

Climate change increases bilateral trade cost

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Abstract

It is well established that climate change affects economic production, but its effects on trade networks, especially trade costs, have not been studied. I use international trade and weather data covering almost 200 years to show that climate change increases trade costs. Estimating a simple augmented gravity framework, I find that rising temperatures at the origin or destination country increase bilateral trade cost. I use a standard trade model to quantify the welfare impact of increased trade cost, finding that the impact of climate change on trade cost over the preceding 100 years reduced welfare in the 2010s by 0.82 percent. This effect depends not only on countries' own climate trends, but importantly on the climate trends of countries they export to and import from. Looking at the distribution of gains, poor and rich countries are equally harmed by trade cost increases due to climate change. Smaller economies, which are more reliant on international trade, are especially affected. My methodology can easily be embedded in studies of the impact of climate change using models of international trade.

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Existing analyses of the effect of climate change on trade usually take trade networks — especially trade costs — as given and focus on the effect on output (Costinot, Donaldson, & Smith, 2016; Nath, 2020). Trade costs, however, are determined by the same economic forces as production activities, for example worker productivity and the availability of labor and capital. It makes sense to assume, therefore, that climate change would affect trade cost as well as output.

I show that over the last 170 years, climate change has pushed the world apart: Rising temperatures increase bilateral trade cost. I estimate a standard gravity framework with one addition, an interaction between distance and decadal averages of temperature at the origin and destination countries. I find a significant, negative impact of climate change on trade cost. I show that the results are robust to various specifications of the effect of distance on bilateral trade.

I embed these estimates in a standard model of international trade (Eaton & Kortum, 2002) to quantify the welfare impacts and find that welfare in the 2010s would have been 0.82 percent higher if climate change had not increased trade over the preceding 100 years, purely due to the resulting reduction in trade costs. Analyzing the welfare gains, I show that this effect depends not only on countries' own climate trends, but importantly on the climate trends of countries they export to and import from. Poor and rich countries benefit equally. The benefits are especially large for smaller economies, which are more reliant on international trade.

Since I only rely on an augmented gravity specification, the effect of climate change on trade cost I demonstrate in this paper can easily be included in estimations of the future impact of climate change using trade models. This is especially true for the broad class of trade models that allow for gravity estimation to be solved separately from the rest of the model.

The remainder of the paper proceeds as follows: Section 1 discusses the data I use and presents descriptive statistics, Section 2 describes the gravity equation framework I use for my reduced form estimation, Section 3 presents results of the reduced form estimation, Section 4 estimates the welfare impacts of trade cost increases due to climate change, and Section 5 concludes.

1 Data and descriptive statistics

I use data on trade flows from the CEPII TRADHIST database of historical trade data (Fouquin & Hugot, 2016). The data cover yearly international bilateral trade flows from 1827 until 2014 and contain additional information necessary for estimating gravity equations. I combine these with Berkeley Earth data on the yearly mean of daily maximum temperatures (Rohde, Muller, Jacobsen,

Muller, Perlmutter, Rosenfeld, Wurtele, Groom, & Wickham, 2013). The temperature data go as far back as 1750 for some areas, achieve significant global coverage starting in 1850 and full global coverage beginning in 1960. I have weather data for more than 90 percent of all countries in the trade data beginning in the 1880s and for all countries in the trade data beginning in the 1910s.

Figure 1 shows average temperature in °C across decades, plus a 95 percent confidence interval. The point at which I have weather data on all countries is indicated in the figure with a vertical line. Changes before that point in time can reflect additions to the sample as well as actual temperature changes. Over time, temperature rises from around 23°C in the 1910s to almost 25°C in the 2010s, with an especially fast change beginning in the 1980s. To underscore the increased speed of warming in recent decades, Figure 2 shows averages of decade-on-decade changes for each country in the sample, again with a 95 percent confidence interval. The average country has seen a significant increase in temperatures for most decades since the 1910s (where I have data for all countries in my sample), except for a brief period of a small decrease or no change from the 1950s to the 1970s. Recent decades' average changes of up to 0.3°C exceed past changes considerably, certainly for the period where I have weather data for all countries in my sample.

2 Gravity estimation framework

Augmented by a time dimension, gravity equations describe trade flows X_{nit} between an origin i and destination n at time t as (Head & Mayer, 2015)

$$X_{nit} = G_t S_{it} M_{nt} \phi_{nit}$$

where S_{it} and M_{nt} are exporter and importer capabilities, also called multilateral resistance terms (Anderson & van Wincoop, 2003), and ϕ_{nit} is a measure of trade cost between the two countries, called a bilateral resistance term.

While different models yield different interpretations of what the multi- and bilateral resistance terms reflect, for the purposes of estimating a gravity equation, the bilateral resistance term is usually modeled as

$$\phi_{nit} = d_{ni}^\alpha e^{\mathbf{C}'_{nit}\beta}$$

with d_{ni} a measure of physical distance between the two countries and \mathbf{C}_{nit} a collection of bilateral

variables that affect trade between the two countries, such as contiguity or colonial history. The elasticity of trade flows with respect to distance α could capture preferences (Anderson & van Wincoop, 2003) or country (Eaton & Kortum, 2002) or firm productivity dispersion (Melitz, 2003). I augment this basic specification by allowing the effect of distance to vary by temperature,

$$\phi_{nit} = d_{ni}^{\alpha + \delta_1 T_{it} + \delta_2 T_{nt}} e^{\mathbf{C}'_{nit} \boldsymbol{\beta}} \quad (1)$$

T_{ct} is a measure of temperature in country c during period t , which is fully interacted with distance. To estimate this, I use origin-period and destination-period fixed effects to model the multilateral resistance terms (Anderson & van Wincoop, 2003). Accordingly, I drop the level effects of T_{ct} which are captured by those fixed effects. Note that this specification could be applied to any trade model that yields a gravity equation, so my estimation results apply to any model in this large class. Dividing by the origin and destination GDPs to obtain normalized trade flows $\pi_{nit} \equiv X_{nit} / (Y_{it} Y_{nt})$ and letting $\tilde{d}_{ni} \equiv \log(d_{ni})$, this yields an estimating equation

$$\begin{aligned} \pi_{nit} &= e^{\gamma_{it} + \xi_{nt} + \log(\phi_{nit})} \\ &= \exp \left\{ \gamma_{it} + \xi_{nt} + \alpha \tilde{d}_{ni} + \delta_1 \tilde{d}_{ni} T_{it} + \delta_2 \tilde{d}_{ni} T_{nt} + \mathbf{C}'_{nit} \boldsymbol{\beta} \right\} \end{aligned} \quad (2)$$

with origin-period and destination-period fixed effects γ_{it} and ξ_{nt} . To deal with instances where trade flows are equal to zero, rather than using a log-linear estimator, this is commonly estimated in its exponentiated form using pseudo-Poisson maximum likelihood estimation (PPML) (Santos Silva & Tenreyro, 2006), which I follow here.

Because I deal with temperature changes over long time horizons, I estimate this model across several periods, each comprising multiple years, rather than using yearly data. In my baseline specification, I use each decade from 1830 to 2020 as a period t . I calculate decadal averages of all variables for each origin-destination pair to estimate the model. Using averages is especially attractive if π_{nit} is interpreted as a (noisy) measure of a true underlying trade share, since decadal averages are closer to the true underlying value than yearly data.

3 Gravity estimation results

Table 1 shows the results of estimating different versions of this gravity equation via PPML, using the R command `hdfepml` (Garrucho, Zylkin, & Cruz, 2023). I use the population-weighted great

circle distance between the origin and destination countries in kilometers to capture d_{ni} . Instead of log distance, I use the de-meaned version $\tilde{d}_{ni}^{\text{dm}} \equiv \log(d_{ni}) - \log(\bar{d}_{ni})$ to center interaction terms at the mean distance. (Note that this does not change the estimated coefficient on distance or its interpretation, it simply makes it easier to interpret estimates for interactions, since those now reflect the effect size when all variables involved are at their respective means.) As temperature measures, I use the yearly mean of daily maximum temperature in °C. I convert these to z -scores \mathcal{T}_{ct} to facilitate the interpretation of effect sizes and to center interactions at mean temperatures. \mathbf{C}_{nit} contains a common language indicator, contiguity indicator and indicators for current and past colonial relationships, taking decadal means for all variables. Standard errors are clustered by country pair with p -values shown in brackets.

The first column shows results for the basic model (2). The second column shows results for a robustness check also interacting the variables included in \mathbf{C}_{nit} with both temperature measures. The third column contains a robustness check which lets α vary over time by interacting \tilde{d}_{ni} with decade indicators; Figure 3 shows the estimated α_t across decades. The fourth column shows a model using an alternative trade flow normalization based solely on the destination GDP, $\tilde{\pi}_{nit} \equiv X_{nit}/Y_{nt}$. The fifth column shows a baseline model without temperature variables. Figure 4 shows coefficient estimates for δ_1 and δ_2 from a model with time-varying coefficients on the temperature variables (that is, δ_1 and δ_2 vary across decades, by interacting their respective variables with decade indicators). Figure 5 shows the corresponding coefficients on distance.

I consistently find a negative effect of distance on trade flows, with a magnitude roughly comparable to the estimates from Santos Silva and Tenreyro (2006). I also consistently find that temperature at the origin and destination decrease trade flows between the two countries. My baseline specification yields that, at the mean origin and destination temperatures, a one percent increase in distance decreases trade flows by 1.155 percent, with an additional 0.083 percent decrease for each one standard deviation increase in temperature at the origin and an additional 0.106 percent decrease for each one standard deviation temperature increase at the destination. Looking at the time-varying effect estimates in Figure 4, in decades with universal data coverage (beginning in the 1910s), temperature at both the origin and especially the destination tends to have a negative and statistically significant impact on trade flows.

A potential mechanism for this effect is that shipping and receiving goods is an industrial task much like many other. It involves both manual and cognitive labor. It is well established that weather shocks and climate change affect the efficiency of both kinds of labor and of industrial

firms more generally (Adhvaryu, Kala, & Nyshadham, 2019; Carleton & Hsiang, 2016; Somanathan, Somanathan, Sudarshan, & Tewari, 2021; Zhang, Dêschenes, Meng, & Zhang, 2018). Through the same mechanisms that climate change affects efficiency in manufacturing firms, it could also affect the efficiency of dock and freight workers and those overseeing freight and port operations.

The components of \mathbf{C}_{nit} interacted with temperatures do not yield a consistent pattern. Being in a current colonial relationship has a lower impact on trade flows when either temperature rises, while shared language and ever having colonial ties have a larger impact when temperature at the destination increases. Since the elements of \mathbf{C}_{nit} are relevant only for a subset of trade relationships, while all countries have to contend with distance, it makes sense that these impacts are noisier. There is also a priori no strong reason to suspect that climate change would interact with any one of the elements of \mathbf{C}_{nit} in a particular way.

4 Welfare impacts

The gravity estimation results allow me to estimate the change in ϕ_{nit} resulting from a change to the climate of a different period $s \neq t$, by plugging temperatures for that period T_{is} into the specification for the bilateral resistance term (1) to obtain a counterfactual ϕ'_{nit} . The change in the bilateral resistance term is

$$\hat{\phi}_{nit} \equiv \frac{\phi'_{nit}}{\phi_{nit}} \stackrel{(1)}{=} d_{ni}^{\delta_1(T_{is}-T_{it})+\delta_2(T_{ns}-T_{nt})}$$

noting that all covariates remain constant — I simply want to estimate the change in bilateral resistance stemming from the changed weather variables.

