

Math, Problem Set #1, Probability Theory

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Due Monday, June 26 at 8:00am

1. **Exercises from chapter.** Do the following exercises in Chapter 3 of ? : 3.6, 3.8, 3.11, 3.12 (watch this movie [clip](#)), 3.16, 3.33, 3.36.

Solutions:

3.6

Assumptions:

(a) $\Omega = \cup_{i \in I} B_i$

(b) $B_i \cap B_j = \emptyset$

Proof:

$$P(A) = P(\Omega \cap A) = P(\cup_{i \in I} B_i \cap A) = P(\cup_{i \in I} (B_i \cap A)) = \sum_{i \in I} P(A \cap B_i)$$

3.8

First, if the events E_k are independent, then their complements E_k^c are independent, too.

$$\begin{aligned} 1 - \prod_{k=1}^n (1 - P(E_k)) &= 1 - \prod_{k=1}^n (P(E_k)^c) \\ &= 1 - P(\cap_{k=1}^n E_k^c) \end{aligned}$$

and by De Morgan's law we get

$$\begin{aligned} &= 1 - P((\cup_{k=1}^n E_k)^c) \\ &= P(\cap_{k=1}^n E_k) \end{aligned}$$

3.11 TODO

3.12

In the first round the unconditional probability of picking the door with the car out of the three doors is just $P(car) = \frac{1}{3}$. After having observed one goat the conditional probability of choosing the door with the car is $P(car | \text{seen one goat}) = \frac{1}{2}$ because one chooses randomly out of two possible choices.

When there are 10 doors $P(car) = \frac{1}{10}$ and $P(car | \text{seen 8 goats}) = \frac{1}{2}$ because, again, you are randomly choosing one out of two possible choices conditional on the 8 goats you have already observed.

3.16

$$\begin{aligned} \text{Var}[X] &= E[(X - \mu)^2] = E[X^2 - 2X\mu + \mu^2] \\ &= E[X^2] - 2E[X]\mu + \mu^2 = E[X^2] - 2\mu^2 + \mu^2 \\ &= E[X^2] - \mu^2 \end{aligned}$$

3.33

Since $B \sim \text{Binom}(n, p)$, $E[B] = np$ and $\text{Var}[B] = \sigma^2 = p(1 - p)$

$$P\left(\left|\frac{B}{n} - p\right| \geq \varepsilon\right) = P(|B - np| \geq n\varepsilon)$$

And by the Chebyshev's Inequality we get

$$\begin{aligned} P(|B - np| \geq n\varepsilon) &\leq \frac{np(1 - p)}{n^2\varepsilon^2} \\ &\leq \frac{p(1 - p)}{n\varepsilon^2} \end{aligned}$$

3.36

Since $X_i \sim \text{Bernoulli}(p)$ for $i \in (1, 6242)$ with $E[x] = p = 0.801$ and $\text{Var}[X] = \sigma^2 = p(1 - p) = 0.199 \times 0.801$, let the number of students actually enrolling be $S = \sum_{i=1}^{6242} X_i$, then the variable $\frac{S - 6242p}{\sigma\sqrt{6242}} \sim N(0, 1)$. Then the probability that more than 5500 students will enroll is given by $1 - P(x < \frac{5500 - 6242p}{\sigma\sqrt{6242}}) = 0$. This probability is practically zero.

2. Construct examples of events A , B , and C , each of probability strictly between 0 and 1, such that

- (a) $P(A \cap B) = P(A)P(B)$, $P(A \cap C) = P(A)P(C)$, $P(B \cap C) = P(B)P(C)$, but $P(A \cap B \cap C) \neq P(A)P(B)P(C)$.
- (b) $P(A \cap B) = P(A)P(B)$, $P(A \cap C) = P(A)P(C)$, $P(A \cap B \cap C) = P(A)P(B)P(C)$, but $P(B \cap C) \neq P(B)P(C)$. (Hint: You can let Ω be a set of eight equally likely points.)

Solutions: TODO

3. Prove that Benford's Law is, in fact, a well-defined discrete probability distribution.

Solutions: TODO

4. A person tosses a fair coin until a tail appears for the first time. If the tail appears on the n th flip, the person wins 2^n dollars. Let the random variable X denote the player's winnings.
- (a) (St. Petersburg paradox) Show that $E[X] = +\infty$.
- (b) Suppose the agent has log utility. Calculate $E[\ln X]$.

