Alie(n)ation: Political Outsiders in the 2016 U.S.

Presidential Election

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March 4, 2021

Abstract

Political alienation describes a feeling of estrangement from the political system. While some have suggested that political alienation can lead to negative voting in national elections—that is, voting against a particular candidate or entity—the mechanism through which alienation affects such behaviors remains unclear. What leads the alienated to cast negative votes and how do we know those votes signal a rejection of the political system? I argue that the politically alienated can channel their feelings of normlessness and powerlessness through their support for 'political outsiders.' Combining open-ended responses from the 2016 American National Election Studies with newly developed text-analysis tools, I show that those harboring feelings of normlessness and powerlessness were more likely to 'like' Trump for being an outsider. I also show that normlessness increased the likelihood of turning out to vote for outsiders like Donald Trump and Bernie Sanders, and that the vote-suppressing effect of powerlessness on turnout was weakened in 2016.

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¹Replication materials are available <u>here</u>.

1 Introduction

Political alienation describes a feeling of estrangement from the political system (Citrin et al. 1975; Olsen 1969). Perhaps unsurprisingly, the politically alienated show disproportionately low rates of political participation and interest (Finifter 1970; Templeton 1966), and tend to hold more cynical views of political institutions and those who control them (Mc Dill and Ridley 1962; Thompson and Horton 1960; Templeton 1966). When they do participate, the alienated have been known to act in ways that convey their discontent, for instance, casting "negative" votes against local referenda (Horton and Thompson 1962; Mc Dill and Ridley 1962). Some have suggested that these effects of alienation have the potential to spillover into national elections: Aberbach (1969) found that the distrustful were more likely to vote for the outsider, Barry Goldwater, in the 1964 presidential election, and Templeton (1966) found that alienation was related to patterns of vote switching across the 1956 and 1960 presidential elections.

And yet, the mechanism through which feelings of alienation might affect attitudes and behaviors in national elections remains unclear. When and why might the alienated cast negative votes in national elections and how can we be sure that their votes are truly intended as a repudiation of the political system? I argue that the politically alienated can channel their attitudes of powerlessness and normlessnss (Finifter 1970) through their support for 'political outsiders,' such as Donald Trump and Bernie Sanders in the 2016 U.S. presidential election. I argue further that these two categories of alienation will have somewhat nuanced effects on candidate evaluations and voting behavior: those harboring feelings of either powerlessness or normlessness will be more favorable towards political outsiders specifically because they represent a challenge to the political system, but only normlessness will be a positive predictor of turn out and vote choice. On the other hand, powerlessness is, by definition, the belief that one lacks the power to effect change through the political process, so there is no expectation that these attitudes will be expressed through turn out or vote choice.

To empirically evaluate my argument, I rely on data from the American National Election Study (ANES). I begin by using a semi-automated text-analysis approach, the Structural Topic Model (Roberts et al. 2014), to explore open-ended survey responses describing what people liked about Trump. A topic emerges from the model that is directly related to Trump's status as a political outsider. The ability of the Structural Topic Model to incorporate co-variates allows me to show that feelings of powerlessness and normlessness additively increase the likelihood that people see candidates such as Trump as a 'political outsider.' Then, I show that both dimensions of alienation affect one's decision to support outsiders at the ballot box: the vote-suppressing effect of powerlessness in national elections was drastically weakened in 2016, while at the same time, normlessness increased the likelihood of turning out to vote for Sanders in the Democratic primary and Trump in the Republican primary and general election. Together these findings help to explain the emergence and surprising success of two rather untraditional presidential candidates.

The findings presented here speak to a growing line of literature that is concerned with the repercussions of a decline in confidence in the American political system (Citrin and Stoker 2018; Hetherington and Rudolph 2015). Such a foundation of support is essential for the health and sustainability of our democracy (Easton 1965), so it is vital that we identify the changes in political attitudes and behaviors that result from the erosion of that support. A crucial change that I identify in this project is alienated individuals' increased willingness to support candidates for office that present themselves as a challenge to our political system. If a sizable portion of the public continues to harbor feelings of alienation, this could drastically alter the types of candidates that choose to run for office, and ultimately, the types of individuals that wield political power.

2 Political Alienation: Definition and Effects

What does it mean to be politically alienated? The definition given by Citrin et al. (1975, 3) closely reflects the popular conceptualization of political alienation as a "relatively enduring sense of estrangement from existing political institutions, values, and leaders." Typically, feelings of alienation are considered "diffuse" (Easton 1965) in nature, meaning they stem from evaluations of the political system in the broadest sense, and not from evaluations of specific political actors or policies. This definition performs well in capturing the essence of political alienation, but the precise ways in which one is estranged from the political system, and how those feelings of estrangement might influence other political attitudes and behaviors, remain unclear. As such, a number of scholars have attempted to delineate the various modes, dimensions, or categories of alienation.

One of the first scholars to dissect the concept of alienation into its constitutive parts was the social psychologist Melvin Seeman (1959), who identified five different modes: powerlessness, normlessness, meaninglessness, isolation, and self-estrangement. Early studies of political alienation often referred to this typology in their analyses (Horton and Thompson 1962; Aberbach 1969; Olsen 1969), with a particular emphasis being placed on the effects of powerlessness and normlessness. Powerlessness is the belief that one's lacks the necessary means to effect change on the political system (Seeman 1959; Olsen 1969; Finifter 1970) and is closely (and inversely) related to the concept of political efficacy (Campbell et al. 1960; Gamson 1961). We might think of powerlessness as representing a feeling that one is incapable of affecting the "inputs" to the political system (Easton 1965; Almond and Verba 1963). Normlessness, on the other hand, is the belief that the social rules or "norms" that guide our political system have been lost. One might feel normless, for instance, if they believed that their representatives were only representing the interests of corporations, and not those of constituents. In this sense, normlessness flows from evaluations of the system's "outputs" (Easton 1965; Almond and Verba 1963).

The primary way in which feelings of alienation, including normlessness and power-

lessness, are known to influence one's political attitudes is that they produce a sense of "negativism" (Horton and Thompson 1962). For instance, Thompson and Horton (1960) found that the politically alienated were more likely to hold unfavorable views toward a local school bond referendum.¹ Citrin et al. (1975) also showed that the politically alienated held more negative evaluations of the current political climate and were more willing to support systemic change. Whether one's feelings of alienation and the negative attitudes that flow from them will translate into political action, however, depends on the dimension of alienation, as well as the context in which those feelings might be expressed.

On the one hand, feelings of powerlessness often appear negatively related to several forms of political participation such as voting (Horton and Thompson 1962; Aberbach 1969) and discussing politics with others (Olsen 1969; Finifter 1970). Given that powerlessness if the belief that one is incapable of influencing the inputs to the political system, it is unsurprising that the politically powerless would not often use the political process to air their grievances. On the other hand, normlessness appears typically unrelated to political participation (Finifter 1970; Olsen 1969), as those harboring such feelings may or may not see the political process as a viable mechanism for signaling their discontent. These two dimensions of alienation, though they may have similar effects on political attitudes, often have unique effects on political behaviors.

