The dangers of using Seasonal Adjustment and other filters in Econometrics

Some economic and environmental examples

44TH INTERNATIONAL SYMPOSIUM ON FORECASTING, DIJON

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Introduction O

1 Introduction

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- When using seasonally unadjusted data, how can we decide what is the optimal seasonal adjustment to use?
 - Not theoretical point of view

Methodologies

- Do we have sensible statistical tools to discriminate among the different available alternatives?
- Knowing that the *estimated* components are not *observable*, is it enough to pay attention to just the component of interest and forget about the remaining ones?
- Is the ideal property of *orthogonality* among the different component reasonably fulfilled?
- How potential *outliers* and other variants of *intervention* analysis affect final estimated components?

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 2
 Traditional approach

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$$y_t = T_t + C_t + S_t + e_t$$

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Four monthly time series pertaining to the Spanish economic CLI used in: http://uam-ucm-economic-indicators.es/

- CAR REGISTRATIONS
- HOUSING STARTS
- CEMENT CONSUMPTION
- TRUCKS

From 1978M01 to 2013M12

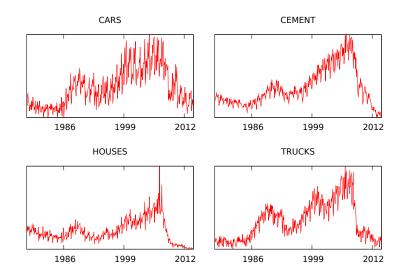
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5 Several signal extraction methodologies

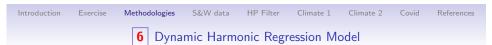
Using several model-based signal extraction methodologies, namely

- SEATS-TRAMO
- X-12 ARIMA
- Linear Dynamic Harmonic Regression (Bujosa et al., 2007)

Disclaimer and explanation of the posterior empirical results



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The DHR model consists of several unobserved components plus an irregular stationary zero mean component $e = \{e_t\}_{t \in \mathbb{Z}}$

$$y = \sum_{j=0}^{R} s^j + e. \tag{1}$$

ullet DHR components $s^j = \{s^j_t\}_{t\in\mathbb{Z}}$ are oscillatory

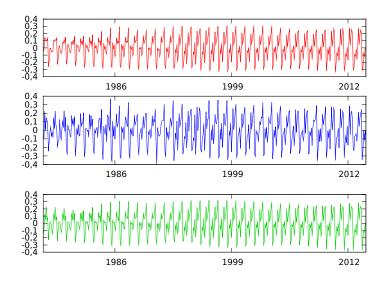
$$s_t^j = a_t^j \cos(\omega_j t) + b_t^j \sin(\omega_j t), \tag{2}$$

where frequency ω_j is associated to the j-th component.

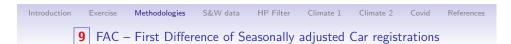
- Oscillations are modulated by two GRW processes $a^j=\{a^j_t\}_{t\in\mathbb{Z}}$ and $b^j=\{b^j_t\}_{t\in\mathbb{Z}}.$
- $\omega_0 = 0$ corresponds to the trend (or zero frequency term).
- The model is fitted in the frequency domain.

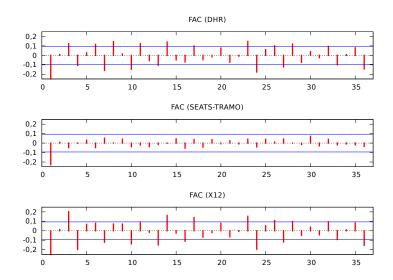


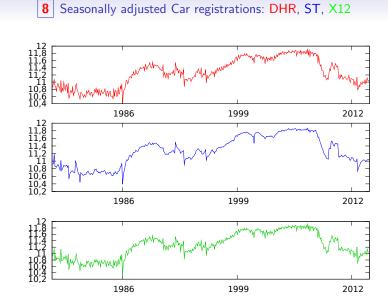
7 Car registrations Seasonal Factors: DHR, ST, X12



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10 Summary of tentative results of the four series

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- Outlier detection plus other interventions as easter effects and calendar effects are crucial in the estimation of unobserved components models
- As a matter of fact when you don't use this option in SEATS-TRAMO there is evidence of seasonality in the SA series
- Using outlier detection plus easter and calendar effects produce considerable reduction in the estimated residual variances ranging from 21% to 31%

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11 Results from a Stock & Watson data base

- Housing starts
- IPI
- Money supply M1
- Retail sales

-0,1 -0,2 -0,3

5

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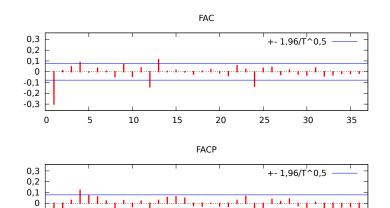
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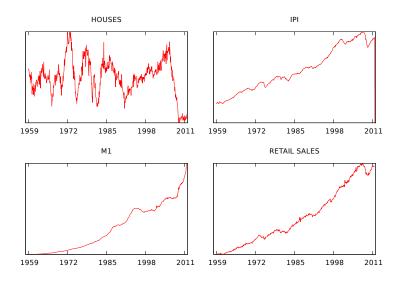
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13 Results from a Stock & Watson data base: Housing starts



