

Milestone 7

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3/28/2020

Abstract

This is an extension of Jessica Trounstone’s “Segregation and Inequality in Public Goods” (2016). I was able to replicate the main results of Trounstone’s paper in R to suggest that racial segregation contributes to political polarization and decreased spending on public goods. Additionally, I extend the analysis by imputing missing data and rerunning Trounstone’s original model as a robustness check.

Introduction

This is my pdf document. Please refer to the Github repository of my final project for further information.¹. You can also access all of the original replication materials made available by Trounstone on Harvard Dataverse here. I make use of Trounstone (2016), Pencharz and Ball (2003), Xie (2020), Wickham (2019), and Xie (2015).

Summary of Trounstone (2016)

Trounstone’s *Segregation and Inequality in Public Goods* attempts to explain differences in public goods provision and political polarization through a racial lens by examining the relationships between polarization, goods provision, and segregation. Trounstone measures segregation with Theil’s H index, which measures the degree to which the diversity of a neighborhood differs from the diversity of the entire city. The main finding of the paper is that segregation, not simply diversity or political views, is an important determinant of both political polarization and spending on public goods. In general, segregation leads to the coincidence of racial and spatial political cleavages, which can make compromise on taxation and public spending difficult and tends to generally drive down the rate of spending on public goods. Because minorities are much more likely to live in racially segregated areas than whites, this suggests that public goods are also segregated across racial lines.

Literature Review

This is my literature review. Sources will be added when a more thorough job is done for milestone #8.

In the United States, residential segregation across racial lines remains a deeply entrenched problem in our society.

¹All sources, analysis, and further information are available on my Github repository for this project

Neighborhood racial isolation has been associated with racial intolerance and increased political competition. Those who live in homogenous neighborhoods are also more likely to believe in negative stereotypes about out-groups.

On the city level, diversity is associated with increased racial tension, lower levels of cooperation, intolerance, and lower spending on public goods.

A combination of homogenous neighborhoods within a diverse city leads to severe segregation and high degrees of racial tension.

Racial segregation has been associated with partisan political divides and a lack of cooperation across groups on city-wide policy.

Replication

Table 1 was able to be replicated exactly. The replication for Table 2 was close, but not exact, as were the replications for Table 3 and 4, which combined replicated Table 3 in the main paper. However, the implications of the main results are essentially the same for all of these regressions. The IV regression was able to be replicated exactly and the results presented in Tables 5 and 6 reconstruct the results from Table 5 in the original paper. So far, I have not been able to successfully recreate the marginal effects; however, I am working to figure out what the problem is. I believe that with a little bit more time I will be able to successfully recreate all of the results given my success in Table 1 and Tables 5 and 6.

As for the paper's Appendix Tables, I was able to successfully recreate Tables A1 and A2. I have not yet been able to successfully recreate Table A3 in Stata because I am having difficulty in figuring out why I have fewer observations and also how they calculated some of their statistics, as it seems that they filter for some conditions only for some rows, which I have yet to figure out how to do in R. I did not attempt to recreate Table A4 due to time constraints and also because it seemed extraneous.

Extension Ideas

There are a variety of ways that I could build upon this analysis. Because I have not narrowed it down to one yet, I will use this as an opportunity to propose a few ideas.

How well does this theory apply to more recent elections, and can we use new data to test it? For example, I could try to hunt down some more recent election data and use some sort of prediction function to see how well the model predicts the actual results. However, this could be quite difficult if the data is too messy or difficult to obtain (or even non-existent).

Another idea could be to redo the results but using the dissimilarity index, the most common measure of segregation, which Trounstine, perhaps controversially, chooses not to use in her analysis in favor of Theil's entropy score. It could be interesting to see whether her results hold up against this sort of robustness check.

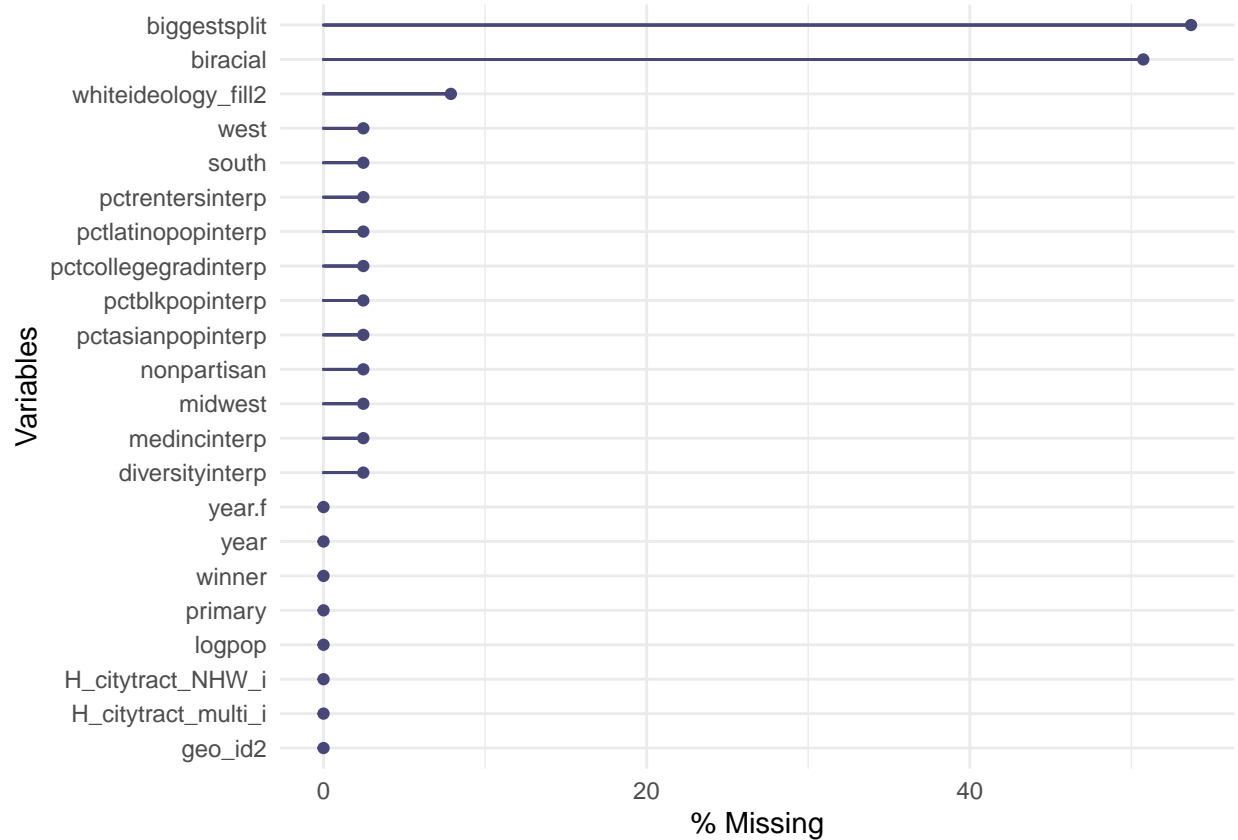
Another interesting thing to do would be to find a city that has recently become less segregated and see if political polarization has decreased and public goods provision has increased.

Finally, another approach to put to practice some more of the skills we've learned in this class would be to try and adopt a Bayesian framework to some of Trounstine's analyses.

Extension 1

One aspect of Trounstine's paper with room for improvement is that there is a large amount of missing data in her datasets upon which she bases her analyses. Using R's mice package (add citation), we can

perform multiple imputation. By performing imputation multiple times, this helps account for the uncertainty inherent in the individual imputations. Before performing the multiple imputations, we will first look at the missing data to see if there are any patterns.

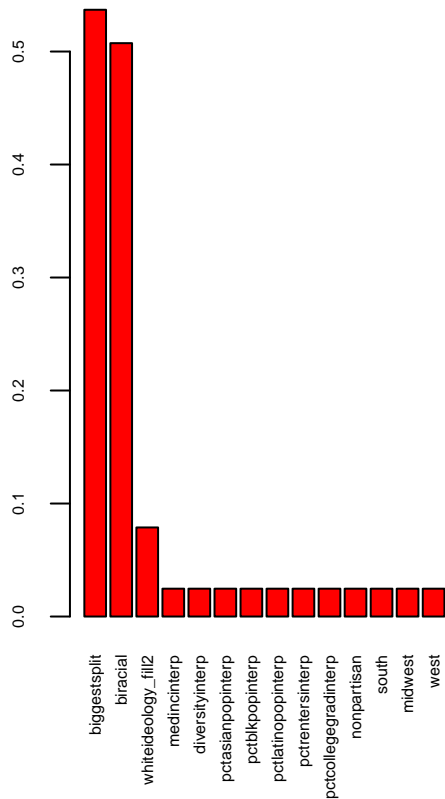


The above plot reveals that biggestsplit and biracial are the two variables with by far the largest percentage of missing values. Both of these are missing approximately 50% of their observations. Since biggestsplit is the dependent variable in our analysis with this dataset, this is an important fact.

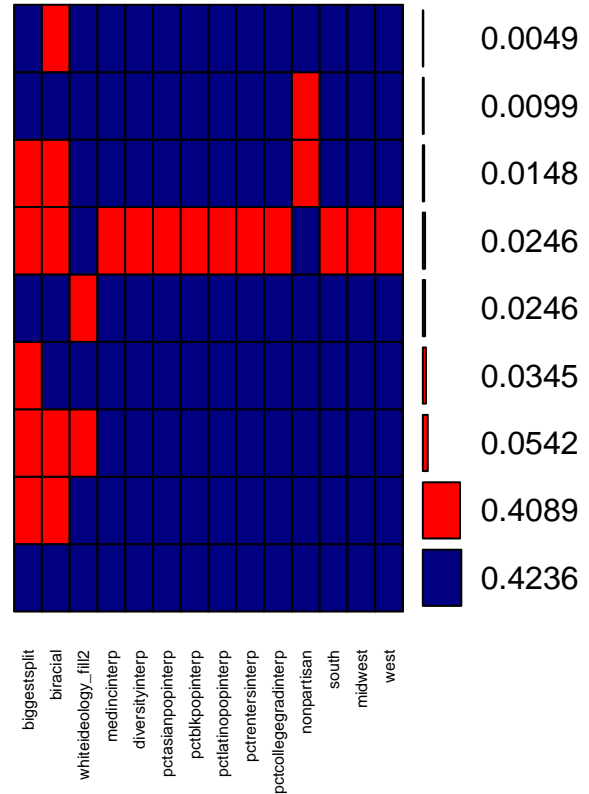
```
rp_aggr <- rp_impute %>%
  select("medincinterp",
         "biggestsplit", "diversityinterp",
         "pctasianpopinterp", "pctblkpopinterp", "pctlatinopopinterp",
         "pctrentersinterp", "pctcollegegradinterp", "biracial", "nonpartisan",
         "south", "midwest", "west", "whiteideology_fill2")

aggr_plot <- aggr(rp_aggr, bars = TRUE, col=c('navyblue','red'), numbers=TRUE, sortVars=TRUE, cex.axis=
```

Histogram of missing data



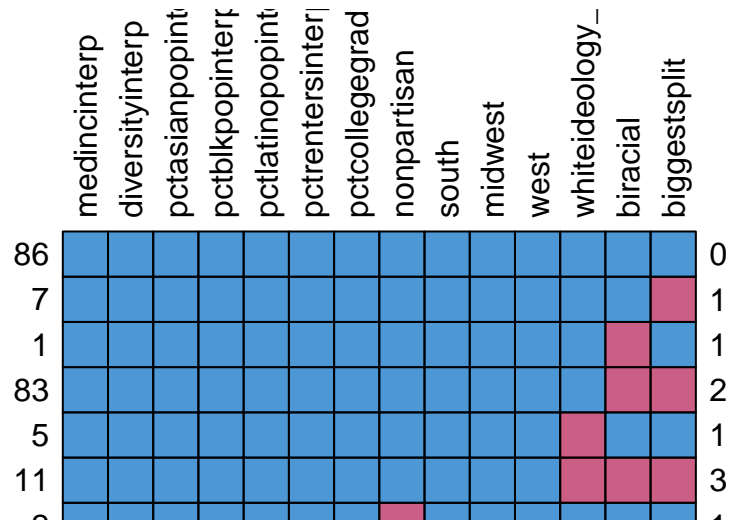
Pattern



Variables sorted by number of missings:

| Variable | Count |
|----------------------|------------|
| biggestsplit | 0.53694581 |
| biracial | 0.50738916 |
| whiteideology_fill2 | 0.07881773 |
| medincinterp | 0.02463054 |
| diversityinterp | 0.02463054 |
| pctasianpopinterp | 0.02463054 |
| pctblkpopinterp | 0.02463054 |
| pctlatinpopinterp | 0.02463054 |
| pctrentersinterp | 0.02463054 |
| pctcollegegradinterp | 0.02463054 |
| nonpartisan | 0.02463054 |
| south | 0.02463054 |
| midwest | 0.02463054 |
| west | 0.02463054 |

```
md.pattern(rp_aggr, rotate.names = TRUE)
```



| | medincinterp | diversityinterp | pctasianpopinterp | pctblkpopinterp | pctlatinpopinterp | pctrentersinterp | pctcollegegradinterp | nonpartisan | south |
|----|--------------|-----------------|-------------------|-----------------|-------------------|------------------|----------------------|-------------|-------|
| 86 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 7 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 83 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 5 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 11 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 2 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 1 |
| 3 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 1 |
| 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| | 5 | 5 | 5 | 5 | 5 | 5 | 5 | 5 | 5 |

| | midwest | west | whiteideology_fill2 | biracial | biggestsplit |
|----|---------|------|---------------------|----------|--------------|
| 86 | 1 | 1 | 1 | 1 | 1 |
| 7 | 1 | 1 | 1 | 1 | 0 |
| 1 | 1 | 1 | 1 | 0 | 1 |
| 83 | 1 | 1 | 1 | 0 | 0 |
| 5 | 1 | 1 | 0 | 1 | 1 |
| 11 | 1 | 1 | 0 | 0 | 0 |
| 2 | 1 | 1 | 1 | 1 | 1 |
| 3 | 1 | 1 | 1 | 0 | 0 |
| 5 | 0 | 0 | 1 | 0 | 0 |
| | 5 | 5 | 16 | 103 | 109 |

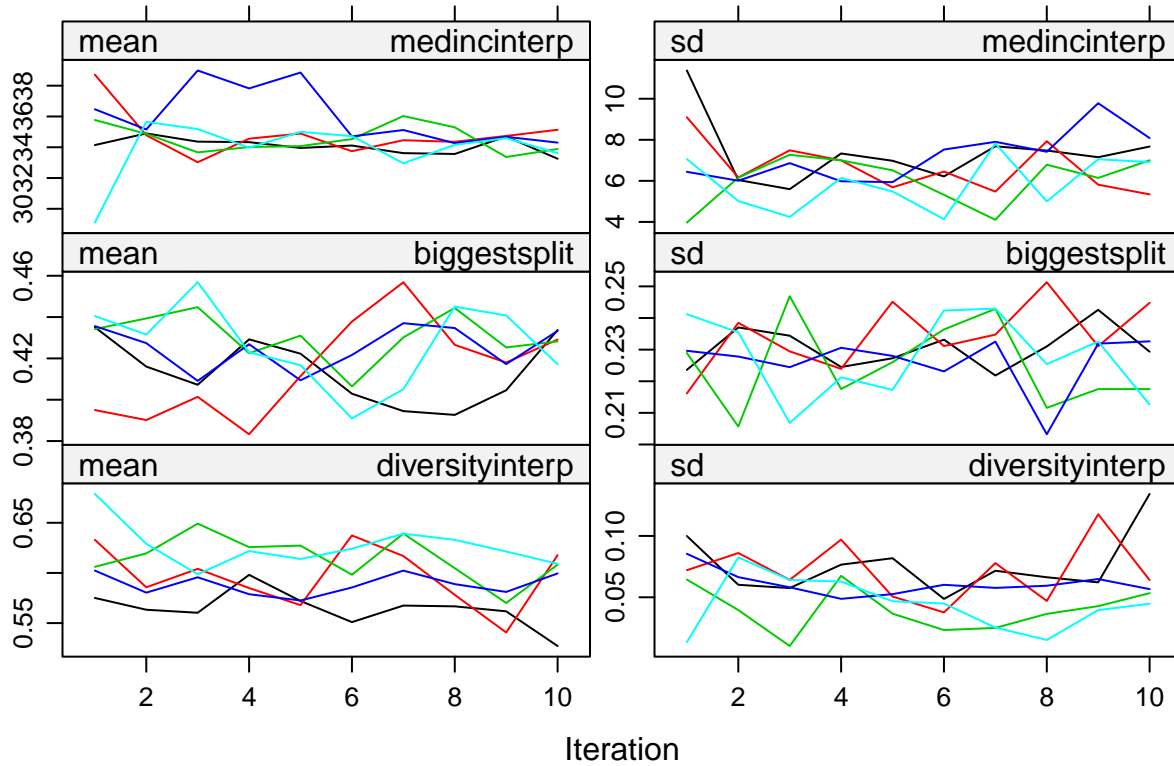
this plot tells us that 42% of the data are not missing anything

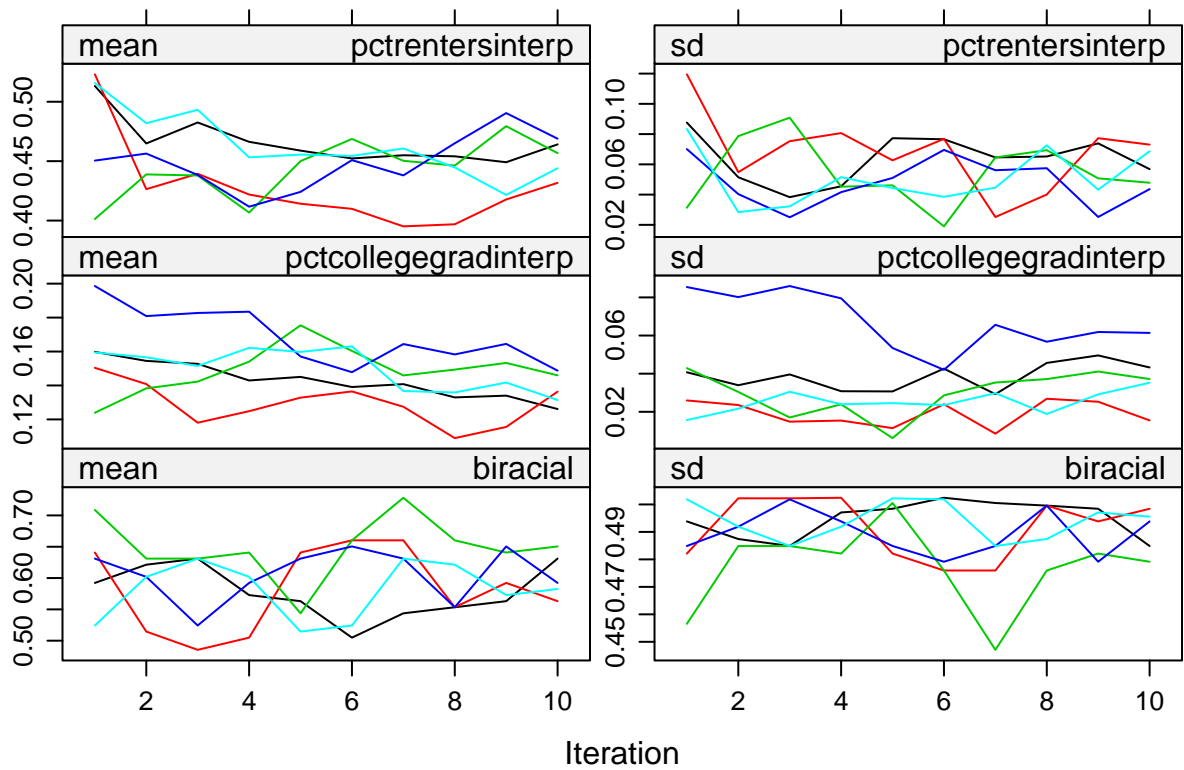
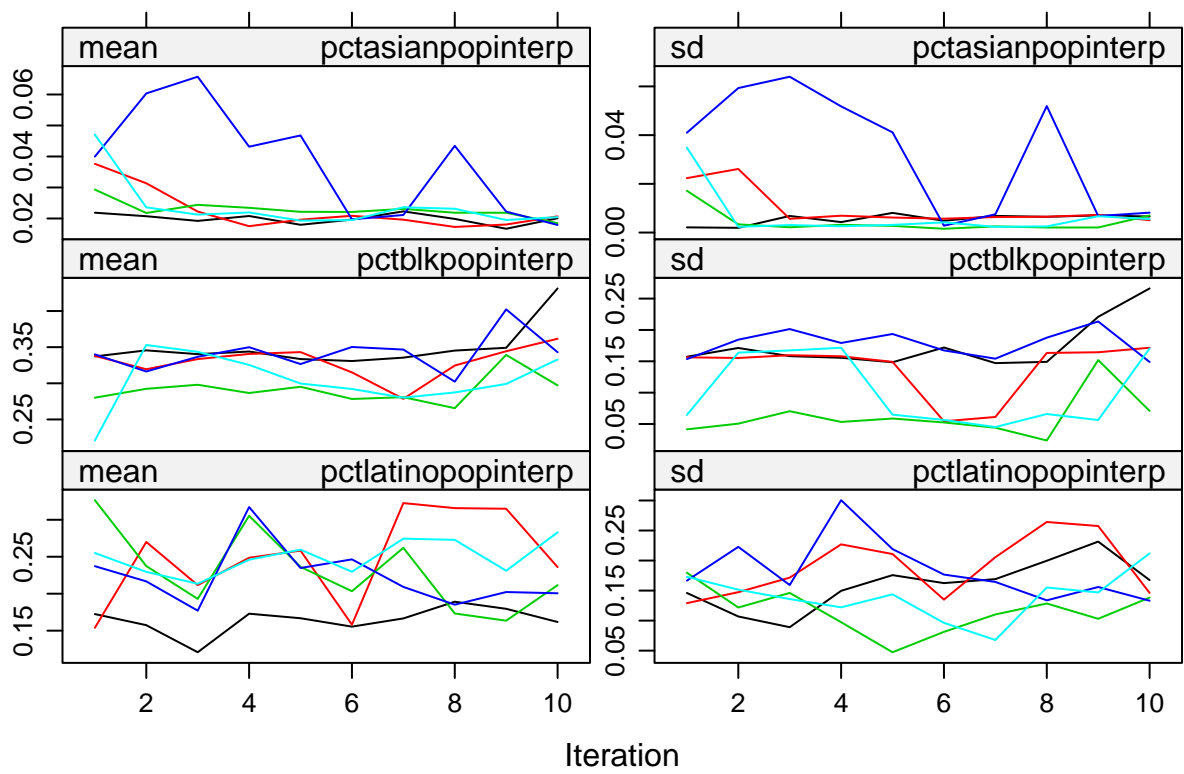
To better understand any potential patterns in missing data, we then plotted the pattern of missingness for only those variables missing values. From the plot on the right, we can see that approximately 42% of

