# Scenario generation for nongaussian time series via Quantile Regression

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Introduction

## Introduction

#### Motivation

- Renewable energy scenarios are important in many fields in Power Systems:
  - 1. Energy trading;
  - 2. unit commitment;
  - 3. grid expansion planning;
  - 4. investment decisions
- In stochastic optimization problems, a set of scenarios is a needed input.
- ▶ Robust optimization requires bounds for probable values.

Change in paradigm: from predicting the conditional mean to predicting the conditional distribution

# Probability Forecasting Approaches

- Parametric Models
  - Assume a distributional shape
  - Low computational costs
  - Faster convergence
  - Examples: Arima-GARCH, GAS
- Nonparametric Models
  - Don't require a distribution to be specified
  - High computational cost
  - Needs more data to produce a good approximation
  - Examples: Quantile Regression, Kernel Density Estimation,
     Artificial Intelligence

# The nongaussianity of Wind Power



- Renewables, such as wind and solar power have reportedly nongaussian behaviour
- Convenience of using a nonparametric approach, which doesn't rely on assuming a distribution
- Puantile regression is the chosen technique available to model this time series dynamics, by estimating a thin grid of  $\alpha$ -quantiles at once and forming a data-driven conditional distribution

Quantile Regression

# Quantile Regression

## Definition of the Conditional Quantile

Let the  $\alpha$ -conditional quantile function of Y for a given value x of the d-dimensional random variable X, i.e.,  $Q_{Y|X}:[0,1]\times\mathbb{R}^d\to\mathbb{R}$ , can be defined as:

$$Q_{Y|X}(\alpha, x) = F_{Y|X}^{-1}(\alpha, x) = \inf\{y : F_{Y|X}(y, x) \ge \alpha\}. \tag{1}$$

# Conditional Quantile from a sample

Let a dataset be composed from  $\{y_t, x_t\}_{t \in \mathcal{T}}$  and let  $\rho$  be the check function

$$\rho_{\alpha}(x) = \begin{cases} \alpha x & \text{if } x \ge 0\\ (1 - \alpha)x & \text{if } x < 0 \end{cases}, \tag{2}$$

The sample quantile function for a given probability  $\alpha$  is then based on a finite number of observations and is the solution to minimizing the loss function  $L(\cdot)$ :

$$\hat{Q}_{Y|X}(\alpha,\cdot) \in \underset{q}{\text{arg min}} L_{\alpha}(q) = \sum_{t \in T} \rho_{\alpha}(y_t - q(x_t)).$$
 (3)

# Conditional Quantile as a Linear Programming Problem

$$\begin{split} \min_{\beta_{0\alpha},\beta_{\alpha},\varepsilon_{t\alpha}^{+},\varepsilon_{t\alpha}^{-}} & \sum_{\alpha \in A} \sum_{t \in T} \left(\alpha \varepsilon_{t\alpha}^{+} + (1-\alpha)\varepsilon_{t\alpha}^{-}\right) \\ \text{subject to} & & & & & & & & \\ \varepsilon_{t\alpha}^{+} - \varepsilon_{t\alpha}^{-} &= y_{t} - \beta_{0\alpha} - \beta_{\alpha}^{T}x_{t}, & \forall t \in T, \forall \alpha \in A, \\ \varepsilon_{t\alpha}^{+}, \varepsilon_{t\alpha}^{-} &\geq 0, & \forall t \in T, \forall \alpha \in A, \\ \beta_{0\alpha} + \beta_{\alpha}^{T}x_{t} &\leq \beta_{0\alpha'} + \beta_{\alpha'}^{T}x_{t}, \\ \forall t \in T, \forall (\alpha, \alpha') \in A \times A, \alpha < \alpha', \end{split}$$

• We apply QR to estimate the conditional distribution  $\hat{Q}_{Y_{t+k}|X_{t+k},Y_t,Y_{t-1},...}(\alpha,\cdot)$  for a k-step ahead forecast of time serie  $\{y_t\}$ , where  $X_{t+k}$  is a vector of exogenous variables at the time we want to forecast.

