

The dangers of using Seasonal Adjustment and other filters in Econometrics

Some economic and environmental examples

44TH INTERNATIONAL SYMPOSIUM ON FORECASTING, DIJON

Antonio García-Ferrer ¹ Marcos Bujosa ²

¹Dpto. de Análisis Económico: Economía Cuantitativa.
Universidad Autónoma de Madrid

²Dpto. de Análisis Económico y Economía Cuantitativa.
Universidad Complutense de Madrid

June 30 – July 3, 2024

1 Introduction

- When using seasonally unadjusted data, how can we decide what is the optimal seasonal adjustment to use?
 - Not theoretical point of view
- Do we have sensible statistical tools to discriminate among the different available alternatives?
- Knowing that the *estimated* components are not *observable*, is it enough to pay attention to just the component of interest and forget about the remaining ones?
- Is the ideal property of *orthogonality* among the different component reasonably fulfilled?
- How potential *outliers* and other variants of *intervention* analysis affect final estimated components?

1 Introduction

- When using seasonally unadjusted data, how can we decide what is the optimal seasonal adjustment to use?
 - Not theoretical point of view
- Do we have sensible statistical tools to discriminate among the different available alternatives?
- Knowing that the *estimated* components are not *observable*, is it enough to pay attention to just the component of interest and forget about the remaining ones?
- Is the ideal property of *orthogonality* among the different component reasonably fulfilled?
- How potential *outliers* and other variants of *intervention* analysis affect final estimated components?

1 Introduction

- When using seasonally unadjusted data, how can we decide what is the optimal seasonal adjustment to use?
 - Not theoretical point of view
- Do we have sensible statistical tools to discriminate among the different available alternatives?
- Knowing that the *estimated* components are not *observable*, is it enough to pay attention to just the component of interest and forget about the remaining ones?
- Is the ideal property of *orthogonality* among the different component reasonably fulfilled?
- How potential *outliers* and other variants of *intervention* analysis affect final estimated components?

2 Traditional approach

$$y_t = T_t + C_t + S_t + e_t$$

3

Small empirical exercise

Four monthly time series pertaining to the Spanish economic CLI used in: <http://uam-ucm-economic-indicators.es/>

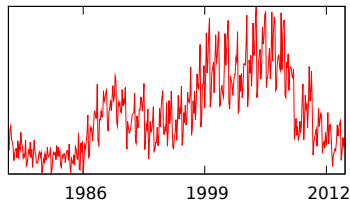
- CAR REGISTRATIONS
- HOUSING STARTS
- CEMENT CONSUMPTION
- TRUCKS

From 1978M01 to 2013M12

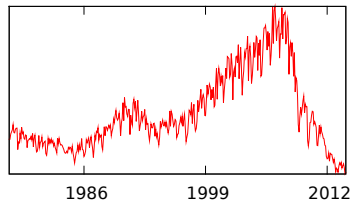
4

Small empirical exercise

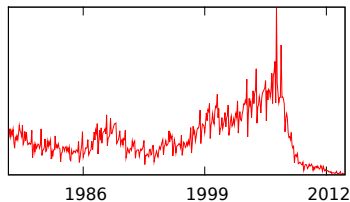
CARS



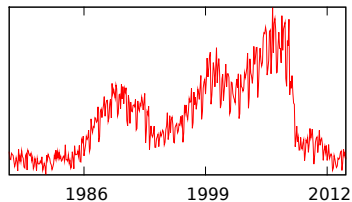
CEMENT



HOUSES



TRUCKS



5

Several signal extraction methodologies

Using several model-based signal extraction methodologies, namely

- SEATS-TRAMO
- X-12 ARIMA
- Linear Dynamic Harmonic Regression ([Bujosa et al., 2007](#))

Disclaimer and explanation of the posterior empirical results

6 Dynamic Harmonic Regression Model

The DHR model consists of several unobserved components plus an irregular stationary zero mean component $e = \{e_t\}_{t \in \mathbb{Z}}$

