

The Delta-Method applied to Sobol' indices

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Abstract

We explore the use of the Delta-method in order to estimate the Sobol' indices.

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1 Convergence in distribution

Definition 1. (Convergence in distribution) Assume that X_1, X_2, \dots is a sequence of real-valued random variables with cumulative distribution functions $\{F_n\}_{n \geq 0}$. Assume that X is a real-valued random variable with cumulative distribution function F . The sequence X_n converges in distribution to X if:

$$\lim_{n \rightarrow \infty} F_n(X_n) = F(x).$$

for any $x \in \mathbb{R}$ at which F is continuous. In this case, we write:

$$X_n \xrightarrow{D} X.$$

The following theorem gives an example of such convergence.

Example 1. (Maximum of uniform random numbers) Assume that X_1, X_2, \dots are independent uniform random numbers such that $X_n \sim \mathcal{U}(0, 1)$. Let Y_n be the maximum:

$$Y_n = \max_{1 \leq i \leq n} X_i.$$

Therefore the sequence $n(1 - Y_n)$ converges in distribution to an exponential random variable, i.e.:

$$n(1 - Y_n) \xrightarrow{D} \mathcal{E}(1).$$

Proof. By definition, the exponential distribution with rate λ has the cumulative distribution function:

$$F(y) = 1 - \exp(-\lambda y),$$

for any real $y \geq 0$. We apply the previous equality to $\lambda = 1$ which shows that we must prove that:

$$F(y) = 1 - \exp(-y),$$

for any real $y \geq 0$. Let F_n be the cumulative distribution function of the random variable $n(1 - Y_n)$. By definition of the cumulative distribution function,

$$\begin{aligned} F_n(z) &= P(n(1 - Y_n) \leq z) \\ &= P(1 - Y_n \leq z/n) \\ &= P(1 - z/n \leq Y_n) \\ &= 1 - P(Y_n \leq 1 - z/n), \end{aligned}$$

for any $z \in \mathbb{R}$. However, the cumulative distribution function of the maximum is:

$$\begin{aligned} P(Y_n \leq y) &= P(X_1 \leq y, X_2 \leq y, \dots, X_n \leq y) \\ &= P(X_1 \leq y)P(X_2 \leq y) \cdots P(X_n \leq y) \text{ (by independence)} \\ &= yy \cdots y \text{ (by definition of the C.D.F. of a uniform random variable)} \\ &= y^n, \end{aligned}$$

for any $y \in [0, 1]$. Therefore,

$$F_n(z) = 1 - (1 - z/n)^n$$

for any $z \in \mathbb{R}$. But

$$\lim_{n \rightarrow \infty} (1 + z/n)^n = \exp(z)$$

for any $z \in \mathbb{R}$. Therefore,

$$\lim_{n \rightarrow \infty} F_n(z) = 1 - \exp(-z)$$

for any $z \in \mathbb{R}$ which concludes the proof. \square

2 Convergence in probability

Definition 2. (Convergence in probability) Assume that X_1, X_2, \dots is a sequence of real-valued random variables. The sequence X_n converges in probability to X if, for any $\epsilon > 0$:

$$\lim_{n \rightarrow \infty} P(|X_n - X| > \epsilon) = 0.$$

In this case, we write:

$$X_n \xrightarrow{P} X.$$

Example 2. (Convergence of an exponential random variable) Assume that X_1, X_2, \dots are independent random numbers such that $X_n \sim \text{Exp}(n)$. Therefore the sequence X_n converges in probability to the zero random variable:

$$X_n \xrightarrow{P} 0.$$

Proof. By definition of the exponential cumulative distribution function, we have

$$P(X_n \leq x) = 1 - \exp(-nx),$$

for any real number $x \geq 0$. Let $X = 0$. For any $\epsilon > 0$, we have:

$$\begin{aligned} P(|X_n - X| > \epsilon) &= P(|X_n| > \epsilon) \text{ (since } X = 0\text{)} \\ &= P(X_n > \epsilon) \text{ (since } X_n \geq 0\text{)} \\ &= 1 - P(X_n \leq \epsilon) \\ &= \exp(-n\epsilon) \end{aligned}$$

by definition of the exponential cumulative distribution function. Since $\epsilon > 0$, this implies:

$$\lim_{n \rightarrow \infty} \exp(-n\epsilon) = 0$$

which implies that

$$\lim_{n \rightarrow \infty} P(|X_n - X| > \epsilon) = 0$$

and concludes the proof. \square

3 Convergence in probability implies convergence in distribution

Theorem 1. Assume that X and Y are random variables. Let $y \in \mathbb{R}$ and let $\epsilon > 0$. Therefore

$$P(Y \leq y) \leq P(X \leq y + \epsilon) + P(|Y - X| > \epsilon) \quad (1)$$

Proof.

$$P(Y \leq y) = P(Y \leq y, X \leq y + \epsilon) + P(Y \leq y, X > y + \epsilon).$$

But $P(Y \leq y, X \leq y + \epsilon) \leq P(X \leq y + \epsilon)$, which implies :

$$\begin{aligned} P(Y \leq y) &\leq P(X \leq y + \epsilon) + P(Y \leq y, X > y + \epsilon) \\ &\leq P(X \leq y + \epsilon) + P(Y - X \leq y - X, y + \epsilon < X) \\ &\leq P(X \leq y + \epsilon) + P(Y - X \leq y - X, y - X < -\epsilon). \end{aligned}$$