To go from this to the implied welfare impacts, however, I need to specify a model of international trade, to discipline how wages and prices would adjust under this counterfactual. I use the well-established model of Eaton and Kortum (2002) combined with the exact hat algebra of Dekle, Eaton, and Kortum (2008) to estimate the welfare change that would occur if the 2010s had instead had the climate of other decades in my data. Under this model, the bilateral resistance term is equal to

$$\phi_{nit} = d_{nit}^{-\theta}$$

where $\theta > 0$ measures productivity dispersion in the Fréchet distribution of technology underlying the Eaton and Kortum (2002) model. As Dekle et al. (2008) show, the counterfactual trade shares

π'_{nit} resulting from a change $\hat{d}_{nit} \equiv d'_{nit}/d_{nit}$ (keeping technology fixed) are

$$\pi'_{nit} = \frac{\pi_{nit} (\hat{d}_{nit} \hat{w}_{nit})^{-\theta}}{\sum_{k=1}^N \pi_{nkt} (\hat{d}_{nkt} \hat{w}_{nkt})^{-\theta}}$$

Further, they show that the resulting welfare change is

$$\hat{W}_{it} \equiv \frac{W'_{it}}{W_{it}} = \left(\frac{\pi'_{iit}}{\pi_{iit}} \right)^{-\frac{1}{\theta}}$$

which is simply the change in own trade share raised to a negative power — if the own trade share decreases, welfare increases. It is straightforward to solve the resulting system for wage changes \hat{w}_{it} that equate counterfactual trade deficits and surpluses with those observed in the data, yielding counterfactual trade shares π'_{nit} which enable me to estimate the welfare change for each country.¹ Following Dekle et al. (2008), I set the only unknown parameter $\theta = 8.28$.

I use the 2010s as my reference period and estimate welfare changes resulting from a shift to each previous decade's climate. I do this for all previous decades in my data for which I have weather observations covering at least 90 percent of the countries in my sample. Figure 7 shows the mean welfare change across decades, as well as the 5th and 95th percentile of welfare changes. (Appendix Table 3 shows the same information in table form.) The impact tends to be larger when switching to earlier climates, since temperatures are increasing over time and higher temperatures increase trade cost. For example, the mean increase for the earliest decade, the 1880s, is estimated to be 0.93 percent, whereas for the 1950s I estimate an average welfare increase of 0.65 percent and for the most recent decade, the 2000s, I estimate an 0.08 percent welfare increase, on average. Across all decades prior to the 2000s, all countries see an increase in welfare, and the 5th percentile of welfare changes is consistently positive. At the 95th percentile, welfare impacts are as high as 1.74 percent in the 1880s counterfactual.²

A core correlate of welfare changes are of course climate trends. Figure 9, Figure 10 and Figure 11 show welfare changes in the 1910s counterfactual across countries' own temperature change between period t and the 2010s, as well as their export and import network temperature changes, respectively. The export network temperature change for country i and period t is a weighted average of temperature changes for all countries i exports to, weighted by their 2010s share

¹ This also requires choosing a normalization; I fix world GDP at its 2010s value.

² Appendix Figure 19 and Appendix Table 4 show versions of these results using population-weighted averages based on countries' 2010s population. Results are very similar.

in country i 's exports:

$$\text{Export network change}_{it} \equiv \frac{1}{\sum_{k \neq i} \pi_{ki, 2010s}} \sum_{k \neq i} \pi_{ki, 2010s} \Delta T_{kt}$$

where ΔT_{nt} is country n 's change in temperature between period t and the 2010s. The import network change is defined analogously. All three measures of climate trends are strongly correlated with welfare gains, which makes sense — countries which see larger temperature increases over the last 100 years also see larger trade cost increases. Reversing those larger trade cost increases leads to larger welfare gains.

These temperature measures are, of course, correlated. Figure 12 highlights this, showing export and import network temperature changes across the change in own temperature between the 1910s and 2010s. Both network changes in temperature are positively correlated with countries' own change in temperature, and export network changes especially so. To understand whether all three temperature measure have some independent predictive power for welfare gains, Table 2 shows results from regression of welfare impacts \hat{W}_{it} across periods. These regressions include period fixed effects to analyze correlates of welfare change within period. Standard errors are clustered at the country level. The first column highlights that all three temperature measures are correlated with welfare changes, even conditional on the other two. The change in import network temperature is especially strongly correlated with welfare changes.

This is in part because, as already seen in Figure 12, import network temperature changes are somewhat less correlated with own temperature changes. The reason is that import networks tend to span a larger distance, as shown in Figure 13. This figure shows a histogram of 2010s import network distance minus export network distance. The import network distance for country i is defined as the distance to all countries i trades with, weighted by their import shares,

$$\text{Import network distance}_i \equiv \frac{1}{\sum_{k \neq i} \pi_{ki, 2010s}} \sum_{k \neq i} \pi_{ki, 2010s} d_{ki}$$

The export network distance is defined analogously. As the figure shows, most countries import from a larger network than they export to.³ This leads to the weaker correlation between own

³ This makes sense in the context of the model of Eaton and Kortum (2002), for example, since countries export goods for which they are especially efficient producers. Each country faces many other competitors for each good, and each of those competitors has a chance of being especially productive at producing any given good. So country i 's chance to be efficient enough to beat out all of those competitors is relatively small for any given good. Country i will therefore be efficient enough to export only for a small number of goods.

temperature change and import network temperature change. Consequently, especially import network temperature changes contain some independent predictive power, even conditional on own temperature changes. If nothing else, this highlights the global nature of climate change, and the limitation of drawing inferences about its impact from purely country-specific information.

To understand the distribution of gains across countries, Figure 14 shows the estimated welfare impacts of returning to the climate of the 1910s across countries' 2010s log GDP. Larger economies tend to benefit less from reversing the impact of climate change on trade cost. As Figure 8 shows, however, welfare gains are not correlated with GDP per capita.⁴ That is, rich and poor countries alike are roughly equally affected by the trade cost impacts of climate change.

To understand why larger economies benefit less from trade cost reductions, the second column of Table 2 shows a regression of welfare gains on log 2010s GDP, highlighting that across periods, GDP and welfare gains are strongly correlated. The third column adds controls for countries' own temperature change between period t and the 2010s, as well as their export and import network temperature changes. These results highlight that the correlation between welfare gains and GDP is not due to the fact that larger economies face different climate trends, since the coefficient on 2010s log GDP remains very similar. This is also underscored by Figure 15, Figure 16 and Figure 17, which show own temperature change, export network change and import network change in the 1910s counterfactual across 2010s log GDP. All correlations are relatively weak, though larger economies tend to see a somewhat larger increase in import network temperature.

As the fourth column of Table 2 shows, the explanation is straightforward. That regression controls for countries' 2010s own trade share. As Figure 18 highlights, larger economies tend to have higher own trade shares — they have larger domestic markets, and are less reliant on international trade. As soon as that control is added to the regression, smaller economies no longer see larger welfare gains. The reason smaller economies benefit is that they are more reliant on international trade. Reversing trade cost increases from climate change is therefore especially valuable for smaller economies.

⁴ To highlight that the distribution of gains is relatively similar across decades, though of course with varying levels over time, Appendix Figures 20 and 21 show the distribution of welfare impacts across log GDP per capita for the 1950s and 1980s counterfactuals, respectively.

5 Conclusion

I show that climate change pushes countries further apart by increasing the cost of trade. In an augmented gravity estimation, rising decade-level average temperatures at the origin or destination country increase bilateral trade cost, an effect that is robust across various specifications. The welfare impacts of this are considerable. Using the Eaton and Kortum (2002) model, I find that average welfare during the 2010s would have been 0.82 percent higher if climate change had not increased trade cost over the preceding 100 years. This effect depends not only on countries' own climate trends, but importantly on the climate trends of countries they export to and import from. Poor and rich countries benefit equally. The benefits are especially large for smaller economies, which are more reliant on international trade.

Since I only rely on an augmented gravity specification, the effect of climate change on trade cost I demonstrate in this paper can easily be included in estimations of the future impact of climate change using trade models. This is especially true for the broad class of trade models that allow for gravity estimation to be solved separately from the rest of the model. I hope this will enrich our analysis of the impact of climate change.

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Tables

Table 1: Gravity estimation results

| Variable | Basic model | Full interaction | Distance \times decade | $\tilde{\pi}_{nit}$ | Benchmark |
|---|-------------------|-------------------|--------------------------|---------------------|-------------------|
| $\tilde{d}_{ni}^{\text{dm}}$ | -1.155 [0.000] | -1.147 [0.000] | | -0.999 [0.000] | -1.192 [0.000] |
| $\tilde{d}_{ni}^{\text{dm}} \mathcal{T}_{it}$ | -0.083 [0.031] | -0.088 [0.012] | -0.067 [0.043] | -0.111 [0.000] | |
| $\tilde{d}_{ni}^{\text{dm}} \mathcal{T}_{nt}$ | -0.106 [0.001] | -0.097 [0.004] | -0.089 [0.004] | -0.033 [0.079] | |
| Language $_{ni} \times \mathcal{T}_{it}$ | | 0.091 [0.220] | | | |
| Language $_{ni} \times \mathcal{T}_{nt}$ | | 0.152 [0.028] | | | |
| Contiguous $_{ni} \times \mathcal{T}_{it}$ | | -0.115 [0.458] | | | |
| Contiguous $_{ni} \times \mathcal{T}_{nt}$ | | 0.040 [0.742] | | | |
| Current colony $_{nit} \times \mathcal{T}_{it}$ | | -0.592 [0.002] | | | |
| Current colony $_{nit} \times \mathcal{T}_{nt}$ | | -0.891 [0.000] | | | |
| Ever colony $_{ni} \times \mathcal{T}_{it}$ | | -0.004 [0.980] | | | |
| Ever colony $_{ni} \times \mathcal{T}_{nt}$ | | 0.257 [0.036] | | | |
| $\tilde{d}_{ni}^{\text{dm}} \times \text{decade}$ | No | No | Yes | No | No |
| \mathbf{C}_{nit} | Yes | Yes | Yes | Yes | Yes |
| Origin-decade FE | Yes | Yes | Yes | Yes | Yes |
| Destination-decade FE | Yes | Yes | Yes | Yes | Yes |

Note: The outcome are trade flows from i to n normalized by dividing by origin and destination GDPs, $\pi_{nit} \equiv X_{nit}/(Y_{it}Y_{nt})$. d_{ni} is the population-weighted great circle distance between the origin and destination countries in km. I subtract the log of the mean distance to center interaction terms at the mean distance, $\tilde{d}_{ni}^{\text{dm}} \equiv \log(d_{ni}) - \log(\bar{d}_{ni})$. (Note that this does not change the estimated coefficient on distance or its interpretation, it simply makes it easier to interpret estimates for interactions, since those now reflect the effect size when all variables involved are at their respective means.) \mathcal{T}_c is the z -score of the yearly mean of daily maximum temperatures in country c in $^{\circ}\text{C}$. \mathbf{C}_{nit} contains a common language indicator, contiguity indicator and two indicators for current and past colonial relationships, taking decade means for all variables within each origin-destination pair. Decades t are the decades from 1830 to 2020. *Distance \times decade* allows the effect of distance to vary over time by interacting distance with decade indicators. $\tilde{\pi}_{nit}$ uses an alternative trade flow normalization based solely on destination GDP, $\tilde{\pi}_{nit} \equiv X_{nit}/Y_{nt}$. Outcomes for this specification are winsorized at the 99th percentile, due to a handful of outliers. *Benchmark* omits the temperature variables. Standard errors clustered by country pair, p -values in brackets.

