

Four fitting models comparison for SARS-nCoV-2 total cumulative cases curves

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May 4, 2020

1 Abstract

Four different models have been compared for fitting SARS-nCoV-2 total cumulative cases curves in 187 countries over a period of 103 days. Evaluated models have been: *Simple Logistic Function* (**SLF**), *Simple Gompertz Function* (**SGF**), *Double Logistic Function* (**DLF**) and *Double Gompertz Function* (**DGF**). **DGF** model showed lower MSE, RMSE, NRMSE, MAE, NAE and higher Pearson R compared to the others. R^2 ($p > .99$), R^2_{adj} ($p > .99$), ΔAIC (> 0) and ΔBIC ($p < .5$) showed higher percentages for **DGF** compared with the others.

Results suggest that *Double Gompertz Function* may be a good fitting model for SARS-nCoV-2 cumulative cases curve.

2 Methods

2.1 Data

SARS-nCoV-2 total cumulative cases data have been gathered from Johns Hopkins University GitHub repository [1] and summed into single countries where regional level was provided [2]. Data have been used “as is” without rejecting any outlier and/or error (negative daily Δ). Data and results have been stored in a **pandas** n-dimensional **DataFrame**.

Raw data contained 187 countries and daily cumulative confirmed cases for 103 days, from 2020-01-22 to 2020-05-04.

Data have been collected from: Afghanistan, Albania, Algeria, Andorra, Angola, Antigua and Barbuda, Argentina, Armenia, Australia, Austria, Azerbaijan, Bahamas, Bahrain, Bangladesh, Barbados, Belarus, Belgium, Belize, Benin, Bhutan, Bolivia, Bosnia and Herzegovina, Botswana, Brazil, Brunei, Bulgaria, Burkina Faso, Burma, Burundi, Cabo Verde, Cambodia, Cameroon, Canada, Central African Republic, Chad, Chile, China, Colombia, Comoros, Congo (Brazzaville), Congo (Kinshasa), Costa Rica, Cote d’Ivoire, Croatia, Cuba, Cyprus, Czechia, Denmark, Diamond Princess, Djibouti, Dominica, Dominican Republic, Ecuador, Egypt, El Salvador, Equatorial Guinea, Eritrea, Estonia, Eswatini, Ethiopia, Fiji, Finland, France, Gabon, Gambia, Georgia, Germany, Ghana, Greece, Grenada, Guatemala, Guinea, Guinea-Bissau, Guyana, Haiti, Holy See, Honduras, Hungary, Iceland, India, Indonesia, Iran, Iraq, Ireland, Israel, Italy, Jamaica, Japan, Jordan, Kazakhstan, Kenya, Korea, South, Kosovo, Kuwait, Kyrgyzstan, Laos, Latvia, Lebanon, Liberia,

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Libya, Liechtenstein, Lithuania, Luxembourg, MS Zaandam, Madagascar, Malawi, Malaysia, Maldives, Mali, Malta, Mauritania, Mauritius, Mexico, Moldova, Monaco, Mongolia, Montenegro, Morocco, Mozambique, Namibia, Nepal, Netherlands, New Zealand, Nicaragua, Niger, Nigeria, North Macedonia, Norway, Oman, Pakistan, Panama, Papua New Guinea, Paraguay, Peru, Philippines, Poland, Portugal, Qatar, Romania, Russia, Rwanda, Saint Kitts and Nevis, Saint Lucia, Saint Vincent and the Grenadines, San Marino, Sao Tome and Principe, Saudi Arabia, Senegal, Serbia, Seychelles, Sierra Leone, Singapore, Slovakia, Slovenia, Somalia, South Africa, South Sudan, Spain, Sri Lanka, Sudan, Suriname, Sweden, Switzerland, Syria, Taiwan*, Tajikistan, Tanzania, Thailand, Timor-Leste, Togo, Trinidad and Tobago, Tunisia, Turkey, US, Uganda, Ukraine, United Arab Emirates, United Kingdom, Uruguay, Uzbekistan, Venezuela, Vietnam, West Bank and Gaza, Western Sahara, Yemen, Zambia, Zimbabwe.

2.2 Models

Models have been defined with `lmfit` (implementation of classical `curve_fit` in `scipy`) using Nelder-Mead method for fitting [2].

Total residual from each function have been initially compared (unsorted, sorted, gaussian distribution) to find the model with residual μ closer to 0 and shorter σ . *Akaike Information Criterion* differences (ΔAIC) in relative probability density space, *Bayesian Information Criterion* differences (ΔBIC), R^2 and R^2_{adj} coefficients mean have been used to find the likely better fitting model that has been finally compared, country by country. See Section 7 for formulae.

Models have been defined as follow:

- Simple Logistic Function (**SLF**):

```
def simple_logistic_function(x, a, b, k, e):
    d = k * (b - np.array(x))
    return (a / (1 + np.exp(d))) + e
```

$$f(t) = \frac{a}{1 + e^{k(b-t)}} + \varepsilon$$

- Double Logistic Function (**DLF**):

```
def double_logistic_function(x, a1, b1, k1, a2, b2, k2, e):
    d1 = k1 * (b1 - np.array(x))
    g1 = a1 / (1 + np.exp(d1))
    d2 = k2 * (b2 - np.array(x))
    g2 = (a2 - a1) / (1 + np.exp(d2))
    return g1 + g2 + e
```

$$f(t) = \frac{a_1}{1 + e^{k_1(b_1-t)}} + \frac{a_2 - a_1}{1 + e^{k_2(b_2-t)}} + \varepsilon$$

- Simple Gompertz Function (**SGF**):

```
def simple_gompertz_function(x, a, b, k, e):
    exp = - np.exp(k * (b - x))
    return a * np.exp(exp) + e
```

$$f(t) = a \cdot e^{-e^{k(b-t)}} + \varepsilon$$

- Double Gompertz Function (**DGF**):

```
def double_gompertz_function(x, a1, b1, k1, a2, b2, k2, e):
    exp1 = - np.exp(k1 * (b1 - x))
    g1 = a1 * np.exp(exp1)
    exp2 = - np.exp(k2 * (b2 - x))
    g2 = (a2 - a1) * np.exp(exp2)
    return g1 + g2 + e
```

$$f(t) = a_1 \cdot e^{-e^{k_1(b_1-t)}} + (a_2 - a_1) \cdot e^{-e^{k_2(b_2-t)}} + \varepsilon$$

3 Model fitting

Fitting has been performed with `lmfit` using Nelder-Mead method

```
model = lmfit.Model(function)
result = model.fit(data=y, params=p, x=x, method='Nelder', nan_policy='omit')
```

initial parameters `p` have been guessed as follows (where y are observed values):

- **SLF** and **SGF**

```
p = model.make_params(
    a=y[-1],
    b=max_y_i,
    k=.1,
    e=y[0]
)
```

$$a = y_{-1}$$

$$b = x_{\max(dy)}$$

$$k = 0.1$$

$$\varepsilon = y_0$$

- **DLF** and **DGF**

```
p = model.make_params(
    a1=y[max_y_i] * 2,
    b1=max_y_i,
    k1=.1,
    a2=max(y),
    b2=len(y),
    k2=.1,
    e=y[0]
)
```

$$a_1 = 2y_{\max(dy)}$$

$$b_1 = x_{\max(dy)}$$

$$k_1 = 0.1$$

$$a_2 = \mathbf{max}(y)$$

$$b_2 = x_{y_{-1}}$$

$$k_2 = 0.1$$

$$\varepsilon = y_0$$

Fitting failed for 1 countries (Saint Lucia with DLF) returning best fit information from 187 countries, for a total of 19261 observed and 76941 predicted values.

Complete `python` backend for data gathering, fitting and analysis along with a `pickle` saved dataframe of all measured data and results is online available [2].

Fitting examples are reported in figures [REF] [REF] [REF].

4 Analysis

Several skill score have been used to evaluate to average skill and skill interpolating extreme values (outliers).

Mean absolute error (**MAE**) is a natural, unambiguous, measure of average error [3]. It shows the errors in the same unit as variables themselves. MAE is bounded below by 0 (best case) and unbounded above. Advantage over Mean Bias Error (**MBE**) is that, taking absolute error values, positive and negative errors can't cancel out [4].

Normalized Absolute Error (**NAE**), bounded below 0 (best case) and “virtually” unbounded above. If more than 1 errors are greater than observed values themselves. Advantage over Normalized Bias Error (**NBE**) is that, taking absolute error values, positive and negative errors can't cancel out. [REF]

Mean Squared Error (**MSE**), variance, taking the square of residual is highly sensitive to large outliers [4]. Bounded below 0 (best case) and unbounded above.

Root Mean Squared Error (**RMSE**) is very commonly used as a measure of deviation from the observed value. Although it has been criticized as being ambiguous [3] and its dependence on the squared error means that it is not resistant to outliers deviating from a Gaussian distribution. It has been included because of its sensitivity to large outliers. Bounded below 0 (best case) and unbounded above.

Normalized Root Mean Squared Error (**NRMSE**) allows to compare **RMSE** of different models, normalized on observed values, on a $(0, 1]$ scale [REF].

Pearson Correlation coefficient (**Pearson R**) depends on squared deviations and so is not a resistant measure. However, this statistic removes the effect of any bias in the interpolated data. Problems with correctly capturing the variance will not be highlighted as the measure normalizes the observed and modeled values by their standard deviations [4]. The statistic is standardized. However,

because of its insensitivity to biases and errors in variance, the correlation coefficient should be considered as a measure of potential skill [5].

