

Multivar_HW4

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Question 1

With the information provided,

$$\begin{aligned}\text{Cov}(X) &= \text{Cov}(\lambda F + e) \\ &= \lambda \lambda^\top \text{Var}(F) + \Psi \\ &= \lambda \lambda^\top + \Psi.\end{aligned}$$

For $i \neq j$

$$= (\lambda \lambda^\top)_{ij} = \lambda_i \lambda_j.$$

Hence the i -th row ignoring its diagonal element is

$$(\lambda_i \lambda_1, \lambda_i \lambda_2, \dots, \lambda_i \lambda_p) = \lambda_i \lambda^\top,$$

which is the vector λ^\top scaled by the constant λ_i .

Therefore every off-diagonal row of Σ_X is proportional to every other, as required.

Question 2

a

```
library(psych)
library(stats)
library(tidyverse)
```

```
## -- Attaching core tidyverse packages ----- tidyverse 2.0.0 --
## v dplyr      1.1.4      v readr      2.1.5
## v forcats    1.0.0      v stringr   1.5.1
## v ggplot2    3.5.1      v tibble    3.2.1
## v lubridate  1.9.4      v tidyr     1.3.1
## v purrr      1.0.4
## -- Conflicts ----- tidyverse_conflicts() --
## x ggplot2::%>%() masks psych::%>%()
## x ggplot2::alpha() masks psych::alpha()
## x dplyr::filter() masks stats::filter()
## x dplyr::lag() masks stats::lag()
## i Use the conflicted package (<http://conflicted.r-lib.org/>) to force all conflicts to become errors
```

```

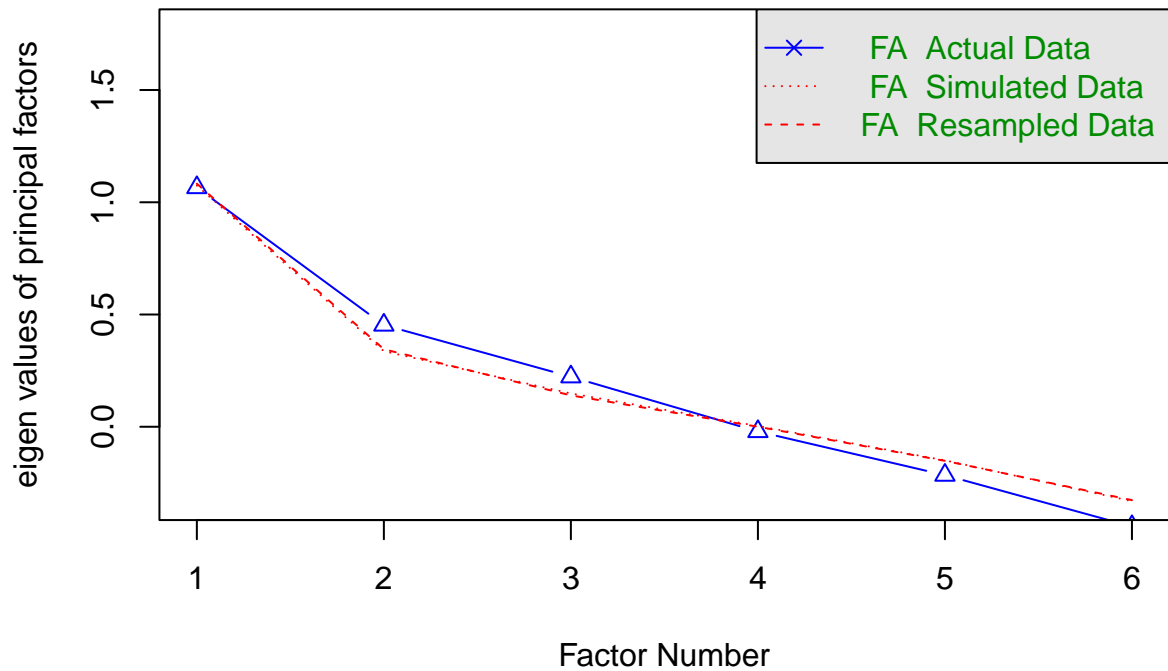
dat_raw <- read.table("Data/testData.txt", header = TRUE,
                      stringsAsFactors = FALSE)

tests <- dat_raw[, 1:6]
colnames(tests) <- paste0("T", 1:6)

fa.parallel(tests, fa = "fa", fm = "ml", n.iter = 100)