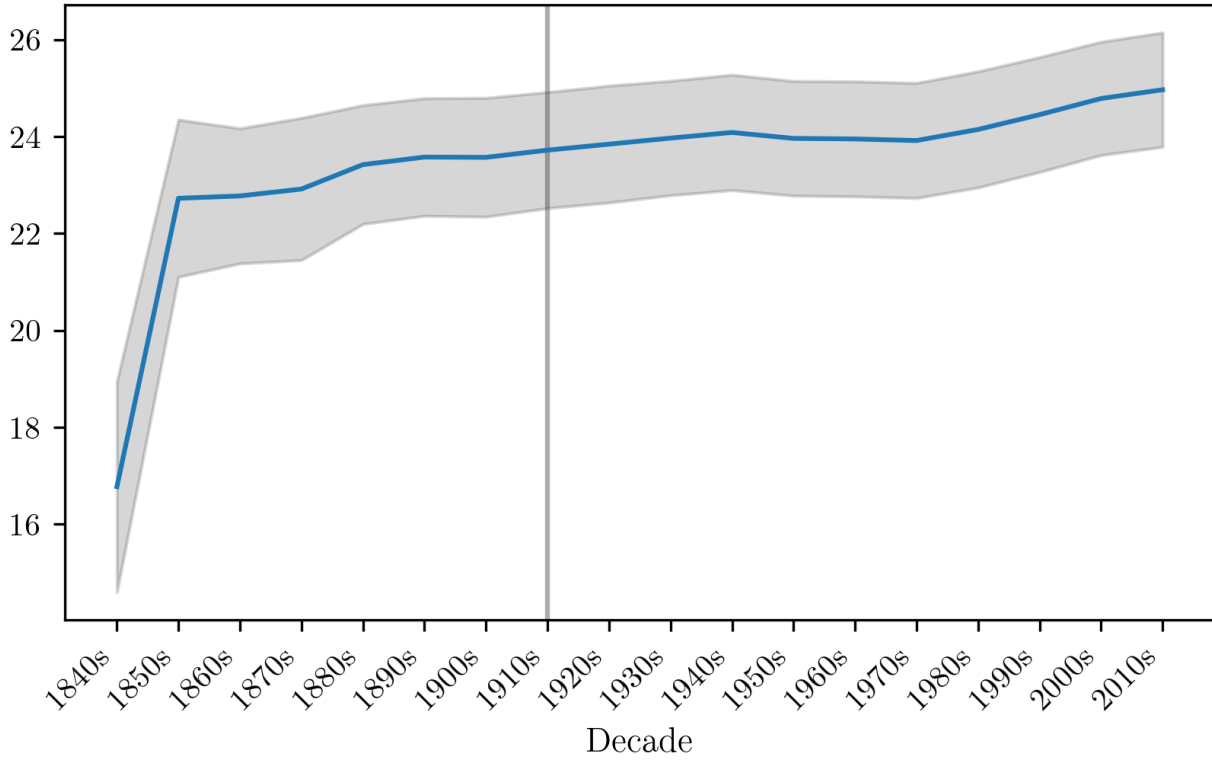
Table 2: Correlates of welfare changes

| Variable | \hat{W}_{it} | \hat{W}_{it} | \hat{W}_{it} | \hat{W}_{it} |
|---------------------------|------------------|-------------------|-------------------|-------------------|
| Log 2010s GDP | | −0.025 [0.023] | −0.033 [0.001] | 0.001 [0.794] |
| Own change | 0.238 [0.000] | | 0.224 [0.000] | |
| Export network change | 0.202 [0.099] | | 0.257 [0.060] | |
| Import network change | 0.757 [0.000] | | 0.788 [0.000] | |
| 2010s own trade share (%) | | | | −0.022 [0.000] |
| Decade FE | Yes | Yes | Yes | Yes |

Note: The outcome \hat{W}_{it} is the welfare change for country i under decade t 's climate counterfactual. *Own change* is each country's own change in temperature between each decade and the 2010s, whereas the two *network change* variables are the average change in temperature between each decade and the 2010s across each country's 2010s trade network. Standard errors clustered by country, p -values in brackets.

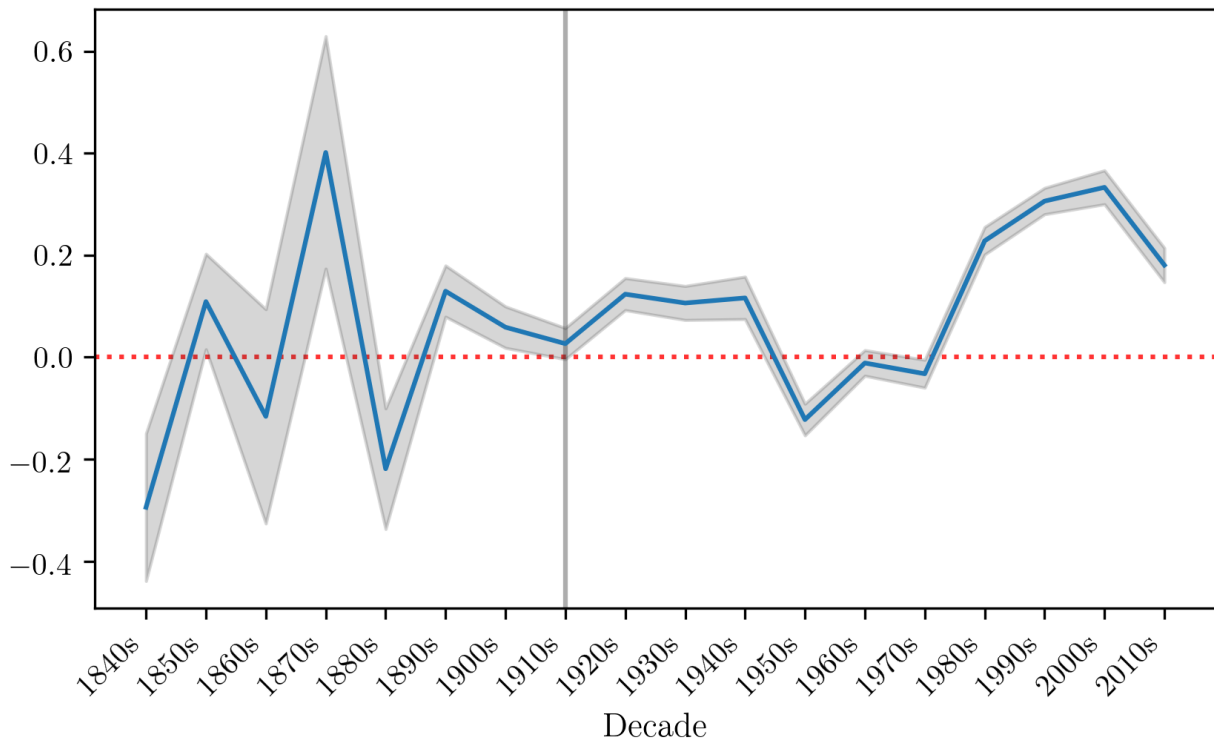
Figures

Figure 1: Average temperature ($^{\circ}\text{C}$) across decades



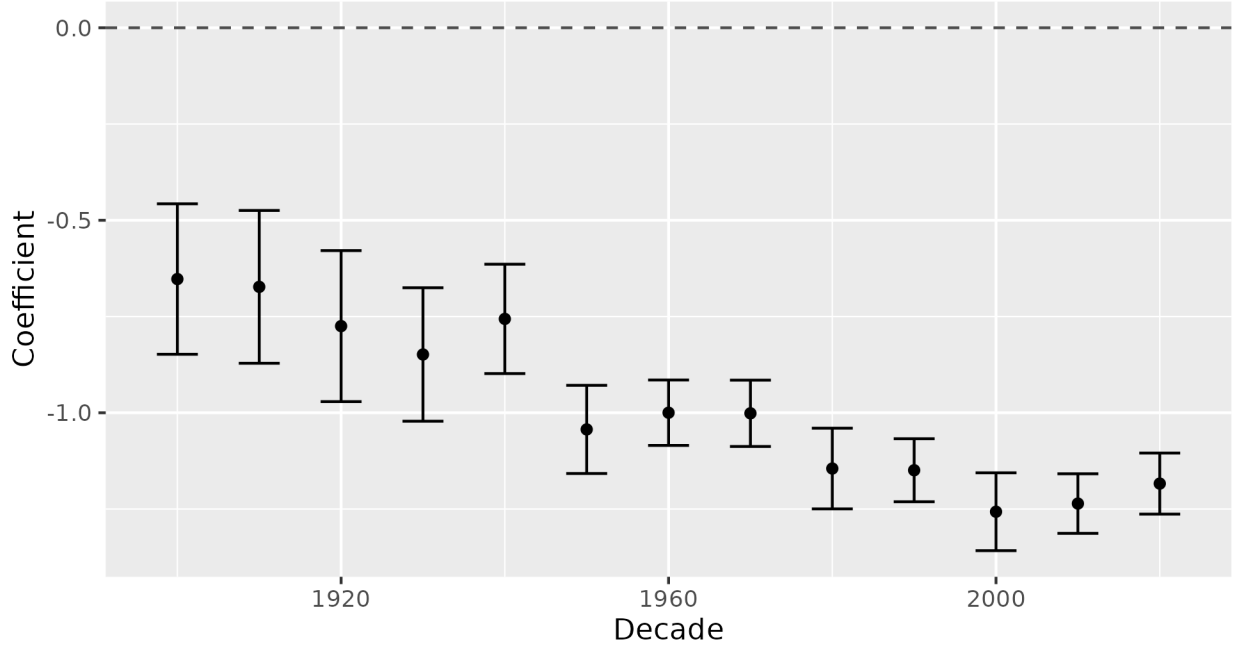
Note: The figure shows the average temperature for each decade. Gray bands show 95 percent confidence intervals. The solid vertical line indicates the first decade for which all countries have non-missing weather data. Results for decades prior to this reflect sample selection as well as changes over time.

Figure 2: Average temperature change ($^{\circ}\text{C}$) across decades



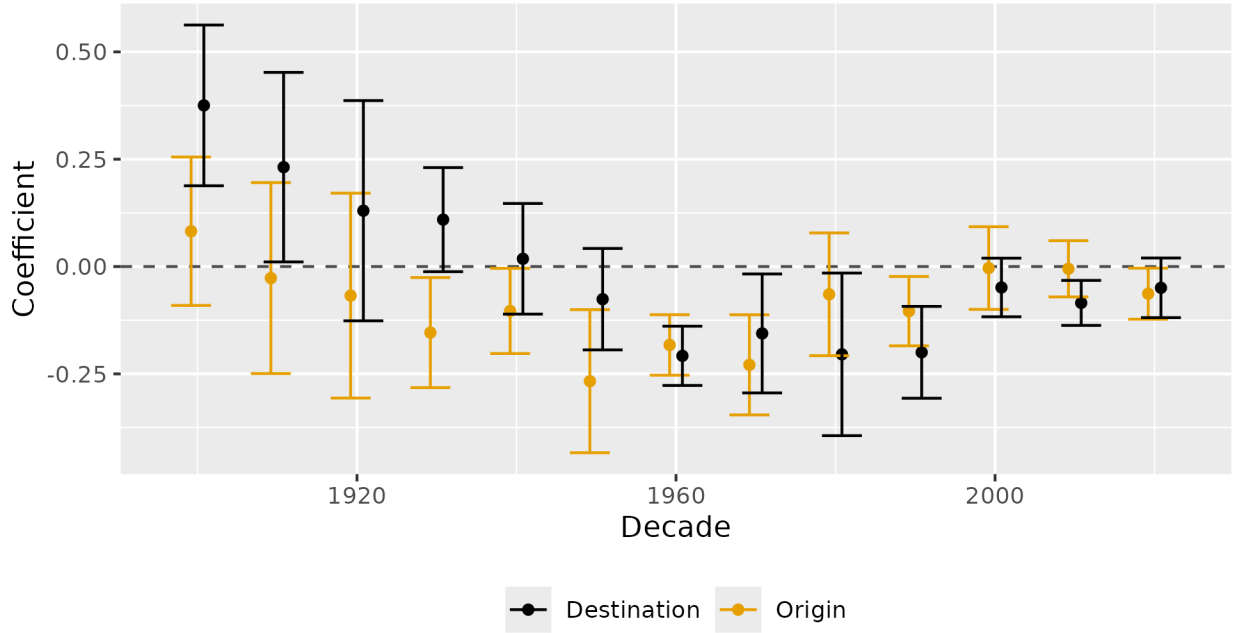
Note: The figure shows the average temperature change compared to the previous decade. Gray bands show 95 percent confidence intervals. The solid vertical line indicates the first decade for which all countries have non-missing weather data. Results for decades prior to this reflect sample selection as well as changes over time. The dotted horizontal line indicates no change.