$$y = \sum_{j=0}^R s^j + e. \quad (1)$$

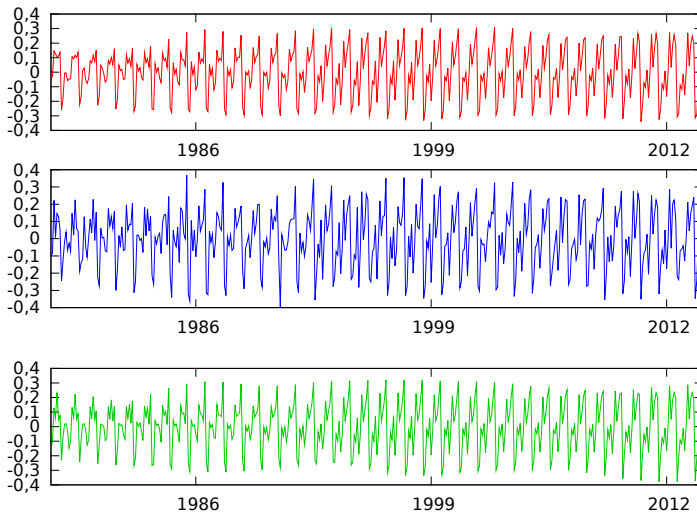
- DHR components $s^j = \{s_t^j\}_{t \in \mathbb{Z}}$ are oscillatory

$$s_t^j = a_t^j \cos(\omega_j t) + b_t^j \sin(\omega_j t), \quad (2)$$

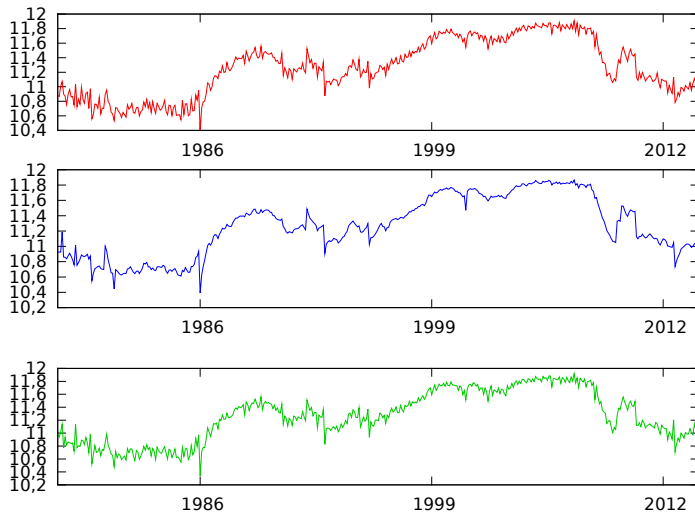
where frequency ω_j is associated to the j -th component.

- Oscillations are modulated by two GRW processes $a^j = \{a_t^j\}_{t \in \mathbb{Z}}$ and $b^j = \{b_t^j\}_{t \in \mathbb{Z}}$.
- $\omega_0 = 0$ corresponds to the trend (or zero frequency term).
- The model is fitted in the frequency domain.

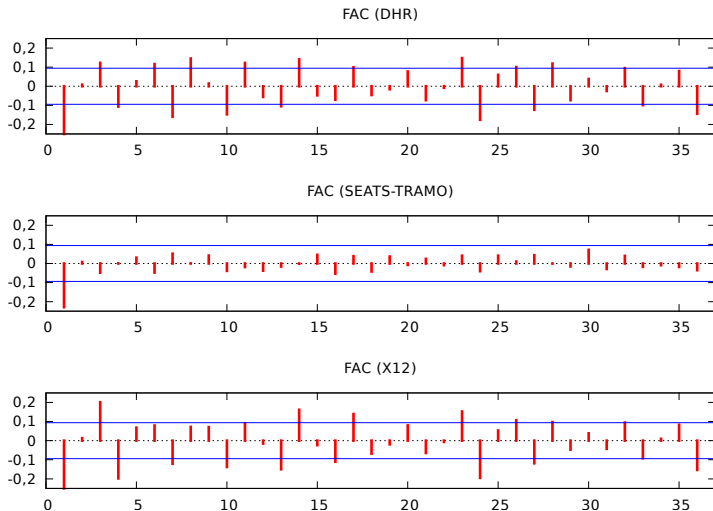
7 Car registrations Seasonal Factors: DHR, ST, X12



8 Seasonally adjusted Car registrations: DHR, ST, X12



9 FAC – First Difference of Seasonally adjusted Car registrations



10

Summary of tentative results of the four series

- Outlier detection plus other interventions as easter effects and calendar effects are crucial in the estimation of unobserved components models
- As a matter of fact when you don't use this option in SEATS-TRAMO there is evidence of seasonality in the SA series
- Using outlier detection plus easter and calendar effects produce considerable reduction in the estimated residual variances ranging from 21% to 31%

