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title: "Appendix: Code"
output: html_notebook
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```{r}
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```
library(rstan)
library(bayesplot)
library(sp)
data(meuse)
library('plotly')
library(lattice)
source("~/Documents/BU/MA578/bslm.R")
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```
compute_distances <- function(x,y){
 n<- length(x)
 D = matrix(0,nrow=n,ncol=n)
 for(i in 1:n){
 for(j in 1:n){
 p1 = c(x[i],y[i])
 p2 = c(x[j],y[j])
 D[i,j] <- sqrt(sum((p1 - p2)^2))
 }
 }
 return(D)
}
```

```
D <- compute_distances(meuse$x,meuse$y)
p <- plot_ly(z=~D)
layout(p,title = "Distances Between Sites", xaxis=list(title =
'Site'),yaxis=list(title = 'Site'))
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```
p<-plot_ly(x= meuse$x,y=meuse$y,z = ~log(meuse$lead))
layout(p,title = "Log lead concentration as a function of location", scene =
list(xaxis=list(title = 'x location'),yaxis=list(title = 'y
location'),zaxis=list(title = 'Log lead')))
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Fix tau2 to be 1 to understand impact of phi.

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phi=1.1
H <- 1*exp(-D/phi)
p <- plot_ly(z=~H)
layout(p,title = "Spatial Correlation for Phi = 1", xaxis=list(title =
'Site'),yaxis=list(title = 'Site'))

phi=300
H <- 1*exp(-D/phi)
p <- plot_ly(z=~H)
layout(p,title = "Spatial Correlation for Phi = 300", xaxis=list(title =
'Site'),yaxis=list(title = 'Site'))

phi=3000
H <- 1*exp(-D/phi)
p <- plot_ly(z=~H)
layout(p,title = "Spatial Correlation for Phi = 3000", xaxis=list(title =
'Site'),yaxis=list(title = 'Site'))
```

```{r}

xyplot(log(lead) ~ dist.m | soil, data = meuse,
 panel = function(...) {
 panel.xyplot(type = "p", ...)
 panel.lmline(lty = 2, ...)
 },xlab='Distance to River (meters)',ylab='Log Lead
Concentration',main='Lead Concentration vs Distance to River by Soil Type')
xyplot(log(lead) ~ dist.m , data = meuse,
 panel = function(...) {
 panel.xyplot(type = "p", ...)
 panel.lmline(lty = 2, ...)
 },xlab='Distance to River (meters)',ylab='Log Lead
Concentration',main='Lead Concentration vs Distance to River')
```

Set up design matrices
```{r}
y <- log(meuse$lead)
X <- model.matrix(~ -1 + soil + soil:dist.m, data = meuse)

J <- length(levels(meuse$soil))
n <- nrow(X)
p <- ncol(X) / J
site_j <- aggregate(. ~ soil, data = meuse, length)
soil_site <- data.frame(group = site_j$soil,
 coef = c(rep("_intercept", J), rep("_distance",
J)))
X_beta <- model.matrix(~ -1 + coef , data = soil_site)

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Metropolis-Gibbs sampling procedure. Use 4 chains, 10000 simulations and warmup length of 2000.

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```{r}
nchains <- 4
nsims <- 10000
warmup <- 2000
gibbs_chain <- matrix(0, ncol = 12, nrow = nsims - warmup)
param_names <- c("BetaA0", "BetaB0", "BetaC0", "BetaA1", "BetaB1", "BetaC1",
 "Alpha1", "Alpha0", "Kappa2", "Sigma2", "Tau2",
 "Phi")
colnames(gibbs_chain) <- param_names
nparams <- length(param_names)
sims <- mcmc_array(nsims - warmup, nchains, param_names)

Initialize starting values for beta as MLE estimates for log(lead) ~
f(distance to river), not controlling for soil type.
n <- length(y)
beta_fit <- lm(y ~ X - 1)
beta <- coef(beta_fit)

Assuming equal weight of deviance for fit of log lead concentration on error
in estimation and error from spatial component.
sigma2 <- deviance(beta_fit) / (2*n)
tau2 <- deviance(beta_fit) / (2*n)

