

# Locally Stable Time-Varying Multi-Seasonal ARMA Models

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# Collaborators

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- **Robert Kohn**, University of New South Wales

# Structure of the talk

- Time-varying multi-seasonal AR models
  - Parameter evolution process
  - Bayesian inference
  - Simulation experiments
  - Application
- 
- Slides: <http://mattiasvillani.com/news>
  - Paper: [arXiv: 2409.18640](https://arxiv.org/abs/2409.18640) (ARMA extension soon on arXiv)

# Motivation

- Box-Jenkins. Global stationarity is restrictive.
- Differencing. Information loss. Overdifferencing. Levels are non-stationary. Poor short-run forecasts.
- Data in most recent regime. Wasteful. No probability for regime change.
  
- Time-varying parameters has a long history. [1], [2], [3]
- Recently: more realistic parameter evolutions. [4], [5], [6], [7]
- Parameters can enter explosive region for some periods.
  
- Seasonality is predictable. Model it! Time-varying?
- Time series with multiple seasonal periods is now common.

# Multi-seasonal AR models

- Seasonal AR( $p, P$ ) with season  $s$

$$\phi_p(L)\Phi_P(L^s)(y_t - \mu) = \varepsilon_t, \quad \varepsilon_t \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

- Lag operator:  $L^j y_t = y_{t-j}$

- Regular lag polynomial

$$\phi_p(L) = 1 - \phi_1 L - \dots - \phi_p L^p$$

- Seasonal lag polynomial

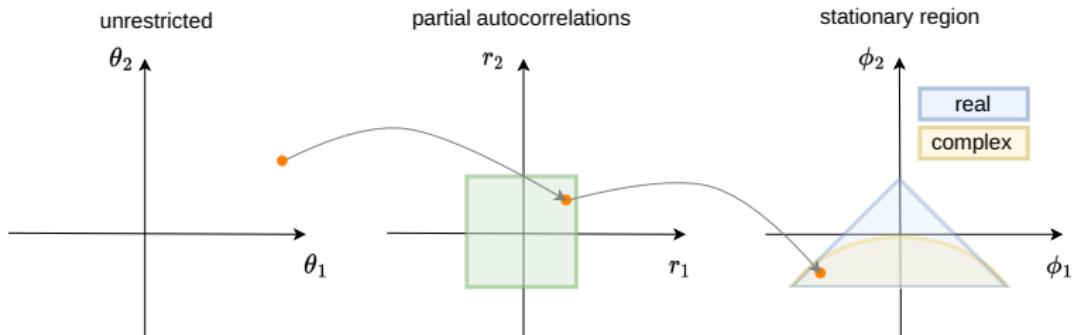
$$\Phi_p(L^s) = 1 - \Phi_1 L^s - \dots - \Phi_P L^{Ps}$$

- Multiple seasonal periods, e.g. hourly data with daily, weekly and yearly cycles.

- Multi-seasonal AR models with  $M$  polynomials

$$\prod_{j=1}^M \phi_j(L^{s_j})(y_t - \mu) = \varepsilon_t$$

# Enforce stability in AR models



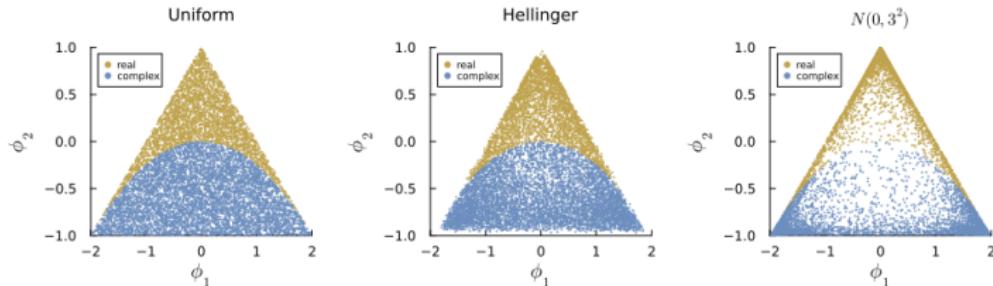
■ Unrestr. params  $\theta_k \in \mathbb{R} \Rightarrow$  Partial autocorr  $r_k \in (-1, 1)^p$

$$r_k = \frac{\theta_k}{(1 + \theta_k^2)^{1/2}}$$

- Partial autocorrs  $r \Rightarrow$  stable AR coeffs  $\phi$  by a recursion [8]
- Same parameterization for all polynomial factors.
- Invertibility in MA by the same parameterization.

# Uniform distribution over stability region

- Unrestricted parameters  $\theta_k$  have no interpretation. Priors?
- Uniform distribution over stability region  $\mathbb{S}_p$  for  $\phi$ .



- Lemma: If, independently,

$$\theta_k \sim \begin{cases} t(k+1, 0, \frac{1}{\sqrt{k+1}}) & \text{if } k \text{ is odd} \\ t_{\text{skew}}\left(\frac{k}{2}, \frac{k+2}{2}, 0, \frac{1}{\sqrt{k+1}}\right) & \text{if } k \text{ is even,} \end{cases}$$

then  $\phi = (\phi_1, \dots, \phi_p)^\top$  is uniformly distributed over  $\mathbb{S}_p$ .

## Time-varying multi-seasonal AR

- Time-varying multi-seasonal AR

$$\prod_{j=1}^M \phi_{jt}(L^{s_j})(y_t - \mu_t) = \varepsilon_t \quad \varepsilon_t \stackrel{\text{iid}}{\sim} N(0, \sigma_t^2)$$

$\phi_{jkt}$  is  $k$ th AR parameter in  $j$ th polynomial at time  $t$ .

- $\phi_t = (\phi_{1t}, \dots, \phi_{p_j t})^\top$  parameters in AR polynomial at time  $t$ .
- $\phi_t = \mathbf{g}(\theta_t)$  map from unrestricted  $\theta_t$  to stable  $\phi_t$ .

# Dynamic shrinkage process priors for TVSAR

## ■ Dynamic shrinkage process prior [6] for TVSAR

$$\phi_t = \mathbf{g}(\boldsymbol{\theta}_t)$$

$$\theta_{kt} = \theta_{k,t-1} + \nu_{kt}, \quad \nu_{kt} \stackrel{\text{indep}}{\sim} N(0, \exp(h_{kt}))$$

$$h_{kt} = \mu_k + \kappa_k(h_{k,t-1} - \mu_k) + \eta_{kt}, \quad \eta_{kt} \stackrel{\text{iid}}{\sim} Z(1/2, 1/2, 0, 1)$$

## ■ Global log-volatility $\mu_k$

## ■ Local log-volatility $\eta_{kt}$ from heavy-tailed $Z$ -distr.

■ Constant periods, periods of rapid change and jumps.

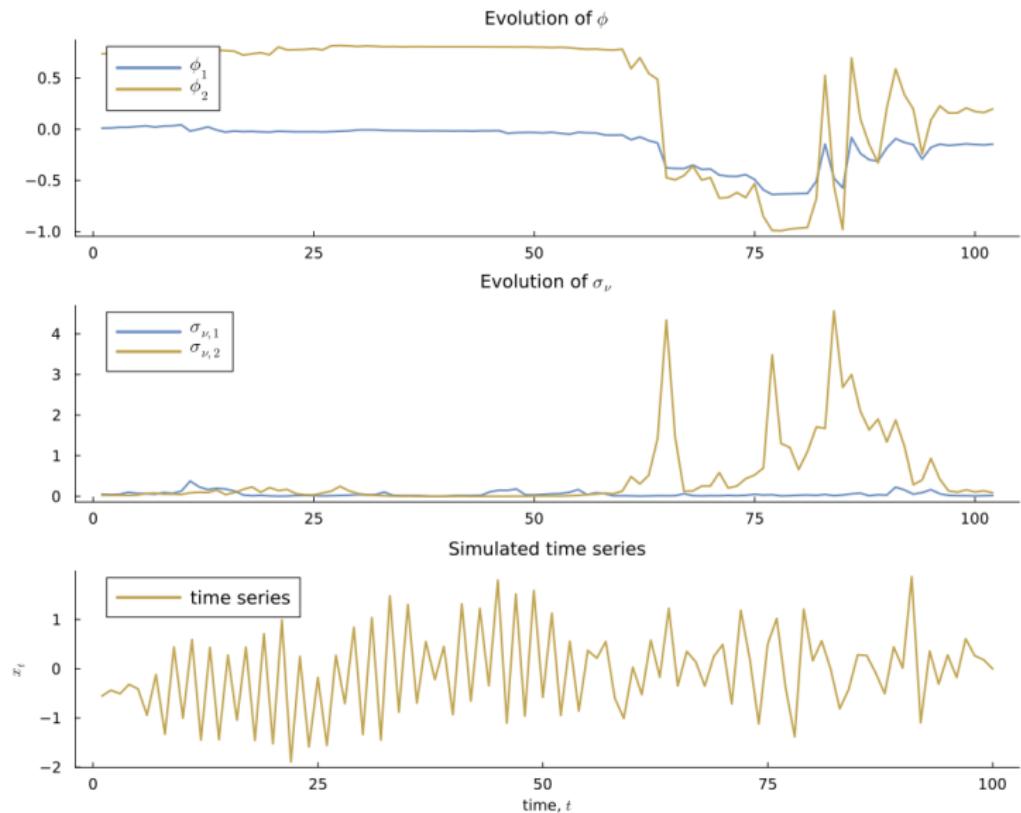
■ Time-series extension of **horseshoe prior** [9], since for  $\kappa = 0$

$$\text{Var}(\nu_t) = \exp(h_t) = \exp(\mu) \exp(\eta_t) = \tau^2 \lambda_t^2$$

and

$$\eta_t \sim Z(1/2, 1/2, 0, 1) \implies \lambda_t = \exp(\eta_t/2) \sim C^+(0, 1)$$

# Dynamic shrinkage process priors for TVSAR



## Conditional likelihood function

- Multi-seasonal TVSAR can be expressed as **regression** with **nonlinear restrictions** by multiplying out polynomials in

$$\prod_{j=1}^M \phi_{jt}(L^{s_j}) y_t = \varepsilon_t$$

- Example (static case):

$$(1 - \phi_1 L)(1 - \Phi_1 L^s) y_t = (1 - \phi_1 L - \Phi_1 L^s + \phi_1 \Phi_1 L^{1+s}) y_t$$

$$y_t = \mathbf{x}_t^\top \tilde{\boldsymbol{\phi}} + \varepsilon_t$$

- ▶  $\mathbf{x}_t = (y_{t-1}, y_{t-s}, y_{t-(1+s)})^\top$
- ▶  $\tilde{\boldsymbol{\phi}} = (\phi_1, \Phi_1, \phi_1 \Phi_1)^\top$ .
- ▶ Lags 2, ...,  $s-1$  have zero coefficients.

# Bayesian inference TVSAR

- TVSAR with conditional likelihood:

$$y_t = \mathbf{x}_t^\top \tilde{\boldsymbol{\phi}}_t + \varepsilon_t \quad \varepsilon_t \stackrel{\text{indep}}{\sim} N(0, \sigma_t^2)$$

$$\tilde{\boldsymbol{\phi}}_t = \tilde{\mathbf{g}}(\boldsymbol{\theta}_t)$$

$$\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} + \boldsymbol{\nu}_t, \quad \boldsymbol{\nu}_t \stackrel{\text{indep}}{\sim} N(0, \text{Diag}(\exp(\mathbf{h}_t)))$$

$$\mathbf{h}_t = \boldsymbol{\mu} + \boldsymbol{\kappa}(\mathbf{h}_{t-1} - \boldsymbol{\mu}) + \boldsymbol{\eta}_t, \quad \boldsymbol{\eta}_t \stackrel{\text{iid}}{\sim} Z(1/2, 1/2, 0, 1)$$

- Aim: posterior conditional on time series  $y_{1:T}$

$$p(\boldsymbol{\theta}_{0:T}, \mathbf{h}_{1:T}, \boldsymbol{\mu}, \boldsymbol{\kappa}, \sigma_{1:T} | y_{1:T})$$

- Gibbs sampling by data augmentation:

- ▶ Mixture of normal for  $\log \chi_1^2$  for volatility  $\mathbf{h}_{0:T}$  [10]
- ▶ Polya-Gamma augmentation for  $Z$ -distribution. [6]

- After augmentation: **transition model is linear Gaussian.**

# Bayesian inference TVSARMA

## ■ TVSARMA with exact likelihood:

$$y_t = \mathbf{x}_t^\top \tilde{\boldsymbol{\phi}}_t + \mathbf{z}_t^\top \tilde{\boldsymbol{\psi}}_t + \varepsilon_t, \quad \varepsilon_t \stackrel{\text{indep}}{\sim} N(0, \sigma_t^2)$$
$$(\tilde{\boldsymbol{\phi}}_t, \tilde{\boldsymbol{\psi}}_t) = \tilde{\mathbf{g}}(\boldsymbol{\theta}_t)$$
$$\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} + \boldsymbol{\nu}_t, \quad \boldsymbol{\nu}_t \stackrel{\text{indep}}{\sim} N(0, \text{Diag}(\exp(\mathbf{h}_t)))$$
$$\mathbf{h}_t = \boldsymbol{\mu} + \boldsymbol{\kappa}(\mathbf{h}_{t-1} - \boldsymbol{\mu}) + \boldsymbol{\eta}_t, \quad \boldsymbol{\eta}_{kt} \stackrel{\text{iid}}{\sim} Z(1/2, 1/2, 0, 1)$$

## ■ Add Gibbs update steps for

- ▶ past disturbances:  $\mathbf{z}_t = (\varepsilon_{t-1}, \varepsilon_{t-s}, \varepsilon_{t-(1+s)})^\top$
- ▶ pre-sample lags:  $y_0, y_{-1}, y_{-2}, \dots$

# Sampling from conditional posterior for $\theta_{0:T}$

- Full conditional posterior for  $\theta_{0:T}$ :
  - ▶ state-space model with state vector  $\theta_t$ .
  - ▶ **nonlinear Gaussian observation model.**
  - ▶ Linear (heteroscedastic) Gaussian state transition.
- **Particle Gibbs with Ancestor Sampling (PGAS)** [11]
  - ▶ simulation consistent
  - ▶ slow
  - ▶ **particle degeneracy when model is near-degenerate**  
(parameters standing still for extended periods).
- **FFBSx - FFBS with extended Kalman filter**
  - ▶ fast
  - ▶ robust to near-degeneracy
  - ▶ approximate, but shown to be very accurate for TVSAR
  - ▶ **automatic differentiation**  $\tilde{\phi}_t = \tilde{\mathbf{g}}(\theta_t)$  makes it all beautiful.

# Extended Kalman filter

## ■ State-space model

$$\theta_t = A\theta_{t-1} + Bu_t + \eta_t, \quad \eta_t \sim N(\mathbf{0}, \Sigma_\eta)$$
$$y_t = C(\theta_t) + \varepsilon_t, \quad \varepsilon_t \sim N(0, \Sigma_\varepsilon)$$

