

Computation of the MASC Estimator

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Given a treated unit treated in period T and a set of control units, The estimator $\hat{\mu}_t^{masc}$ fit to outcomes up to period $t, t < T$ is defined as:

$$\hat{\mu}_t^{masc}(\phi, m) \equiv \phi \hat{\mu}_t^{ma}(m) + (1 - \phi) \hat{\mu}_t^{sc}$$

where $\hat{\mu}_t^{ma}(m)$ is an m -nearest neighbor estimator and $\hat{\mu}_t^{sc}$ is a standard synthetic control estimator, both fit to the path of outcomes up to period t . The model averaging parameter ϕ governs the weight placed on matching versus synthetic control, $\phi \in [0, 1]$.

Our implementation uses a rolling-origin cross-validation procedure to resolve the implicit trade off between interpolation bias versus extrapolation bias in choosing the tuning parameters m and ϕ . This involves making a series of one-step ahead forecasts, each of which is estimated only using data from periods prior to the forecast date. Mathematically, the criterion is:

$$Q(\phi, m) = \frac{1}{|\mathcal{F}|} \sum_{f \in \mathcal{F}} (y_{1,f+1} - \hat{\mu}_{f+1}^{masc}(\phi, m))^2$$

where $y_{1,f+1}$ is the outcome of the treated unit in period $f + 1$, and \mathcal{F} is a subset of time periods taken before the treatment period. There is a natural bias-variance trade off which drives the choice of folds f to include in \mathcal{F} . Later periods are preferred because they will likely be more relevant to the post-treatment period and will use more data. However, including earlier periods may reduce the variance in $Q(m, \phi)$. We recommend that the analyst pick a cutoff value f^* , defining \mathcal{F} to include all $f \geq f^*$ in the pre-treatment period. The analyst must also choose the set of candidate nearest neighbor estimators from which to pick m .

Computationally, we solve the cross-validation problem in two steps. First, we fix the candidate nearest neighbor estimator m . Conditional on m , the cross-validation criterion has an analytic solution for ϕ using least square algebra:

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$$\phi^*(m) = \frac{\sum_{f \in \mathcal{F}} (\hat{\mu}_{f+1}^{ma}(m) - \hat{\mu}_{f+1}^{sc})(y_{1,f+1} - \hat{\mu}_{f+1}^{sc})}{\sum_{f \in \mathcal{F}} (\hat{\mu}_{f+1}^{ma}(m) - \hat{\mu}_{f+1}^{sc})^2}$$

where the the real solution $\hat{\phi}(m)$ is then defined to respect the bounds on ϕ :

$$\hat{\phi}(m) \equiv \begin{cases} 0, & \text{if } \phi^*(m) \leq 0 \\ 1, & \text{if } \phi^*(m) \geq 1 \\ \phi^*(m) & \text{otherwise} \end{cases}$$

For the second step, we then set $\hat{m} \equiv \underset{m}{\operatorname{argmin}} Q(\hat{\phi}(m), m)$ and set $\hat{\phi} = \hat{\phi}(\hat{m})$. The cross-validated MASC estimator is a weighted average of $\hat{\mu}_T^{sc}$ and $\hat{\mu}_T^{ma}(\hat{m})$ with weights $(1 - \hat{\phi})$ and $\hat{\phi}$ respectively.