

## Gaussian Mixture Models

It is a model with a PDF of the form :

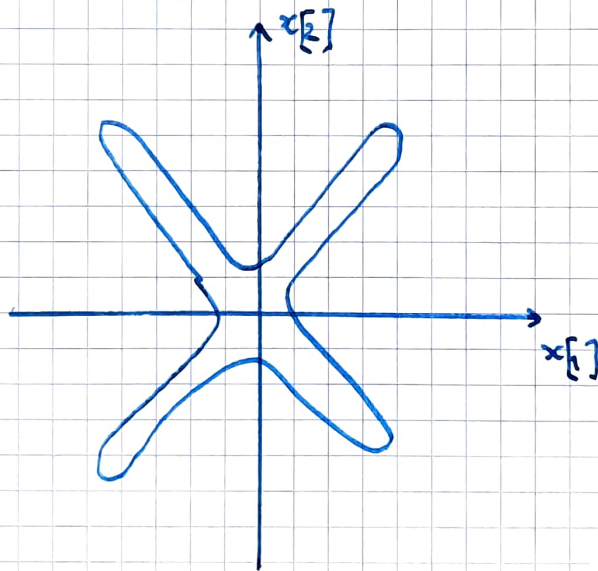
$$\sum_{j=1}^J w_j \mathcal{N}(\mu_j, \Sigma_j)$$

where  $\mathcal{N}(x | \mu_j, \Sigma_j) = \frac{1}{\sqrt{(2\pi)^m \det \Sigma_j}} \exp\left(-\frac{1}{2}(x - \mu_j)^T \Sigma_j^{-1} (x - \mu_j)\right)$

$\in \mathbb{R}^m$   $\rightarrow$   $m$  is often 2 because we often work in  $\mathbb{R}^2$  (2 Dimensions)

First case : we want to build a model of this type

Example : we want to create a density in  $\mathbb{R}^2$  such that the iso-densities look like this :



According to Gaussian - Multivariate - Covariance - Matrix / Gaussian.pdf we can define :

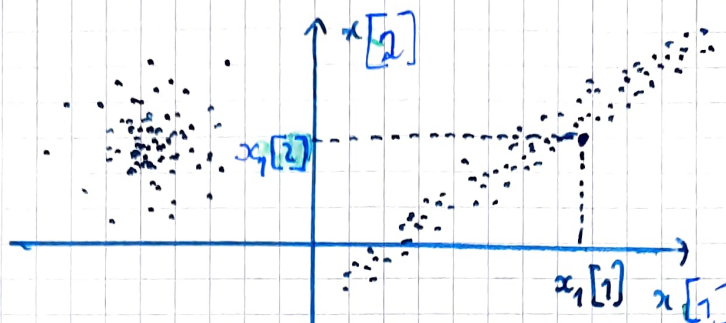
$$\text{PDF}(x) = \frac{1}{2} \mathcal{N}(0, \Sigma_1) + \frac{1}{2} \mathcal{N}(0, \Sigma_2)$$

with  $\Sigma_1 = Q D_1 Q^T$      $\Sigma_2 = Q D_2 Q^T$      $Q = \begin{pmatrix} -\frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \end{pmatrix}$

$$D_1 = \begin{pmatrix} 10 & 0 \\ 0 & 1 \end{pmatrix}$$

$$D_2 = \begin{pmatrix} 1 & 0 \\ 0 & 10 \end{pmatrix}$$

Case two: We have a cloud of points and want to find the best GMM to model these points (real life)



These "observed" points are denoted by  $x_1, x_2, \dots, x_n$ . For each point  $x_i$  ( $1 \leq i \leq n$ ),  $x_i[1]$  is his first coordinate and  $x_i[2]$  the second.

Step 1: choose J

Algorithms won't do it for us

Now we have a model:

$$PDF(x) = \sum_{j=1}^J w_j \mathcal{N}(\mu_j, \Sigma_j)$$

We have  $\underbrace{J}_{w_j} + \underbrace{mJ}_{\mu_j} + \underbrace{\frac{J(J-1)}{2}}_{\Sigma_j} = \frac{J^2}{2} + \underbrace{\left(m + \frac{1}{2}\right)}_{\substack{m=2 \\ \text{in the graphics above} \\ \text{(because 2 Dimensional)}}} J$  parameters to optimize

Step 2: Use EM algorithm:

At step  $t$ :

① Compute  $\gamma_{ij}^{(t)} = \frac{w_j^{(t)} \mathcal{N}(x_i | \mu_j^{(t)}, \Sigma_j^{(t)})}{\sum_{k=1}^J w_k^{(t)} \mathcal{N}(x_i | \mu_k^{(t)}, \Sigma_k^{(t)})}$

and  $w_j^{(t+1)} = \frac{1}{n} \sum_{i=1}^n \gamma_{ij}^{(t)}$

② Compute  $\mu_j^{(t+1)} = \frac{\sum_{i=1}^n \gamma_{ij}^{(t)} x_i}{\sum_{i=1}^n \gamma_{ij}^{(t)}}$  for all  $j \in [1, J]$

and  $\Sigma_j^{(t+1)} = \frac{\sum_{i=1}^n \gamma_{ij}^{(t)} (x_i - \mu_j^{(t+1)})(x_i - \mu_j^{(t+1)})^T}{\sum_{i=1}^n \gamma_{ij}^{(t)}}$  for all  $j \in [1, J]$



This algorithm converges to a local maximum likelihood.

See the folder GMM - EM - Proof because the proof is particularly interesting in itself, with the use of:

- Jensen inequality
- Deriving the Lagrangian
- Kullback-Leibler divergence
- Some nice differential calculations
- Conditional probabilities

Without proving anything, note that the EM algorithm is intuitive because:

$$\gamma_{ij}^{(t)} = \frac{P(z_i=j) \mathcal{N}(x_i|z_i=j)}{\mathcal{N}(x_i)} = \underbrace{P}_{\text{prior}}(z_i=j | x_i=x_i)$$

is the probability of belonging to the  $j$ -th Gaussian given  $x_i$ :

- This probability is used to weight each new average parameter

$$\mu_j^{(t+1)} = \frac{\sum_i \gamma_{ij}^{(t)} x_i}{\sum_i \gamma_{ij}^{(t)}}$$

- Similarly for the expression of the covariance.