

A Correlated Signals Algorithmic Game Theory Model of Litigation Bargaining

Michael Abramowicz
abramowicz@law.gwu.edu
George Washington University Law School

March 18, 2025

Abstract

This article adapts a linear programming algorithm to model litigation games in which each party receives a signal of the quality of the litigation. The games are represented in extensive-form game trees of approximately 45,000 nodes each. The 1,250 permutations of game parameters solved allow for two-sided asymmetric information, litigant decisions whether to file and answer, options to quit if settlement fails, asymmetric risk aversion, asymmetric costs, and asymmetric information quality, among other variables, in cases concerning either damages or liability. Welfare variables of interest, concerning accuracy and expenditures, can be directly measured. The results illustrate how fee shifting can change bargaining dynamics and social welfare, assuming litigants play rational equilibrium strategies. In simulations with risk neutrality and mild risk aversion, the British rule has a minimal effect on the chance that a case will be contested at all but often leads to a lower incidence of trial, in contrast to the conventional wisdom. Increasing the cost of litigation in the model adversely affected social welfare but generally did not meaningfully change the advisability of fee shifting. Asymmetries in risk preferences, information, and litigation costs can significantly affect case outcomes, but do not have unambiguous implications for the fee-shifting debate. To confirm the robustness of its conclusions, multiple equilibria are generated, first for default parameter values using the baseline game tree size and then for all models using a smaller game tree. Numerous visualizations of equilibria are offered, and over 20,000 more are included in the Supplemental Materials.

1 Introduction

REMAINING ISSUES: (1) Supplemental Information must contain the detailed analysis of American vs British; must check the description here. (2) Same

mini axes on Errors and expenditures. (3) Note insert cite in introduction. (4) Insertion in conclusion (5) Fix description of disposition varying noise multiplier. Get rid of Figure 21 (Errors and expenditures based on player noise levels) or rerun it. Update up to Costs Structure (which has up-to-date data).

A natural assumption for settlement bargaining models is that the plaintiff and defendant receive correlated signals of the quality of the litigation. In a tort case, for example, each party may possess a combination of shared and private information about the events leading to an accident, and so each party can be understood as receiving a noisy signal of the true merits of the case. Even in a case in which facts are unambiguous and the only dispute is as to the law, each party will assess the case based on its lawyer’s analysis and hunches, and the parties’ estimates of case strength based on such legal advice will be correlated but not necessarily identical. Most of the literature modeling settlement bargaining, however, does not feature correlated signals. In the classic models of Landes (1971) [18], Posner (1973) [22], and Gould (1973) [11], parties’ estimates of litigation quality are entirely exogenous. Later Bayesian models, such as Bebchuk (1984) [3] and Polinsky and Rubinfeld (1998) [20], assume that one party knows the quality of litigation while the other party knows the distribution of all cases but does not receive a signal of the particular case’s quality.

To substitute for correlated information, some models grant the plaintiff some information and the defendant some other information. The aggregate of these pieces of information then determines case quality. For example, Daughety and Reinganum (1994) [8] assign one party information on liability and another party information on damages. Similarly, in Friedman and Wittman (2006) [9], Dari-Mattiacci and Saraceno (2020) [6], and Abramowicz (2025) [1], the plaintiff and defendant each receive independent random signals from a known distribution, and the true value of the case is an average of those signals. Thus, in these models, each of the plaintiff’s and defendant’s signals correlates with the merits, but unrealistically, the signals do not correlate with one another. The game thus becomes like a game of poker in which there are no common cards. It is difficult to develop intuitions about whether the optimal strategy and resulting equilibria in such a game will resemble those in a game of correlated signals. One might intuitively expect discovery in litigation to produce closely correlated signals and that such correlated signals in turn would lead to high rates of settlement and relatively limited opportunities for strategic negotiation behavior.

In only one pair of articles do the plaintiff and defendant each estimate the quality of the litigation based on correlated signals. Lee and Klerman (2016) [19] and Klerman, Lee, and Liu (2018) [17] use correlated signals to update the canonical selection model of Priest and Klein (1984) [23]. In these updated models, the true value of the litigation is Y (distributed uniformly over the entire real line), and the plaintiff and defendant receive signals of $Y + \epsilon_p$ and $Y + \epsilon_d$, respectively, where each ϵ is a normally distributed random variable. The papers differ in that the older adopts the Landes-Posner-Gould assumption that the cases settle whenever there exists a settlement amount at which each party prefers to settle, while the more recent explicitly models the bargaining struc-

ture. The later paper’s bargaining model is that of Chatterjee and Samuelson (1983) [5], in which each party separately makes an offer and settlement occurs at the midpoint if the plaintiff’s offer is less than or equal to the defendant’s.

These papers represent significant advances in the realism of litigation models. But they have several limitations that make it difficult to assess how correlated signals would impact issues other than the selection of cases for litigation. First, even in the more recent paper, the litigation quality is constrained to a range where each party has a credible threat to take the case to trial (pp. 392-93). In effect, the plaintiff always files, the defendant always answers, and neither party ever quits before trial. Yet we have known since Shavell (1982) and P’ng (1983) about the importance of filing and answer decisions and since Huang (2004) [13] and Hubbard (2016) [14] that even if it is rare for the plaintiff or defendant, respectively, to quit just before trial, the possibility of such quitting may have significant strategic ramifications. Second, the entire cost of litigation is thus borne all at once, rather than sequentially. Third, although case quality varies, there is a simple cutoff, above which the plaintiff has a high fixed probability of winning and below which the plaintiff the probability is fixed at a low value. Fourth, these papers consider bargaining only under the American rule in which each party bears its own costs and ignore the British rule in which the loser bears the other party’s fees.

In this article, I relax these assumptions through use of a computational model that identifies Nash equilibria in large extensive form game trees, applying the linear programming algorithm identified in von Stengel, van den Elzen, and Talman (2002) [26]. Abramowicz (2025) applies this algorithm to model settlement bargaining but its models do not include correlated signals of litigation quality. Instead, Abramowicz seeks to establish a proof of concept for algorithmic game theory modeling by using a setup similar to that of Dari-Mattiacci and Saraceno, in which the parties have independent signals that aggregate to determine case quality. The largest models in that paper included game trees of 16,111 nodes each; the increased complexity associated with correlated signals means that game trees in this article are as large as 45,211 nodes. As in Abramowicz (2025), the game trees identified in this article represent perfect Bayesian Nash equilibria. That is, the plaintiff and defendant each assign a mixed strategy to each information set, and each player’s mixed strategy is a best response for each player given that player’s information. That is, neither player could improve its strategy even by a small ϵ , holding the other player’s strategy constant. The mixed strategies are exact, defined by rational numbers rather than floating point approximations. The algorithmic game theory model thus follows in the Bayesian tradition of recent settlement bargaining scholarship, but with enhanced strategic complexity, including multiple levels of case quality, multiple signal levels, variable levels of fee-shifting, and litigant options to file, answer, and quit before trial.

This approach has three principal payoffs. First, the models allow assessment, in a theoretical but more realistic setting than past models, of how much strategic behavior one might rationally expect with closely correlated signals. One might expect risk-neutral actors to engage in relatively little strategic be-

havior. The initial expense of filing or answering will tend to remove cases where a party’s information indicates that it has an especially weak case, and both parties benefit from achieving settlement rather than trial. Surprisingly, even when parties have strong (though imperfect) signals of litigation quality, strategic behavior can be sufficiently high as to make settlement low or nonexistent. Many cases, of course, do settle, though as Chiang and Klerman (2023) [4] demonstrate, settlement rates vary highly around the world. Relatively high rates of settlement relative to the model must be explained by a factor such as nonequilibrium play or by model parameters such as high litigation costs or risk aversion.

Second, by permitting variation of these parameters, the model allows for an analysis of the degree to which changes in parameters might affect social welfare. For example, although models since Landes (1973) have recognized that settlement and trial costs are positively related, they have not allowed for analysis of the social welfare consequences of the relative expense of litigation. Increases in settlement because of high trial costs in principle might even be sufficient to offset those cost increases. The effects on litigation accuracy, meanwhile, seem indeterminate. On one hand, high trial costs might mean that settlement levels can deviate a great deal from expectations; on the other, greater settlement might produce more consistent outcomes. In fact, the correlated signals models show that higher trial costs produce worse results, both in terms of litigation expenditure and in terms of accuracy. This suggests that reforms to lower the cost of litigation may be worthwhile even if they reduce accuracy to some degree. Meanwhile, the richness of the modeling also allows assessment of how asymmetry in factors such as cost and risk aversion may affect welfare.

Third, the models provide the first comparison of the American rule and British rule with correlated signals. A reasonable conclusion from articles such as [INSERT CITEs] is that the British rule deters frivolous suits (and defenses), thus leading to fewer suits filed (and answered), but because litigation generally selects for mutually optimistic parties, the British rule might decrease settlement. The results here are considerably more nuanced and in many cases push in the opposite direction with respect to both filing and settlement. Although a potential danger of a computational model is that it might be a black box, in fact we can observe precisely how adding fee-shifting changes players’ strategic incentives with respect to different decisions and then in turn how changes in those decisions affect opponents’ incentives. A tentative conclusion is that any effects of fee-shifting on either expenditures or inaccuracy are likely to be relatively modest. Even with fee-shifting, plaintiffs will often bring cases with relatively low quality, and if litigation fees are sufficiently high to deter the vast majority of disputes, results may still be highly inaccurate.

2 Litigation as an Extensive Form Game

This section frames the fee-shifting problem in explicitly game theoretic terms by describing the extensive-form game tree, and it then offers an overview of

the algorithm the article uses to solve it. The last subsection reports on the performance of the algorithm in practice when applied to the game tree.

2.1 The Game Tree

To use a computational approach that finds equilibria from extensive form game trees, the model must encompass a discrete number of values representing the strength of the plaintiff’s case, the signal that each party receives about the strength of the case, and the settlement values that each party can offer the other. In most of the simulations, we will allow for $n_{LS} = 10$ quality levels on the issue of liability, $n_{LS}^P = n_{LS}^D = 10$ signals of liability strength for each of the plaintiff and defendant on the issue of liability (damages will be considered separately later), and $n_o = 10$ offers for each party.

A model dependent on discrete variables is admittedly less general than a model based on continuous quantities, and a virtue of the mathematical literature on settlement bargaining is that some variables can be modeled continuously. In Bebchuk (1984), for example, the defendant knows the probability that the court will find liability, which can be any value from a mutually known probability density function. But not all variables in mathematical settlement models are continuous. In their review of the settlement bargaining literature, Daughety and Reinganum (2012) [7] include both two-types and continuous types models. Even models that offer continuous types in one respect feature only two types in another respect. For example, Klerman, Lee, and Liu (2018) [17] allow each party to obtain a continuous signal of the quality of litigation, but the quality itself can take on only one of two values. Moreover, the one-sided asymmetric information models of litigation can be considered one-type models, at least with respect to the party without information. Thus, the conventions of the existing literature do not mandate continuous variables. The 10 quality levels and signals of liability strength for each party allow a considerably greater degree of granularity than two-types models.

To describe the litigation game, however, it is easier to focus on a much smaller game tree, where $n_{LS} = n_{LS}^P = n_{LS}^D = n_o = 2$. This smaller game tree, including probabilities corresponding to a perfect equilibrium of the game (produced to three decimal places), is presented in Figure 1. This is presented in full at the expense of legibility to illustrate that even with these low parameter values, the game tree is fairly large. (The full figure is available in the Supplemental Materials.) Thus, we will simplify further, focusing on particular parts of the game tree.

Figure 2 illustrates the beginning of the litigation game. The first decision in the game belongs neither to the plaintiff nor the default, but to Chance, abbreviated C, a nonstrategic player. This common device, known as the Harsanyi transformation after Harsanyi (1967) [12], effectively creates a random number generator, transforming a game of perfect but incomplete information (where each player knows all moves that have been made but does not know payoffs) into one of complete but imperfect information (where a player may not know of some moves, such as Chance decisions revealed only to another player, but

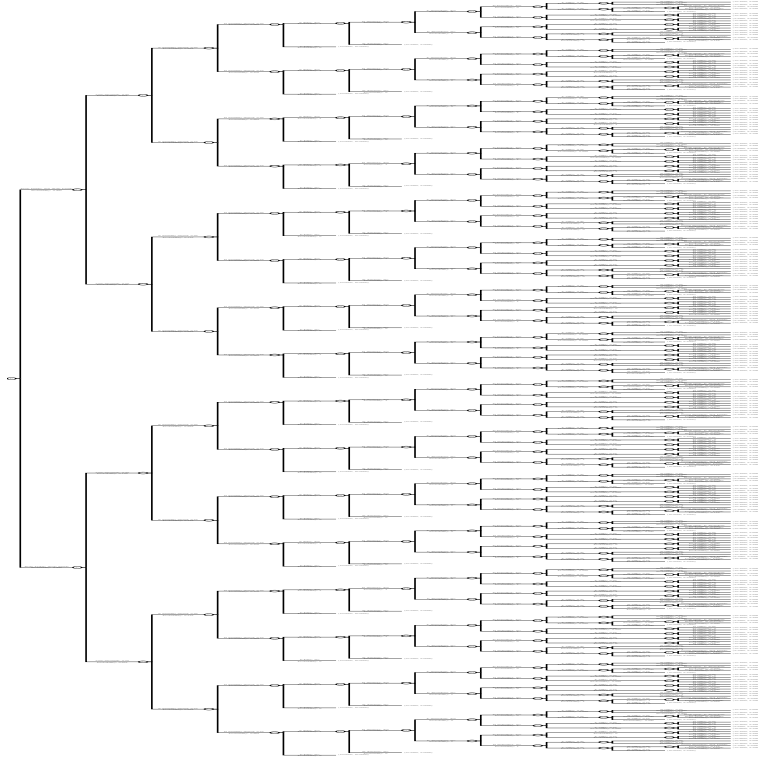


Figure 1: A bird's eye view of a small game tree

knows the complete game structure). Chance's initial decision in this game is whether the case is one in which the defendant is truly liable or one in which the defendant is not truly liable. Chance chooses true liability with a probability p_{TL} , here $\frac{1}{2}$.

