Name:

M339J: Probability models

University of Texas at Austin

Solution: Practice Problems for In-Term One

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Notes: This is a closed book and closed notes exam. The maximal score on this exam is ??

points.

Time: 50 minutes

1.1. TRUE/FALSE QUESTIONS. Please, note your final answer on the front page of this exam.

Problem 1.1. Let X denote the outcome of a roll of a fair, regular icosahedron (a polyhedron with 20 faces) with numbers $1, 2, \dots, 20$ written on its sides. Then $\mathbb{E}[X] = 15/2$. True or false? Why?

Solution: FALSE

Straight from the definition of expectation, we have that

$$\mathbb{E}[X] = \frac{1}{20}(1+2+\cdots+20) = \frac{1}{20} \cdot \frac{20\cdot 21}{2} = \frac{21}{2}.$$

Problem 1.2. (2 pts) Let X be an exponential random variable. Then, its mean and its standard deviation are equal. *True or false?*

Solution: TRUE

Let $X \sim Exponential(\theta)$. From our tables, we get

$$\mathbb{E}[X] = \theta,$$

$$Var[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \theta^2 \cdot 2! - \theta^2 = \theta^2.$$

Since the standard deviation is the square root of the variance, we are done!

Problem 1.3. (2 points) For a random variable X and for a positive constant d, in our usual notation, we have

(1.1)
$$\mathbb{E}[X] = e_X(d)S_X(d) + \mathbb{E}[X \wedge d].$$

True or false?

Solution: TRUE

1.2. **Free-response problems.** Please, explain carefully all your statements and assumptions. Numerical results or single-word answers without an explanation (even if they're correct) are worth 0 points.

Problem 1.4. Let the random variable X have a two-parameter Pareto distribution with $\alpha = 2$ and $\theta = 10$. What is the 75th percentile of this distribution?

Solution: Let F_X denote the cumulative distribution function of X. We need to solve for x in

$$F_X(x) = \frac{3}{4} \, .$$

From the STAM tables, we learn that

$$F_X(x) = 1 - \left(\frac{\theta}{x+\theta}\right)^{\alpha}.$$

So, we solve for x:

$$\frac{3}{4} = 1 - \left(\frac{\theta}{x+\theta}\right)^{\alpha} \quad \Leftrightarrow \quad \frac{1}{4} = \left(\frac{10}{x+10}\right)^{2} \quad \Leftrightarrow \quad \frac{1}{2} = \frac{10}{x+10} \quad \Leftrightarrow \quad x+10 = 20 \quad \Leftrightarrow \quad x = 10$$

Problem 1.5. (10 points) A population of insureds consists of three types of people: α , β and γ . There is an equal number of Type α and Type β people in the population. The number of Type γ people is equal to the total number of the remaining two types of people. The probability that a Type α person makes at least one claim in a year is 1/5. The probability that a Type β person makes at least one claim in a year is 2/5. The probability that a Type γ person makes at least one claim in a year is 3/5.

At the end of the year, a person is chosen at random from the population. It is observed that that person had at least one claim. What is the probability that the person was of Type β ?

Solution: From the given breakdown of the population, we conclude that

(1.2)
$$\mathbb{P}[\alpha] = \mathbb{P}[\beta] = 1/4, \quad \mathbb{P}[\gamma] = 1/2.$$

Let E denote the event that there was at least one claim. By Bayes' Theorem, we have that

(1.3)
$$\mathbb{P}[\beta \mid E] = \frac{\mathbb{P}[E \mid \beta] \times \mathbb{P}[\beta]}{\mathbb{P}[E \mid \alpha] \times \mathbb{P}[\alpha] + \mathbb{P}[E \mid \beta] \times \mathbb{P}[\beta] + \mathbb{P}[E \mid \gamma] \times \mathbb{P}[\gamma]}$$
$$= \frac{(2/5)(1/4)}{(1/5)(1/4) + (2/5)(1/4) + (3/5)(1/2)} = \frac{2}{9} .$$

Problem 1.6. (15 points) Losses X follow a Pareto distribution with parameters $\alpha > 1$ and θ unspecified. For a positive constant c, determine the ratio of the mean excess loss function evaluated at $c\theta$ to the mean excess loss function evaluated at θ .

