#### **Resampling Techniques and their Application**

#### -Class 4-

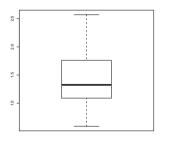
#### Frank Konietschke

Institut für Biometrie und Klinische Epidemiologie Charité - Universitätsmedizin Berlin, Berlin frank.konietschke@charite.de



## **Motivation and Examples-II**

The diameter of cork of a Champagne bottle is supposed to be 1.5 cm. If the cork is either too large or too small, it will not fit in the bottle. The manufacturer measures the diameter in a random sample of **n=36** bottles and obtains:

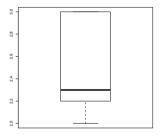


```
x=c(
0.59, 1.23, 1.00, 0.84, 0.88, 1.71,
1.81, 1.84, 2.03, 1.39, 1.30, 1.31,
1.96, 1.33, 2.57, 1.19, 1.01, 2.06,
1.32, 1.55, 1.28, 0.93, 1.63, 1.24,
1.83, 1.81, 0.94, 1.46, 1.25, 1.56,
0.61, 0.83, 1.17, 2.24, 1.68, 1.51)
```

Data Analysis: Confidence interval and t-test

## **Motivation and Examples-III**

A researcher measures the reaction time of n = 10 mice to signal pain when a stitch is applied to their tail.



Data Analysis: Confidence interval and t-test

The diameter of cork of a Champagne bottle is supposed to be 1.5 cm. If the cork is either too large or too small, it will not fit in the bottle. The manufacturer measures the diameter in a random sample of n=36 bottles. Is there evidence at 1% level that the true mean diameter has moved away from the target?

• Hypothesis:  $H_0: \mu = 1.5 \text{ vs. } H_1: \mu \neq 1.5$ 

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mx <- mean(x)
vx <- var(x)
T <- sqrt(36)*(mx-1.5)/sqrt(vx)
p <- 2*min(pt(T,35),1-pt(T,35))</pre>
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- Estimator:  $\widehat{\mu} = 1.4136$
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- Quality of the estimator? Is the method valid? Is p = 0.27 a good estimate?

A researcher measures the reaction time of n = 10 mice to signal pain when a stitch is applied to their tail. Compute a 95% confidence interval.

```
x=c(2.4, 3.0, 3.0, 2.2, 2.2,
2.2, 2.2, 2.8, 2.0, 3.0)
mx <- mean(x)
vx <- var(x)
crit <- qt(0.975,9)
lower <- mx-crit/sqrt(10)*sqrt(vx)
upper <- mx+crit/sqrt(10)*sqrt(vx)</pre>
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• Estimator:  $\widehat{\mu} = 2.5$ 

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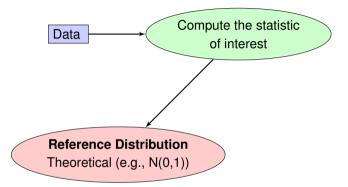
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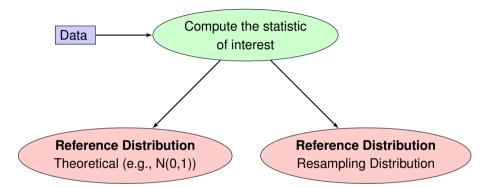
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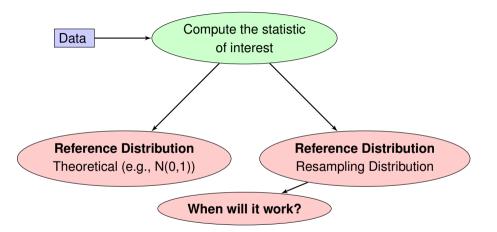
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- Quality of the confidence interval?

Data









The resampling test will control the type-I error  $\alpha$  if and only if the resampling distribution of the statistic mimics the distribution of the test, at least asymptotically.

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  - Drawing with replacement is just one of them
  - Today, we will study few more

• Randomly sample/generate from data  $\mathbf{X} = (X_1, \dots, X_n)$ 

$$X_1^*,\ldots,X_n^*$$

such that

$$T^* = \sqrt{n} \frac{\overline{X}_{\cdot}^* - E(\overline{X}_{\cdot}^* | \mathbf{X})}{\widehat{\sigma}^*}$$

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• Compute p-values, critical values etc from the **Resampling Distribution** 

• Data  $X = (X_1, \dots, X_n)$  (fixed values)

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In R: sample(x,replace=TRUE)

#### **Permutation**

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- Note that  $E(X_k^*)=0$ ,  $Var(X_k^*)=\widehat{\sigma}^2$  and  $\mu_3(X_k^*)=\widehat{\mu}_3$ . (Why?)

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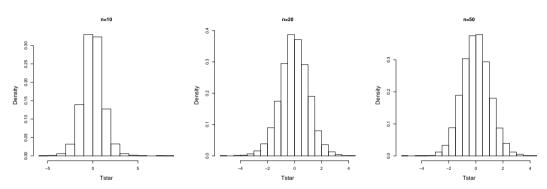
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- 4. Safe the value of  $T^*$  in  $T_{\ell}$
- 5. Repeat the steps above a large number of times, e.g.  $n_{boot} = 10,000$
- 6. Compute the critical value  $c^*(1-\alpha/2)$  or the p-value as

$$p = 2 \min \left\{ \frac{1}{n_{boot}} \sum_{\ell=1}^{n_{boot}} \mathcal{I}(T_{\ell} \leq T), \frac{1}{n_{boot}} \sum_{\ell=1}^{n_{boot}} \mathcal{I}(T_{\ell} \geq T) \right\}$$

 $\mathcal{I}(x)$  is an indicator

## **Illustration Resampling Distribution**



Compute the  $(1 - \alpha)$ -quantile

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# What is $E(\overline{X}^* | \mathbf{X})$ ?-II

- Compute the following terms

  - $E(X_1^*|\mathbf{X})$   $E(X_1^{2,*}|\mathbf{X}) =$
  - $E(\overline{X}^*|\mathbf{X}) =$
  - $E(\widehat{\sigma}^{2,*}|\mathbf{X}) =$
- For all of the different resampling variables

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  - Generate a matrix of resampling variables (example next slide)

```
skew<-function(x){
n<-length(x)
mx<-mean(x)
sdx<-sd(x)
n/((n-1)*(n-2))*sum(((x-mx)/sdx)^3)}</pre>
```

```
myboot <- function(n,Distribution,nboot,nsim){</pre>
T=Tboot = Tbootsp =c()
#-----#
if(Distribution == "Normal"){
x <- matrix(rnorm(n*nsim),ncol=nsim)}</pre>
if(Distribution == "Exp"){
x <- matrix(rexp(n*nsim)-1,ncol=nsim)}</pre>
mx <- colMeans(x)</pre>
vx <- (colSums(x^2)-n*mx^2)/(n-1)
T=sqrt(n)*(mx)/sqrt(vx)
```

```
------Simulate the skew parametric Bootstrap---#
mu3=skew(x[.i])
f=8/mu3^2
xstarsp=matrix(sign(mu3)*sqrt(vx)*((rchisq(n*nboot,f)-f)/sqrt(2*f)),ncol=nboot)
mxbootsp=colMeans(xstarsp)
vxbsp=(colSums(xstarsp^2) - n*mxbootsp^2)/(n-1)
Tstarnp = sqrt(n)*mxbootsp/sqrt(vxbsp)
p1sp = mean(Tstarnp >= T[i])
p2sp = mean(Tstarnp <= T[i])
Thootsp[i] = (2*min(p1sp,p2sp)<0.05)}
```

```
result = data.frame(n=n,nsim=nsim,nboot=nboot,Dist=Distribution,
T=mean(abs(T)>qt(0.975,n-1)),
Tboot = mean(Tboot), Tbootsp = mean(Tbootsp))
result}
myboot(10,"Exp",1000,1000)
```

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  - Overdispersion parameters, variance of their estimators,...

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- How to estimate  $f(\theta)$ ?

Estimation of $\theta$	Properties
Maximum-Likelihood	F must be known
	Algorithm can be difficult
	Algorithm might not converge
	Large sample for distribution
Moment based	F can be unknown
	Computation usually feasible
	Usually exist (no converging issues)
	Small sample approximations
Resampling Methods	

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• Basically, we simulate data from  $\widehat{F}_n$ 

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  - Estimate the parameter of interest using the values of  $\widehat{ heta}_1,\dots,\widehat{ heta}_{n_{boot}}$

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```
X=c(2.4, 3.0, 3.0, 2.2, 2.2,
2.2, 2.2, 2.8, 2.0, 3.0)
n < -10
nboot <-10000
B<- apply(matrix(1:n,
ncol=nboot,nrow=n),
2, sample, replace=TRUE)
xstar <- matrix(X[B],</pre>
ncol=nboot.nrow=n)
mxstar <- colMeans(xstar)</pre>
tauhat2 <- var(mxstar)</pre>
```

#### **Project: How Good is the Estimator?**

- Use computer simulations to assess the quality of the estimator  $\hat{\tau}^2$
- Compare with  $\hat{\tau}_{emp}^2 = \hat{\sigma}^2/n$
- Compare the bias and MSE of the estimators
  - Generate  $n_{sim}$  random samples from different distributions
  - ullet Compute the estimator upon the sample and safe the value in  $\widehat{ heta}_s$
  - Assess the bias and MSE

$$Bias = rac{1}{n_{sim}} \sum_{s=1}^{n_{sim}} (\widehat{ heta}_s - heta) \text{ and } MSE = rac{1}{n_{sim}} \sum_{s=1}^{n_{sim}} (\widehat{ heta}_s - heta)^2$$

## **Project**

- Task: Estimate the variance of correlation coefficients using resampling strategies
- How to get an idea about the true variance?
- Is bootstrap a good way?