Figure 3: Coefficients on log distance across decades (only distance varies)



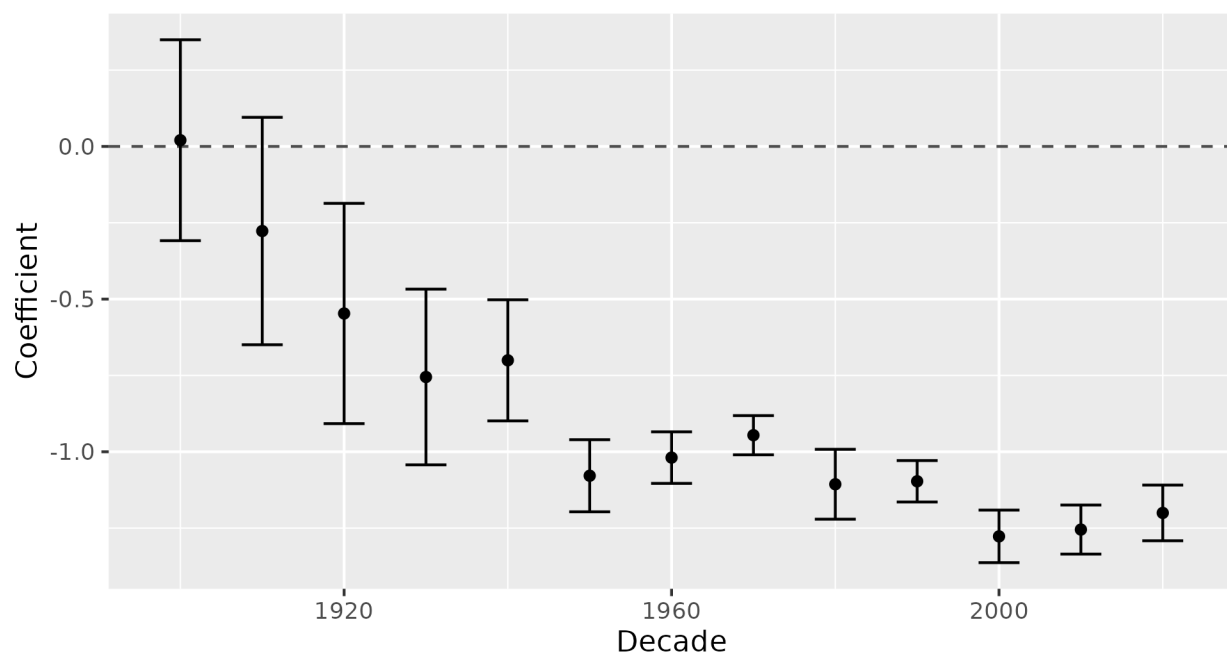
Note: Results are from a gravity framework for decade-level average trade flows between countries, estimated via Poisson pseudo-maximum likelihood to deal with zero flows. Coefficients are for distance between origin-destination pairs interacted with decade indicators. Vertical lines and whiskers indicate 95 percent confidence intervals. Other coefficients in the model, including those on origin and destination temperature, do not vary across decades.

Figure 4: Coefficients on temperature times log distance across decades (distance effect also varies)



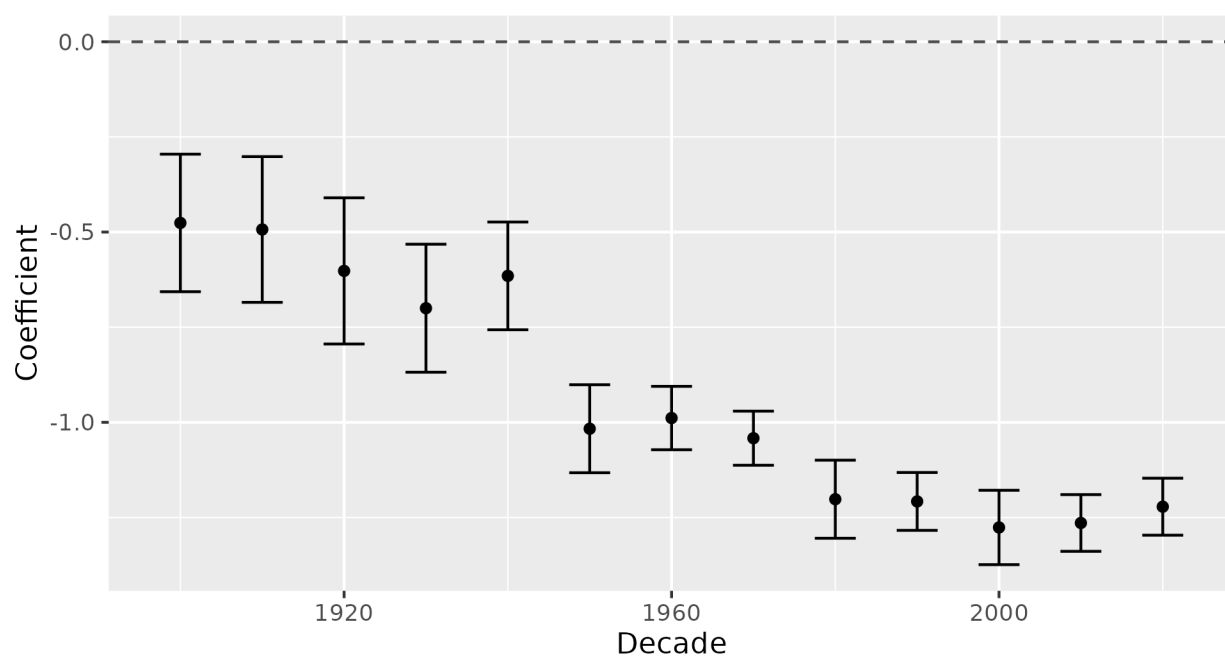
Note: Results are from a gravity framework for decade-level average trade flows between countries, estimated via Poisson pseudo-maximum likelihood to deal with zero flows. Coefficients are for temperature (in °C) at the origin and destination country. Vertical lines and whiskers indicate 95 percent confidence intervals. The effect of log bilateral distance on trade flows is also allowed to vary by decade.

Figure 5: Coefficients on log distance across decades (temperature effect also varies)



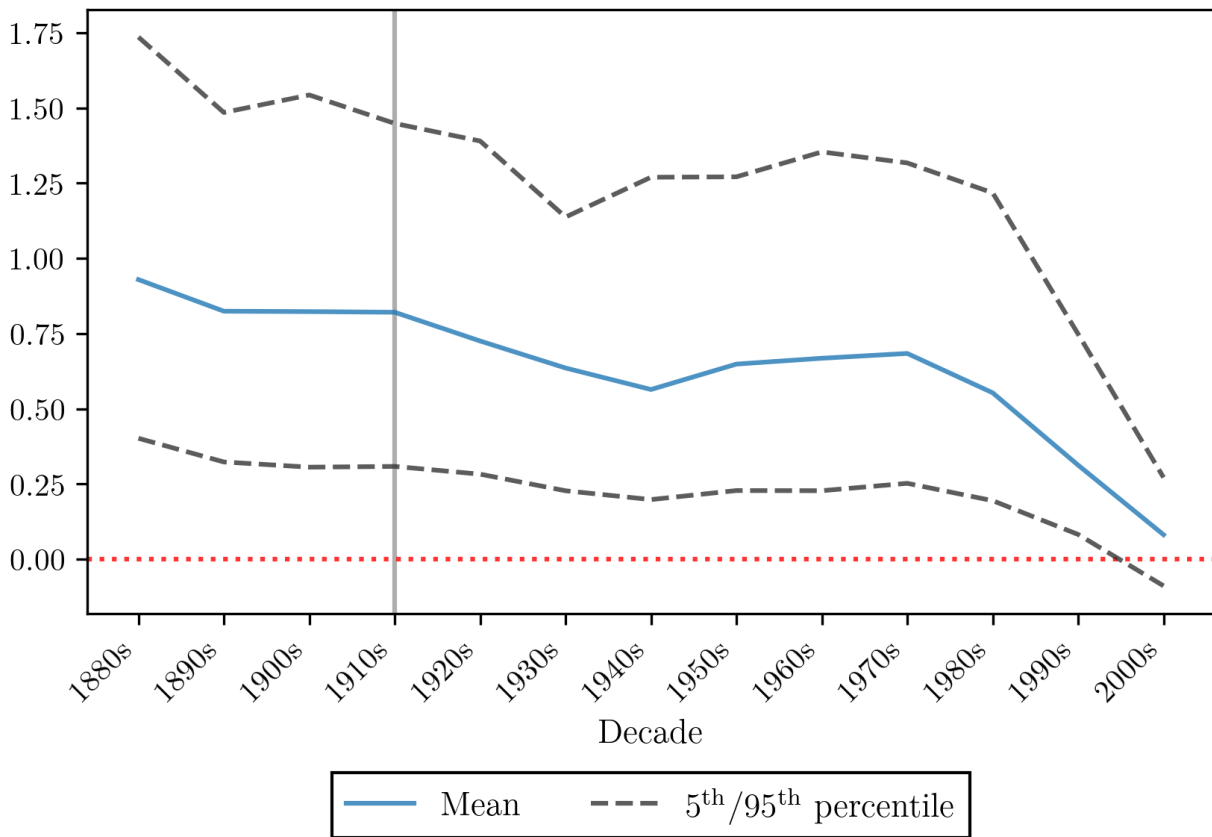
Note: Results are from a gravity framework for decade-level average trade flows between countries, estimated via Poisson pseudo-maximum likelihood to deal with zero flows. Coefficients are for distance between origin-destination pairs interacted with decade indicators. Vertical lines and whiskers indicate 95 percent confidence intervals. The effect of origin and destination temperatures on trade flows is also allowed to vary by decade.

