

Computational Statistics Project 2



The code for this project is available under
<https://github.com/max607/computational-statistics-em>.

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1 Maximum Likelihood estimation of θ

1.1 Notation

In the following, random variables (RV) are denoted with capital letters, e.g. Y , their realizations with lowercase letters, e.g., y_i , where always $i = 1, \dots, n$. Log-likelihoods are written, e.g., as $\ell(\theta)$. For parameters, Greek letters are used and their estimators are distinguished with a hat, e.g., $\hat{\theta}$. In the context of this project, derivatives are only taken when functions are viewed as functions of one variable, denoted, e.g., as $\ell'(\theta)$.

1.2 Likelihood

150 one dimensional data are given. Figure 1.1 shows a histogram of them. They exhibit a positive skew.

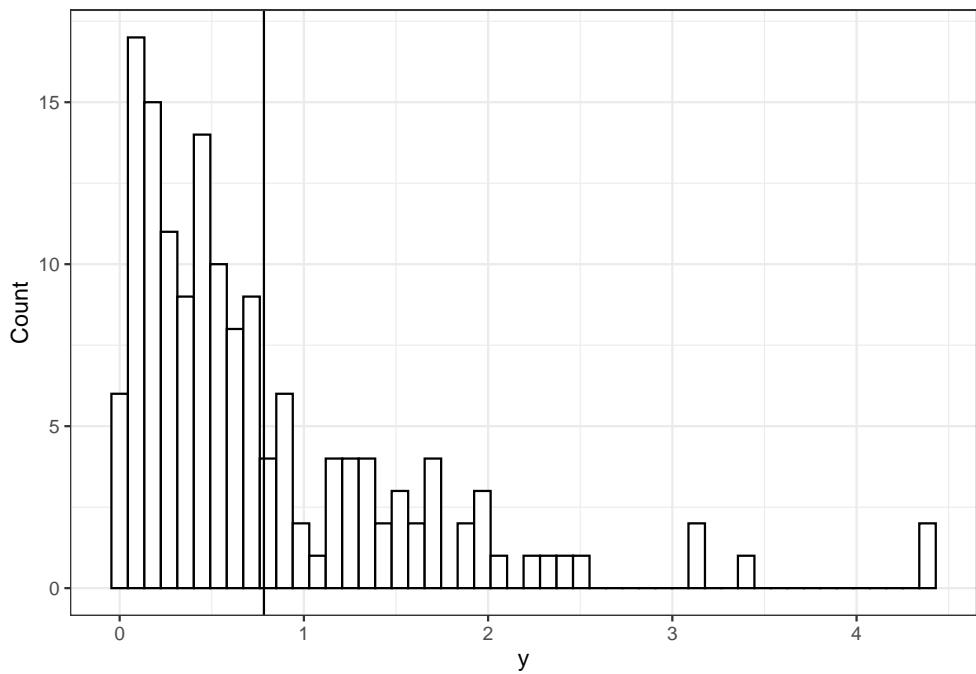


Figure 1.1: Histogram of provided data. The black line indicates the sample mean.

It is assumed they are independently identically distributed (i.i.d.) realizations y_i of a RV Y with probability density function (PDF)

$$f_\theta(y_i) = \frac{\theta^2}{\theta + 1} (1 + y_i) \exp(-\theta y_i), \quad y_i, \theta > 0. \quad (1.1)$$

The goal is to estimate θ via maximum likelihood (ML). Its log-likelihood

$$\begin{aligned}
\ell(\theta) &= \log \left(\prod_{i=1}^n f_\theta(y_i) \right) \\
&= 2n \log(\theta) - n \log(\theta + 1) - \theta \sum_{i=1}^n y_i + c \\
&\propto 2 \log(\theta) - \log(\theta + 1) - \theta \bar{y} + c,
\end{aligned} \tag{1.2}$$

where c is a constant, which does not depend on θ , and thus is irrelevant for subsequent calculations. The last term (1.2) results by dividing the log-likelihood by the number of observations n , so \bar{y} denotes the sample mean.

