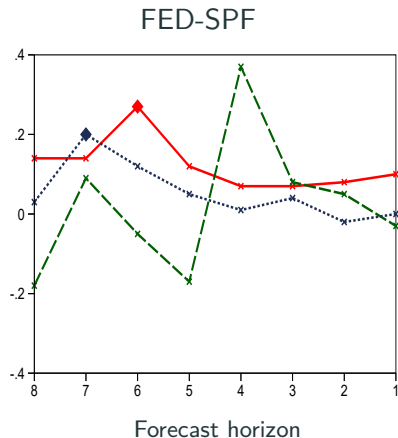
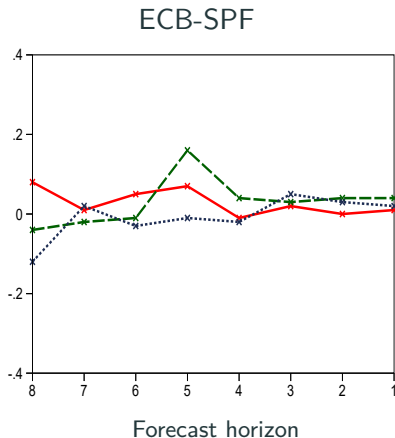
Notes: For each h , the graphs depict the estimates of β_h for **inflation**, **output growth** and **unemployment** in the ECB- (left) and FED-SPF (right) when $K_{i,t,h}$ is considered as the dependent variable. A diamond “ \diamond ” indicates that the estimate is significantly different from zero at the 5% level. A cross “ \times ” marks an insignificant estimate. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Deviations in $\sigma_{i,t,h}^2$



Notes: For each h , the graphs depict the estimates of β_h for inflation, output growth and unemployment in the ECB- (left) and FED-SPF (right) when $\sigma_{i,t,h}^2$ is considered as the dependent variable. A diamond “ \diamond ” indicates that the estimate is significantly different from zero at the 5% level. A cross “ \times ” marks an insignificant estimate. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Deviations in $e_{i,t,h}^2$



Notes: For each h , the graphs depict the estimates of β_h for inflation, output growth and unemployment in the ECB- (left) and FED-SPF (right) when $e_{i,t,h}^2$ is considered as the dependent variable. A diamond “ \diamond ” indicates that the estimate is significantly different from zero at the 5% level. A cross “ \times ” marks an insignificant estimate. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

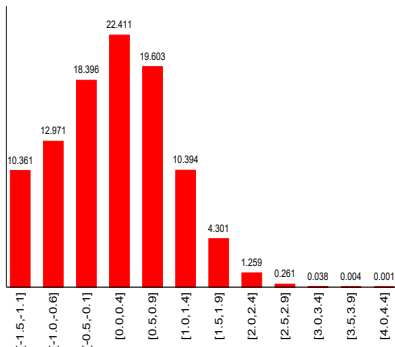
Discussion – Direct effect of rounding

Counterfactual rounding of non-rounded histograms

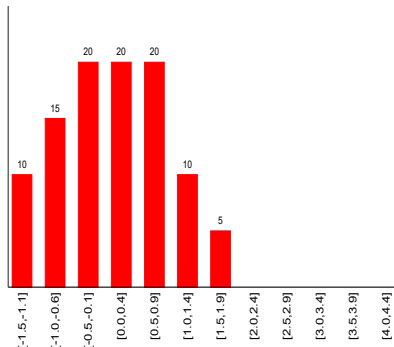
- We have seen: Rounding indicates deficiencies in ex-ante uncertainty.
- Rounded forecasts are often forecasts with very few bins.
- Next: Isolate direct effect of rounding.
- Strategy:
 - Consider forecasts that are non-rounded.
 - Round the reported probabilities in all bins to the nearest multiple of 5.
 - Obtain counterfactual uncertainty $\sigma_{i,t,h}^2(Co)$.
 - Compare reported ex-ante uncertainty $\sigma_{i,t,h}^2$ with $\sigma_{i,t,h}^2(Co)$.