```

Parallel Analysis Scree Plots



Parallel analysis suggests that the number of factors = 0 and the number of components = NA

```

m <- 3

fa_raw <- fa(tests,
              nfactors = m,
              rotate = "none",
              scores = TRUE)

print(fa_raw$loadings, cutoff = .30, digits = 3)

```

```

##
## Loadings:
##   MR1   MR2   MR3
## T1  0.717  0.595 -0.363

```

```
## T2                -0.530
## T3
## T4  0.520         0.612
## T5 -0.600  0.765
## T6  0.468
##
##                MR1   MR2   MR3
## SS loadings    1.437 0.976 0.906
## Proportion Var 0.240 0.163 0.151
## Cumulative Var 0.240 0.402 0.553
```

I selected 3 but unfortunately this only explains 55% of the variance. I chose this because two factors just isn't enough and more than 3 makes interpretation hard.

Factor 1 covers 24% of variance and seems to contrast t5 against t 1,4,6 as it has positive weights for those and a negative weight for t5.. Factor 2 accounts for 16% of variance is largely accounting for T5 and T1. Factor 3 accounts for another 15% of variance, and contrasts t4 against t1 and 2. Altogether these are hard to interpret.

```
communalities <- fa_raw$communality
specificities <- 1 - communalities

cat("\nCommunalities:\n"); print(round(communalities, 3))
```

```
##
## Communalities:

##      T1      T2      T3      T4      T5      T6
## 0.999 0.297 0.123 0.647 0.997 0.256
```

```
cat("\nSpecificities:\n"); print(round(specificities, 3))
```

```
##
## Specificities:

##      T1      T2      T3      T4      T5      T6
## 0.001 0.703 0.877 0.353 0.003 0.744
```

Tests 1 and 5 are almost redundant with the factors. Tests 2,3, and 6 carry substantial test specific variance as shown with high specificity. This is especially true for T3 as it has the highest specificity.

b

```
fa_rot <- factanal(tests, factors = m,
                   rotation = "varimax", scores = "regression")
print(fa_rot, digits = 3, cutoff = 0.30)
```

```
##
## Call:
## factanal(x = tests, factors = m, scores = "regression", rotation = "varimax")
```

```
##
## Uniquenesses:
##   T1   T2   T3   T4   T5   T6
## 0.005 0.703 0.866 0.378 0.005 0.726
##
## Loadings:
##   Factor1 Factor2 Factor3
## T1  0.996
## T2      -0.328 -0.387
## T3      0.345
## T4      0.784
## T5      0.977
## T6  0.367      0.366
##
##               Factor1 Factor2 Factor3
## SS loadings      1.178   1.078   1.061
## Proportion Var    0.196   0.180   0.177
## Cumulative Var    0.196   0.376   0.553
##
## The degrees of freedom for the model is 0 and the fit was 0.0165
```

```
obs26_scores <- fa_rot$scores[26, ]
obs26_scores
```

```
##   Factor1   Factor2   Factor3
## -3.5288297 -0.4527419  0.6954494
```

This is now easier to interpret because there are clear factors for certain items.

Factor 1 is clearly accountign fo T1 variance wiht some T6, Factor 2, is focused on T5 with some negative T2 and factor 3 is spread across several. Now, T1, T5 and T4 are only tied to one matrix which makes things simpler.

c

```
scores_all <- fa_rot$scores      # n × 3 matrix (n = number of students)

obs26 <- scores_all[26, ]        # a named vector
print(obs26, digits = 3)
```

```
## Factor1 Factor2 Factor3
## -3.529  -0.453   0.695
```

Student 26 is very weak for factor 1 focused on T1 and T6. They are slightly below average for T5 (with some signal from T2), and moderately above average for T4 with some loading from T3 and T6. Altogether it is hard to give a text interpretation here without knowing whwat each test asks.