Solution: By definition, the *mean excess loss function* of the random variable X at a positive constant d such that $\mathbb{P}[X > d] > 0$ is given by

$$e_X(d) = \mathbb{E}[X - d \mid X > d].$$

According to our class notes, we also have that

(1.5)
$$e_X(d) = \frac{\mathbb{E}[X] - \mathbb{E}[X \wedge d]}{S_X(d)}.$$

Using the STAM tables for the Pareto distribution, we get that in the present problem

(1.6)
$$e_X(d) = \frac{\frac{\theta}{\alpha - 1} - \frac{\theta}{\alpha - 1} \left[1 - \left(\frac{\theta}{d + \theta} \right)^{\alpha - 1} \right]}{\left(\frac{\theta}{d + \theta} \right)^{\alpha}} = \frac{\frac{\theta}{\alpha - 1}}{\frac{\theta}{d + \theta}} = \frac{d + \theta}{\alpha - 1}.$$

Finally, our answer is

(1.7)
$$\frac{\frac{c\theta+\theta}{\alpha-1}}{\frac{\theta+\theta}{\alpha-1}} = \frac{c+1}{2}.$$

Problem 1.7. Let $\{X_n, n \geq 1\}$ be a sequence of independent random variables. Assume that all the variables in the sequence have the two-parameter Pareto distribution with $\theta = 10$ and $\alpha = 3$. For each n, define the random variable

$$Y_n = \frac{X_1^2 + X_2^2 + \dots + X_n^2}{n} \, .$$

Does the limit of the sequence $\{Y_n, n \geq 1\}$ as $n \to \infty$ exist? If so, how much is it? If not, why not?

Solution: Since the random variables $\{X_n, n \geq 1\}$ are independent and identically distributed, the random variables $\{X_n^2, n \geq 1\}$ are also independent and identically distributed. If the mean of the random variable X_1^2 exists and is finite, we can apply the Law of Large Numbers. Using the STAM tables, we get that

$$\mathbb{E}[X_1^2] = \frac{2\theta^2}{(\alpha - 1)(\alpha - 2)} = \frac{200}{2} = 100.$$

We see that the expected value is finite. So, not only does the Law of Large Numbers apply, but invoking it we can conclude that the limit of Y_n is 100.

Problem 1.8. (10 points) Let $X \sim Pareto(\alpha = 3, \theta = 3000)$. Assume that there is a deductible of d = 5000. Find the loss elimination ratio.

Solution: Using the tables, we get

$$\mathbb{E}[X] = \frac{\theta}{\alpha - 1} = \frac{3000}{3 - 1} = 1500,$$

$$\mathbb{E}[X \wedge d] = \frac{\theta}{\alpha - 1} \left[1 - \left(\frac{\theta}{d + \theta}\right)^{\alpha - 1} \right] = \frac{3000}{3 - 1} \left[1 - \left(\frac{3000}{5000 + 3000}\right)^{3 - 1} \right] = 1500[1 - \left(\frac{3}{8}\right)^{2}].$$

Finally, the loss elimination ratio is

$$[1 - (\frac{3}{8})^2] = \frac{55}{64}.$$

Problem 1.9. (10 points) Let the ground-up loss X be exponentially distributed with mean \$800.

An insurance policy has an ordinary deductible of \$100 and the maximum amount payable per loss of \$2500.

Find the expected value of the amount paid (by the insurance company) per positive payment.

Solution: We are given $X \sim Exponential(\theta = 800)$, the deductible d = 100 and the policy limit u - d = 2500. We need to calculate $\mathbb{E}[Y^P]$ where $Y^P = Y^L \mid Y^L > 0$ and

$$Y^{L} = \begin{cases} (X - d)_{+}, & X < u, \\ u - d, & X \ge u \end{cases}$$
$$= (X \wedge u - d)_{+}.$$

By the memoryless property of the exponential distribution, we have that

$$Y = X - d \mid X > d$$

is also exponential with mean 800. So, using our tables, we get

$$\mathbb{E}[Y^P] = \mathbb{E}[Y \land (u - d)] = \mathbb{E}[Y \land 2500] = 800(1 - e^{-2500/800}) \approx 764.85.$$

Problem 1.10. (10 points) Losses in year y follow a two-parameter Pareto distribution with parameters $\alpha = 4$ and $\theta = 10$.

Losses in year y + 1 are uniformly 10% higher than those in year y.

An insurance covers each loss subject to a deductible d = 20. Calculate the **loss elimination** ratio for year y + 1.