Counterfactual rounding of non-rounded histograms

As reported:



Artificially rounded:

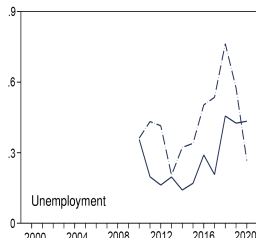
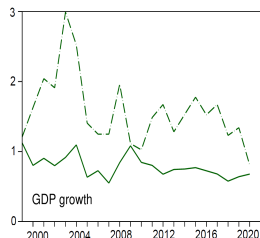
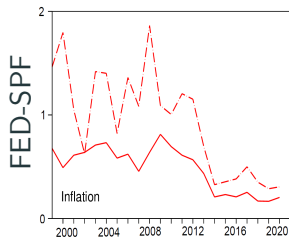
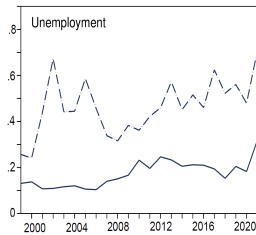
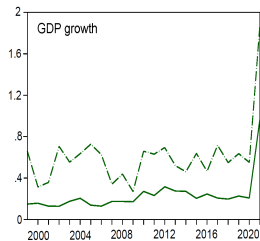
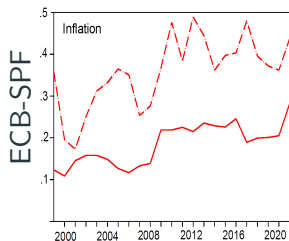


⇒ Both $K_{i,t,h}$ and $\sigma_{i,t,h}^2$ decline: $\sigma_{i,t,h}^2 = 0.72$, $\sigma_{i,t,h}^2(Co) = 0.67$.

⇒ Holds for the full sample as well: Reduction of 7%-10% (ECB-SPF) and 10%-11% (FED-SPF), depending on outcome variables.

Discussion – SPF forecasts in the COVID-19 environment

Average ex-ante uncertainty for $h = 4$ until 2021Q1



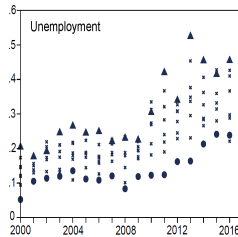
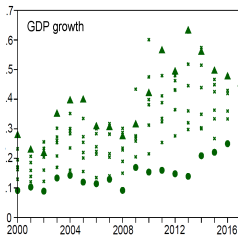
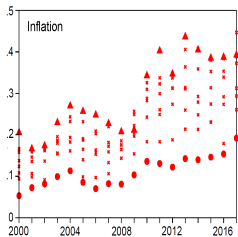
Target period

Dashed lines: uncertainty from non-rounded forecasts,
solid lines: uncertainty from rounded forecasts.

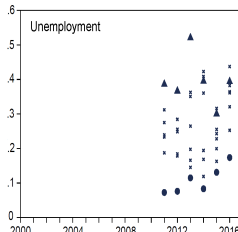
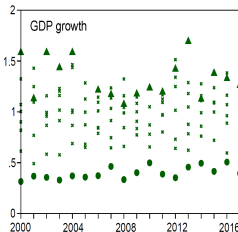
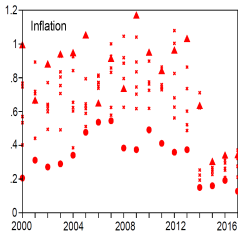
Discussion – Bin width and ex-ante uncertainty

Consensus ex-ante uncertainty

ECB-SPF



FED-SPF



Target period

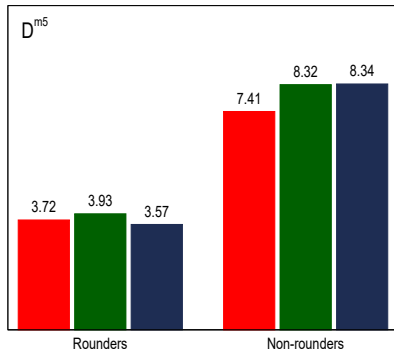
Notes: Graphs depict the cross-sectional averages over the variances derived from the individual histograms, i.e. $\sigma_{i,t,h}^2$, at forecast horizons $h \in \{8, 7, \dots, 1\}$. Triangles " \triangle " depict the 8-step-ahead variance forecasts and bullets " \bullet " the 1-step-ahead variance forecast. The sample period is 1999Q1-2017Q4.