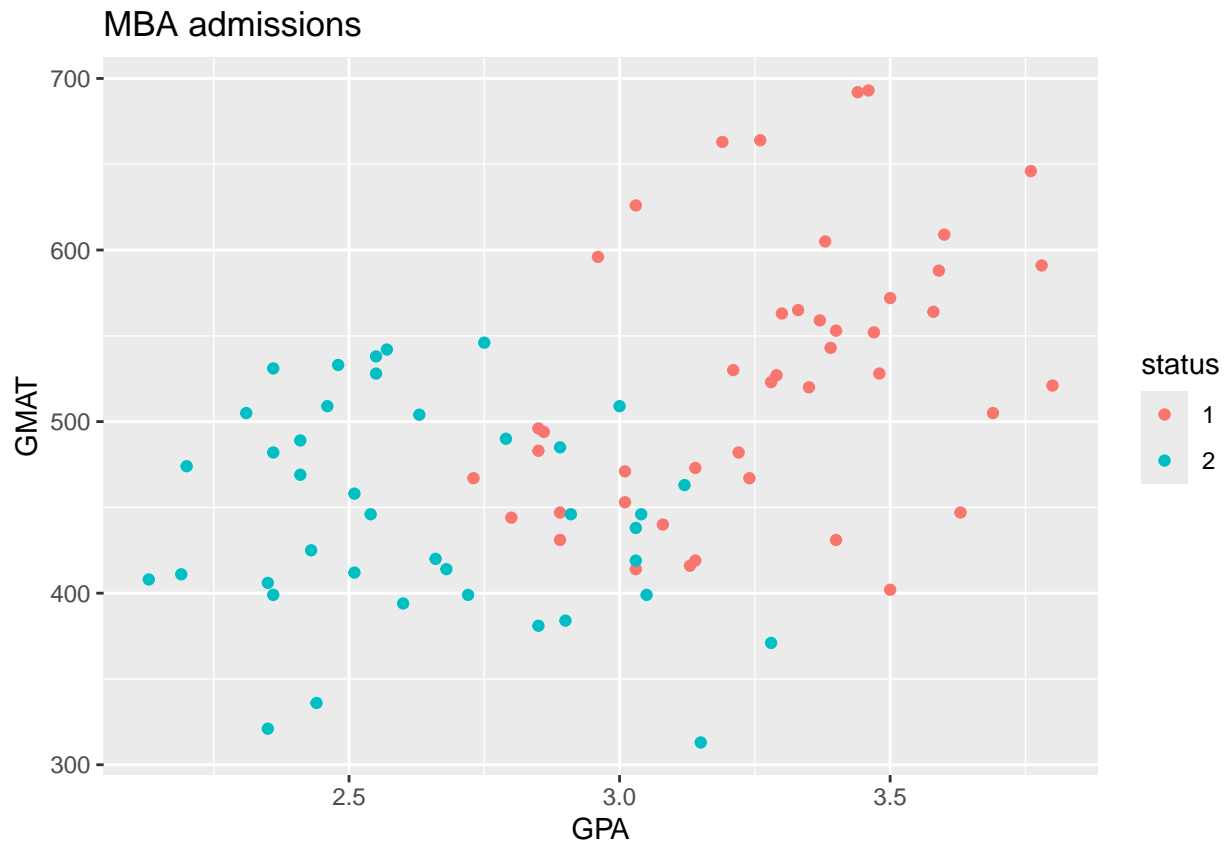
Question 2

##a

```
load("Data/admission.RData")
colnames(admission) <- c("GPA", "GMAT", "status")
admission <- as_tibble(admission) %>% mutate(status = factor(status))

g1 <- filter(admission, status == 1)
g2 <- filter(admission, status == 2)

ggplot(admission, aes(GPA, GMAT, colour = status)) + geom_point() +
  labs(title = "MBA admissions")
```



```
xbar1 <- colMeans(dplyr::select(g1, GPA, GMAT))
xbar2 <- colMeans(dplyr::select(g2, GPA, GMAT))
S1 <- cov(dplyr::select(g1, GPA, GMAT))
S2 <- cov(dplyr::select(g2, GPA, GMAT))
Spooled <- ((nrow(g1)-1)*S1 + (nrow(g2)-1)*S2)/(nrow(g1)+nrow(g2)-2)

cat("\nGroup means:\n")
```

```
##
## Group means:
```

```
print(round(rbind(Admit = xbar1, Deny = xbar2), 2))
```

```
##      GPA  GMAT
```

```
## Admit 3.27 526.11
## Deny 2.64 446.08
```

```
cat("\nGroup covariance matrices:\n\nS1 (Admit):\n")
```

```
##
## Group covariance matrices:
##
## S1 (Admit):
```

```
print(round(S1, 3))
```

```
##          GPA      GMAT
## GPA  0.080      9.216
## GMAT 9.216 6215.374
```

```
cat("\nS2 (Deny):\n")
```

```
##
## S2 (Deny):
```

```
print(round(S2, 3))
```

```
##          GPA      GMAT
## GPA  0.088     -3.268
## GMAT -3.268 3823.763
```

```
cat("\nPooled covariance (Spooled):\n")
```

```
##
## Pooled covariance (Spooled):
```

```
print(round(Spooled, 3))
```

```
##          GPA      GMAT
## GPA  0.084      3.350
## GMAT 3.350 5091.605
```