## Standard Kalman filter

```
# Prior propagation step
```

$$\bar{\mu} = A*\mu .+ B*u$$

$$\bar{\Omega} = A*\Omega*A' + \Sigma_n$$

```
# Measurement update
```

$$K = \bar{\Omega}*C' / (C*\bar{\Omega}*C' .+ \Sigma_e)$$

$$\mu = \bar{\mu} + K*(y .- C*\bar{\mu})$$

$$\Omega = (I - K*C)*\bar{\Omega}$$

## Extended Kalman filter

```
# Prior propagation step
```

$$\bar{\mu} = A*\mu + B*u$$

$$\bar{\Omega} = A*\Omega*A' + \Sigma_n$$

$$\bar{C} = \partial C(\bar{\mu}, \text{Cargs})$$

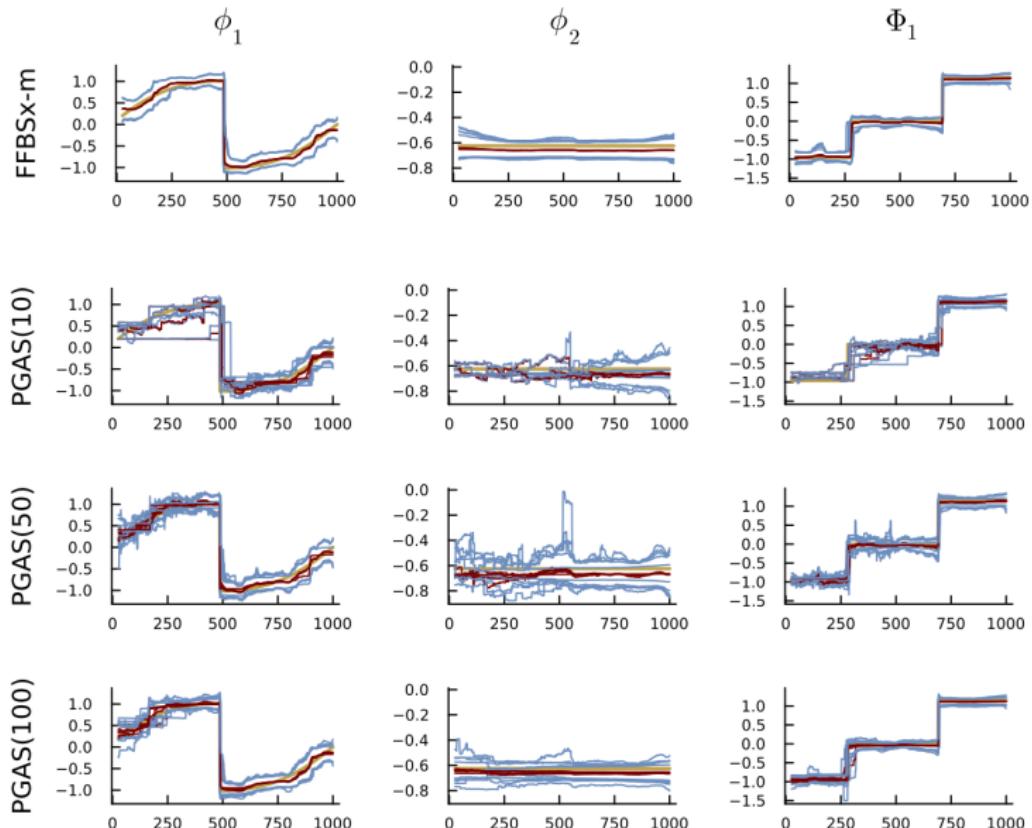
```
# Measurement update
```

$$K = \bar{\Omega}*\bar{C}' / (\bar{C}*\bar{\Omega}*\bar{C}' .+ \Sigma_e)$$

$$\mu = \bar{\mu} + K*(y .- C(\bar{\mu}, \text{Cargs}))$$

$$\Omega = (I - K*\bar{C})*\bar{\Omega}$$

# Simulation TVSAR with $p = (2, 2)$ and $s = (1, 12)$



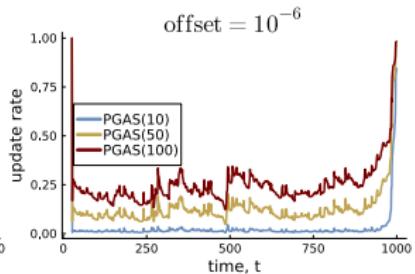
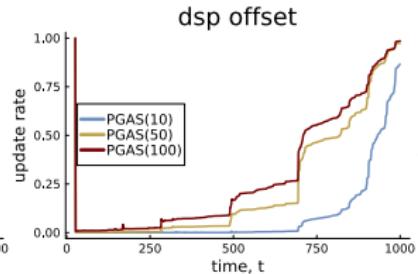
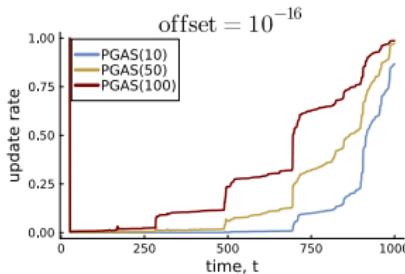
# PGAS - larger offset improves convergence