Explicit modeling of true liability will allow us ultimately to measure the accuracy of outcomes explicitly. A less ambitious approach would be to define accuracy solely based on whether settlements mimic resulting judgments, but that would ignore the imperfections of judgments themselves. A more ambitious approach, to be pursued in future work, would be to model the emergence of disputes endogenously in some particular legal context, as in Hylton [15], in which a potential tort defendant initially makes a decision about taking care. The approach here, however, is more general, and by varying Chance's probabilities, we will be able to change the distribution of true liability.

Chance then immediately faces a second decision, what the strength of the plaintiff's case on liability should be. Using the baseline parameters (to be described below), when the defendant is not truly liable, there is an 85.1% chance that the liability strength will be low, and when the defendant is truly liable, there is an 85.1% chance that the liability strength will be high. This

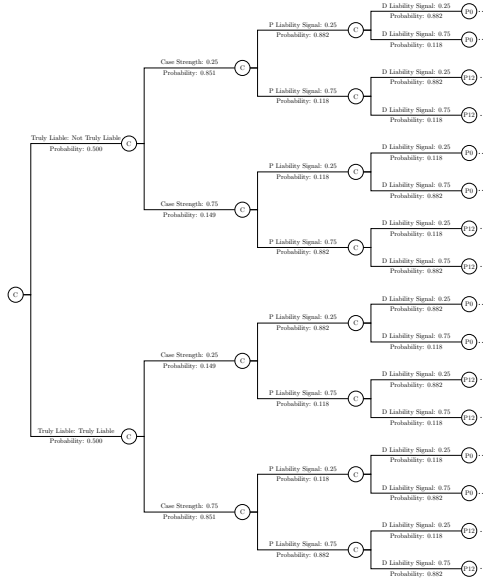


Figure 2: The beginning of the small game tree

approach recognizes that although there will generally be a correlation between true liability and the strength of a plaintiff's case, there may be cases in which a plaintiff has a weak case despite true liability, or a strong case despite absence of true liability. By separating true liability from the strength of the liability case, we can assess how the strength of the legal system, defined loosely here as the correlation of true liability with case strength, affects the case for fee-shifting.

Each case has a liability strength that is a noisy signal of the true liability value. Then, each party in turn receives a noisy signal of the liability strength. As illustrated in Figure 2, if the liability strength is low, then the plaintiff is more likely to receive the low signal than the high signal, and vice versa if the liability strength is high. The defendant's signal is determined independently of the plaintiff's, but the two are correlated, because each is determined in part by the true liability level. Although the game tree imposes an ordering, making it appear that the Chance decision determining the plaintiff's signal occurs before the Chance decision determining the defendant's signal, that is irrelevant and has no effect on equilibrium determination, because the defendant does not learn of the plaintiff's signal and vice versa.

Figure 3 provides a simple visualization of the probabilities of different liability strengths given the true liability status, and of the probabilities that a player will receive different liability signals. These probability values correspond to the baseline values that this article will use in its simulations. The degree of noise is admittedly somewhat arbitrary. A future project would be to calibrate signals based on empirical data for some particular category of cases, but the observation of Gelbach (2018) [10] that many very different models may share

a common reduced form suggests that this will be challenging. An informal assessment based on Figure 3 might conclude that the assumed noise values chosen correspond to a legal system that is pretty good. It is relatively rare, for example, for a truly liable case to have a below-average liability strength signal, or for a party to receive a signal indicating likely liability when the actual liability strength is below average. This point highlights that the highly strategic play and relatively low settlement rates that we will observe with risk neutrality are not the result of weak party information. We will, however, also consider the effects of both stronger and weaker information, as well as of asymmetric information.

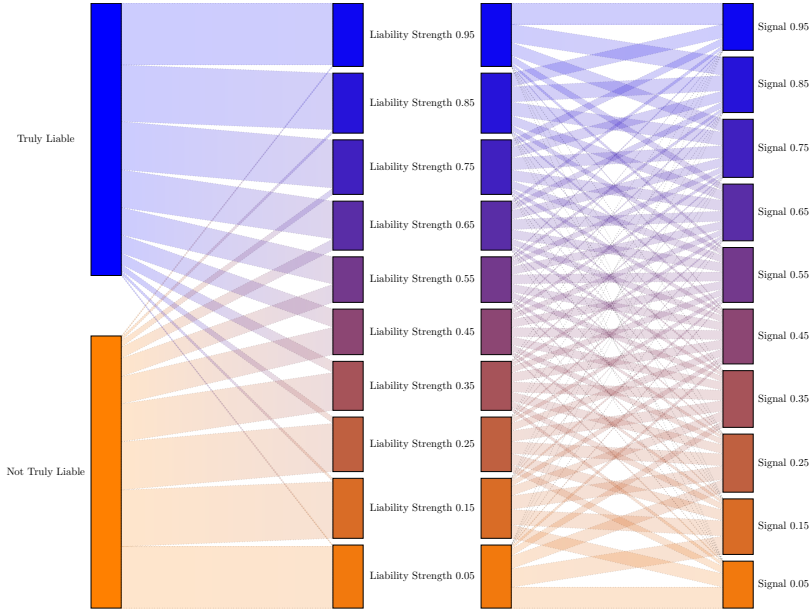


Figure 3: Liability strength given true liability and party's signal given liability strength

A litigation game equilibrium can be calculated for any distribution of signals given true values, so a modeler could simply exogenously specify the conditional probabilities illustrated in Figure 3. It is useful, however, to encapsulate the degree of noise into simple parameters. The probabilities of liability strengths in Figure 3 depend on a parameter σ_{LS} , which represents the standard deviation of a normal distribution from which noise values are drawn. A noise value is summed with a value taken from the true liability Bernoulli distribution, so the greater σ_{LS} , the more attenuated the connection between whether a defendant is truly liable and what the liability strength is. The signals of liability strength that the plaintiff and defendant receive depend on the similar parameters σ_{LS}^P and σ_{LS}^D . In the baseline simulations represented in Figure 3, $\sigma_{LS} = 0.35$ and $\sigma_{LS}^P = \sigma_{LS}^D = 0.20$.

More formally, the unit interval is divided into n_{LS} equal segments, corresponding to the n_{LS} values. The probability of each liability strength level is equal to the relative probability that the sum falls into the corresponding segment, ignoring sums less than 0 or greater than 1. That is, let $TL = \{0, 1\}$ represent the possible true liability values, corresponding to not truly liable and truly liable, respectively; let $LS_i = \frac{i - \frac{1}{2}}{n_{LS}}$ for $\forall i \in \{1, 2, \dots, n_{LS}\}$ represent the set of liability strength values; and let $N_{LS,k} = \langle \sigma_{LS} \Phi^{-1}(\frac{i}{k}) \rangle_{i=1}^{k-1}$ represent a sequence of $k - 1$ noise values drawn from the inverse normal distribution with standard deviation σ_{LS} . Then, $\forall t \in TL$, let

$$p_{LS_i|t} = \lim_{k \rightarrow \infty} \frac{|\{t + \nu | \nu \in N_{LS,k}, LS_i - \frac{1}{2n_{LS}} < t + \nu \leq LS_i + \frac{1}{2n_{LS}}\}|}{|\{t + \nu | \nu \in N_{LS,k}, 0 < t + \nu \leq 1\}|} \quad (1)$$

represent the probability that the signal with index i is received given a truly liable value of t . The probability of a liability signal conditional on the case not being truly liable is defined analogously. Note that the unconditional probability of each liability strength level is not necessarily equal, and indeed with this formulation, the probability that the litigation strength is at a level in the middle of the probability spectrum is lower than the probability that the litigation strength is at a level near the ends.

For player j , the probability of receiving a particular signal of liability strength can be calculated as

$$p_{LS_i^j|LS_i} = \lim_{k \rightarrow \infty} \frac{|\{LS_i + \nu | \nu \in N_{LS,k}, LS_i^j - \frac{1}{2n_{LS}} < LS_i + \nu \leq LS_i^j + \frac{1}{2n_{LS}}\}|}{|\{LS_i + \nu | \nu \in N_{LS,k}, 0 < t + \nu \leq 1\}|} \quad (2)$$

The volume of the flows from the “not truly liable” and “truly liable” nodes to each of the liability strength nodes correspond to the $p_{LS_i|t}$ probability values, and the volume of the flows from the liability strength nodes to the signal nodes correspond to the $p_{LS_i^j|LS_i}$ values.

Returning to our smaller game tree with just two liability strengths and two signals for each party in Figure 2, each path leads to a decision to be made by the plaintiff. An information set number is indicated next to the “P” indicating plaintiff. The information sets labeled “P0” correspond to the cases in which the plaintiff receives the low signal of liability strength, while the information sets labeled “P12” correspond to the cases in which the plaintiff receives the high signal. (The missing information set numbers between 0 and 12 appear in the endgame, to be presented momentarily.) This signifies that the plaintiff cannot distinguish among all of the cases within each information set. That is, the plaintiff does not know whether the case is one that is truly liable or not truly liable, and the plaintiff does not know whether the defendant receives a low signal or a high signal. The plaintiff’s equilibrium strategy must be optimal given the constraint that the probabilities that the plaintiff assigns to its next move must be the same, regardless of the true liability and of the defendant’s signal, and thus conditional only on the plaintiff’s signal.

[illegible]

Assuming that the plaintiff files and the defendant answers, then the plaintiff faces a decision whether the plaintiff will abandon the litigation rather than incur trial costs if settlement fails, and the defendant similarly faces a decision whether to default in the event settlement fails. Neither party's decision has any immediate effect, and neither party finds out about the other party's decision. The placement of these decisions before settlement rather than after is admittedly counterintuitive. But it greatly speeds up the algorithm for calculating equilibria, relative to a model in which each player takes into account both its own settlement offer and its opponent's in deciding whether to quit. With the latter approach, there would be one information set for each player for each triple of a liability signal, a settlement offer, and an opponent's settlement offer. By deciding abandonment or defaulting in advance, we greatly reduce the number of information sets in the game, with just one information set concerning quitting for each liability signal for a player. Although it might be interesting to study the strategic interaction in a model in which a party's offer may affect the

opponent’s decision whether to quit, that is beyond our scope here. Including decisions whether a party will abandon or quit litigation continues to serve its primary role of affecting bargaining indirectly, by forcing each party to consider in making a settlement offer whether its opponent has a credible threat to take a case to trial.

Whatever the future plans the parties make about whether to abandon or default, each party makes a settlement offer to the other party. In the diagram, it appears that the plaintiff makes an offer before the defendant, but this timing is again irrelevant. Note that the defendant faces the same information set regardless of whether the plaintiff makes a low offer or a high offer. That is, the defendant does not learn the plaintiff’s decision before it makes its own, so the decisions are effectively simultaneous.

The bargaining protocol used is that introduced by Chatterjee and Samuelson (1983) [5] and used in Friedman and Wittman [9], Dari-Mattiacci and Saraceno (2020) [6], and Abramowicz (2025) [1]. Each party submits simultaneous offers, and the case settles if and only if the defendant’s offer exceeds the plaintiff’s, with the settlement value equal to the midpoint. While future work might employ some other more complex bargaining protocol, such as back-and-forth offers between the parties, that would exponentially increase the number of information sets and the amount of time needed for the algorithm to run. Chatterjee-Samuelson bargaining avoids the simplification in the Landes-Gould-Posner bargaining models of assuming that a case will settle whenever that is in the mutual interest of the parties, given their expectations. With Chatterjee-Samuelson bargaining, parties have some incentive to shade their offers in their own favor in the hope of capturing more of the settlement surplus, though they must balance this objective with the risk of bargaining failure. The approach thus allows for an important dimension of strategic bargaining, while avoiding the simplification adopted by models that arbitrarily choose a player to make a take-it-or-leave-it offer to the other player.

Finally, at the far right end of Figure 4 are the utilities received by the players in cases that settle, as well as the further game play in cases that do not settle. If a case does not settle, but one player previously decided to quit in the event of settlement failure, then the game immediately ends without trial. If both players decided to quit, which in practice is likely to occur only when the cost of trial is much greater than the level in Figure 4, then a chance decision assigns each a 50% chance of quitting, as demonstrated near the top right of the diagram. In the more important case where neither party quits after settlement failure, trial occurs. Each party is charged a cost of $c_{trial} = 0.15$, and the court must determine whether to impose liability. The court makes its decision based on a noisy assessment of liability strength, must as the players do; thus, the modeling allows not only for liability strengths that differ from truth but also for different levels of judicial idiosyncrasy. The parameter determining the noisiness of the signal that the court faces is the same as that for the players, $\sigma_{LS}^C = 0.2$. The court, however, observes one of only two liability signal levels. If it observes the higher level, the court imposes liability. In this game tree, no fee shifting occurs, so liability results in a transfer from the defendant to the

plaintiff of 1.0. Note that the court is making an inference based on its signal, but this is not a fully Bayesian inference. The court does not take into account the initial distribution of disputes, and it does not refine its inferences based on game play, for example by determining whether a failure of settlement is more likely when the defendant is or is not truly liable. A project for future work, more in the spirit of the mechanism design literature such as Spier (1994) [25], would be to make the court an additional player in the game that can determine the level of damages and seeks to optimize social welfare.

The game tree can be simplified when calculating equilibria. The Chance decisions determining true liability and the liability strength do not directly result in any information being added to the plaintiff's or defendant's information set. The only information from a Chance decision that matters to the litigants is the signal of liability strength. We can thus skip the first two Chance decisions and jump directly to the stage where the plaintiff receives a signal, so long as the probabilities of each signal are the same for the plaintiff and so long as the probabilities that the defendant receives each signal conditional on any signal received by the plaintiff remain the same. Figure 5 shows a reduced form version of the beginning of the game tree in the small game where there are only two signals. The probability that the defendant receives the same signal as the plaintiff, 79.2%, can be computed in the larger game tree by weighting the probability of different branches. Using this reduced form will produce the same equilibria as will result from following Figure 2.

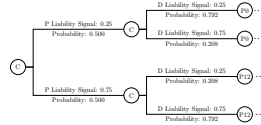


Figure 5: A reduced-form version of the beginning of the small game tree

The end of the game can also be simplified. The court's eventual decision on liability affects the players, but a model in which a player receives the expected value of the court's decision with certainty produces identical equilibria. Meanwhile, in the rare circumstances when both parties decide to quit, resolving the case at an average of the utilities that would obtain from each quitting produces the same result as flipping a coin to determine whether the plaintiff wins or loses. In both cases, this works regardless of risk aversion level, so long as the final outcomes are utilities, rather than dollars. Figure 6 illustrates the reduced-form version of the end of the game tree in the very small game.

