Resampling Techniques and their Application

-Class 12-

Frank Konietschke

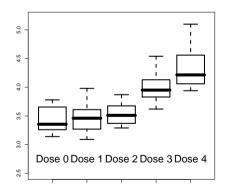
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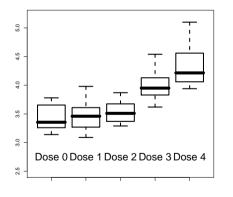
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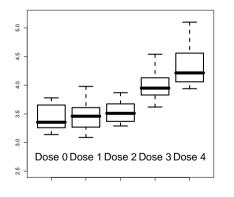


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x2 = c(3.46, 3.98, 3.09, 3.49, 3.31, 3.73, 3.23)
x3 = c(3.71, 3.36, 3.38, 3.64, 3.41, 3.29, 3.61, 3.87)
x4 = c(3.86, 3.80, 4.14, 3.62, 3.95, 4.12, 4.54)
x5 = c(4.19, 4.16, 3.94, 4.26, 4.86, 3.96, 4.24, 5.10)
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- Toxicity trial: 40 rats were randomized into 5 dose groups (Dose 0 Dose 4)
- After treatment: relative liver weight of each animal
- Question: Which dose(s) differ from control? Trend?



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Multiple hypotheses

In general

$$H_0: \mathbf{C}oldsymbol{\mu} = \mathbf{0}$$
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- It is on us to define the alternative H₁
- The contrast matrix **C** is nothing but the pattern of the alternative H_1
- In general, **C** is a $q \times a$ matrix. Each row vector is a contrast.

$$m{C} = \left(egin{array}{c} m{c}_1' \ dots \ m{c}_q' \end{array}
ight) \;\; = \;\; \left(egin{array}{cccc} \mu_1 & \mu_2 & \dots & \mu_a \ c_{11} & c_{12} & \cdots & c_{1a} \ c_{21} & c_{22} & \cdots & c_{2a} \ dots & dots & dots & dots \ c_{q1} & c_{q2} & \cdots & c_{qa} \end{array}
ight); \; \sum_{i=1}^a c_{\ell i} = 0, \; \ell = 1, \dots, q \ \end{array}$$

Example 1: Many-to-one comparisons (Dunnett):

$$H_1: \left\{ egin{array}{ll} \mu_1
eq \mu_2 \\ \mu_1
eq \mu_3 \\ \vdots \\ \mu_1
eq \mu_a \end{array}
ight. \Leftrightarrow oldsymbol{C} = \left(egin{array}{ccccccc} -1 & 1 & 0 & \cdots & 0 \\ -1 & 0 & 1 & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ -1 & 0 & \cdots & \cdots & -1 \end{array}
ight)$$

General Contrasts (II)

Example 2: Trend (Williams)

$$H_{1}: \left\{ \begin{array}{c} \mu_{1} \neq \mu_{a} \\ \mu_{1} \neq \mu_{a-1} = \mu_{a} \\ \vdots \\ \mu_{1} \neq \mu_{2} = \ldots = \mu_{a} \end{array} \right. \Leftrightarrow \boldsymbol{C} = \left(\begin{array}{cccc} -1 & 0 & 0 & \cdots & 1 \\ -1 & 0 & 0 & \frac{n_{a-1}}{n_{a-1} + n_{a}} & \frac{n_{a}}{n_{a-1} + n_{a}} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ -1 & \frac{n_{2}}{n_{2} + \ldots + n_{a}} & \cdots & \cdots & \frac{n_{a}}{n_{2} + \ldots + n_{a}} \end{array} \right)$$

General Contrasts (III)

Example 3: All-pairs (Tukey):

$$H_{1}: \left\{ \begin{array}{c} \mu_{1} \neq \mu_{2} \\ \mu_{1} \neq \mu_{3} \\ \vdots \\ \mu_{1} \neq \mu_{a} \\ \mu_{2} \neq \mu_{3} \\ \vdots \\ \mu_{a-1} \neq \mu_{a} \end{array} \right. \Leftrightarrow \boldsymbol{C} = \left(\begin{array}{cccc} -1 & 1 & 0 & \cdots & 0 \\ -1 & 0 & 1 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & \cdots & -1 & 1 \end{array} \right)$$

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And many more

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- And many more
- See contrMat function in multcomp package

Individual hypothesis

$$H_0^{(\ell)}: \mathbf{c}'_\ell \mu = 0, \ \ell = 1, \ldots, q$$

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- Global null hypothesis

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$$H_0: \mathbf{C}oldsymbol{\mu} = \mathbf{0}$$

• Reject H_0 , if any $H_0^{(\ell)}$ is rejected

Multiple Contrast Tests

Estimators

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 - $\overline{\mathbf{X}}_{\cdot} = (\overline{X}_{1\cdot}, \dots, \overline{X}_{a\cdot})'$ (vector of means)
- Variance of contrasts in means

$$\mathbf{S} = Cov(\overline{\mathbf{X}}.) = diag\left(\frac{\sigma_1^2}{n_1}, \dots, \frac{\sigma_a^2}{n_a}\right)$$

$$\Gamma = Cov(\mathbf{C}\overline{\mathbf{X}}.) = \mathbf{CSC}' = \mathbf{C}diag\left(\frac{\sigma_1^2}{n_1}, \dots, \frac{\sigma_a^2}{n_a}\right)\mathbf{C}'$$

$$\sigma_\ell^2 = Var(\mathbf{c}'_\ell \overline{\mathbf{X}}.) = \sum_{i=1}^a \sigma_i^2 \frac{\mathbf{c}_{\ell i}^2}{n_i} = \mathbf{c}'_\ell \mathbf{Sc}_\ell$$