Figure 6: Coefficients on log distance across decades (benchmark excluding temperature variables)



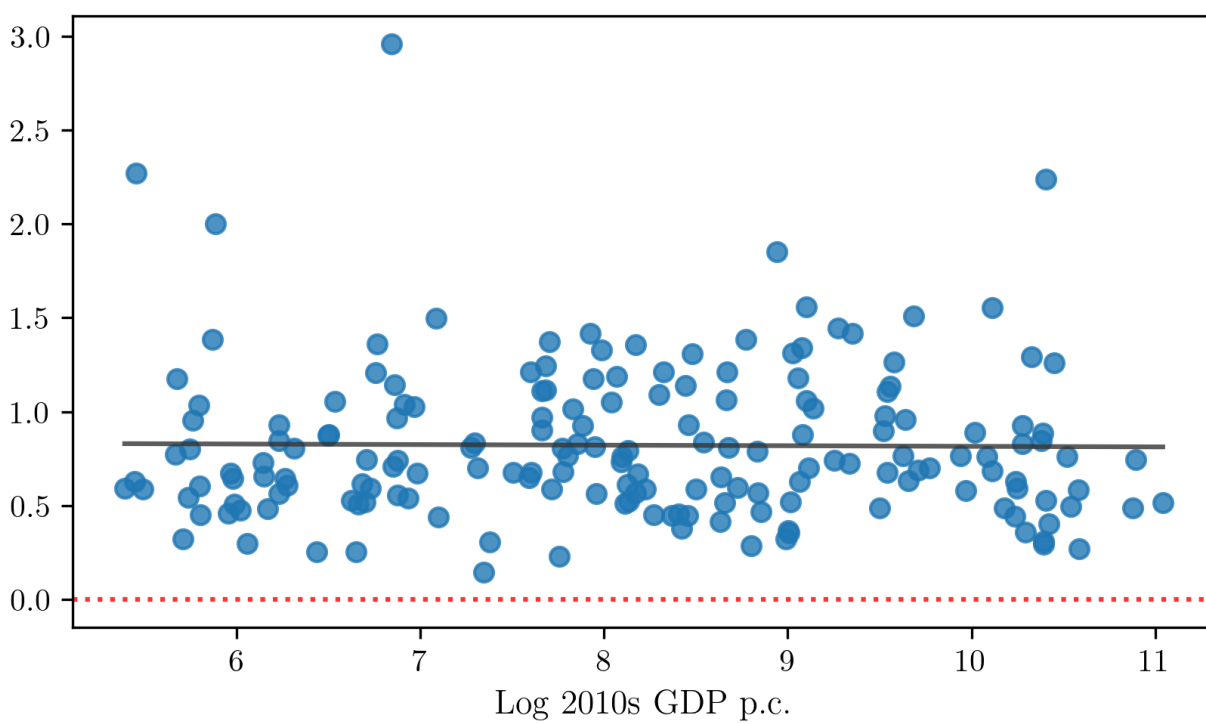
Note: Results are from a gravity framework for decade-level average trade flows between countries, estimated via Poisson pseudo-maximum likelihood to deal with zero flows. Coefficients are for distance between origin-destination pairs interacted with decade indicators. Vertical lines and whiskers indicate 95 percent confidence intervals. Other coefficients in the model do not vary across decades. This benchmark specification does not include origin and destination temperatures.

Figure 7: Welfare change (percent) across decades



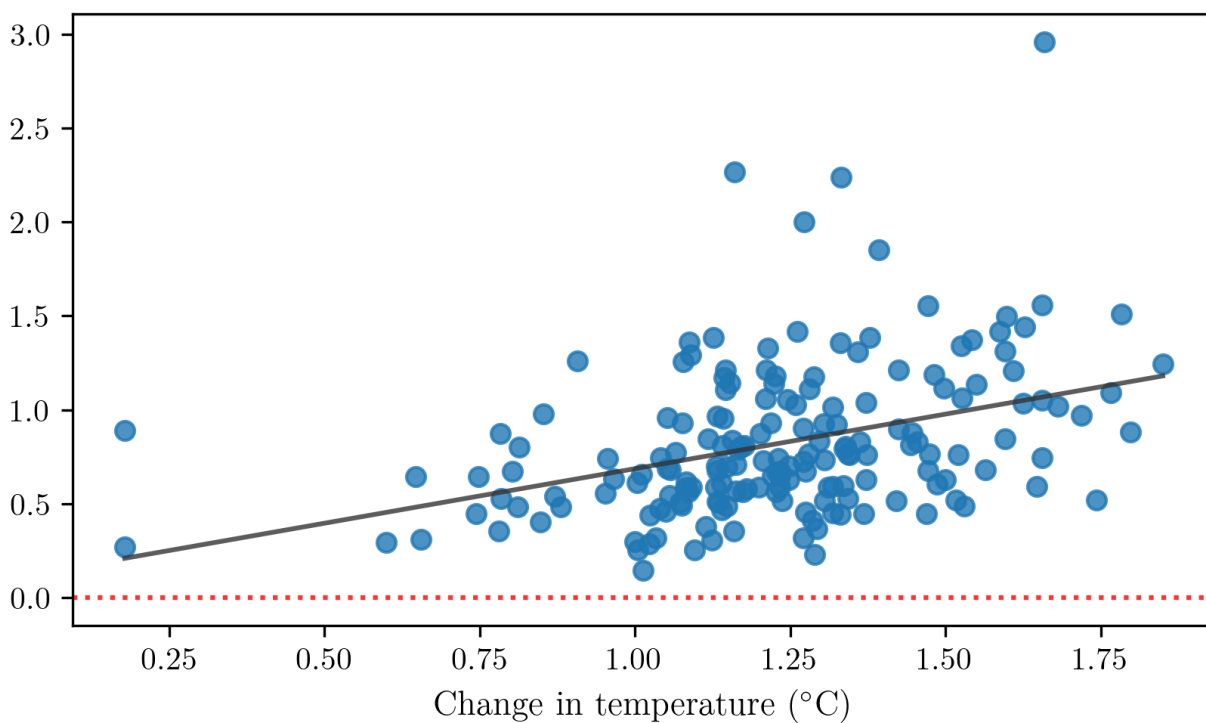
Note: The graph starts for the first decade where I have more than 90 percent non-missing weather data for countries in my sample. The solid vertical line indicates the first decade for which all countries have non-missing weather data.

Figure 8: Welfare change (percent) in 1910s climate counterfactual across 2010s GDP per capita



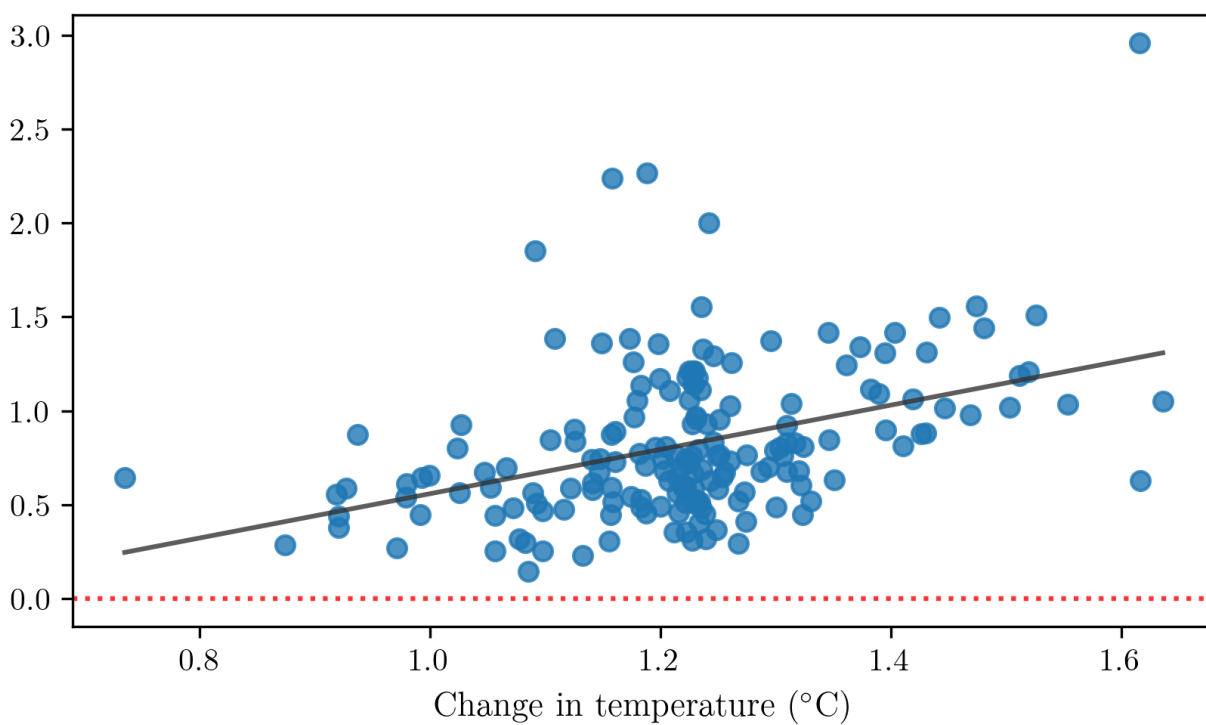
Note: The solid line shows a linear fit.

Figure 9: Welfare change (percent) in 1910s climate counterfactual across change in own temperature between the 1910s and 2010s



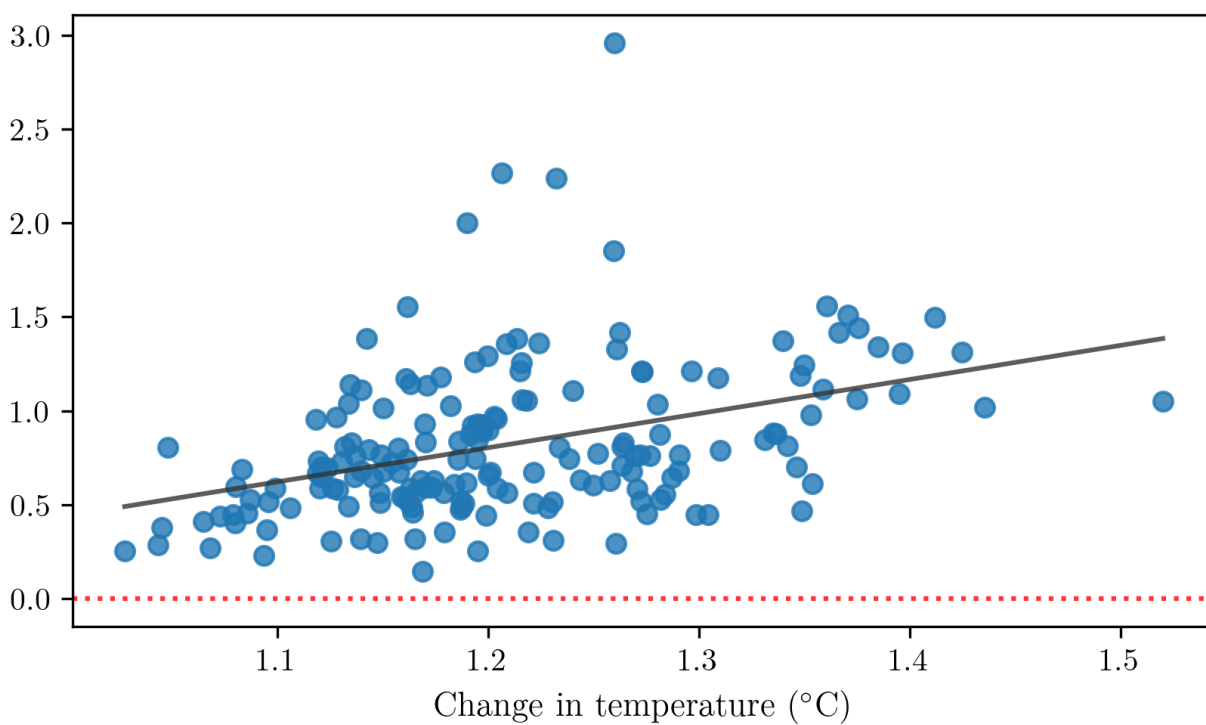
Note: Change in own temperature is the change in country i 's own temperature between the 1920s and 2010s. The solid line shows a linear fit.

