State-Space Models Models, Applications and State inference

Mattias Villani

Statistiska institutionen Stockholms universitet

Institutionen för datavetenskap Linköpings universitet











Lecture overview

- Time varying parameter models
- State space models
- The Bayes filter
- The Kalman filter

Autoregressive time series models

Autoregressive process (AR) for time series

$$y_t = \rho y_{t-1} + \varepsilon_t, \qquad \varepsilon_t \stackrel{iid}{\sim} N(0, \sigma^2)$$

The joint distribution for the whole time sequence $y_1, y_2, ..., y_T$ factorizes as

$$p(y_1,...,y_T) = p(y_1)p(y_2|y_1)\cdots p(y_T|y_{T-1})$$

where

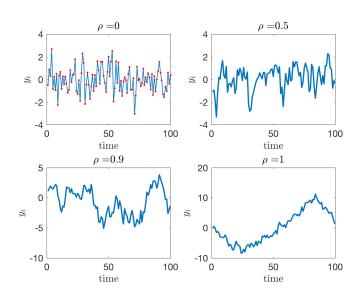
$$y_t|y_{t-1} \sim N\left(\rho y_{t-1}, \sigma^2\right)$$
.

 \blacksquare AR(p) process

$$y_t | y_{t-1}, ..., y_{t-p} \sim N \left(\sum_{j=1}^p \rho_j y_{t-j}, \sigma^2 \right).$$

 \blacksquare ARIMA(p, q).

Autoregressive time series models



Hidden Markov models

Two **regimes** defined by **latent** (hidden) variable $x_t \in \{1, 2\}$

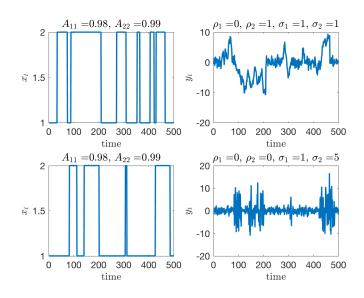
$$y_t = \begin{cases} \rho_1 y_{t-1} + \varepsilon_t, & \quad \varepsilon_t \stackrel{\textit{iid}}{\sim} N(0, \sigma_1^2) & \textit{if } z_t = 1\\ \rho_2 y_{t-1} + \varepsilon_t, & \quad \varepsilon_t \stackrel{\textit{iid}}{\sim} N(0, \sigma_2^2) & \textit{if } z_t = 2 \end{cases}$$

 \blacksquare x_t follows a Markov chain. Transition from state $j \to k$

$$\Pr\left(x_t = k | x_{t-1} = j\right) = A_{jk}$$

But what if changes in parameters appear more gradual?

Hidden Markov models



Time varying parameter models

Smoothly time varying parameter model

$$y_{t} = \rho_{t} y_{t-1} + \varepsilon_{t} \qquad \varepsilon_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\varepsilon}^{2}\right)$$
$$\rho_{t} = \rho_{t-1} + \nu_{t} \qquad \nu_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\nu}^{2}\right)$$

- The persistence parameter ρ is a **latent** (hidden) continuous variable that evolves over time (random walk).
- More generally, for some $-1 \le a < 1$,

$$y_{t} = \rho_{t} y_{t-1} + \varepsilon_{t} \qquad \varepsilon_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\varepsilon}^{2}\right)$$

$$\rho_{t} = a \rho_{t-1} + \nu_{t} \qquad \nu_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\nu}^{2}\right)$$

Time varying variance

$$\begin{aligned} y_t &= \rho y_{t-1} + \varepsilon_t & \quad \varepsilon_t \overset{\textit{iid}}{\sim} \textit{N}\left(0, \sigma_{\varepsilon, t}^2\right) \\ \ln \sigma_{\varepsilon, t}^2 &= \ln \sigma_{\varepsilon, t-1}^2 + \nu_t & \quad \nu_t \overset{\textit{iid}}{\sim} \textit{N}\left(0, \sigma_{\nu}^2\right) \end{aligned}$$

Time varying parameter models

Smoothly time varying parameter regression

$$\begin{aligned} y_t &= \mathbf{x}_t^\mathsf{T} \boldsymbol{\beta}_t + \boldsymbol{\varepsilon}_t & \quad \boldsymbol{\varepsilon}_t \stackrel{iid}{\sim} N\left(0, \sigma_\varepsilon^2\right) \\ \boldsymbol{\beta}_t &= \boldsymbol{\beta}_{t-1} + \boldsymbol{\nu}_t & \quad \boldsymbol{\nu}_t \stackrel{iid}{\sim} N\left(0, \sigma_v^2\right) \end{aligned}$$