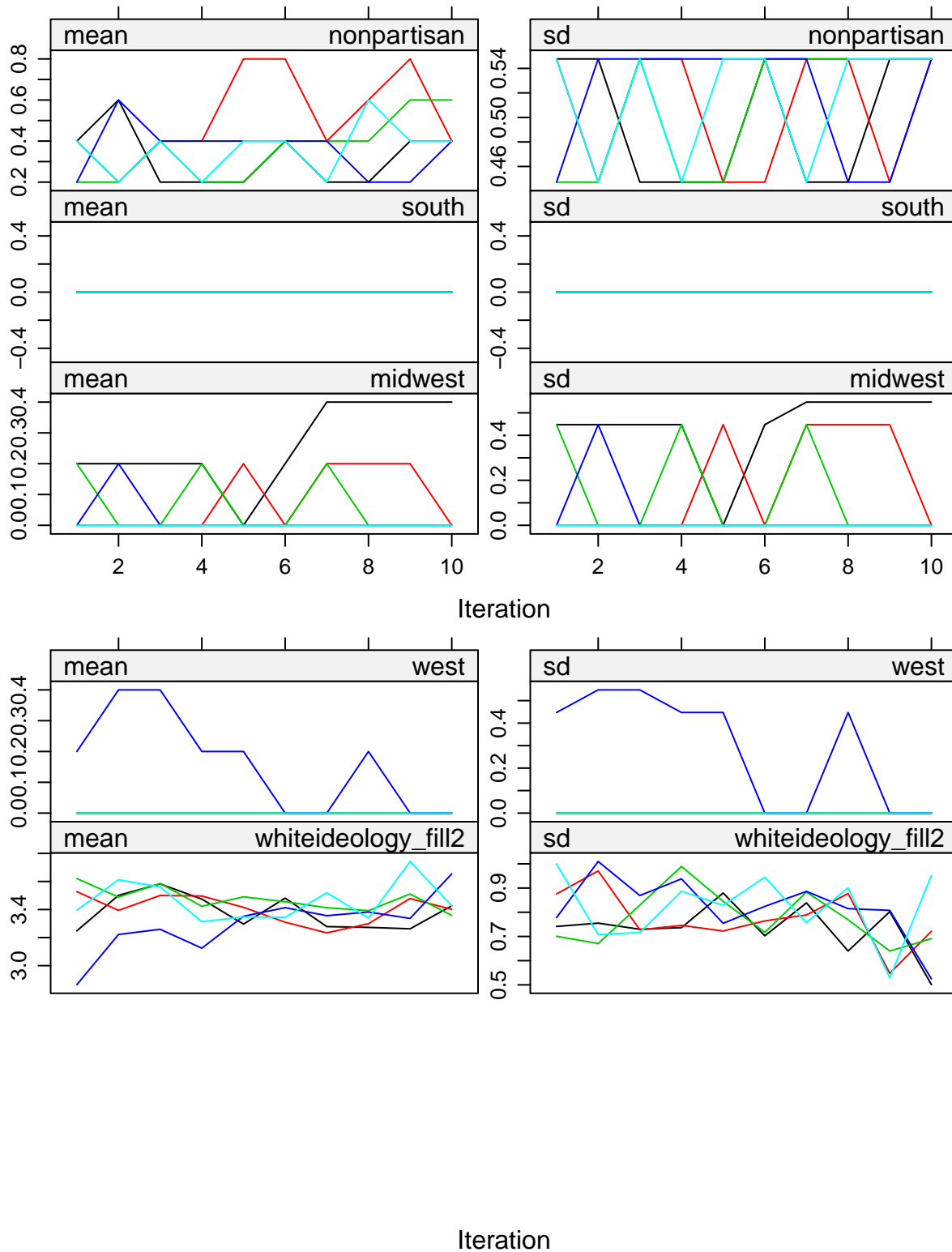
observations are complete. There seem to be a correspondence between missing a value for biggest split and missing biracial. There also seem to be about 2% of values for which most of the variables are missing. However, most observations are not missing more than 2-3 values.

Finding no clear patterns in the missing data, I next performed multiple imputations (with 10 iterations) on the dataset. A non-stochastic imputation method, Classification and Regression Trees (CART), was used instead of the default because of an error with matrix inversion caused by the data. Before examining the results of Trounstein's model using the imputed data, I first run some diagnostic tests of the imputation results to make sure that everything is running as expected.

First, I check the convergence of the algorithm used within mice() for each of the variables. For the most part, the fits intertwine and do not exhibit any trends at later iterations, as desired. There are some issues with the results for some of the variables with very few missing values (such as West and South), which makes sense as we would expect the mean to be less reliable due to the law of large numbers, and convergence is more difficult. Nevertheless, since there are so few of these values in the actual dataset, this is not a major concern.

