Advanced Macroeconometrics – Assignment 1

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The executable code that was used in compiling the assignment is available on GitHub at $\frac{https://github.com/maxmheinze/macrometrics}{}.$

Exercise 1

First, we read in the data set, removing the first column indicating the recommended transformation.

```
# Header -----
rm(list = ls())
gc()

pacman::p_load(
    tidyverse,
    urca,
    vars
)

# Read in Data ------

fred <- read.csv("./assignment1/data/fred.csv")[-1,]

kpdata <- read.table("./assignment1/data/data_kilian_park_2009.txt", header = FALSE)
colnames(kpdata) <- c("oil_prod_change", "econ_act_real", "oil_price_real",
    "div_growth_change_real")</pre>
```

Creating the Function

Next, we create the desired function ts_explode(). Along with the vector to be transformed, it asks for a specification of whether the data is ordered with the latest or earliest value first. It takes the earliest value first as default.

```
# Create the function --------
ts_explode <- function(input_vector, start_with_latest = FALSE) {</pre>
  # The function is called ts_explode because a single vector explodes into an entire data
  # Package dplyr required for lag() function
  require(dplyr)
  # Reverse input vector if user specifies it is sorted latest to earliest
  input_vector <- if (start_with_latest == FALSE) {</pre>
   input_vector
  } else {
   rev(input_vector)
  # Do the transformations, assign transformed vectors
 original <- input_vector</pre>
 log_transformed <- log(input_vector)</pre>
 mom_growth <- input_vector/dplyr::lag(input_vector, 1) - 1</pre>
 mom_growth_log <- log(input_vector/dplyr::lag(input_vector, 1))</pre>
 yoy_growth <- input_vector/dplyr::lag(input_vector, 12) - 1</pre>
 yoy_growth_lagged <- dplyr::lag(input_vector, 12)/dplyr::lag(input_vector, 24) - 1</pre>
  # Create a data frame to export, reverse ordering back to original in case
  start_with_latest = TRUE was specified
  export_df <- if (start_with_latest == FALSE) {</pre>
   data.frame(original, log_transformed, mom_growth, mom_growth_log, yoy_growth,
   yoy_growth_lagged)
  } else {
   data.frame(original = rev(original),
```

Preparing the Data Frame

Using ts_explode(), we create the data frame ind_prod including all transformations of the INDPRO variable. We bind the data frame together with the date column, which we transform from character to date. All other changes in the resulting data frame are of cosmetic nature.

```
# Prepare Industrial Production Data Frame -----

ind_prod <- fred$sasdate %>%
   cbind(ts_explode(fred$INDPRO)) %>%
   as_tibble() %>%
   mutate(date = lubridate::mdy(`.`)) %>%
   dplyr::select(-`.`) %>%
   relocate(date, .before = original)
```

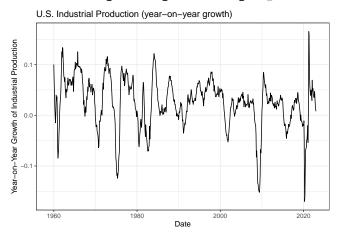
Plots

Next, we plot both the logged variable and the year-on-year growth rate.

Warning: Removed 1 row containing missing values (`geom_line()`).



Warning: Removed 13 rows containing missing values (`geom_line()`).



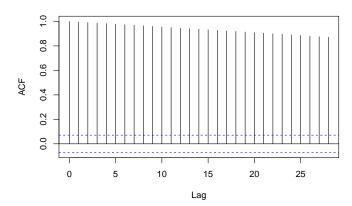
The plots above depict the logged time series of U.S. industrial production and its year-on-year growth rates. The logged time series of U.S. industrial production appears to exhibit a generally upward trend over time. The logged time series of U.S. industrial production also shows evidence of cyclical patterns, with periods of expansion and contraction occurring in a cyclical manner. There are also periodic fluctuations in the year-on-year growth rates, which suggests the presence of seasonality in U.S. industrial production.

Assessing Properties

Next we assess the properties of both logged industrial production and its yearly growth rate. We start with logged industrial production:

```
# ACF plot of Logged Industrial Production
stats::acf(ind_prod$log_transformed[!is.na(ind_prod$log_transformed)], main = "ACF of Logged
Industrial Production")
```

ACF of Logged Industrial Production



The ACF of Logged Industrial Production reveals a highly persistent nature of the time series with significant autocorrelation coefficients, particularly between Logged Industrial Production and its first five lags, which are extremely close to 1. The autocorrelation coefficients remain highly significant in the displayed laggs. This highly persistent nature suggests potential non-stationarity in the series. For a more appropriate procedure, we apply the Augmented Dickey-Fuller (ADF) unit root tests.

Type: "trend" (urtest1a) H0: The time series is random walk around a trend H1: The time series is trend-stationary

```
#ADF test 1a) of Logged Industrial Production
urtest1a = ur.df(ind_prod$log_transformed[!is.na(ind_prod$log_transformed)], type = "trend",
selectlags = "AIC")
summary(urtest1a)
```

```
##
## # Augmented Dickey-Fuller Test Unit Root Test #
##
## Test regression trend
##
##
## Call:
## lm(formula = z.diff \sim z.lag.1 + 1 + tt + z.diff.lag)
##
## Residuals:
                  1Q
                       Median
                                             Max
## -0.131711 -0.003784 0.000414
                              0.004158
                                        0.054957
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
              1.923e-02 1.006e-02
                                   1.911
                                           0.0564 .
## (Intercept)
## z.lag.1
             -4.932e-03
                        2.998e-03
                                  -1.645
                                           0.1004
## tt
              6.338e-06
                        6.121e-06
                                   1.036
                                           0.3007
## z.diff.lag
              2.889e-01
                        3.458e-02
                                   8.353 3.12e-16 ***
## ---
                 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 0.009501 on 763 degrees of freedom
## Multiple R-squared: 0.0963, Adjusted R-squared: 0.09275
## F-statistic: 27.1 on 3 and 763 DF, p-value: < 2.2e-16
##
##
## Value of test-statistic is: -1.6449 7.7617 3.7276
##
## Critical values for test statistics:
##
        1pct 5pct 10pct
## tau3 -3.96 -3.41 -3.12
```

```
## phi2 6.09 4.68 4.03
## phi3 8.27 6.25 5.34
```

As the coefficient of trend (tt) is insignificant, we never apply the ADF test with drift and trend. Instead, we move on to the second equation including only the drift component.

Type: "drift" (urtest2a) H0: The time series is random walk around a drift H1: The time series is (log-level) stationary around a drift

```
#ADF test 2a) of Logged Industrial Production
urtest2a = ur.df(ind_prod$log_transformed[!is.na(ind_prod$log_transformed)], type = "drift",
selectlags = "AIC")
summary(urtest2a)
```

```
##
## # Augmented Dickey-Fuller Test Unit Root Test #
##
## Test regression drift
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 + 1 + z.diff.lag)
##
## Residuals:
##
       Min
                 1Q
                       Median
                                   30
                                           Max
                                       0.055859
## -0.131010 -0.003791 0.000361 0.004097
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.0093366 0.0031640
                                 2.951 0.00327 **
## z.lag.1
            -0.0019295 0.0007638 -2.526 0.01173 *
## z.diff.lag
            0.2866319 0.0345154
                                 8.304 4.55e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.009501 on 764 degrees of freedom
## Multiple R-squared: 0.09503,
                                Adjusted R-squared: 0.09266
## F-statistic: 40.11 on 2 and 764 DF, p-value: < 2.2e-16
##
##
## Value of test-statistic is: -2.5263 11.1054
##
## Critical values for test statistics:
       1pct 5pct 10pct
## tau2 -3.43 -2.86 -2.57
## phi1 6.43 4.59 3.78
```

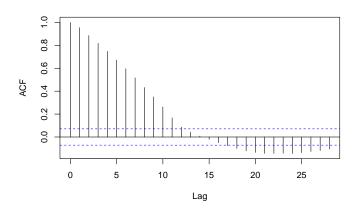
As the drift component is significant, and the absolute value of the test-statistic (2.5263) is below all the critical values for conventional significance levels, we conclude that we fail to reject the null hypothesis that the time series is random walk around a drift.

Thus, we conclude that Logged Industrial Production is I(1).