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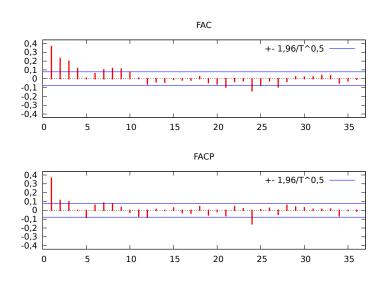
12 Results from a Stock & Watson data base



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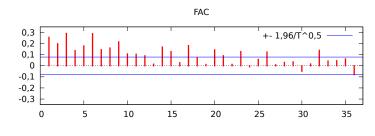
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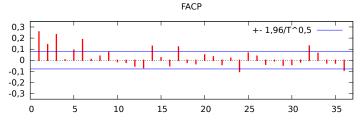
14 Results from a Stock & Watson data base: IPI



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Results from a Stock & Watson data base: Money supply





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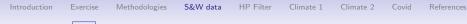
Hodrick and Prescott (1981, 1997); Whittaker (1922)

$$y_t = \tau_t + c_t + \epsilon_t$$

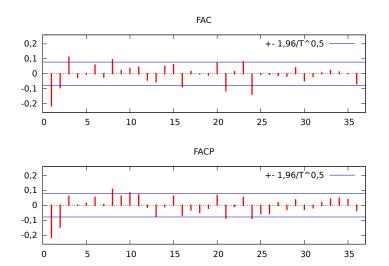
Given a positive λ , there is a trend component τ that solves

$$\min_{\tau} \left(\sum_{t=1}^{T} (y_t - \tau_t)^2 + \lambda \sum_{t=2}^{T-1} \left[(\tau_{t+1} - \tau_t) - (\tau_t - \tau_{t-1}) \right]^2 \right)$$

Why
$$\lambda = 1600$$
?



Results from a Stock & Watson data base: Retail sales



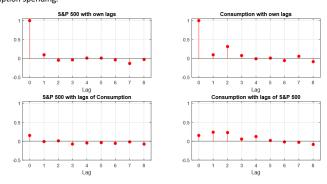
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Methodologies

Figure 2. Autocorrelations and cross-correlations for first-difference of stock prices and real consumption spending.

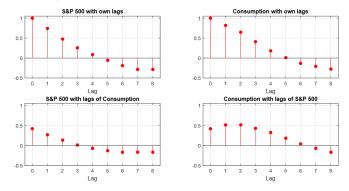


Notes to Figure 2. Upper left: autocorrelations of log growth rate of end-of-quarter value for S&P 500. Upper right: autocorrelations of log growth rate of real consumption spending. Lower panels: cross correlations.

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19 Why You Should Never Use the Hodrick-Prescott Filter (Hamilton, 2018)

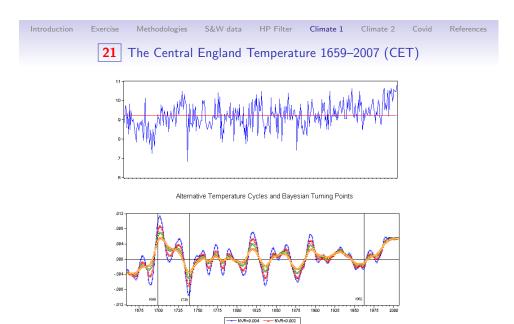
Figure 3. Autocorrelations and cross-correlations for HP cyclical component of stock prices and real consumption spending.



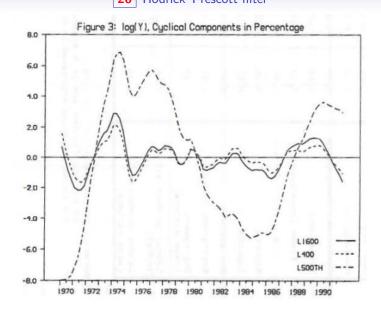
Notes to Figure 3. Upper left: autocorrelations of HP cycle for log of end-of-quarter value for S&P 500. Upper right: autocorrelations of HP cycle for log of real consumption spending. Lower panels: cross correlations.

