

Problem from Sample Exam.

Consider two **independent** r.v.s  $X$  and  $Y$  w/ cdfs  $F_X$  and  $F_Y$ , resp.

Define:

$$U := \max(X, Y)$$

and

$$V := \min(X, Y)$$

Express  $F_U$  and  $F_V$  in terms of  $F_X$  and  $F_Y$ .

→: By def'n: for all  $x \in \mathbb{R}$

$$\begin{aligned} F_U(x) &= \mathbb{P}[U \leq x] \\ &= \mathbb{P}[\max(X, Y) \leq x] = \\ &= \mathbb{P}[X \leq x, Y \leq x] = \\ &= \underbrace{\mathbb{P}[X \leq x]}_{F_X(x)} \cdot \underbrace{\mathbb{P}[Y \leq x]}_{F_Y(x)} \quad \text{independence} \\ &= F_X(x) \cdot F_Y(x) \end{aligned}$$

By def'n: for all  $x \in \mathbb{R}$

$$\begin{aligned} F_V(x) &= \mathbb{P}[V \leq x] \\ &= \mathbb{P}[\min(X, Y) \leq x] \quad \mathbb{P}[E] = 1 - \mathbb{P}[E^c] \\ &= 1 - \mathbb{P}[\min(X, Y) > x] = \\ &= 1 - \mathbb{P}[X > x, Y > x] = \\ &= 1 - \underbrace{\mathbb{P}[X > x]}_{1 - F_X(x)} \cdot \underbrace{\mathbb{P}[Y > x]}_{1 - F_Y(x)} \quad \text{independence} \\ &= 1 - (1 - F_X(x))(1 - F_Y(x)) \\ &= 1 - 1 + F_X(x) + F_Y(x) - F_X(x)F_Y(x) \\ &= F_X(x) + F_Y(x) - F_X(x) \cdot F_Y(x) \end{aligned}$$

□

## More about the Normal Dist'n.

Recall:  $Z \sim N(0,1)$

Any normal  $X \sim N(\mu, \sigma^2)$  can be expressed as

$$X = \mu + \sigma \cdot Z \quad \text{for some } Z \sim N(0,1)$$

$$\bullet \mathbb{E}[X] = \mathbb{E}[\mu + \sigma \cdot Z] = \mu + \sigma \cdot \underbrace{\mathbb{E}[Z]}_{=0} = \mu$$

↑  
linearity

$$\bullet \text{Var}[X] = \text{Var}[\mu + \sigma \cdot Z] = \sigma^2 \cdot \underbrace{\text{Var}[Z]}_{=1} = \sigma^2$$

↑  
shift

Problem. Let  $X$  be normally distributed w/ mean  $\mu$  and std dev  $\sigma$ .  
What is the probability that  $X$  falls within  $\frac{1}{10}$  of the std dev from its mean?

$$\rightarrow: X \sim \text{Normal}(\text{mean} = \mu, \text{sd} = \sigma)$$

$$\mathbb{P}\left[|X - \mu| < \frac{1}{10}\sigma\right] = ?$$

$$\mathbb{P}\left[-\frac{1}{10}\sigma < X - \mu < \frac{1}{10}\sigma\right] =$$

$$= \mathbb{P}\left[-0.1 < \frac{X - \mu}{\sigma} < 0.1\right] =$$

$\frac{X - \mu}{\sigma} \sim N(0,1) \sim Z$

$$= \mathbb{P}[-0.1 < Z < 0.1] =$$

$$= \Phi(0.1) - \underbrace{\Phi(-0.1)}_{1 - \Phi(0.1)} = 2\Phi(0.1) - 1 = 2 \cdot 0.5398 - 1$$

$$= 1.0796 - 1$$

$$= 0.0796$$

□

# The Central Limit Theorem.

Let  $\{X_1, X_2, \dots\}$  be a sequence of i.i.d. r.v.s w/  $\mu_X = \mathbb{E}[X_i] < \infty$  and  $\sigma_X^2 = \text{Var}[X_i] < \infty$ .

Define for all  $n \in \mathbb{N}$ :

$$S_n = X_1 + X_2 + \dots + X_n$$

$$\frac{S_n - n \cdot \mu_X}{\sigma_X \sqrt{n}} \xrightarrow[n \rightarrow \infty]{\mathcal{D}} N(0, 1)$$

Problem. Automobile losses are independent and uniformly dist'l'd between 0 and 15,000. Calculate the (approximate) probability that the total amount for 200 losses lies between 1,450,000 and 1,700,000.

$$\rightarrow: X_i \sim U(0, 15000) \quad i = 1, \dots, 200$$

$$S_{200} = X_1 + X_2 + \dots + X_{200}$$

$$\text{We need } P[1.45 \cdot 10^6 < S_{200} < 1.7 \cdot 10^6] = ?$$

$$\mu_X = \mathbb{E}[X_i] = 7500$$

$$\text{Var}[X_i] = \frac{(15000)^2}{12} = 18750000$$

$$\sigma_X = 4330.127$$

In general:  $X \sim U(a, b)$

$$\mathbb{E}[X] = \frac{a+b}{2}$$

$$\text{Var}[X] = \frac{(b-a)^2}{12}$$

$$P\left[\frac{1.45 \cdot 10^6 - 7500 \cdot 200}{4330.127 \cdot \sqrt{200}} < Z < \frac{1.7 \cdot 10^6 - 7500 \cdot 200}{4330.127 \cdot \sqrt{200}}\right] =$$

$$\approx P[-0.82 < Z < 3.27] = \Phi(3.27) - \Phi(-0.82)$$

$$= 0.9995 - 0.2061 = 0.7934$$

□

## Section 4.2.

### Exponential Distribution.

A random variable  $T$  has the exponential distribution with rate  $\lambda$ , written as

$$T \sim \text{Exp}(\lambda)$$

If  $T$  has the pdf given by

$$f_T(t) = \lambda e^{-\lambda t} \quad \text{for } t \geq 0$$

Q: What is the cdf of  $T$ ?

$$\rightarrow: F_T(t) = \int_0^t f_T(u) du = \int_0^t \lambda e^{-\lambda u} du = \lambda \cdot \left(-\frac{1}{\lambda}\right) e^{-\lambda u} \Big|_{u=0}^t$$

$$F_T(t) = 1 - e^{-\lambda t}$$

Def'n. The survival function of any random variable  $X$  is  
 $S_X: \mathbb{R} \rightarrow [0,1]$   
given by

$$S_X(x) = 1 - F_X(x) \quad \text{for all } x \in \mathbb{R}$$

Note: For  $T \sim \text{Exp}(\lambda)$ :

$$S_T(t) = \mathbb{P}[T > t] = e^{-\lambda t}$$

Without proof:

$$\mathbb{E}[T] = SD[T] = \frac{1}{\lambda}$$