Solutions:

Let X be the payoff from the St. Petersburg game, then its expected value can be calculated by

$$E[X] = \sum_{k=1}^{\infty} \left(\frac{1}{2}\right)^k 2^k = 1 + 1 + 1 + 1 + \dots = +\infty$$

When the payoff of one game is weighted with the utility function $u(x) = \ln(x)$, the expected utility is given by

$$E[u(X)] = \sum_{k=1}^{\infty} \left(\frac{1}{2}\right)^k \ln(2^k) = \ln(2) \sum_{k=1}^{\infty} \frac{k}{2^k} = \ln(2)S$$

where

$$\begin{aligned} S &= \sum_{k=1}^{\infty} \frac{k}{2^k} = S - \frac{1}{2}S + \frac{1}{2}S = S - \sum_{k=1}^{\infty} \frac{1}{2} \frac{k}{2^k} + \frac{1}{2}S = S - \sum_{k=1}^{\infty} \frac{k}{2^{k+1}} + \frac{1}{2}S \\ &= \sum_{k=1}^{\infty} \frac{k}{2^k} - \sum_{k=2}^{\infty} \frac{k-1}{2^k} + \frac{1}{2}S = \sum_{k=1}^{\infty} \frac{1}{2^k} + \frac{1}{2}S = 1 + \frac{1}{2}S \Rightarrow S = 2 \end{aligned}$$

so,

$$E[X] = 2\ln(2)$$

5. (Siegel's paradox) Suppose the exchange rate between USD and CHF is 1:1. Both a U.S. investor and a Swiss investor believe that a year from now the exchange rate will be either 1.25 : 1 or 1 : 1.25, with each scenario having a probability of 0.5. Both investors want to maximize their wealth in their respective home currency (a year from now) by investing in a risk-free asset; the risk-free interest rates in the U.S. and in Switzerland are the same. Where should the two investors invest?
6. Consider a probability measure space with $\Omega = [0, 1]$.

- (a) Construct a random variable X such that $E[X] < \infty$ but $E[X^2] = \infty$.
 - (b) Construct random variables X and Y such that $P(X > Y) > \frac{1}{2}$ but $E[X] < E[Y]$.
 - (c) Construct random variables X , Y , and Z such that $P(X > Y)P(Y > Z)P(X > Z) > 0$ and $E(X) = E(Y) = E(Z) = 0$.
7. Let the random variables X and Z be independent with $X \sim N(0, 1)$ and $P(Z = 1) = P(Z = -1) = \frac{1}{2}$. Define $Y = XZ$ as the product of X and Z . Prove or disprove each of the following statements.
- (a) $Y \sim N(0, 1)$.
 - (b) $P(|X| = |Y|) = 1$.
 - (c) X and Y are not independent.
 - (d) $Cov[X, Y] = 0$.
 - (e) If X and Y are normally distributed random variables with $Cov[X, Y] = 0$, then X and Y must be dependent.
8. Let the random variables X_i , $i = 1, 2, \dots, n$, be i.i.d. having the uniform distribution on $[0, 1]$, denoted $X_i \sim U[0, 1]$. Consider the random variables $m = \min\{X_1, X_2, \dots, X_n\}$ and $M = \max\{X_1, X_2, \dots, X_n\}$. For both random variables m and M , derive their respective cumulative distribution (cdf), probability density function (pdf), and expected value.
9. You want to simulate a dynamic economy (e.g., an OLG model) with two possible states in each period, a “good” state and a “bad” state. In each period, the probability of both shocks is $\frac{1}{2}$. Across periods the shocks are independent. Answer the following questions using the Central Limit Theorem and the Chebyshev Inequality.
- (a) What is the probability that the number of good states over 1000 periods differs from 500 by at most 2%?
 - (b) Over how many periods do you need to simulate the economy to have a probability of at least 0.99 that the proportion of good states differs from $\frac{1}{2}$ by less than 1%?
10. If $E[X] < 0$ and $\theta \neq 0$ is such that $E[e^{\theta X}] = 1$, prove that $\theta > 0$.