# State-Space Models Parameter inference and software

#### Mattias Villani

Statistiska institutionen Stockholms universitet

Institutionen för datavetenskap Linköpings universitet











#### Lecture overview

- **E**stimating model parameters
- Bayesian inference for the LGSS model
- Live demo of some R packages

The linear Gaussian state-space (LGSS) model

$$\begin{split} \text{Measurement eq:} \quad & \mathbf{y}_t = \mathbf{C} \mathbf{x}_t + \boldsymbol{\epsilon}_t \\ \text{State eq:} \quad & \boldsymbol{\epsilon}_t \overset{iid}{\sim} \textit{N}\left(\mathbf{0}, \boldsymbol{\Omega}_{\epsilon}\right) \\ \\ \text{State eq:} \quad & \boldsymbol{\epsilon}_t = \mathbf{A} \mathbf{x}_{t-1} + \mathbf{B} \mathbf{u}_t + \boldsymbol{\nu}_t \\ \end{split}$$

- The elements in A, B, C,  $\Omega_{\varepsilon}$  and  $\Omega_{\nu}$  may be unknown.
- Example: time-varying regression with p covariates  $z_t$   $(p \times 1)$

$$y_{t} = \mathbf{z}_{t}^{T} \boldsymbol{\beta}_{t} + \boldsymbol{\varepsilon}_{t}, \qquad \boldsymbol{\varepsilon}_{t} \stackrel{iid}{\sim} \boldsymbol{N}(0, \Omega_{\varepsilon})$$

$$\boldsymbol{\beta}_{1t} = \boldsymbol{a}_{1} \cdot \boldsymbol{\beta}_{1,t-1} + \boldsymbol{\nu}_{t} \qquad \boldsymbol{\nu}_{t} \stackrel{iid}{\sim} \boldsymbol{N}(0, \Omega_{\nu})$$

$$\vdots$$

$$\boldsymbol{\beta}_{pt} = \boldsymbol{a}_{p} \cdot \boldsymbol{\beta}_{p,t-1} + \boldsymbol{\nu}_{t} \qquad \boldsymbol{\nu}_{t} \stackrel{iid}{\sim} \boldsymbol{N}(0, \Omega_{\nu})$$

- ▶ Here  $C = \mathbf{z}_t^T$ ,  $\mathbf{x}_t = \beta_t$  and  $\mathbf{A} = \text{Diag}(a_1, ..., a_p)$ .
- The state space model's matrices (A etc) are parametrized by  $\theta = (\theta_1, ..., \theta_s)$ . To be explicit:  $A(\theta)$ ,  $B(\theta)$ , ...,  $\Omega_{\nu}(\theta)$ .

- Two options: Maximum likelihood estimate (MLE) or Bayesian.
- Likelihood function

$$\rho(y_1, ..., y_T | \theta) = \prod_{t=1}^{T} \rho(y_t | y_{1:t-1}, \theta)$$

How compute  $p(y_t|y_{1:t-1},\theta)$ ? The trick: i) condition on  $x_t$ , ii) exploit conditional independencies, iii) get rid of  $x_t$  by integrating it out:

$$p(y_t|y_{1:t-1}, \theta) = \int p(y_t|y_{1:t-1}, x_t, \theta) p(x_t|y_{1:t-1}, \theta) dx_t$$
  
= 
$$\int p(y_t|x_t, \theta) p(x_t|y_{1:t-1}, \theta) dx_t$$

- Note:
  - $p(x_t|y_{1:t-1},\theta) = \overline{bel}(x_t)$  is Gaussian
  - $\triangleright p(y_t|x_t,\theta)$  is Gaussian
  - $ightharpoonup p\left(\mathbf{y}_{t}|\mathbf{y}_{1:t-1}, \theta\right)$  is then also Gaussian [not obvious, but

- Remember: we are looking for the Gaussian  $p(y_t|y_{1:t-1}, \theta)$ .
- Mean by law of iterated expectations (E = EE)

$$\mathbb{E}\left(y_{t}|y_{1:t-1},\theta\right) = C\mathbb{E}\left(x_{t}|y_{1:t-1},\theta\right) = C\bar{\mu}_{t}$$

lacksquare Variance by conditional variance formula ( $V=\mathit{EV}+\mathit{VE}$ )

$$\begin{split} \mathbb{V}\left(\mathbf{y}_{t}|\mathbf{y}_{1:t-1},\theta\right) &= \mathbb{E}_{\mathbf{x}_{t}|\mathbf{y}_{1:t-1},\theta}\left[\mathbb{V}\left(\mathbf{y}_{t}|\mathbf{x}_{t},\mathbf{y}_{1:t-1},\theta\right)\right] \\ &+ \mathbb{V}_{\mathbf{x}_{t}|\mathbf{y}_{1:t-1},\theta}\left[\mathbb{E}\left(\mathbf{y}_{t}|\mathbf{x}_{t},\mathbf{y}_{1:t-1},\theta\right)\right] \\ &= \Omega_{\varepsilon} + \mathbb{V}_{\mathbf{x}_{t}|\mathbf{y}_{1:t-1},\theta}\left(\mathsf{C}\mathbf{x}_{t}\right) = \Omega_{\varepsilon} + \mathsf{C}\bar{\Sigma}_{t}\mathsf{C}^{T} \end{split}$$

In summary, the likelihood function is

$$p(\mathbf{y}_{1},...,\mathbf{y}_{T}|\theta) = \prod_{t=1}^{T} N\left(\mathbf{y}_{t}|\mathsf{C}\bar{\mu}_{t},\mathsf{C}\bar{\Sigma}_{t}\mathsf{C}^{T} + \Omega_{\varepsilon}\right)$$

where C,  $\Omega_{\varepsilon}$ ,  $\bar{\mu}_t$  and  $\bar{\Sigma}_t$  all depend on  $\theta$  generally.