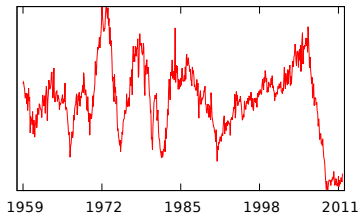
11

Results from a Stock & Watson data base

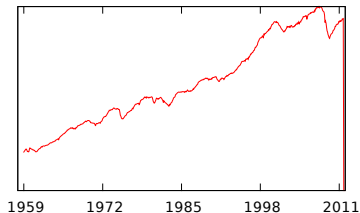
- Housing starts
- IPI
- Money supply – M1
- Retail sales

12 Results from a Stock & Watson data base

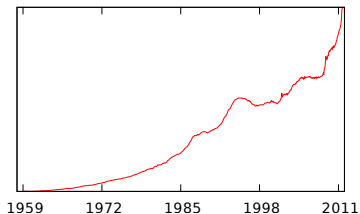
HOUSES



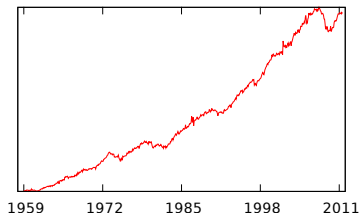
IPI



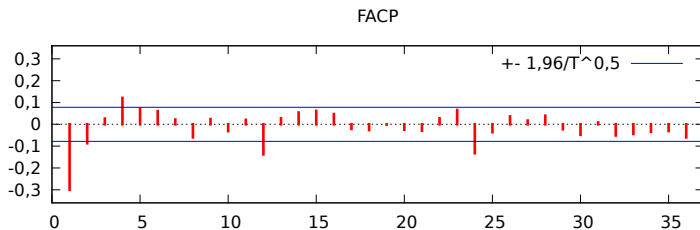
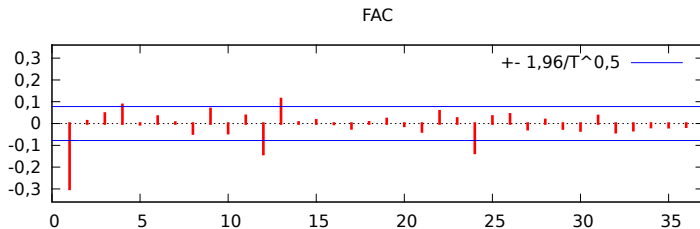
M1



RETAIL SALES

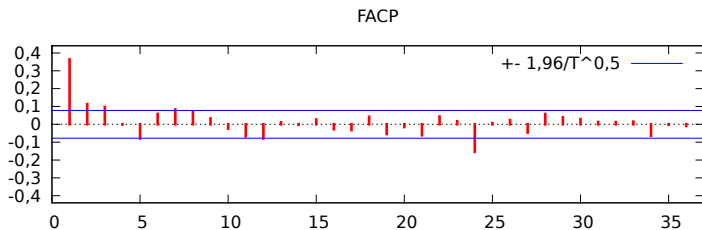
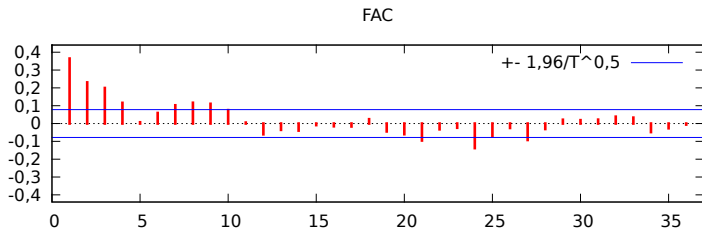


13 Results from a Stock & Watson data base: Housing starts

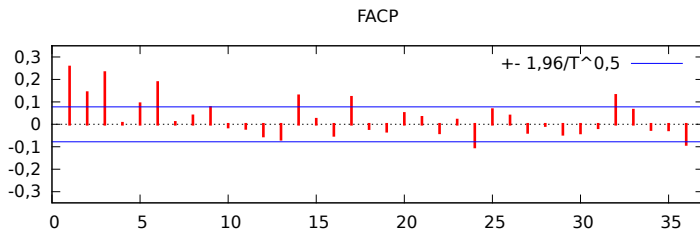
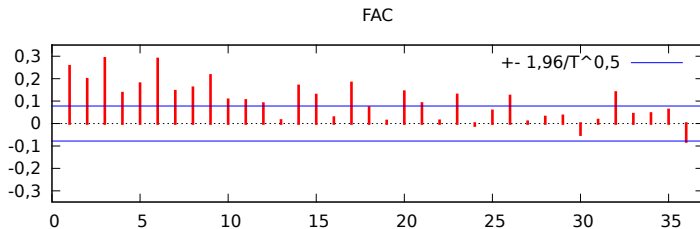


