

Problem. Midterm and final.

Midterm and final scores in a large class have an (approximately) bivariate normal dist'n w/ parameters:

	<u>mean</u>	<u>sd</u>
<u>midterm</u>	65	18
<u>final</u>	60	20

correlation: 0.76

Q: What is the "estimated" mean final score of the students who were above the mean on the midterm?

→: Let U be the midterm score, and let V be the final score

Our task is to find:

$$\mathbb{E}[V \mid U > \mu_U] = ?$$

Let X and Z be U and V in standard units, resp.

Our ancillary task is to find:

$$\mathbb{E}[Z \mid X > 0] = ?$$

$$\mathbb{E}[Z \mid X > 0] = \frac{\mathbb{E}[Z \cdot \mathbb{I}_{[X > 0]}]}{\mathbb{P}[X > 0]}$$

by def'n.

$$\mathbb{E}[Z \mid X > 0] = \int_{-\infty}^{\infty} \mathbb{E}[Z \mid X=x] \cdot f_X(x \mid X > 0) dx$$

Law of Total Probability

$Z \mid X=x \sim \text{Normal}(px, 1-p^2)$

for $x > 0$:

$$\begin{aligned} f_X(x \mid X > 0) dx &= \mathbb{P}[X \in (x, x+dx) \mid X > 0] \\ &= \frac{\mathbb{P}[X \in (x, x+dx) \text{ and } X > 0]}{\mathbb{P}[X > 0]} \\ &= \frac{\mathbb{P}[X \in (x, x+dx)]}{\frac{1}{2}} = \frac{f_X(x) dx}{\frac{1}{2}} \\ &= 2f_X(x) = 2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[Z \mid X > 0] &= \int_0^{\infty} px \cdot 2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \frac{2p}{\sqrt{2\pi}} \int_0^{\infty} x e^{-\frac{x^2}{2}} dx = \left[\begin{array}{l} u = \frac{x^2}{2} \\ du = \frac{1}{2} \cdot 2 \cdot x dx \\ = x dx \end{array} \right] \\ &= \frac{2p}{\sqrt{2\pi}} \int_0^{\infty} e^{-u} du = \frac{2p}{\sqrt{2\pi}} \cdot 1 \end{aligned}$$

In this problem:

$$\frac{1.52}{\sqrt{2\pi}} = \underline{0.6063923}$$

$$\Rightarrow \text{Our final answer: } 60 + 20 \cdot 0.6063923 =$$

$$= \underline{72.12785}$$



Bivariate Normal in Matrix Notation.

Consider a bivariate normal pair (U, V) .

In 2D: we place the means into a vector

$$\mu := \begin{pmatrix} \mu_U \\ \mu_V \end{pmatrix} \quad \text{anything in } \mathbb{R}^2$$

The variances/covariances are placed into a matrix:

$$\Sigma := \begin{bmatrix} \sigma_U^2 & \sigma_U \sigma_V \rho \\ \sigma_U \sigma_V \rho & \sigma_V^2 \end{bmatrix} \quad (\text{positive definite})$$

Then, the joint density of (U, V) can be written as:

In 1D:

$$f_U(u) = \frac{1}{\sigma_U \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{u - \mu_U}{\sigma_U} \right)^2}$$

$$f_{U,V}(u, v) = \frac{1}{2\pi} \cdot \frac{1}{(\det(\Sigma))^{1/2}} \exp \left(-\frac{1}{2} \begin{pmatrix} u - \mu_U \\ v - \mu_V \end{pmatrix}_{1 \times 2}^T \begin{matrix} \Sigma^{-1} \\ 2 \times 2 \end{matrix} \begin{pmatrix} u - \mu_U \\ v - \mu_V \end{pmatrix}_{2 \times 1} \right)$$

Multivariate Normal Density.

Let $\mathbf{X} = (X_1, X_2, \dots, X_p)^T$ be

Normal (mean = $\mu = (\mu_1, \mu_2, \dots, \mu_p)^T$, $\Sigma = \begin{bmatrix} \sigma_1^2 & \text{Cov} & \text{Cov} & \text{Cov} \\ \text{Cov} & \ddots & \ddots & \ddots \\ \text{Cov} & \ddots & \ddots & \ddots \\ \text{Cov} & \ddots & \ddots & \sigma_n^2 \end{bmatrix}$)
w/ Σ positive definite.

Then,

$$f_{\mathbf{X}}(\underbrace{x_1, \dots, x_p}_{\mathbf{x}}) = \frac{1}{(2\pi)^{p/2}} \cdot \frac{1}{(\det(\Sigma))^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu)\right)$$

for all $\mathbf{x} \in \mathbb{R}^p$