Context becomes a critical factor, then, as there may be situations in which the potential for expressing one's feelings of alienation through their political action becomes clearer, leading the powerless and normless to deviate from their typical behavioral patterns. One important context in which this can occur is elections, where the politically alienated may be motivated to cast "negative" votes—that is, voting against particular candidates, policies, or political entities. For this reason, several early studies of political alienation examined the effects of alienation in the context of local referenda where, unlike typical elections for office, voters are given the option to vote against a particular measure (Mc Dill and Ridley

¹To measure alienation, Thompson and Horton (1960) use an index that taps into both normlessness and powerlessness.

1962; Horton and Thompson 1962; Thompson and Horton 1960). These studies were consistent in their finding that the politically alienated were disproportionately more likely to cast negative votes against the referendum.

Less is known about whether the alienated can be primed to cast negative votes in the context of national elections. One reason for this gap in our knowledge is that opportunities for the alienated to cast such votes rarely arise, largely because referenda and initiatives do not occur at the national level in the United States. However, there are occasions in which the appeals made by particular candidates for office may signal to members of the mass public that a vote for that candidate is a vote against the political system. These types of appeals are typically made by so-called 'political outsiders,' and their presence in U.S. presidential elections is rare. Prior to 2016, the closest example in recent memory was Barry Goldwater's candidacy on the Republican ticket in the 1964 election. Goldwater's strong conservative positions, particularly his opposition to the Civil Rights Act, often placed him at odds with the other candidates, including member's of his own party. This observation lead Aberbach (1969) to suggest that part of Goldwater's support may have derived from the politically alienated—indeed, he found that (only) normlessness was related to voting for Goldwater among both Democrats and Republicans.²

Although Goldwater was outcast from his political party his status as U.S. Senator from Arizona preempted him from being a political outsider in the truest sense. Furthermore, while Aberbach's (1969) analysis may have demonstrated a link between normlessness and voting for Goldwater, the motivation underlying the vote of the politically alienated remained unclear. In the next section, I will argue that the presence of two political outsiders in the 2016 election—Donald Trump and Bernie Sanders—provides an opportunity to re-examine the effects of alienation on one's attitudes and voting behaviors.

²Aberbach (1969) also investigated the interactive effects of powerlessness and normlessness to test the proposition that those alienated on both dimensions would be *more* likely to have been motivated to vote for Goldwater, but no support for an interactive relationship was found. A body of work further explored this interactive relationship—often referred to as the "efficacy-trust mobilization" hypothesis—but the totality of evidence has suggested that the appropriate specification for these two variables is additive (e.g., Sigelman and Feldman 1983; Fraser 1970).

3 Alienation and Outsiders in the 2016 Election

The 2016 presidential election cycle provides a unique context in which to examine whether political alienation can lead to negative voting in a national election. As Templeton (1966) noted long ago, most typical presidential elections feature establishment-type candidates from either party, and the debates tend to center around prominent political issues of the day. In these elections, feelings of alienation are likely to play only a minor role: feelings of powerlessness may dampen participation in the political process as they typically do, while feelings of normlessness may take a backseat to partisan or ideological considerations (Finifter 1970). The 2016 election, however, featured two political outsiders whose presence may have motivated the politically alienated to alter their typical behavior patterns to cast negative votes: Donald Trump, a New York businessman with no prior office-holding experience, infiltrated the ranks of the Republican Party and would go on to win the presidency over the Democratic candidate, Hillary Clinton, who many would consider the ultimate 'political insider.' Bernie Sanders, as one of only a handful of independents to ever hold a seat in the US Senate, put up a serious fight in the 2016 Democratic primary. Compared to Goldwater, these two candidates come much closer to meeting the definition of a 'political outsider.'

How, specifically, will the effects of political alienation manifest in this context? As I have previously discussed, the two dimensions of political alienation—normlessness and powerlessness—have unique effects on attitudes and behaviors. Regarding one's attitudes, both dimensions of alienation are known to produce a sense of "negativism"—that is, an orientation against the existing political structures that entails a sense of cynicism or skepticism. These sort of negative appeals were made repeatedly throughout the campaign by both Donald Trump and Bernie Sanders. Consider the following statement from Trump who is tapping into the feelings of normlessness when speaking at a campaign rally in Sioux City, Iowa (Jackson 2016):

At the heart of this election is a simple question: will our country be governed by the people or will it be governed by the corrupt political class?

Compare this to a statement from Democratic hopeful Bernie Sanders who said the following at the Brookings Institute the same day he announced his intention to seek the nomination (Dews 2015):

There is a lot of sentiment that enough is enough, that we need fundamental changes, that the establishment – whether it is the economic establishment, the political establishment, or the media establishment – is failing the American people.

Although Sanders and Trump sought nominations from opposing political parties, the message being conveyed by either candidate was the same—both were attempting to show that they were the "anti-" candidates by describing the political system as failing and corrupt. I believe that these appeals rang loudest with the politically alienated, who are rarely given a chance to support a candidate that shares their negative orientation toward the political system. As a result, I expect that the politically alienated will hold favorable attitudes towards candidates like Trump and Sanders due specifically to their outsider statuses.

The effects of alienation on voting behaviors in the 2016 election are more nuanced. I argue that the presence of political outsiders in this election provided sufficient incentive for the powerless and normless to deviate from their typical behavior patterns. For the politically powerless, the typically negative relationship with participation may evaporate if the powerless joined the 2016 primary and general elections at rates higher than usual due to their willingness to support these political outsiders at the ballot box. We shouldn't necessarily expect the powerless to become the more likely than the powerful to turn out, but a comparison of the effects of powerlessness across several elections may reveal that its effect in the 2016 election was unique. For the politically normless, the rare opportunity to express one's feelings of discontentment through a negative vote may have been more apparent with Trump and Sanders in the race, thus increasing the likelihood that those feeling normless turned out to vote for them. Initial evidence to support this expectation comes from the 2016 presidential primary elections in New Hampshire where trust, which falls under the veil or normlessness, was related to voting for Sanders and Trump among Democrats and Republicans, respectively (Dyck, Pearson-Merkowitz, and Coates 2018). Although this

partially supports my expectations, it remains unclear if this trend held in other primary elections, as well as in the general election. Furthermore, the motivation for casting those votes remains to be identified. The next section reveals my empirical strategy for testing these expected effects of political alienation on candidate evaluations and voting behaviors in the 2016 election.

4 Data and Methods

To examine the effects of political alienation in the context of the 2016 U.S. Presidential Election, I rely upon data from the American National Election Studies (ANES). The ANES data are well-suited for my purposes, as they contain items that capture several dimensions of political alienation (Mason, House, and Martin 1985), as well as information on vote choice, party affiliation, and various demographics. Additionally, the data include open-ended responses about the things that respondents 'like' about the two major parties' candidates for president, allowing me to examine, in their own words, the reasons that alienated voters may have felt attracted to a political outsider such as Trump. The remainder of this section will focus on describing these data in greater detail, along with my approach to model them.

4.1 Measures of Alienation

There are two primary independent variables of interest in the analyses to follow: *cynicism* and *election unresponsiveness*. The first, *cynicism*, is derived from the 'No Trust' and 'Big Interests' items that are part of the ANES' 'Cynicism' index (Mason, House, and Martin 1985). The 'No Trust' item asks:

with possible answers being (1) "Always," (2) "Most of the time," (3) "About half the time," (4) "Some of the time," or (5) "Never." The 'Big Interests' item asks:

"How often can you trust the federal government in Washington to do what is right?"