R^2 coefficient test ... Bounded from 0 to 1, $R^2 > .99$ has been fixed to evaluate model H_0 against alternative model H_1 [6].

R^2_{adj} coefficient test ... Bounded from 0 to 1, $R^2_{adj} > .99$ has been fixed to evaluate model H_0 against alternative model H_1 [6].

Akaike Information Criterion (**AIC**), its score (or weight) and its relative probability distribution space ... Within the probability distribution space, bounded from 0 to 1, a ΔAIC_p value less than 0.5 means model H_0 has more chances to be better fitting than alternative model H_1 [7].

Bayesian Information Criterion (**BIC**) and its delta ... A $\Delta\text{BIC} > 0$ means model H_0 has evidence to be better fitting than alternative model H_1 . Look in Appendix for formulae [8].

4.1 Analysis Example

Example of analysis of two fitting models for a noisy observed sample.

Observed o values are generated by

$$o = f(x) = a \cdot x^2 + b + \varepsilon_x$$

where $x = (0, 100]$, $a = 2.3$, $b = 1000$ and ε_x is a random noise $(0, 3000]$.

Observed are fitted with two models using `scipy.optimize.curve_fit` (least squares method):

$$e_1(x) = a \cdot x^3$$

and

$$e_2(x) = a \cdot x^2 + b$$

Model e_2 is taken as null hypothesis H_0 .

```
import numpy as np
from scipy import stats as sts
from scipy.optimize import curve_fit
import sklearn.metrics as skl
from matplotlib import pyplot as plt

def AICp(diff):
    try:
        return np.exp(-.5 * diff) / (1 + np.exp(-.5 * diff))
    except Exception as err:
        if diff > 0:
            return 1.
        return 0.
```

```

def func_noise(x, a, b, err=False):
    noise = np.random.normal(0, 3e3, len(x))
    return a * np.array(x) ** 2 + b + (noise if err else 0)

def func1(x, a):
    return a * np.array(x) ** 3

def func2(x, a, b):
    return a * np.array(x) ** 2 + b

def report(o, E, P):
    cT = []
    mae, nae, mbe, nbe, mse, rmse, nrmse, r, r2, r2a, aic, bic = (
        [] for _ in range(12)
    )
    for i, e in enumerate(E):
        mae.append(skl.mean_absolute_error(o, e))
        nae.append(np.sum(np.abs(o - e)) / np.sum(o))
        mbe.append(np.sum(o - e) / len(o))
        nbe.append(np.sum(o - e) / np.sum(o))
        _mse = skl.mean_squared_error(o, e)
        mse.append(_mse)
        _rmse = np.sqrt(_mse)
        rmse.append(_rmse)
        _nrmse = _rmse / np.sum(o)
        nrmse.append(_nrmse)
        r.append(sts.pearsonr(o, e)[0])
        r2.append(skl.r2_score(o, e))
        rss = np.sum(np.abs(o - e) ** 2)
        tss = np.sum(np.abs(o - np.mean(o)) ** 2)
        r2a.append(1 - (rss / tss) * ((len(o) - 1) / (len(o) - P[i] - 1)))
        aic.append(len(o) * np.log(rss / len(o)) + 2 * 1)
        bic.append(len(o) * np.log(rss / len(o)) + np.log(len(o)) * P[i])
    cT.append([f"{mae[0]:.0f}", f"{mae[1]:.0f}", f"{(mae[0]-mae[1]):.0f}"])
    cT.append([f"{nae[0]}", f"{nae[1]}", f"{(nae[0]-nae[1])}"])
    cT.append([f"{mbe[0]:.0f}", f"{mbe[1]}", f"{(mbe[0]-mbe[1]):.0f}"])
    cT.append([f"{nbe[0]}", f"{nbe[1]}", f"{(nbe[0]-nbe[1])}"])
    cT.append([f"{mse[0]:.0f}", f"{mse[1]:.0f}", f"{(mse[0]-mse[1]):.0f}"])
    cT.append([f"{rmse[0]:.0f}", f"{rmse[1]:.0f}", f"{(rmse[0]-rmse[1]):.0f}"])
    cT.append([f"{nrmse[0]}", f"{nrmse[1]}", f"{(nrmse[0]-nrmse[1])}"])
    cT.append([f"{r[0]}", f"{r[1]}", f"{(r[0]-r[1])}"])
    cT.append([f"{r2[0]}", f"{r2[1]}", f"{(r2[0]-r2[1])}"])
    cT.append([f"{r2a[0]}", f"{r2a[1]}", f"{(r2a[0]-r2a[1])}"])
    cT.append([f"{aic[0]:.0f}", f"{aic[1]:.0f}", f"$p$: {AICp(aic[0]-aic[1])}"])
    cT.append([f"{bic[0]:.0f}", f"{bic[1]:.0f}", f"{bic[0] - bic[1]:.0f}"])
    return cT

```

```

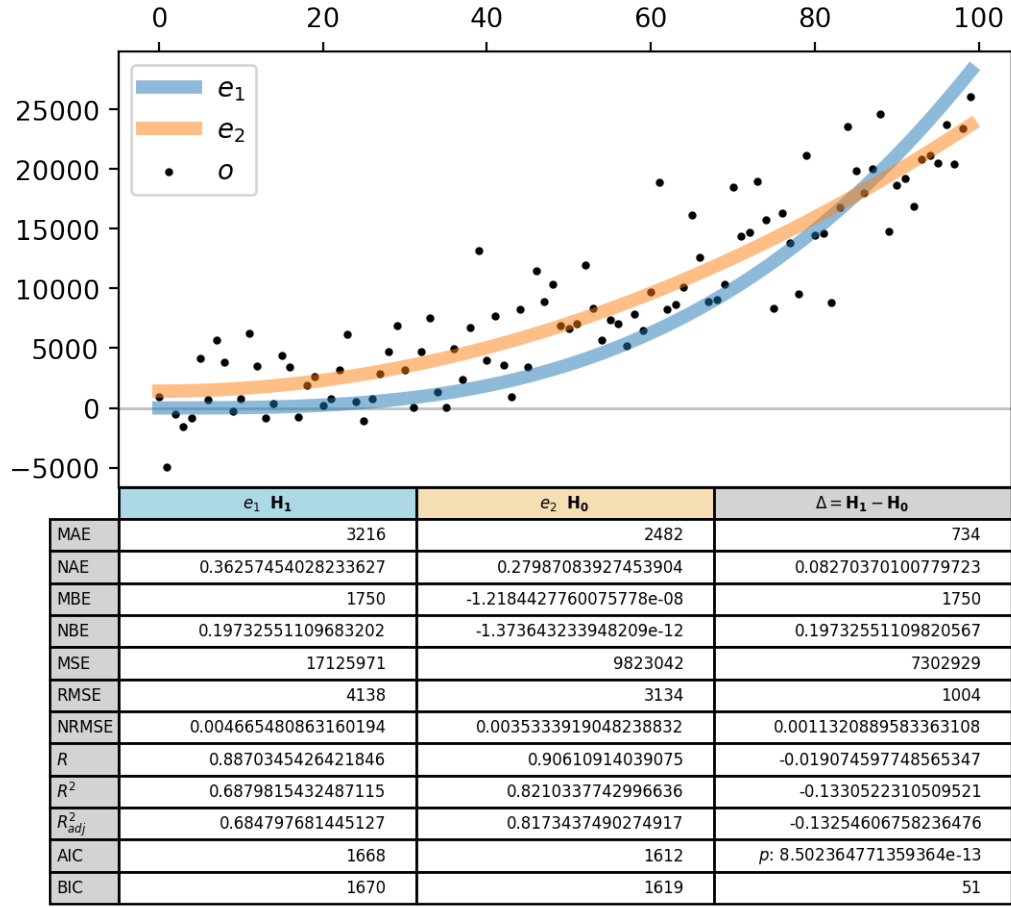
x = np.arange(0, 100)
o = func_noise(x, 2.3, 1e3, err=True)
popt1, _ = curve_fit(func1, x, o)
e1 = func1(x, *popt1)
popt2, _ = curve_fit(func2, x, o)
e2 = func2(x, *popt2)

cellText = report(o, [e1, e2], [1, 2])

fig, ax = plt.subplots(figsize=(6, 3), dpi=200)
ax.xaxis.tick_top()
ax.axhline(0, lw=1, alpha=.25, c="k")
ax.plot(x, e1, label="$e_1$", lw=5, alpha=.5)
ax.plot(x, e2, label="$e_2$", lw=5, alpha=.5)
ax.scatter(x, o, label="$o$", s=4, c="k")
ax.legend(loc="best")
ax.table(
    cellText=cellText,
    colLabels=[
        "$e_1$  $\mathbf{H_1}$", "$e_2$  $\mathbf{H_0}$",
        "$\Delta = \mathbf{H_1} - \mathbf{H_0}$"
    ],
    colColours=["lightblue", "wheat", "lightgrey"],
    rowColours=["lightgrey" for _ in range(12)],
    rowLabels=[
        "MAE", "NAE", "MBE", "NBE", "MSE", "RMSE", "NRMSE",
        "$R$", "$R^2$", "$R^2_{adj}$", "AIC", "BIC"
    ],
)
plt.show()

