

# Chapter 3 - Multi-parameter models: Exercise solutions

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Click on the arrow to see a solution.

## Exercise 3.1

Let  $X_1, \dots, X_n \mid \theta, \sigma^2 \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$ , where  $\theta$  is assumed known. Show that the  $\text{Inv-}\chi^2$  distribution is a conjugate prior for  $\sigma^2$ .

### Solution

The normal density function for a single observation  $x_i$  is

$$p(x_i \mid \theta, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x_i - \theta)^2\right) \propto \frac{1}{(\sigma^2)^{1/2}} \exp\left(-\frac{1}{2\sigma^2}(x_i - \theta)^2\right)$$

The likelihood for the iid normal model with known mean  $\theta$  is therefore the product of  $n$  such density functions:

$$\begin{aligned} p(x_1, \dots, x_n \mid \theta, \sigma^2) &= \prod_{i=1}^n p(x_i \mid \theta, \sigma^2) \propto \frac{1}{(\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta)^2\right) \\ &\propto \frac{1}{(\sigma^2)^{n/2}} \exp\left(-\frac{ns^2}{2\sigma^2}\right) \end{aligned}$$

where we have defined  $s^2 = \frac{\sum_{i=1}^n (x_i - \theta)^2}{n}$  as the sample standard deviation (dividing by  $n$  instead of  $n - 1$  since the mean  $\theta$  is assumed known).

The density of the  $\sigma^2 \sim \text{Scaled-Inv-}\chi^2(\nu_0, \sigma_0^2)$  prior is of the form

$$p(\sigma^2) \propto \frac{1}{(\sigma^2)^{1+\nu_0/2}} \exp\left(-\frac{\nu_0\sigma_0^2}{2\sigma^2}\right)$$

The posterior distribution from using the  $\sigma^2 \sim \text{Scaled-Inv-}\chi^2(\nu_0, \sigma_0^2)$  prior is given by Bayes' theorem as (to avoid cluttering the notation, we do not write out the conditioning

on  $\theta$  since it is known)

$$\begin{aligned}
p(\sigma^2|x_1, \dots, x_n) &\propto p(x_1, \dots, x_n|\sigma^2)p(\sigma^2) \\
&\propto \frac{1}{(\sigma^2)^{n/2}} \exp\left(-\frac{ns^2}{2\sigma^2}\right) \frac{1}{(\sigma^2)^{1+\nu_0/2}} \exp\left(-\frac{\nu_0\sigma_0^2}{2\sigma^2}\right) \\
&\propto \frac{1}{(\sigma^2)^{1+(\nu_0+n)/2}} \exp\left(-\frac{1}{2\sigma^2}(\nu_0\sigma_0^2 + ns^2)\right) \\
&= \frac{1}{(\sigma^2)^{1+(\nu_0+n)/2}} \exp\left(-\frac{\nu_0 + n}{2\sigma^2} \frac{\nu_0\sigma_0^2 + ns^2}{\nu_0 + n}\right)
\end{aligned}$$

which is proportional to a  $\sigma^2 \sim \chi^2(\nu_0 + n, \frac{\nu_0\sigma_0^2 + ns^2}{\nu_0 + n})$  density. Note how the location parameter in the posterior

$$\frac{\nu_0\sigma_0^2 + ns^2}{\nu_0 + n} = \frac{\nu_0}{\nu_0 + n}\sigma_0^2 + \frac{n}{\nu_0 + n}s^2$$

is a weighted average of prior location  $\sigma_0^2$  and the data estimate  $s^2$ , with more weight placed on the strongest information source (prior with  $\nu_0$  imaginary sample data points versus the data with  $n$  actual data points).