Regularization

# Regularization

#### Best Subset selection via MILP

- Mixed Integer Linear Programming (MILP) models allow only K variables to be used for each  $\alpha$ -quantile. This means that only K coefficients  $\beta_{p\alpha}$  may have nonzero values, for each  $\alpha$ -quantile. It must be guaranteed by constraints on the optimization problem.
- ▶ We present three forms of regularization using MILP

## MILP - One model for each $\alpha$ -quantile

$$\min_{\substack{\beta_{0\alpha},\beta_{\alpha},z_{p\alpha}\varepsilon_{t\alpha}^{+},\varepsilon_{t\alpha}^{-}\\ }} \sum_{\alpha\in A} \sum_{t\in \mathcal{T}} \left(\alpha\varepsilon_{t\alpha}^{+} + (1-\alpha)\varepsilon_{t\alpha}^{-}\right) \tag{4}$$

s.t 
$$\varepsilon_{t\alpha}^+ - \varepsilon_{t\alpha}^- = y_t - \beta_{0\alpha} - \sum_{p=1}^P \beta_{p\alpha} x_{t,p}, \quad \forall t \in T, \forall \alpha \in A,$$

$$\varepsilon_{t\alpha}^+, \varepsilon_{t\alpha}^- \geq 0,$$

$$-Mz_{p\alpha} \leq \beta_{p\alpha} \leq Mz_{p\alpha}$$

$$\sum_{p=1}^{P} z_{p\alpha} \leq K,$$

$$z_{p\alpha} \in \{0, 1\},\$$

$$\beta_{0\alpha} + \beta_{\alpha}^T x_t \leq \beta_{0\alpha'} + \beta_{\alpha'}^T x_t,$$

$$\forall t \in T, \forall \alpha \in A,$$
 (5)

$$\forall t \in T, \forall \alpha \in A,$$
 (6)

$$\forall \alpha \in A, \forall p \in P,$$
 (7)

$$\forall \alpha \in A,$$
 (8)

$$\forall \alpha \in A, \forall p \in P, \tag{9}$$

$$\forall t \in T, \forall (\alpha, \alpha') \in A \times A, \alpha < \alpha',$$

(10)

# MILP - Defining groups for $\alpha$ -quantiles

$$\min_{\beta_{0\alpha},\beta_{\alpha},z_{p\alpha}\varepsilon_{t\alpha}^{+},\varepsilon_{t\alpha}^{-}} \qquad \sum_{\alpha\in A} \sum_{t\in T} \left(\alpha\varepsilon_{t\alpha}^{+} + (1-\alpha)\varepsilon_{t\alpha}^{-}\right) \tag{11}$$

s.t 
$$\varepsilon_{t\alpha}^+ - \varepsilon_{t\alpha}^- = y_t - \beta_{0\alpha} - \sum_{p=1}^{P} \beta_{p\alpha} x_{t,p}, \quad \forall t \in T, \forall \alpha \in A,$$
 (12)

$$\varepsilon_{t\alpha}^+, \varepsilon_{t\alpha}^- > 0, \qquad \forall t \in T, \forall \alpha \in A,$$
 (13)

$$-Mz_{p\alpha g} \leq \beta_{p\alpha} \leq Mz_{p\alpha g}, \qquad \forall \alpha \in A, \forall p \in P, \forall g \in G$$
 (14)

$$z_{p\alpha g} := 2 - (1 - z_{pg}) - I_{g\alpha}$$
 (15)

$$\sum_{p=1}^{P} z_{pg} \le K, \qquad \forall g \in G, \tag{16}$$

$$\beta_{0\alpha} + \beta_{\alpha}^{\mathsf{T}} x_{t} \leq \beta_{0\alpha'} + \beta_{\alpha'}^{\mathsf{T}} x_{t}, \qquad \forall t \in \mathsf{T}, \forall (\alpha, \alpha') \in \mathsf{A} \times \mathsf{A}, \alpha < \alpha',$$