"Would you say the government is pretty much run by a few big interests looking out for themselves or that it is run for the benefit of all people?"

with the possible responses being (0) "For the benefit of all people" or (1) "Run by a few big interests." To form the cynicism scale, respondents are given a point for each cynical answer (italicized) that they provide, creating a 3-point scale that ranges from 0 (No cynical answers) to 2 (All cynical answers). From their analysis of the various measures of political alienation contained in the ANES, Mason, House, and Martin (1985) conclude that the 'No Trust' and 'Big Interests' items that I am using here form "a single latent construct reflecting a lack of trust in the ability of the federal government to act in ways that people regard as right and fair." (p. 145). From this definition, it is clear that this measure of cynicism closely reflects attitudes of normlessness (Finifter 1970).

The second variable, *election unresponsiveness*, comes from a single item that asks:

"How much do you feel that having elections makes the government pay attention to what people think?"

to which individuals may respond (0) "A good deal," (1) "Some," or (2) "Not much." This question captures attitudes of *powerlessness*, specifically as it relates to elections. While others have used more general measures of political efficacy when operationalizing *powerlessness* (e.g., Aberbach 1969), the *election unresponsiveness* item is appropriately focused on the context in which I expect alienation to have an effect (i.e., elections).³

My argument regarding the effects of political alienation in national elections leads to three hypotheses that can be tested using these measures of alienation from the ANES. First, I expect that both *cynicism* and *election unresponsiveness* will increase the likelihood that a respondent 'likes' Trump for being a political outsider (Hypothesis 1). As I have argued, alienation on either dimension—powerlessness or normlessness—makes individuals more likely to look favorably upon candidates that present themselves as a challenge to the political system. Next, I expect that *cynicism* will increase the likelihood that individuals

³Unfortunately, I am unable to perform further analyses using both *cynicism* and the *political efficacy* index as no respondents in the 2016 study were asked both sets of questions.

turn out to vote for Sanders in the Democratic primary and for Trump in the Republican primary and general election (Hypothesis 2). As scholars have noted, the two major parties' candidates for presidents are typically establishment figures (Templeton 1966), so with Trump and Sanders in the race, those that were unsatisfied with the outputs of our political system—that is, those feeling normless—were given a rare opportunity to use the national election process to signal their disaffection. Finally, I expect that election unresponsiveness will be negatively related to turnout in national elections, but will have no relationship with turnout or vote choice in 2016. While the politically powerless are no doubt expected to admire candidates with anti-establishment qualities, their belief that elections are ineffective tools for sparking political change means that we should not expect the politically incapable to be more likely to turnout and vote for outsiders. Instead, I expect that the (typical) vote-suppressing effects of incapability were weakened, and certainly not reversed, in 2016 due to the presence of 'political outsiders' in the race.

Hypothesis 1: Cynicism and Election Unresponsiveness both increase the likelihood of 'liking' Trump because he is a political outsider

Hypothesis 2: Cynicism increases the likelihood of turning out and voting for Sanders in the Democratic primary and for Trump in the Republican primary and general election

Hypothesis 3: Election Unresponsiveness is negatively related to turnout in typical national elections, and unrelated to turnout and vote choice in the 2016 election

4.2 Open-Ended 'Likes' about Trump and the Structural Topic Model

For decades, the ANES has included a number of open-ended questions that ask respondents what they like and dislike about the two major parties' candidates for president. To understand whether the politically alienated were attracted to Trump due to his status as a

political outsider (Hypothesis 1), I utilize the open-ended 'likes' about Trump from the 2016 study.⁴ The question that I am interested in asked respondents:

"Is there anything in particular about Donald Trump that might make you want to vote for him?"

If respondents provide a first thought, the interview follows up by asking "anything else?" until the respondent says 'no.'⁵ Of the 4,270 respondents in the 2016 sample, approximately 43% (1,850 respondents) provided a response to this question. Unfortunately, this question was not asked about candidates in the primary election, so I am unable to explore what voters liked about Bernie Sanders.

In the past, the ANES has used human coders to sort the open-ended 'likes' and 'dislikes' into a fixed set of categories, but as of the writing of this manuscript the 2016 open-ended responses concerning the two presidential candidates have yet to be coded. Fortunately, several forms of (semi-) automated content analysis have emerged to assist researchers in categorizing large bodies of text (Grimmer and Stewart 2013). I use one such approach, the Structural Topic Model (Roberts, Stewart, and Tingley 2019), to assist me in categorizing these open-ended 'likes' about Trump. Roberts et al. (2014) has previously shown that the topics that emerge from a Structural Topic Model performed on open-ended responses the the 'most important problem in Washington' question from the ANES closely mimic the ANES' manually-coded categories. As my results show, the STM also performs quite well with the open-ended 'likes' about Trump.

The intuition behind the Structural Topic Model is simple: identify clusters of words that tend to co-occur (i.e., topics). This is the same basic intuition underlying more common forms of topic models such as LDA (Blei, Ng, and Jordan 2003), but the STM is unique in

⁴Redacted versions of these open-ended responses—which are used in this analysis—are publicly available from the website of the American National Election Studies.

⁵For each respondent, their initial response and all follow-ups to the interviewers prime are contained in a single document (or cell) per respondent. There is no unique identifier to separate each respondents 'likes' into different cells, so all mentions from a respondent must be analyzed together. Respondents that provided no 'likes' about Trump are coded as Not Applicable (NA), and their data is not used in the estimation of the Structural Topic Model.

the sense that it allows researchers to include covariates that they suspect will affect 1) the use of certain topics (prevalence), or 2) the use of specific words within a topic (content) (Roberts et al. 2014). My expectation is that feelings of political alienation will increase the frequency with which respondents use the 'political outsider' topic, so I choose to specify cynicism and election unresponsiveness as prevalence covariates.

One advantage of the Structural Topic Model is that the researcher need not provide a set of 'training' documents from which each topic should be built. Instead, the STM takes a 'bottom-up' approach, allowing the machine to infer topics from the data. However, the STM does still require a small amount of supervision, as the researcher must decide the number of topics (K) that are to be found. Roberts, Stewart, and Tingley (2019) note that there isn't necessarily a "correct" number of topics for a given set of documents, and advise researchers to explore models that range in their number of topics, selecting one's model based on measures such as semantic coherence or exclusivity. Following this advice, I undertake an iterative process to narrow-in on a model with 26 topics. 6

4.3 Models of Voting Behavior

My examination of the effects of alienation on voting behavior in the 2016 election centers largely on the results of three models. I begin by using a multinomial logit to model the 2016 primary elections. Here, I restrict my sample to respondents from states that hosted open primaries on Super Tuesday in 2016 and use 'vote choice' as the dependent variable with the options being "Sanders, "Trump," "Other," and "Did Not Vote." 7,8 Restricting my sample in this way allows me to avoid the possibility of unobserved time-dependencies that could arise from pooling respondents that voted at different points in the election cycle. Additionally, focusing on open primaries allows me to include both Democrats and

⁶Appendix B.1 provides a more detailed discussion of the process of selecting the number of topics.