```

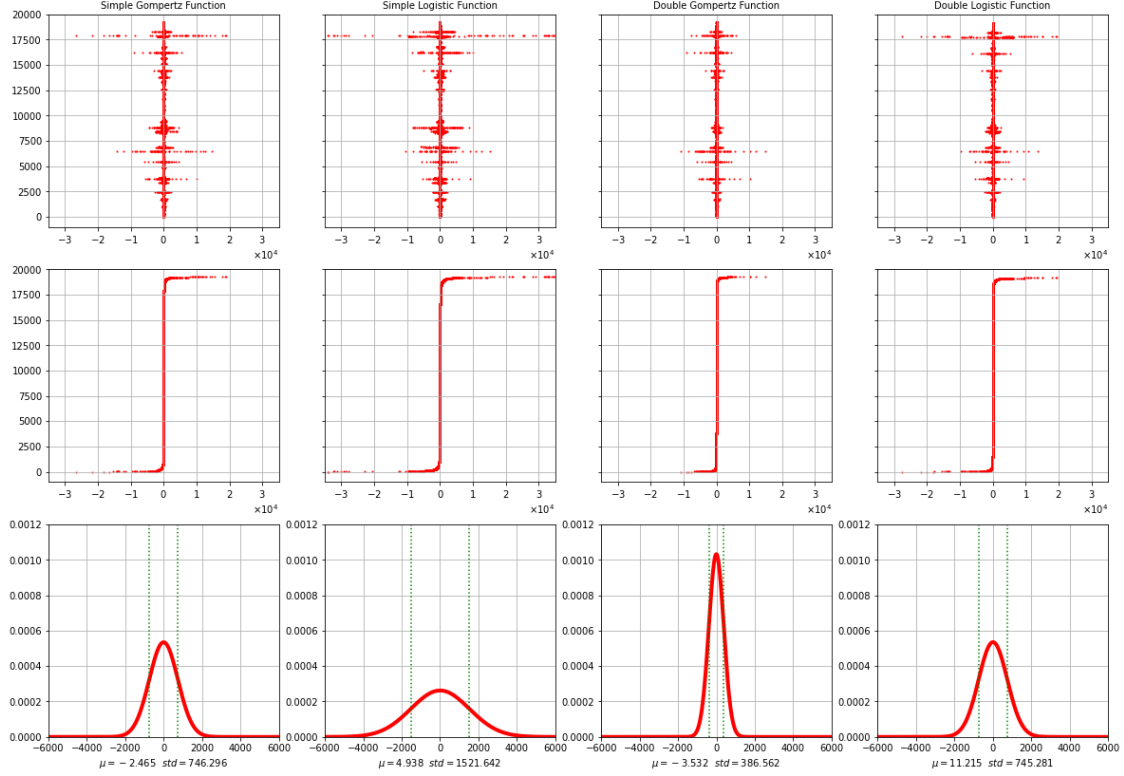
Fitting and analysis report (figure [REF]).



4.2 Residual

Total residual from each model have been collected and compared to get a first “rough” evidence of the most likely better fitting model [FIG].

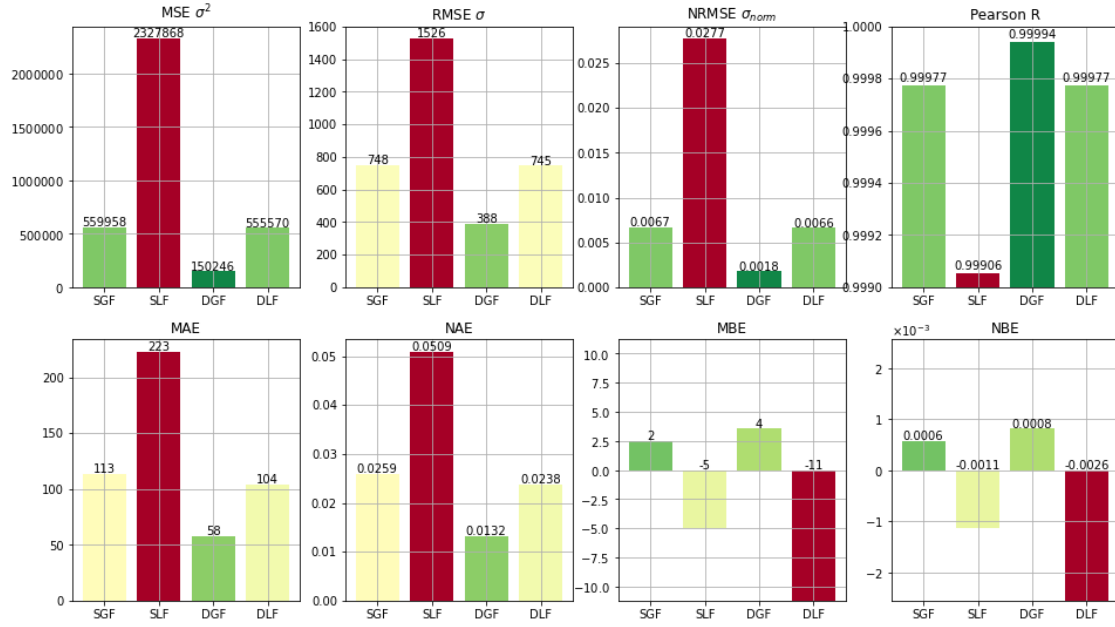
Models Residual



DGF showed the lower residual standard deviation while the mean of all four models has been found very close to 0.

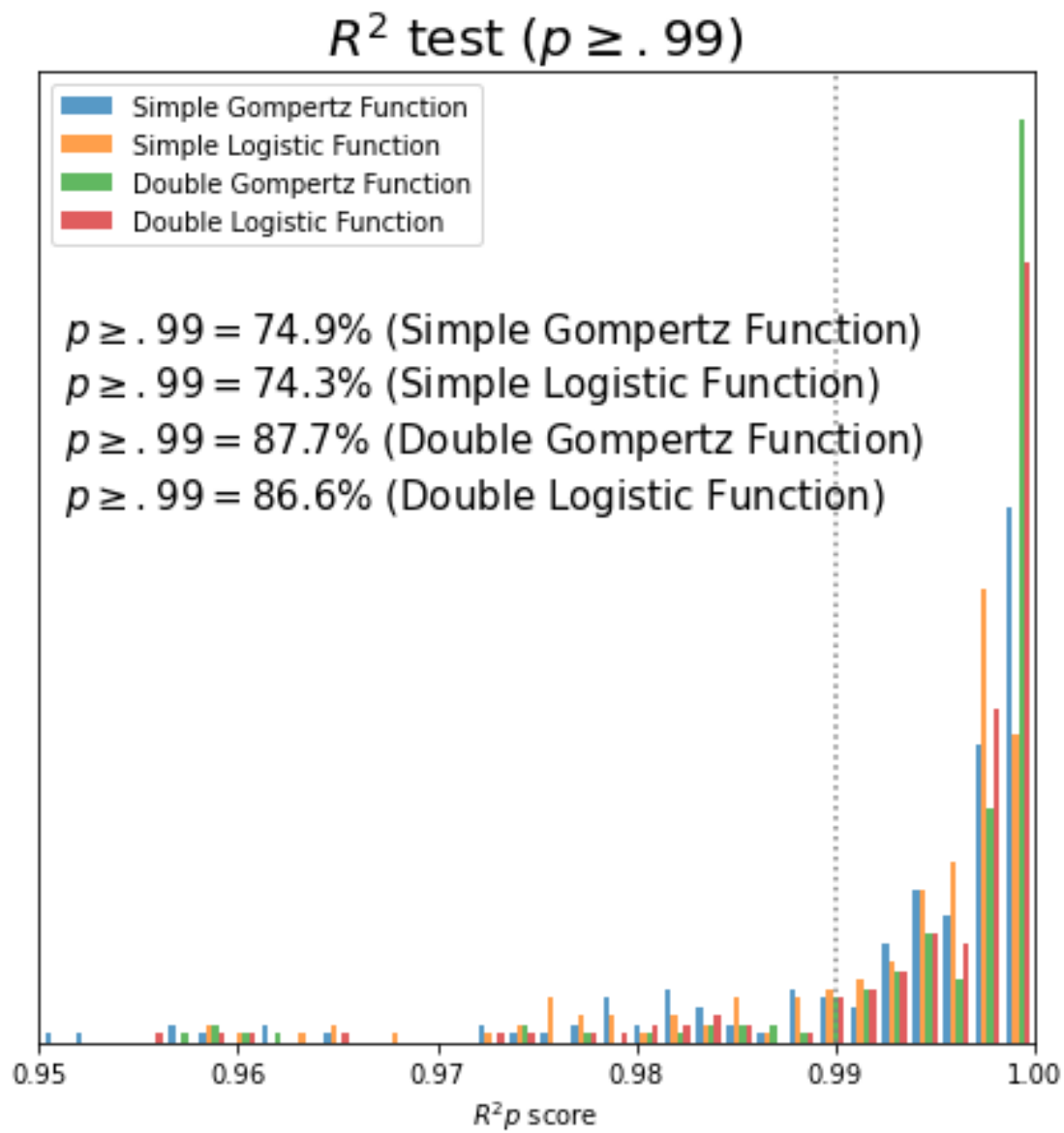
Mean Squared Error (**MSE**, Variance), Root Mean Squared Error (**RMSE**, Standard Deviation), Normalize Root Mean Squared Error (**NRMSE**, Normalized Standard Deviation), Mean Absolute Error (**MAE**), Normalized Absolute Error (**NAE**), Mean Bias Error (**MBE**), Normalized Bias Error (**NBE**) and Pearson Correlation Coefficient (**Pearson R**) have been computed for all models residual Section 4.3 (see Appendix for formulae Section 7). **DGF** showed the best results for all values, but MBE and NBE, confirming the first null hypothesis that could have been the best fitting model among the chosen ones.