Solution: Pareto is a scale distribution with the scale parameter θ . So, the losses in year y+1, denoted by X again have the two-parameter Pareto distribution with $\alpha=4$ and $\theta=11$. By definition, the loss elimination ratio is

$$\frac{\mathbb{E}[X \wedge d]}{\mathbb{E}[X]} .$$

In our case, using the provided tables, we get

$$\frac{\mathbb{E}[X \wedge d]}{\mathbb{E}[X]} = \frac{\frac{\theta}{\alpha - 1} (1 - \left(\frac{\theta}{d + \theta}\right)^{\alpha - 1})}{\frac{\theta}{\alpha - 1}} = 1 - \left(\frac{\theta}{d + \theta}\right)^{\alpha - 1} = 1 - \left(\frac{11}{20 + 11}\right)^{4 - 1} = 1 - \left(\frac{11}{31}\right)^3 \approx 0.9553.$$

Problem 1.11. (10 points) Assume that the severity random variable X is uniform on the interval (0, 1000). There is an insurance policy to cover this loss. The insurance policy has a deductible of 200 per loss and the maximum amount payable by the insurer per loss of 700.

Calculate the expected value of the per payment random variable Y^P under this policy.

Solution: Note that

$$X - 200 \mid X > 200 \sim U(0, 800).$$

So,

$$(X - 200) \wedge 700 \mid X > 200 \sim$$

$$\begin{cases} U(0, 700) & \text{with probability } 7/8 \\ 700 & \text{with probability } 1/8 \end{cases}$$

$$\mathbb{E}[Y^P] = \frac{7}{8} \cdot 350 + \frac{1}{8} \cdot 700 = 393.75.$$

Problem 1.12. Assume that the severity random variable X is exponentially distributed with mean 1400. There is an insurance policy to cover this loss. This insurance policy has

- (1) a deductible of 200 per loss,
- (2) the coinsurance factor of $\alpha = 0.25$, and
- (3) the maximum amount payable by the insurer per loss of 700.

Calculate the expected value of the per payment random variable Y^P under this policy.

Solution: Due to the memoryless property of the exponential distribution, we have

$$X - 200 \mid X > 200 \sim Exponential(\theta = 1400).$$

Due to the fact that the exponential distribution is a scale distribution, when we introduce the coinsurance factor, we get

$$0.25(X - 200) \mid X > 200 \sim Exponential(\theta^* = 0.25 * 1400 = 350).$$

Hence, using our tables with $Y \sim Exponential(\theta = 350)$,

$$\mathbb{E}[Y^P] = \mathbb{E}[Y \land 700] = 350(1 - e^{-700/350}) \approx 302.63.$$

1.3. MULTIPLE CHOICE QUESTIONS.

Problem 1.13. (5 pts) Let X be exponential with variance 225. Let $a = \mathbb{E}[|20 - X|]$. Then,

- (a) $0 \le a < 50$
- (b) $50 \le a < 150$
- (c) $150 \le a < 325$
- (d) $325 \le a < 550$
- (e) None of the above.

Solution: (a)

From the given information, $X \sim Exponential(\theta = 15)$. So,

$$\begin{split} a &= \mathbb{E}[|20 - X|] = \int_0^\infty |20 - x| f_X(x) \, dx \\ &= \int_0^{20} (20 - x) f_X(x) \, dx + \int_{20}^\infty (-20 + x) f_X(x) \, dx \\ &= \int_0^\infty (20 - x) f_X(x) \, dx + 2 \int_{20}^\infty (x - 20) f_X(x) \, dx \\ &= 20 \int_0^\infty f_X(x) \, dx - \int_0^\infty x \, f_X(x) \, dx + 2 \int_0^\infty y f_X(y + 20) \, dy \\ &= 20 - \theta + 2e^{-20/\theta} \int_0^\infty y f_X(y) \, dy \\ &= 20 - \theta + 2e^{-20/\theta} \theta \\ &= 20 - 15 + 30e^{-20/15} \approx 12.9079. \end{split}$$

Problem 1.14. Let E and F be two events on the same probability space. You know that $\mathbb{P}[E \cup F] = 0.75$ and $\mathbb{P}[E \cup F^c] = 0.85$. What is the probability of the event E?

- (a) 0.5
- (b) 0.6
- (c) 0.65
- (d) 0.7
- (e) None of the above.

Solution: (b)

By the basic properties of probability, we know that

$$0.75 = \mathbb{P}[E \cup F] = \mathbb{P}[E] + \mathbb{P}[F] - \mathbb{P}[E \cap F],$$

$$0.85 = \mathbb{P}[E \cup F^c] = \mathbb{P}[E] + \mathbb{P}[F^c] - \mathbb{P}[E \cap F^c].$$

We can sup up the above two equalities to get

$$1.6 = 2\mathbb{P}[E] + (\mathbb{P}[F] + \mathbb{P}[F^c]) - (\mathbb{P}[E \cap F] + \mathbb{P}[E \cap F^c]) = 2\mathbb{P}[E] + 1 - \mathbb{P}[E] = \mathbb{P}[E] + 1.$$

Finally, $\mathbb{P}[E] = 0.6$.