⁷ "Other" includes all candidates other than Trump or Sander—regardless of partisanship—that ran in the 2016 primary elections.

⁸States hosting open primaries on Super Tuesday in 2016 include: Alabama, Arkansas, Georgia, Minnesota, Tennessee, Texas, Vermont, and Virginia.

Republicans in the same model, while also allowing for the possibility that some individuals voted for candidates of the out-party.

Next, I estimate a multinomial logit for the general election, where the dependent variable is once again 'vote choice' with the options being "Clinton," "Trump," "Other," and "Did Not Vote." Unlike the model of the primary election, the sample for the general election is not restricted in any way. Both models include several co-variates that are known to influence both vote choice and turnout including partisanship (treated as categorical with 'Independent' as the reference category), ideology (ranges from 1 = Extremely Liberal to 7 = Extremely Conservative), partisan strength (0 = pure independent to 3 = strong partisan), income (in quartiles), white (1 if white, 0 otherwise), female (1 if female, 0 otherwise), age (17-99 continuous), political interest (1 = not much interested to 3 = very much interested), and education (1 = less than high school to 5 = Masters or above).

Finally, I estimate a series of logistic regressions to model the effects of *cynicism* and *election unresponsiveness* on turnout in the last eight elections (including 2016). This allows me to show that the effects of alienation I find in 2016 are largely unique to that particular election cycle. These models also include a number of co-variates that confound the relationship between either category of alienation and turnout, including partisan strength, education, income, white, age, and an indicator for Independents (all coded the same as above).

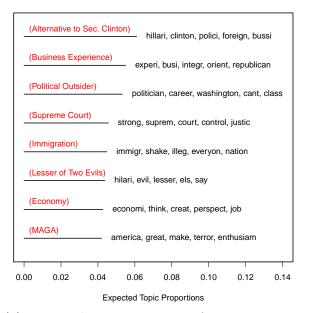
5 Results

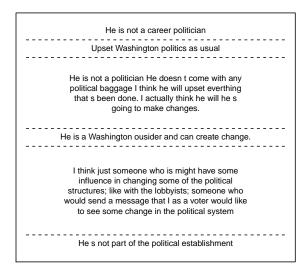
I begin by exploring the results of the Structural Topic Model. Figure 1 presents the top 8 topics from the 26-topic Structural Topic Model along with their expected topic proportions. ¹⁰ I have assigned a label to each topic based on the FREX words that are closely associated with each topic. ¹¹ The topics that arise are both coherent and insightful: many

⁹See Appendix A for a list of all ANES variables used in this analysis.

¹⁰Figure 8 in Appendix B.2 shows the expected topic proportions for all 26 topics.

¹¹FREX words are those that are both Frequent and Exclusive to a topic (Bischof and Airoldi 2012).





- (a) Expected Topic Proportion for Top 8 Topics, Labels in Parentheses
- (b) Exemplary Texts from 'Political Outsider' Topic

Figure 1: Exploring Topics from STM

respondent's stated that they liked Trump for his positions on immigration, the Supreme Court, and the economy; Trump's status as the alternative to his challenger, Hillary Clinton, was the topic with the highest expected proportion; and, of course, the 'Make America Great Again' slogan was parroted by a fair number of respondents. Of particular interest, however, is the topic labeled 'Political Outsider.' This topic is characterized by such words as 'politician,' 'career,' and 'washington.' These words alone do not fully reveal the topic's contents, so I also provide several responses that dedicate a large portion of their content to this topic as shown in Figure 1b. These responses show that Trump was viewed by some as a political outsider that was "not part of the political establishment." The second to last response in Figure 1b is particularly telling, as the respondent is indicating that a vote for Trump was a way of "send[ing] a message that I as a voter would like to see some change in the political system." These sentiments seem to suggest that part of Trump's appeal was rooted in his lack of political experience and his willingness to challenge the political system.

Were the politically alienated particularly inclined to view Trump as a political outsider

(Hypothesis 1)? To answer this question, I regress the proportion of a document dedicated to the 'Political Outsider' topic on the measures of cynicism and election unresponsiveness, as well as partisanship and political interest, and present coefficient plots of the results in Figure 2c. 12 In Figures 2a, 2b, and 2d I also present the results of similar models where the dependent variables is the document-topic proportion dedicated to other popular topics ('Alternative to Sec. Clinton,' 'Business Experience,' and 'Supreme Court,' respectively). I do not have specific expectations regarding the relationship between alienation and the use of these topics, and as we see, not such relationships exist. However, a clear relationship does exist between use of the 'Political Outsider' topic and both cynicism ($\hat{\beta} = 0.013, p < 0.001$) and election unresponsiveness ($\hat{\beta} = 0.013, p < 0.001$), as indicated by the positive and statistically significant coefficient estimates for these variables in Figure 2c. ¹³ Substantively, this means that those in either category of alienation—that is, those with feelings of normlessness or powerlessness—were more likely to say that they like Trump because he is a political outsider. It is important to note here that neither partisanship nor political interest—both of which are among the most influential predictors of political attitudes and behaviors—were related to the perception of Trump as an outsider. From this exercise, I have found clear support for Hypothesis 1.

My next task is to examine the effects of cynicism and election unresponsiveness on voting behaviors in national elections, but especially the 2016 presidential election, with the expectation that cynicism will increase the likelihood of voting for Sanders in the 2016 primary elections and Trump in both the primary and general elections (Hypothesis 2), and that election unresponsiveness will be negatively related to turnout in typical national elections, but unrelated to turnout in 2016 (Hypothesis 3). As previously noted, I use multinomial logits to model vote choice. Interpretation of the coefficients from a multinomial logit is notoriously tricky, however, as the coefficients represent the change in the log-odds

¹²Results are also presented in Table 1 of Appendix B.2.

¹³These effects are robust to the inclusion of a more extensive battery of co-variates including ideology, sex, race, education, and age (see Table 2 of Appendix B.2).

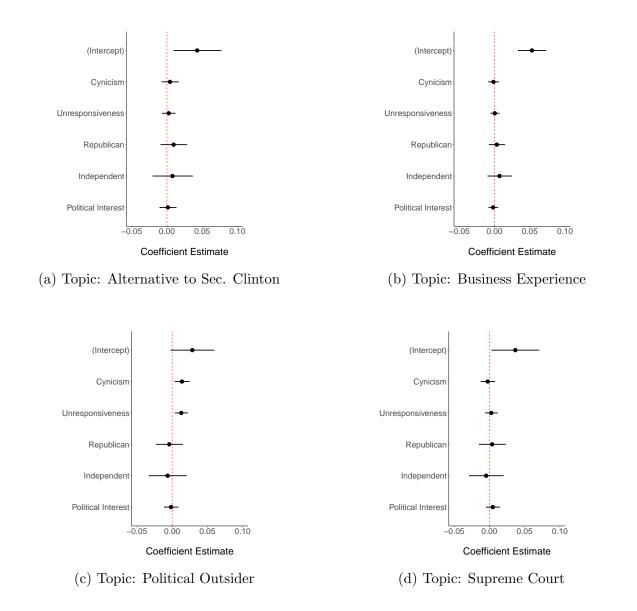


Figure 2: Coefficient Plots from Models of Topic Usage on Cynicism and Unresponsiveness

Note: Estimates from Table ?? in Appendix B.2

of selecting a particular outcome over some baseline category as the result of a one-unit increase in the predictor. Instead, researchers are encouraged to calculate and interpret more substantively meaningful quantities of interest (King, Tomz, and Wittenberg 2000; Paolino 2020), so for both elections, I simulate the predicted probability of selecting each outcome (along with 90% confidence intervals) as cynicism and election unresponsiveness move from their lowest (0) to their highest (2) values using the observed values approach (Hanmer and Ozan Kalkan 2013).