4.3 Errors tests

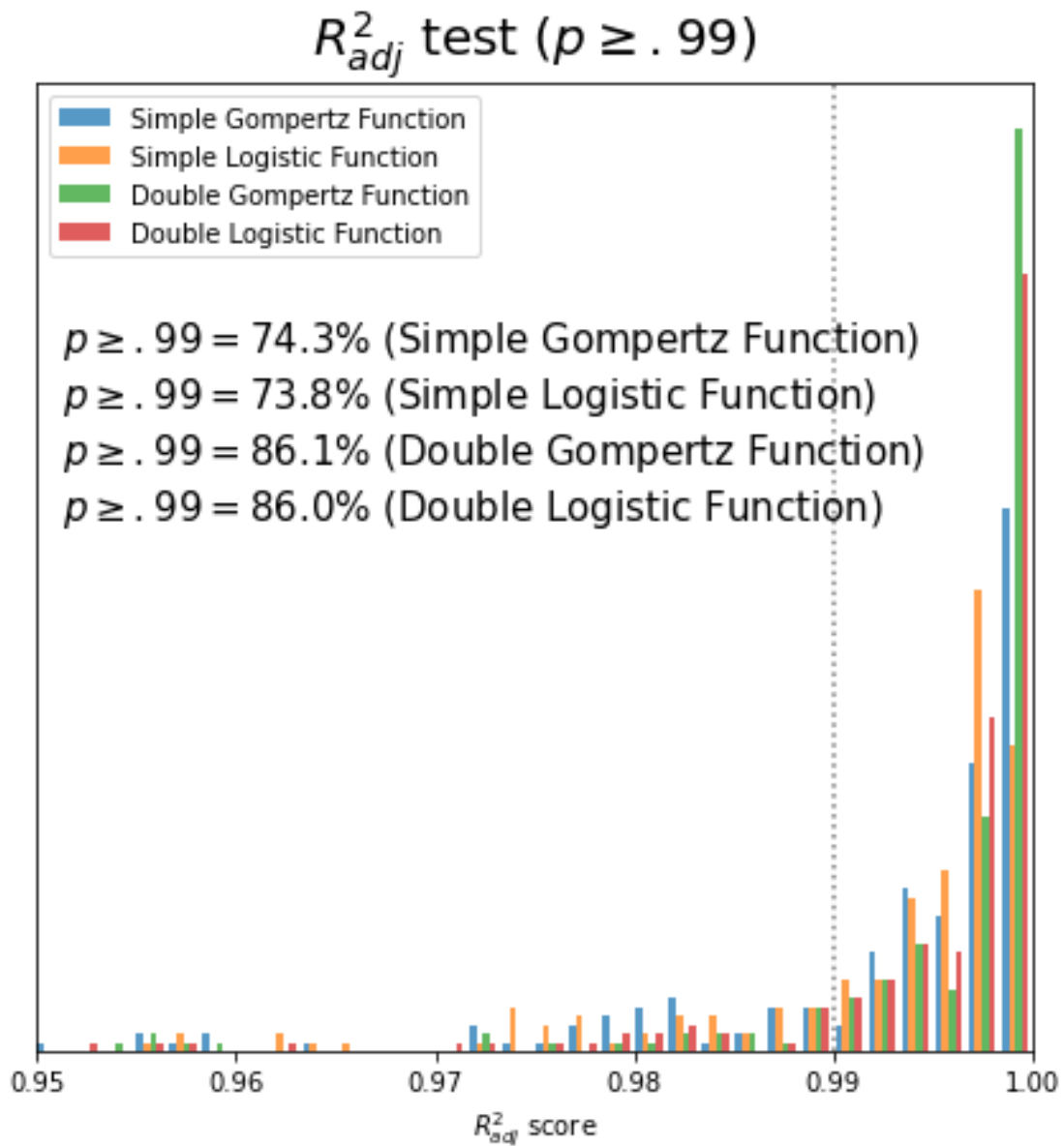


Coefficient of Determination R^2 , Adjusted Coefficient of Determination R^2_{adj} , Akaike Information Criterion (**AIC**) and Bayesian Information Criterion (**BIC**) have computed and collected from all fits and compared with each other.