Problem 1.15. The time until the next bus arrives is a continuous random variable T with the density

$$f_T(t) = \begin{cases} \kappa(10 - t) & 0 < t < 10\\ 0 & \text{otherwise} \end{cases}$$

for some constant κ . Given that you have already waited for 4 minutes, what is the probability that you will wait for at least another 4 minutes?

- (a) 1/25
- (b) 1/9
- (c) 1/8
- (d) 1/3
- (e) None of the above.

Solution: (b)

For any t between 0 and 10, the survival function of the random variable T is

$$S_T(t) = \frac{\kappa}{2} (10 - t)^2$$

So, the conditional probability in the problem is

$$\mathbb{P}[T > 8 \mid T > 4] = \frac{\mathbb{P}[T > 8, T > 4]}{\mathbb{P}[T > 4]} = \frac{\mathbb{P}[T > 8]}{\mathbb{P}[T > 4]} = \frac{\frac{\kappa}{2}(10 - 8)^2}{\frac{\kappa}{2}(10 - 4)^2} = \frac{4}{36} = \frac{1}{9}.$$

Problem 1.16. Let X_1 , X_2 , and X_3 be independent, identically distributed random variables with the probability mass function

$$p_X(x) = \begin{cases} \frac{1}{4} & x = 0\\ \frac{1}{2} & x = 1\\ \frac{1}{4} & x = 2 \end{cases}$$

Find $\mathbb{P}[X_1X_2X_3=0]$.

- (a) 27/64
- (b) 1/8
- (c) 31/64
- (d) 37/64
- (e) None of the above.

Solution: (d)

The product $X_1X_2X_3$ equals zero if at least one of the random variables X_1, X_2 and X_3 equals zero. So,

$$\mathbb{P}[X_1 X_2 X_3 = 0] = 1 - \mathbb{P}[X_1 \neq 0, X_2 \neq 0, X_3 \neq 0].$$

Since the random variables are independent, the above probability equals

$$1 - \mathbb{P}[X_1 \neq 0] \mathbb{P}[X_2 \neq 0] \mathbb{P}[X_3 \neq 0] = 1 - \left(\frac{3}{4}\right)^3 = 1 - \frac{27}{64} = \frac{37}{64}.$$

Problem 1.17. A recent study indicates that the annual cost of fertilizing a Japanese plum tree in Austin has a mean 100 with a variance of 20. A tax of 10% is introduced on fertilizer, i.e., fertilizer is made 10% more expensive. What is the variance of the new annual cost of fertilizing a Japanese plum tree in Austin after the tax is introduced?

- (a) 20
- (b) 22
- (c) 23.1
- (d) 24.2
- (e) None of the above.

Solution: (d)

Let X be the cost before the tax. Then, the cost after the tax equals 1.1X. So, the new variance is

$$Var[1.1X] = 1.21Var[X] = 1.21(20) = 24.2.$$

Problem 1.18. Let the random variable X have the cumulative distribution function given by

$$F_X(x) = \begin{cases} 0 & \text{for } x < 1\\ \frac{1}{2}(x^2 - 2x + 2) & \text{for } 1 \le x < 2\\ 1 & \text{for } 2 \le x \end{cases}$$

What is the expectation of X?

- (a) 2/3
- (b) 5/6
- (c) 7/6
- (d) 4/3
- (e) None of the above.

Solution: (d)

This is a mixed random variable with $\mathbb{P}[X=1] = 1/2$. The pdf is $f_X(x) = x - 1$ for 1 < x < 2. So,

$$\mathbb{E}[X] = \frac{1}{2}(1) + \int_{1}^{2} x(x-1) \, dx = \frac{1}{2} + \left[\frac{x^{3}}{3} - \frac{x^{2}}{2} \right]_{x=1}^{2}$$
$$= \frac{1}{2} + \frac{8}{3} - \frac{4}{2} - \frac{1}{3} + \frac{1}{2} = 4/3.$$

Problem 1.19. Let the independent random variables X and Y have the same mean. You are given that coefficient of variation of X equals 2 and the coefficient of variation of Y equals 4. What is the coefficient of variation of the average of X and Y?

- (a) 3/2
- (b) $\frac{\sqrt{13}}{2}$
- (c) 5/2
- (d) There is not enough information to answer this problem.
- (e) None of the above.