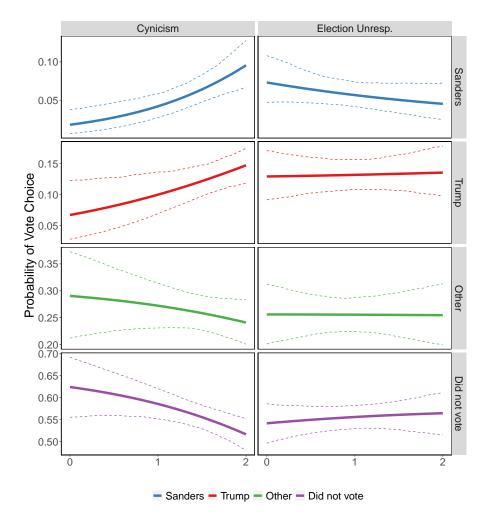


Figure 3: Predicted Probabilities - Vote Choice in 2016 Primary Election

Note: Estimates come from Multinomial Logit in Table 4

I begin by examining the effects of alienation in the open primary elections held on Super Tuesday in 2016, with the model's output presented in Table 4 of Appendix C and the predicted probabilities of vote choice presented in Figure 3. Looking at the plots in the left column of Figure 3 we see that cynicism is clearly related to voting for Sanders or Trump in the primaries: moving from the lowest (0) to the highest (2) value of cynicism significantly increases the probability of voting for Sanders 7.8% (90% CI [0.038,0.115]) and for Trump by 8.5% (90% CI [0.013,0.131]). Figure 3 also shows that cynicism was negatively related to abstention, as well as voting for some candidate than Trump or Sanders: moving across the range of the cynicism scale reduced the probability of not voting by 10.7% (90% CI[-

0.190,-0.022]) and the probability of voting for some other candidate by 4.9% [90% CI[-0.149, 0.049]), though this effect is only significant in the former case.

In the right panel of Figure 3 we see that the effects of election unresponsiveness on vote choice in the primaries are far less pronounced: increases in election unresponsiveness appear to reduce the probability of voting for Sanders by 2.7% (90% CI[-0.075,0.014]) and increase the probability of not voting by 2.3% (90% CI[-0.049,0.99]), though the confidence intervals for these estimates overlap zero. There is essentially no effect of election unresponsiveness on the probability of voting for Trump or some other candidate. My findings that cynicism primed turnout in favor of Trump and Sanders in the primary elections and that election unresponsiveness was unrelated to turnout and vote choice are thus far consistent with Hypotheses 2 and 3.

Now I turn to examining the effects of cynicism and election unresponsiveness in the general election. Table 5 of Appendix C shows the output from the multinomial logit and Figure 4 shows the predicted probability of voting for Clinton, Trump, some other candidate, or not voting at all as cynicism and election unresponsiveness vary from their minimum to their maximum values. In the left column of Figure 4, we see that going from the least to the most cynical attitudes towards government significantly increases the probability of voting for Trump by 7.5% (90% CI[0.033,0.115]) and significantly decreases the probability of not voting (i.e., increases turnout) by 5.7% (90% CI[-0.093,-0.018]). Cynicism also appears to reduce the probability of voting for Clinton by 3.3% (90% CI[-0.071,0.006]) and for some other (third-party) candidate by by 1.6% (90% CI[-0.009,0.038]), but the confidence intervals for these differences in predicted probabilities clearly contain zero. Thus it appears that cynicism motivated turnout in favor of the political outsider in the general election (Trump) just as it did in the primaries; these findings are clearly consistent with Hypothesis 2.

Turning now to the right column of Figure 4, we see that the effects of election unresponsiveness are again far less pronounced than the effects of cynicism: moving across the range of the election unresponsiveness scale produces a slight decrease in the probability of

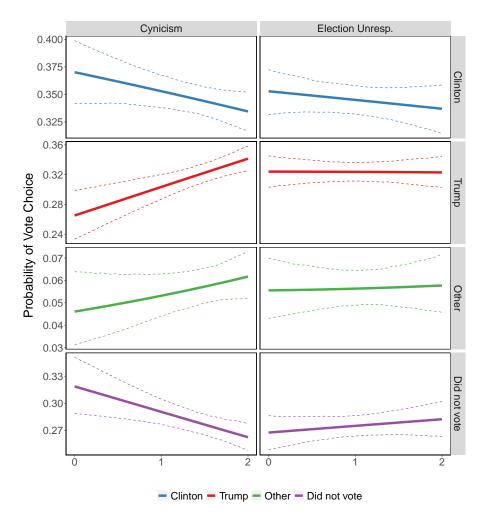


Figure 4: Predicted Probabilities - Vote Choice in 2016 General Election

Note: Estimates come from Multinomial Logit in Table 5

voting for Clinton (1.7%, 90% CI[-0.050,0.018]) and a slight increase in the probability of not voting at all (1.5%, 90% CI[-0.017,0.044]), though neither of these differences are statistically significant. Election unresponsiveness also appears to have essentially no relationship with voting for Trump or some other (third-party) candidate. Here again I have found support for part of Hypothesis 3 that election unresponsiveness would be unrelated to one's decision to turnout, and consequently, unrelated to one's decision to vote for Trump in 2016.

In order to fully support the expectations of Hypothesis 3, however, I must also show that election unresponsiveness is negatively related to turnout in typical national elections. To do this, I use a series of logistic regressions to model turnout in each of the last eight

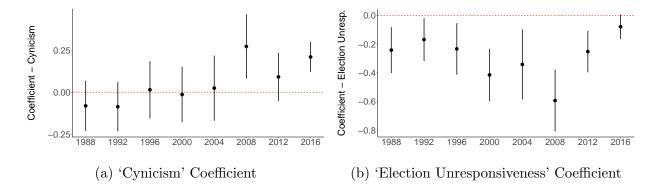


Figure 5: Coefficient Plots from Turnout Models, Estimated in Each Presidential Election Year, 1988-2016

Note: Estimates from Table 3 in Appendix C

presidential elections (1988-2016). These models include cynicism and election unresponsiveness as the primary predictors along with a range of co-variates including partisan strength, eduction, income, age, and indicators for 'independent' and 'white.' I have not hypothesized about the effects of cynicism in typical national elections, but if my argument is correct that the politically alienated channel their frustrations with the political system through their support for political outsiders—and given that Trump is the only outsider to compete in a general election in recent memory—then we might expect cynicism to only be related to turnout in 2016. From these models, I plot the *cynicism* and *election unresponsiveness* logit coefficients in Figures 5a and 5b, respectively.¹⁴

The coefficient estimates for cynicism shown in Figure 5a suggest that it is rarely a motivator for turnout in general elections: in six of the last eight elections, the *cynicism* coefficient is statistically indistinguishable from zero. This is not the case in 2008 and 2016, however, as cynicism is positive and significantly related to turnout in these elections. At the same time, election unresponsiveness appears to consistently depress turnout, with the coefficient on election unresponsiveness being negative and statistically distinguishable from zero in every presidential election between 1988 and 2012 (Figure 5b). It is only during the 2016 general election that election unresponsiveness has no discernible effect on turnout,

¹⁴Full model results are given in Table 3 of Appendix C.

as the 90% confidence interval on the estimate overlaps with zero.¹⁵ This finding provide finding provides further support to Hypothesis 3 that election unresponsiveness would be negatively related to turnout in typical national elections.