Next we assess the ACF of the yearly growth rate of Instustrial production:

```
#ACF plot of Year-on-Year Growth Rate
stats::acf(ind_prod$yoy_growth[!is.na(ind_prod$yoy_growth)], main = "ACF of YoY Growth of
Industrial Production")
```

ACF of YoY Growth of Industrial Production



The ACF of YoY Growth rate of Industrial Production shows positive persistence, but to a lesser degree than the former series. The correlation coefficient decreases relatively quickly and even turns significantly negative after the 17th lag, indicating the presence of a pattern or cycle in the data at specific time lags. Now we apply the Augmented Dickey-Fuller (ADF) unit root tests.

Type: "trend" (urtest1b) H0: The time series is random walk around a trend H1: The time series is trend-stationary

```
#ADF test 1b) of Year-on-Year Growth Rate
urtest1b = ur.df(ind_prod$yoy_growth[!is.na(ind_prod$yoy_growth)], type = "trend",
selectlags = "AIC")
summary(urtest1b)
```

```
##
## # Augmented Dickey-Fuller Test Unit Root Test #
##
## Test regression trend
##
##
## lm(formula = z.diff ~ z.lag.1 + 1 + tt + z.diff.lag)
##
## Residuals:
                       Median
                  10
                                    30
                                            Max
## -0.112205 -0.005982 0.000166 0.005728
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
              2.846e-03
                        1.088e-03
                                   2.615
                                          0.0091 **
             -6.130e-02
                        1.047e-02
                                  -5.852 7.24e-09 ***
## z.lag.1
## tt
             -3.507e-06
                        2.295e-06
                                  -1.528
                                          0.1268
              3.272e-01
                        3.435e-02
                                   9.527
## z.diff.lag
                                         < 2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.01324 on 751 degrees of freedom
## Multiple R-squared: 0.1286, Adjusted R-squared: 0.1251
## F-statistic: 36.94 on 3 and 751 DF, p-value: < 2.2e-16
##
##
## Value of test-statistic is: -5.8523 11.4203 17.1258
##
## Critical values for test statistics:
##
        1pct 5pct 10pct
## tau3 -3.96 -3.41 -3.12
## phi2 6.09 4.68 4.03
```

```
## phi3 8.27 6.25 5.34
```

As the coefficient of trend (tt) is insignificant, we never apply the ADF test with drift and trend. Instead, we move on to the second equation including only the drift component.

Type: "drift" (urtest2b) H0: The time series is random walk around a drift H1: The time series is stationary around a drift

```
#ADF test 2b) of Year-on-Year Growth Rate
urtest2b = ur.df(ind_prod$yoy_growth[!is.na(ind_prod$yoy_growth)], type = "drift",
selectlags = "AIC")
summary(urtest2b)
```

```
##
## # Augmented Dickey-Fuller Test Unit Root Test #
##
## Test regression drift
##
##
## Call:
## lm(formula = z.diff \sim z.lag.1 + 1 + z.diff.lag)
##
## Residuals:
##
       Min
                 1Q
                       Median
                                   3Q
                                           Max
## -0.113192 -0.006224 0.000028 0.005802 0.133866
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.0014075 0.0005468
                                  2.574
                                         0.0102 *
## z.lag.1
             -0.0570068 0.0100998
                                 -5.644 2.35e-08 ***
## z.diff.lag
              0.3246308 0.0343372
                                  9.454 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.01325 on 752 degrees of freedom
## Multiple R-squared: 0.1259, Adjusted R-squared: 0.1235
## F-statistic: 54.14 on 2 and 752 DF, p-value: < 2.2e-16
##
##
## Value of test-statistic is: -5.6443 15.934
##
## Critical values for test statistics:
        1pct 5pct 10pct
## tau2 -3.43 -2.86 -2.57
## phi1 6.43 4.59 3.78
```

As the drift component is significant, and the absolute value of the test-statistic (5.6443) is above all the critical values for conventional significance levels, we conclude that we reject the null hypothesis that the time series is random walk around a drift.

Thus, we conclude that the YoY Growth of Industrial Production is I(0).