The full game tree will still be needed to calculate variables such as party win rates at trial and accuracy of the system of litigation. Thus, after the equilibria are calculated using the simplified version of the tree, the full tree can be explored to determine the exact probability of every outcome. None of the results in this article depends on Monte Carlo sampling, and so no error bars are needed, given the exact equilibrium identified by the algorithm.

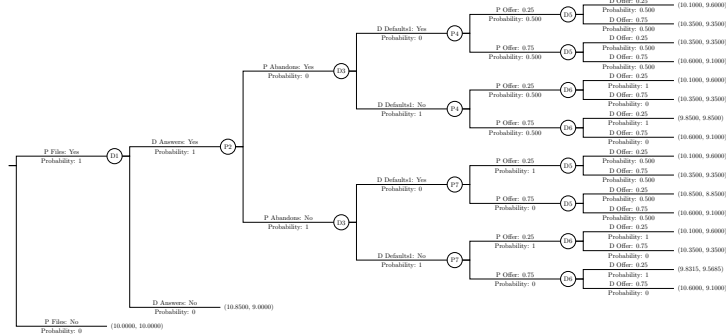


Figure 6: A reduced-form version of end of the game tree

3 Results

In addition to furnishing a much more complex game tree than theorists have been able to achieve with pure math models, the computational game theory approach applied here enables many changes to the litigation game to be made with no increase in algorithmic complexity. Any change that affects only the players' utilities leaves the game tree the same size, and so the time needed for modeling needs to be multiplied only by a constant factor representing the number of different permutations of parameter values to be tested. Some considerations that are difficult to model mathematically are trivial computationally. Risk aversion, for example, is ordinarily complex to model, and different forms of utility function may require separate proofs, as in Park and Lee (2019). The approach here easily accommodates changes in the degree or form of risk aversion. Similarly, changes to the costs the parties bear and the quality of their information add no complications. Different fee-shifting rules and intensity of fee shifting, including requirements that the loser pay a multiple of legal fees, can easily be integrated in the model.

This section thus considers a wide range of potential variations of its settlement bargaining model. The analysis begins with the baseline parameter values defined in the game illustrated in Figure 1 (but with $n_{LS} = n_{LS}^P = n_{LS}^D = n_a = 10$), while varying costs and fee shifting multipliers. The parameter values are admittedly not yet calibrated to actual data, but real world parameter values are likely to vary from place to place and case category to case category in any event. Probably the most important parameter, other than the fee-shifting rule, that may vary across suits is the cost relative to the stakes. The litigation may play out very differently in a small claims case, in which total costs to litigants including lawyers and their own time will generally be high as a percentage of the stakes, than in a multi-billion dollar suit in which even the priciest lawyers will charge only a small fraction of what is at issue. Thus, even when varying other parameters, simulations considering various permutations of costs and the fee-shifting multiplier are included. Space limitations prevent us from reporting

all findings here, but spreadsheets and diagrams for each simulation are available in the Supplemental Materials, including over 20,000 diagrams in PDF or Latex form, and simple changes in the source code can be used to test combinations of parameters not already simulated.

3.1 Baseline comparison

Setting costs at $c_{file} = c_{answer} = c_{trial} = 0.15$ with stakes normalized at 1, we may scrutinize the results of two simulations, corresponding to the American rule and the British rule. In the latter, if a case goes to trial, the loser pays the winner’s costs, i.e. 0.30 with these baseline parameter values. The model does not differentiate whether these costs are in the form of attorney fees, court costs, or litigants’ time. Figures 7 and 8 illustrate the plaintiff’s decision whether to file suit and the defendant’s decision whether to answer (conditional on filing) in the identified equilibria. Although models like Shavell (1982) [24] suggest that the British rule will reduce the likelihood that a plaintiff with a low likelihood of winning will file suit, no such effect is visible in these figures, as the plaintiff’s decisions are identical. The defendant’s decisions are almost identical too, except that when the defendant receives a signal of 0.65, the defendant has a 4% chance of answering under the British rule, in contrast to the American rule, where the defendant never answers.

The unexpected stasis helps illustrate the power of computed equilibria to identify dynamics not apparent in simpler models. This result is a consequence of multiple competing pressures that balance one another, pressures that would not be evident in a game that did not feature file/answer decisions, settlement offers, and decisions whether to give up after the failure of settlement. Consider a plaintiff who receives a signal of 0.35 and is thus near the margin of deciding whether to file suit. The British rule does indeed provide further impetus in the direction of not filing suit. On the other side of the liability signals spectrum, the defendant faces similar pressure not to answer. Fee shifting, however, also gives this plaintiff a greater incentive to drop the suit rather than face trial costs if settlement fails. That change in turn makes it more attractive for the defendant to fight, toughening the defendant’s settlement negotiation stance. These changes can lead to further hydraulic pressures that just happen in this example to approximately cancel out.

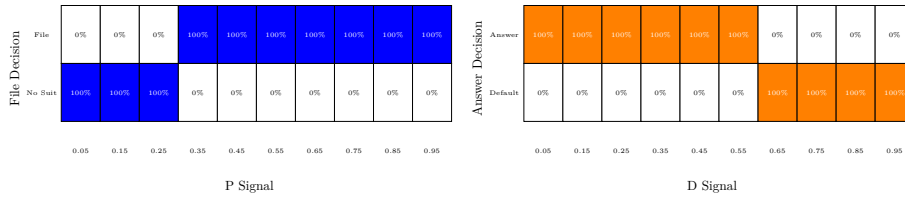


Figure 7: File and answer decisions (American rule)

It might seem that this is a “just so” story and that one could construct

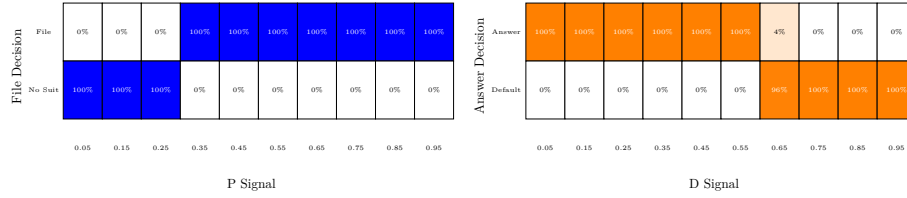


Figure 8: File and answer decisions (British rule)

many narratives to explain how changing a parameter leads to some change (or in this case, no change) in some decisions by the parties. Indeed, the equilibrium dynamics are sufficiently complicated that we should resist the temptation of overinterpretation. But we can directly observe at least the immediate pressures on the file and answer decisions that result from changing from the American to the British rule. Specifically, we can change the players' utility payoffs to those that obtain with fee shifting, but retain the strategies that they play under the American rule equilibrium. Then, we can examine the utilities that each player can expect to receive at each information set. Because an information set may correspond to multiple nodes of the game tree, these utility values are the expected utilities from choosing each action at a node, weighted by the probability of reaching that node in the game tree. At equilibrium, the actions with positive probability must all have the same, maximal expected utility (otherwise, we would not be at perfect equilibrium), but changing the payoffs can alter this. We can thus identify actions that the player would choose more often, assuming that all other aspects of both parties' strategies remained fixed.

This analysis (contained in the Supplemental Materials under the Supplemental information folder) demonstrates, for example, that, holding all else equal, the plaintiff's nonsuit rate would increase by 6.4 percentage points, and the defendant's default rate by 4.4 percentage points, if changing to the British rule while holding the other party's strategy constant. It also demonstrates how changing these values in turn would lead to the other effects described in the previous paragraph. The technique thus provides some intuition for why switching to the British rule might not lead the parties to contest litigation less often. The words "might not" highlight not that we should be unsure about whether the equilibria partially depicted in Figures 7 and 8 really exist—they do, and the source code confirmed that they are in fact perfect Nash equilibria, though of course we cannot be sure that human beings will play according to the equilibria—but that we must be attentive to the possibility that different results might obtain with other parameter values.

We will return to exploring other parameter values, but first let us consider the offers that the parties make, conditional on the signals that they received. This is illustrated in Figure 9 for the American rule, and Figure 10, for the British. Some columns feature no offers; these are the same columns for which the parties did not contest litigation. What is striking about the offers that

the parties make is that they are stingy. When the plaintiff receives a signal of 0.35, the plaintiff insists on a settlement of no less than 0.75. (Recall that under Chatterjee-Samuelson bargaining, the case will settle only if the defendant's simultaneous offer is greater than or equal to the plaintiff's.) The defendant, meanwhile, when receiving a signal of 0.55, will offer no more than a nuisance settlement of 0.15. And the parties are stingier still when moving toward the extremes, at which each party insists on the stingiest possible settlement. One might have the intuition that settlement offers would rise roughly linearly with signals and that close cases would settle for middle-range values. After all, the parties do have a monetary interest in avoiding trial.

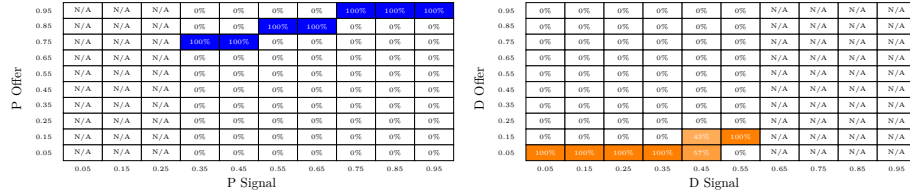


Figure 9: Offer decisions (American rule)

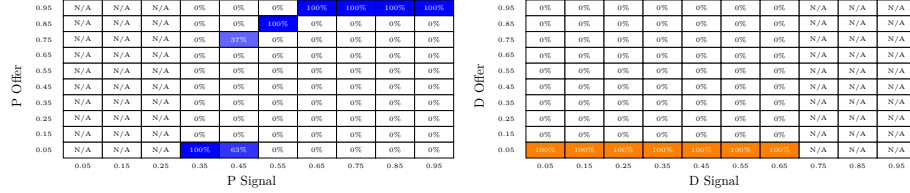


Figure 10: Offer decisions (British rule)

That strategic bargaining might occur is not surprising. All of the Bayesian models of settlement that use Chatterjee-Samuelson bargaining (including Friedman and Wittman (2007) [9]) recognize that each party faces a trade-off: Greater generosity increases the chance of avoiding trial, but greater stinginess increases the portion of the bargaining surplus that the offeror will receive if a settlement does occur. But strategic bargaining plays a considerably larger role than one might expect. Recall from Figure 3 that the amount of noise obscuring each party's evaluation of the merits does not seem all that large, and yet considerable strategic behavior manifests. The pattern of offers observed here is hardly invariant, but settlement behavior of this sort occurs under a wide range of parameter values, especially when the parties are risk neutral. Real-world bargaining includes many features not accounted for in this model, but this result tentatively suggests that litigants might rationally often fail to reach settlements even when their actual assessments of the litigation value are not far apart. At least, these results should make us cautious about accepting at face value the results of Landes-Posner-Gould models that assume settlement occurs whenever

settlement would produce social surplus.

The considerable extent of strategic bargaining does have an upside for this project. If the analysis revealed that parties' settlement offers were often very close to each other, then one might worry that the discreteness of action choices in this game theoretical model might be affecting results. Suppose, for example, that a change in some parameter would lead in a continuous model to slight changes in the parties' settlement functions that nonetheless had large effects on the settlement rate. It would be hard for a discrete model, in which each party must choose among only ten offer levels, to capture that nuance. The discreteness of the model with respect to some variables that may be continuous in mathematical models may indeed be the most serious weakness of the approach described here relative to a mathematical model, though in Abramowicz (2025), discreteness has only a minimal effect. In a world in which parties make either stingy or generous offers, but with little in between, a model with ten offer level choices seems likely to be able to capture, at least approximately, the bargaining dynamics.

Figures 11 and 12 offer a visualization of how the plaintiff's file decision, the defendant's answer decision, the parties' offers decisions, and to a much lesser extent their decisions whether to quit after settlement failure, directly affect the variables of ultimate interest: accuracy and expenditures. Inaccuracy is defined in two different (though generally highly correlated) ways: as false negatives or false positives. A false negative inaccuracy is measured from the perspective of the meritorious plaintiff. Every dollar by which the plaintiff's net recovery—that is, damages received minus plaintiff's litigation expenditures—falls short of what the plaintiff ideally would receive from a truly liable defendant in a hypothetical perfectly accurate and costless legal system counts as a false negative dollar. A false positive inaccuracy is measured from the perspective of the truly not liable defendant. Every dollar that such a defendant spends, whether on litigation or in damages, counts as a false positive dollar. These values can be calculated directly, because the model explicitly calculates the proportion of cases in which the defendant is truly liable or not. The mathematical modeler, as we will see later, may need to resort to much less direct measures of inaccuracy that fail to take into account actual litigation expenditures.

Figures 11 and 12 provide a clean visual depiction of both error costs and expenditures, which have social welfare significance independent of their effects on accuracy. Nonetheless, these diagrams may require some explanation. Every case is accounted for on the vertical axis, and cases are ordered by their type of disposition, with cases in which the plaintiff doesn't file suit and the defendant does not answer at the bottom, and cases that settle (under these parameters, only with the British rule), that result in a party quitting after settlement failure (just a few cases with the British rule, too narrow to fit a label), or that result in a trial with the plaintiff losing or winning pile on toward the top. In areas of the diagrams where the false negatives and false positives are empty, this is because those cases do not produce any false positives. For example, the vast majority of cases in which the plaintiff does not file suit are cases in which the defendant is truly not liable, so this resolution is nirvana, producing zero error costs and zero

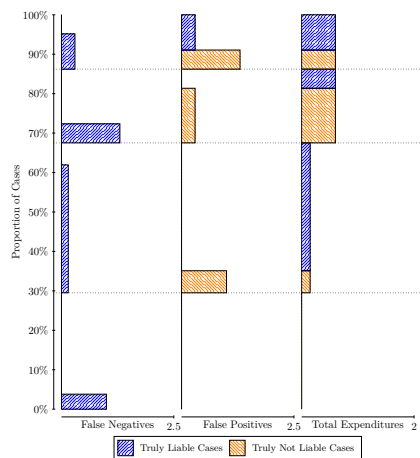


Figure 11: Errors and expenditures (American)

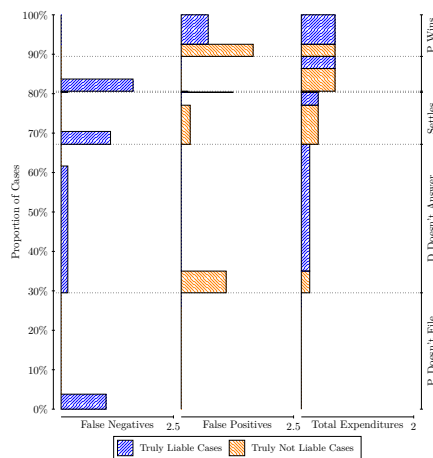


Figure 12: Errors and expenditures (British)

expenditures. A few cases, however, are cases in which the defendant is truly liable, resulting in false negatives, represented by the rectangle with north east lines in the bottom left of the figure. (All diagrams are designed with patterns to be presentable in black and white, but the color versions, in which this box is blue, are available in the Supplemental Materials.) The height of this rectangle represents the proportion of cases of this type, and the width represents the size of the false negatives, large but not as great as would be if the parties spent money on litigating first.