6 Conclusion

The feelings of estrangement that characterize political alienation often produce a sense of "negativism" (Horton and Thompson 1962). For the powerless, this typically results in a lack of participation, and for the normless, opportunities to translate one's negative attitudes into action are rarely available. However, I have argued and demonstrated that the 2016 U.S. presidential election cycle was one such opportunity for the politically alienated to cast negative votes at the national level due to the presence of two 'outsider' candidates: Donald Trump and Bernie Sanders. In my analysis, I showed that the politically alienated—whether they harbored feelings of normlessness or powerlessnes—were more likely to say they liked Trump specifically because he stood opposed to the political system. Then, I showed that normlessness had a substantial impact on voting behavior in the 2016 election, increasing the chance that one turned out to vote for Trump or Sanders in the primary and for Trump in the general election. And while powerlessness appeared to have no effect on turnout (and consequently, vote choice) in 2016, I showed that this is an unusual occurrence, as powerlessness has depressed turnout in the last seven presidential elections.

My finding that alienation can be channeled through support for 'political outsiders' is consequential given that typical indicators of alienation, such as the level of trust in government, have fallen rapidly in recent decades (Citrin and Stoker 2018). If feelings of alienation continue to swell in the American public, more opportunistic 'outsiders' may emerge to try and meet the demand. My results also help us to understand that the way in which feelings of alienation will be expressed depends on whether they are feelings of

¹⁵The 90% confidence interval's overlap with zero suggests that this estimate is not significant using a one-tailed test at $\alpha = 0.05$. However, this effect does become significant if one applies a one-tailed test at $\alpha = 0.10$.

normlessness or powerlessness: both dimensions of alienation can change peoples *attitudes* about the political world, but have different effects on how we *behave* in it. Understanding this distinction is crucial to determining exactly how feelings of alienation will manifest in the American public.

An important contribution of this work is that it helps to round-out our understanding of how a candidate like Trump, with no prior office-holding experience or political skills, could become the nominee for a major American political party, and eventually be elected president. Previous explanations have largely pointed to the power of identity as key drivers of Trump's success (Sides, Tesler, and Vavreck 2018) and, to be fair, identity was clearly a key component of the 2016 election. However, my argument regarding the effects of alienation is purely attitudinal as I have no reason to believe that the alienated share a cohesive sense of identity, especially considering that these attitudes are held among members of opposing parties.

The lack of a cohesive identity is also what distinguishes alienation from another label that was commonly used to describe the appeal of both Trump and Sanders: populism. Populism is typically characterized as a struggle between the people vs. the elites or us vs. them (Lee 2019). As I have defined it, alienation certainly entails a disdain for ruling elites, but does not necessarily require one to identify with 'the people.' In fact, the *isolation* dimension of alienation may even entail an outright rejection of the will of the people. Populism and alienation may share similar features (i.e. resentment for the political system), but they are not one and the same.

Up to this point, I have remained agnostic about the specific sources of peoples' feelings of political alienation. As I have defined it, true feelings of alienation are unlikely to have developed rapidly in the lead-up to election day. These attitudes are supposed to be directed toward the political system as a whole, and given that our political system typically changes at a rather glacial pace, it follows that feelings of alienation should be slow-developing, as well. This is not to say, however, that feelings of political alienation that lie dormant cannot

be activated. In fact, I believe that this was likely the case in 2016; Trump and Sanders both used their platforms to make clear to voters that they were the "anti-" candidates in the race, and such appeals, as I have demonstrated, modified the typical relationship between alienation and voting behavior. Identifying the specific source of such attitudes remain beyond the scope of this project, but provide fertile ground for future research.

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Alie(n)ation: Political Outsiders in the 2016 U.S. Presidential Election

Appendix

Maxwell B. Allamong

Appendix A ANES Variables and Coding

For each of the variables below, the codes presented in parentheses are the variable codes for the 2016, 2012, and CDF data files, respectively.

• Cynicism

- 'No Trust'
 - * Variables: V161215; trustgov_trustgrev; VCF0604
 - * Question Wording: "How often can you trust the federal government in Washington to do what is right?"
 - * Coding: 1 = Always, 2 = Most of the time, 3 = About half of the time, 4 = Some of the time, 5 = Never
- 'Big Interests'
 - * Variables: V161216; trustgov_bigintrst; VCF0605
 - * Question Wording: "Would you say the government is pretty much run by a few big interests looking out for themselves or that it is run for the benefit of all the people?"
 - * Coding: 1 = Run by a few big interests, 0 = For the benefit of all the people

• Election Unresponsiveness

- Variables: V161220; respons_elections; VCF0624
- Question Wording: "How much do you feel that having elections makes the government pay attention to what the people think?"
- Coding: 1 = A good deal, 2 = Some, 3 = Not much

• Partisanship

- Variables: V161158x; pid_self; VCF0301
- Coding: indicators (1 = True, 0 otherwise) created for 'Democrat,' 'Independent,'
 and 'Republican'
- Note: 'Democrat' is the reference category for several models included in this paper, therefore, only the 'Republican' and 'Independent' indicators are seen in the tables/figures the present the results of these models

Ideology

- Variables: V161126; libcpre_self; VCF0803
- Coding: 1 = Extremely liberal, 2 = Liberal, 3 = Slightly liberal, 4 = Moderate;
 middle of the road, 5 = Slightly conservative, 6 = Conservative, 7 = Extremely
 conservative

• White

- Variables: V161310x; dem_racecps_white; VCF0105a
- Coding: 1 = White, 0 otherwise

• Income

- Variables: V161361x; inc_incgroup_pre; VCF0114
- Coding:

• Female

- Variables: V161342; gender_respondent_x; VCF0104
- Coding: 1 = Female, 0 otherwise

• Age

- Variables:
- Coding:

• Education

- Variables: V161270; dem_edugroup_x; VCF0110
- Coding: 1 = Less than high school diploma, 2 = High school diploma or equivalent, 3 = Some college but no degree, 4 = Associates degree, 5 = Bachelors degree, 6 = Masters degree, 7 = Professional or doctorate degree

• Political Interest

- Variables:
- Coding:

Appendix B Structural Topic Model

B.1 Model Selection

Roberts et al. (2014) note that there is not necessarily a correct number of topics for any given corpus. Instead, researcher are advised to generate a number of models while varying the number of topics, and then visually inspect the results.