14

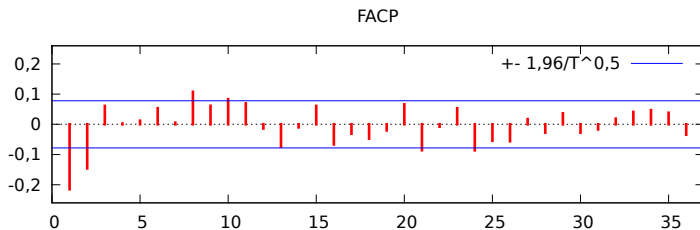
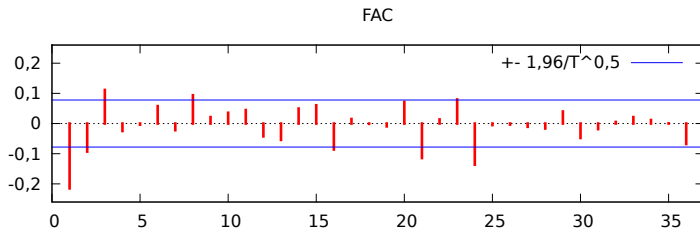
Results from a Stock & Watson data base: IPI



15 Results from a Stock & Watson data base: Money supply



16 Results from a Stock & Watson data base: Retail sales



17 Hodrick–Prescott filter

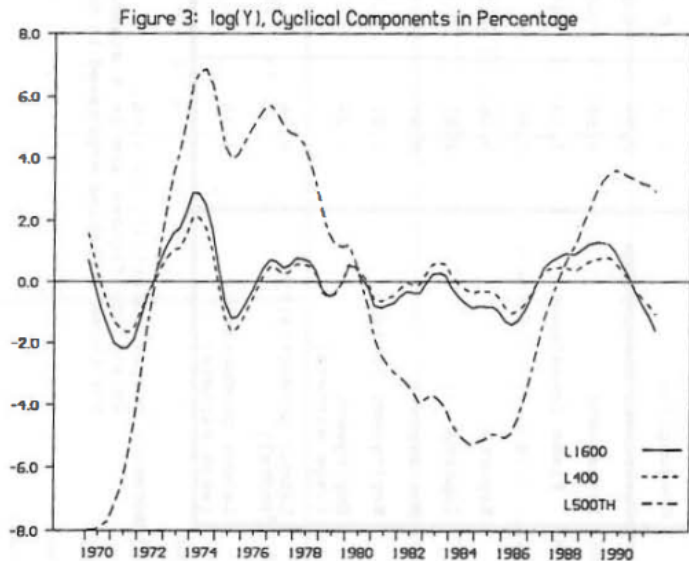
Hodrick and Prescott (1981, 1997); Whittaker (1922)

$$y_t = \tau_t + c_t + \epsilon_t$$

Given a positive λ , there is a trend component τ that solves

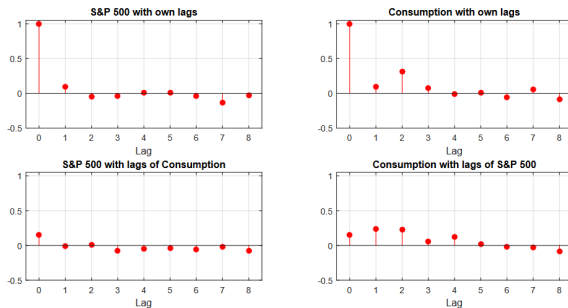
$$\min_{\tau} \left(\sum_{t=1}^T (y_t - \tau_t)^2 + \lambda \sum_{t=2}^{T-1} [(\tau_{t+1} - \tau_t) - (\tau_t - \tau_{t-1})]^2 \right)$$

Why $\lambda = 1600$?

18 Hodrick–Prescott filter

19 Why You Should Never Use the Hodrick-Prescott Filter (Hamilton, 2018)

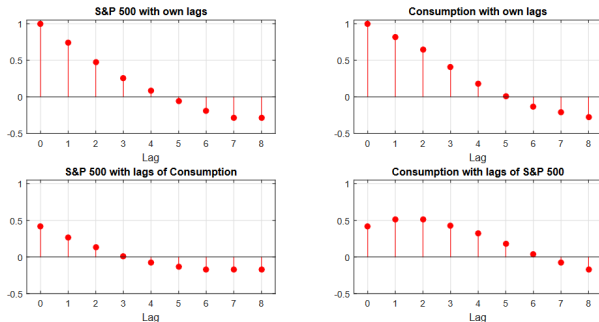
Figure 2. Autocorrelations and cross-correlations for first-difference of stock prices and real consumption spending.