The cases in which the defendant does not answer produce similarly low error rates, but not as low, because the plaintiff bears the filing costs. When cases settle, there are some false positives and false negatives, as well as a moderate amount of expenditures from both parties' initial expenditures. Finally, when the cases go to trial, expenditures are at their maximum, and because the court is not perfect, some errors result: false positives when a nonmeritorious plaintiff wins and false negatives when a meritorious plaintiff loses. Expenditures may cause some inaccuracy even in cases in which the correct party wins. For example, when the plaintiff loses in a case in which the defendant truly is not liable, under the American rule, the defendant suffers a false positive as a result of litigation expenditures. Similarly, when the plaintiff wins in a case in which the defendant truly is liable under the British rule, no false negatives occur because the plaintiff receives full compensation, but the defendant suffers false positives from paying even more than damages in total.

In short, the area of the rectangles represents the total size of the false negatives, false positives, and total expenditures. Thus, one can at a quick glance appreciate the social welfare effects of fee shifting, holding all other parameter values constant. In this example and in most others, the effects are not all that great. Still, the British rule may have a slight edge. The British rule fails to

increase the chance that the plaintiff will not file or that the defendant will not answer, a change that at least at the margins would be beneficial here given the blankness of these regions, but it is no worse than the American rule. The key difference is that the slightly greater party generosity in settlement under the British rule (visible in Figures 9 and 10) leads to more settlement. It is not obvious that this should be the result. While a party that expects to lose more likely than not should be more generous, a party that expects to win should be less generous. (The analysis for now ignores that parties may spend more when the British rule applies, an issue to which we will return.) The increased settlement naturally decreases expenditures in the British rule, and it also modestly decreases false negatives and false positives. This initial face-off might qualify as a narrow win for the British, but, like the Battle of Bunker Hill, it might yet give some encouragement to the American side.

3.2 Costs and fee-shifting multipliers

To develop a more complete picture of how fee-shifting intensity may affect case disposition and thus accuracy and expenditures, we can vary the costs level and introduce a wider range of fee-shifting levels. Recall that in the baseline $c_{file} = c_{answer} = c_{trial} = 0.15$, so if a case goes to trial, each party spends 0.3, or 30% of the amount at stake. Yet there may be some cases in which costs, broadly conceived, may be much lower relative to stakes and others in which costs are higher. Thus, in Figure 3.1, the baseline costs levels are multiplied by each element of $\{\frac{1}{4}, \frac{1}{2}, 1, 2, 4\}$, and five fee shifting multipliers, ranging from 0 (representing the American rule) to 2 (where the loser pays twice the winner’s fees) are included. Although multiple fee shifting certainly would be unusual and is not considered in the earlier algorithmic game theory model of Abramowicz (2025), Polinsky and Rubinfeld (1996) [21] recognize that it is unclear what the optimal multiple is. After all, without a multiple, a party with a 75% chance of winning would expect on average to bear half the per-party costs (one quarter of the parties’ combined costs). Because costs are exogenous in the model, the fee shifting multiple can be viewed more generally as a range of penalties that the loser must pay to the winner.

The middle row of Figure 3.1 confirms, interpolates, and extrapolates what we determined above. The stacked bars represent the disposition of cases, with each group of stacked bars reflecting the results of one simulation and different dispositions coded with different patterns. The amount of fee shifting has relatively minimal effects on the number of cases that the plaintiff does not bring or the defendant does not answer. But fee shifting does affect settlement, increasing it at least up to a multiplier of 1. For multipliers above this level, some of the settled cases do not settle and instead become cases in which the plaintiff or the defendant gives up before trial. There is less pressure to settle if there is a good chance that the other party will quit rather than fight. Finally, the rate of trial is lower, the greater the amount of fee-shifting.

Dari-Mattiacci and Saraceno (2020) note that the “common wisdom” is that the English rule discourages settlement. This observation dates to early papers

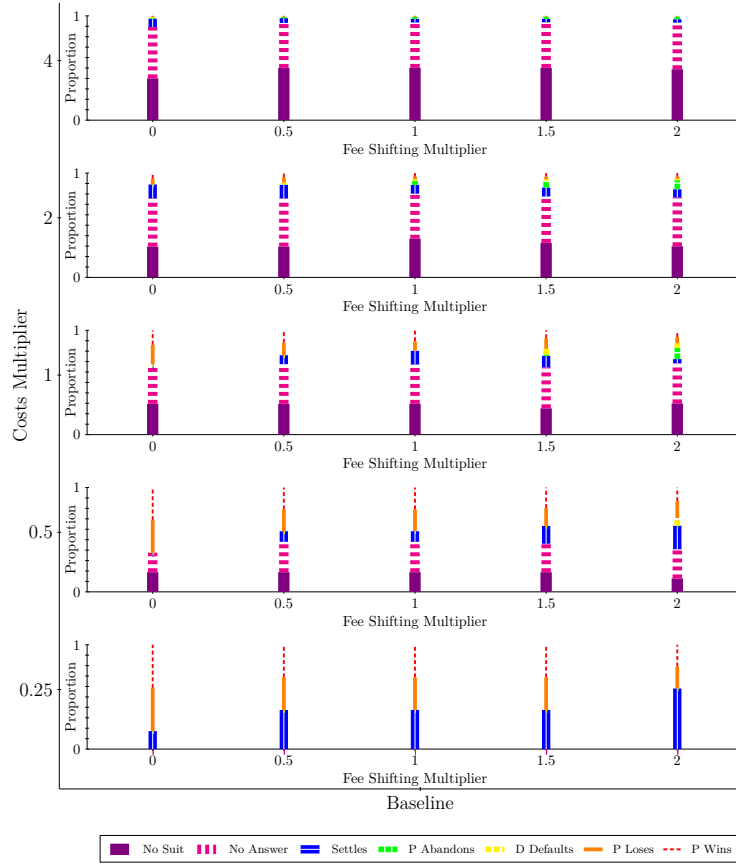


Figure 13: Dispositions with baseline parameters, varying costs and fee shifting multipliers

on fee-shifting, such as Shavell (1982). The intuition is that the cases that tend to go to trial will be those in which the parties are mutually optimistic, and if the parties are sufficiently mutually optimistic, each will expect to benefit from fee-shifting, so fee-shifting makes settlement less likely. A problem with this logic is that a rational Bayesian should recognize that if trial occurs, that will be the result of mutual optimism. Dari-Mattiacci and Saraceno counter the common wisdom, in a model in which damages are uncertain, by deriving a result “that the rate of litigation [meaning trial] is independent of the fee-shifting rule.” (p. 14) Abramowicz (2025) largely reproduces this finding, at least at relatively low costs values. The data here further undermines the common wisdom by demonstrating that, at least under some parameters, settlement will become more common and trial will become less common with greater amounts of fee-shifting. This conclusion reflects not only settlement dynamics but also decisions

whether to file and answer, which Dari-Mattiaci and Saraceno consider only informally (p. 13).

The result that trial levels fall with fee-shifting survives an examination of the other rows of the chart, which consider other costs levels. Moreover, the effect is even more stark if one reads the chart by moving not just to the right to observe the effect of a higher level of fee shifting, but simultaneously also up to observe the effect of a higher level of costs. Making a comparison in the style of a chess bishop may be justified because when fee shifting is greater, parties may spend more on litigation. Katz (1987) [16], for example, estimates that changing from the American rule to the British rule might increase per case expenditures in tried cases by as much as 125%. Thus, one might approximately assess the effect of such a change by comparing the results associated with a 0 fee shifting multiplier and a costs multiplier of 1 with those for multipliers of 1 and 2, respectively. If fee shifting in fact makes litigation more expensive, the model here suggests that it will also make trial considerably less common.

When comparing vertically or diagonally, even more apparent than the change in the trial rate is the change in the rates at which suits are filed and contested. With the lowest cost cases, suit is always filed and contested, and no one quits after settlement failure. In these cases, higher fee shifting multiples generally lead to more settlement, although there is little difference among the middle range of multipliers. With the highest costs cases, where the combined cost of trial exceeds the stakes, cases are rarely brought, and those brought are rarely defended. The remainder are mostly settled to avoid the trial costs, and if not settled, abandoned, with very few cases going to trial. One should thus expect very few contested cases to have such high costs relative to stakes. Such an absence may not mean that the litigation system always succeeds at keeping costs to a reasonable percentage of stakes, but instead that such cases are rarely litigated. This need not necessarily signify a failure of the legal system; note that about half of the uncontested cases are nonsuits and half are defaults, and many of these will end in the correct outcome with minimal if any litigation expenditures. Indeed, one might tentatively venture that error costs and litigation costs are paradoxically lower when the cost of the legal system is higher.

On these parameters, though, that inference would be wrong. Figure 14 illustrates false negative and false positive inaccuracy, as well as average expenditures on litigation, including in the denominator all potential cases, not just those actively litigated. As costs rise, all of these lines generally rise with them. With the highest costs, false negatives are very high relative to false positives, especially with fee shifting. The timing of litigation affords the defendant an informational advantage; there are some suits that the defendant would drop that are never brought by the plaintiff. Enough of the cases are still litigated that expenditures amount to around 35% of the stakes with the ordinary British rule. Still, this might be less than one would expect, only around three times greater than the level of total expenditures that occurs when costs are $\frac{1}{16}$ as large. The parties' greater hesitance to fight partially but incompletely offsets the increased relative cost of litigation. Still, the results strongly contradict the idea that higher expenditures per trial might lead to greater accuracy and

perhaps lower costs overall. As trial costs rise, both total accuracy and total expenditures unmistakably increase.

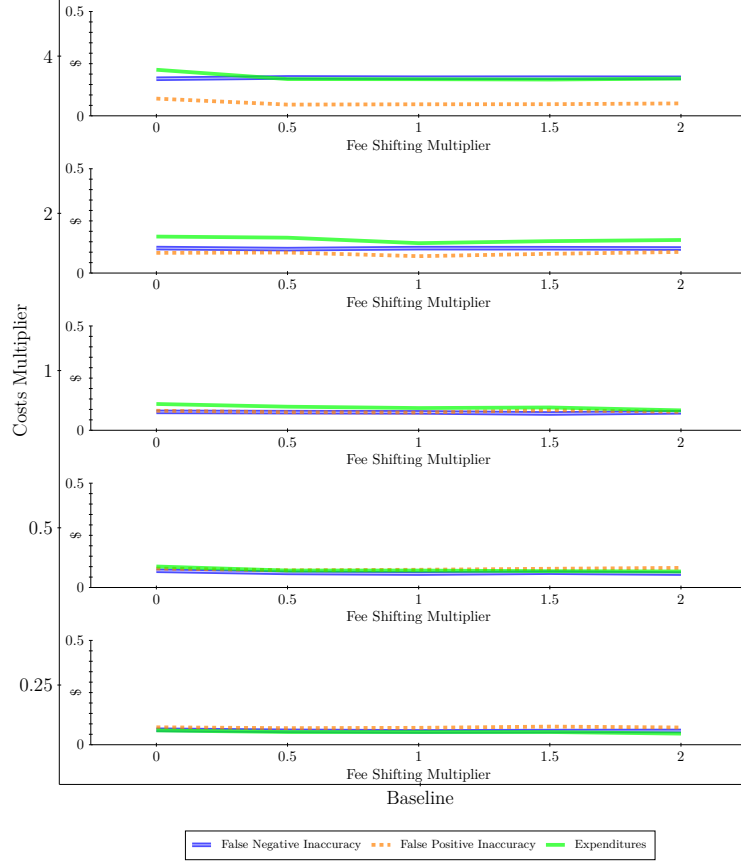


Figure 14: Inaccuracy and expenditures with baseline parameters

Within each graph comprising Figure 14, the error costs and expenditure lines are relatively, though not perfectly, flat. This may seem at odds with another observation of Dari-Mattiacci and Saraceno, that fee-shifting better promotes accuracy in lower-cost litigation systems, perhaps explaining the pattern of different rules on opposite sides of the pond. The level of costs affects the level of errors, but increasing cost does not much weaken the case for fee-shifting. With correlated signals, close observation provides at most limited support for their claim. Error costs fall ever so slightly at the baseline costs level. (Total expenditures, a variable that they do not consider, also fall slightly.) Still, we do not observe deteriorating inaccuracy with greater costs, as Abramowicz (2025, fig. 10) found in many models when using a definition of accuracy similar to that adopted in this article.

3.3 Risk aversion

A computational model also can easily consider the effects of risk aversion. Utility with risk aversion is modeled as a function of wealth, $U = -e^{-\alpha W}$, and the values of α used to model risk aversion in these simulations are 1, 2, and 4. Recall that each party's initial wealth is set to 10.0, and stakes are 1.0. The shape of the risk aversion curves, with the x axis depicting wealth on the interval $[9, 11]$, are displayed in Figure 15. The y-axis utility values are omitted, because the utility values can be (and are) linearly transformed without affecting the equilibrium.