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Variance of a contrast

$$\hat{\mathbf{S}} = diag\left(\frac{s_1^2}{n_1}, \dots, \frac{s_a^2}{n_a}\right)$$

$$\hat{\Gamma} = \mathbf{C}\hat{\mathbf{S}}\mathbf{C}' = \mathbf{C}diag\left(\frac{s_1^2}{n_1}, \dots, \frac{s_a^2}{n_a}\right)\mathbf{C}'$$

$$\hat{\sigma}_{\ell}^2 = \mathbf{c}_{\ell}'\hat{\mathbf{S}}\mathbf{c}_{\ell}$$

Estimators

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Variance estimators

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Variance of a contrast

$$\begin{split} \widehat{\mathbf{S}} &= \operatorname{diag}\left(\frac{s_1^2}{n_1}, \dots, \frac{s_a^2}{n_a}\right) \\ \widehat{\boldsymbol{\Gamma}} &= \mathbf{C}\widehat{\mathbf{S}}\mathbf{C}' = \mathbf{C}\operatorname{diag}\left(\frac{s_1^2}{n_1}, \dots, \frac{s_a^2}{n_a}\right)\mathbf{C}' \\ \widehat{\sigma}_{\ell}^2 &= \mathbf{c}_{\ell}'\widehat{\mathbf{S}}\mathbf{c}_{\ell} \end{split}$$

```
library(multcomp)
C <-contrMat(rep(1,5), "Dunnett")</pre>
X < -c(x1, x2, x3, x4, x5)
n < -c(8.7.8.7.8)
N < -sum(n)
grp<-factor(c(rep(1:5,n)))</pre>
a<-5
Dat<-data.frame(X=X,grp=grp)
Xbar<-aggregate(X~grp,data=Dat,mean)[,2]</pre>
si2 <-aggregate(X~grp,data=Dat,var)[,2]
Shat <- diag(si2/n)
Gammahat<-C%*%Shat%*%t(C)
```

Multiple Comparisons

ullet For $H_0^{(\ell)}: \mathbf{c}_\ell' oldsymbol{\mu} = 0$

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• Are the test statistics $T_{(\ell)}$ and $T_{(\ell')}$ independent?

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• Covariance of $\mathbf{c}'_{\ell}\overline{\mathbf{X}}$ and $\mathbf{c}'_{m}\overline{\mathbf{X}}$:

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- Computation of c orrelation matrix R:

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- ullet $\widehat{f R} = diag(\widehat{f \Gamma})^{-1/2} \Gamma diag(\widehat{f \Gamma})^{-1/2}$

Test decisions

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 - Reject $H_0^{(\ell)}: \mathbf{c}_\ell' \mu = 0$ if $|T_\ell| \geq t_{1-lpha}(\widehat{m{R}})$

$$\boldsymbol{c}_{\ell}'\overline{\mathbf{X}}_{\cdot}\pm t_{1-\alpha}(\widehat{\boldsymbol{R}})\widehat{\sigma}_{\ell},\ \ell=1,\ldots,q$$

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- $t_{1-\alpha}(\widehat{\mathbf{R}})$: $(1-\alpha)$ -quantile of the multivariate $T(\mathbf{0}, \nu, \widehat{\mathbf{R}})$ distribution
- Or, compute adjusted p-values using the $T(\mathbf{0}, \nu, \widehat{\mathbf{R}})$ distribution

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```
library(multcomp)
```

```
nu=round(min(nul))
Rhat<-cov2cor(Gammahat)
set.seed(1)
tmax=qmvt(0.95,tail="both",corr=Rhat,
df=nu)$quantile
T0>=tmax
$
```

```
pv<-sapply(1:4,function(j)
1-pmvt(-abs(Tl[j]),abs(Tl[j]),df=nu,
delta=rep(0,4),corr=Rhat)[1])</pre>
```

Properties

- Method is a multiple t-test
- In case of small samples, method might be liberal or conservative
- Resampling methods to improve the approximation /method
- Goal: Approximate the distribution of the maximum

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- 4. Compute $\overline{X}_{i\cdot}^*$ and $s_{i\cdot}^{2,*}$ (means and variances)
- 5. Compute means $\overline{\mathbf{X}}^*$ and variance estimator $\widehat{\sigma}_{\ell}^{2,*}$ with the resampling variables
- 6. Compute the vector of test statistics

$$T_{\ell}^* = rac{oldsymbol{c}_{\ell}'(\overline{f X}_{\cdot}^* - E(\overline{f X}_{\cdot}^* | f X))}{\widehat{\sigma}_{\ell}^*}$$

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- 8. Reject H_0 , if the p-value

$$\frac{1}{n_{boot}} \sum_{s=1}^{n_{boot}} \mathcal{I}(A_s^* \ge T_0) < \alpha$$

 Many different ways of generating X* are possible

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