(17)

$$\sum I_{g\alpha} = 1, \qquad \forall \alpha \in A, \tag{18}$$

$$I_{g\alpha}, z_{pg} \in \{0, 1\}, \qquad \forall p \in P, \forall g \in G,$$
 (19)

#### MILP - Penalization of derivative

$$\min_{\beta_{0\alpha},\beta_{\alpha},z_{p\alpha}\varepsilon_{t\alpha}^{+},\varepsilon_{t\alpha}^{-}}$$

$$\sum_{\alpha \in A} \sum_{t \in T} \left( \alpha \varepsilon_{t\alpha}^{+} + (1 - \alpha) \varepsilon_{t\alpha}^{-} \right) + \gamma \sum_{\alpha \in A'} D_{2\rho\alpha}$$
 (20)

s.t 
$$\varepsilon_{t\alpha}^+ - \varepsilon_{t\alpha}^- = y_t - \beta_{0\alpha} - \sum_{\rho=1}^P \beta_{\rho\alpha} x_{t,\rho}, \quad \forall t \in T, \forall \alpha \in A,$$
 (21)

$$\varepsilon_{t\alpha}^{+}, \varepsilon_{t\alpha}^{-} \geq 0, \qquad \forall t \in T, \forall \alpha \in A,$$

$$-Mz_{p\alpha} \leq \beta_{p\alpha} \leq Mz_{p\alpha}, \qquad \forall \alpha \in A, \forall p \in P,$$
(22)

$$-Mz_{p\alpha} \le \beta_{p\alpha} \le Mz_{p\alpha}, \qquad \forall \alpha \in A, \forall p \in P, \tag{23}$$

$$\sum_{p=1}^{P} z_{p\alpha} \le K, \qquad \forall \alpha \in A, \tag{24}$$

$$z_{p\alpha} \in \{0,1\},$$
  $\forall \alpha \in A, \forall p \in P,$  (25)

$$\tilde{D}_{\rho\alpha'}^2 = \frac{\left(\frac{\beta_{\rho\alpha''} - \beta_{\rho\alpha'}}{\alpha'' - \alpha'}\right) - \left(\frac{\beta_{\rho\alpha'} - \beta_{\rho\alpha}}{\alpha' - \alpha}\right)}{\alpha'' - 2\alpha' + \alpha}$$
(26)

$$D2_{p\alpha'} > \tilde{D}^2_{p\alpha'}$$
  $\forall \alpha' \in A', \forall p \in P$  (27)

$$D2_{p\alpha'} > -\tilde{D}^2_{p\alpha'}$$
  $\forall \alpha' \in A', \forall p \in P$  (28)

$$\beta_{0\alpha} + \beta_{\alpha}^T x_t \leq \beta_{0\alpha'} + \beta_{\alpha'}^T x_t, \qquad \forall t \in T, \forall (\alpha, \alpha') \in A \times A, \alpha < \alpha',$$

(29)

#### Variable Selection via LASSO

- ▶ Another way of doing regularization is including the coefficients  $\ell_1$ -norm on the objective function
- In this method, coefficients are shrunk towards zero by changing a continuous parameter  $\lambda$ , which penalizes the size of the  $\ell_1$ -norm.
- ▶ When the value of  $\lambda$  gets bigger, fewer variables are selected to be used.
- ► The optimization problem for a single quantile is presented below:

$$\min_{\beta_0,\beta} \sum_{t \in T} \alpha |y_t - q(x_t)|^+ + \sum_{t \in T} (1 - \alpha) |y_t - q(x_t)|^- + \lambda \|\beta\|_1,$$