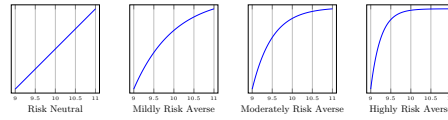


Figure 15: Levels of risk aversion

3.3.1 Shared risk preferences

Figure 16 confirms that settlement becomes considerably more common and trial becomes considerably more rare when the parties are risk averse, especially if risk aversion is pronounced. Anticipating the increased likelihood of achieving a settlement, parties are more likely to contest claims. In the middle row, for example, as risk aversion becomes greater, the proportion of cases in which the plaintiff does not file or the defendant does not answer falls. The effect is greater when fee shifting is absent. With moderate or high risk aversion and no fee shifting, litigation is always contested at the baseline costs level, and a very high percentage of cases settle. Introduction of fee shifting, however, greatly reduces the chance of settlement success. Risk aversion exacerbates the downside risk of fee shifting, and when settlement fails, one party or the other frequently quits rather than face trial. Still, the relationships are quite complex. In the rightmost panel of the middle row, for example, the proportion of cases resulting in settlement initially falls but then rises with increased settlement. When both parties are risk averse, each player seeks to avoid trial but also recognizes that the other player is also risk averse, making trial unlikely. At the highest costs levels, however, risk aversion makes trial and even settlement very unlikely. With high risk aversion and a costs multiplier of 4, fee shifting means that the merits cease to differentiate outcomes, as either the plaintiff never files or the defendant never answers.

Sufficiently high risk aversion makes the litigation game similar to the famous game theory game of chicken (or, for biologists, hawk-dove). In contrast to the classic game of chicken, however, the parties are not identically situated, and which party has the more credible threat to go to trial may vary based on their signals, the costs levels, and the rules. Given these complex dynamics, it is not surprising that the effects on inaccuracy and expenditures are more complex as

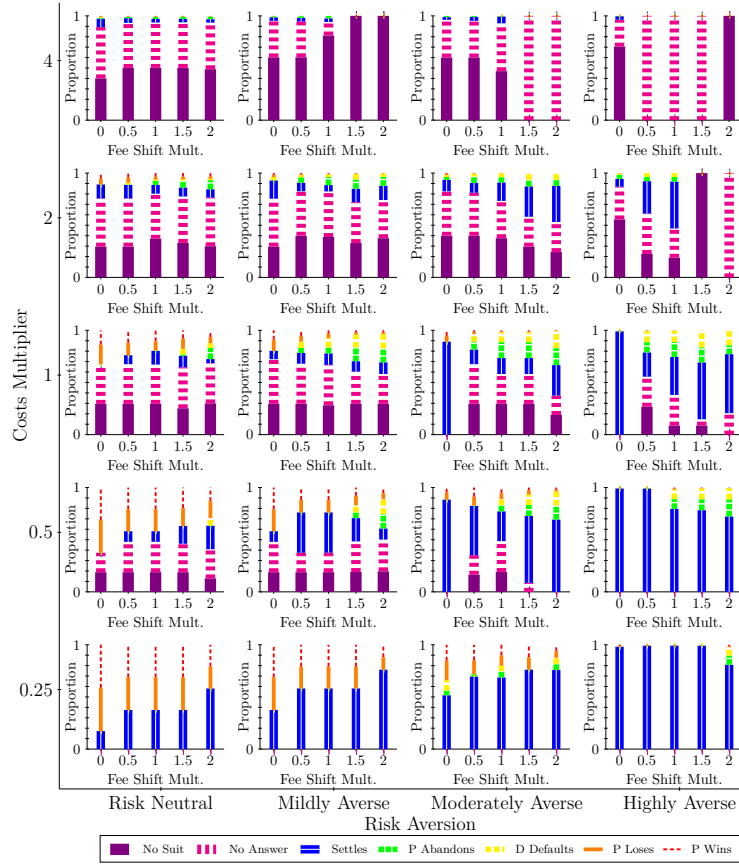


Figure 16: Dispositions varying risk aversion

well, especially with high levels of risk aversion, high levels of costs, and high degrees of fee shifting. The bottom three rows of Figure 17 suggest that mild risk aversion has relatively small effects. Overall, there is less inaccuracy and there are lower expenditures with mild risk aversion than with risk neutrality, though of course the false negative inaccuracy and false positive inaccuracy may have a greater effect on welfare. With moderate risk aversion and cost multipliers of 0.5 or 1, the benefits of fee shifting are accentuated. But with the highest levels of risk aversion, the picture becomes more complex at those costs multipliers, with moderate fee shifting multipliers performing considerably better than higher multipliers. At the highest costs level, the two inaccuracy measures sharply diverge in some places, as the party too cowardly to contest litigation suffers. All in all, the data suggests that with mutual mild risk aversion, ordinary fee shifting will generally produce benefits, but the dynamics are more complex and predictions more uncertain with high ratios of costs to stakes.

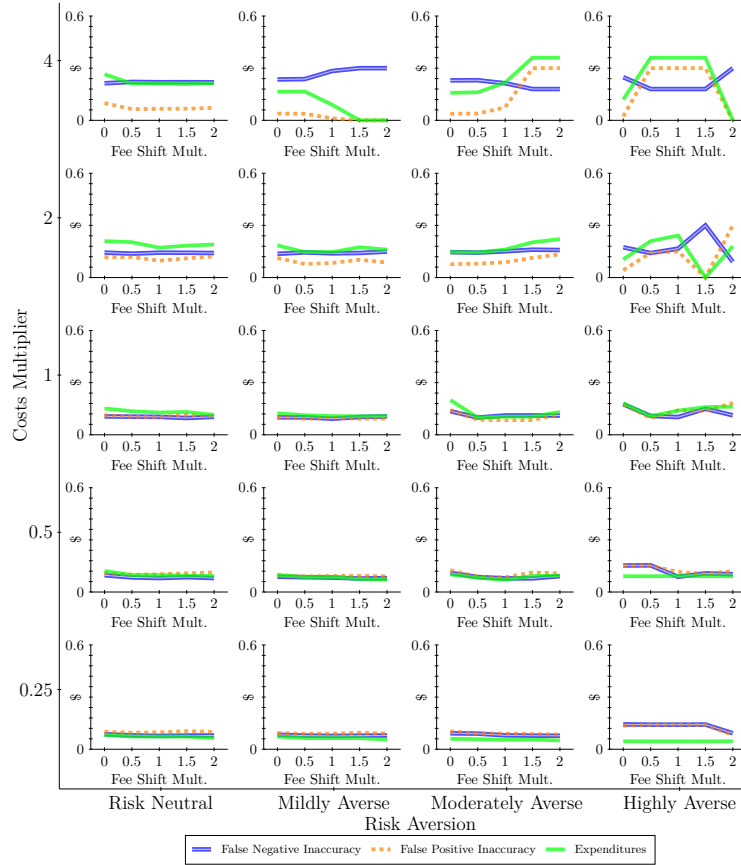


Figure 17: Inaccuracy and expenditures with risk aversion

3.3.2 Asymmetric risk preferences

The above analysis assumed that both parties had the same level of risk aversion. What happens if one party is more risk averse than the other? Predictably, that has significant effects on willingness to contest litigation. Figure 18 illustrates this. When the plaintiff is risk averse and the defendant is risk neutral, or the plaintiff is moderately risk averse and the defendant is only mildly risk averse, there are many cases in which the plaintiff does not file, and very few in which the defendant does not answer. The reverse pattern obtains when the defendant is more risk averse.

The result is a divergence between the false negative inaccuracy and false positive inaccuracy rates. When the plaintiff is less likely to file, there are high risks of false negatives, and the reverse is true when the defendant is less likely to file. As shown in Figure 19, these effects are relatively small when costs are relatively low and neither party is more than mildly risk averse. The

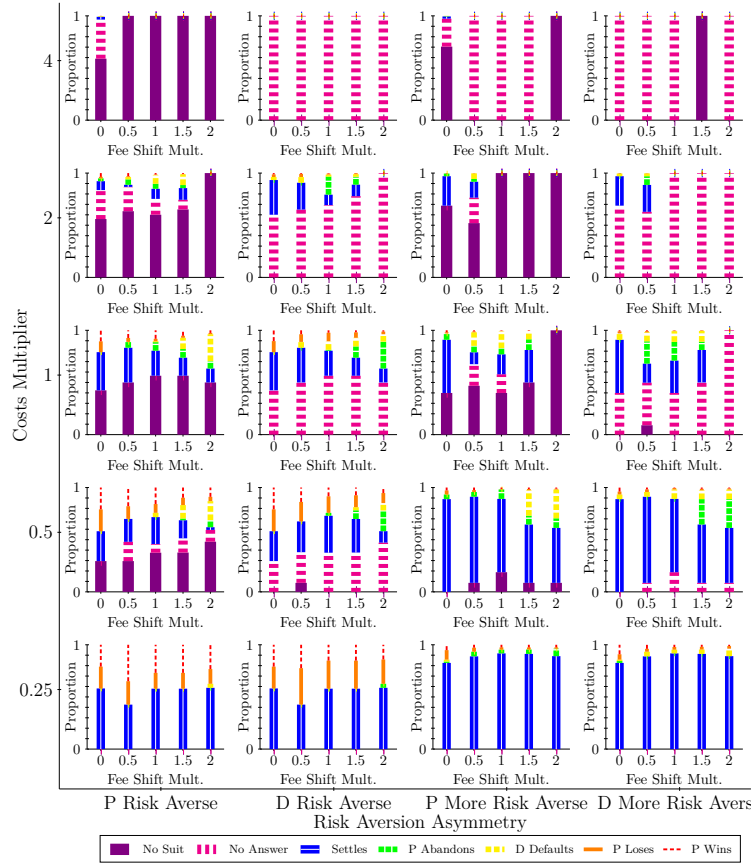


Figure 18: Disposition of cases with asymmetric risk aversion

divergence becomes greater with higher costs and greater levels of risk aversion. It also generally becomes greater within the individual graph panels when the fee shifting multipliers increase. The intuition is that fee shifting increases the stakes, and it thus exacerbates the effects of asymmetric risk aversion. This is not true in all cases. Consider, for example, the row where the costs multiplier is 0.5 and both parties exhibit at least some risk aversion. Still, the general point is that risk aversion is not a healthy trait in a poker player. A striking pattern in Figure 18 is that when parties quit after settlement failure with high levels of fee shifting, the quitting litigant is more likely to be the *less* risk averse party. Contesting a litigation when one has no intention of going to trial is a form of bluff, and the less risk averse party is in a better position to bluff.

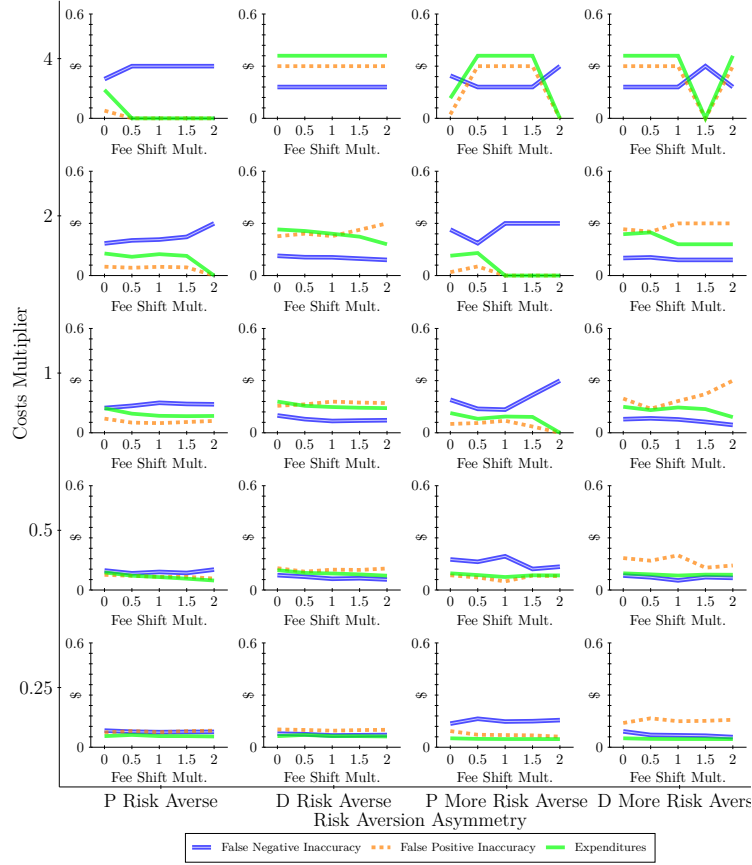


Figure 19: Inaccuracy and expenditures with asymmetric risk aversion

3.4 Informational endowments

We will now return to our assumption of risk neutrality, though the Supplemental Materials contain a version of every simulation not focusing on risk run under conditions of mild risk aversion. Perhaps as important as risk aversion is the degree to which players are able to ascertain the strength of the case. Recall that σ_{LS}^P and σ_{LS}^D represent the standard deviation of the normal distribution used to obfuscate the signal of liability strength received by the plaintiff and defendant, respectively. In these simulations, the noise affecting the court, σ_{LS}^C is set to the noise of the party with stronger information. The baseline value of these variables is 0.2. Meanwhile, σ_{LS} , with a baseline value of 0.35, governs the relationship between true liability status and case strength, with lower values leading to a tighter connection. Figure 3 illustrates the baseline values, and the Supplemental Materials include diagrams illustrating the other signal values used in this section.