Figure 10: Welfare change (percent) in 1910s climate counterfactual across change in export network temperature between the 1910s and 2010s



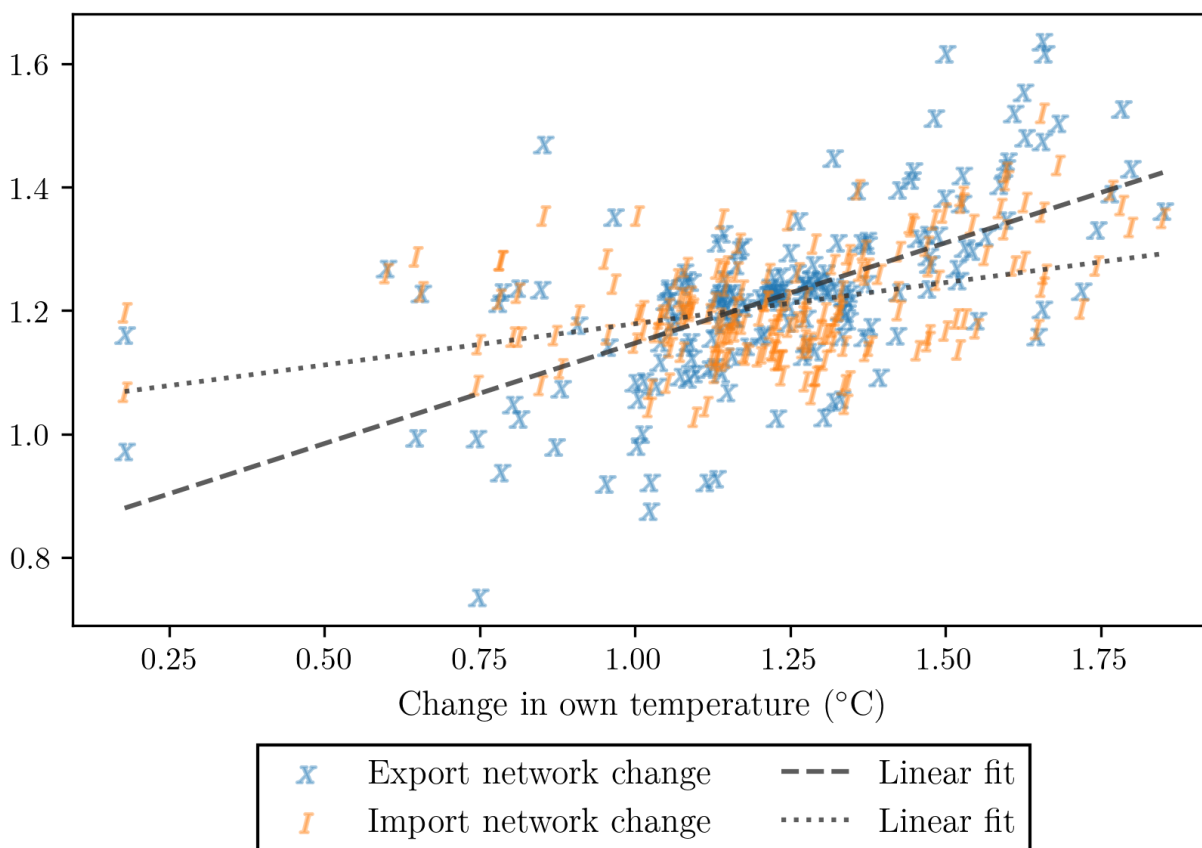
Note: The solid line shows a linear fit. The change in export network temperature for country i is the change in temperature across all other countries weighted by their share in country i 's total exports.

Figure 11: Welfare change (percent) in 1910s climate counterfactual across change in import network temperature between the 1910s and 2010s



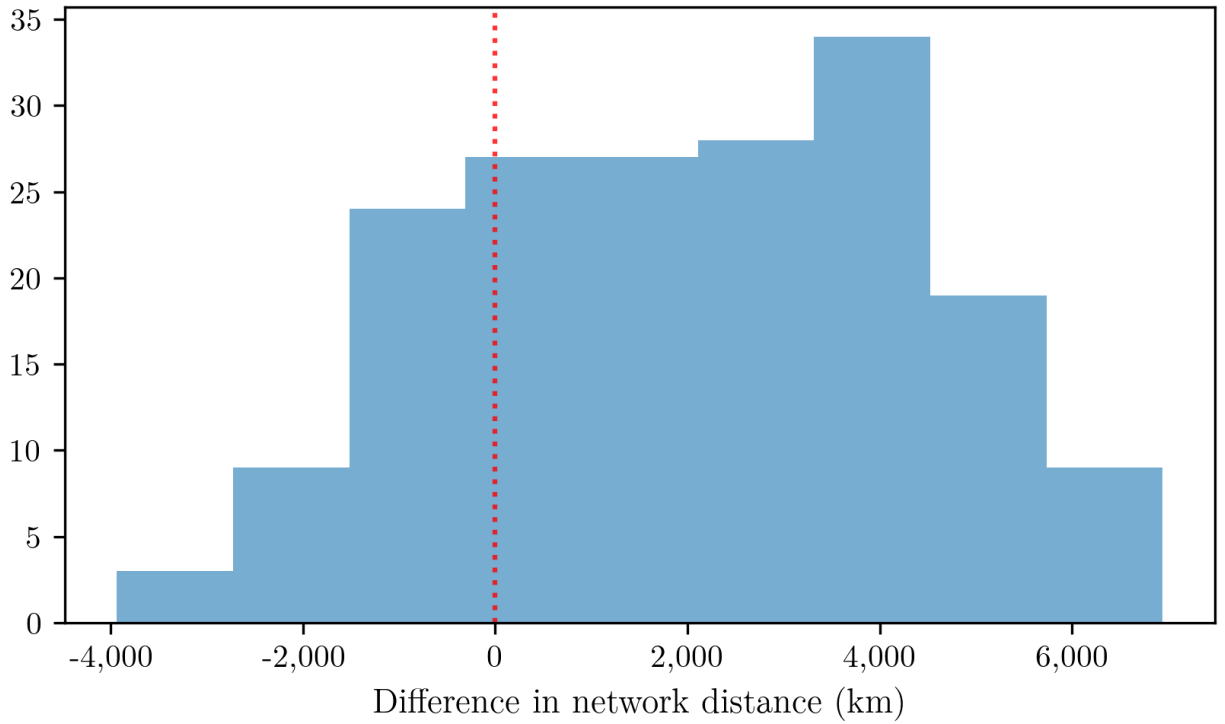
Note: The solid line shows a linear fit. The change in import network temperature for country i is the change in temperature across all other countries weighted by their share in country i 's total imports.

Figure 12: Change in trade network temperature between the 1910s and 2010s across change in own temperature between the 1910s and 2010s



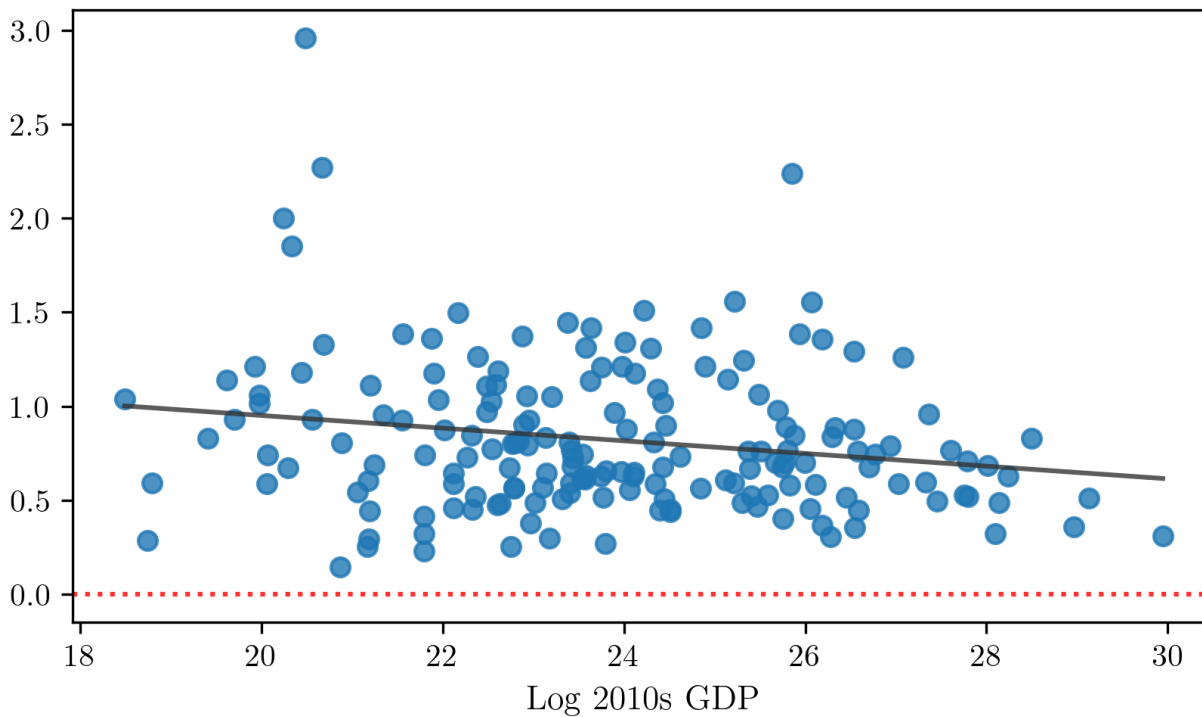
Note: The *import network change* for country i is the change in temperature across all other countries weighted by their share in country i 's total exports. The *export network change* for country i is the change in temperature across all other countries weighted by their share in country i 's total imports. The dashed line shows a linear fit for average temperature changes across the export network. The dotted line shows a linear fit for average temperature changes across the import network.

Figure 13: Histogram of import network distance minus export network distance



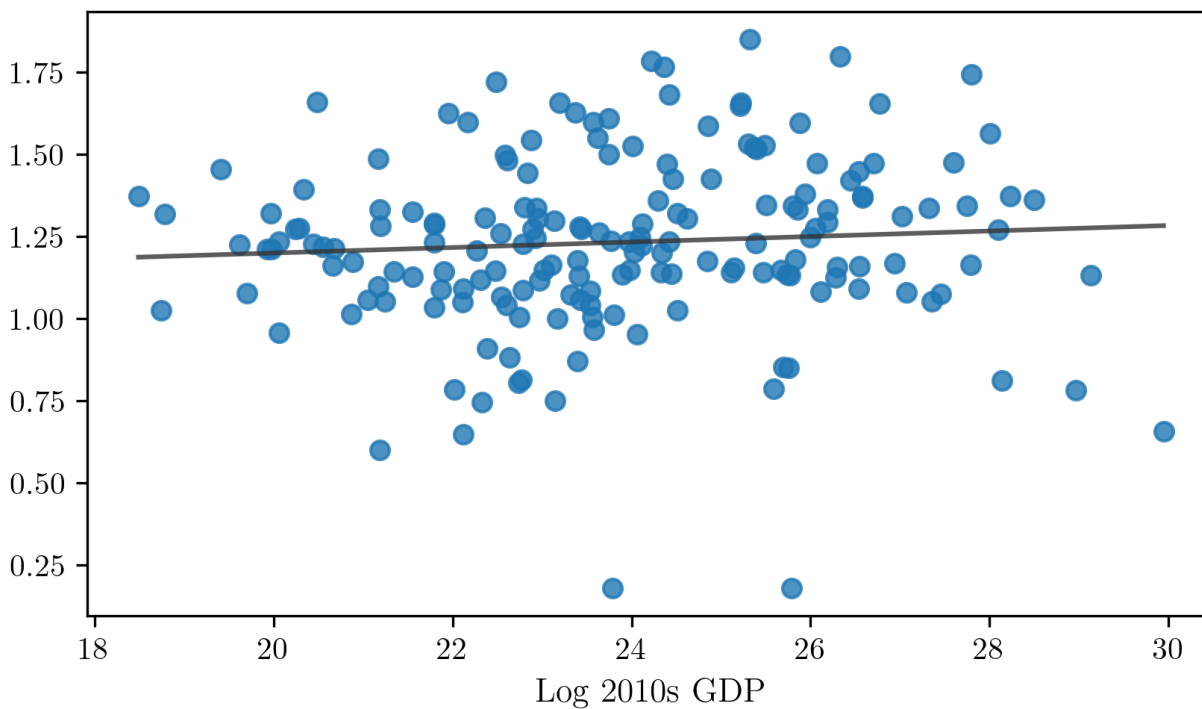
Note: The distance of the import network for country i is the distance to all other countries weighted by their share in country i 's total imports. The distance of the export network for country i is the distance to all other countries weighted by their share in country i 's total exports. The dotted line shows equal-distance networks. For observations to the right of that line, the import network spans a greater distance than the export network.