The first derivative of (1.2)

$$\ell'(\theta) = \frac{2}{\theta} - \frac{1}{\theta + 1} - \bar{y}. \tag{1.3}$$

There is no analytical solution available for equating (1.3) to zero and solving for θ .

1.3 Newton-Raphson

The maximum of $\ell(\theta)$ has to be found numerically, here via Newton-Raphson, i.e., iteratively applying

$$\theta^* = \theta - \frac{\ell'(\theta)}{\ell''(\theta)}, \tag{1.4}$$

where $\ell''(\theta) = -\frac{2}{\theta^2} + \frac{1}{(\theta+1)^2}$, θ is the value at iteration t and θ^* is the updated value at iteration $t + 1$, until the update gets very close to zero. The final value is taken as the ML estimate.

For the given data $\hat{\theta} = 1.7424899$.

2 Bootstrapping for standard error of $\hat{\theta}$

2.1 Sampling from $f_{\hat{\theta}}(y_i)$

The goal is to quantify the uncertainty of $\hat{\theta}$. For this, parametric bootstrap is applied, for which sampling from $f_{\hat{\theta}}(y_i)$ is necessary. The starting point is restating (1.1) as

$$f_{\theta}(y_i) = \frac{\theta}{\theta+1} \theta \exp(-\theta y_i) + \frac{1}{\theta+1} \theta^2 y_i \exp(-\theta y_i) \quad (2.1)$$

and recognizing this as a mixture of two Gamma distributions in shape and rate parameterization, where θ is the rate and the shapes are equal to one and two, respectively.

Starting from first principals, it is assumed only RVs $U \stackrel{iid}{\sim} U(0, 1)$ are available. This is not too much of a hassle, as Gamma RVs with shape j are the sum of j Exponential RVs, which in turn can be easily obtained via inversion

$$f^{-1}(u; \theta) = -\frac{\log(u)}{\theta}, \quad (2.2)$$

where θ already is the desired rate.

Sampling from $f_{\hat{\theta}}(y_i)$ is implemented in the following steps:

- 1) Draw u_i n -times.
- 2) Calculate temporary $y_i = f^{-1}(u_i; \hat{\theta})$.
- 3) Draw the number of shape two Gamma RVs n_2 .
 - 1) Draw u_j n -times.
 - 2) $n_2 = \#\{u_j | u_j < \frac{1}{\hat{\theta}+1}\}$.
- 4) Draw u_k n_2 -times.
- 5) Calculate temporary $y_k = f^{-1}(u_k; \hat{\theta})$.
- 6) Add y_k s component-wise to the fist n_2 y_i s.
- 7) Return y_i s.

Returned is a sample of size n , which can be seen as realizations of Y . n_2 observations are realizations of a Gamma RV with shape two and $n - n_2$ observations are realizations of a Gamma RV with shape one.

For the purpose of estimating θ with the estimator of Section 1.3 it is of no importance that the sample is sorted by shape.

2.2 Bootstrap standard error

Given B samples of size n produced by the sampler of the previous section, and corresponding bootstrap estimates $\hat{\theta}_b^*$, $b = 1, \dots, B$, the bootstrap standard error

$$\hat{se}_B(\hat{\theta}) = \sqrt{\frac{1}{B} \sum_{b=1}^B (\hat{\theta}_b^* - \hat{\theta}^*)^2}, \quad (2.3)$$

according to the lecture slides, where $\hat{\theta}^*$ is the mean of all $\hat{\theta}_b^*$.

For this simulation $B = 10^4$ and the resulting $\hat{se}_B(\hat{\theta}) = 0.1142654$.

3 EM

3.1 Augmented data

Consider a more complex PDF for the given data

$$f_\pi(y_i) = \pi f_\theta(y_i) + (1 - \pi) f_\lambda(y_i), \quad y_i > 0, \pi \in [0, 1], \quad (3.1)$$

where $f_\theta(y_i)$ is the previous PDF and $f_\lambda(y_i) = \lambda \exp(-\lambda y_i)$, $\lambda > 0$, i.e., $f_\pi(y_i)$ is a mixture of two PDFs.