- Smoothly time varying parameter survival model
- The hazard function (conditional probability of death at time t):

$$\begin{split} \lambda(t|\mathbf{x}) &= \lambda_0(t) \cdot \exp\left(\mathbf{x}^T \boldsymbol{\beta}_t\right) \\ \boldsymbol{\beta}_t &= \boldsymbol{\beta}_{t-1} + \boldsymbol{\nu}_t \qquad \boldsymbol{\nu}_t \overset{\textit{iid}}{\sim} \textit{N}\left(\mathbf{0}, \sigma_{\boldsymbol{\nu}}^2\right) \end{split}$$

And so on ...

Unobserved components models

- Model a time series as components: mean, trend, season, cycles etc.
- Local level model

$$y_{t} = \mu_{t} + \varepsilon_{t} \qquad \varepsilon_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\varepsilon}^{2}\right)$$

$$\mu_{t} = \mu_{t-1} + \nu_{t} \qquad \nu_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\nu}^{2}\right)$$

Local trend model

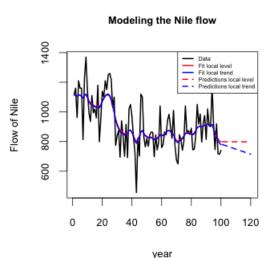
$$y_{t} = \mu_{t} + \varepsilon_{t} \qquad \varepsilon_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\varepsilon}^{2}\right)$$

$$\mu_{t} = \mu_{t-1} + \beta_{t-1} + \nu_{t} \qquad \nu_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\nu}^{2}\right)$$

$$\beta_{t} = \beta_{t-1} + \eta_{t} \qquad \eta_{t} \stackrel{iid}{\sim} N\left(0, \sigma_{\eta}^{2}\right)$$

Unobserved components models

■ See my code UnobservedComponentsModel.R



State-space models

Basic state-space model

Measurement eq:
$$y_t = Cx_t + \varepsilon_t$$
 $\varepsilon_t \stackrel{iid}{\sim} N\left(0, \sigma_{\varepsilon}^2\right)$
State eq: $x_t = Ax_{t-1} + \nu_t$ $\nu_t \stackrel{iid}{\sim} N\left(0, \sigma_{\nu}^2\right)$

- Measurements y_t are driven by an underlying unobserved state x_t .
- Time-varying parameter models: $x_t = \rho_t$.
- Hidden Markov models are state space models with a discrete state variable.
- Example 1: x_t is employment at time t. y_t are labor force survey estimates.
- Example 2: x_t is democrats' voting share. y_t are results from poll.
- Example 3: x_t is the position of flying vehicle at time t. y_t are sensor measurements.

Local trend model is a state space model

■ The linear Gaussian state-space (LGSS) model

$$\begin{split} \text{Measurement eq:} \ \ \mathbf{y}_t &= \mathbf{C}\mathbf{x}_t + \boldsymbol{\varepsilon}_t \qquad \boldsymbol{\varepsilon}_t \stackrel{iid}{\sim} \textit{N}\left(\mathbf{0}, \Omega_{\boldsymbol{\varepsilon}}\right) \\ \text{State eq:} \ \ \mathbf{x}_t &= \mathbf{A}\mathbf{x}_{t-1} + \mathbf{B}\mathbf{u}_t + \boldsymbol{\nu}_t \qquad \quad \boldsymbol{\nu}_t \stackrel{iid}{\sim} \textit{N}\left(\mathbf{0}, \Omega_{\boldsymbol{\nu}}\right) \end{split}$$

Local trend model

$$\begin{aligned} y_t &= \mu_t + \varepsilon_t & \varepsilon_t \stackrel{iid}{\sim} N\left(0, \sigma_\varepsilon^2\right) \\ \mu_t &= \mu_{t-1} + \beta_{t-1} + \nu_t & \nu_t \stackrel{iid}{\sim} N\left(0, \sigma_\nu^2\right) \\ \beta_t &= \beta_{t-1} + \eta_t & \eta_t \stackrel{iid}{\sim} N\left(0, \sigma_\eta^2\right) \end{aligned}$$

State space formulation

$$\mathbf{x}_t = \left(\begin{array}{c} \mu_t \\ \beta_t \end{array} \right), \mathbf{A} = \left(\begin{array}{cc} 1 & 1 \\ 0 & 1 \end{array} \right), \mathbf{C} = \left(\begin{array}{cc} 1 & 0 \end{array} \right), \Omega_{\epsilon} = \sigma_{\epsilon}^2, \Omega_{\nu} = \left(\begin{array}{cc} \sigma_{\nu 1}^2 & 0 \\ 0 & \sigma_{\nu 2}^2 \end{array} \right)$$

The posterior distribution of the state

■ The linear Gaussian state-space (LGSS) model

Aim: the posterior distribution of the state at time t

$$p(x_t|y_1, ..., y_T, u_1, ..., u_T)$$

- Also called the **smoothing distribution**.
- The joint smoothing distribution

$$p(x_1, ..., x_T | y_1, ..., y_T, u_1, ..., u_T)$$

More on this later.

Model structure

■ The linear Gaussian state-space (LGSS) model

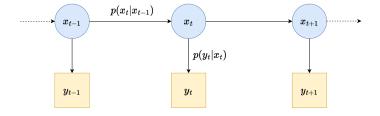
Note 1: x_t is first order Markov:

$$p(x_t|x_{t-1},...,x_1) = p(x_t|x_{t-1})$$

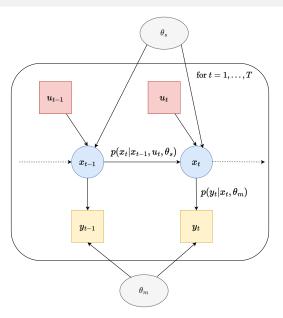
Note 2: **Conditional** on x_t , y_t is independent of past observations and states.



State-space model as graphical model



State-space model as graphical model



The filtering distribution

- Short hand notation: $x_{1:t} = \{x_1, ..., x_t\}$.
- Aim: the filtering distribution of the state at time t

$$p(\mathsf{x}_t|\mathsf{y}_{1:t},\mathsf{u}_{1:t})$$

Filtering = Posterior for x_t , after observing y_t

$$\pi(\mathsf{x}_t) \equiv p(\mathsf{x}_t|\mathsf{y}_{1:t},\mathsf{u}_{1:t})$$

Prior for x_t , **before** observing y_t

$$\bar{\pi}(\mathsf{x}_t) \equiv p(\mathsf{x}_t|\mathsf{y}_{1:t-1},\mathsf{u}_{1:t})$$

Recall: smoothing distribution conditions on all data

$$p(x_t|y_{1:T}, u_{1:T})$$

The Bayes filter

- We are now at time t.
- We have just given the control command u_t .
- We have not yet observed y_t.
- Prior beliefs at this stage:

$$\bar{\pi}(\mathsf{x}_t) = \int \rho(\mathsf{x}_t|\mathsf{u}_t,\mathsf{x}_{t-1}) bel(\mathsf{x}_{t-1}) d\mathsf{x}_{t-1}$$

- Now comes the observation y_t .
- Update to posterior using Bayes' theorem:

$$\pi(\mathbf{x}_t) \propto p(\mathbf{y}_t|\mathbf{x}_t)\bar{\pi}(\mathbf{x}_t).$$

The Bayes filter

Prediction step (control update)

$$\bar{\pi}(\mathsf{x}_t) = \int \rho(\mathsf{x}_t|\mathsf{u}_t,\mathsf{x}_{t-1})\pi(\mathsf{x}_{t-1})d\mathsf{x}_{t-1}$$

■ Measurement update step

$$\pi(\mathbf{x}_t) \propto p(\mathbf{y}_t|\mathbf{x}_t)\bar{\pi}(\mathbf{x}_t).$$

The Kalman filter

- The Kalman filter is the special case of the Bayes filter for the linear Gaussian state-space (LGSS) model.
- Under linearity and Gaussianity:
 - we can compute the integral in the prediction step analytically
 - ▶ the posterior in the measurement update becomes Gaussian
- Prediction update

$$\bar{\pi}(\mathsf{x}_t) = N(\bar{\mu}_t, \bar{\Sigma}_t)$$

Measurement update

$$\pi(\mathbf{x}_t) = N(\mu_t, \Sigma_t)$$

The Kalman filter tells us how to iteratively compute the sequences $\{\mu_t, \Sigma_t\}$ throughout time t = 1, ..., T.