-none- numeric

Estimating an AR Model

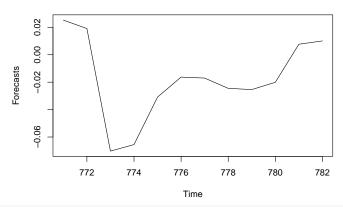
1

order

```
# Estimate AR model -----
ar_model_yoygrowth = ar.ols(ind_prod$yoy_growth[!is.na(ind_prod$yoy_growth)])
summary(ar_model_yoygrowth)
## Length Class Mode
```

```
## ar
               26
                   -none- numeric
## var.pred
               1 -none- numeric
               1
                     -none- numeric
## x.mean
## x.intercept 1
                     -none- numeric
## aic
               29
                     -none- numeric
## n.used
                1
                     -none- numeric
## n.obs
                1
                     -none- numeric
## order.max
               1 -none- numeric
## partialacf 0 -none- NULL
## resid
             757 -none- numeric
## method
               1
                     -none- character
## series
                     -none- character
                1
## frequency
                1
                     -none- numeric
## call
                2
                     -none- call
## asy.se.coef 2
                     -none- list
# Generate forecasts for the next year
forecasts <- predict(ar_model_yoygrowth, n.ahead = 12)</pre>
print(forecasts)
## $pred
## Time Series:
## Start = 758
## End = 769
## Frequency = 1
## [1] 0.025269455 0.019107977 -0.070284530 -0.065514117 -0.030876699
## [6] -0.016306685 -0.017024809 -0.024510326 -0.025405957 -0.020109857
## [11] 0.007656505 0.010083272
##
## $se
## Time Series:
## Start = 758
## End = 769
## Frequency = 1
## [1] 0.01016206 0.01654370 0.02111975 0.02485038 0.02826507 0.03122660
## [7] 0.03401918 0.03654871 0.03859429 0.04052347 0.04271988 0.04482780
# Extract the forecasted values and the corresponding time period
forecasted_values <- as.vector(forecasts$pred)</pre>
time_period <- seq_along(forecasted_values) + length(ind_prod$yoy_growth)</pre>
# Plot the forecasts
plot(time_period, forecasted_values, type = "l",
    main = "AR Model Forecasts", xlab = "Time", ylab = "Forecasts")
```

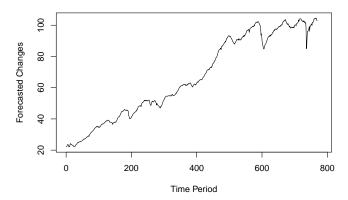
AR Model Forecasts



```
# Average YoY growth rate
mean(ind_prod$yoy_growth[!is.na(ind_prod$yoy_growth)])
```

```
## [1] 0.02552905
```

Forecasted Changes in Original Time Series



The lag order is determined by default in the ar.ols() function using the Akaike Information Criterion (AIC) as the method for order selection. AIC is a model selection criterion that balances model goodness of fit and complexity. Lower AIC values are preferred. The ar.ols() function computes the AIC value for different lag orders and selects the lag order that minimizes the AIC as the optimal lag order for the AR model.

If the AR model is capturing the underlying patterns wel, then as the YoY growth rate of industrial production is stationary, the forecasts are likely to converge towards their long-run average of around 0.026 (2.6%).

Bonus Question

```
# Remove missing values from predicted_values and realized_values
  predicted_values <- na.omit(predicted_values)</pre>
  realized_values <- na.omit(realized_values)</pre>
  # Check and adjust length of predicted_values and realized_values
  if (length(predicted_values) > length(realized_values)) {
    predicted_values <- predicted_values[1:length(realized_values)]</pre>
  } else if (length(realized_values) > length(predicted_values)) {
    realized_values <- realized_values[1:length(predicted_values)]</pre>
  # Compute RMSE
  rmse <- sqrt(mean((predicted_values - realized_values)^2))</pre>
  return(rmse)
# Specify lag orders to compare
lag_orders \leftarrow c(1, 2, 3, 4, 5, 6, 7, 8)
# Specify holdout periods to compare
holdout_periods <- c(6, 12, 24)
# Initialize matrix to store RMSE results
results <- matrix(NA, nrow = length(lag_orders), ncol = length(holdout_periods))
# Iterate over lag orders and holdout periods to calculate RMSE for each model
for (i in 1:length(lag_orders)) {
  for (j in 1:length(holdout_periods)) {
    rmse <- calculate_rmse(lag_orders[i], holdout_periods[j], ind_prod)</pre>
    results[i, j] <- rmse</pre>
 }
}
# Print RMSE results
print(results)
##
                           [,2]
               [,1]
                                       [,3]
## [1,] 0.01083095 0.009360575 0.05575435
## [2,] 0.01020996 0.007806726 0.07005326
## [3,] 0.01017002 0.008160002 0.06738367
## [4,] 0.01018277 0.008012830 0.06806190
## [5,] 0.01018192 0.007866010 0.06822266
## [6,] 0.01007486 0.008113162 0.06834463
## [7,] 0.01016670 0.008140459 0.07030717
## [8,] 0.01027124 0.008264455 0.07092213
```