Figure 14: Welfare change (percent) in 1910s climate counterfactual across 2010s GDP



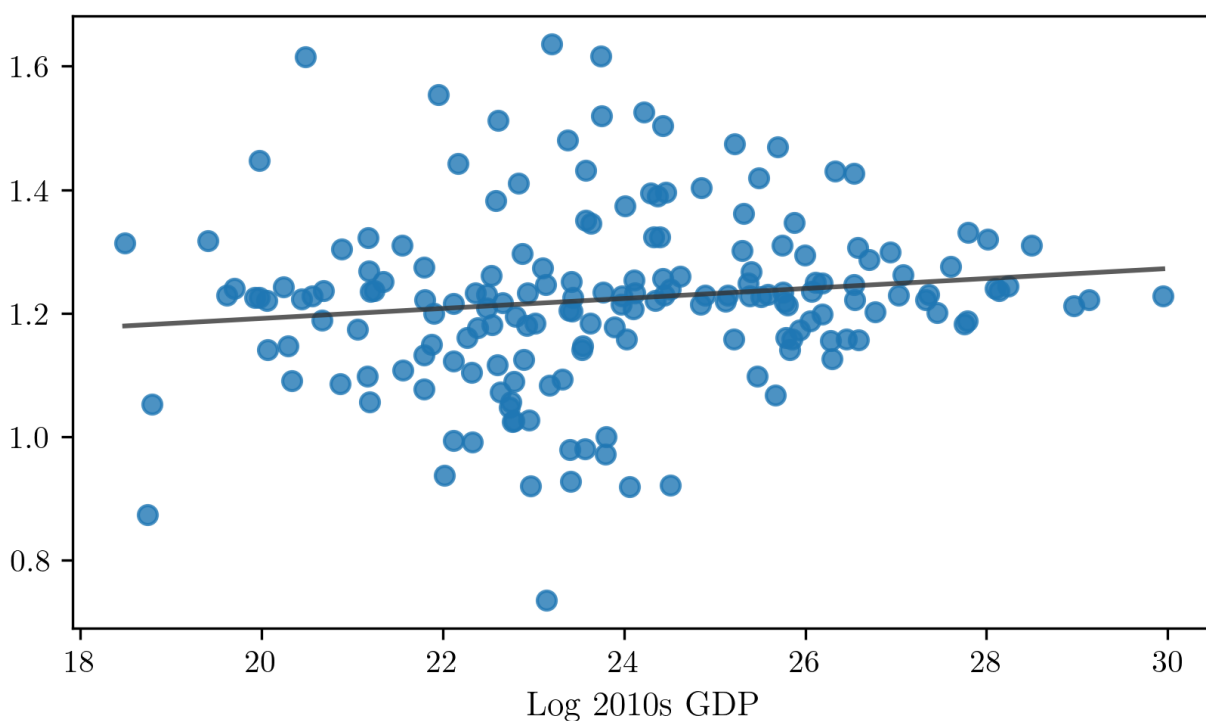
Note: The solid line shows a linear fit.

Figure 15: Change in own temperature between the 1910s and 2010s across 2010s GDP



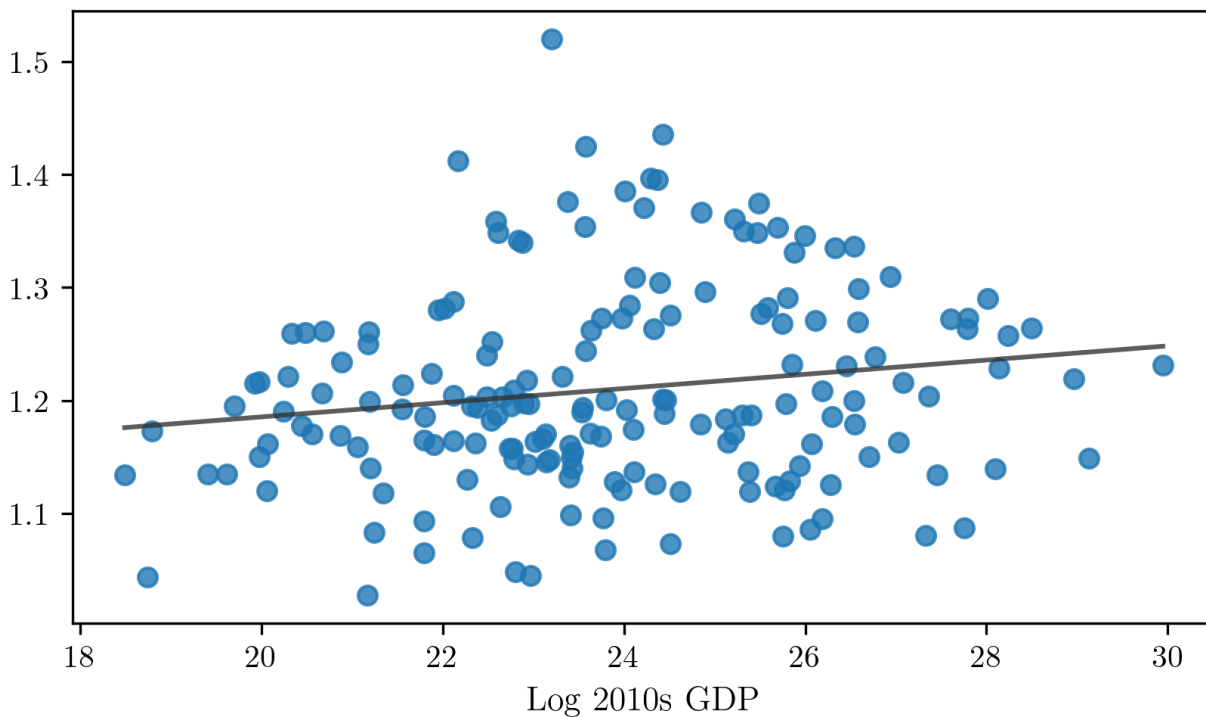
Note: Change in own temperature is the change in country i 's own temperature between the 1920s and 2010s. The solid line shows a linear fit.

Figure 16: Change in export network temperature between the 1910s and 2010s across 2010s GDP



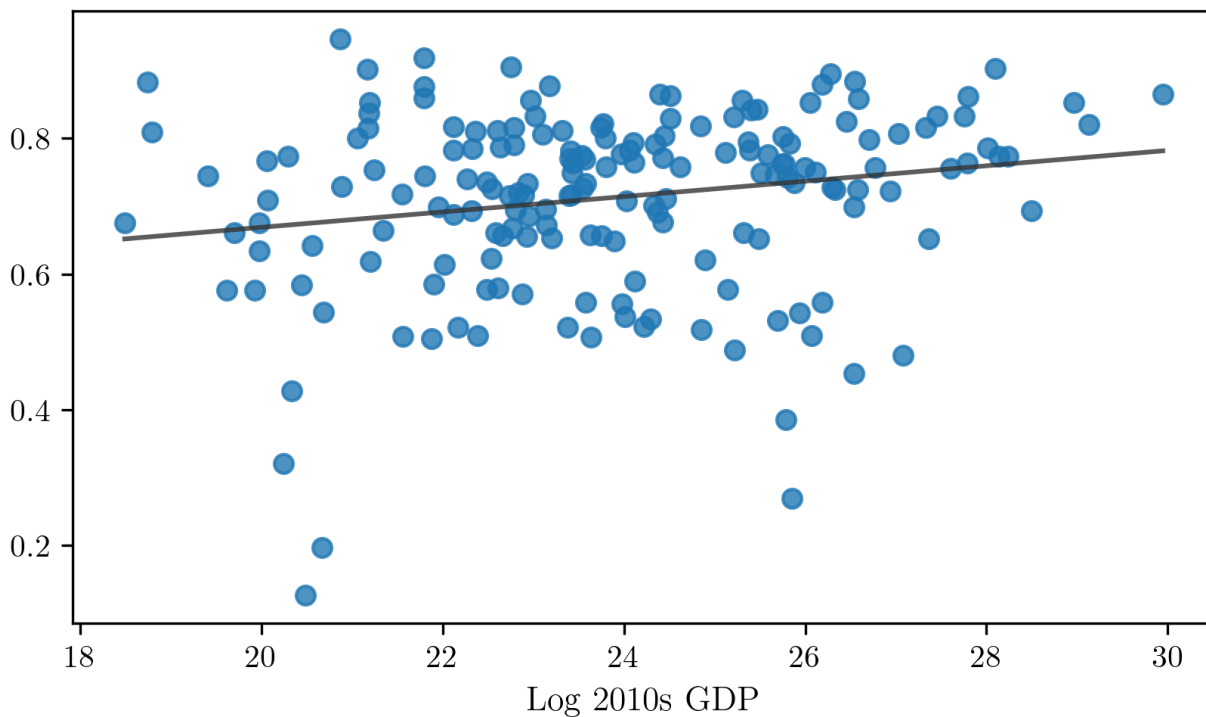
Note: The change in export network temperature for country i is the change in temperature across all other countries weighted by their share in country i 's total exports. The solid line shows a linear fit.

Figure 17: Change in import network temperature between the 1910s and 2010s across 2010s GDP



Note: The change in import network temperature for country i is the change in temperature across all other countries weighted by their share in country i 's total imports. The solid line shows a linear fit.

Figure 18: 2010s own trade share across 2010s GDP



Note: The solid line shows a linear fit.

Appendix A Additional tables

Table 3: Welfare change (percent) across decades

| Statistic | 1880s | 1890s | 1900s | 1910s | 1920s | 1930s | 1940s | 1950s | 1960s | 1970s | 1980s | 1990s | 2000s |
|-----------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|--------|
| Mean | 0.929 | 0.824 | 0.823 | 0.821 | 0.725 | 0.635 | 0.564 | 0.649 | 0.668 | 0.684 | 0.553 | 0.313 | 0.081 |
| p_5 | 0.402 | 0.323 | 0.306 | 0.308 | 0.283 | 0.227 | 0.198 | 0.228 | 0.228 | 0.252 | 0.195 | 0.082 | -0.089 |
| p_{10} | 0.486 | 0.411 | 0.406 | 0.435 | 0.354 | 0.309 | 0.226 | 0.286 | 0.314 | 0.324 | 0.238 | 0.101 | -0.055 |
| p_{25} | 0.606 | 0.571 | 0.525 | 0.550 | 0.482 | 0.433 | 0.348 | 0.418 | 0.426 | 0.446 | 0.334 | 0.175 | 0.035 |
| p_{50} | 0.829 | 0.730 | 0.719 | 0.741 | 0.652 | 0.588 | 0.491 | 0.564 | 0.604 | 0.587 | 0.459 | 0.253 | 0.080 |
| p_{75} | 1.205 | 1.047 | 1.030 | 1.037 | 0.871 | 0.782 | 0.672 | 0.781 | 0.840 | 0.846 | 0.687 | 0.413 | 0.136 |
| p_{90} | 1.531 | 1.360 | 1.413 | 1.331 | 1.224 | 0.990 | 1.039 | 1.096 | 1.087 | 1.177 | 1.063 | 0.582 | 0.193 |
| p_{95} | 1.736 | 1.485 | 1.543 | 1.449 | 1.390 | 1.137 | 1.270 | 1.271 | 1.354 | 1.317 | 1.218 | 0.749 | 0.272 |

Note: The table summarizes the estimated percent change in welfare under climate change counterfactuals for each decade. *Mean* reports the average welfare change for each decade, while p_x reports the x^{th} percentile of welfare changes for each decade.