## Volatility models

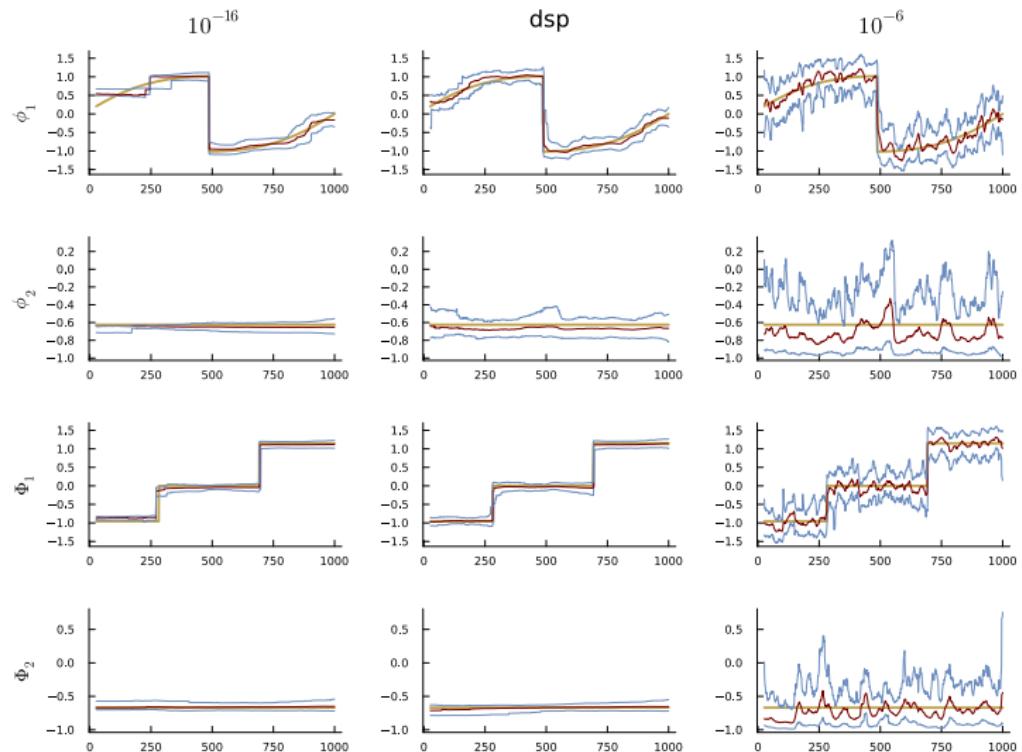
$$\nu_t = \exp(h_t/2)\epsilon_t, \quad \epsilon_t \sim N(0, 1)$$

## The usual trick

$$\log(\nu_t^2 + \text{offset}) = h_t + \log \epsilon_t^2$$



# PGAS - but makes parameters more wiggly



# Spectral density ARMA

## Spectral density ARMA(p,q)×(P,Q)

seasonal period

no. of AR

no. of MA

no. of seasonal AR

no. of seasonal MA

### Parameter settings

$\phi_1$

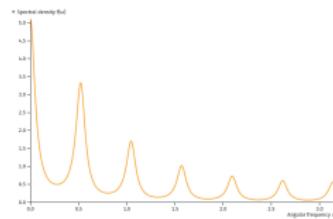
No regular MA

$\Phi_1$

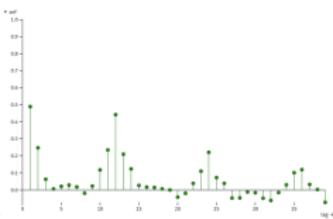
No seasonal MA

plot log f(u)