To begin, I noted that the 2008 Likes/Dislikes about Candidates were manually coded by the ANES into roughly 30 categories, so I expect that roughly the same amount of categories in the 2016 data. As such, I use the searchK function from the stm package in R to generate performance diagnostics on models that range in the number of topics from 20 to 40. These diagnostics suggested that the most appropriate number of topics was somewhere in the mid to high 20s, so use the searchK function once again to generate diagnostics on models ranging from 22 to 28. The diagnostic values of these models are shown in Figure ??.

Diagnostic Values by Number of Topics Held-Out Likelihood Residuals 4.3 Held-Out Likelihood Residuals 4. -5.45 23 24 25 26 27 23 24 25 26 27 Number of Topics (K) Number of Topics (K) Semantic Coherence **Lower Bound** -218 Semantic Coherence Lower Bound -34500-222 -226 -3480022 23 24 25 26 27 24 25 26 27 Number of Topics (K) Number of Topics (K)

Figure 6: Determining the Number of Topics to Model, Diagnostics

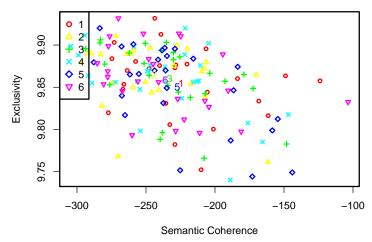


Figure 7: Comparing Semantic Coherence and Exclusivity of Models with Various Initializations

In selecting the number of topics, we are looking for the held-out likelihood and semantic coherence to be high while the residuals should be low. The model with 25 topics seems to fit this pattern quite well. The residuals are clearly lowest in this model, and the held-out likelihood and semantic coherence are quite high. Therefore, I settle on the model with 25 topics.

Because the results of the STM are sensitive to initialization, I then use the selectModel function to generate several models, all with 25 topics. From each of the model runs, I plot the semantic coherence and exclusivity, shown in Figure ??. Notice that models 1 through 6 all show roughly the same values of semantic coherence and exclusivity. Because the models performed so similarly, I manually inspected the topic content from several of the models, and selected the model where the FREX (Frequent-Exclusive) words logically went together and a common theme could be discerned from exemplar texts.

B.2 Model Results

Top Topics

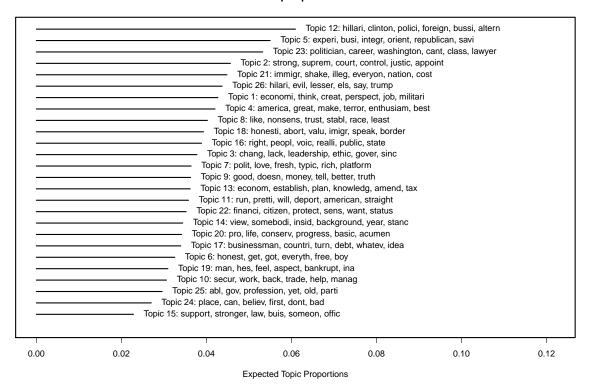


Figure 8: Expected Topic Proportion for All 26 Topics

Table 1: Effect of Cynicism and Election Unresponsiveness on Topic Use (Top 4 Topics)

| | Topic | | | |
|--------------------|-----------------------------|------------------------|-----------------------|--------------------|
| | Alternative to Sec. Clinton | Business Experience | Political Outsider | Supreme Court |
| Cynicism | 0.004 (0.006) | -0.002 (0.003) | 0.013 (0.005) | -0.000 (0.005) |
| Election Unresp. | $0.003 \\ (0.006)$ | -0.000 (0.003) | $0.013 \\ (0.004)$ | $0.001 \\ (0.004)$ |
| Independent | $0.007 \\ (0.015)$ | $0.006 \\ (0.009)$ | -0.002 (0.015) | -0.002 (0.012) |
| Republican | $0.011 \\ (0.010)$ | $0.002 \\ (0.006)$ | -0.002 (0.009) | $0.006 \\ (0.008)$ |
| Political Interest | $0.002 \\ (0.005)$ | -0.003 (0.003) | -0.001 (0.005) | $0.005 \\ (0.005)$ |
| Constant | $0.039 \\ (0.018)$ | $0.060 \\ (0.010)$ | $0.030 \\ (0.016)$ | $0.033 \\ (0.016)$ |
| Observations | 1,375 | 1,375 | 1,375 | 1,375 |

^{*}p<0.1; **p<0.05; ***p<0.01; two-tailed tests
Standard errors in parentheses
Reference category for 'Republican' and 'Independent' is 'Democrat'.
Leaners are included as partisans.

Table 2: Effects of Cynicism and Election Unresponsiveness on Topic Usage

| | Political Outsider (1) |
|--------------------|------------------------------|
| Cynicism | 0.0130 (0.005) |
| Election Unrest. | $0.012 \\ (0.004)$ |
| Independent | -0.006 (0.015) |
| Republican | $0.001 \\ (0.011)$ |
| Political Interest | -0.002 (0.005) |
| Ideology | -0.002 (0.003) |
| Education | $0.004 \\ (0.003)$ |
| White | $0.007 \\ (0.008)$ |
| Female | -0.005 (0.006) |
| Age | $0.000 \\ (0.000)$ |
| Observations | 1,375 |

Observations 1,375

*p<0.1; **p<0.05; ***p<0.01; two-tailed tests
Standard errors in parentheses
Reference category for 'Republican' and
'Independent' is 'Democrat'.
Leaners are included as partisans.

Appendix C Models of Voting Behavior

Table 3: Turnout in the 1988-2016 U.S. Presidential Elections

| | Dependent variable: | | | | | | | |
|--------------------------------|------------------------|------------------------|------------------------|--------------------------|--------------------------|---------------------------|--------------------------|------------------------|
| | 1988 | 1992 | 1996 | 2000 | 2004 | 2008 | 2012 | 2016 |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Cynicism | -0.081 (0.090) | -0.087 (0.089) | 0.014 (0.103) | -0.014 (0.100) | 0.024 (0.117) | 0.272*** (0.115) | 0.091 (0.086) | 0.210*** (0.053) |
| Gov. Unresponsive | -0.242^{***} (0.096) | -0.169^{**} (0.089) | -0.233^{**} (0.108) | -0.415^{***} (0.110) | -0.342^{**} (0.147) | -0.594^{***} (0.130) | -0.252^{***} (0.087) | -0.079^* (0.052) |
| Partisan Strength | 0.619*** (0.091) | 0.382*** (0.087) | 0.565*** (0.103) | 0.424*** (0.103) | 0.552*** (0.126) | 0.680*** (0.124) | 0.493*** (0.081) | 0.143*** (0.047) |
| Education | 0.895*** (0.090) | 0.851*** (0.086) | 0.773*** (0.099) | 0.713*** (0.098) | 0.608*** (0.124) | 0.757*** (0.129) | 0.502*** (0.060) | 0.336*** (0.035) |
| Independent | 0.460** (0.269) | -0.128 (0.232) | 0.039 (0.301) | -0.341 (0.281) | -0.036 (0.361) | 0.054 (0.321) | -0.197 (0.219) | -0.614^{***} (0.142) |
| Income | 0.502*** (0.066) | 0.447*** (0.059) | 0.435*** (0.072) | 0.341*** (0.075) | 0.292*** (0.084) | 0.144** (0.086) | 0.189*** (0.047) | 0.082*** (0.028) |
| White | -0.185 (0.162) | 0.248** (0.141) | -0.005 (0.177) | 0.084 (0.180) | 0.381** (0.200) | -0.044 (0.188) | -0.228^* (0.141) | 0.143** (0.081) |
| Age | 0.038*** (0.004) | 0.030*** (0.004) | 0.034*** (0.005) | 0.027*** (0.005) | 0.014*** (0.006) | 0.023*** (0.006) | 0.038*** (0.004) | 0.015*** (0.002) |
| Constant | -4.810^{***} (0.448) | -3.848^{***} (0.404) | -4.250^{***} (0.511) | -2.952^{***} (0.502) | -2.399^{***} (0.552) | $-2.540^{***} \\ (0.559)$ | -2.501^{***} (0.345) | -1.911^{***} (0.198) |
| Observations Akaike Inf. Crit. | 1,492 1,442.129 | 1,914 1,709.360 | 1,336 1,219.054 | 1,233 1,120.179 | 901 775.303 | 921 808.480 | 2,387 1,909.729 | 3,910 4,687.162 |