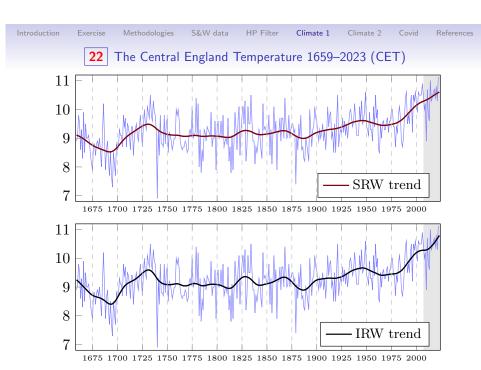
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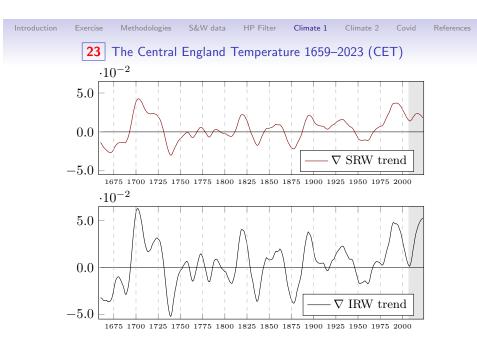






(Moreno et al., 2013; García-Ferrer et al., 2008)

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Persuade or inform?

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Global Temperature Anomaly (GTA)

O.5

O.0

Atlantic Multidecadal Oscillation (AMO)

O.5

O.0

O.5

O.0

Atlantic Multidecadal Oscillation (AMO)

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25 Have AMO and GTA a common 63-years cycle?

DHR components for GTA

Trend, Cycle (shifted +0.7 units) and irregular (shifted -1 units) DHR components

0.5

0.5

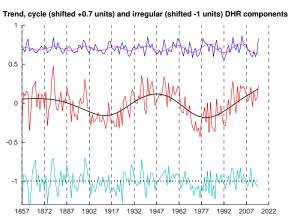
1850 1865 1880 1895 1910 1925 1940 1955 1970 1985 2000 2015

$$\label{eq:GTA} \begin{aligned} \textit{GTA} &= T + S^{63} + S^{21} + \sum (\text{other harmonics}) + Irreg \end{aligned}$$

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26 Have AMO and GTA a common 63-years cycle?

DHR Trend-cycle component for AMO



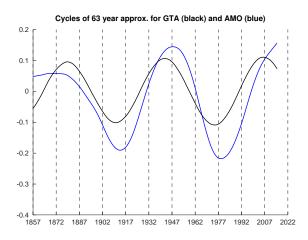
$$AMO = T + S^{21} + \sum$$
 (other harmonics) $+ Irreg$

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27 Have AMO and GTA a common 63-years cycle?

Not clear GTA has a periodic cycle, but not AMO



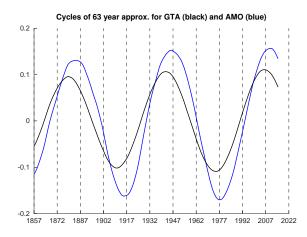
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29 Have the "original" AMO and GTA a common cycle?

They seem to have a common cycle

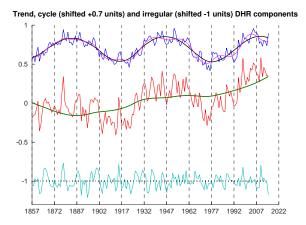
(as suggested in Professor Young's article)



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28 Have original AMO and GTA a common 63-years cycle?

DHR components for "original" AMO data

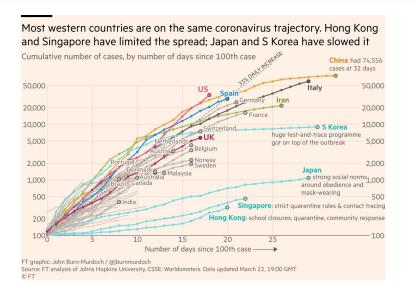


$$AMO_{
m with\ trend} = T + S^{63} + S^{21} + \sum ({
m other\ harmonics}) + Irreg$$

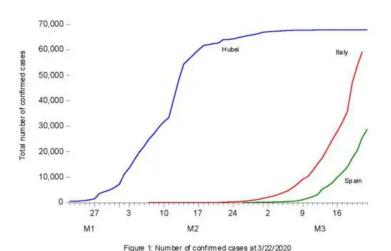
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30 Coronavirus trajectories



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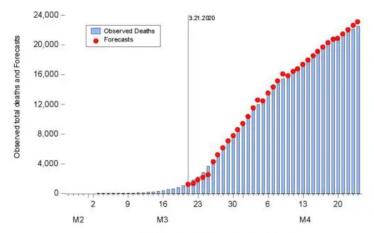


Figure 3: Observed Deaths and Forecasts in Spain

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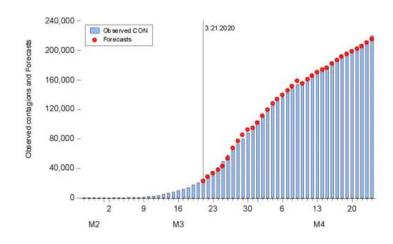


Figure 2: Observed contagions and Forecasts in Spain

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