Spectral density of an ARMA(1,0)×(1,0) process with season s=12

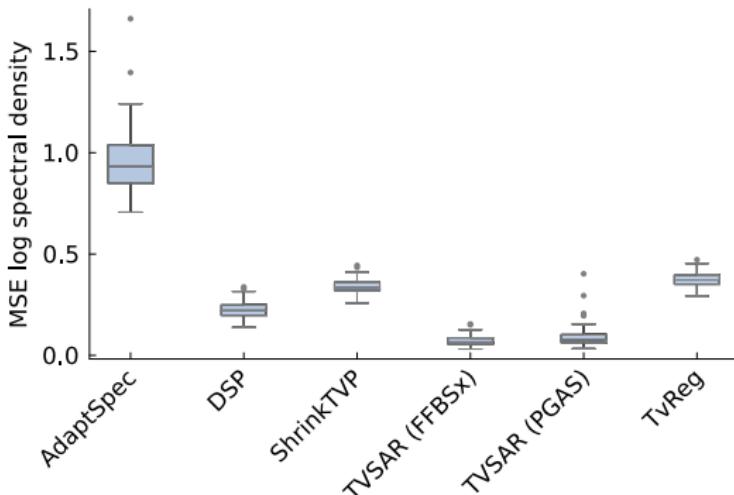


Estimated ACF of ARMA(1,0)×(1,0) process with season s=12

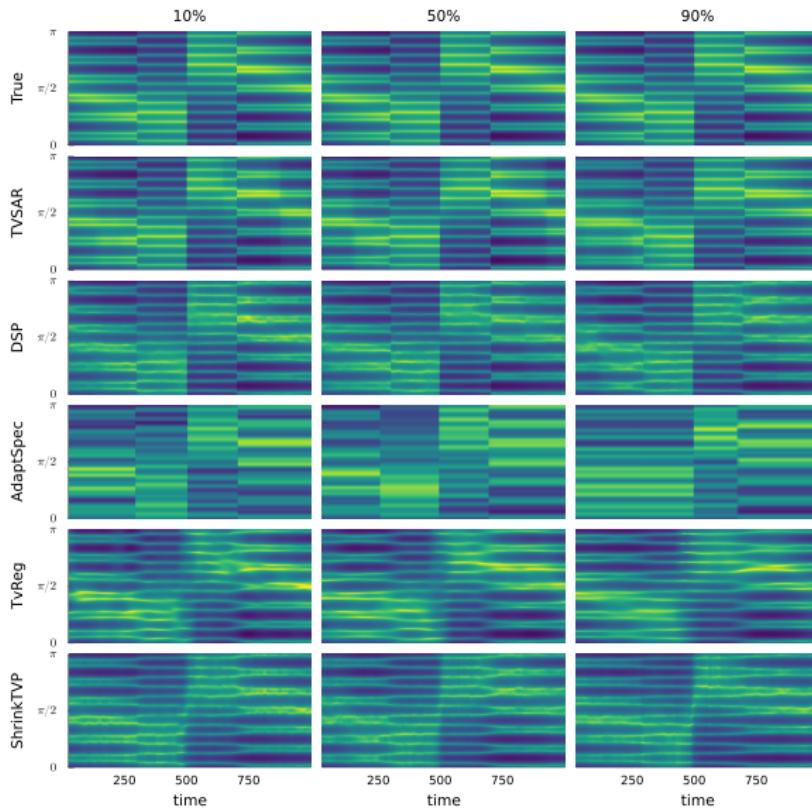


# Fitting the time-varying spectral density

- Benchmark models are linear:
  - ▶ Fit  $y_t = \mathbf{z}_t^\top \tilde{\phi}_t + \varepsilon_t$  with zero restrictions
  - ▶ Coeff on lag  $x_{t-(1+s)}$  fitted without multiplicative restriction.
- Time-varying spectral density  $f(\omega, t)$ . Spectrogram.
- MSE of log spectrogram estimate:  $\widehat{\log f(\omega, t)}$  integrated over frequency  $\omega \in (0, \pi]$  and time  $t = 1, \dots, 1000$ .

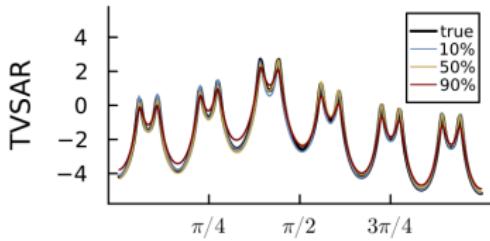


# Fitted log spectrogram

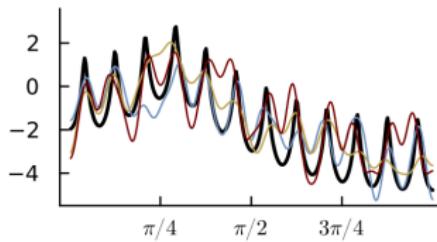
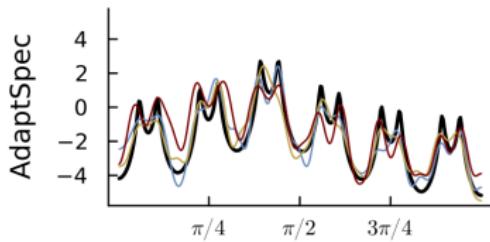
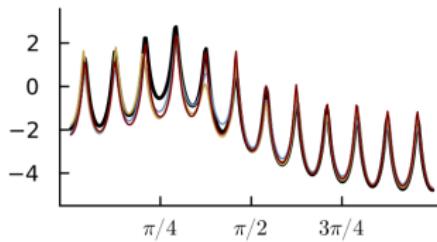
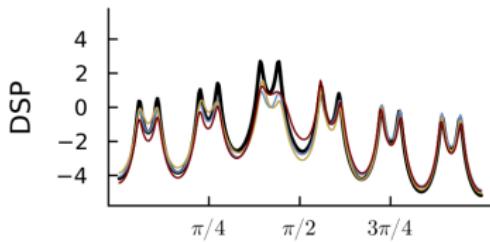
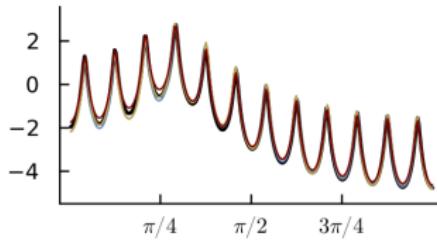


# Fitted log spectrogram snapshots

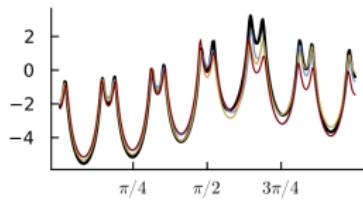
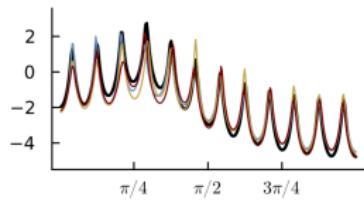
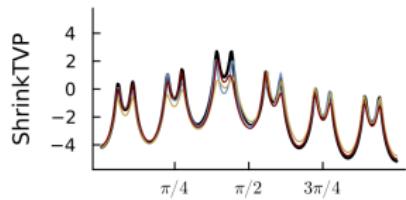
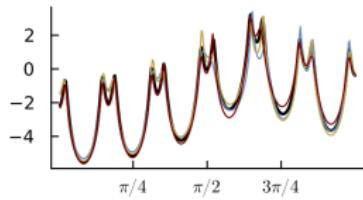
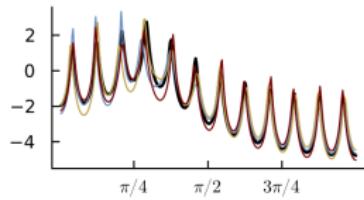
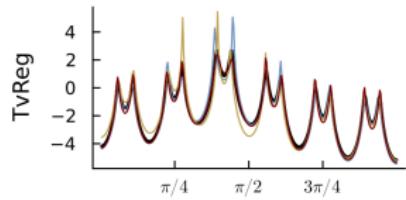
$t = 100$



