Home Assignment 1

Matthew Zhao

1/14/2023

Lecture 1 HW

Problem 1

$$\hat{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

$$= \frac{1}{n} \sum_{i=1}^n X_i^2 - 2X_i \bar{X}_n + \bar{X}_n^2$$

$$= \frac{1}{n} \sum_{i=1}^n (X_i^2) - 2(\frac{1}{n} \sum_{i=1}^n X_i \bar{X}_n) + \frac{1}{n} (\sum_{i=1}^n \bar{X}_n)^2$$

$$= \frac{1}{n} \sum_{i=1}^n (X_i^2) - 2(\frac{1}{n} \sum_{i=1}^n X_i)^2 + \frac{1}{n} (\sum_{i=1}^n X_i)^2$$

$$= \frac{1}{n} \sum_{i=1}^n (X_i^2) - (\frac{1}{n} \sum_{i=1}^n X_i)^2$$

By LLN,
$$\frac{1}{n} \sum_{i=1}^{n} (X_i^2) \stackrel{p}{\to} E_P(X^2)$$
 and $\frac{1}{n} \sum_{i=1}^{n} X_i \stackrel{p}{\to} E_P(X)$

By CMT,
$$\hat{\sigma}_n^2 \xrightarrow{p} E_P(X^2) - E_P(X)^2 = Var_P(X)$$

Problem 2

Intuitively, as $n \to \infty$, the probability that $X_1, ..., X_n$ does not contain θ approaches 0 by definition of uniform distribution. Hence, the probability that our estimator for θ (max $\{X_1, ..., X_n\}$) does not equal θ also goes to zero, thus $P(|\hat{\theta} - \theta| > \epsilon) \to 0$ as $n \to \infty$.

Formally, we can derive the CDF F_{θ} of our estimator and use it to show that as $n \to \infty$, our estimator converges in probability to θ . Specifically, we can use order statistics to show that $F_{\theta}(x) = F(x)^n$ where F(x) is the CDF for each individual X_i drawn from $U[0, \theta]$:

$$F_{\theta}(x) = \mathbb{P}(X_1, ..., X_n \le x)$$
$$= \mathbb{P}(X_1 \le x) \mathbb{P}(X_2 \le x) ... \mathbb{P}(X_n \le x)$$
$$= F(x)^n$$

Plugging this result in we show:

$$\mathbb{P}(|\hat{\theta}_n - \theta| > \epsilon) = \mathbb{P}(\theta - \epsilon < \hat{\theta}_n < \theta + \epsilon)$$

$$= F_{\theta}(\theta + \epsilon) - F_{\theta}(\theta - \epsilon)$$
$$= F(\theta + \epsilon)^n - F(\theta - \epsilon)^n$$
$$= \frac{(\theta + \epsilon)^n}{\theta} - \frac{(\theta - \epsilon)^n}{\theta} \to 0$$

as $\epsilon \to 0$, thus $\hat{\theta}_n \stackrel{p}{\to} \theta$

Problem 3

```
library(tidyverse)
library(haven)
library(stargazer)
library(xtable)
options(scipen=999)
make_ci <- function(xbar,var){</pre>
  c(xbar - 1.96*var, xbar + 1.96*var)
}
ci coverage <- function(n){</pre>
  X <- rnorm(n)
  xbar <- mean(X)</pre>
  sd \leftarrow sd(X)
  known <- make_ci(xbar,1/sqrt(n))</pre>
  est <- make_ci(xbar,sd/sqrt(n))</pre>
  c(0 \ge known[1] & 0 \le known[2],
    0 >= est[1] & 0 <= est[2])
known_results <- c()</pre>
est_results <- c()
N \leftarrow c(30,100,500)
for (n in N){
  known_temp <- c()</pre>
  est temp <- c()
  for (x in 1:1000){
    known_temp <- c(known_temp,ci_coverage(n)[1])</pre>
    est_temp <- c(est_temp,ci_coverage(n)[2])</pre>
  known_results <- c(known_results,mean(known_temp))</pre>
  est_results <- c(est_results,mean(est_temp))</pre>
data.frame(N,est_results,known_results)
```

```
## N est_results known_results
## 1 30 0.951 0.962
## 2 100 0.954 0.949
## 3 500 0.946 0.950
```

Lecture 2 HW

Problem 1

- (1)
- (2)
- (3)
- **(4)**
- **(5)**
- **(6)**

Problem 2

- (0)
- (1)
- **(2)**
- (3)