Notes to Figure 2. Upper left: autocorrelations of log growth rate of end-of-quarter value for S&P 500. Upper right: autocorrelations of log growth rate of real consumption spending. Lower panels: cross correlations.

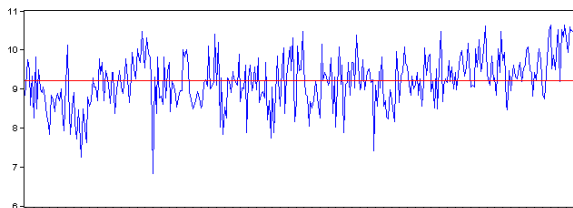
20 Why You Should Never Use the Hodrick-Prescott Filter (Hamilton, 2018)

Figure 3. Autocorrelations and cross-correlations for HP cyclical component of stock prices and real consumption spending.

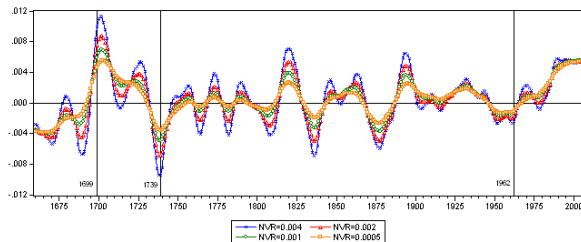


Notes to Figure 3. Upper left: autocorrelations of HP cycle for log of end-of-quarter value for S&P 500. Upper right: autocorrelations of HP cycle for log of real consumption spending. Lower panels: cross correlations.

21 The Central England Temperature 1659–2007 (CET)

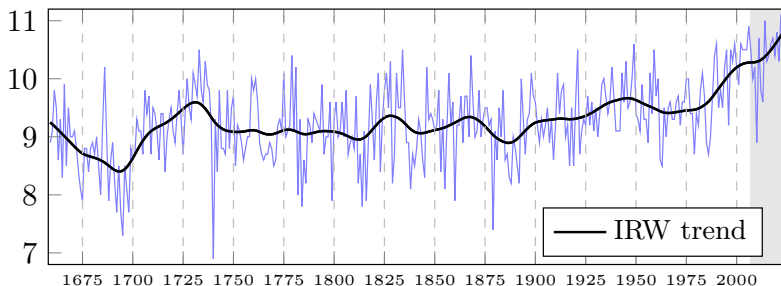
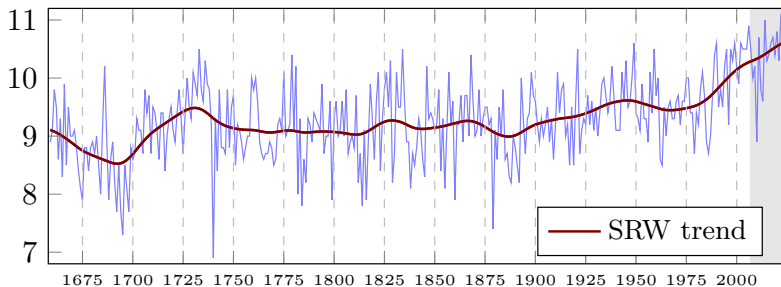


Alternative Temperature Cycles and Bayesian Turning Points

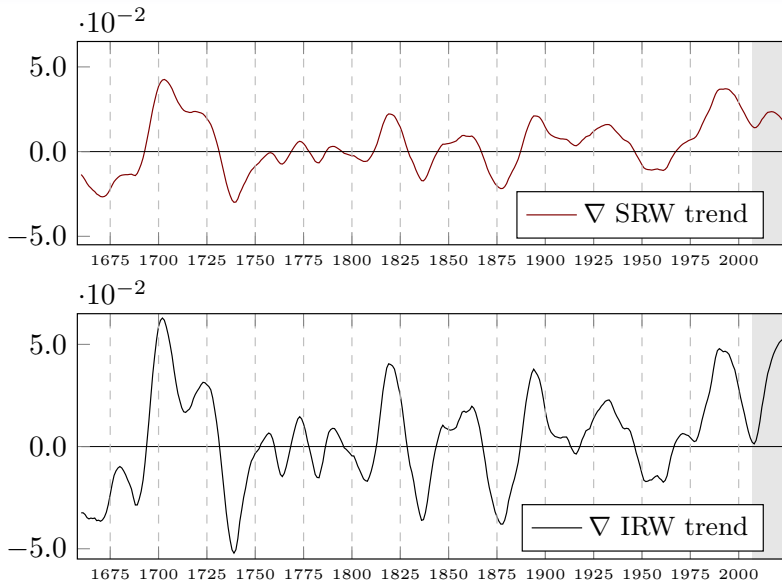


(Moreno et al., 2013; García-Ferrer et al., 2008)