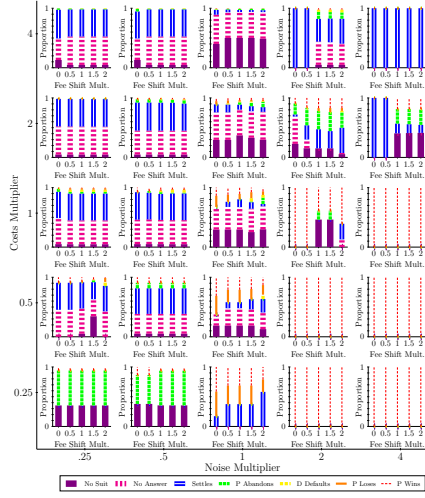


Figure 20: Disposition based on player noise levels

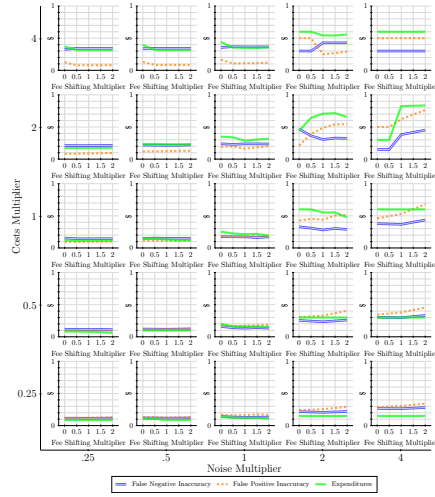


Figure 21: Errors and expenditures based on player noise levels

3.4.1 Player noise level

If the litigants both have better information, they are more likely to be able to avoid litigation, either through settlement or because one party does not contest the litigation. This is illustrated in Figures 20 and 21. With the lowest levels of noise (that is, the best information) and the lowest costs, cases are always filed, because a settlement is likely, but with higher costs, even at half of our baseline, nuisance suits may not be worth the cost of filing for the plaintiff, and nuisance defenses may not be worth the cost for the defendant. As a result, total error costs and expenditures are virtually unchanged in the lower three rows of the left column, with the cost savings of cheaper litigation entirely erased by greater pursuit of litigation. At the highest cost levels, an asymmetry develops in favor of the defendant, once again because on occasion the plaintiff decides not to bring a suit that the defendant would not have contested. Meanwhile, fee shifting plays an important role when player information is strong, though its overall effect is limited, because the vast majority of potential cases are not contested at all. For costs in the middle range, there is a decline in trial as a result of fee-shifting, albeit from already relatively low levels. If a case appears to be in the middle of the probability spectrum, without fee-shifting it might be worth it to pursue the case to trial, but with fee shifting, if settlement fails, at least one player is likely to make a Bayesian inference that victory is unlikely and give up.

With high noise and bad information, the picture is quite different. Litigation becomes a crapshoot, and cases are likely to go to trial when costs are relatively low. When costs are higher, a wide range of outcomes are possible. With the highest level of cost and the highest level of noise, the plaintiff takes

advantage by always filing. The defendant then realizes that if the defendant answers, both will still have sufficient incentive to negotiate aggressively that trial is a real possibility, a disaster resulting in each party paying more than the amount at stake altogether. This example is a special, extreme case of a negative expected value suit that can be rational for a plaintiff to bring. Slight changes in assumptions can lead to significant changes in equilibria; for example, in the equivalent to the top-right cell with risk aversion (not shown but available in the Supplemental Materials), the plaintiff never files. We cannot rule out the possibility that there are multiple equilibria here, as in the game of chicken. Overall, with high noise, the effect of fee shifting is relatively low. At the limit, if neither party has any information, then the expected value of fee shifting for a risk neutral party is 0.

3.4.2 Information quality asymmetry

Figures 20 and 21 assume that the players have equally high-quality information, whatever the noise level might be. In this section, we consider the possibility that one party has higher quality information. In Figures 22 and 23, the middle column reflects a situation in which the plaintiff experiences half as much noise ($\sigma_{LS}^P = 0.1$) while the defendant experiences twice as much ($\sigma_{LS}^D = 0.4$), and that is reversed in the column further to the right. This change makes our model quite close to conventional one-sided asymmetric information models, where one party knows the precise probability of liability and the other party knows only the distribution. Placing aside the row with very low costs, fee shifting in such a situation tends to reduce the incidence of settlement, as most models suggest. The story, however, is more complex than the typical signaling model (in which the party with the information makes an offer) or screening model (in which the party who lacks the information makes the offer), and the outcome is neither a pure pooling or pure separating equilibrium. Rather, as in most of the models here, the parties are sometimes able to enter into settlements, but often fail to reach agreement.

Settlement becomes quite rare with medium to high costs and a fee shifting multiplier of 1 or greater. If the party with high-quality information knows that it has a strong case, then it has little reason to make concessions in settlement, because it is likely to benefit from fee shifting. Thus, the party with high-quality information will make concessions only if its case is relatively weak. The party with low-quality information thus faces a lemons problem: If the opponent makes sufficient concessions to allow a reasonable settlement, it is likely the case that the opponent knew its case was weak, in which case settlement was inadvisable for the party with low-quality information. Settlement thus becomes rare or nonexistent. Meanwhile, the party with high-quality information may abandon the litigation assuming settlement fails if the information indicates that its case is actually weak. The party with low-quality information also has some reason to quit. After all, if the party with high-quality information continues to trial despite a high level of fee shifting, that probably means that the party with high-quality information knows that it has a strong case.

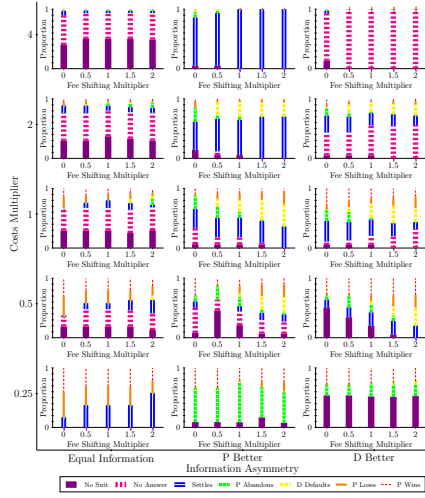


Figure 22: Disposition based on information quality asymmetry

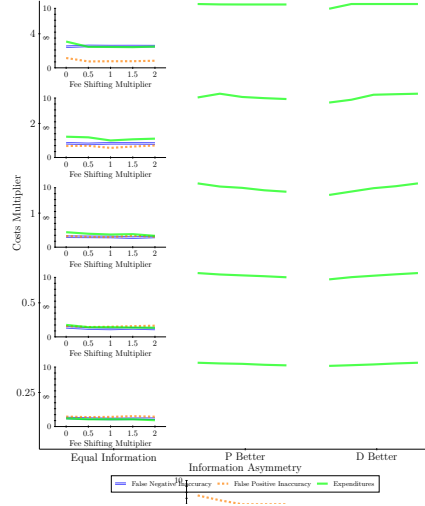


Figure 23: Errors and expenditures based on information quality asymmetry

Thus, the party with high-quality information will continue to trial if it has a reasonably favorable signal, and the party with low-quality information will continue to trial only if it has a very favorable signal. The result is among cases that are tried, the party with better information wins a high percentage of the time. This is, of course, not inconsistent with the famous result of Priest and Klein (1984) [23]. They showed that the plaintiff would win half of all cases only in a limit equilibrium, as both parties' information became sufficiently strong. The analysis here suggests that with high information quality asymmetry, plaintiff win rates may be considerably higher than 50%, if the plaintiff has higher-quality information, or considerably lower than 50%, in the reverse situation. Anticipating these bargaining dynamics, the party with low-quality information becomes less likely to contest suit as the fee shifting multiplier increases, at least when costs are relatively high, and the party with high-quality information will almost always contest, unless its information is especially bad, because of the chance that the opponent will quit.

The informational advantage produces small reductions in the danger that the party that possesses it will suffer the adverse effects of inaccuracy, while imposing such adverse effects on the other party. For example, in the column in which the plaintiff has better information, false negative inaccuracy, which occurs when a meritorious plaintiff receives too little on net, is slightly lower than in the baseline simulation. But false positive inaccuracy is much worse. Fee shifting, meanwhile, modestly amplifies these effects, reducing the danger of inaccuracy for the party with better information but increasing inaccuracy for the other party. Meanwhile, the effect of fee shifting on expenditures is indeter-

minate, falling in some panels while rising in others. The attractiveness of fee shifting when information is highly asymmetric will thus depend on which type of inaccuracy looms larger in the social loss function. If the two forms of inaccuracy are given equal weight and the effect of expenditures is indeterminate or small, the American rule is likely to be superior with highly asymmetric risk aversion.

3.4.3 Quality of the legal system

So far, this section has focused on variations to the noise that obfuscates the liability quality when observed by a party as a liability signal. We may also vary σ_{LS} , the noise level that determines the liability signal given that a case is one in which the defendant is or is not truly liable. This can be seen as a measure of the quality of the legal system. (Varying σ_{LS}^C , the noise obfuscating the court's determination, might have a similar effect, but we simply assume that the court's information is as good as the better informed party's. Altering this might usefully be varied to assess the effects of judicial bias if the parties can predict the court's bias, but we leave this for future work.) Figure 24 provides a visualization of the effect of choosing a lower or higher value for case strength noise, using the baseline level of costs. The full diagram, which is available in the Supplemental Materials, shows similar effects. The primary effect of increasing the case strength noise is to increase inaccuracy and expenditures, and the reverse is true for decreasing the case strength noise. This is as one would expect. If cases in which the defendant is truly liable are often cases in which the defendant is likely to win, the legal system will not be as successful in its aims. But the data suggest that such a change does not have major consequences for an analysis of the relative effect of changing the fee shifting multiplier. This provides some reassurance that the admittedly arbitrary choice of a case strength noise level is not likely to be driving results.

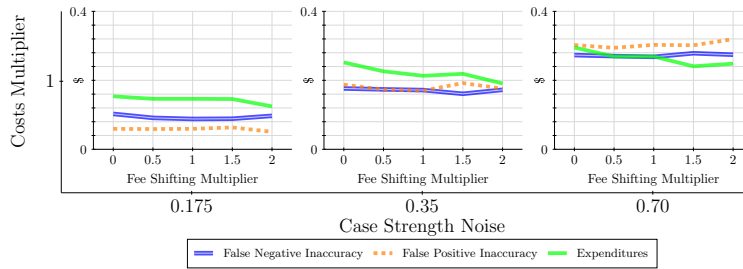


Figure 24: Inaccuracy and expenditures varying case strength noise

3.5 Costs structure

In each section, we have considered the possibility of different ratios of costs to the stakes of the case. Yet in our examples, the plaintiff's cost of filing has

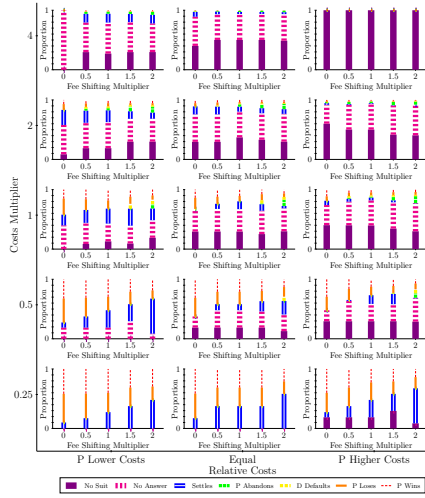


Figure 25: Disposition with asymmetric costs

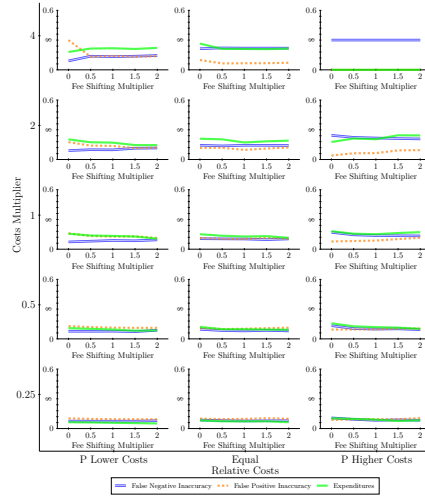


Figure 26: Errors and expenditures with asymmetric costs

equaled the defendant's cost of answering, and the trial cost has been that same amount for each party. In this section, we consider the possibility that the parties might bear different costs of litigation, or that the timing of costs might be weighted either toward the initiation of the case or toward trial.

3.5.1 Relative costs

A party might bear higher costs than its adversary because it chooses a pricier lawyer. The choice of attorney, which would affect likelihood of winning at trial, is beyond our scope here. But other parties also might bear different levels of costs for reasons outside their control. For example, the burden of discovery might fall more heavily on one party or the other. In Figures 25 and 26, we consider, in addition to the baseline scenario, a scenario in which the plaintiff's costs, both for filing and for trial, are half what they otherwise would be, and one in which the plaintiff's costs are twice what they ordinarily would be. In both scenarios, the defendant's costs are constant.

As one would expect, facing higher costs than one's opponent leads one to face a greater danger of inaccuracy. In the left column, false positives (adversely affecting the defendant) are generally greater in magnitude than false negatives, given the plaintiff's lower costs. And in the right column, the plaintiff's higher costs generally translate into a greater magnitude of false negatives than false positives. The effects start at the pleading stages. A plaintiff is more likely not to file a lawsuit at any level of costs than in the condition in which the plaintiff has lower costs or the parties have the same costs. The effect on the defendant is more muted. Because the plaintiff does not file relatively often when the plaintiff faces high costs, when the plaintiff does file, that provides a credible

signal to the defendant that the plaintiff has relatively favorable information. Meanwhile, greater degrees of fee shifting tend to mitigate the advantage of asymmetric costs. The false positives and false negatives become closer together in general to the right of the minigraphs in Figure 26. With fee shifting of at least 1, the party that bears the cost is not necessarily the party that spends the money. But the sum of false negative and false positive costs may not decline. Meanwhile, with asymmetric costs, expenditures tend to decline with higher fee shifting levels, but this is not true in all panels.

3.5.2 Time allocation of costs

Settlement bargaining dynamics depend considerably on when costs are incurred. Figures 27 and 28 hold constant the total costs that the parties will bear if a case goes to trial, but vary the proportion of those costs that are borne at the beginning or the end. In the leftmost column, there is no cost to filing or answering, so of course the parties always file and answer. The dynamics are akin to models that consider only settlement and ignore the prior decisions of whether to file and defend. Especially with moderate to high costs, a party often gives up if settlement fails. Trial is relatively more expensive (because total costs are fixed), and so a party that thinks it has a weak case will likely not proceed to trial. Meanwhile, the column furthest on the right assumes that all of the costs are borne at filing, and trial is free. This extreme seems unlikely, but there may be cases where the amount of work that precedes settlement negotiation is great, and the expense of trial is low in relative terms. Unsurprisingly, in this circumstance, it is common for parties not to contest litigation, and fee shifting magnifies this effect when costs are relatively low. Once a suit is filed, it is pursued to trial absent settlement, which rarely succeeds. Overall, fee shifting again reduces the incidence of trial.