The Kalman filter

■ The linear Gaussian state-space (LGSS) model

$$\begin{split} \text{Measurement eq:} \quad & \mathbf{y}_t = \mathbf{C}\mathbf{x}_t + \boldsymbol{\varepsilon}_t \\ \text{State eq:} \quad & \boldsymbol{\varepsilon}_t \overset{iid}{\sim} \textit{N}\left(\mathbf{0}, \Omega_{\boldsymbol{\varepsilon}}\right) \\ \text{V}_t \overset{iid}{\sim} \textit{N}\left(\mathbf{0}, \Omega_{\boldsymbol{v}}\right) \end{split}$$

Algorithm KalmanFilter(μ_{t-1} , Σ_{t-1} , u_t , y_t)

Prediction update:
$$\begin{cases} \bar{\mu}_t = \mathsf{A}\mu_{t-1} + \mathsf{Bu}_t \\ \bar{\Sigma}_t = \mathsf{A}\Sigma_{t-1}\mathsf{A}^T + \Omega_{\nu} \end{cases}$$

- $\text{Measurement update}: \begin{cases} \mathsf{K}_t = \bar{\Sigma}_t \mathsf{C}^T \left(\mathsf{C}\bar{\Sigma}_t \mathsf{C}^T + \Omega_\varepsilon \right)^{-1} \\ \mu_t = \bar{\mu}_t + \mathsf{K}_t (\mathsf{y}_t \mathsf{C}\bar{\mu}_t) \\ \Sigma_t = (\mathsf{I} \mathsf{K}_t \mathsf{C}) \bar{\Sigma}_t \end{cases}$
- ▶ Return $μ_t, Σ_t$

The Kalman filter code - single update

```
function kalmanfilter_update(\mu, \Sigma, u, y, A, B, C, Q, R)

# Prediction step - moving state forward without new measurement \bar{\mu} = A*\mu + B*u; \bar{\Sigma} = A*\Sigma*A' + R;

# Measurement update - updating the N(\bar{\mu}, \bar{\Sigma}) prior with the new data point K = \bar{\Sigma}*C' / (C*\bar{\Sigma}*C' + Q); # Kalman Gain \mu = \bar{\mu} + K*(y - C*\bar{\mu}); \Sigma = (I(length(\mu)) - K*C)*\bar{\Sigma};

return \mu, \Sigma end
```

The Kalman filter code

```
function kalmanfilter (U, Y, A, B, C, Q, R, \mu_o, \Sigma_o)
     # Prelims
     T = size(Y,1)
     n = length(\mu_o)
     \mu_{all} = zeros(T,n)
     \Sigma_{all} = zeros(n,n,T)
     \mu = \mu_0
     \Sigma = \Sigma_{o}
          \mu, \Sigma = kalmanfilter update(\mu, \Sigma, U[t,:]', Y[t,:]', A, B, C, Q, R)
          \mu_all[t,:] = \mu
          \Sigma all[:,:,t] = \Sigma
     return \mu all, \Sigma all
end
```

Kalmar filter intuition

Assume everything is univariate and no control:

$$\begin{array}{ll} \text{Measurement eq:} \quad y_t = c x_t + \varepsilon_t & \quad \varepsilon_t \stackrel{\textit{iid}}{\sim} \textit{N}\left(0, \omega_\varepsilon^2\right) \\ \text{State eq:} \quad x_t = a x_{t-1} + \nu_t & \quad \nu_t \stackrel{\textit{iid}}{\sim} \textit{N}\left(0, \omega_\nu^2\right) \end{array}$$

- ▶ Algorithm KalmanFilter(μ_{t-1} , σ_{t-1}^2 , y_t)
- Prediction update: $\begin{cases} \bar{\mu}_t = a\mu_{t-1} \\ \bar{\sigma}_t = a^2\sigma_{t-1}^2 + \omega_{\nu}^2 \end{cases}$

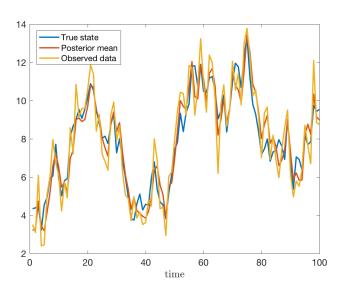
A simulated example

■ The linear Gaussian state-space (LGSS) model

Measurement eq:
$$y_t = x_t + \varepsilon_t$$
 $\varepsilon_t \stackrel{iid}{\sim} N\left(0,1\right)$ State eq: $x_t = 0.9x_{t-1} + u_t + v_t$ $v_t \stackrel{iid}{\sim} N\left(0,0.5\right)$

- Control: $u_t \sim |r_t|$ where $r_t \sim N(0, 1)$.
- T = 100.
- Initial state value: $x_0 \sim N(0, 10^2)$.

Data, state and posterior of state



Posterior intervals for the state