4.4 R²

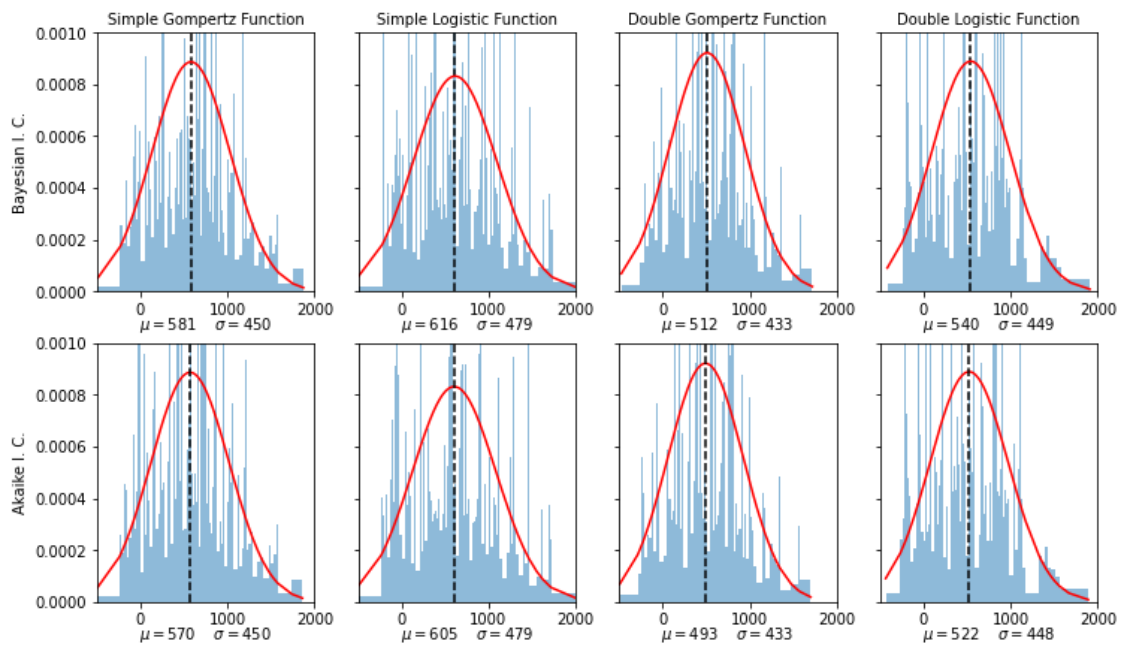


4.5 Adjusted R2



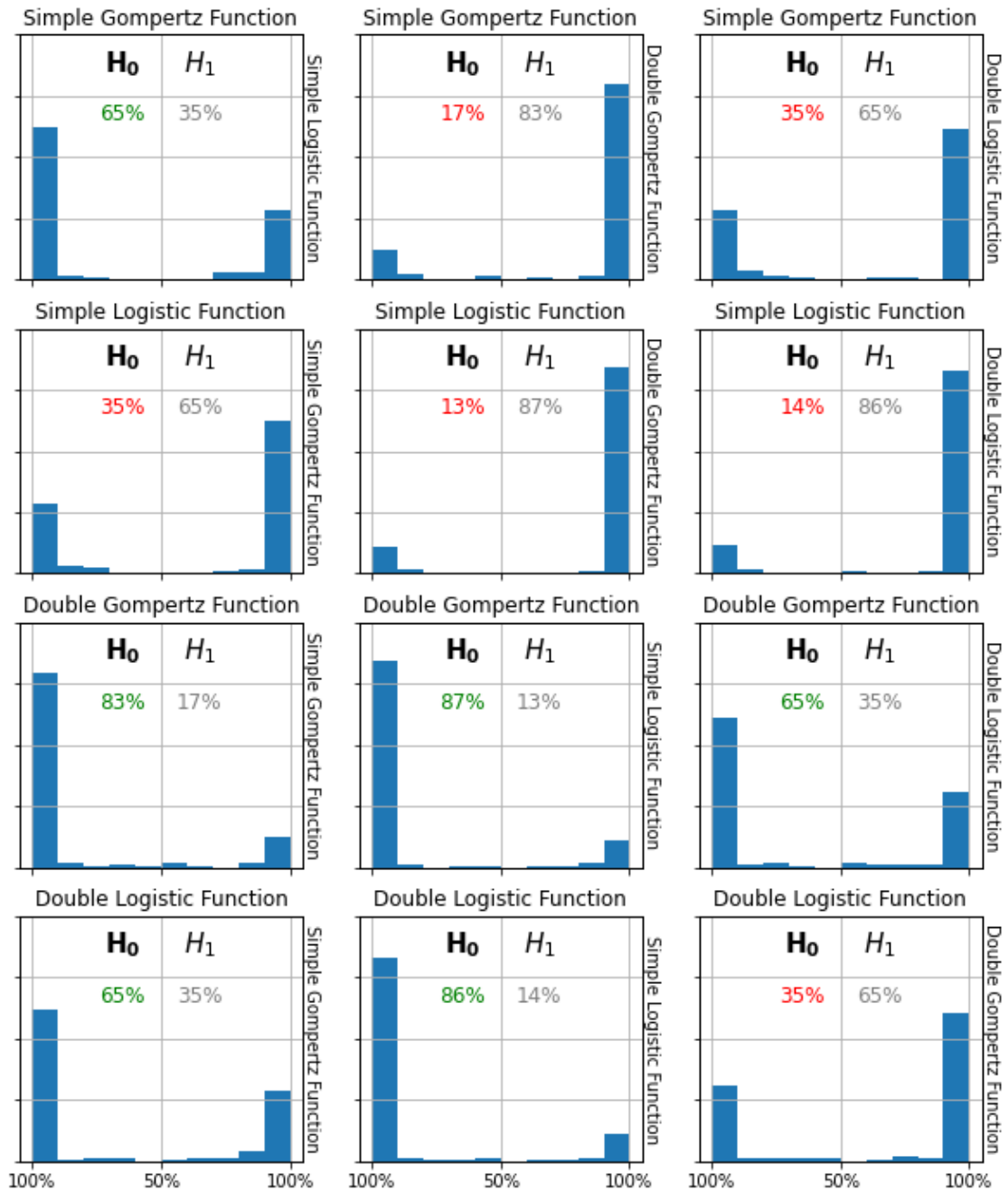
4.6 Information Criteria

Models Statistical Tests



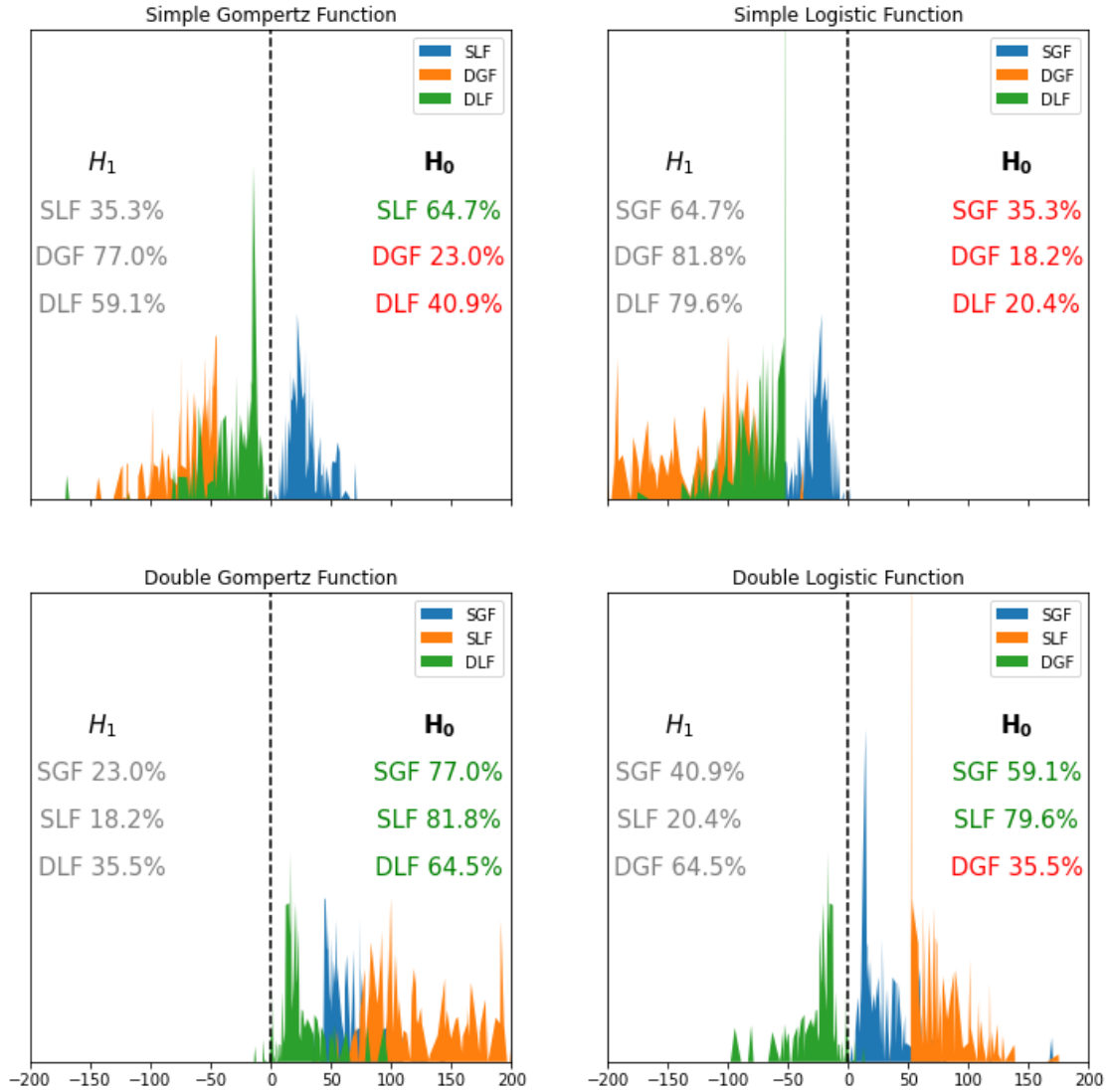
4.6.1 AIC

ΔAIC p density



4.6.2 BIC

ΔBIC test



Double Gompertz Function showed $\Delta\text{BIC} < 0$ for:

- **43 countries** compared with *Simple Gompertz Function*. 24 have less than 1000 cases: Botswana, Burundi, Central African Republic, Chad, Comoros, Congo (Kinshasa), Gabon, Guatemala, Guinea-Bissau, Liberia, Liechtenstein, MS Zaandam, Saint Kitts and Nevis, Sao Tome and Principe, Sierra Leone, Somalia, South Sudan, Suriname, Tajikistan, Tanzania, Timor-Leste, Trinidad and Tobago, Western Sahara, Yemen. **19 have more than 1000 cases:** Bangladesh, Bolivia, China, Cote d'Ivoire, Croatia, Ecuador, Finland, Ghana, Kazakhstan, Lithuania, Malaysia, Nigeria, Norway, Oman, Peru, Russia, Saudi Arabia, Spain, Ukraine.
- **34 countries** compared with *Simple Logistic Function*. 23 have less than 1000 cases: Antigua and Barbuda, Burundi, Central African Republic, Chad, Comoros, Congo (Brazzaville),

H_0	MAE	NAE	MBE	NBE	MSE	RMSE	NRMSE	R	$p > .99$ R^2	$p > .99$ R^2_{adj}	$p < .5$ ΔAIC	$\Delta > 0$ ΔBIC	H_1
SGF	113	0.025888	2.49	0.000568	559958	748	0.006672	0.999773	74.87%	74.33%	64.71%	64.71%	SLF
											16.58%	22.99%	DGF
											35.48%	40.86%	DLF
SLF	223	0.050858	-4.95	-0.001130	2327868	1526	0.027737	0.999056	74.33%	73.80%	35.29%	35.29%	SGF
											13.37%	18.18%	DGF
											13.98%	20.43%	DLF
DGF	58	0.013196	3.56	0.000812	150246	388	0.001790	0.999939	87.70%	86.10%	83.42%	77.01%	SGF
											86.63%	81.82%	SLF
											64.52%	64.52%	DLF
DLF	104	0.023764	-11.21	-0.002559	555570	745	0.006620	0.999775	86.56%	86.02%	64.52%	59.14%	SGF
											86.02%	79.57%	SLF
											35.48%	35.48%	DGF

Diamond Princess, Fiji, Gabon, Guatemala, Guinea-Bissau, MS Zaandam, Maldives, Mauritania, Nepal, Papua New Guinea, Sao Tome and Principe, Suriname, Tajikistan, Tanzania, Timor-Leste, Vietnam, Yemen. **11 have more than 1000 cases:** Bolivia, China, Djibouti, Ecuador, France, Ghana, Iceland, Nigeria, Oman, Qatar, Saudi Arabia.

- **66 countries** compared with *Double Logistic Function*. 39 have less than 1000 cases: Albania, Antigua and Barbuda, Bahamas, Belize, Bhutan, Cabo Verde, Cambodia, Central African Republic, Chad, Congo (Brazzaville), Diamond Princess, Equatorial Guinea, Eswatini, Fiji, Georgia, Guatemala, Guinea-Bissau, Kosovo, Laos, Liechtenstein, MS Zaandam, Monaco, Montenegro, Mozambique, Namibia, Nepal, Niger, Papua New Guinea, Paraguay, Saint Vincent and the Grenadines, Sao Tome and Principe, Sudan, Tajikistan, Tanzania, Togo, Vietnam, West Bank and Gaza, Western Sahara, Yemen. **27 have more than 1000 cases:** Belarus, Belgium, Bolivia, China, Djibouti, Ecuador, France, Greece, Honduras, Iceland, Indonesia, Iraq, Israel, Korea, South, Morocco, Nigeria, Norway, Oman, Peru, Qatar, Saudi Arabia, Singapore, Spain, Sweden, Tunisia, Ukraine, Uzbekistan.