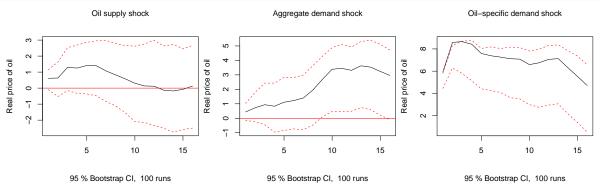
The AR Model with a holdout period of 12 months and a lag order of 2 performs the based among the alternatives as it minimizes the RMSE. When specifying a holdout period of 6 (24) months, we would, based on minimizing the RMSE, select an AR model with 6 (1) lags. Thus, the preferred specification (i.e. optimal lag order) depends on the selected holdout period.

Exercise 2

```
# Estimating the VAR ----
var_model <- VAR(kpdata, p = 24, type = "const")
coefficients <- coef(var_model)
residuals <- resid(var_model)</pre>
```

Replicating Figure 1

```
# Replicating Figure 1 -
# Computing impulse response function with recursive-design wild bootstrap
set.seed(123) # for reproducibility
nboot <- 1000 # number of bootstrap replications</pre>
irf1 <- irf(var_model, impulse = "oil_prod_change", response = "oil_price_real", n.ahead =</pre>
15, boot = TRUE, nboot = nboot, ci = 0.95, boot.type = "rdwb")
irf1$irf$oil_prod_change <- irf1$irf$oil_prod_change * (-1) # Switching signs of shock
irf1$Lower$oil_prod_change <- irf1$Lower$oil_prod_change * (-1) # and of the CIs
irf1$Upper$oil_prod_change <- irf1$Upper$oil_prod_change * (-1)</pre>
irf2 <- irf(var_model, impulse = "econ_act_real", response = "oil_price_real", n.ahead = 15,
boot = TRUE, nboot = nboot, ci = 0.95, boot.type = "rdwb")
irf3 <- irf(var_model, impulse = "oil_price_real", response = "oil_price_real", n.ahead =</pre>
15, boot = TRUE, nboot = nboot, ci = 0.95, boot.type = "rdwb")
plot(irf1, main="Oil supply shock", ylab="Real price of oil")
plot(irf2, main="Aggregate demand shock", ylab="Real price of oil")
plot(irf3, main="Oil-specific demand shock", ylab="Real price of oil")
```

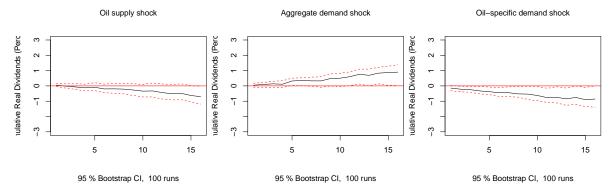


The IRF plots show the responses of the real price of oil to the oil supply shock, the aggregate demand shock, and the oil-market specific demand shock. The oil supply shock is represented by a negative one standard deviation shock, while the other two shocks are positive and would raise the real price of oil. Differently from Kilian and Park (2009), the plots display only the two-standard error bands. The confidence intervals were constructed using a recursive-design wild bootstrap.

Similar to the results in Kilian and Park (2009), the key finding is that the three shocks have different impacts on the real price of oil.