$t = 400$

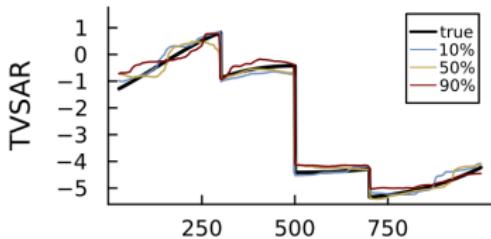


# Fitted log spectrogram snapshots

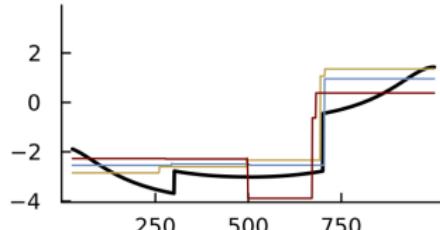
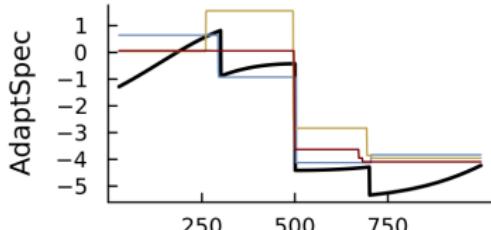
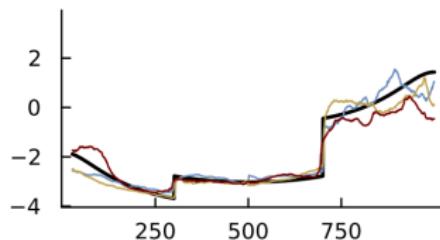
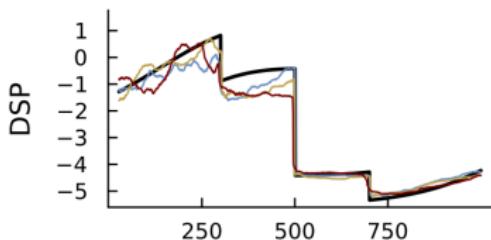
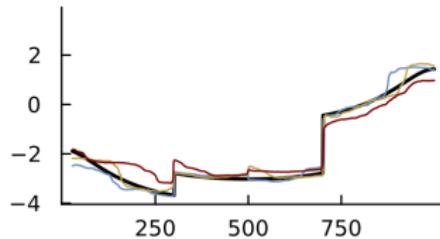


# Fitted log spectral density over time

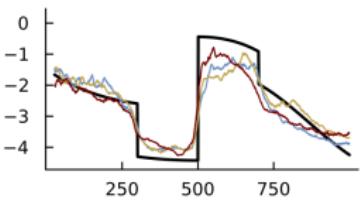
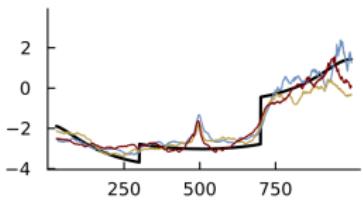
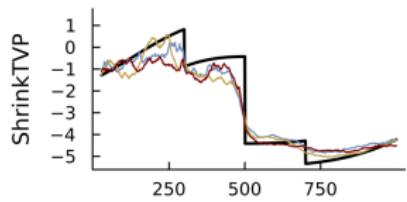
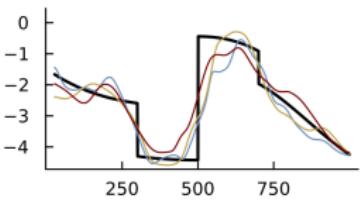
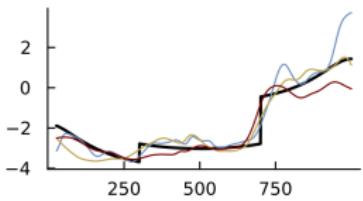
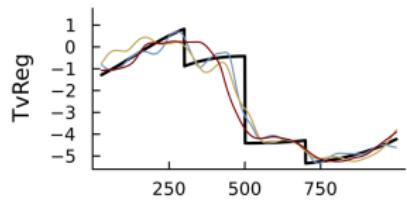
$$\omega = \pi/4$$



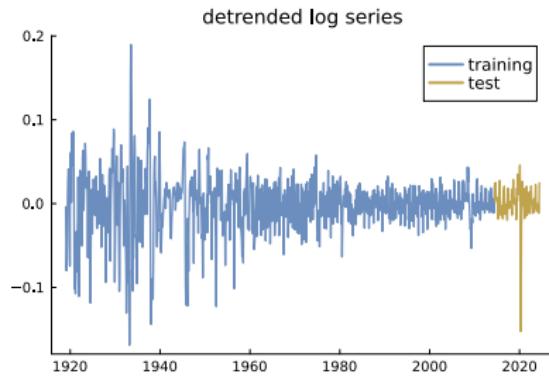
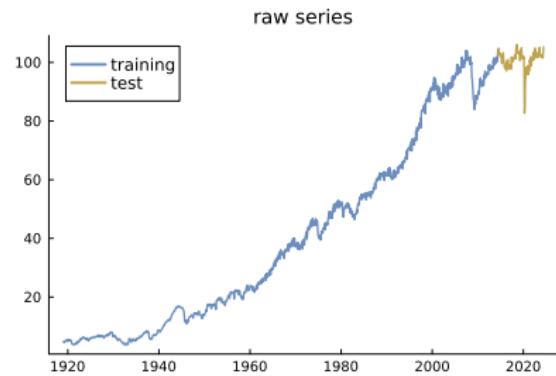
$$\omega = \pi/2$$