Solution: (b)

Let $\mu = \mathbb{E}[X] = \mathbb{E}[Y]$. Then, $\sigma_X = SD[X] = 2\mu$ and $\sigma_Y = SD[Y] = 3\mu$. The variance of the average of the two random variables is

$$Var\left[\frac{1}{2}(X+Y)\right] = \frac{1}{4}(Var[X] + Var[Y]).$$

In terms of μ , the variance of the average can be rewritten as

$$Var\left[\frac{1}{2}(X+Y)\right] = \frac{1}{4}(4\mu^2 + 9\mu^2) = \frac{13\mu^2}{4}.$$

So, the standard deviation of the average can be expressed as $\frac{\mu\sqrt{13}}{2}$. Hence, the coefficient of variation of the average equals $\frac{\sqrt{13}}{2}$.

Problem 1.20. (5 points) Let X be the ground-up loss random variable. Assume that X has the two-parameter Pareto distribution with $\theta = 4,000$ and $\alpha = 3$.

Let B denote the expected payment per loss on behalf of an insurer which wrote a policy with no deductible and with a policy limit of 5,000. Then,

- (a) $B \approx 1,000$
- (b) $B \approx 1,200$
- (c) $B \approx 1,400$
- (d) $B \approx 1,600$
- (e) None of the above

Solution: (d)

Using our tables,

$$B = \mathbb{E}[X \wedge 5000] = \frac{\theta}{\alpha - 1} \left[1 - \left(\frac{\theta}{5000 + \theta}\right)^{\alpha - 1}\right] = \frac{4000}{2} \left[1 - \left(\frac{4000}{9000}\right)^{2}\right] \approx 1600.$$

Problem 1.21. (5 points) Let X be the ground-up loss random variable. Assume that X has the exponential distribution with mean 1,000.

Let B denote the expected payment per loss on behalf of an insurer who wrote a policy with a deductible of 1,500 and with the **maximum payment by the insurer** equal to 2,500. Then,

- (a) $B \approx 714$
- (b) $B \approx 816$
- (c) $B \approx 918$

- (d) $B \approx 1020$
- (e) None of the above

Solution: (e)

Using our tables, we get

$$B = \mathbb{E}[X \land 4000] - \mathbb{E}[X \land 1500] = 1000(1 - e^{-4000/1000}) - 1000(1 - e^{-1500/1000})$$
$$= 1000(e^{-1.5} - e^{-4}) \approx 204.8.$$

Problem 1.22. (5 points) The ground-up loss X is modeled by a two-parameter Pareto distribution with parameters $\alpha = 2$ and $\theta = 200$. For an insurance policy on the above loss, there is a <u>franchise</u> deductible of 200. Find the expected value of the *per payment* random variable.

- (a) 200
- (b) 400
- (c) 600
- (d) 800
- (e) None of the above.

Solution: (c)

In our usual notation, as we have shown in class,

$$\mathbb{E}[Y^P] = d + \frac{\mathbb{E}[X] - \mathbb{E}[X \wedge d]}{S_X(d)}.$$

In this problem, the ground-up loss is $X \sim Pareto(\alpha = 2, \theta = 200)$. So, using the STAM tables, we have

$$S_X(d) = \left(\frac{\theta}{d+\theta}\right)^{\alpha},$$

$$\mathbb{E}[X] = \frac{\theta^1 \cdot 1!}{\alpha - 1} = \frac{\theta}{\alpha - 1},$$

$$\mathbb{E}[X \wedge d] = \frac{\theta}{\alpha - 1} \left(1 - \left(\frac{\theta}{d+\theta}\right)^{\alpha - 1}\right).$$

So,

$$\mathbb{E}[Y^P] = d + \frac{\frac{\theta}{\alpha - 1} - \frac{\theta}{\alpha - 1} \left(1 - \left(\frac{\theta}{d + \theta}\right)^{\alpha - 1}\right)}{\left(\frac{\theta}{d + \theta}\right)^{\alpha}} = d + \frac{\frac{\theta}{\alpha - 1} \left(\frac{\theta}{d + \theta}\right)^{\alpha - 1}}{\left(\frac{\theta}{d + \theta}\right)^{\alpha}} = d + \frac{\frac{\theta}{\alpha - 1}}{\frac{\theta}{d + \theta}} = d + \frac{d + \theta}{\alpha - 1}$$

Using the parameter values and the deductible from this problem, we get

$$\mathbb{E}[Y^P] = 200 + \frac{200 + 200}{2 - 1} = 600.$$