4.7 Results

All tests strongly confirmed *Double Gompertz Function* as the better fitting model for SARS-nCoV-2 cumulative cases curve fitting. Results also showed that **DGF** is not only much more fitting than models with less parameters (**SLF** and **SGF**) as expected but also compared to *Double Logistic Function* with the same degrees of freedom.

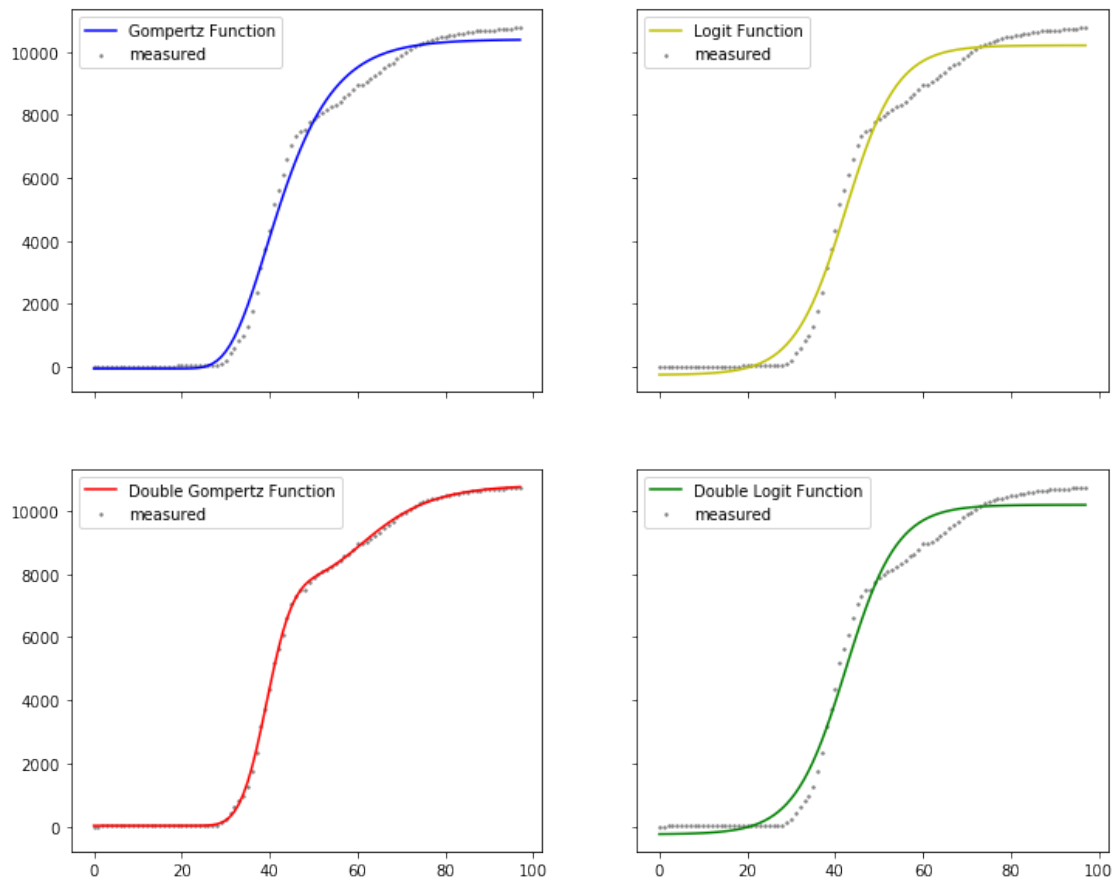
5 Conclusions

Among the compared models *Double Gompertz Function* has showed the best results and scores fitting data of SARS-nCoV-2 cumulative cases, suggesting that this model should be studied more deeply (possibly improved) and compared to other existing models for further analysis, including forecasting capabilities.

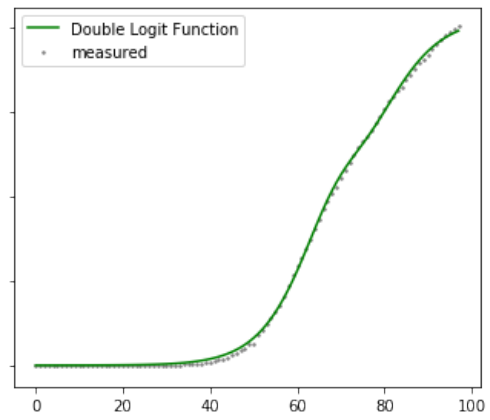
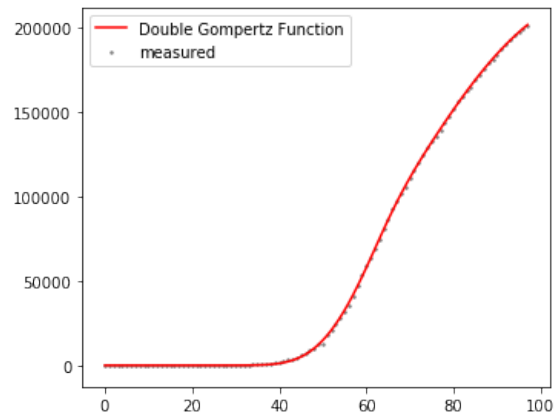
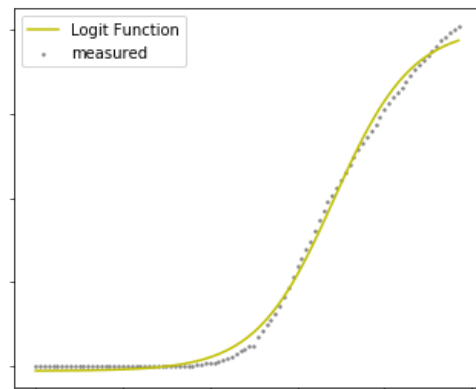
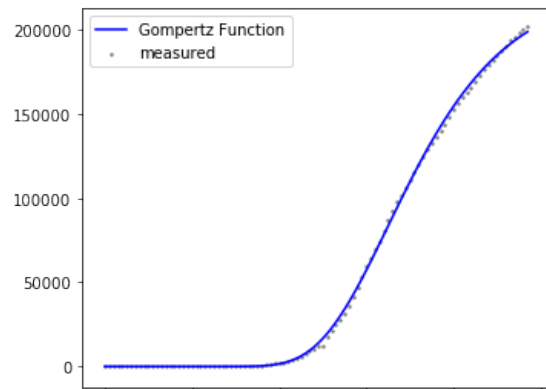
6 Plots

