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## LECTURE 1: MAXIMUM LIKELIHOOD ESTIMATION

SET UP: You observe  $X$  (possibly a vector)  
 from a parametric family of distributions:  
 $X$  has pdf (or pmf if discrete)

$$f(x; \theta)$$

where  $f$  is completely specified, although it  
 depends on the parameter  $\theta$  ( $\theta$  possibly  
 a vector)

EXAMPLE:  $X = (X_1, \dots, X_n)$  where each  $X_i$  has  
 a Normal  $(\mu, \sigma)$  density, and the  $X_i$ 's are independent

$$f_{X_i}(x_i) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(x_i - \mu)^2\right)$$

The joint density of  $(X_1, \dots, X_n)$  is obtained as a  
 product, by independence:

$$f_X(x_1, \dots, x_n) = (2\pi\sigma^2)^{-n/2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2\right)$$

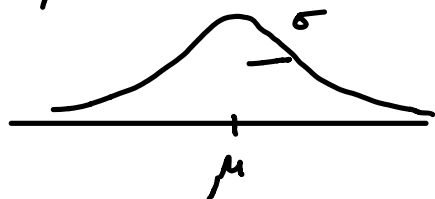
The parameters,  $\mu, \sigma^2$ , are typically unknown

UNKNOWN, NOT OBSERVED

OBSERVED

$\mu, \sigma^2$

$X_1, \dots, X_n$



BASIC PROBLEM: ESTIMATE THE PARAMETER  $\theta$  ②  
USING THE DATA  $x_1, \dots, x_n$

FIND A "GOOD" FUNCTION  $T(x_1, \dots, x_n)$   
SO THAT  $T(x_1, \dots, x_n)$  IS A REASONABLE  
GUESS AT  $\theta$ ; SUCH FUNCTIONS ARE ESTIMATORS

A METHOD THAT OFTEN PRODUCES REASONABLE  
ESTIMATORS IS MAXIMUM LIKELIHOOD:

THE LIKELIHOOD PRINCIPLE SAYS TO FIND  
VALUE  $\theta$  WHICH MAXIMIZES

$$L(\theta; x) = f(x; \theta)$$

HERE, the observed value  $x$  of  $X$  is held fixed,  
and the above is optimized over  $\theta$

Example  $x_1, \dots, x_n$  IID  $N(\mu, \sigma^2)$   
(Independent, Identically Distributed)

NOTE: maximizing  $L(\theta)$  is equivalent to maximizing  
 $l(\theta) = \log L(\theta)$ , since  $\log$  is monotone function

$l$  is often easier to work with, especially  
when  $f$  has a product form, as it  
does in the case of IID

$$l(\mu, \sigma^2) = -\frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 - \frac{n}{2} \log 2\pi \quad (3)$$

The final term is irrelevant since it does not depend on either  $\mu$  or  $\sigma^2$

Since everything is smooth, we can optimize by looking for critical points

$$\frac{\partial l}{\partial \mu} = \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)$$

$$\frac{\partial l}{\partial \sigma^2} = -\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum (x_i - \mu)^2$$

The simultaneous solutions are

$$\mu = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}$$

$$\sigma^2 = \frac{1}{n} \sum (x_i - \bar{x})^2$$

Note that  $\bar{X} \sim N(\mu, \sigma^2/n)$



It has many nice properties:

$$E[\bar{X}] = \mu, \quad \text{also } \bar{X} \rightarrow \mu \text{ as } n \rightarrow \infty$$

The first property is called UNBIASED

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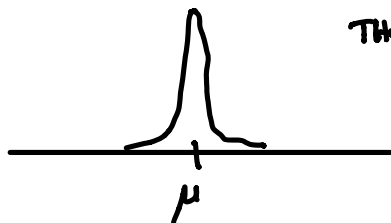
DEFN Suppose  $X \sim f(x; \theta)$ . Let  $T = T(X)$  be an estimator of  $\theta$  (That is,  $T$  is any function of  $X$  which does not depend on  $\theta$ ).  $T$  is called UNBIASED if

$$E[T] = \theta$$

Note: The distribution of  $T$  depends of  $\theta$  since  $T$  is a function of  $X$ , and  $X \sim f(x; \theta)$ . So  $E[T]$  is in general a function of  $\theta$ .  $T$  is unbiased if this function is the identity.

This means the distribution of  $T$  is centered about the quantity we are trying to recover. However,  $T$  may still be spread out around  $\theta$ .

Ex Normal as above.  $\bar{X}$  is unbiased, and  $\text{Var}(\bar{X}) = \sigma^2/n$ , so the dispersion of the distribution about the unknown  $\mu$  is shrinking as  $n \rightarrow \infty$ .



THIS THE ACCURACY  
OF THE ESTIMATOR  
INCREASES WITH  $n$

Exercise: Let  $X_1, \dots, X_n$  be IID Poisson( $\lambda$ ) (5)

Find the Maximum Likelihood Estimator (MLE)

of  $\lambda$ . Note

$$p(k|\lambda) = \frac{e^{-\lambda} \lambda^k}{k!}, \quad k=0,1,2,\dots$$

BE CAREFUL:

Let  $X_1, \dots, X_n$  be IID UNIF[0,  $\theta$ ]

$$f(x; \theta) = \begin{cases} \frac{1}{\theta} & \text{if } 0 \leq x \leq \theta \\ 0 & \text{otherwise} \end{cases}$$

$$= \frac{1}{\theta} I_{[0, \theta]}(x)$$

$$L(\theta; x_1, \dots, x_n) = \prod_{i=1}^n \frac{1}{\theta} I_{[0, \theta]}(x_i)$$

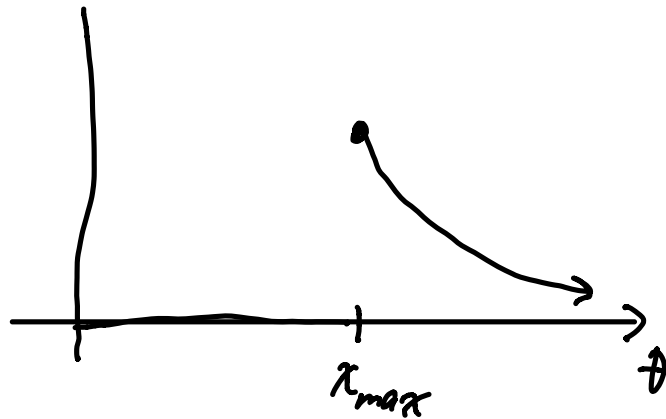
$$= \theta^{-n} \prod_{i=1}^n I_{[0, \theta]}(x_i)$$

$$= \begin{cases} \theta^{-n} & \text{if } 0 \leq x_1, \dots, x_n \leq \theta \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \theta^{-n} & \text{if } x_{\max} \leq \theta \\ 0 & \end{cases}$$

$$L(\theta) = \theta^{-n} I_{[\chi_{\max}, \infty)}(\theta)$$

⑥



The value maximizing  $L(\theta)$  is  $\chi_{\max}$

Exercise: Find  $c$  so that  $c\chi_{\max}$  is unbiased estimator of  $\theta$

Note that  $\chi_{\max}$  itself is BIASED:

