Statistics 705 Notes 4

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Random number generation and univariate distributions

Anyone who considers arithmetical methods of producing random digits is, of course, in a state of sin. — John von Neumann (1951)

Anyone who has not seen the above quotation in at least 100 places is probably not very old. — D. V. Pryor (1993)

Random number generators should not be chosen at random. — Donald Knuth (1986)

Background reading

Becker, R. A., Chambers, J. M. and Wilks, A. R. (1988) The New S Language. Wadsworth & Brooks/Cole. –reference for ${\bf runif}$ () function in R

There are a plethora of references listed if you type help(".Random.seed") into R

True randomness?

In R, the Random Number Generators are only pseudorandom—the numbers aren't random, but should be indistinguishable from truly random numbers

True sources of randomness include

- Atmospheric noise (random.org)
- Thermal noise
- In general, measuring a physical process that we think to be random (and then adjusting for the error of measurement)

What does R do?

Pseudorandom number generation

- Make a deterministic sequence that should be indistinguishable from truly random numbers if we don't know the starting point of the sequence
- How?

Example using runif()

```
runif(5)
## [1] 0.8754247 0.1411678 0.1758072 0.9235542 0.1449683
set.seed(1)
runif(5)
## [1] 0.2655087 0.3721239 0.5728534 0.9082078 0.2016819
set.seed(1)
runif(5)
## [1] 0.2655087 0.3721239 0.5728534 0.9082078 0.2016819
```

LCGs

The simplest and most common pseudorandom number generators: Linear-Congruential Generators (LCG's):

$$x_{n+1} = a \cdot x_n + b \mod m$$

- a is called the multiplier
- b the addend
- m the modulus, usually close to a word-size, e.g. 2^{32} or $2^{31}-1$

With carefully chosen a, the period (length of the cycle of the LCG) will be

- m if m is a power of 2
- m-1 if m is prime

```
LCG <- function (a, b, m, x0, n) {
    x <- numeric(n)
    # x %% y gives x (mod y)
    for(i in 1:n){
    x[i] <- (a*x0 + b) %% m
    x0 <- x[i]}
    return(x)
}</pre>
```

Divide by m if we want Unif[0,1]

LCGs

A good triplet according to S. K. Park and K. W. Miller, "Random Number Generators: Good Ones are Hard to Find," Transactions of the ACM, Nov. 1988 is

$$m = 2^{31} - 1$$
; $a = 7^5$; $b = 0$

```
# generate 10 from this LCG
LCG(7^5,0,(2^(31)-1),x0=1,10)
```

```
## [1] 16807 282475249 1622650073 984943658 1144108930
## [7] 101027544 1457850878 1458777923 2007237709
```

There are famously bad choices

Ex:

$$a = 65539$$
; $m = 2^{31}$; $b = 0$; x_0 odd

Knuth called this "truly horrible!"

"its very name RANDU is enough to bring dismay into the eyes and stomachs of many computer scientists!" Knuth in Art of Computer Programming, Vol. 2

```
Too short of a period
```

```
Ex: a = 13; b = 0; m = 16
```

LCG(13,0,16,1,9)

[1] 13 9 5 1 13 9 5 1 13

period is 4!

Ex: a = 17; b = 7; m = 16

```
LCG(17,7,16,1,18)

## [1] 8 15 6 13 4 11 2 9 0 7 14 5 12 3 10 1 8 15

period is 16!
```

Full period obtained when p = m

By the Hull-Dobell Theorem, this occurs iff

- gcd(b, m) = 1
- If q|m and q is prime, then q|(a-1)
- If 4|m, then 4|(a-1)

Back to our ex:

$$a = 17$$
; $b = 7$; $m = 16$

Full period!

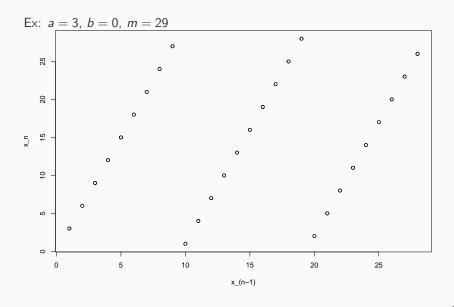
Ex: Numerical Recipes uses

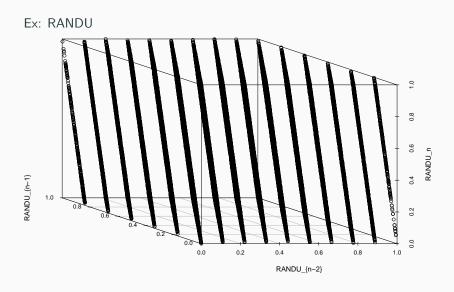
$$m = 2^{32}$$
; $a = 1664525$; $b = 1013904223$

- gcd(b,m)=1 as b is odd
- 2 is only prime that divides m, and 2|1664524
- $4|m \text{ and } \frac{1664524}{4} = 416131$

Other main issue is a number-theoretic property spotted by G. Marsaglia in a famous article ("Random numbers fall mainly in planes", 1968 Proc. Nat. Acad. Sci.):

- LCGs result in sequences which fall along hyperplanes in some number of dimensions; at most (but sometimes much less than) \sqrt{m} .
- There are several tests of randomness (mentioned e.g. by Knuth) which test how finely spaced these hyperplanes are (e.g., "spectral test" of Coveyou - Macpherson)
- Idea: further apart planes are, the worse the LCG
- RANDU fails the spectral test in dimension 3





To test the randomness of a given sequence, we can use the χ^2 goodness of fit test for the mulitnomial distribution

Used to assess the equidistribution of the non-overlapping K-tuples

$$(x_n, \ldots, x_{n+K-1}), n = 0, K, 2K, \ldots$$

by tabulating the counts of these K-tuples falling in sets $A \subset [0,1]^K$ of the form

$$\prod_{i=1}^K \left[a_i/L, \left(a_i+1\right)/L\right)$$

for $a_i \in \{0, 1, ..., L-1\}$

Idea: Divide the unit K-cube into L^K equal sized bins, let p_i denote the probability of an observation falling into bin i

Test

$$H_0: p_i = rac{1}{L^K}$$
 for all i; $H_a: p_i
eq rac{1}{L^K}$ for some i

Use Pearson's χ^2 test:

under
$$H_0$$
, $T = \sum_i \frac{(n_i - np_i)^2}{np_i} \sim (approx.)\chi^2_{L^K-1}$

A quick R code is given by (ncoord corresponds to K and nquant corresponds to L)

```
FitNtupl<- function(ncoord, nquant, indata)</pre>
{
  ## assumes block of pseudo-random uniform[0,1)
  ## numbers in indata; to be tested for fit based on
  ## empirically generated contingency-table of nquant
  ## equal-length intervals in each of ncoord
  ## consecutive coordinates
  ntup = length(indata) %/% ncoord
  idata = c(matrix(trunc(nquant * indata - 1e-11),
            ncol = ncoord) %*% nquant^(0:(ncoord - 1))) + 1
  cellexp = ntup/(nquant^ncoord)
  cells = table(idata)
```

```
code con...
diagind = 1 + (0:(nquant - 1)) * sum(nquant^(0:(ncoord - 1)))
# overall balance
chistat = sum((cells - cellexp)^2)/cellexp
# fraction falling in diagonal cells
# test for coord agreement
diagstat = (sum(cells[diagind]) - nquant *
              cellexp)^2/(nquant *cellexp *
              (1 - ((cellexp * nquant)/ntup)))
list(chisq = chistat, pval = 1 - pchisq(chistat,
        nquant^ncoord - 1), diagstat = diagstat,
        diagPval = 1-pchisq(diagstat, 1), CountTbl = cells)
```

```
set.seed(123)
x <- FitNtupl(2,4,runif(1.e4))
x$chisq
## [1] 14.4576
x$pval
## [1] 0.4911446
x$diagstat
## [1] 0.05226667
x$diagPval
## [1] 0.819165
```

Try RANDU

```
x <- LCG(65539,0,2^(31),1,2e5)
x <- x/(2^(31))
unlist(FitNtupl(2,10,x))[1:4]</pre>
```

```
## chisq pval diagstat diagPval
## 101.1480000 0.4211794 0.2250000 0.6352563
```

No wonder it was used so much!

Fails more complex tests (ex: spectral test)

For a variety of distributions, R has built in pseudogenerators for

- r: generate RV from the given distribution
- d: density functions (for continuous distributions) or probability mass functions (for discrete distributions)
- **p**: cumulative distribution function $F(x) = \mathbb{P}(X \le x)$
- **q**: quartile function, $Q(u) = F^{-1}(u)$ (For discrete distributions the quantile is the smallest integer m such that $F(m) \ge u$.)

Ex: For the normal distribution, calls are rnorm, dnorm, pnorm, qnorm

```
rnorm(5, mean=0,sd=1)
## [1] -0.4941739 1.1275935 -1.1469495 1.4810186 0.9161912
dnorm(1, mean=0,sd=1)
## [1] 0.2419707
pnorm(1, mean=0,sd=1)
## [1] 0.8413447
qnorm(0.84, mean=0, sd=1)
## [1] 0.9944579
qnorm(pnorm(1, mean=0,sd=1), mean=0,sd=1)
## [1] 1
```

 Table 5.1: S function names and parameters for standard probability distributions.

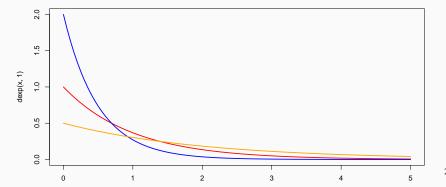
Distribution	S name	Parameters
beta	beta	shape1, shape2
binomial	binom	size, prob
Cauchy	cauchy	location, scale
chi-squared	chisq	df
exponential	exp	rate
F	f	df1, df2
gamma	gamma	shape, rate
geometric	geom	prob
hypergeometric	hyper	m, n, k
log-normal	lnorm	meanlog, sdlog
logistic	logis	location, scale
negative binomial	nbinom	size, prob
normal	norm	mean, sd
Poisson	pois	lambda
T	t	df
uniform	unif	min, max
Weibull	weibull	shape, scale
Wilcoxon	wilcox	m, n

Example: Exponential distribution

```
dexp(0:4,1) # rate * e^{-rate* x}

## [1] 1.00000000 0.36787944 0.13533528 0.04978707 0.01831564

x <- (0:100)/20
plot(x,dexp(x,1),lwd=2,col="red",type='l',ylim=c(0,2))
points(x,dexp(x,2),lwd=2,col="blue",type='l')
points(x,dexp(x,0.5),lwd=2,col="orange",type='l')</pre>
```

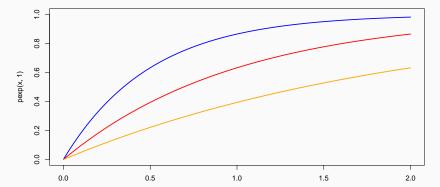


Example: Exponential distribution

```
pexp(0:4,1) # 1 - e^{-rate* x}

## [1] 0.0000000 0.6321206 0.8646647 0.9502129 0.9816844

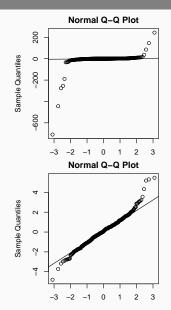
x <- (0:100)/50
plot(x,pexp(x,1),lwd=2,col="red",type='l',ylim=c(0,1))
points(x,pexp(x,2),lwd=2,col="blue",type='l')
points(x,pexp(x,0.5),lwd=2,col="orange",type='l')</pre>
```

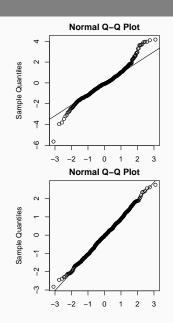


Ex: Q-Q plots

```
par(mfrow=c(2,2),mar=c(2, 2, 2, 2))
dof <- c(1,5,10,100)
for(i in 1:4){
x <- rt(250, df = dof[i])
par(pty = "s")
qqnorm(x); qqline(x)}</pre>
```

The greater spread of the extreme quantiles for the data is indicative of a long- tailed distribution.





```
Ex: MLE (more on this later!)
x <- rnorm(300,0,3)
nll <- function(m,s) -sum(dnorm(x, mean = m, sd = s, log = TRUE)
library(MASS)
fitdistr(x, "normal")

## mean sd
## 0.1793423 2.8662987
## (0.1654858) (0.1170162)</pre>
```

```
Ex: MLE
library(stats4)
M1 <- mle(nll, start=list(m=-5, s=1), method="L-BFGS-B")
M2 <- mle(nll, start=list(m=-5, s=1), method="CG")
summary(M1)
## Maximum likelihood estimation
##
## Call:
## mle(minuslog1 = nll, start = list(m = -5, s = 1), method = "L
##
## Coefficients:
## Estimate Std. Error
## m 0.1793422 0.1654859
## s 2.8662992 0.1170161
##
## -2 log L: 1483.176
```

```
Fx: MIF
summary(M2)
## Warning in sqrt(diag(object@vcov)): NaNs produced
## Maximum likelihood estimation
##
## Call:
## mle(minuslog1 = nll, start = list(m = -5, s = 1), method = "C
##
## Coefficients:
## Estimate Std. Error
## m 143.9835 54.15994
## s 979.6245
                    NaN
##
## -2 log L: 4689.765
```