*p<0.1; **p<0.05; ***p<0.01; one-tailed tests

 $\it Note:$ These are the regression estimates used to produce Figures 5a and 5b

Table 4: Effect of Cynicism and Election Unresponsiveness on Vote Choice in 2016 Primary Election, Multinomial Logit

| | Sanders | Trump | Did Not Vote |
|-----------------------|-----------|------------|--------------|
| Cynicism | 1.086*** | 0.612** | 0.029 |
| | (0.369) | (0.304) | (0.170) |
| Unresponsive to Elec. | -0.330 | 0.014 | -0.027 |
| • | (0.312) | (0.215) | (0.156) |
| Republican | -1.879** | 16.297*** | 0.255 |
| • | (0.790) | (0.556) | (0.330) |
| Independent | -0.012 | 16.919*** | 0.402 |
| • | (0.864) | (0.513) | (0.541) |
| Ideology | -0.340** | 0.047 | 0.015 |
| G. | (0.173) | (0.159) | (0.094) |
| Income | 0.063 | 0.005 | 0.192** |
| | (0.156) | (0.117) | (0.086) |
| Partisan Strength | 0.359 | 0.719*** | 0.469*** |
| O | (0.259) | (0.208) | (0.137) |
| Education | 0.323 | 0.686*** | 0.354*** |
| | (0.218) | (0.163) | (0.110) |
| White | 0.335 | 0.234 | -0.085 |
| | (0.458) | (0.451) | (0.258) |
| Female | -0.649 | -0.153 | 0.324 |
| | (0.420) | (0.307) | (0.217) |
| Age | -0.018 | 0.051*** | 0.037*** |
| - | (0.014) | (0.010) | (0.007) |
| Constant | -3.087** | -25.216*** | -5.679*** |
| | (1.362) | (0.825) | (0.713) |
| Akaike Inf. Crit. | 1,103.091 | 1,103.091 | 1,103.091 |
| Observations = | 602 | 602 | 602 |

^{*}p<0.1; **p<0.05; ***p<0.01; one-tailed tests

Reference category for dependent variables is 'Any candidate other than Trump or Clinton'

Reference category for 'Independent' and 'Republican' is 'Democrat' Analysis limited to Super Tuesday states with open primaries

Table 5: Effect of Cynicism and Election Unresponsiveness on Vote Choice in 2016 General Election, Multinomial Logit

| | Clinton | Trump | Other |
|-----------------------|-----------|-----------|-----------|
| Cynicism | -0.032 | 0.328*** | 0.266** |
| | (0.084) | (0.093) | (0.143) |
| Unresponsive to Elec. | -0.070 | -0.020 | -0.004 |
| | (0.081) | (0.081) | (0.127) |
| Independent | -2.023*** | 1.928*** | 0.150 |
| | (0.175) | (0.189) | (0.238) |
| Republican | -1.019*** | 0.840*** | -1.175*** |
| | (0.230) | (0.273) | (0.341) |
| Ideology | -0.275*** | 0.331*** | 0.021 |
| | (0.048) | (0.055) | (0.081) |
| Income | 0.124*** | 0.029 | 0.082 |
| | (0.044) | (0.045) | (0.069) |
| Partisan Strength | 0.287*** | 0.118* | -0.610*** |
| | (0.073) | (0.075) | (0.125) |
| Education | 0.409*** | 0.149*** | 0.232*** |
| | (0.056) | (0.057) | (0.090) |
| White | -0.138 | 0.900*** | 0.267 |
| | (0.124) | (0.153) | (0.213) |
| Female | 0.016 | 0.054 | 0.163 |
| | (0.112) | (0.113) | (0.175) |
| Age | 0.016*** | 0.022*** | -0.011** |
| | (0.003) | (0.003) | (0.006) |
| Constant | -1.077*** | -5.997*** | -1.805*** |
| | (0.350) | (0.396) | (0.542) |
| Akaike Inf. Crit. | 5,247.994 | 5,247.994 | 5,247.994 |
| Observations = | 2,996 | 2,996 | 2,996 |

*p<0.1; **p<0.05; ***p<0.01; one-tailed tests Reference category for the dependent variables is 'Any candidate other than Trump or Clinton'

Reference category for 'Independent' and 'Republican' is 'Democrat'

Table 6: Effect of Cynicism and Election Unresponsiveness on Vote Preference in 2016 General Election, Multinomial Logit

| | Trump | Other/Third-Party |
|-----------------------|----------------|-------------------|
| Cynicism | 0.489* | 0.696** |
| | (0.297) | (0.329) |
| Unresponsive to Elec. | 0.415^{*} | 0.358 |
| - | (0.272) | (0.293) |
| Independent | 3.319*** | 0.552 |
| - | (0.820) | (0.812) |
| Republican | 3.742*** | 1.587*** |
| - | (0.524) | (0.555) |
| Ideology | 0.438** | 0.113 |
| 3. | (0.185) | (0.179) |
| Income | -0.054 | -0.195 |
| | (0.155) | (0.169) |
| Partisan Strength | 0.209 | -0.599** |
| | (0.285) | (0.325) |
| Education | -0.532*** | 0.062 |
| | (0.190) | (0.219) |
| White | 1.388*** | 0.210 |
| | (0.428) | (0.451) |
| Female | -1.074^{***} | -0.844** |
| | (0.408) | (0.440) |
| Age | -0.008 | -0.068*** |
| | (0.012) | (0.018) |
| Constant | -3.801*** | 0.707 |
| | (1.186) | (1.188) |
| Akaike Inf. Crit. | 387.221 | 387.221 |
| Observations = | | |

^{*}p<0.1; **p<0.05; ***p<0.01; one-tailed tests

Sample limited to non-voters in 2016 General Election

Reference category for the dependent variables is 'Clinton'

than Trump or Clinton'

Reference category for 'Independent' and 'Republican' is 'Democrat'

References

Roberts, Margaret E., Brandon M. Stewart, Dustin Tingley, Christopher Lucas, Jetson Leder-Luis, Shana Kushner Gadarian, Bethany Albertson, and David G. Rand. 2014. "Structural Topic Models for Open-Ended Survey Responses". *American Journal of Political Science* 58, no. 4 (): 1064–1082. ISSN: 00925853. doi:10.1111/ajps.12103.