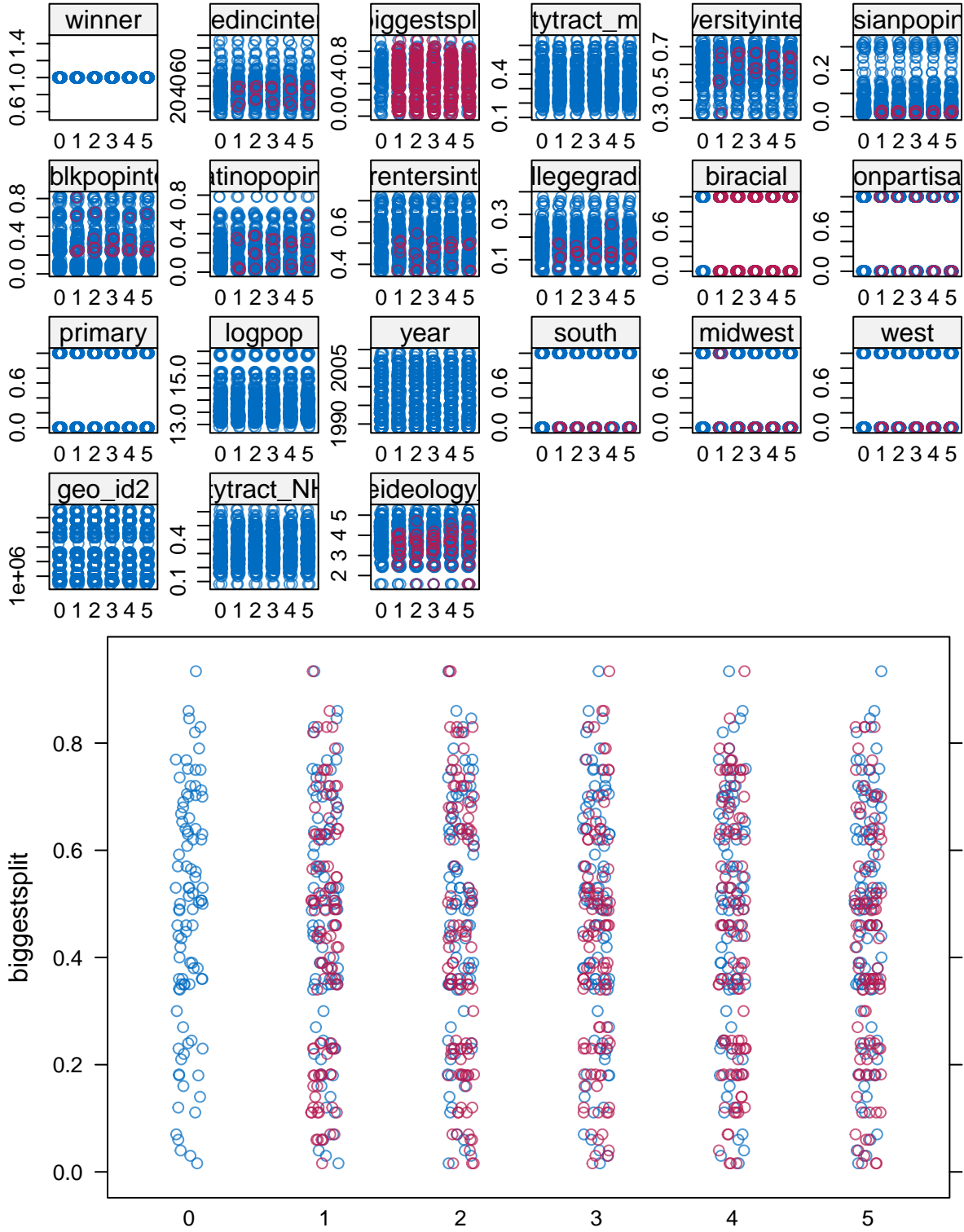




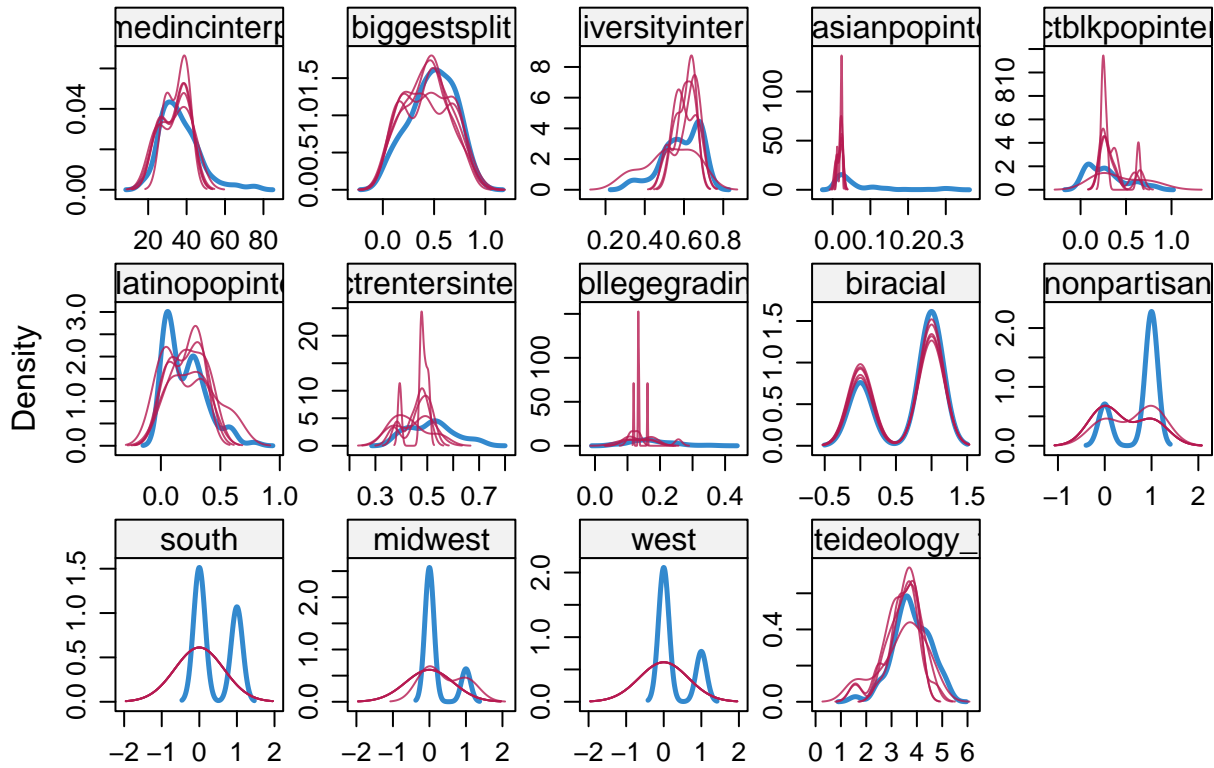


We can also check the imputed values against the original values using `stripplot()`. Each column in each subplot represents a separate iteration. The magenta points represent the imputed data. The values of the variable in questions are along the y axis. We expect the spread of the data to be similar if the imputations were done well. If the data were missing completely at random, then the imputed data should have the same distribution as the original data. In particular, we want to be sure that the imputations are within a plausible

range of the data. This is the case for all of our imputed variables, and there does not seem to be cause for alarm from these results.



Finally, we can also look at the density plots for each variable's actual data and for their imputed data from each of the iterations, which are represented in magenta. Overall, the density plots align quite well for the variables with the most missing data, biracial and biggestsplit, and relatively well for most of the other variables with missing data. Again, the fit is less good for variables with fewer missing data points.



While the imputations are not perfect, there do not seem to be any major problems so far. Thus, we can now proceed with our analysis. We fit each of our ten imputed datasets to Trounstone's 3 models using this data and then pool the results for each. The results are as follows:

Extension 1 Racial Polarization Table

Regression 1

| | term | estimate | std.error | p.value |
|----|----------------------|--------------|-------------|--------------|
| 1 | (Intercept) | 0.021289728 | 0.428314633 | 9.609281e-01 |
| 2 | H_citytract_multi_i | 0.424903753 | 0.386881542 | 2.974014e-01 |
| 3 | diversityinterp | 0.157450011 | 0.336414986 | 6.466862e-01 |
| 4 | pctasianpopinterp | -0.073034854 | 0.444854156 | 8.720015e-01 |
| 5 | pctblkpopinterp | 0.178639425 | 0.221495595 | 4.373528e-01 |
| 6 | pctlatinpopinterp | 0.164075368 | 0.214956001 | 4.610619e-01 |
| 7 | medincinterp | -0.001305195 | 0.003899519 | 7.399154e-01 |
| 8 | pctrentersinte | -0.505540820 | 0.245655070 | 4.293737e-02 |
| 9 | pctcollegegradinterp | 0.492826278 | 0.480209115 | 3.137296e-01 |
| 10 | biracial | 0.222204013 | 0.028445949 | 1.156828e-08 |
| 11 | nonpartisan | -0.084784453 | 0.063719326 | 2.135170e-01 |
| 12 | primary | -0.068569601 | 0.022538375 | 2.947573e-03 |
| 13 | logpop | 0.021215866 | 0.030728917 | 4.943477e-01 |

The results of the model with the new dataset are slightly different in comparison to the original results from Trounstone (2016). With the imputed data, we now have a total of 203 observations in our model as compared to the original model, which had only 91 observations. Interestingly, while the sign of the coefficient on the main variable of interest, the Theil's H segregation index is the same and the standard error has decreased slightly, the result has become statistically insignificant and the 95% confidence interval contains zero. Thus while these results still suggest that segregation may be associated with increased political polarization, they

confer a lesser degree of certainty than Trounstein's original analysis. The coefficients for pctblkpopinterp, pctlatinpopinterp, and medincinterp have also switched signs, and all coefficients except indicators for a biracial and primary election are also statistically insignificant, as was the case in the original analysis. In general, the standard error on the coefficients has decreased slightly.

In the second and third regressions, the results are similar in comparison to the unimputed data to the first. Standard errors have reduced slightly. The main coefficient on segregation index has decreased in magnitude, although in these specifications, it remains statistically significant. The other coefficients except for biracial and primary are insignificant, and the coefficients on percent Asian, Black, and Latino have all switched signs.

Regression 2

| | term | estimate | std.error | p.value |
|----|----------------------|--------------|-------------|--------------|
| 1 | (Intercept) | -0.028021685 | 0.429391912 | 9.487522e-01 |
| 2 | H_citytract_NHW_i | 0.503812393 | 0.274260488 | 8.465817e-02 |
| 3 | diversityinterp | 0.128086138 | 0.278403460 | 6.497486e-01 |
| 4 | pctasianpopinterp | 0.096327684 | 0.448519030 | 8.328082e-01 |
| 5 | pctblkpopinterp | 0.200753636 | 0.176609378 | 2.745499e-01 |
| 6 | pctlatinpopinterp | 0.208709231 | 0.217173836 | 3.581982e-01 |
| 7 | medincinterp | -0.002197686 | 0.003739137 | 5.598618e-01 |
| 8 | pctrentersinterp | -0.590022278 | 0.239549408 | 1.527298e-02 |
| 9 | pctcollegegradinterp | 0.646050477 | 0.466141649 | 1.746476e-01 |
| 10 | biracial | 0.220179445 | 0.028209600 | 1.265350e-08 |
| 11 | nonpartisan | -0.073206205 | 0.063241354 | 2.746007e-01 |
| 12 | primary | -0.067852253 | 0.022558381 | 3.347131e-03 |
| 13 | logpop | 0.024484858 | 0.030505455 | 4.274870e-01 |