The net effects on error costs and total expenditures are relatively small. In the middle baseline column, unless costs are very high, total expenditures tend to fall modestly with greater amounts of fee shifting, and the same is true in the other four columns. Whether the moment of truth occurs before or after settlement negotiations, increased fee shifting, especially double fee shifting, decreases trial rates. Accuracy effects are more equivocal, and the false positive and false negative lines are flatter than they are in many of the interventions. The greater the proportion of costs at the beginning, the greater the tendency of false positives and false negatives to be minimized in the middle of the fee shifting spectrum, with the British rule. But there are exceptions, especially at higher costs levels, where the differential timing of plaintiff and defendant decisions plays a larger role.

3.6 Game structure

In all of the simulations so far, the game tree has remained the same. This section considers three possible changes to the game structure. It first considers case in which damages are at issue instead of liability. That provides an op-

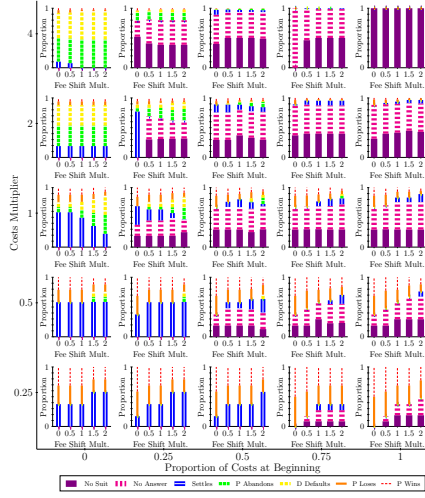


Figure 27: Disposition based on timing of costs

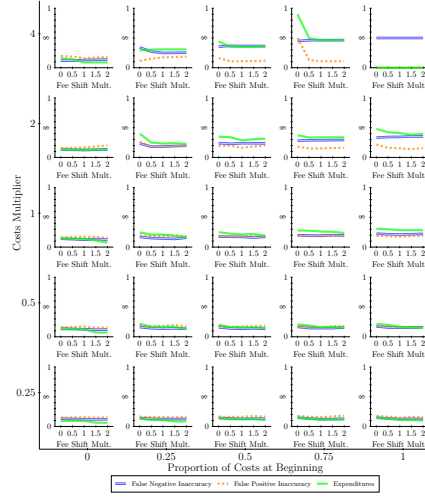


Figure 28: Errors and expenditures based on timing of costs

portunity to examine the fee shifting mechanism embodied by Federal Rule of Civil Procedure 68, which has more relevance to damages cases, as well as fee shifting rules in liability cases that resemble Rule 11 in imposing fees only when the margin of victory is high. Finally, we consider the effects of eliminating the right to quit if settlement fails, a variation of an intervention that some have suggested.

3.6.1 Damages cases

When damages are at issue, the model assumes that the plaintiff is always liable and that there is some true level of damages of which the parties receive a signal. The true level of damages is assumed to be uniformly distributed across 10 values, and the plaintiff's and defendant's signals of true damages quality are the same as their signals of true liability quality, $\sigma_{DS}^P = \sigma_{DS}^D = 0.2$. Figure 29 provides the damages analogue to Figure 3.

Because the defendant is always liable, fees are always shifted against the defendant if the case goes to trial. The result, illustrated in Figures 30 and 31 is that fee shifting greatly decreases the defendant's willingness to go to trial. With very low costs, fee shifting results in more settlement, but in all but the lowest row, fee shifting substantially increases the proportion of cases in which the defendant decides not to contest liability, thus paying the full damages demanded by the plaintiff. The result is that false positives greatly increase, considerably more than false negatives decrease. The analysis suggests the obvious conclusion that courts must be careful in determining when fee shifting applies. Applying fee shifting whenever the plaintiff prevails on the issue of liability will not make sense if the defendant has conceded liability

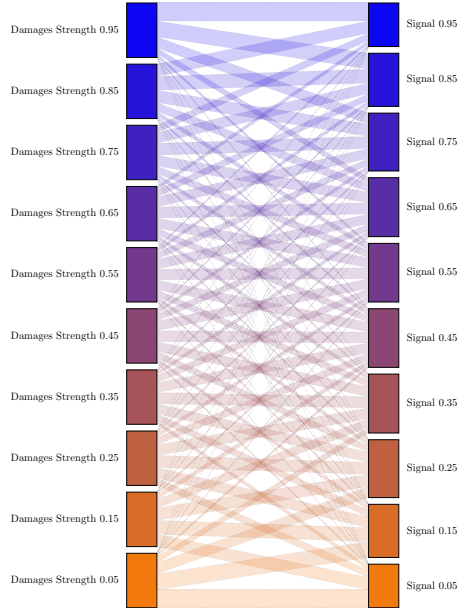


Figure 29: Damages signals given true damages strength

and contests damages. Indeed, courts applying the British rule often explicitly recognize that the party entitled to attorneys' fees is the party that prevailed on the significant issues in a case. *See, e.g., Trytek v. Gale Indus.*, 3 So. 3d 1194 (Fla. 2009). This highlights, however, a need for models that consider the possibility that liability *and* damages are at issue.

3.6.2 Fee-shifting rules

In cases in which damages are at issue, Federal Rule of Civil Procedure 68 or mechanisms similar to it may be useful. Under Rule 68, if a defendant makes a settlement offer to a plaintiff, and the defendant is found liable but the damages are less than the plaintiff was ordered, the plaintiff must pay the defendant's costs. Under *Delta Air Lines, Inc. v. August*, 450 U.S. 346 (1981), the plaintiff is not required to pay the defendant's costs when the defendant wins the lawsuit. If the plaintiff were required to do so, Rule 68 would amount to a simple one-way cost-shifting rule, because the defendant would always offer a nominal settlement, such as one dollar, and then seek shifting if the defendant prevailed. Thus, Figures 32 and 33 reflect assumptions that the defendant is always truly liable and is found liable. The amount that the plaintiff is required to pay is equal to the defendant's initial fees times the fee multiplier, but only when the rule is triggered.

The figures also reflects simulations employing a variation on Rule 68, for which we use the label "Reverse 68." Under this Rule, the British rule is the

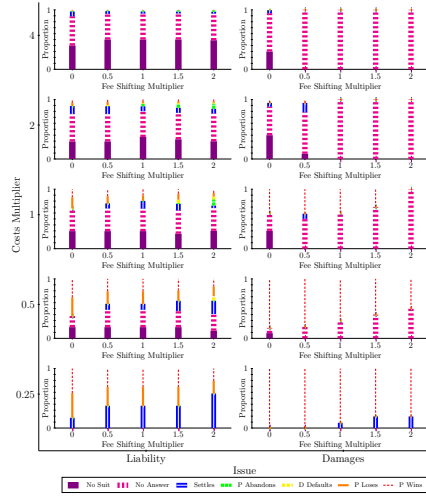


Figure 30: Disposition based on issue in case

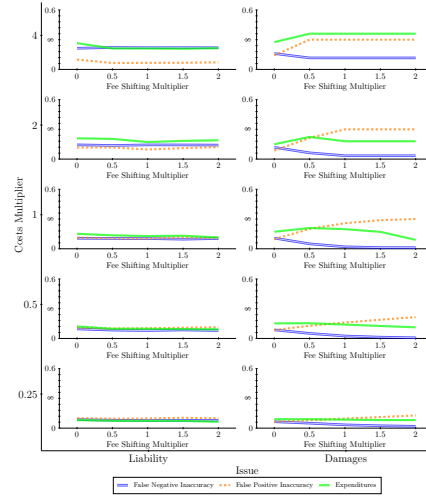


Figure 31: Errors and expenditures based on issue in case

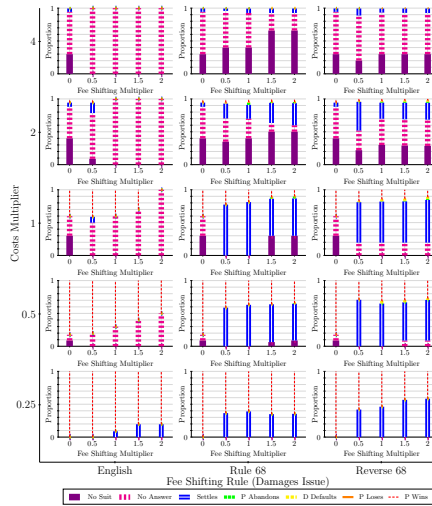


Figure 32: Disposition based on fee-shifting rule

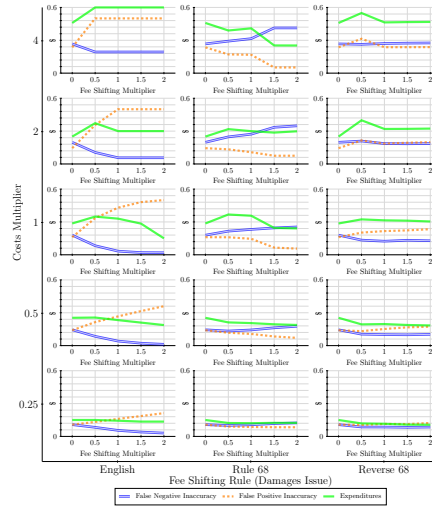


Figure 33: Errors and expenditures based on fee-shifting rule

default, so the plaintiff may be entitled to fee shifting, but if Rule 68 is triggered, the plaintiff suffers a double whammy. Not only does the plaintiff not receive the fee shifting to which it ordinarily would be entitled, but in fact must pay the defendant for having rejected its offer. Arguably, Reverse 68 is closer to how Rule 68 works in practice, because the prevailing party ordinarily is entitled to costs under Rule 67, and Rule 68 sometimes changes that entitlement. This has virtually no effect in practice, however, because costs are generally small relative to attorneys' fees, which are not allowed under Rule 68. We thus use the phrase "Rule 68" to describe a hypothetical version of Rule 68 that applies to all costs and fees, and the phrase "Reverse 68" to describe the effect of that rule if the United States hypothetically simultaneously changed to the British rule default.

Different fee-shifting rules also may be used when the issue is liability rather than damages. We simulated fee-shifting based on the margin of victory, as discussed in Bebchuk and Chang (1996) [2]. Specifically, we considered the possibility that the court might impose fee shifting only if the court's signal of liability were at least equal to 0.80. The results (omitted here, but available in the Supplemental Materials) show only small differences in dispositions and costs from baseline. Further work might examine a variety of different possible thresholds at which to apply margin-of-victory fee shifting.

3.6.3 Quitting rules

Finally, we consider the effect of removing the option that players have to quit after settlement failure. It is useful in part because one can imagine a proposal under which parties are allowed to settle but not to quit. The intuition for such a proposal is that quitting allows a party to evade the fee shifting rule, escaping the need to pay for the other party's expenses incurred before trial. More importantly, however, this exercise is useful to facilitate assessment of theories, such as Huang (2004) [13] and Hubbard (2016) [14], that highlight the importance of this option on settlement dynamics. We can assess whether optionality is more or less important with fee shifting.

Figures 34 and 35 show that, at least with moderate to high costs, the rule barring settlement is counterproductive in the absence of fee shifting. The rule leads to more contested litigation. Each side recognizes that settlement is the only alternative to trial. Thus, a party that might not settle ordinarily, confident that the other side would abandon litigation, will be more inclined to settle, and thus a party with a slightly weak case will be more inclined to contest in the first place. Relative to a world in which those parties do not contest litigation, this is more expensive, and so litigation is more expensive, with little benefit in terms of false positives and negatives. When fee-shifting exists, however, the error costs and expenditures in the right column are fairly close to those in the left column. That does not create a strong argument in favor of instituting such a rule. Indeed, it illustrates that even though fee shifting can lead to bluffing followed sometimes by quitting, removing the ability to bluff would not significantly alter outcomes. The more general lesson is that

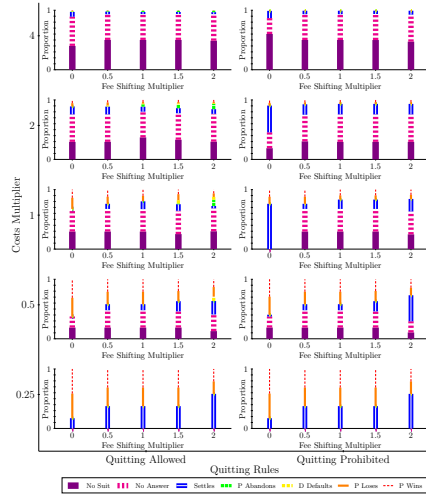


Figure 34: Disposition based on permissibility of quitting

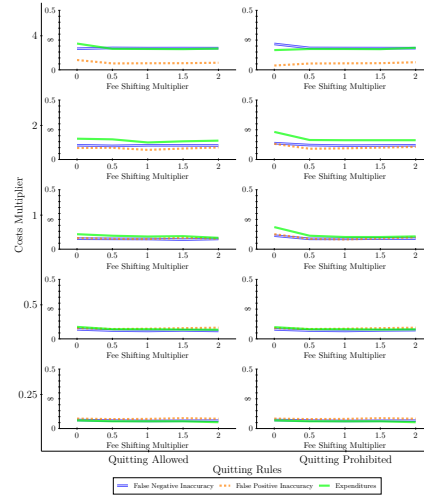


Figure 35: Errors and expenditures based on permissibility of quitting

litigation optionality, so critical to assessing incentives under the American rule, may matter less under the British. This is so even though parties are more likely to quit to avoid trial under the British rule, especially with multiple fee shifting, than under the American rule.

4 Analysis

This article's methodology has made it possible to generate exact equilibria for a wide variety of game permutations, 1,250 in all. But this leaves two questions about robustness. First, are the results dependent on the happenstance of stumbling upon some equilibria rather than others? Second, if multiple equilibria do exist and the parties play different equilibria, is that likely to significantly affect the results? And third, are the results dependent on the variables left fixed across simulations, namely the number of liability (or damages) strength values, the number of signals, and the number of offers? We address these issues in turn.

