

$$\text{Var } \bar{X} = \text{Var} \left[\frac{1}{n} \sum_{i=1}^n X_i \right] = \frac{1}{n^2} \text{Var} \sum_{i=1}^n X_i =$$

$$= \frac{1}{n^2} n \sigma^2 = \frac{\sigma^2}{n}$$

$$\sqrt{\text{Var } \bar{X}} = \frac{\sigma}{\sqrt{n}}$$

Confidence intervals

An interval I_θ that covers θ with probability $1 - \alpha$ is called a confidence interval for θ with conf. level $1 - \alpha$.

Let $\mathbf{x} = (x_1, x_2, \dots, x_n)$ be a sample from distribution $X \sim P_{\mathbf{x}}$ that depends on the parameter θ and $\mathbf{X} = (X_1, \dots, X_n)$ corresponding theoretical sample.

Given definition means that there are functions of theoretical sample $a_1(\mathbf{X})$ and $a_2(\mathbf{X})$ so that

$$P(a_1(\mathbf{X}) \leq \theta \leq a_2(\mathbf{X})) = 1 - \alpha$$

Then values of these RVs $a_1(\mathbf{x})$ and $a_2(\mathbf{x})$ are called lower and upper conf. limits

$$I_\theta = (a_1(\mathbf{x}), a_2(\mathbf{x}))$$

is conf. interval at the conf. limit $1 - \alpha$

Requirements for conf. interval:

- as narrow as possible
- conf. level as high as possible (closer to 1)

A very wide conf. interval gives the message that there is a great deal of uncertainty concerning the value of what we are estimating

if conf. limits $a_1(x)$ and $a_2(x)$ are finite, the interval is then called two sided, otherwise one-sided $(-\infty, a_2)$, (a_1, ∞)

Classical conf. intervals are 90%, 95%, 99%

Example

If $X \sim N(\mu, \sigma^2)$ then from our prev. RVs we are interested in such λ_α

$$P(X < \mu + \lambda_\alpha \sigma) = 1 - \alpha$$

$$P(X < \mu + \lambda_\alpha \sigma) = P\left(\frac{X - \mu}{\sigma} < \lambda_\alpha\right) = \Phi(\lambda_\alpha) = 1 - \alpha$$

So λ_α is the $(1 - \alpha)$ -quantile of $N(0, 1)$ and is also called the complement α -quantile

Some logic applies to:

$$P(\mu - \lambda_{\alpha/2} \sigma < X < \mu + \lambda_{\alpha/2} \sigma) = 1 - \alpha$$

Example 2

$$X \sim N(\mu, \sigma^2)$$

$$P(\mu - \lambda_{\alpha/2} \sigma < X < \mu + \lambda_{\alpha/2} \sigma) = 1 - \alpha$$

where $\lambda_{\alpha/2}$ is $N(0, 1)$ complement
 $\alpha/2$ quantile

Remarks

Example 2

Single Random Sample . Conf. interval for the Mean

Let x_1, \dots, x_n be a random sample from $N(\mu, \sigma^2)$. We want to construct a conf. interval for the mean μ

Theorem

Let x_1, \dots, x_n be a random sample from $N(\mu, \sigma^2)$ where μ is unknown. Then:

$$I_\mu = \bar{X} \pm \lambda_{\alpha/2} \frac{\sigma}{\sqrt{n}} \quad \text{if } \sigma \text{ is known}$$

$$I_\mu = \bar{X} \pm t_{\alpha/2, f} \frac{s}{\sqrt{n}} \quad \text{if } \sigma \text{ unknown}$$

where s is stand. deviation of the sample, $\lambda_{\alpha/2}$ and $t_{\alpha/2, f}$ are $\alpha/2$ complement quantiles of $N(0, 1)$ and $t(f)$, $f = n - 1$

Chi-square and t-distribution

A random variable X with a density function of the form

$$k x^{\frac{f}{2}-1} e^{-x/2} \quad (x \geq 0)$$

is said to have a χ^2 distr. with f degrees of freedom and k being a normalizing constant

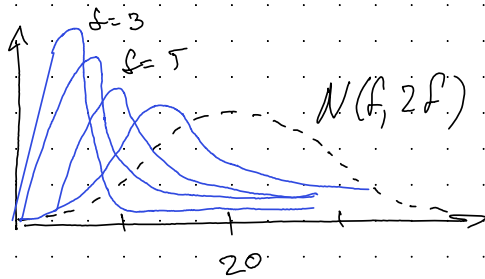
By setting $f = 1, 2, \dots$ we obtain a whole family of distributions.

All distributions in the family are skewed, but the larger is f the more symmetric is the distribution.

if $X \sim \chi^2(f)$ then $EX = f$ and $\text{Var}X = 2f$

Remark

if f is large, then $\chi^2(f) \approx N(f, 2f)$



Chi-square distribution

if the random variables X_1, \dots, X_n are independent and have distributions $\chi^2(f_1), \chi^2(f_2), \dots, \chi^2(f_n)$ respectively then

$$Y = \sum_{i=1}^n X_i \sim \chi^2\left(\sum_{i=1}^n f_i\right)$$

Theorem

If X_i $i=1 \dots n$ are independent random var-s and $X_i \sim N(0, 1) \Rightarrow$

$$\sum_{i=1}^n X_i^2 \sim \chi^2(n)$$

and

$$\sum_{i=1}^n (X_i - \bar{X})^2 \sim \chi^2(n-1)$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

Student's t-distribution

Def a random var X with density function

$$k \left(1 + \frac{x^2}{f}\right)^{-(f+1)/2}, \quad (-\infty < x < \infty)$$

where

$$k = \frac{\Gamma\left(\frac{f+1}{2}\right)}{\sqrt{f\pi} \Gamma\left(\frac{f}{2}\right)}, \quad f \in \{1, 2, \dots\}$$

is said to have a t-distrib.
with f degree of freedom

Property

If $X \sim N(0, 1)$ and $Y \sim \chi^2(f)$

Where X and Y are independent

$$Z = \sqrt{f} \frac{X}{\sqrt{Y}} \sim t(f)$$

Theorem

if $X_i \sim N(\mu, \sigma^2)$ are independent random
var-s $i = 1, 2, \dots, n$ then

$$\sqrt{n} \frac{\bar{X} - \mu}{S} \sim t(n-1)$$

where $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

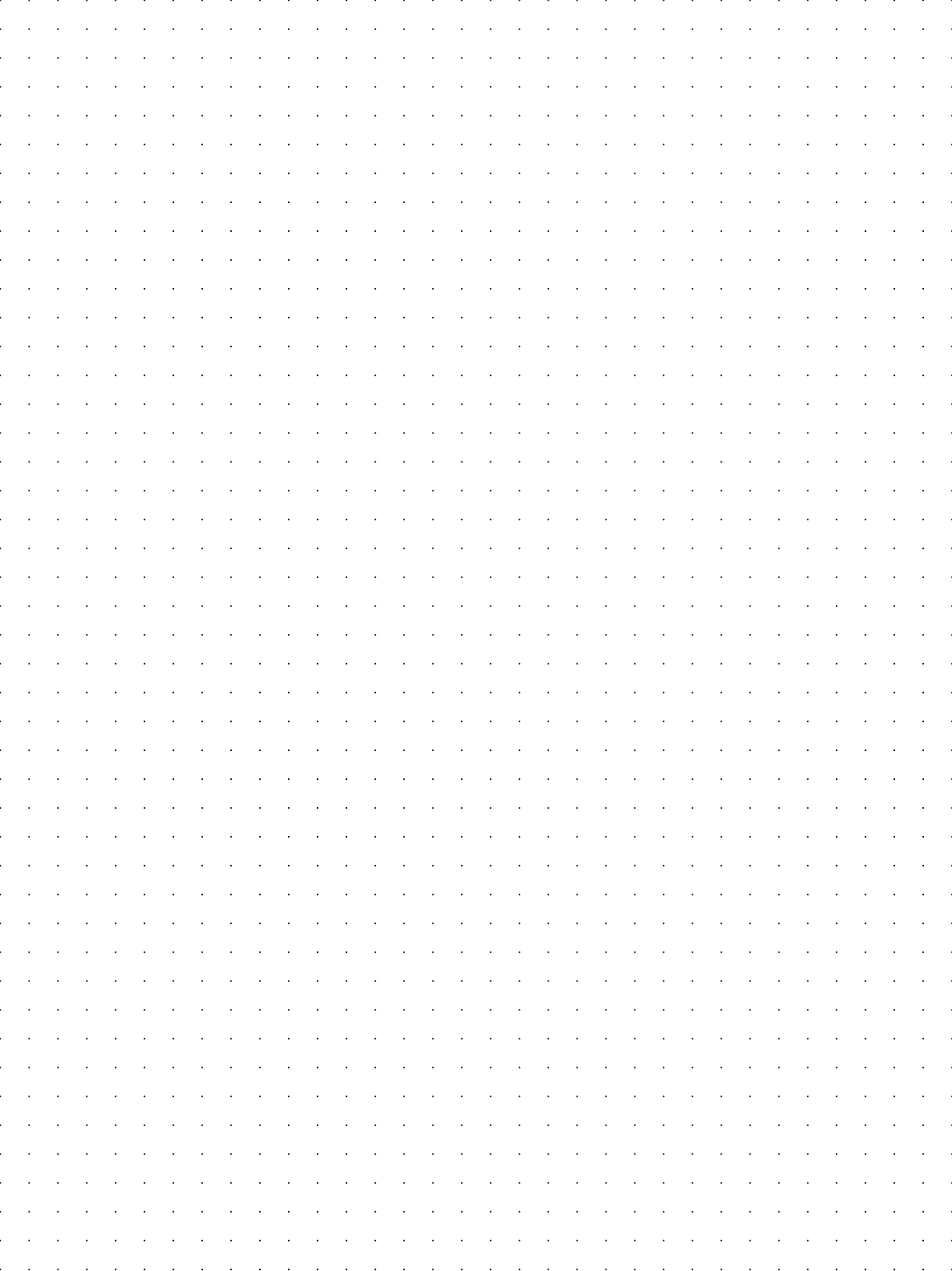
$$S = \left[\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 \right]^{1/2}$$

Conf. Interval for the Mean

$$P\left(\bar{X} - t_{\alpha/2, f} \frac{s}{\sqrt{n}} < \mu < \bar{X} + t_{\alpha/2, f} \frac{s}{\sqrt{n}}\right) = 1 - \alpha$$

in other words

$$P\left(-t_{\alpha/2, f} < \sqrt{n} \frac{\bar{X} - \mu}{s} < t_{\alpha/2, f}\right) = 1 - \alpha$$



Two samples. Conf. Interval for difference between mean

The following model is often employed in practice: Two indep. samples have been collected

x_1, \dots, x_n from $N(\mu_1, \sigma_1^2)$

y_1, \dots, y_n from $N(\mu_2, \sigma_2^2)$

Theorem

Let x_1, \dots, x_{n_1} and y_1, \dots, y_{n_2} be independent random samples from $N(\mu_1, \sigma_1^2)$ and $N(\mu_2, \sigma_2^2) \Rightarrow$

$$I_{\mu_1 - \mu_2} = \bar{x} - \bar{y} \pm \lambda_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

if σ_1, σ_2 are known

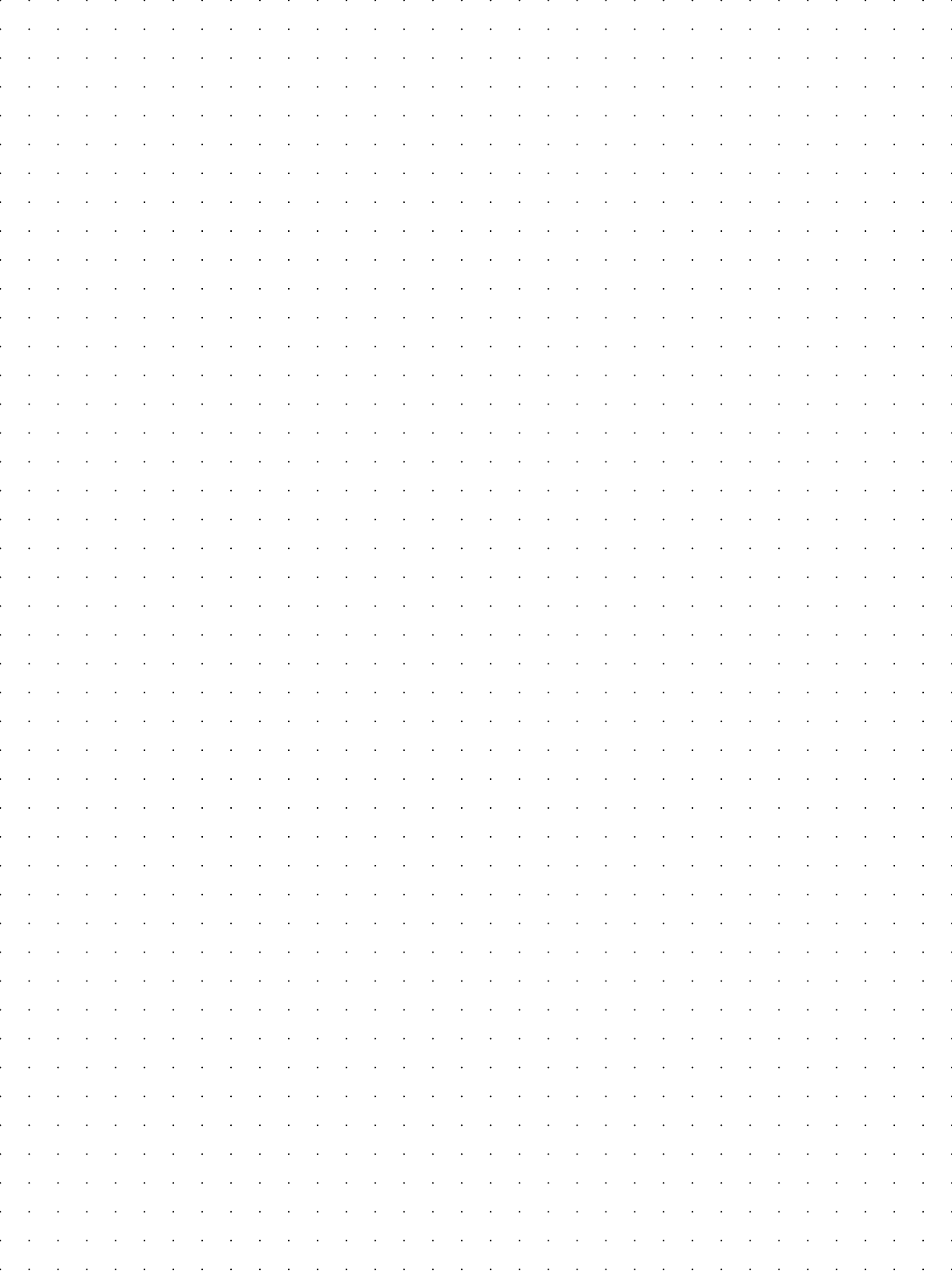
$$I_{\mu_1 - \mu_2} = \bar{x} - \bar{y} \pm t_{\alpha/2, f} S \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

if $\sigma_1 = \sigma_2 = \sigma$ are unknown

are two-sided conf. intervals for $\mu_1 - \mu_2$ with conf. level

1- α where $S^2 = \frac{\sum_{i=1}^{n_1} (x_i - \bar{x})^2 + \sum_{i=1}^{n_2} (y_i - \bar{y})^2}{n_1 + n_2 - 2}$

and $f = n_1 + n_2 - 2$



Paired Samples

Dependent (paired) samples - repeated measurements on the same object (for example before and after using some treatment medication etc)

Before: X_1, \dots, X_n ; $X_i \leftarrow X_i \sim N(\mu_1, \sigma_1^2)$

After: Y_1, \dots, Y_n ; $Y_i \leftarrow Y_i \sim N(\mu_1 + \Delta, \sigma_2^2)$

derive a new variable $W_i =$

$= Y_i - X_i \sim N(\Delta, \sigma_w^2)$ where

$$\sigma_w^2 = \sigma_1^2 + \sigma_2^2$$

Theorem (from previous section) is used (where both of the parameters are unknown):

$$I_\mu = \bar{X} \pm t_{\alpha/2, n-1} \frac{s}{\sqrt{n}} \quad \text{if } \sigma \text{ unknown}$$

Theorem