Regression 3

| | term | estimate | std.error | p.value |
|----|----------------------|--------------|------------|--------------|
| 1 | (Intercept) | -0.057209577 | 0.42854984 | 8.954317e-01 |
| 2 | H_citytract_NHW_i | 0.476314380 | 0.27242266 | 9.894688e-02 |
| 3 | diversityinterp | 0.167218557 | 0.27603962 | 5.502128e-01 |
| 4 | pctasianpopinterp | 0.120181421 | 0.45108654 | 7.936525e-01 |
| 5 | pctblkpopinterp | 0.223218134 | 0.17199570 | 2.129575e-01 |
| 6 | pctlatinpopinterp | 0.189338868 | 0.21100140 | 3.881837e-01 |
| 7 | medincinterp | -0.001439268 | 0.00383561 | 7.097029e-01 |
| 8 | pctrentersinterp | -0.629882625 | 0.23710123 | 8.897578e-03 |
| 9 | pctcollegegradinterp | 0.486846914 | 0.48119419 | 3.187958e-01 |
| 10 | biracial | 0.215510583 | 0.02819784 | 1.520803e-08 |
| 11 | nonpartisan | -0.064834230 | 0.06181155 | 3.178115e-01 |
| 12 | primary | -0.067327075 | 0.02246490 | 3.475962e-03 |
| 13 | logpop | 0.033340422 | 0.03063427 | 2.828310e-01 |
| 14 | whiteideology_fill2 | -0.027578781 | 0.02151460 | 2.075705e-01 |
| 15 | year.f1990 | -0.009450522 | 0.13352336 | 9.452010e-01 |
| 16 | year.f1991 | -0.004676689 | 0.05683288 | 9.347988e-01 |
| 17 | year.f1992 | -0.080365497 | 0.11793485 | 5.040218e-01 |
| 18 | year.f1993 | 0.001730653 | 0.05964367 | 9.768955e-01 |
| 19 | year.f1994 | -0.051833262 | 0.10874620 | 6.409745e-01 |
| 20 | year.f1995 | -0.031970530 | 0.07422024 | 6.725307e-01 |
| 21 | year.f1996 | 0.005376359 | 0.13244501 | 9.679063e-01 |
| 22 | year.f1997 | -0.035283231 | 0.07631290 | 6.472195e-01 |
| 23 | year.f1998 | -0.166828180 | 0.15569417 | 3.012218e-01 |

```

24      year.f1999 -0.049963327 0.08868351 5.830423e-01
25      year.f2000  0.163501084 0.14243864 2.633043e-01
26      year.f2001 -0.053189517 0.08390707 5.339416e-01
27      year.f2002 -0.150864838 0.11003594 1.800189e-01
28      year.f2003 -0.130987638 0.08515400 1.358167e-01
29      year.f2004 -0.103246293 0.11041594 3.676207e-01
30      year.f2005 -0.153377918 0.09521167 1.225700e-01
31      year.f2006 -0.198930433 0.17307381 2.697358e-01
32      year.f2007 -0.161639818 0.09795083 1.166759e-01
33      year.f2008 -0.159312984 0.27620087 5.801481e-01
34      year.f2009 -0.182999459 0.16244234 2.935900e-01
35          south  0.134289154 0.05100857 1.199686e-02
36      midwest  0.125276354 0.05718915 3.438830e-02
37          west  0.028609786 0.05452091 6.020324e-01

```

Appendix

Table 1

Table 2

Table 3

Table 3

Main Analysis 4

Main Analysis 5

Table 5

Appendix

TABLE A2 Cities Included in Racial Polarization Data

| City Name | Replication | | | | |
|---------------|-----------------------------|-----------|--|--------------|--------------|
| | Segregation: Mean H Index | | Largest Racial Divide, Number of Elections | | |
| | Multigroup | Two-Group | Black/White | Latino/White | Black/Latino |
| Austin, TX | 0.204 | 0.208 | 1 | 0 | 0 |
| Baltimore, MD | 0.510 | 0.516 | 3 | 1 | 0 |
| Charlotte, NC | 0.269 | 0.287 | 2 | 0 | 0 |
| Chicago, IL | 0.572 | 0.460 | 7 | 0 | 1 |
| Cleveland, OH | 0.558 | 0.531 | 2 | 0 | 0 |
| Columbus, OH | 0.316 | 0.284 | 3 | 0 | 1 |
| Dallas, TX | 0.359 | 0.339 | 4 | 0 | 1 |

| | | | | | |
|-------------------|-------|-------|---|---|---|
| Denver, CO | 0.289 | 0.254 | 1 | 2 | 0 |
| Detroit, MI | 0.398 | 0.255 | 1 | 0 | 1 |
| Houston, TX | 0.339 | 0.308 | 7 | 0 | 2 |
| Indianapolis, IN | 0.292 | 0.293 | 0 | 0 | 1 |
| Jacksonville, FL | 0.233 | 0.222 | 2 | 0 | 0 |
| Los Angeles, CA | 0.351 | 0.366 | 3 | 0 | 5 |
| Memphis, TN | 0.470 | 0.474 | 2 | 0 | 0 |
| Milwaukee, WI | 0.423 | 0.360 | 3 | 0 | 0 |
| New York, NY | 0.468 | 0.474 | 5 | 3 | 1 |
| Oklahoma, OK | 0.231 | 0.165 | 1 | 0 | 0 |
| Philadelphia, PA | 0.492 | 0.487 | 5 | 0 | 0 |
| Phoenix, AZ | 0.255 | 0.270 | 0 | 1 | 0 |
| San Antonio, TX | 0.237 | 0.225 | 0 | 4 | 0 |
| San Diego, CA | 0.255 | 0.266 | 3 | 0 | 1 |
| San Francisco, CA | 0.223 | 0.161 | 3 | 0 | 1 |
| San Jose, CA | 0.186 | 0.198 | 0 | 2 | 1 |
| Tucson, AZ | 0.185 | 0.192 | 1 | 0 | 0 |
| Washington, DC | 0.464 | 0.491 | 3 | 0 | 0 |

TABLE A3 Summary Statistics: Census of Government Finance and Population

| Variable | Obs | Mean | SD | Min | Max |
|---|-------|------------|-------------|------------|---------------|
| Direct General Expenditure per Capita | 13742 | 1.186 | 1.220 | 0.019 | 70.457 |
| Highways per Capita | 13603 | 0.081 | 0.053 | 0.000 | 1.106 |
| Parks per Capita | 12905 | 0.061 | 0.061 | 0.000 | 1.111 |
| Police per Capita | 13626 | 0.181 | 0.094 | 0.000 | 1.546 |
| Sewers per Capita | 11223 | 0.092 | 0.077 | 0.000 | 1.591 |
| Welfare, Health, and Housing per Capita | 10871 | 0.057 | 0.131 | 0.000 | 4.984 |
| Own Source Revenue per Capita | 13741 | 0.942 | 1.118 | 0.021 | 76.123 |
| Two-Group H Index | 13742 | 0.076 | 0.099 | 0.000 | 0.767 |
| Diversity | 13742 | 0.309 | 0.188 | 0.007 | 0.772 |
| % Black | 13742 | 0.097 | 0.151 | 0.000 | 0.980 |
| % Asian | 13742 | 0.032 | 0.054 | 0.000 | 0.674 |
| % Latino | 13742 | 0.104 | 0.161 | 0.000 | 0.987 |
| 5Y Change, % Black | 11194 | 0.007 | 0.019 | −0.101 | 0.229 |
| 5Y Change, % Latino | 11194 | 0.016 | 0.020 | −0.171 | 0.207 |
| 5Y Change, % Asian | 11194 | 0.005 | 0.011 | −0.056 | 0.128 |
| Median Income | 13742 | 54,520.132 | 22,081.359 | 15,642.802 | 240,938.047 |
| % Local Gov. Employees | 13742 | 3.359 | 0.951 | 0.677 | 8.365 |
| % Renters | 13742 | 0.360 | 0.140 | 0.014 | 0.871 |
| % Over 65 | 13742 | 0.125 | 0.050 | 0.012 | 0.771 |
| % College Degree | 13742 | 0.160 | 0.099 | 0.003 | 0.587 |
| Population (logged) | 13742 | 10.132 | 1.016 | 6.071 | 15.921 |
| City Ideology | 2130 | 4.023 | 0.780 | 1.000 | 7.000 |
| Population | 13742 | 53,723.022 | 208,143.791 | 433.000 | 8,214,426.000 |

Bibliography

Pencharz, Paul B., and Ronald O. Ball. 2003. “Different Approaches to Define Individual Amino Acid Requirements.” *Annual Review of Nutrition* 23. Annual Reviews: 101–16.