4.1 Multiple equilibria

The von Stengel, van den Elzen, and Talman (2002) algorithm is guaranteed to furnish a single exact equilibrium given any arbitrary initialization of the parties' information sets to fully mixed strategies. It does not provide a means of generating all of the perfect Nash equilibria, in contrast to the Lemke-Howson algorithm, which generates all Nash equilibria but is many orders of magnitude too slow for this problem. This raises the question whether the equilibria iden-

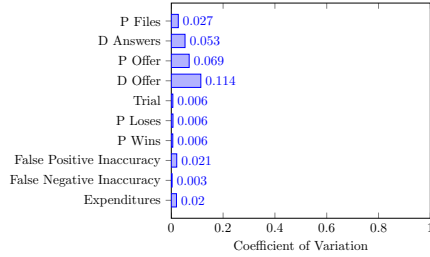


Figure 36: Coefficient of variation for 29 American rule equilibria

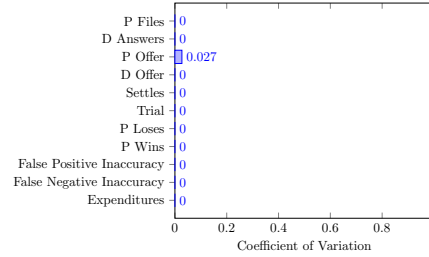


Figure 37: Coefficient of variation for 9 British rule equilibria

tified here are representative of the broader class of equilibria for each game type. We can address this by randomizing the parties' initial information sets to determine whether that generates multiple equilibria and, if so, how much variation there is among these multiple equilibria.

We performed this for the baseline simulation, both with and without fee shifting, by attempting to generate up to 50 equilibria for each. This produced 29 equilibria for the American rule, because 21 of the equilibria were repeats, and 9 for the British rule. We cannot from this exercise determine how many equilibria exist. In principle, these equilibria could be the entire populations or they could be a small sample. Figures 36 and 37 thus reports the coefficient of variation (the standard deviation divided by the mean) for various outcome variables among the American and British rule equilibria, respectively. The results are relatively reassuring. There is some variation with the American rule, particularly in the answer and offer decisions, but this results in only very small variation for our primary outcome variables of interest, error costs and expenditures. (The “Settles” variable is omitted; there were no settlements in all but one equilibria, and in that equilibria, 10.7% of cases settled.) The results for the British rule are even more reassuring. The only outcome variable with a nonzero coefficient of variation was the plaintiff offer, at 0.03; this variation was insufficient to actually change whether any cases settled.

4.2 Off-equilibrium play

Chicken highlights a central concern about multiple equilibria, that if they exist, the parties might not coordinate on any one equilibrium. The only pure Nash strategies in chicken are those in which one party swerves and the other does not. Either way, no one dies. But if each player believes that they are coordinating on the equilibrium in which the other swerves, then both die. This outcome is also possible, of course, if they coordinate on the mixed strategy, in which each has a high probability of swerving, but is far more likely if there is a potential failure to coordinate on pure strategies. In the context of litigation, if the plaintiff believes that the equilibrium is one in which the plaintiff drives a hard bargain and the defendant thinks that the equilibrium is one in which the defendant

drives a hard bargain, then they may end up with a nonequilibrium settlement failure that would not occur if both parties were coordinated.

Given the low coefficient of variation for the baseline scenarios, the noncoordination problem appears insignificant, but that does not rule out the possibility that off-equilibrium play could be considerably different from equilibrium play in other scenarios. It was infeasible to generate a large number of equilibria for each of the 1,250 scenarios to assess the implications of coordination failure, given available computational resources. As a substitute, we generated up to 50 equilibria per scenario on a smaller game tree size, where $n_{LS} = n_{LS}^P = n_{LS}^D = n_o = 5$. For each equilibrium attempt, we initialized the mixed strategies at random in a way that allowed for the possibility that one action might have a much higher probability than others, in the hope of creating highly diverse starting points for the algorithm. For efficiency, we then attempted to find an equilibrium using floating point arithmetic. That is not guaranteed to succeed because of numerical stability issues, so whenever it failed, we switched to seeking an equilibrium using exact arithmetic. We then collected all the unique equilibria produced.

We then compared outcomes in a correlated equilibrium with outcomes in an average equilibrium. A correlated equilibrium, developed by Aumann (1974), is one in which each party, observing some signal of which equilibrium to play, has the incentive to follow that signal on the assumption that the other party will as well. Any set of Nash equilibria can be combined to create a correlated equilibrium, because if one knows that one's opponent is playing a particular Nash equilibrium, by definition one will want to play the strategy corresponding to that equilibrium as well. The average equilibrium is calculated simply by averaging the parties' mixed strategies at each information set. For example, if in half of equilibria, plaintiff files 100% of the time conditional on a signal and in the other half, plaintiff files 50% of the time conditional on that signal, then in the average equilibrium, plaintiff files 75% of the time. The average equilibrium will thus have support over at least as many pure strategies as any of the original equilibria strategies.

For each of the 1,250 parameter sets, we created a correlated equilibrium strategy and an average equilibrium strategy and computed the probability or average value of game outcomes for each. We then calculated how correlated the variables were for a number of variables of interest. The result is illustrated in Figure 38, which shows the correlation coefficient across the 1,250 values for each outcome variable. This suggests that lack of ability to coordinate on any single strategy does not dramatically change the story.

4.3 Game tree size

The equilibria on the smaller game tree also can be compared directly to the equilibria on the original game tree. This provides a test of whether results are highly sensitive to the granularity of the strategies. The diagrams illustrating equilibrium results for both the smaller and original game tree are available in the Supplemental Materials. We will offer here a partial visual comparison.

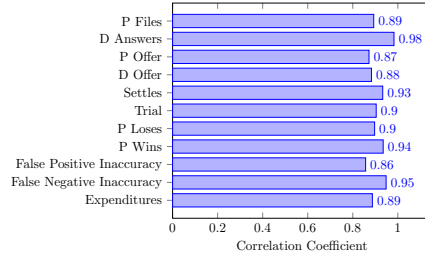
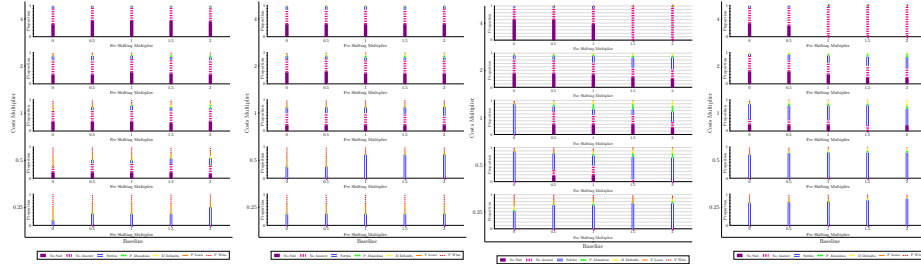


Figure 38: Correlations in outcome variables between correlated and average equilibria



which litigation is contested, even with asymmetric risk aversion, or with different levels of party information, at least unless the parties are significantly risk averse. Parties with greater risk aversion or lower-quality information suffered considerably more from litigation inaccuracy than their adversaries. When litigation becomes more expensive, fewer cases are contested, but the greater per-case expense and gamesmanship in file/answer and quitting decisions generally worsens social welfare outcomes. At very high levels of costs, especially when costs are backloaded, litigation becomes akin to a game of chicken. One or both parties may be bluffing by filing or answering, settlement becomes rare, and inaccuracy is high.

Perhaps the most decisive conclusion that can be drawn from the constellation of models explored in the paper is that litigation may play out quite differently with different costs levels. An advantage of using a common model in different scenarios is that we can be more comfortable attributing the difference to the changes in parameters than we would be if comparing the implications of pairs of models making multiple different assumptions. The analysis illustrates more clearly than is possible under other models the complex interactions among file/answer decisions, settlement offers, and quit options, as well as how these change depending on case stakes, the degree of fee shifting, risk aversion, and a number of other variables. The comparisons suggest that the long-term goal must be either to develop different rules for different case types or to develop alternative litigation rules that tend to maximize social welfare under a wide range of parameters.

With respect to fee-shifting, the project sheds light on how the presence or degree of fee-shifting affects the incidence of trial and social welfare measures. [INSERT]

At least three forms of skepticism about the methodology are warranted. First, litigants may not be rational Bayesians who play equilibrium strategies. But, as the literature generally presumes, models assuming rationality may provide a first guess at how litigants behave, and economic forces may gradually push actors in any given legal context toward equilibrium strategies. Moreover, though more complicated to model, two-sided asymmetric information seems like a more accurate description of most litigation than one-sided information, and a model that recognizes the possibility that litigants may not contest litigation or may quit after settlement failure illustrates dynamics that some models assume away.

Second, models' usefulness may depend on parameter values, and the parameters here are not calibrated based on data, either statistics from actual litigation or on microfoundations of individual behavior. Indeed, the heterogeneity of results that obtain from this model's admittedly arbitrary parameter choices highlight both the futility of assuming that any calibration will fit all litigation but also that calibration may be useful for particular types of cases. Risk aversion, for example, has a profound effect on the results, and even casual observation suggests that the high trial rates observed with risk neutrality are less plausible than the rates observed with mild risk aversion.

Third, although the model offers more bells and whistles than its math-

ematical predecessors, many chimes are still missing, including learning and bargaining over time, as well as the role of lawyers and agency costs. Indeed, this is but a first step toward building models that can incorporate even more complex features of the litigation game. Because the game tree size may increase exponentially with additional features, other computational game theory approaches that seek approximate equilibria may be needed to make larger game models solvable.

References

- [1] Michael Abramowicz. “Modeling Fee Shifting with Computational Game Theory”. In: *Journal of Legal Studies* 54 (2025), pp. 165–198.
- [2] Lucian Arye Bebchuk and Howard F. Chang. “An Analysis of Fee Shifting Based on the Margin of Victory: On Frivolous Suits, Meritorious Suits, and the Role of Rule 11”. In: *Journal of Legal Studies* 25 (1996), pp. 371–403.
- [3] Lucian Ayre Bebchuk. “Litigation and Settlement Under Imperfect Information”. In: *RAND Journal of Economics* 15 (1984), pp. 404–415.
- [4] Yun-chien Chang and Daniel Klerman. “Settlement Around the World: Settlement Rates in the Largest Economies”. In: *Journal of Legal Analysis* 14 (2023), pp. 80–175.
- [5] Kalyan Chatterjee and William Samuelson. “Comparison of Arbitration Procedures: Models with Complete and Incomplete Information”. In: *IEEE Transactions on Systems, Man and Cybernetics* 31 (1983), pp. 835–851.
- [6] Giuseppe Dari-Mattiacci and Margherita Saraceno. “Fee Shifting and Accuracy in Adjudication”. In: *International Review of Law and Economics* 63 (2020), pp. 1–15.
- [7] Andrew F. Daughety and Jennifer F. Reinganum. “Settlement”. In: *Procedural Law and Economics*. Ed. by Chris William Sanchirico. 2012, pp. 386–471.
- [8] Andrew F. Daughety and Jennifer F. Reinganum. “Settlement Negotiations with Two-Sided Asymmetric Information: Model Duality, Information Distribution, and Efficiency”. In: *International Review of Law and Economics* 14 (1994), pp. 283–298.
- [9] Daniel Friedman and Donald Wittman. “Litigation with Symmetric Bargaining and Two-Sided Incomplete Information”. In: *Journal of Law, Economics, and Organization* 23 (2006), pp. 98–126.
- [10] Jonah P. Gelbach. “The Reduced Form of Litigation Models and the Plaintiff’s Win Rate”. In: *Journal of Law and Economics* 61 (2018), pp. 125–157.
- [11] John P. Gould. “The Economics of Legal Conflicts”. In: *Journal of Legal Studies* 2 (1973), pp. 279–300.

- [12] John C Harsanyi. “Games with incomplete information played by ‘Bayesian’ players, I–III Part I. The basic model”. In: *Management science* 14.3 (1967), pp. 159–182.
- [13] Peter H. Huang. “Lawsuit Abandonment Options in Possibly Frivolous Litigation Games”. In: *Review of Litigation* 23 (2004), pp. 47–113.
- [14] William H.J. Hubbard. “Sinking Costs to Force or Deter Settlement”. In: *Journal of Law, Economics, and Organization* 32 (2016), pp. 545–585.
- [15] Keith N. Hylton. “An Asymmetric-Information Model of Litigation”. In: *International Review of Law and Economics* 22 (2002), pp. 153–175.
- [16] Avery Katz. “Measuring the Demand for Litigation: Is the English Rule Really Cheaper?” In: *Journal of Law, Economics, and Organization* 3 (1987), pp. 143–176.
- [17] Daniel Klerman, Yoon-Ho Alex Lee, and Lawrence Liu. “Litigation and Selection with Correlated Two-Sided Incomplete Information”. In: *American Law and Economics Review* 20 (2018), pp. 382–459.
- [18] William M. Landes. “An Economic Analysis of the Courts”. In: *Journal of Law and Economics* 14 (1971), pp. 61–107.
- [19] Yoon-Ho Alex Lee and Daniel Klerman. “The Priest-Klein Hypotheses: Proofs and Generality”. In: *International Review of Law and Economics* 48 (2016), pp. 59–76.
- [20] A. Mitchell Polinsky and Daniel Rubinfeld. “Does the English Rule Discourage Low Probability of Prevailing Plaintiffs?” In: *Journal of Legal Studies* 27 (1998), pp. 519–535.
- [21] A. Mitchell Polinsky and Daniel Rubinfeld. “Optimal Awards and Penalties When the Probability of Prevailing Varies Among Plaintiffs”. In: *RAND Journal of Economics* 27 (1996), pp. 269–280.
- [22] Richard A. Posner. “An Economic Approach to Legal Procedure and Judicial Administration”. In: *Journal of Legal Studies* 2 (1973), pp. 399–458.
- [23] George L. Priest and Benjamin Klein. “The Selection of Disputes for Litigation”. In: *Journal of Legal Studies* 13 (1984), pp. 1–56.
- [24] Steven Shavell. “Suit, Settlement, and Trial: A Theoretical Analysis Under Alternative Methods for the Allocation of Legal Costs”. In: *Journal of Legal Studies* 11 (1982), pp. 55–81.
- [25] Kathryn E. Spier. “Pretrial Bargaining and the Design of Fee-Shifting Rules”. In: *RAND Journal of Economics* 25 (1994), pp. 197–214.
- [26] Bernhard von Stengel, Antoon van den Elzen, and Dolf Talman. “Computing Normal Form Perfect Equilibria for Extensive Two-Person Games”. In: *Econometrica* 70 (2002), pp. 693–715.