Replicating the Lower Panel of Figure 3

```
# Replicating Figure 3 -
set.seed(123) # for reproducibility
nboot <- 1000 # number of bootstrap replications</pre>
irf4 <- irf(var_model, impulse = "oil_prod_change", response = "div_growth_change_real",</pre>
            cumulative = TRUE, n.ahead = 15, boot = TRUE, nboot = nboot, ci = 0.95,
            boot.type = "rdwb")
irf4$irf$oil_prod_change <- (-1)*irf4$irf$oil_prod_change # Switching signs of shock
irf4$Lower$oil_prod_change <- (-1)*irf4$Lower$oil_prod_change # and of the CIs
irf4$Upper$oil_prod_change <- (-1)*irf4$Upper$oil_prod_change</pre>
irf5 <- irf(var_model, impulse = "econ_act_real", response = "div_growth_change_real",
            cumulative = TRUE, n.ahead = 15, boot = TRUE, nboot = nboot, ci = 0.95,
            boot.type = "rdwb")
irf6 <- irf(var_model, impulse = "oil_price_real", response = "div_growth_change_real",</pre>
            cumulative = TRUE, n.ahead = 15, boot = TRUE, nboot = nboot, ci = (0.95),
            boot.type = "rdwb")
plot(irf4, ylim = c(-3,3), main="Oil supply shock", ylab="Cumulative Real Dividends
(Percent)")
plot(irf5, ylim = c(-3,3), main="Aggregate demand shock", ylab="Cumulative Real Dividends
(Percent)")
plot(irf6, ylim = c(-3,3), main="Oil-specific demand shock", ylab="Cumulative Real Dividends
(Percent)")
```



The above code replicates the cumulative responses of the dividend-growth rate to each shock. Differently from Kilian and Park (2009), the plots display only the two-standard error bands. The confidence intervals were constructed using a recursive-design wild bootstrap. From the plots, we observe that an oil supply shock result in a significant and negative impact on real dividends, which becomes apparent after five months. Positive aggregate demand shocks have a persistent positive effect on real dividends, with the effect being significant in most horizons. Conversely, positive shocks to precautionary demand lead to a persistent negative effect on real dividends, with the impact being significant across all horizons.

Replicating Table 2

```
# Replicating Table 2 -----
table_2 <- fevd(var_model, n.ahead = 1000)$div_growth_change_real %>%
   magrittr::multiply_by(100) %>%
   cbind(horizon = 1:1000) %>%
   as_tibble() %>%
```

```
dplyr::relocate(horizon, .before = oil_prod_change) %>%
  dplyr::filter(horizon %in% c(1, 2, 3, 12, 1000)) %>%
  mutate(horizon = ifelse(horizon == 1000, Inf, horizon))
knitr::kable(table_2)
```

horizon	oil_prod_change	$econ_act_real$	oil_price_real	$div_growth_change_real$
1	0.1999480	0.1635769	1.691515	97.94496
2	0.5509699	0.3610968	2.089788	96.99815
3	0.7611886	0.4840789	2.119006	96.63573
12	2.7957267	6.8331769	4.532478	85.83862
Inf	6.6274588	8.3817303	7.932067	77.05874

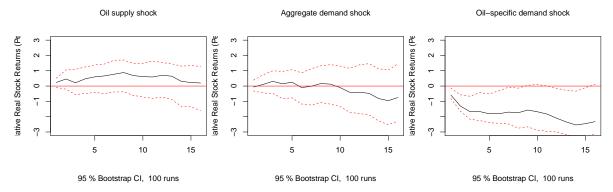
The dividend growth shock has the highest contribution to the forecast error variance, with a percentage of 77.1% at an infinite horizon, while the oil supply shock, oil demand shock, and oil-specific demand shock have smaller contributions. This suggests that the dividend growth shock is the most important factor driving the variability of U.S. real dividend growth over the long run, while oil-related shocks play a relatively smaller role. Specifically, as noted by Kilian and Park (2009), in the long run, shocks affecting the crude oil market account for 23% of the variability in real divident growth, and more than two-thirds of this variability is attributed to oil demand shock. However, the combined impact of these shocks is only 2%.