6.1 Fit examples

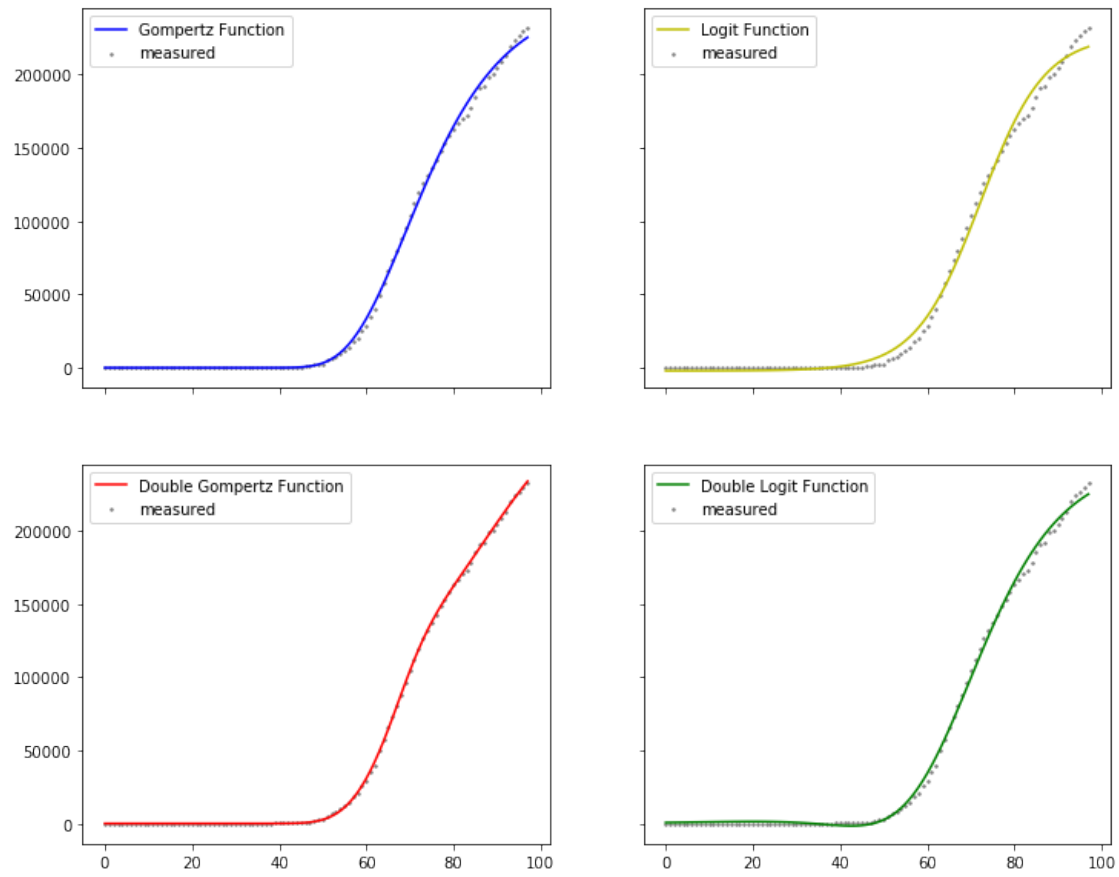
Korea, South



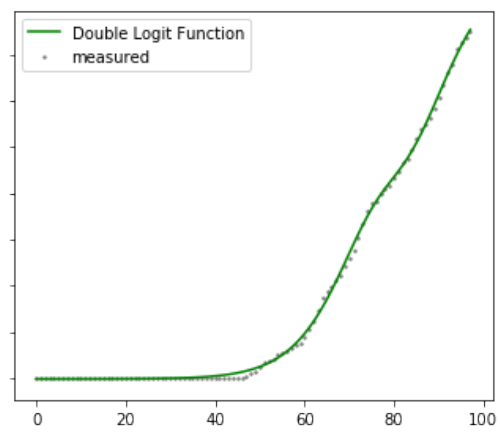
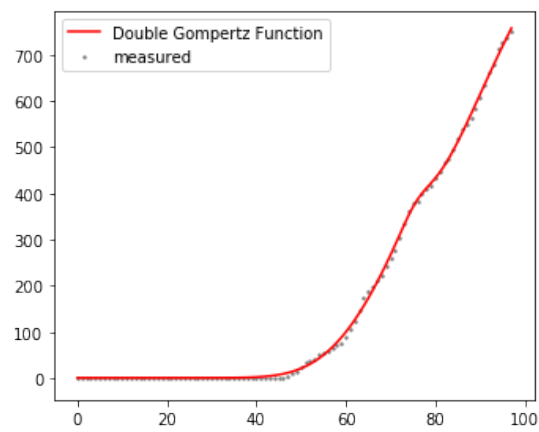
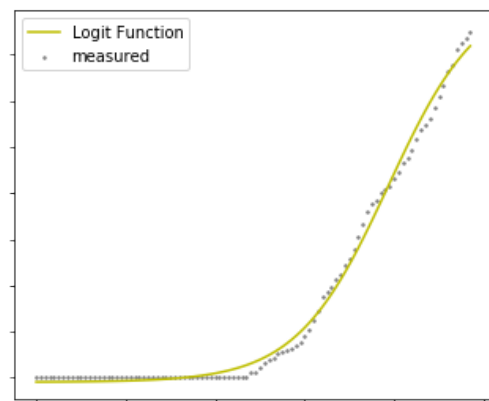
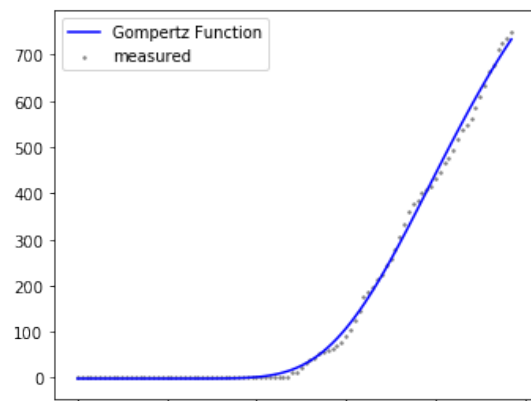
Italy



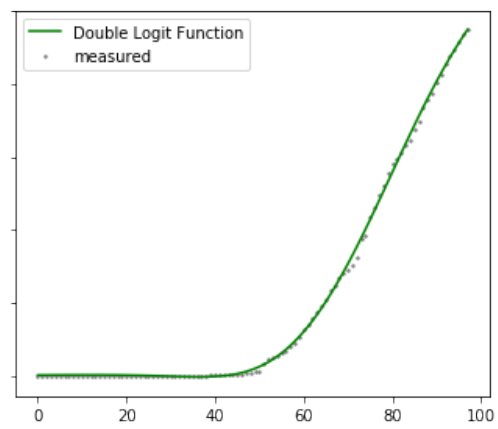
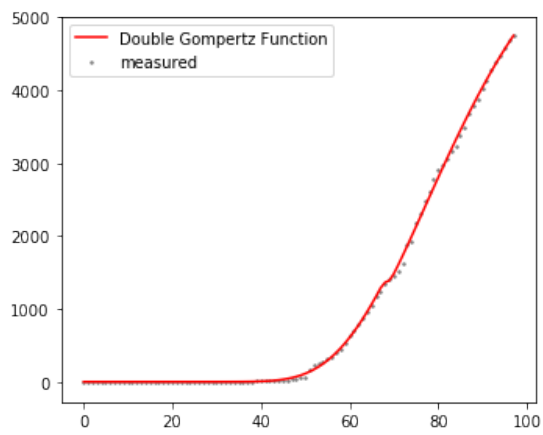
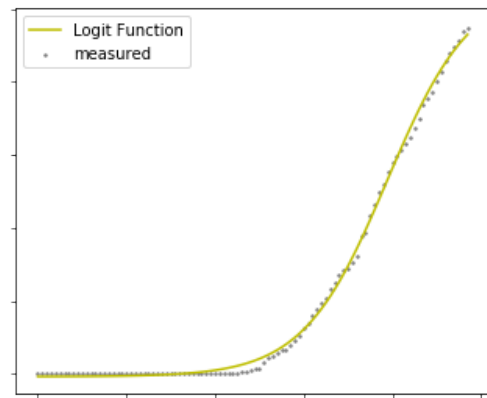
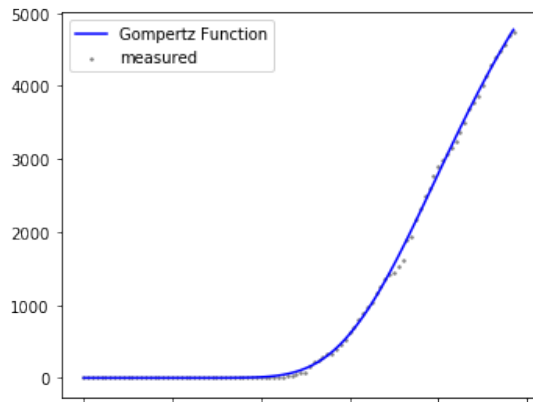
Spain



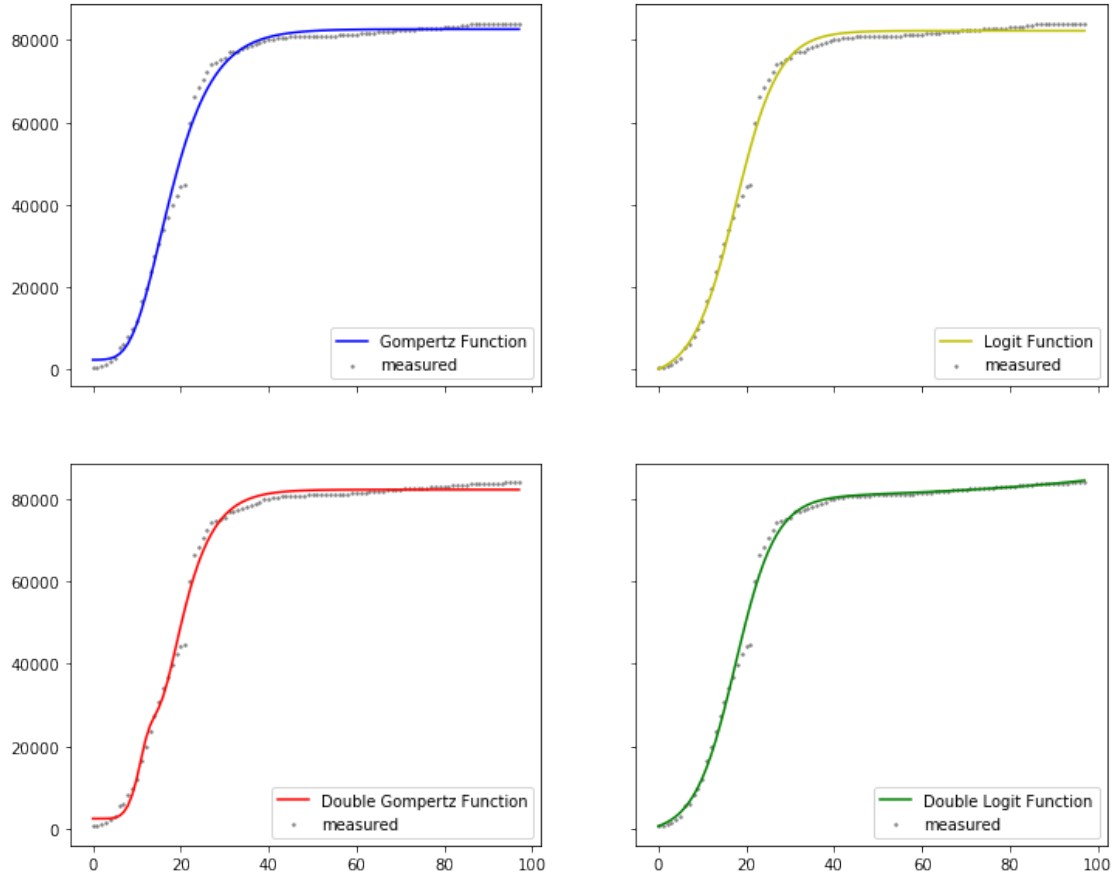
Albania



Finland



China



7 Appendix

7.1 Statistics formulae

In all formulae we assume: y as the observed (measured) values and \hat{y} as expected (predicted by fitting model) values; n is the number of values and n_{var} the number of model's variable parameters.

