M378K Introduction to Mathematical Statistics

Homework assignment #5

Please, provide your **complete solutions** to the following problems. Final answers only, even if correct will earn zero points for those problems.

Problem 5.1. (10 points) Let X be a continuous random variable with the cumulative distribution function denoted by F_X and the probability density function denoted by f_X .

Express the cumulative distribution function and the density of the random variable $\tilde{X} = X^2$ in terms of F_X and f_X .

Solution: The range/support of \tilde{X} is within $[0,\infty)$. For every $x\geq 0$, the cumulative distribution function is

$$\begin{split} F_{\tilde{X}}(x) &= \mathbb{P}[\tilde{X} \leq x] = \mathbb{P}[X^2 \leq x] = \mathbb{P}[X \leq \sqrt{x}] - \mathbb{P}[X < -\sqrt{x}] \\ &= \mathbb{P}[X \leq \sqrt{x}] - \mathbb{P}[X \leq -\sqrt{x}] = F_X(\sqrt{x}) - F_X(-\sqrt{x}). \end{split}$$

As for the probability density function, we have that for all x > 0,

$$f_{\tilde{X}}(x) = F'_{\tilde{X}}(x) = \frac{1}{2\sqrt{x}}(f_X(\sqrt{x}) + f_X(-\sqrt{x})).$$

Problem 5.2. (10 points) Let Y be lognormal with parameters $\mu=1$ and $\sigma=2$, i.e., let $Y\stackrel{(d)}{=}e^X$ with $X\sim N(\mu,\sigma)$.

Define $\tilde{Y} = 3Y$.

Find the median of \tilde{Y} , i.e., find the value m such that $\mathbb{P}[\tilde{Y} \leq m] = 1/2$.

Solution: It can be shown that \tilde{Y} is lognormal with parameters $\mu^* = \mu + \ln(3)$ and $\sigma^* = \sigma$. So, \tilde{Y} can be written as $\tilde{Y} = e^{\tilde{X}}$ where $\tilde{X} \sim N(\mu^*, (\sigma^*)^2)$. Hence, with m denoting the median of \tilde{Y} , we have

$$\begin{aligned} 1/2 &= \mathbb{P}[\tilde{Y} \leq m] \\ &= \mathbb{P}[e^{\tilde{X}} \leq m] \\ &= \mathbb{P}[\tilde{X} \leq \ln(m)]. \end{aligned}$$

Since \tilde{X} is normal with mean μ^* (and the mean and the median of a normal r.v. are one and the same), we conclude that

$$ln(m) = 1 + ln(3) \Rightarrow m = 3e \approx 8.15.$$

Problem 5.3. (10 points) Let T denote the time for a call center employee to respond to any single telephone call. We model the random variable T by uniform distribution on the interval (48,72) with the time being measured in seconds. Let R denote the **rate** at which the call center employee responds to queries expressed in the number of customers per minute.

Does the random variable R have a density? If so, find the density of R.

Solution:

(i)

$$f_T(t) = \frac{1}{24} \mathbb{I}_{(48,72)} = \begin{cases} \frac{1}{24} & \text{for } 48 < x < 72 \\ 0 & \text{otherwise.} \end{cases}$$

(ii)

$$R = 60/T$$
.

(iii)

$$(60/72, 60/48) = (5/6, 5/4)$$

(iv) Method of transformations: For 5/6 < r < 5/4,

$$f_R(r) = \frac{1}{24} \times \frac{60}{r^2} = \frac{5}{2r^2} \,.$$

Problem 5.4. (30 points) Let X, Y and Z be independent and uniformly distributed on (0,1). Find the density function of W = X + Y + Z.

Solution:

Method I: The cdf-method. Let us introduce the random variable T = X + Y. We already showed in class that its density function looks like this

$$f_T(a) = \begin{cases} a & \text{for } a \in [0, 1] \\ 2 - a & \text{for } a \in [1, 2] \\ 0 & \text{otherwise.} \end{cases}$$

Next, we note that W = T + Z. Let's look at its cdf.

$$F_{W}(w) = \mathbb{P}[W \le w] = \mathbb{P}[T + Z \le w] = \int \int f_{T,Z}(t,z) \mathbf{1}_{\{0 \le t + z \le w\}} dt dz$$
$$= \int \int f_{T,Z}(t,z) \mathbf{1}_{\{0 \le t \le w - z\}} dt dz$$
$$= \int_{0}^{w} \int_{0}^{w-z} f_{T,Z}(t,z) dt dz.$$

Since T and Z are independent, we know that

$$f_{T,Z}(t,z) = f_T(t)f_Z(z)$$

So, using the expression we obtained in class and the density of a unit uniform, we get

$$F_W(w) = \int_0^w \int_0^{w-z} f_T(t) f_Z(z) \, dt \, dz$$

Thus, since $Z \sim U(0,1)$, we get that

$$F_W(w) = \int_0^{1 \wedge w} \int_0^{w-z} f_T(t) dt dz$$

For $w \in [0,1]$, we know that w-z within the bounds of integration above is also in [0,1]. So,

$$F_W(w) = \int_0^w \int_0^{w-z} t \, dt \, dz = \int_0^w \frac{(w-z)^2}{2} \, dz = \frac{w^3}{6} \, .$$

As an alternative, we could have used a **geometric** argument in 3D similar to the one we used in class in 2D.