Creating a Stock Returns Variable

```
# Stock Market Returns --
fred <- read.csv("./assignment1/data/fred.csv")[-1,]</pre>
fred_kp <- fred[169:576,] %>% # only choosing 1.1973 to 12.2006
  as_tibble() %>%
  dplyr::select(S.P.500, CPIAUCSL)
sp_kp <- ts_explode(fred_kp$S.P.500) %>%
 magrittr::set_colnames(c("sp500", "sp500_log", "sp500_mom", "sp500_momlog", "sp500_yoy",
"sp500_yoylag")) %>%
 magrittr::extract(-1,) %>%
  as_tibble()
## Warning in ts_explode(fred_kp$S.P.500): By default, ts_explode() assumes that
## values are ordered from earliest to latest. If your vector is ordered from
## latest to earliest, specify `start_with_latest = TRUE`!
## Warning in ts_explode(fred_kp$S.P.500): Growth rates are given in decimals, not
## in percent!
cpi_kp <- ts_explode(fred_kp$CPIAUCSL) %>%
  magrittr::set_colnames(c("cpi", "cpi_log", "cpi_mom", "cpi_momlog", "cpi_yoy",
"cpi_yoylag")) %>%
  magrittr::extract(-1,) %>%
  as_tibble()
## Warning in ts explode(fred kp$CPIAUCSL): By default, ts explode() assumes that
## values are ordered from earliest to latest. If your vector is ordered from
## latest to earliest, specify `start_with_latest = TRUE`!
## Warning in ts_explode(fred_kp$CPIAUCSL): Growth rates are given in decimals, not
## in percent!
```

```
real_stock_returns <- sp_kp$sp500_momlog-cpi_kp$cpi_mom</pre>
kpdata_2 <- cbind(kpdata, real_stock_returns = 100*real_stock_returns)</pre>
# Re-estimating the model -----
var_model_2 <- VAR(kpdata_2, p = 24, type = "const")</pre>
# Replicating Figure 1 -
set.seed(123) # for reproducibility
nboot <- 1000 # number of bootstrap replications</pre>
irf7 <- irf(var_model_2, impulse = "oil_prod_change", response = "oil_price_real", n.ahead =</pre>
15, boot = TRUE, nboot = nboot, ci = 0.95, boot.type = "rdwb")
irf7$irf$oil_prod_change <- irf7$irf$oil_prod_change * (-1) # Switching signs of shock
irf7$Lower$oil_prod_change <- irf7$Lower$oil_prod_change * (-1) # and of the CIs
irf7$Upper$oil_prod_change <- irf7$Upper$oil_prod_change * (-1)</pre>
irf8 <- irf(var_model_2, impulse = "econ_act_real", response = "oil_price_real", n.ahead =</pre>
15, boot = TRUE, nboot = nboot, ci = 0.95, boot.type = "rdwb")
irf9 <- irf(var_model_2, impulse = "oil_price_real", response = "oil_price_real", n.ahead =
15, boot = TRUE, nboot = nboot, ci = 0.95, boot.type = "rdwb")
plot(irf7, ylim = c(-6,12), main="Oil supply shock", ylab="Real price of oil")
plot(irf8, ylim = c(-6,12), main="Aggregate demand shock", ylab="Real price of oil")
plot(irf9, ylim = c(-6,12), main="Oil-specific demand shock", ylab="Real price of oil")
                  Oil supply shock
                                                                               Oil-specific demand shock
                                               Aggregate demand shock
       9
                                       9
                                                                       9
    Real price of oil
                                    price of oil
                                                                    Real price of oil
       5
                        10
                                15
                                                        10
                                                                15
                                                                                        10
                                                                                                15
              95 % Bootstrap CI, 100 runs
                                              95 % Bootstrap CI, 100 runs
                                                                              95 % Bootstrap CI, 100 runs
```

Replicating the Upper Panel of Figure 3



The upper panel of figure 3 displays the cumulative impulse responses of the created variable real stock returns to the three demand and supply shocks in the crude oil market. The results are qualitatively similar to the ones from Kilian and Park (2009). Oil supply shocks still have little effect on cumulative U.S. stock returns while a positive shock in global demand has a smaller effect, and turns negative sooner. Similar to the results in Kilian and Park (2009), an increase in the precautionary demand for oil causes persistently negative U.S. stock returns.

Replicating Table 1

horizon	oil_prod_change	$econ_act_real$	oil_price_real	${\rm real_stock_returns}$
1	0.4825399	0.0341007	3.012370	96.45241
2	0.8239363	0.2569623	6.751464	92.05562
3	1.2693408	0.5531023	7.658378	90.37831
12	2.3320103	2.9279609	8.213784	83.77642
Inf	5.9049610	6.9272703	9.827450	72.04405

Table 1 reveals that, similar to the results in Kilian and Park (2009), in the short run, the impact of oil supply,

aggregate demand, and oil-specific demand shocks is very small with around 96% of the variation in U.S. stock returns associated with other shocks than those in the global crude oil market. However, we also see that as the horizon increases, the explanatory power of the oil shocks increases.