Bin widths in percentage points

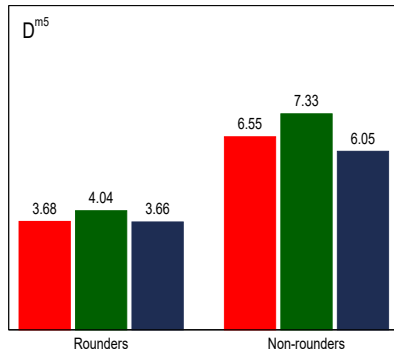
	Inflation	GDP growth	Unemployment rate
ECB-SPF	0.4	0.4	0.4
FED-SPF	0.9	0.9	0.4
	0.4 from 2014Q1		

Average numbers of bins

ECB-SPF



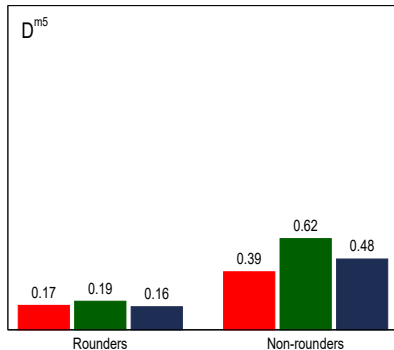
FED-SPF



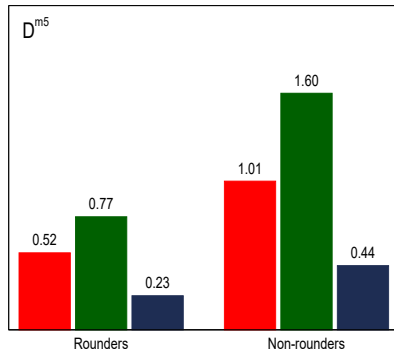
Notes: $\bar{K} = \frac{1}{NTH} \sum_{i=1}^N \sum_{t=1}^T \sum_{h=1}^H K_{i,t,h}$. Results are depicted for the ECB- (first row) versus FED-SPF (second row) and variables **inflation**, **output growth** and unemployment. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Average variance

ECB-SPF



FED-SPF



Notes: $\overline{\sigma^2} = \frac{1}{NTH} \sum_{i=1}^N \sum_{t=1}^T \sum_{h=1}^H \sigma_{i,t,h}^2$. Results are depicted for the ECB- (first row) versus FED-SPF (second row) and variables **inflation**, **output growth** and **unemployment**. The sample period is 1999Q1-2017Q4, except for the unemployment rate forecasts from the FED-SPF, which are available since 2010Q1 for our purposes.

Pooled ex-post uncertainty

Pooled MSE for the ECB- and the FED-SPF			
	Inflation	GDP growth	Unemployment
Full sample MSE			
ECB-SPF	0.46	1.59	0.48
FED-SPF	0.41	1.28	0.26
Excluding 2009			
ECB-SPF	0.37	0.77	0.37
FED-SPF	0.36	0.79	0.26

Cells contain pooled MSE statistics computed as

$$\overline{MSE} = \frac{1}{NTH} \sum_{i=1}^N \sum_{t=1}^T \sum_{h=1}^H e_{i,t,h}^2.$$

Conclusion

- Rounding as a reliable indicator of misalignment between ex-ante and ex-post variances
- Empirically, we confirm that rounders in the SPF...
 - ...use fewer of the available bins and report smaller ex-ante variances than the non-rounders
 - ...do not differ in terms of forecast errors from the non-rounders
- Consensus ex-ante uncertainty based solely on non-rounders is about twice as high as overall consensus, also distinct dynamics.
- Understanding behavior of rounders means explaining data features after 2020/2021.
- Role of bin width should be investigated – Not the aim of this paper.