- **MBE:** Mean Bias Error

$$\mathbf{MBE} = \frac{\sum (y - \hat{y})}{n}$$

- **MAE:** Mean Absolute Error (aka Mean Deviation)

$$\mathbf{MAE} = \frac{\sum |y - \hat{y}|}{n}$$

- **NBE**: Normalized Bias Error

$$\mathbf{NBE} = \frac{\sum (y - \hat{y})}{\sum y}$$

- **NAE**: Normalized Absolute Error (aka Normalized Mean Deviation)

$$\mathbf{NAE} = \frac{\sum |y - \hat{y}|}{\sum y}$$

- **RSS**: Residual Sum of Squares

$$\mathbf{RSS} = \sum (y - \hat{y})^2$$

- **TSS**: Total Sum of Squares

$$\mathbf{TSS} = \sum \left(y - \frac{\sum y}{n} \right)^2$$

- **MSE**: Mean Squared Error (aka Variance)

$$\mathbf{MSE} = \sigma^2 = \frac{\mathbf{RSS}}{n}$$

- **RMSE**: Root Mean Squared Error (aka Standard Deviation)

$$\mathbf{RMSE} = \sigma = \sqrt{\mathbf{MSE}}$$

- **NRMSE**: Normalized Root Mean Squared Error (aka Normalized Standard Deviation)

$$\mathbf{NRMSE} = \sigma_\nu = \frac{\mathbf{RMSE}}{\sum y}$$

- R^2 : Coefficient of Determination:

$$R^2 = 1 - \frac{\mathbf{RSS}}{\mathbf{TSS}}$$

- R_{adj}^2 : Adjusted Coefficient of Determination:

$$R_{adj}^2 = 1 - \frac{\mathbf{RSS}}{\mathbf{TSS}} \left(\frac{n-1}{n-n_{var}-1} \right)$$

- **Pearson R**: Pearson Correlation Coefficient

$$\mathbf{R} = \frac{\sum y\hat{y} - \frac{1}{n} \sum y \sum \hat{y}}{\sqrt{\sum y^2 - \frac{1}{n} (\sum y)^2} \sqrt{\sum \hat{y}^2 - \frac{1}{n} (\sum \hat{y})^2}}$$

- **AIC**: Aikake Information Criterion

$$\mathbf{AIC} = n \ln \left(\frac{\mathbf{RSS}}{n} \right) + 2n_{var}$$

- **AIC p**: Aikake Information Criterion score (or weight) in **AIC** relative probability distribution space [FIG]:

$$\mathbf{AIC}_p = \frac{e^{-0.5 \cdot (\mathbf{AIC}_1 - \mathbf{AIC}_0)}}{1 + e^{-0.5 \cdot (\mathbf{AIC}_1 - \mathbf{AIC}_0)}}$$

$$\mathbf{AIC}_p < .5 \Rightarrow P(H_0) > P(H_1)$$

- **BIC**: Bayesian Information Criterion

$$\mathbf{BIC} = n \ln \left(\frac{\mathbf{RSS}}{n} \right) + \ln(n)n_{var}$$

- $\Delta\mathbf{BIC}$: Bayesian Information Criterion difference [FIG]:

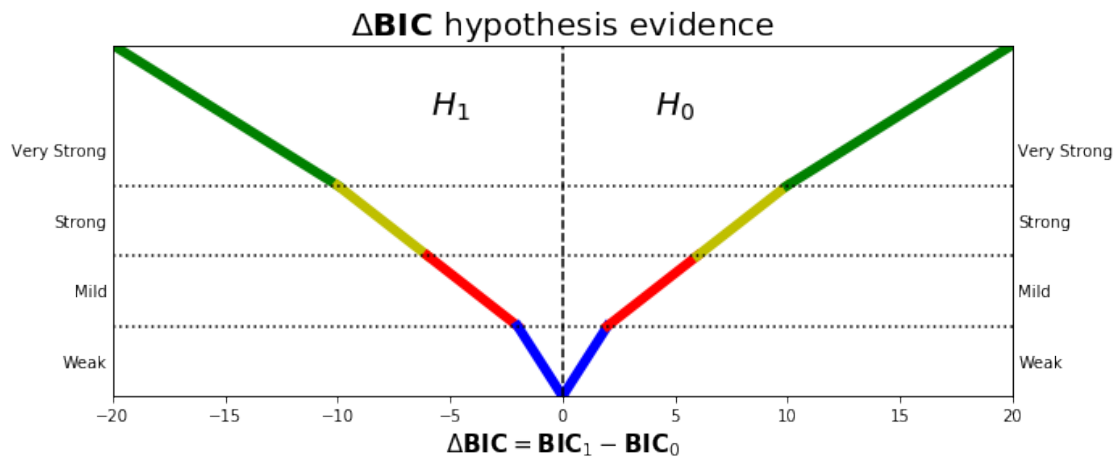
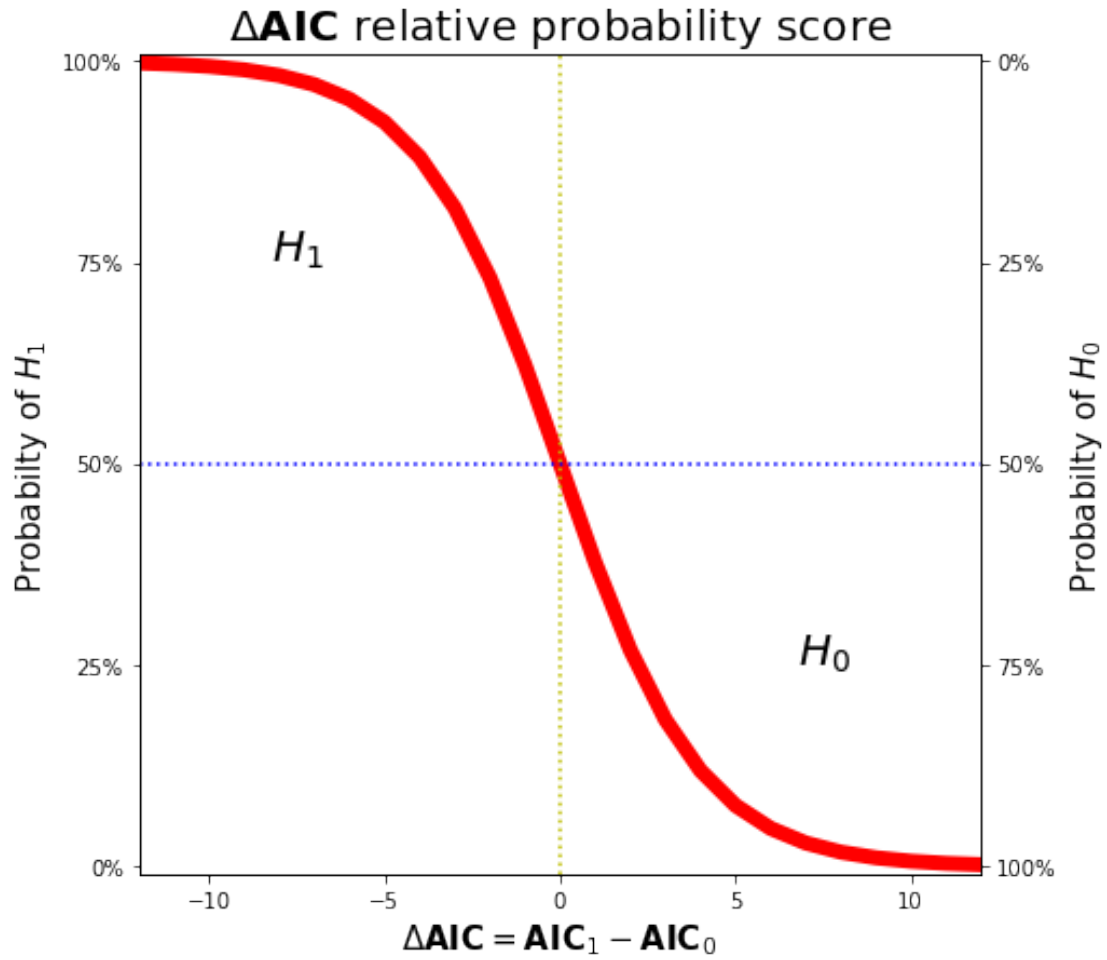
$$\Delta\mathbf{BIC} = \mathbf{BIC}_1 - \mathbf{BIC}_0$$

$$\Delta\mathbf{BIC} > 10 \Rightarrow H_0 \text{ Very strong evidence}$$

$$\Delta\mathbf{BIC} = (6, 10] \Rightarrow H_0 \text{ Strong evidence}$$

$$\Delta\mathbf{BIC} = (2, 6] \Rightarrow H_0 \text{ Mild evidence}$$

$$\Delta\mathbf{BIC} = (0, 2] \Rightarrow H_0 \text{ Weak evidence}$$



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