Moreover, we have $P(Y - X \leq y - X, y - X < -\epsilon) \leq P(Y - X < -\epsilon)$ which implies :

$$P(Y \leq y) \leq P(X \leq y + \epsilon) + P(Y - X < -\epsilon).$$

By definition of a probability, we have $P(Y - X > \epsilon) \geq 0$. Therefore,

$$P(Y \leq y) \leq P(X \leq y + \epsilon) + P(Y - X < -\epsilon) + P(Y - X > \epsilon)$$

which leads to the equation 1 and concludes the proof. \square

Theorem 2. (Convergence in probability implies convergence in distribution) Assume that X_1, X_2, \dots is a sequence of real-valued random variables. If

$$X_n \xrightarrow{P} X$$

then

$$X_n \xrightarrow{D} X.$$

Proof. Assume that $\{F_n\}_{n \geq 0}$ are the cumulative distribution functions of X_1, X_2, \dots . Assume that X is a real-valued random variable with cumulative distribution function F and that $X_n \xrightarrow{P} X$. Let $x \in \mathbb{R}$ be a point at which F is continuous. By the definition 1 we must prove that

$$\lim_{n \rightarrow \infty} F_n(X_n) = F(x). \quad (2)$$

Let $\epsilon > 0$.

First, we first apply the theorem 1 to the random variables X_n and X . The equation 1 implies:

$$P(X_n \leq x) \leq P(X \leq x + \epsilon) + P(|X_n - X| > \epsilon).$$

By hypothesis, the sequence X_1, X_2, \dots converges in probability to X , which implies:

$$\lim_{n \rightarrow \infty} P(X_n \leq x) \leq P(X \leq x + \epsilon).$$

since $\lim_{n \rightarrow \infty} P(|X_n - X| > \epsilon) = 0$. By definition of the cumulative distribution function F , this implies:

$$\lim_{n \rightarrow \infty} P(X_n \leq x) \leq F(x + \epsilon). \quad (3)$$

Secondly, we consider the theorem 1 and derive a similar inequality. In the equation 1, we exchange the random variables X and Y :

$$P(X \leq y) \leq P(Y \leq y + \epsilon) + P(|X - Y| > \epsilon)$$

and we apply the inequality to the real $y - \epsilon$; this leads to:

$$P(X \leq y - \epsilon) \leq P(Y \leq y) + P(|X - Y| > \epsilon).$$

Back to the theorem 2, we apply the previous inequality to the random variables X and X_n , at the point x . We obtain:

$$P(X \leq x - \epsilon) \leq P(X_n \leq x) + P(|X - X_n| > \epsilon).$$

By hypothesis, the sequence X_1, X_2, \dots converges in probability to X , which implies:

$$P(X \leq x - \epsilon) \leq \lim_{n \rightarrow \infty} P(X_n \leq x).$$

By definition of the cumulative distribution function F , this implies:

$$F(x - \epsilon) \leq \lim_{n \rightarrow \infty} P(X_n \leq x). \quad (4)$$

Combining the inequalities 3 and 4, we obtain:

$$F(x - \epsilon) \leq \lim_{n \rightarrow \infty} P(X_n \leq x) \leq F(x + \epsilon).$$

The function F is, by hypothesis, continuous at the point x which implies that we can take the limit of the previous inequality when $\epsilon \rightarrow 0$. This implies:

$$F(x) \leq \lim_{n \rightarrow \infty} P(X_n \leq x) \leq F(x)$$

which leads to the equation 2 and concludes the proof. \square

4 Delta method

Theorem 3. (Slutsky's theorem) Let X_n and Y_n be two sequences of real random variables such that:

$$X_n \xrightarrow{D} X \quad \text{and} \quad Y_n \xrightarrow{P} c,$$

where c is a non-random constant. Therefore,

$$X_n Y_n \xrightarrow{D} cX.$$

Proof. TODO □

Theorem 4. (Continuous mapping) Let X_n be a sequence of real random variables such that:

$$X_n \xrightarrow{D} X.$$

Assume that g is a continuous function on \mathbb{R} . Therefore:

$$g(X_n) \xrightarrow{D} g(X).$$

Proof. TODO □

Theorem 5. (Delta-method) Assume that X_1, X_2, \dots is a sequence of real-valued random variables so that

$$\sqrt{n}(X_n - \theta) \xrightarrow{D} \mathcal{N}(0, \sigma^2). \quad (5)$$

Assume that g is a real function. Let $\theta \in \mathbb{R}$. Suppose that $g \in C^1(\mathbb{R})$ and that $g'(\theta) \neq 0$. Therefore,

$$\sqrt{n}(g(X_n) - g(\theta)) \xrightarrow{D} \mathcal{N}(0, \sigma^2 g'(\theta)^2).$$

Proof. *TODO : fix this proof, which is wrong.*

The Taylor expansion of g at the point θ implies that there exists a $\tilde{\theta}$ between X_n and θ such that:

$$g(X_n) = g(\theta) + g'(\tilde{\theta})(X_n - \theta) + O((X_n - \theta)^2), \quad X_n \rightarrow \theta.$$

This implies:

$$\sqrt{n}(g(X_n) - g(\theta)) = g'(\tilde{\theta})\sqrt{n}(X_n - \theta) + O((X_n - \theta)^2), \quad X_n \rightarrow \theta.$$

Let Y_n and Z_n be the sequences of random variables defined by the equations:

$$Y_n = g'(\tilde{\theta})\sqrt{n}(X_n - \theta) \text{ and } Z_n = O((X_n - \theta)^2).$$

The properties of the gaussian distribution and the equation 5 imply:

$$Y_n = g'(\tilde{\theta})\sqrt{n}(X_n - \theta) \xrightarrow{D} \mathcal{N}(0, \sigma^2 g'(\tilde{\theta})^2).$$

Furthermore,

$$Z_n = O((X_n - \theta)^2) \xrightarrow{D} 0.$$

In order to conclude the proof, we apply Slutsky's theorem to Y_n and Z_n . □

5 References

Most of the results in this document can be found in [?].