$$q(x_t) = \beta_0 - \sum_{p=1}^P \beta_p x_{t,p}.$$

#### Variable Selection via LASSO

At first, we select variables using LASSO

$$\underset{\beta_{0},\beta,\varepsilon_{t\alpha}^{+},\varepsilon_{t\alpha}^{-}}{\operatorname{arg\,min}} \sum_{\alpha\in A} \sum_{t\in T} \left(\alpha\varepsilon_{t\alpha}^{+} + (1-\alpha)\varepsilon_{t\alpha}^{-}\right) + \lambda \sum_{p=1}^{P} \xi_{p\alpha} + \gamma \sum_{\alpha\in A'} D2_{p\alpha} \tag{30}$$

s.t. 
$$\varepsilon_{t\alpha}^+ - \varepsilon_{t\alpha}^- = y_t - \beta_{0\alpha} - \sum_{p=1}^P \beta_{p\alpha} \tilde{x}_{t,p}, \quad \forall t \in T, \forall \alpha \in A,$$
 (31)

$$\varepsilon_{t\alpha}^{+}, \varepsilon_{t\alpha}^{-} \geq 0, \qquad \forall t \in T, \forall \alpha \in A, 
\xi_{p\alpha} \geq \beta_{p\alpha}, \qquad \forall p \in P, \forall \alpha \in A,$$
(32)

$$\xi_{p\alpha} \ge \beta_{p\alpha}, \qquad \forall p \in P, \forall \alpha \in A,$$
 (33)

$$\tilde{D}_{\rho\alpha'}^2 = \frac{\left(\frac{\beta_{\rho\alpha''} - \beta_{\rho\alpha'}}{\alpha'' - \alpha'}\right) - \left(\frac{\beta_{\rho\alpha'} - \beta_{\rho\alpha}}{\alpha' - \alpha}\right)}{\alpha'' - 2\alpha' + \alpha}$$
(34)

$$D2_{p\alpha'} > \tilde{D}^2_{p\alpha'}$$
  $\forall \alpha' \in A', \forall p \in P$  (35)

$$D2_{p\alpha'} > -\tilde{D}^2_{p\alpha'}$$
  $\forall \alpha' \in A', \forall p \in P$  (36)

$$\xi_{p\alpha} \ge -\beta_{p\alpha}, \quad \forall p \in P, \forall \alpha \in A.$$
 (37)

#### Variable Selection via LASSO

▶ We then define  $S_{\theta}$  (where  $\theta = \begin{bmatrix} \lambda & \gamma \end{bmatrix}^T$ ) as the set of indexes of selected variables given by

$$S_{\theta} = \{ p \in \{1, \dots, P\} | |\beta_{\theta, p}^{*LASSO}| \neq 0 \}.$$

Hence, we have that, for each  $p \in \{1, \dots, P\}$ ,

$$\beta_{\theta,p}^{*LASSO} = 0 \Longrightarrow \beta_{\theta,p}^{*} = 0.$$

 $\blacktriangleright$  On the second stage, we estimate coefficients using a regular QR where input variables are only the ones which belonging to  $\mathcal{S}_\lambda$ 

Estimation and Evaluation

## Estimation and Evaluation

#### **Evaluation Metrics**

▶ We use a performance measurement which emphasizes the correctness of each quantile. For each probability  $\alpha \in A$ , a loss function is defined by

$$L(\alpha) = \sum_{t \in T} \rho_{\alpha}(y_t - q_{\alpha}(x_t)).$$

The loss score  $\mathcal{L}$ , which is the chosen evaluation metric to optimize, aggregates the score function over all elements of A:

$$\mathcal{L} = \frac{1}{|A|} \sum_{\alpha \in A} L(\alpha).$$

#### Time-series Cross-Validation

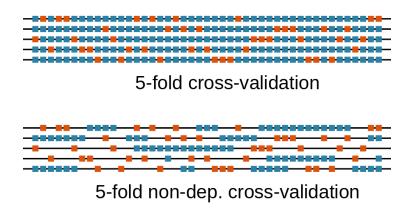


Figure 1:  $\mathcal{K}$ -fold CV and  $\mathcal{K}$ -fold with non-dependent data. Observations in blue are used to estimation and in orange for evaluation. Note that non-dependent data doesn't use all dataset in each fold.