Set parameters of inverse gamma for tau2
alpha_tau <- 10
beta_tau <- 4

Compute alpha MLE estimates and use deviance to get initial starting values
for kappa.
alpha_fit <- lm(beta ~ X_beta - 1)
kappa2 <- deviance(alpha_fit) / nrow(X_beta)

Set parameters of inverse gamma for kappa
alpha_kappa <- 10
beta_kappa <- 4

Set parameters of inverse gamma for alpha phi
alpha_phi <- 10
phi <- 1/rgamma(1, alpha_phi, alpha_phi)

Flat prior on alpha: normal distribution with 0 mean and infinite variance.
alpha_prior <- list(mean=rep(0,2), precision=rep(0,2))
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Function for H(phi)
```{r}
H <- function(phi){
 return(exp(-D/phi))
}
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Function for random walk on phi.
```{r}
phi_jumping <- function(phi_mn){
 phi <- rnorm(1,mean=phi_mn, sd=0.1)
 return(phi)
}
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Function for log posterior.
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logposterior <- function(phi,sigma2,delta,beta,kappa2,tau2,alpha) {
 logpost <-0
 logposterior<- logpost - log(sqrt(sigma2^n)) + (-1/(2*sigma2))*crossprod(y-X
%% beta-delta) - log(sqrt(det(tau2*(H(phi))))) - (1/2) * t(delta) %%
solve(tau2 * H(phi)) %% delta -
 log(sigma2) - (alpha_phi+1)*log(phi) - (alpha_phi/phi) -
log(sqrt(kappa2^6)) - (1/(2*kappa2)) * crossprod(beta- X_beta %% alpha)
- (alpha_tau+1)*log(tau2) - (beta_tau/tau2) - (alpha_kappa+1)*log(kappa2) -
(beta_kappa/kappa2)

 return(logposterior)
}
```

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```{r}
for (chain in 1:nchains) {
 gibbs_chain <- matrix(0,ncol = 12,nrow = nsims - warmup)
 for (it in 1:nsims) {
 # 1. Sample alpha
 alpha <- bsml_sample1(beta, X_beta, sqrt(kappa2), alpha_prior)

 # 2. Sample delta
 delta <- rnorm(n,mean=rep(0,n),sd = chol((tau2) * H(phi)))

 # 3. Sample beta
 beta <- bsml_sample1(y - delta, X, sqrt(sigma2),
 list(mean = X_beta %% alpha, precision = 1 /
kappa2))

 # 4. Sample sigma^2

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sigma2 <- (n * (1/n * crossprod(y - X %%% beta - delta))/rchisq(1,n))
sigma2 <- as.vector(sigma2)

5. Sample tau^2
nu <- alpha_tau + n/2
s2 <- (2*beta_tau + t(delta) %%% solve(H(phi)) %%% delta)/(n+2*alpha_tau)
tau2 <- (nu * (s2/rchisq(1,nu)))
tau2 <- as.vector(tau2)

6. Sample kappa^2
nAlpha <-6
nu <- alpha_kappa + nAlpha/2
s2 <- (2*beta_kappa+ crossprod(beta - X_beta %%% alpha))/
(nAlpha+2*alpha_kappa)
kappa2 <- (nu * (s2/rchisq(1,nu)))
kappa2 <- as.vector(kappa2)

7. Sample phi using random walk
phi_star <- phi_jumping(phi)

logR <- logposterior(phi_star,sigma2,delta,beta,kappa2,tau2,alpha) -
logposterior(phi,sigma2,delta,beta,kappa2,tau2,alpha)
if (logR >= 0 || log(runif(1)) <= logR) {
 phi <- phi_star
}

Save iterations
if(it > warmup){
 sims[it - warmup, chain,] <- c(beta,alpha,kappa2,sigma2,tau2,phi)
}

}
}

...

```{r}
monitor(sims)
mcmc_trace(sims)
mcmc_acf(sims)
mcmc_dens_overlay(sims)
...

```{r}
v<-rbind(sims[,1,],sims[,2,],sims[,3,],sims[,4,])
t(apply(v,2, function(x) quantile(x, c(.025,.25,.5,.75,.975))))
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