Generating random data

Some important commands are

- sample(n) select a random permutation of 1, · · · , n
- sample(x) randomly permute the elements of x
- sample(x, replace = T) generate a bootstrap sample of x
- sample(x, n) sample n items from x without replacement
- sample(x, n, replace = T) sample n items from x with replacement
- sample(x, n, replace = T, prob=p) probability sample of n items from x with replacement

Data summaries

If x is a data vector, some important commands are (note, if missing data, can often add **na.rm=TRUE** as an argument)

- mean(x) computes the mean of x
- sd(x) computes the sample standard deviation of x (note, uses

$$\sqrt{\frac{\sum_{i=1}^{n}(x_i-\bar{x})^2}{n-1}}\right)$$

- var(x) computes the sample variance (variance-covariance matrix when given a matrix input)
- median(x) computes the median of x
- quantile(x) quantiles at (0, 0.25, 0.5, 0.75, 1) by default; interpolates

$$\begin{aligned} \textit{quantile}(x,p) = & [1 - (p(n-1) - \lfloor p(n-1) \rfloor)] x_{(1 + \lfloor p(n-1) \rfloor)} \\ & + [p(n-1) - \lfloor p(n-1) \rfloor] x_{(2 + \lfloor p(n-1) \rfloor)} \end{aligned}$$

summary(x) gives min, Q1, median, mean, Q3, and max of a vector

Classical statistical tests

Many classical univariate (and multivariate!) tests are included in R Examples include:

```
binom.test chisq.test cor.test fisher.test friedman.test kruskal.test mantelhaen.test mcnemar.test prop.test t.test var.test wilcox.test chisq.gof ks.gof
```

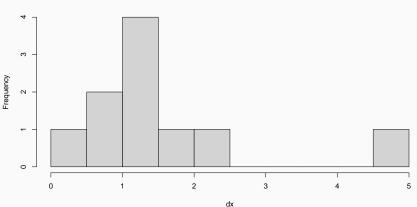
with many, many more (Google is your friend here!)

```
# using the sleep data set
library(ggplot2)
## Warning: package 'ggplot2' was built under R version 4.2.3
data(sleep)
head(sleep)
## extra group ID
## 1 0.7 1 1
## 2 -1.6 1 2
## 3 -0.2 1 3
## 4 -1.2 1 4
## 5 -0.1 1 5
## 6 3.4 1 6
x1 <- sleep[sleep$group==1, 1]</pre>
x2 <- sleep[sleep$group==2, 1]
dx < -x2 - x1
```

visualize the difference

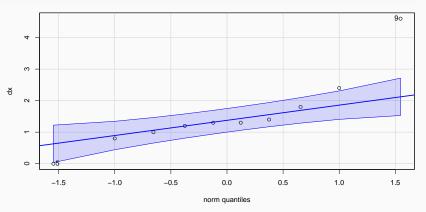
hist(dx,breaks=7)





```
# visualize the difference
dx.df <- data.frame(dx = sleep$extra[sleep$group == 2] -</pre>
                          sleep$extra[sleep$group == 1])
ggplot(dx.df, aes(x = dx)) + geom_histogram(bins=7)
 3-
count
 2-
```

visualize the difference
qqp(dx)



[1] 9 5

("confidence envelope is based on the SEs of the order statistics of an independent random sample from the comparison distribution")

```
# paired t-test
tp <- t.test(x1,x2, paired = TRUE,</pre>
             alternative = "less")
tp
##
## Paired t-test
##
## data: x1 and x2
## t = -4.0621, df = 9, p-value = 0.001416
## alternative hypothesis: true mean difference is less than 0
## 95 percent confidence interval:
       -Inf -0.8669947
##
## sample estimates:
## mean difference
##
             -1.58
```

```
(warnings suppressed)
# Wilcoxon ('Mann-Whitney' test)
wp <- wilcox.test(x2,x1, paired = TRUE,
             alternative = "greater")
wр
##
##
    Wilcoxon signed rank test with continuity correction
##
## data: x2 and x1
## V = 45, p-value = 0.004545
## alternative hypothesis: true location shift is greater than 0
```