- The Kalman filter gives us everything we need for  $p(y_1, ..., y_T | \theta)!$
- Numerical optimization (e.g. optim in R) to find MLE  $\hat{\theta}_{MLE}$ .
- Approximate  $\mathbb{V}\left(\hat{\theta}_{\textit{MLE}}\right)$  from the numerical Hessian.
- Sampling from the posterior distribution

$$p(\theta|y_1, ..., y_T) \propto p(y_1, ..., y_T|\theta) p(\theta)$$

by Metropolis-Hastings.

## State smoothing

Filtering (real time):

$$p(\mathbf{x}_t|\mathbf{y}_{1:t})$$

■ Smoothing (retrospective):

$$p(\mathbf{x}_t|\mathbf{y}_{1:T})$$

- Start at the end t = T. We already have  $p(x_T|y_{1:T})$  from the last iteration of the Kalman filter. Work yourself backward in time to obtain  $p(x_{T-1}|y_{1:T}), ..., p(x_1|y_{1:T})$ .
- Note: the end result are the **marginal** densities at any t,  $p(x_t|y_{1:T})$ . More work to do if one also wants  $p(x_{t_1}, x_{t_2}|y_{1:T})$  for some times  $t_1$  and  $t_2$ .

# State smoothing

- Algorithm Smoothing( $s_{t+1}, S_{t+1}, \mu_t, \Sigma_t, \bar{\mu}_{t+1}, \bar{\Sigma}_{t+1}$ )
  - Mean update:

$$s_t = \mu_t + \Sigma_t A^T \bar{\Sigma}_{t+1}^{-1} (s_{t+1} - \bar{\mu}_{t+1})$$

Covariance update:

$$S_t = \Sigma_t + \Sigma_t A^T \bar{\Sigma}_{t+1}^{-1} \left( S_{t+1} - \bar{\Sigma}_{t+1} \right) \bar{\Sigma}_{t+1}^{-1} A \Sigma_t$$

Return s<sub>t</sub>, S<sub>t</sub>

#### Bayesian inference for the state

- How to sample from posterior of the state  $p(x_1, ..., x_T | y_{1:T}, \theta)$ ?
- Simulate the state trajectory **backward in time** starting with  $x_T$ :

$$p(x_{1:T}|y_{1:T},\theta) = p(x_T|y_{1:T},\theta)p(x_{T-1}|x_T,y_{1:T},\theta)\cdots p(x_1|x_{T-1:2},y_{1:T},\theta)$$

- Forward Filtering Backward Sampling (FFBS):
  - ▶ Run the Kalman filter forward in time t = 1, ..., T.
  - ▶ Simulate  $x_T$  from  $N(\mu_T, \Sigma_T)$ .
  - ▶ Simulate states backward in time t = T 1, T 2, ..., 1:

$$x_t | x_{t+1:T}, y_{1:T}, \theta \sim N(h_t, H_t)$$

$$h_t = \mu_t + \Sigma_t A^T \bar{\Sigma}_{t+1}^{-1} (\mathbf{x}_{t+1} - \bar{\mu}_{t+1})$$

and

$$\mathsf{H}_t = \Sigma_t - \Sigma_t \mathsf{A}^T \bar{\Sigma}_{t+1}^{-1} \mathsf{A} \Sigma_t.$$

- Note: FFBS distributions conditions on x<sub>t+1:T</sub>.
- So the FFBS gives the *joint* (smoothing) posterior for  $x_{1:T}$ ,

### dlm package in R

The linear Gaussian state-space (LGSS) model

$$\begin{split} \text{Measurement eq:} \quad & \mathbf{y}_t = \mathbf{C}\mathbf{x}_t + \boldsymbol{\varepsilon}_t \\ \text{State eq:} \quad & \boldsymbol{\varepsilon}_t \overset{iid}{\sim} \textit{N}\left(\mathbf{0}, \Omega_{\boldsymbol{\varepsilon}}\right) \\ \text{State eq:} \quad & \boldsymbol{v}_t \overset{iid}{\sim} \textit{N}\left(\mathbf{0}, \Omega_{\boldsymbol{\nu}}\right) \end{split}$$

In the dlm package

$$\begin{array}{ll} \text{Measurement eq:} & Y_t = F\theta_t + v_t \\ & \qquad \qquad \epsilon_t \stackrel{\textit{iid}}{\sim} \textit{N}\left(0, \textit{V}\right) \\ \\ \text{State eq:} & \quad \theta_t = G\theta_{t-1} + w_t \\ & \qquad \qquad v_t \stackrel{\textit{iid}}{\sim} \textit{N}\left(0, \textit{W}\right) \\ \end{array}$$

- $\theta_t$  is the state vector in dlm.  $Y_t$  are the measurements (a vector).
- The dlm notation goes back to West and Harrison's yellow DLM bible.
- The state is an unknown, so it is a greek letter.
- Measurements is a random variable so it is a capital letter.
- $\blacksquare$  dlm can also handle when F, G, V, W vary of over time.

# dlm package in R

DLM

$$\begin{array}{ll} \text{Measurement eq:} \quad Y_t = F\theta_t + v_t & \qquad \epsilon_t \stackrel{\textit{iid}}{\sim} \textit{N}\left(0,\textit{V}\right) \\ \\ \text{State eq:} \quad \theta_t = G\theta_{t-1} + w_t & \qquad \nu_t \stackrel{\textit{iid}}{\sim} \textit{N}\left(0,\textit{W}\right) \end{array}$$

- Main functions:
  - ▶ dlm creates the dlm model object
  - dlmFilter Kalman filtering
  - dlmSmooth State smoothing
  - dlmLL computes the log-likelihood