#### Time-series Cross-Validation

▶ The CV score is given by the sum of the loss function for each fold. The optimum value of *t* in this criteria is the one that minimizes the CV score:

$$\theta^* = \operatorname{argmin}_{\theta} CV(\theta) = \sum_{k \in \mathcal{K}} \sum_{\alpha \in A} L(\alpha).$$

To optimize CV function in  $\theta$ , we use the Nelder-Mead algorithm, which is very efficient for searching in a two-dimmension parametric space.

Nonparametric model

# Nonparametric model

## Nonparametric model - Formulation

$$\begin{aligned} & \underset{q_{\alpha t}, \delta_{t}^{+}, \delta_{t}^{-}, \xi_{t}}{\min} & \sum_{\alpha \in A} \sum_{t \in T'} \left(\alpha \delta_{t \alpha}^{+} + (1 - \alpha) \delta_{t \alpha}^{-}\right) \\ & + \lambda_{1} \sum_{t \in T'} \gamma_{t \alpha} + \lambda_{2} \sum_{t \in T'} \xi_{t \alpha} \\ & s.t. & \delta_{t}^{+} - \delta_{t \alpha}^{-} = y_{t} - q_{t \alpha}, & \forall t \in T', \forall \alpha \in A, \\ & D_{t \alpha}^{1} = \frac{q_{\alpha t + 1} - q_{\alpha t}}{x_{t + 1} - x_{t}}, & \forall t \in T', \forall \alpha \in A, \\ & D_{t \alpha}^{2} = \frac{\left(\frac{q_{\alpha t + 1} - q_{\alpha t}}{x_{t + 1} - x_{t}}\right) - \left(\frac{q_{\alpha t} - q_{\alpha t - 1}}{x_{t} - x_{t - 1}}\right)}{x_{t + 1} - 2x_{t} + x_{t - 1}}. & \forall t \in T', \forall \alpha \in A, \\ & \gamma_{t \alpha} \geq D_{t \alpha}^{1}, & \forall t \in T', \forall \alpha \in A, \\ & \gamma_{t \alpha} \geq D_{t \alpha}^{1}, & \forall t \in T', \forall \alpha \in A, \\ & \xi_{t \alpha} \geq D_{t \alpha}^{2}, & \forall t \in T', \forall \alpha \in A, \\ & \xi_{t \alpha} \geq D_{t \alpha}^{2}, & \forall t \in T', \forall \alpha \in A, \\ & \xi_{t \alpha} \geq D_{t \alpha}^{2}, & \forall t \in T', \forall \alpha \in A, \\ & \delta_{t \alpha}^{+}, \delta_{t \alpha}^{-}, \gamma_{t \alpha}, \xi_{t \alpha} \geq 0, & \forall t \in T', \forall \alpha \in A, \\ & q_{t \alpha} \leq q_{t \alpha'}, & \forall t \in T', \forall \alpha \in A, \end{aligned}$$

## Nonparametric vs. Linear Model

► The nonparametric approach is more flexible to capture heteroscedasticity.

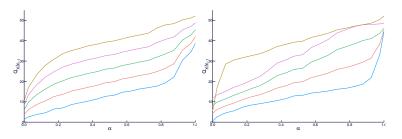


Figure 2: Estimated quantile functions, for different values of  $y_{t-1}$ . On the left using a linear model and using a nonparametric approach on the right.

# Nonparametric vs. Linear Model

► This flexibility might lead to overfitting, if we don't select a proper penalty, as shown below:

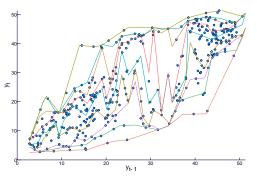


Figure 3: Example of a overfitted quantile function