The goal is to estimate θ, λ and π using an EM algorithm.

Start by introducing augmented data x_i and assume they are i.i.d. realizations from $X \sim Ber(\pi)$. An equivalent formulation of (3.1) thus is

$$f(\theta, \lambda | x_i, y_i) = f_\theta(y_i)^{x_i} f_\lambda(y_i)^{1-x_i}, \quad (3.2)$$

where the likelihood of θ and λ is computed given x_i, y_i .

I.e., simulation from $f_\pi(y_i)$ given θ, λ and π is possible, by first drawing x_i and then drawing from $f_\theta(y_i)$ if $x_i = 1$ or from $f_\lambda(y_i)$ if $x_i = 0$.

(3.2) is called complete likelihood, while (3.1) is called observed likelihood.

3.2 Expectation

For the first part of the EM algorithm, $\mathbb{E}(X | \pi, \theta, \lambda, y_i)$ needs to be calculated. An analytical solution can be obtained by applying Bayes' theorem

$$\begin{aligned} f(x_i | \theta, \lambda, \pi, y_i) &= \frac{f(\theta, \lambda | x_i, y_i) f(x_i | \pi)}{f(\theta, \lambda | y_i)} \\ &= \frac{f_\theta(y_i)^{x_i} f_\lambda(y_i)^{1-x_i} \pi^{x_i} (1-\pi)^{1-x_i}}{(1-\pi) f_\lambda(y_i) + \pi f_\theta(y_i)}. \end{aligned} \quad (3.3)$$

The second line is implied by the distribution of X and the fact that $f(\theta, \lambda | y_i) = f(x_i = 0 | \pi) f(\theta, \lambda | x_i = 0, y_i) + f(x_i = 1 | \pi) f(\theta, \lambda | x_i = 1, y_i)$.

One can immediately see

$$\mathbb{E}(X | \pi, \theta, \lambda, y_i) = \frac{f_\theta(y_i) \pi}{(1 - \pi) f_\lambda(y_i) + \pi f_\theta(y_i)}. \quad (3.4)$$

3.3 Maximization

The second step is simple maximum likelihood estimation given the augmented data. The log-likelihoods of the respective parameters are

$$\ell(\pi) = \log(\pi) \sum_{i=1}^n x_i + \log(1 - \pi) \sum_{i=1}^n (1 - x_i), \quad (3.5)$$

$$\ell(\lambda) = n \log(\lambda) - \lambda \sum_{i=1}^n y_i - \log(\lambda) \sum_{i=1}^n x_i + \lambda \sum_{i=1}^n x_i y_i + c, \quad (3.6)$$

$$\begin{aligned} \ell(\theta) &= 2 \log(\theta) \sum_{i=1}^n x_i - \log(\theta + 1) \sum_{i=1}^n x_i - \theta \sum_{i=1}^n x_i y_i + c \\ &\propto 2 \log(\theta) - \log(\theta + 1) - \theta \tilde{y} + c, \end{aligned} \quad (3.7)$$

where $\tilde{y} = \sum_{i=1}^n x_i y_i / \sum_{i=1}^n x_i$.

The maximizer of (3.5) is $\hat{\pi} = \frac{1}{n} \sum_{i=1}^n x_i$.

It is also possible to maximize (3.6) analytically, via $\hat{\lambda} = (n - \sum_{i=1}^n x_i) / (\sum_{i=1}^n y_i - \sum_{i=1}^n x_i y_i)$, which can be interpreted as the inverse of the sample mean of the y_i , where $x_i = 0$. (3.7) has no closed form and is maximized using Newton-Raphson (see Section 1.3), substituting \bar{y} with \tilde{y} , which can be interpreted as the sample mean of the y_i , where $x_i = 1$.