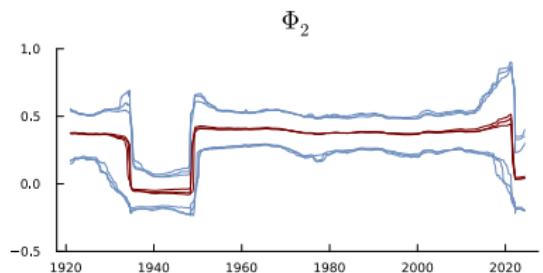
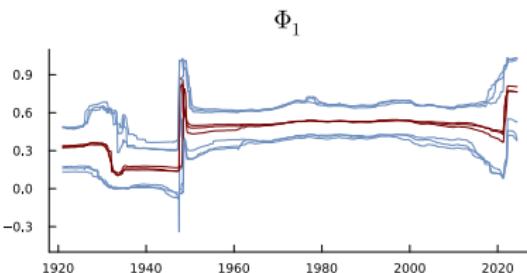
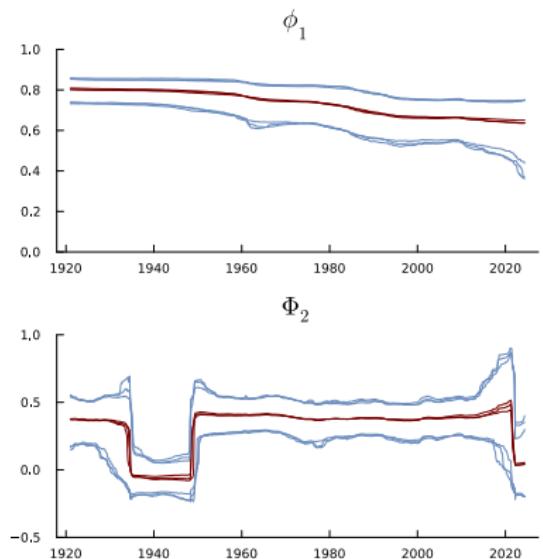
# Fitted log spectral density over time



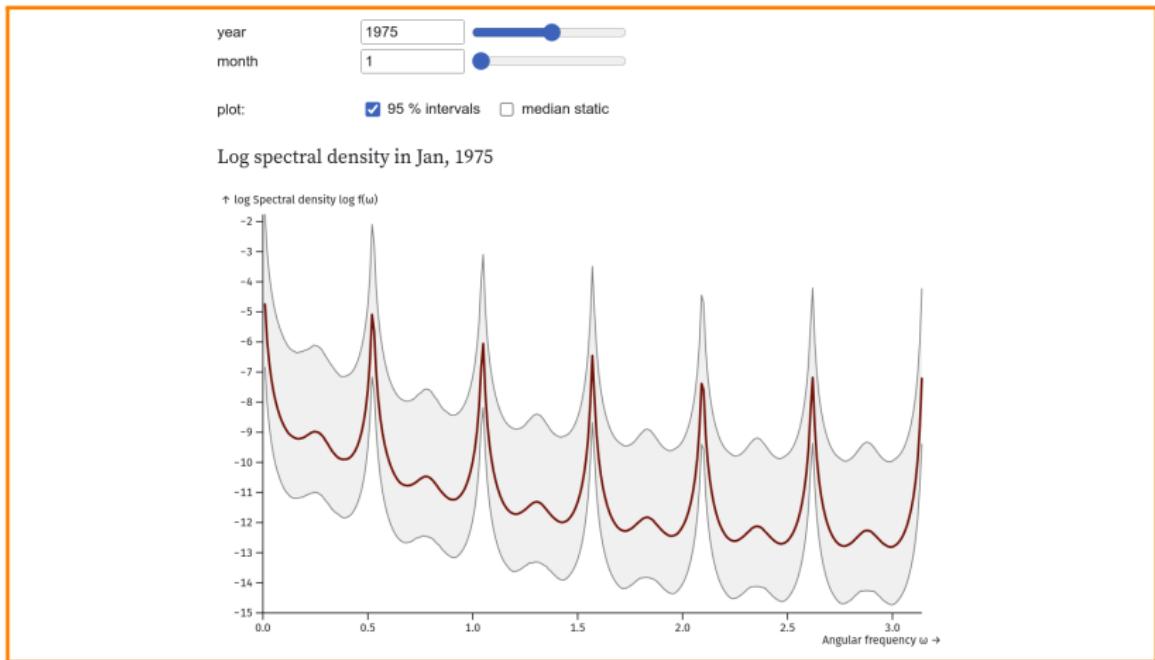
# US industrial production 1919-2024



# SAR(1,2) - FFBSx three random initial values



# SAR(1,2) spectral density snapshots



# Conclusions

- Multi-seasonal AR model with **time-varying** parameters.
- Stability restrictions at every time point. **Locally stable**.
- Parameter evolution by **dynamic shrinkage processes**.
- Challenging inference problem:
  - ▶ **non-linear** - stability restrictions and multiplicative seasonality
  - ▶ dynamic shrinkage prior tends to give **near-degeneracy**.
- Gibbs sampler with **particle/extended Kalman update** step.
- In progress:
  - ▶ Beyond Extended Kalman.
  - ▶ Better prior elicitation and posterior sampling for DSP prior.
  - ▶ **julia** package for arbitrary number of polynomials.

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