b

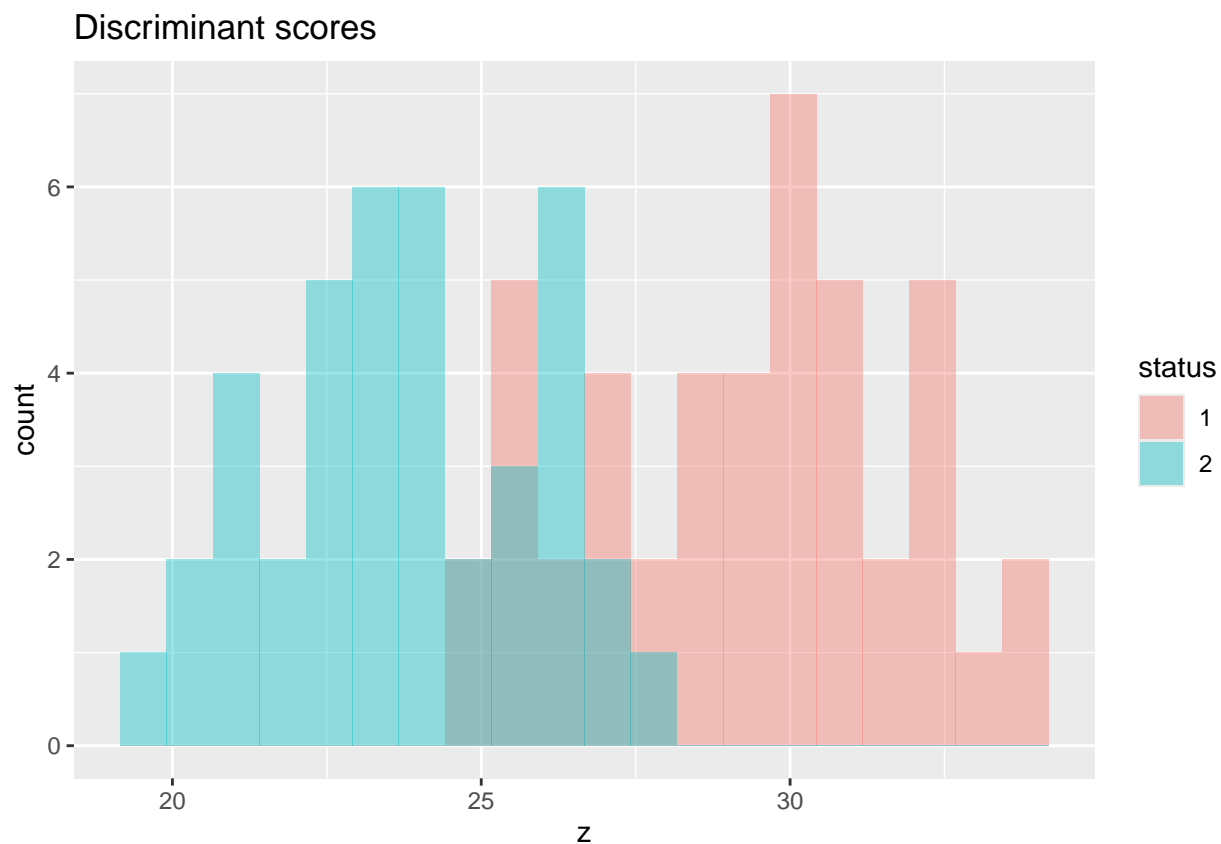
```
a <- solve(Spooled, xbar1 - xbar2)
a
```

```
##          GPA      GMAT
## 7.14806424 0.01101639
```