3.4 Application

- Starting from initial guesses, $\pi_0 = 0.5, \theta_0 = 1, \lambda_0 = 1$, a first version of x_i s is generated using the result from Section 3.2.
- Next, the parameters are updated using the results from Section 3.3.
- Now iterate between the expectation and the maximization step while monitoring the observed log-likelihood given the current parameter estimates.
- Stop when there no longer is a relative increase in the observed log-likelihood and return $\hat{\pi}, \hat{\theta}, \hat{\lambda}$, the maximum likelihood estimates.

In this case $\hat{\pi} = 0.2251571$, $\hat{\theta} = 1.2431197$ and $\hat{\lambda} = 1.4864174$.

A augmented datum x_i can be seen as the probability of observation i pertaining to a cluster, where the PDF is $f_\theta(y_i)$. As can be seen in Figure 3.1, the probability of datum i to be from cluster c_θ increases linearly with y .

Figure 3.2 and 3.3 show the paths taken by the parameters and artificial data during optimization. All seem to have converged.

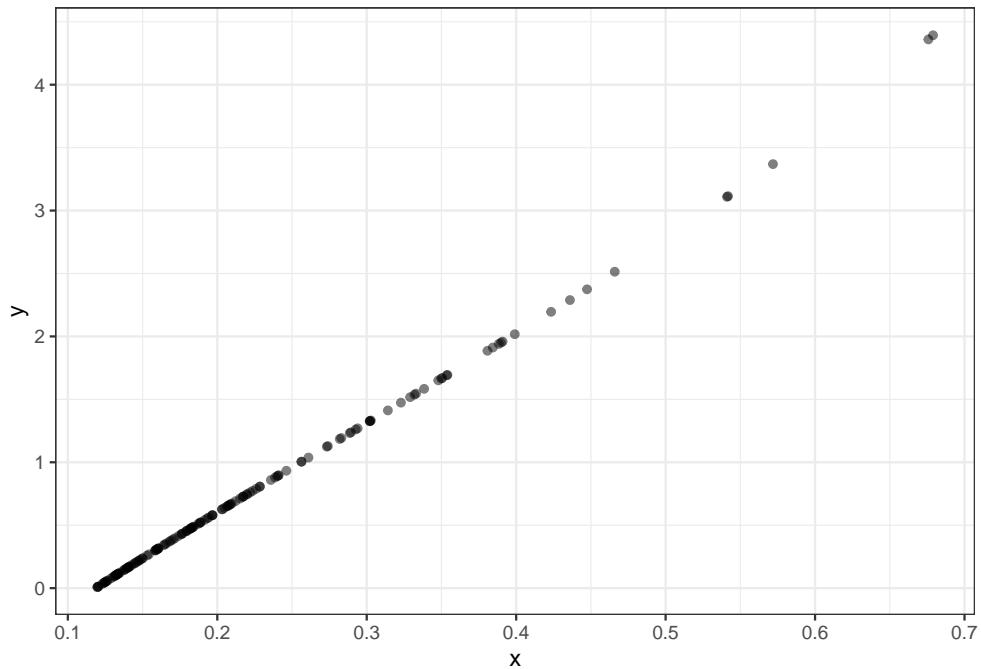


Figure 3.1: Scatterplot of x_i and y_i .