For $w \in [1, 2]$, we have that

$$F_W(w) = \int_0^1 \int_0^{w-z} f_T(t) dt dz$$

$$= \int_0^{w-1} \int_0^{w-z} f_T(t) dt dz + \int_{w-1}^1 \int_0^{w-z} f_T(t) dt dz$$
(5.1) mac

Let's focus on the first integral in the sum above. Due to the fact that f_T is piecewise defined, it's prudent to split the inner integral as follows

$$\int_0^{w-z} f_T(t) dt = \int_0^1 f_T(t) dt + \int_1^{w-z} f_T(t) dt$$

$$= \int_0^1 t dt + \int_1^{w-z} (2-t) dt$$

$$= \frac{1}{2} + \int_{2-w+z}^1 u du$$

$$= \frac{1}{2} + \frac{1}{2} (1 - (2-w+z)^2)$$

$$= \frac{1}{2} (2 - (4+w^2+z^2-4w+4z-2wz))$$

$$= \frac{1}{2} (-2-w^2-z^2+4w-4z+2wz).$$

Now, the entire first integral in (5.1) reads as

$$\int_0^{w-1} \int_0^{w-z} f_T(t) dt dz = \frac{1}{2} \int_0^{w-1} \left(-2 - w^2 - z^2 + 4w - 4z + 2wz \right) dz$$

$$= \frac{1}{2} \left(-2(w-1) - w^2(w-1) - \frac{(w-1)^3}{3} + 4w(w-1) - 4\left(\frac{(w-1)^2}{2}\right) + 2w\left(\frac{(w-1)^2}{2}\right) \right)$$

$$= \frac{1}{2} \left(-2w + 2 - w^3 + w^2 - \frac{1}{3}(w^3 - 3w^2 + 3w - 1) + 4w^2 - 4w - 2(w^2 - 2w + 1) + w(w^2 - 2w + 1) \right)$$

$$= \frac{1}{2} \left(-\frac{1}{3}w^3 + 2w^2 - 2w + \frac{1}{3} \right).$$

The second integral in (5.1) is much easier. We get

$$\int_{w-1}^{1} \int_{0}^{w-z} f_{T}(t) dt dz = \int_{w-1}^{1} \int_{0}^{w-z} t dt dz$$

$$= \frac{1}{2} \int_{w-1}^{1} (w-z)^{2} dz$$

$$= \frac{1}{2} \int_{w-1}^{1} u^{2} du$$

$$= \frac{1}{6} (1 - (w-1)^{3})$$

$$= \frac{1}{6} (1 - (w^{3} - 3w^{2} + 3w - 1))$$

$$= \frac{1}{6} (2 - w^{3} + 3w^{2} - 3w).$$

Combining the two parts of (5.1), we obtain

$$F_W(w) = -\frac{1}{3}w^3 + \frac{3}{2}w^2 - \frac{3}{2}w + \frac{5}{6}$$

Finally, for $w \in [2, 3]$, we have that by symmetry with the first case

$$F_W(w) = \frac{(3-w)^3}{6}$$
.

Upon differentiation, we get that

$$f_W(w) = \begin{cases} \frac{w^2}{2} & \text{for } w \in [0, 1] \\ -w^2 + 3w - \frac{3}{2} & \text{for } w \in [1, 2] \\ \frac{(3-w)^2}{2} & \text{for } w \in [2, 3] \end{cases}$$

The density f_W is zero otherwise.

Method II: Convolution. Let us introduce the random variable T=X+Y, and find its density function. For every $a \in \mathbb{R}$

$$f_T(a) = \int_{-\infty}^{+\infty} f_X(y) f_Y(a - y) \, dy$$
$$= \int_{-\infty}^{+\infty} \mathbf{1}_{[0,1]}(y) \mathbf{1}_{[0,1]}(a - y) \, dy.$$

Now, we can conclude that $f_T(a) = 0$ for every $a \notin [0, 2]$. For all $a \in [0, 1]$, we have

$$f_T(a) = \int_0^a dy = a.$$

For all $a \in [1, 2]$, we have

$$f_T(a) = \int_{a-1}^1 dy = 1 - (a-1) = 2 - a.$$

Written all in one expression:

$$f_T(a) = \begin{cases} a & \text{for } a \in [0, 1] \\ 2 - a & \text{for } a \in [1, 2] \\ 0 & \text{otherwise.} \end{cases}$$

Now, we can express the random variable W as W=T+Z. Since X,Y and Z were assumed independent, we can conclude that T and Z are independent as well. So, we can write the density of W as

$$f_W(a) = \int_{-\infty}^{+\infty} f_T(y) f_Z(a - y) \, dy$$

for all $a \in [0, 3]$. It is evident that f_W vanishes outside of this interval. It is convenient to partition the above integral into regions as follows:

$$f_W(a) = \int_0^1 f_T(y) f_Z(a-y) \, dy + \int_1^2 f_T(y) f_Z(a-y) \, dy$$
$$= \int_0^1 y \mathbf{1}_{[0,1]}(a-y) \, dy + \int_1^2 (2-y) \mathbf{1}_{[0,1]}(a-y) \, dy.$$

This representation leads us to consider different cases of the values of a separately. For $a \in [0,1]$, we have

$$f_W(a) = \int_0^a y \, dy = \frac{1}{2} a^2 \, .$$

For $a \in [1, 2]$, we have

$$f_W(a) = \int_{a-1}^1 y \, dy + \int_1^a (2-y) \, dy$$
$$= \frac{1}{2} \left[(1 - (a-1)^2) - (2-a)^2 + 1 \right]$$
$$= \frac{1}{2} \left[2a - a^2 - 4 + 4a - a^2 + 1 \right]$$
$$= \frac{1}{2} \left[-2a^2 + 6a - 3 \right].$$

For $a \in [2, 3]$, we have

$$f_W(a) = \int_{a-1}^2 (2-y) \, dy = -\frac{1}{2} \left[0 - (3-a)^2 \right] = \frac{1}{2} \left(a - 3 \right)^2.$$