c

```
Z <- as.matrix(dplyr::select(admission, GPA, GMAT)) %*% a # n×1
admission$z <- drop(Z)

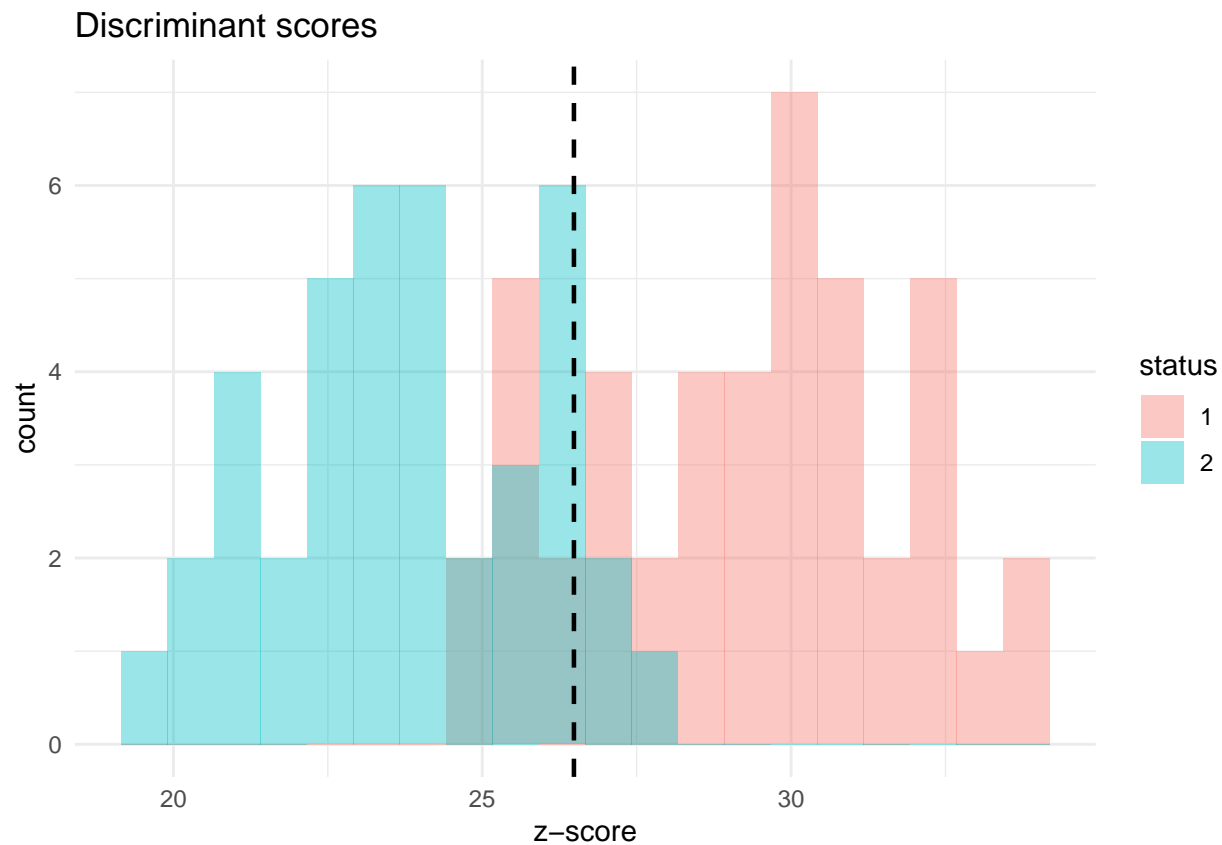
ggplot(admission, aes(x = z, fill = status)) +
  geom_histogram(alpha = 0.4, bins = 20, position = "identity") +
  labs(title = "Discriminant scores")
```



d

```
z_cut <- as.numeric(crossprod(a, (xbar1 + xbar2)/2)) # midpoint

ggplot(admission, aes(x = z, fill = status)) +
  geom_histogram(alpha = 0.4, bins = 20, position = "identity") +
  geom_vline(xintercept = z_cut, linetype = "dashed", linewidth = 0.8) +
  labs(title = "Discriminant scores", x = "z-score") +
  theme_minimal()
```



e

```
## ----- (e) classify (3, 500) -----
new_pt <- c(3, 500)
z_new <- crossprod(a, new_pt)
group <- ifelse(z_new > z_cut, "Admit", "Reject") # 1 = admit
print(group)
```

```
##      [,1]
## [1,] "Admit"
```

This person should be Admitted.