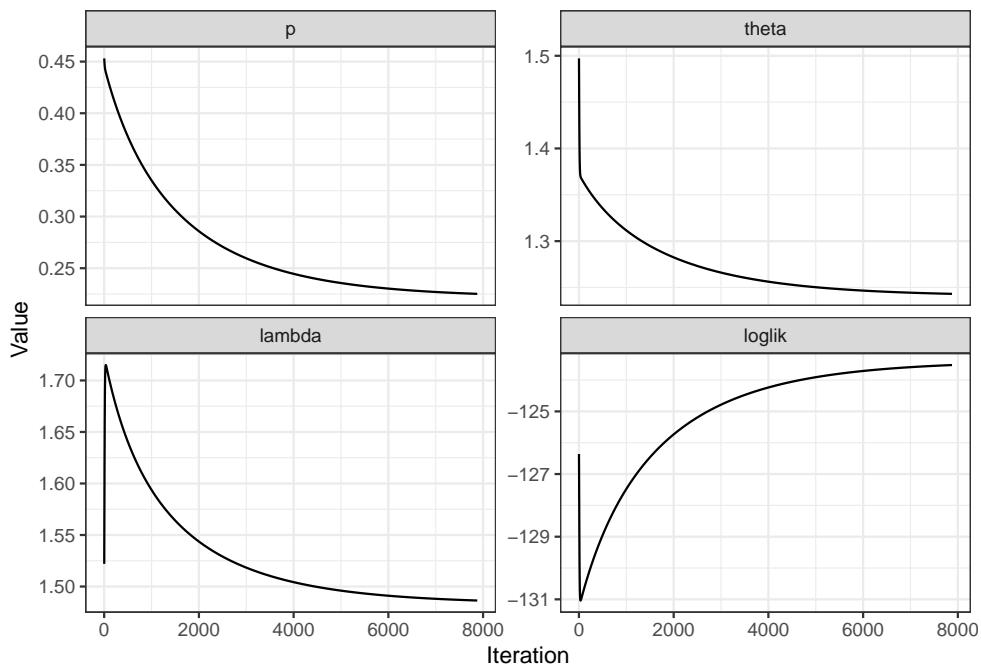


Figure 3.2: Convergence of π , θ , λ and observed log-likelihood.

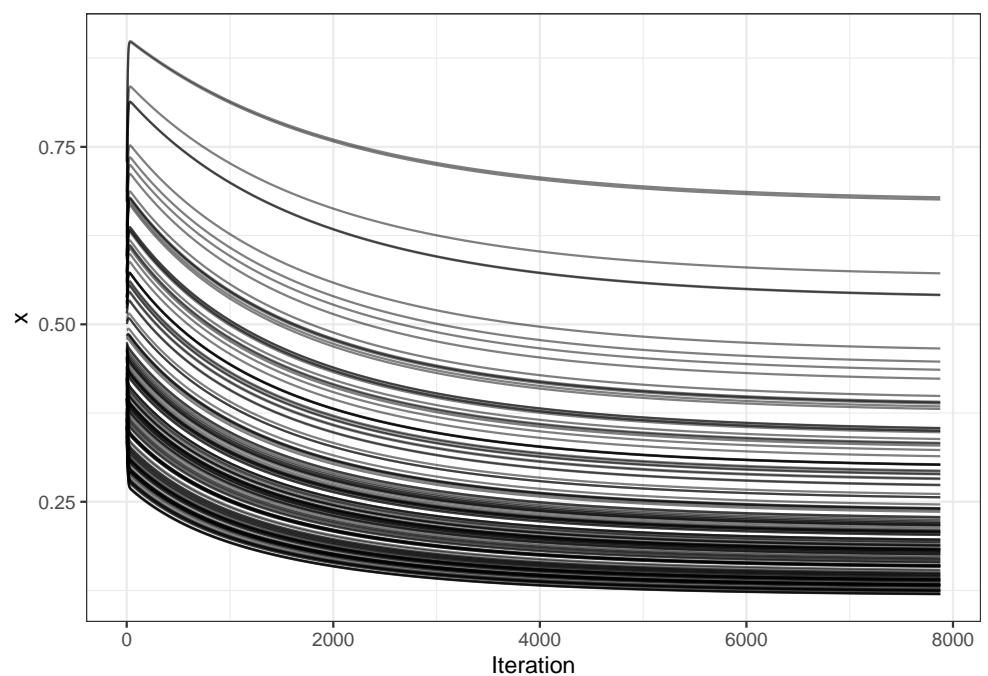


Figure 3.3: Convergence of artificial data x_i .