22 The Central England Temperature 1659–2023 (CET)

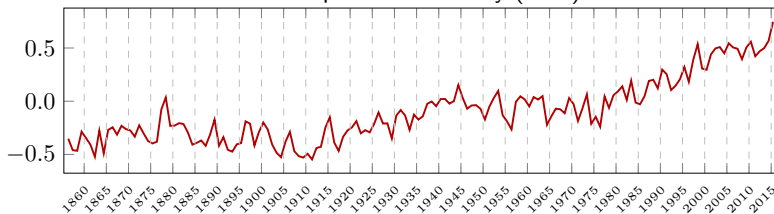


23 The Central England Temperature 1659–2023 (CET)

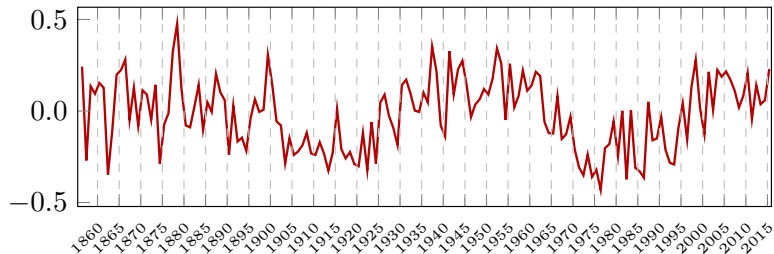


24 Modelling of Global Climate Change (Young et al., 2021)

Global Temperature Anomaly (GTA)



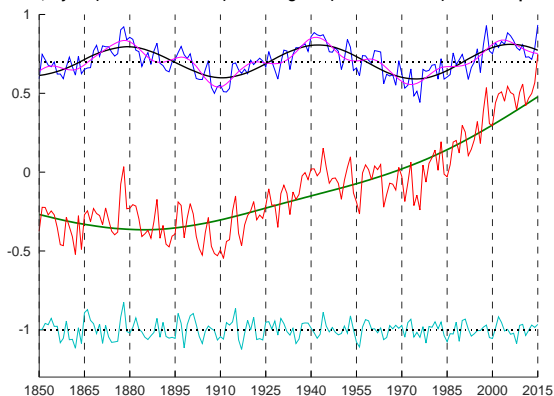
Atlantic Multidecadal Oscillation (AMO)



25 Have AMO and GTA a common 63-years cycle?

DHR components for GTA

Trend, Cycle (shifted +0.7 units) and irregular (shifted -1 units) DHR components

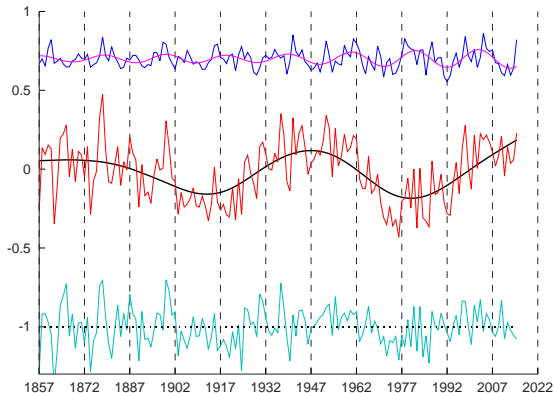


$$GTA = T + S^{63} + S^{21} + \sum(\text{other harmonics}) + Irreg$$

26 Have AMO and GTA a common 63-years cycle?

DHR Trend-cycle component for AMO

Trend, cycle (shifted +0.7 units) and irregular (shifted -1 units) DHR components