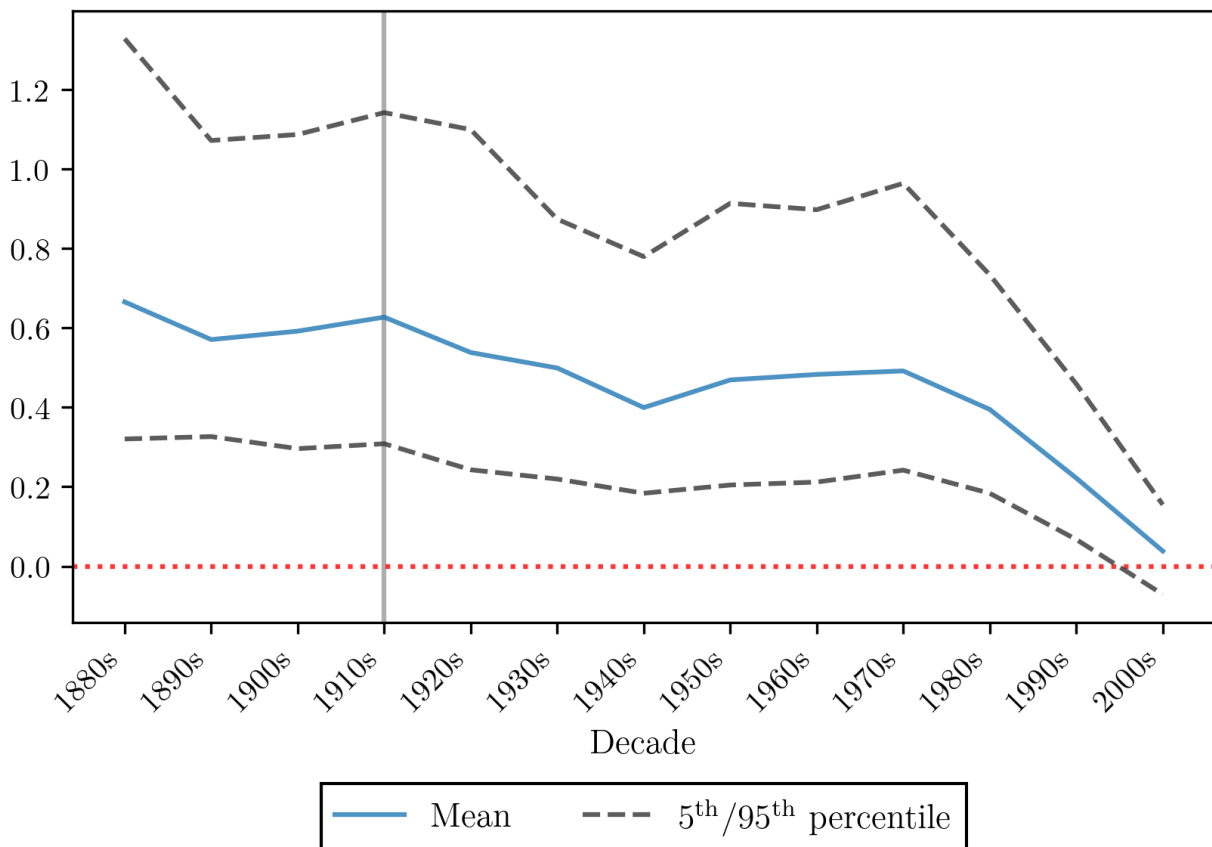
Table 4: Population-weighted welfare change (percent) across decades

| Statistic | 1880s | 1890s | 1900s | 1910s | 1920s | 1930s | 1940s | 1950s | 1960s | 1970s | 1980s | 1990s | 2000s |
|-----------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|--------|
| Mean | 0.666 | 0.571 | 0.592 | 0.627 | 0.538 | 0.499 | 0.400 | 0.469 | 0.483 | 0.492 | 0.395 | 0.222 | 0.038 |
| p_5 | 0.321 | 0.327 | 0.296 | 0.309 | 0.243 | 0.220 | 0.184 | 0.205 | 0.212 | 0.242 | 0.184 | 0.067 | -0.071 |
| p_{10} | 0.384 | 0.341 | 0.405 | 0.320 | 0.286 | 0.236 | 0.203 | 0.229 | 0.228 | 0.269 | 0.228 | 0.094 | -0.055 |
| p_{25} | 0.499 | 0.408 | 0.447 | 0.512 | 0.407 | 0.380 | 0.203 | 0.314 | 0.309 | 0.318 | 0.296 | 0.108 | -0.038 |
| p_{50} | 0.602 | 0.562 | 0.523 | 0.587 | 0.509 | 0.474 | 0.381 | 0.430 | 0.449 | 0.444 | 0.347 | 0.195 | 0.057 |
| p_{75} | 0.683 | 0.610 | 0.607 | 0.709 | 0.565 | 0.617 | 0.517 | 0.567 | 0.580 | 0.556 | 0.425 | 0.309 | 0.077 |
| p_{90} | 1.054 | 0.885 | 0.944 | 0.875 | 0.810 | 0.708 | 0.648 | 0.702 | 0.712 | 0.781 | 0.663 | 0.363 | 0.127 |
| p_{95} | 1.329 | 1.072 | 1.087 | 1.143 | 1.100 | 0.874 | 0.780 | 0.914 | 0.898 | 0.965 | 0.734 | 0.459 | 0.155 |

Note: The table summarizes the estimated percent change in welfare under climate change counterfactuals for each decade. *Mean* reports the average welfare change for each decade, while p_x reports the x^{th} percentile of welfare changes for each decade.

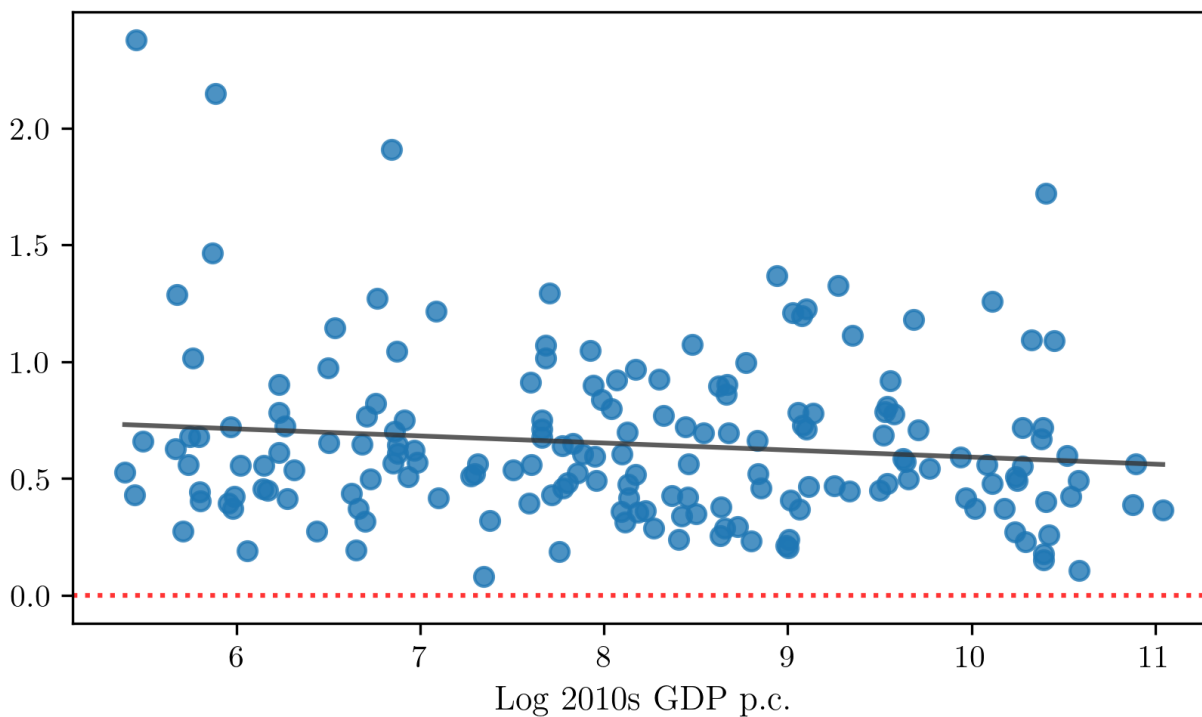
Appendix B Additional figures

Figure 19: Population-weighted welfare change (percent) across decades



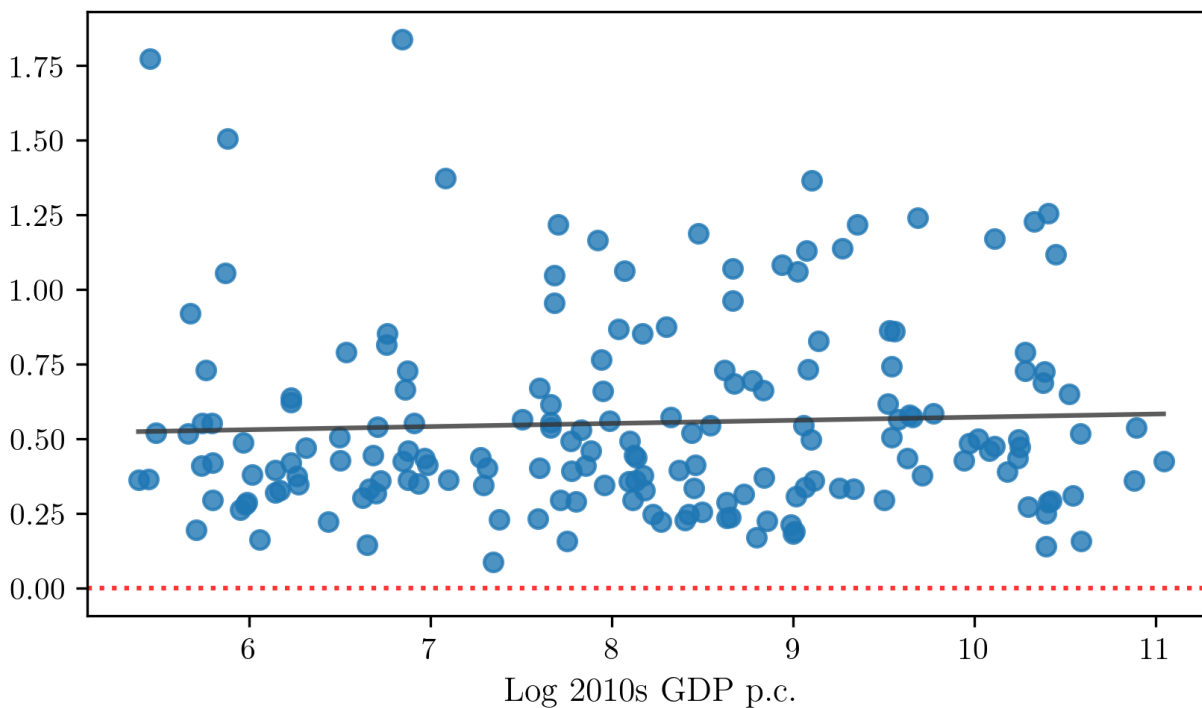
Note: The graph starts for the first decade where I have more than 90 percent non-missing weather data for countries in my sample. The solid vertical line indicates the first decade for which all countries have non-missing weather data.

Figure 20: Welfare change (percent) in 1950s climate counterfactual across 2010s GDP per capita



Note: The solid line shows a linear fit.

Figure 21: Welfare change (percent) in 1980s climate counterfactual across 2010s GDP per capita



Note: The solid line shows a linear fit.