Table 1: **Racial Polarization in Segregated Cities**

| | <i>Dependent variable:</i> | | |
|--------------------------|----------------------------|-------------------|------------------|
| | biggestsplit | | |
| | (1) | (2) | (3) |
| Multigroup H Index | 0.932** (0.394) | | |
| White/Nonwhite H Index | | 0.756** (0.297) | 0.835*** (0.296) |
| Diversity | 0.385 (0.362) | 0.518 (0.323) | 0.584* (0.323) |
| Percent Asian | −0.115 (0.527) | 0.120 (0.558) | −0.004 (0.522) |
| Percent Black | −0.432 (0.269) | −0.237 (0.216) | −0.133 (0.212) |
| Percent Latino | −0.191 (0.257) | −0.059 (0.254) | 0.095 (0.278) |
| Median HH Income (1000s) | −0.004 (0.007) | −0.007 (0.007) | −0.002 (0.006) |
| Percent Renters | −0.580 (0.422) | −0.806* (0.431) | −0.419 (0.454) |
| Percent College Degree | 0.328 (0.711) | 0.723 (0.729) | 0.123 (0.869) |
| Biracial Contest | 0.210*** (0.037) | 0.208*** (0.037) | 0.192*** (0.036) |
| Nonpartisan Election | −0.090 (0.066) | −0.089 (0.066) | −0.034 (0.065) |
| Primary Election | −0.092*** (0.032) | −0.090*** (0.032) | −0.071** (0.030) |
| Population (logged) | 0.035 (0.055) | 0.048 (0.055) | −0.011 (0.061) |
| White Ideology | | | −0.051 (0.032) |
| Constant | −0.242 (0.569) | −0.393 (0.563) | 0.236 (0.605) |
| Wald Chi Squared | a | b | c |
| Observations | 91 | 91 | 86 |
| Akaike Inf. Crit. | −55.548 | −56.381 | −64.922 |
| Bayesian Inf. Crit. | 32.332 | 31.499 | 20.981 |

Note:

*p<0.1; **p<0.05; ***p<0.01

Table 2: **Effect of Segregation on Overall per Capita City Expenditures**

| | <i>Dependent variable:</i> | | |
|-------------------------|----------------------------|----------------------|----------------------|
| | dgepercap_cpi | | |
| | (1) | (2) | (3) |
| H_citytract_NHW_i | −1.153*** (0.221) | −1.011*** (0.254) | −1.733*** (0.437) |
| diversityinterp | 0.106 (0.134) | | −0.063 (0.246) |
| pctblkpopinterp | 0.681*** (0.167) | 0.741*** (0.161) | 0.164 (0.523) |
| pctasianpopinterp | −0.385 (0.302) | −0.852** (0.348) | 0.197 (0.706) |
| pctlatinpopinterp | 1.543*** (0.186) | 1.577*** (0.205) | 1.622*** (0.390) |
| chn5pctblk | | −1.778*** (0.644) | |
| chn5pctlatino | | −2.055** (0.823) | |
| chn5pctasian | | −0.800 (1.093) | |
| medinc_cpi | 0.002* (0.001) | 0.001 (0.002) | 0.004 (0.003) |
| pctlcorgovworker_100 | 0.014 (0.016) | 0.006 (0.018) | −0.030 (0.046) |
| pctrentersinterp | 0.527 (0.333) | 0.547 (0.385) | 0.336 (0.656) |
| pctover65 | 0.093 (0.643) | 0.487 (0.451) | −0.865 (0.816) |
| pctcollegegradinterp | 5.395*** (0.403) | 6.260*** (0.419) | 6.527*** (1.029) |
| logpop | −0.243*** (0.044) | −0.290*** (0.068) | −0.447*** (0.088) |
| ideology_fill | | | −0.012 (0.034) |
| Observations | 13,742 | 11,194 | 2,130 |
| R ² | 0.863 | 0.897 | 0.882 |
| Adjusted R ² | 0.830 | 0.865 | 0.855 |
| Residual Std. Error | 0.503 (df = 11094) | 0.465 (df = 8544) | 0.405 (df = 1741) |

Note:

*p<0.1, **p<0.05; ***p<0.01

Table 3: Effect of Segregation on Public Goods A

| | <i>Dependent variable:</i> | | |
|-------------------------|----------------------------|----------------------|----------------------|
| | highwayspercapNC_cpi | policepercapNC_cpi | parkspcapNC_cpi |
| | (1) | (2) | (3) |
| H_citytract_NHW_i | −0.039** (0.016) | −0.215*** (0.023) | −0.046*** (0.018) |
| diversityinterp | 0.005 (0.010) | 0.059*** (0.013) | 0.001 (0.013) |
| pctblkpopinterp | 0.052*** (0.014) | 0.142*** (0.018) | 0.031* (0.018) |
| pctasianpopinterp | −0.036 (0.026) | −0.055 (0.035) | −0.067*** (0.023) |
| pctlatinpopinterp | 0.025* (0.014) | 0.335*** (0.019) | 0.049*** (0.014) |
| medinc_cpi | 0.0003** (0.0001) | 0.00004 (0.0001) | −0.00002 (0.0001) |
| pctlocalgovworker_100 | −0.0003 (0.001) | −0.001 (0.002) | 0.001 (0.001) |
| pctrentersinterp | 0.011 (0.023) | 0.075*** (0.028) | 0.018 (0.021) |
| pctover65 | 0.140*** (0.032) | 0.147*** (0.045) | 0.127*** (0.040) |
| pctcollegegradinterp | 0.218*** (0.026) | 0.793*** (0.038) | 0.444*** (0.038) |
| logpop | −0.015*** (0.004) | −0.054*** (0.004) | −0.005* (0.003) |
| Observations | 13,603 | 13,626 | 12,905 |
| R ² | 0.571 | 0.837 | 0.750 |
| Adjusted R ² | 0.467 | 0.798 | 0.688 |
| Residual Std. Error | 0.039 (df = 10958) | 0.042 (df = 10991) | 0.034 (df = 10321) |

Note:

*p<0.1; **p<0.05; ***p<0.01

Table 4: **Effect of Segregation on Public Goods B**

| | <i>Dependent variable:</i> | | |
|-------------------------|----------------------------|----------------------|----------------------|
| | sewerspercapNC_cpi | welfhoushealthNC_cpi | genrevownpercap_cpi |
| | (1) | (2) | (3) |
| H_citytract_NHW_i | −0.148*** (0.022) | −0.138*** (0.049) | −0.768*** (0.155) |
| diversityinterp | 0.039*** (0.015) | −0.033 (0.025) | 0.091 (0.085) |
| pctblkpopinterp | 0.012 (0.017) | 0.016 (0.056) | 0.272** (0.120) |
| pctasianpopinterp | −0.124*** (0.044) | 0.130 (0.090) | −0.147 (0.233) |
| pctlatinpopinterp | 0.091*** (0.019) | 0.140*** (0.028) | 1.202*** (0.120) |
| medinc_cpi | 0.001*** (0.0002) | −0.0003 (0.0003) | 0.004*** (0.001) |
| pctlocalgovworker_100 | −0.004* (0.002) | −0.007** (0.003) | 0.002 (0.013) |
| pctrentersinterp | 0.174*** (0.034) | 0.079* (0.046) | 0.569** (0.263) |
| pctover65 | 0.104* (0.053) | −0.058 (0.070) | 0.443 (0.471) |
| pctcollegegradinterp | 0.286*** (0.043) | 0.421*** (0.080) | 4.331*** (0.349) |
| logpop | −0.023*** (0.003) | −0.012* (0.007) | −0.126*** (0.032) |
| Observations | 11,223 | 10,871 | 13,741 |
| R ² | 0.675 | 0.828 | 0.886 |
| Adjusted R ² | 0.586 | 0.777 | 0.859 |
| Residual Std. Error | 0.049 (df = 8805) | 0.062 (df = 8380) | 0.420 (df = 11093) |

Note:

*p<0.1; **p<0.05; ***p<0.01

Table 5: **Effect of Segregation on Public Goods**

| | <i>Dependent variable:</i> | | | | |
|-------------------------|----------------------------|----------------------|----------------------|----------------------|----------------------|
| | highwayspercapNC_cpi | policepercapNC_cpi | parkspercapNC_cpi | sewerspercapNC_cpi | welthoushealthNC_cpi |
| | (1) | (2) | (3) | (4) | (5) |
| H_citytract_NHW_i | -0.039** (0.016) | -0.215*** (0.023) | -0.046*** (0.018) | -0.148*** (0.022) | -0.138*** (0.049) |
| diversityinterp | 0.005 (0.010) | 0.059*** (0.013) | 0.001 (0.013) | 0.039*** (0.015) | -0.033 (0.025) |
| pctblkpopinterp | 0.052*** (0.014) | 0.142*** (0.018) | 0.031* (0.018) | 0.012 (0.017) | 0.016 (0.056) |
| pctasianpopinterp | -0.036 (0.026) | -0.055 (0.035) | -0.067*** (0.023) | -0.124*** (0.044) | 0.130 (0.090) |
| pctlatinpopinterp | 0.025* (0.014) | 0.335*** (0.019) | 0.049*** (0.014) | 0.091*** (0.019) | 0.140*** (0.028) |
| medinc_cpi | 0.0003** (0.0001) | 0.00004 (0.0001) | -0.00002 (0.0001) | 0.001*** (0.0002) | -0.0003 (0.0003) |
| pctlocalgovworker_100 | -0.0003 (0.001) | -0.001 (0.002) | 0.001 (0.001) | -0.004* (0.002) | -0.007** (0.003) |
| pctrentersinterp | 0.011 (0.023) | 0.075*** (0.028) | 0.018 (0.021) | 0.174*** (0.034) | 0.079* (0.046) |
| pctover65 | 0.140*** (0.032) | 0.147*** (0.045) | 0.127*** (0.040) | 0.104* (0.053) | -0.058 (0.070) |
| pctcollegegradinterp | 0.218*** (0.026) | 0.793*** (0.038) | 0.444*** (0.038) | 0.286*** (0.043) | 0.421*** (0.080) |
| logpop | -0.015*** (0.004) | -0.054*** (0.004) | -0.005* (0.003) | -0.023*** (0.003) | -0.012* (0.007) |
| Observations | 13,603 | 13,626 | 12,905 | 11,223 | 13,871 |
| R ² | 0.571 | 0.837 | 0.750 | 0.675 | 0.828 |
| Adjusted R ² | 0.467 | 0.798 | 0.688 | 0.586 | 0.777 |
| Residual Std. Error | 0.039 (df = 10958) | 0.042 (df = 10991) | 0.034 (df = 10321) | 0.049 (df = 8805) | 0.062 (df = 8380) |

Note:

*p<0.1; **p<0.05;

Table 6: Effect of Segregation on City Expenditures, IV Approach A

| | <i>Dependent variable:</i> | | | |
|-------------------------|----------------------------|----------------------|----------------------|-----------------------|
| | dgepercap_cpi | highwayspercapNC_cpi | policepercapNC_cpi | parkspcapNC_cpi |
| | (1) | (2) | (3) | (4) |
| H_citytract_NHW_i | −2.676*** (0.935) | −0.363*** (0.056) | −0.350*** (0.109) | −0.034* (0.019) |
| dgepercap_cpilag | 1.472*** (0.007) | | | |
| highwayspercapNC_cpilag | | 0.477*** (0.003) | | |
| policepercapNC_cpilag | | | 0.955*** (0.004) | |
| parkspcapNC_cpilag | | | | 0.869*** (0.006) |
| diversityinterp | 0.264 (0.355) | −0.032 (0.022) | −0.020 (0.042) | 0.004 (0.007) |
| pctblkpopinterp | 0.376 (0.325) | 0.085*** (0.020) | 0.096** (0.038) | 0.003 (0.007) |
| pctasianpopinterp | 0.143 (0.940) | −0.111** (0.056) | −0.064 (0.110) | −0.022 (0.019) |
| pctlatinpopinterp | 0.087 (0.284) | 0.088*** (0.017) | 0.088*** (0.033) | 0.007 (0.006) |
| medincinterp | −0.004 (0.003) | 0.001*** (0.0002) | 0.001*** (0.0004) | 0.0002*** (0.0001) |
| pctlocalgovworker_100 | −0.104*** (0.032) | 0.021*** (0.002) | 0.026*** (0.004) | 0.003*** (0.001) |
| pctrentersinterp | −0.553 (0.350) | 0.165*** (0.021) | 0.187*** (0.041) | 0.035*** (0.007) |
| pctover65 | 0.301 (0.730) | 0.267*** (0.044) | 0.238*** (0.086) | 0.072*** (0.015) |
| pctcollegegradinterp | 0.248 (0.567) | −0.044 (0.034) | −0.101 (0.067) | 0.028** (0.012) |
| Constant | 0.328 (0.278) | −0.125*** (0.017) | −0.183*** (0.033) | −0.029*** (0.006) |
| Observations | 21,145 | 20,704 | 20,627 | 19,056 |
| R ² | 0.685 | 0.615 | 0.789 | 0.540 |
| Adjusted R ² | 0.685 | 0.615 | 0.789 | 0.539 |
| Residual Std. Error | 4.877 (df = 21125) | 0.290 (df = 20684) | 0.566 (df = 20607) | 0.093 (df = 19036) |

Note:

*p<0.1; **p<0.05; ***p<0.0

Table 7: **Effect of Segregation on City Expenditures, IV Approach B**

| | <i>Dependent variable:</i> | | |
|-------------------------|----------------------------|---------------------|----------------------|
| | sewerspercapNC_cpi | genrevownpercap_cpi | welfhoushealthNC_cpi |
| | (1) | (2) | (3) |
| H_citytract_NHW_i | −0.363*** (0.060) | −1.873** (0.789) | −0.115** (0.054) |
| sewerspercapNC_cpilag | 0.064*** (0.008) | | |
| genrevownpercap_cpilag | | 1.235*** (0.006) | |
| welfhoushealthNC_cpilag | | | 0.893*** (0.005) |
| diversityinterp | 0.080*** (0.024) | 0.047 (0.300) | −0.047** (0.022) |
| pctblkpopinterp | 0.058** (0.025) | 0.360 (0.274) | 0.076*** (0.023) |
| pctasianpopinterp | −0.223*** (0.068) | −0.029 (0.793) | 0.009 (0.053) |
| pctlatinpopinterp | −0.050*** (0.019) | 0.206 (0.240) | 0.078*** (0.017) |
| medincinterp | 0.0002 (0.0003) | 0.0003 (0.003) | 0.001** (0.0002) |
| pctlocalgovworker_100 | 0.001 (0.002) | −0.003 (0.027) | 0.016*** (0.002) |
| pctrentersinterp | 0.073*** (0.024) | 0.263 (0.295) | 0.098*** (0.023) |
| pctover65 | 0.287*** (0.051) | 0.782 (0.616) | 0.127** (0.050) |
| pctcollegegradinterp | 0.029 (0.040) | −0.035 (0.478) | −0.038 (0.038) |
| Constant | 0.004 (0.019) | −0.174 (0.234) | −0.093*** (0.018) |
| Observations | 16,616 | 21,148 | 14,711 |
| R ² | 0.006 | 0.681 | 0.699 |
| Adjusted R ² | 0.005 | 0.681 | 0.698 |
| Residual Std. Error | 0.284 (df = 16596) | 4.115 (df = 21128) | 0.252 (df = 14691) |

Note:

*p<0.1; **p<0.05; ***p<0.01

Table 8: **TABLE A1 Summary Statistics: Racial Polarization Data**

| Statistic | N | Mean | St. Dev. | Min | Max |
|-----------------------|----|--------|----------|--------|--------|
| Largest Racial Divide | 91 | 0.481 | 0.213 | 0.016 | 0.934 |
| H Index: Multigroup | 91 | 0.376 | 0.119 | 0.183 | 0.635 |
| H Index: Two-Group | 91 | 0.353 | 0.114 | 0.156 | 0.614 |
| Diversity | 91 | 0.623 | 0.088 | 0.323 | 0.736 |
| % Asian | 91 | 0.067 | 0.074 | 0.008 | 0.318 |
| % Black | 91 | 0.275 | 0.181 | 0.030 | 0.815 |
| % Latino | 91 | 0.229 | 0.155 | 0.009 | 0.605 |
| Median HH Income | 91 | 36.725 | 10.114 | 17.267 | 75.982 |
| % Renters | 91 | 0.535 | 0.092 | 0.368 | 0.718 |
| % College Degree | 91 | 0.167 | 0.056 | 0.049 | 0.359 |
| Biracial Contest | 91 | 0.725 | 0.449 | 0 | 1 |
| Nonpartisan Election | 91 | 0.714 | 0.454 | 0 | 1 |
| Primary Election | 91 | 0.352 | 0.480 | 0 | 1 |
| Population (logged) | 91 | 14.166 | 0.826 | 13.065 | 15.921 |
| White Ideology | 86 | 3.835 | 0.648 | 2.667 | 5.250 |

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