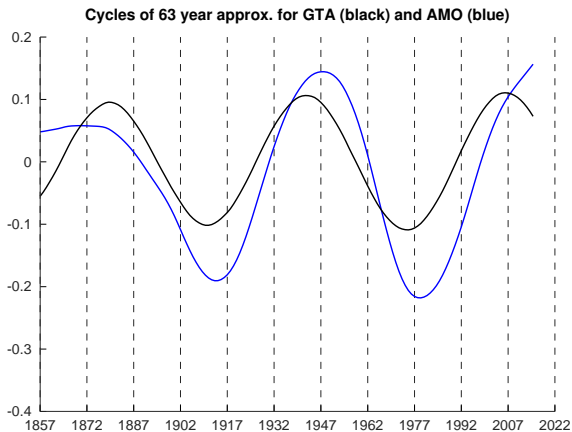


$$AMO = T + S^{21} + \sum(\text{other harmonics}) + Irreg$$

27 Have AMO and GTA a common 63-years cycle?

Not clear

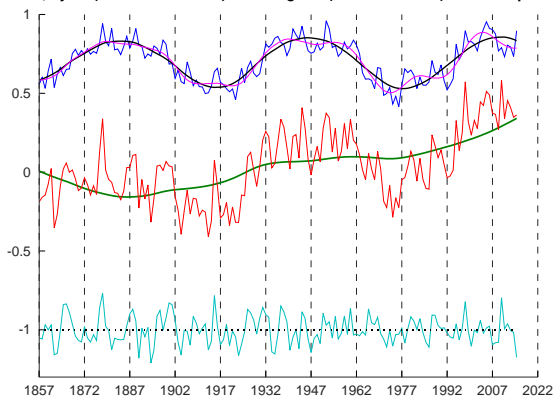
GTA has a periodic cycle, but not AMO



28 Have original AMO and GTA a common 63-years cycle?

DHR components for “original” AMO data

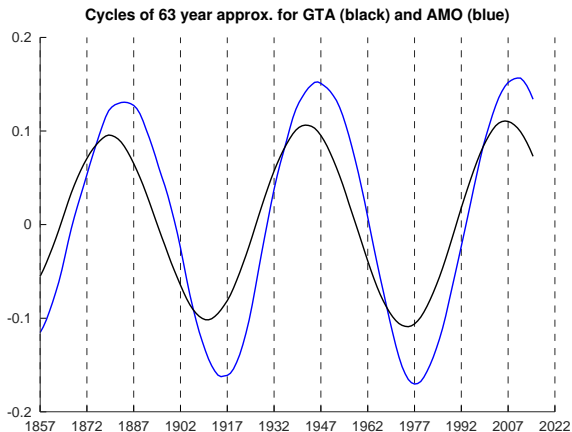
Trend, cycle (shifted +0.7 units) and irregular (shifted -1 units) DHR components



$$AMO_{\text{with trend}} = T + S^{63} + S^{21} + \sum(\text{other harmonics}) + Irreg$$

29 Have the “original” AMO and GTA a common cycle?

They seem to have a common cycle
(as suggested in Professor Young’s article)