f

```
# direction vector
slope <- -a[1]/a[2]
intercept <- z_cut/a[2]

ggplot(admission, aes(GPA, GMAT, colour = status)) +
  geom_point() +
  geom_abline(slope = slope, intercept = intercept, lty = 2) +
  geom_point(aes(x = new_pt[1], y = new_pt[2]), colour = "black",
```



```

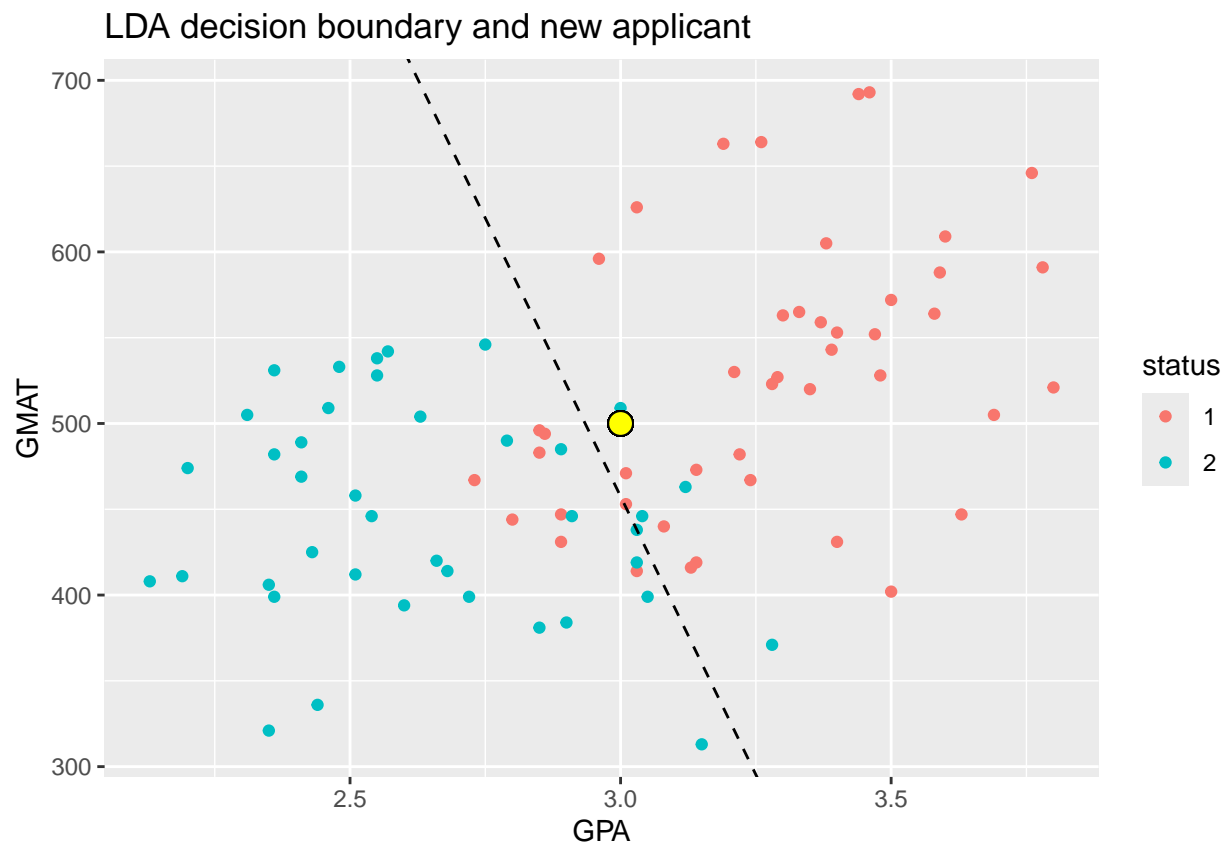
    shape = 21, size = 4, fill = "yellow") +
  labs(title = "LDA decision boundary and new applicant")

```

```

## Warning in geom_point(aes(x = new_pt[1], y = new_pt[2]), colour = "black", : All aesthetics have length 1
## i Please consider using 'annotate()' or provide this layer with data containing
## a single row.

```



g

```

library(MASS)

```

```

##
## Attaching package: 'MASS'

## The following object is masked from 'package:dplyr':
##
## select

```

```

lda_fit <- lda(status ~ GPA + GMAT, data = admission)
lda_fit$scaling

```

```
##          LD1
## GPA  -3.071443853
## GMAT -0.004733621
```

```
cat("scaling", a/lda_fit$scaling)
```

```
## scaling -2.327265 -2.327265
```

This appears to be a scaled and flipped version of my coefficients. specifically, my coefficients are equal to -2.327 times the LDA library coefficients. This is because of the way variance scaling is managed in LDA.