30 Number of confirmed cases at 3/22/2020

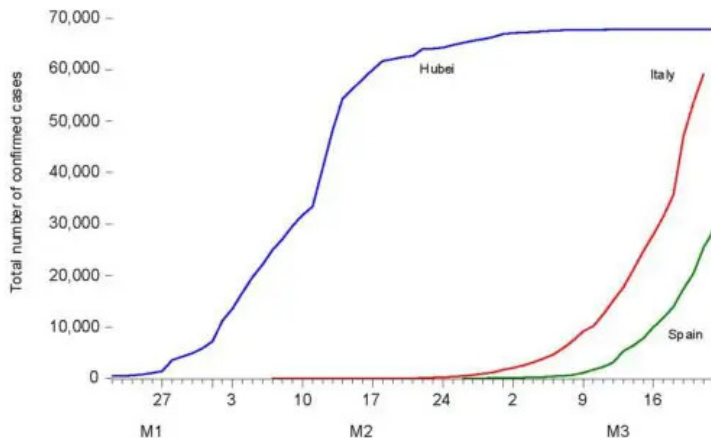


Figure 1: Number of confirmed cases at 3/22/2020

31 Observed contagions and forecasts in Spain

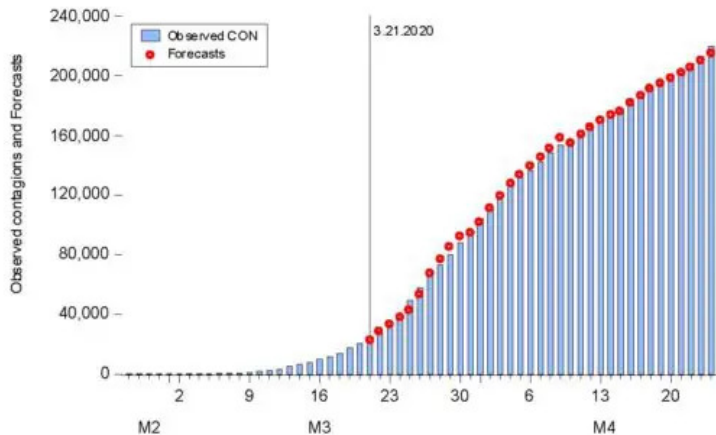


Figure 2: Observed contagions and Forecasts in Spain

32 Observed deaths and forecasts in Spain

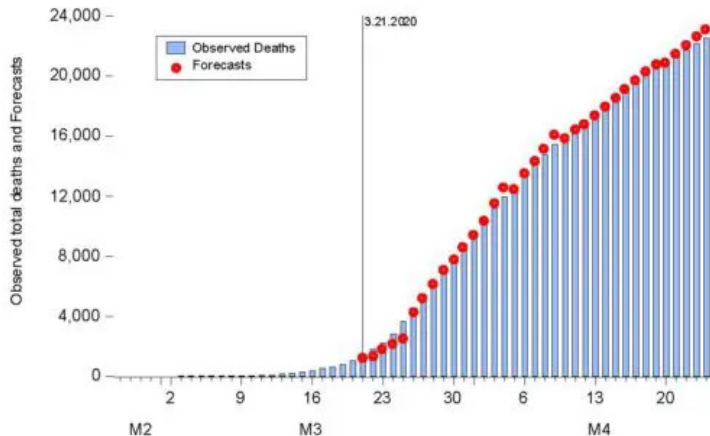


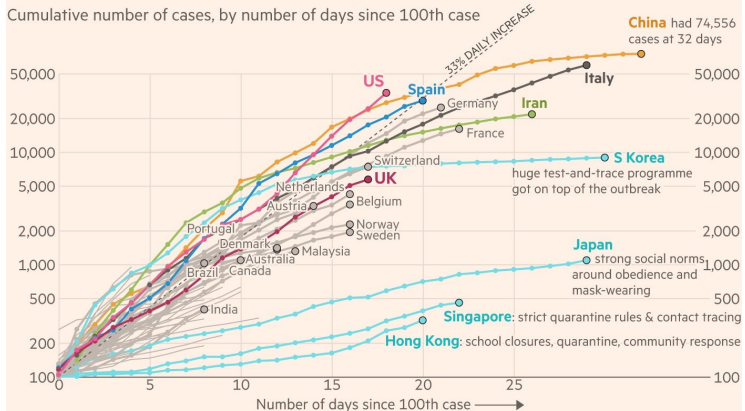
Figure 3: Observed Deaths and Forecasts in Spain

33

Coronavirus trajectories

Most western countries are on the same coronavirus trajectory. Hong Kong and Singapore have limited the spread; Japan and S Korea have slowed it

Cumulative number of cases, by number of days since 100th case



FT graphic: John Burn-Murdoch / @jburnmurdoch

Source: FT analysis of Johns Hopkins University, CSSE; Worldometers. Data updated March 22, 19:00 GMT

© FT

Bujosa, M., García-Ferrer, A., and Young, P. C. (2007). Linear dynamic harmonic regression. *Comput. Stat. Data Anal.*, **52**(2), 999–1024. ISSN 0167-9473.

García-Ferrer, A., Young, P., and Bujosa, M. (2008). Central england temperature: Analysis and forecasting. In *The 28th International Symposium on Forecasting*.

Hamilton, J. D. (2018). Why You Should Never Use the Hodrick-Prescott Filter. *The Review of Economics and Statistics*, **100**(5), 831–843. ISSN 0034-6535.

URL https://doi.org/10.1162/rest_a_00706

Hodrick, R. J. and Prescott, E. (1981). Post-War U.S. Business Cycles: An Empirical Investigation. Discussion Papers 451, Northwestern University, Center for Mathematical Studies in Economics and Management Science.

Hodrick, R. J. and Prescott, E. C. (1997). Postwar u.u. business cycles: An empirical investigation. *Journal of Money, Credit and Banking*, **29**, 1–16.

Moreno, E., Javier Girón, F., and García-Ferrer, A. (2013). A consistent on-line bayesian procedure for detecting change points. *Environmetrics*, **24**(5), 342–356.

Whittaker, E. T. (1922). On a new method of graduation. *Proceedings of the Edinburgh Mathematical Society*, **41**, 63–75.

Young, P. C., Allen, P. G., and Bruun, J. T. (2021). A re-evaluation of the earth's surface temperature response to radiative forcing. *Environmental Research Letters*, **16**(5), 054068.

URL <https://dx.doi.org/10.1088/1748-9326/abfa50>