

M339 J: April 2nd, 2021.

Poisson Thinning.

Thm. Let $N \sim \text{Poisson}(\lambda)$ be a counting random variable for some events of interest.

Suppose that independently of N , each event falls into a particular category indexed by $i = 1 \dots m$ w/ probability p_i ($i = 1 \dots m$)

Let N_i be the number of events of type i , for all $i = 1 \dots m$

Then: • $N_i \sim \text{Poisson}(\lambda_i = \underline{p_i \cdot \lambda})$ for all $i = 1 \dots m$

• N_1, N_2, \dots, N_m are independent random variables

Sample STAM

$$N \sim \text{Poisson}(\lambda = 12)$$

111. The number of accidents follows a Poisson distribution with mean 12. Each accident generates 1, 2, or 3 claimants with probabilities $1/2$, $1/3$, and $1/6$, respectively.

ASSUME INDEPENDENCE BETWEEN N and # of claimants.
Calculate the variance of the total number of claimants.

S .. the total # of claimants
 $\text{Var}[S] = ?$

(A) 20

(B) 25

(C) 30

(D) 35

(E) 40

for $i=1,2,3$
 N_i ... # of accidents w/ i claimants

$$N_i \sim \text{Poisson}(\lambda_i = \lambda \cdot p_i)$$

$$\Rightarrow \begin{aligned} N_1 &\sim \text{Poisson}(\lambda_1 = 6) \\ N_2 &\sim \text{Poisson}(\lambda_2 = 4) \\ N_3 &\sim \text{Poisson}(\lambda_3 = 2) \end{aligned}$$

112. In a clinic, physicians volunteer their time on a daily basis to provide care to those who are not eligible to obtain care otherwise. The number of physicians who volunteer in any day is uniformly distributed on the integers 1 through 5. The number of patients that can be served by a given physician has a Poisson distribution with mean 30.

Determine the probability that 120 or more patients can be served in a day at the clinic, using the normal approximation with continuity correction.

(A) $1 - \Phi(0.68)$

(B) $1 - \Phi(0.72)$

(C) $1 - \Phi(0.93)$

(D) $1 - \Phi(3.13)$

(E) $1 - \Phi(3.16)$

$$\Rightarrow S = 1 \cdot N_1 + 2 \cdot N_2 + 3 \cdot N_3$$

$$\text{Var}[S] = \text{Var}[N_1 + 2N_2 + 3N_3] = \boxed{N_1, N_2, N_3 \text{ independent by our thinning thm}}$$

$$= \text{Var}[N_1] + \text{Var}[2N_2] + \text{Var}[3N_3]$$

$$= \text{Var}[N_1] + 4 \text{Var}[N_2] + 9 \text{Var}[N_3]$$

$$= \lambda_1 + 4 \lambda_2 + 9 \lambda_3$$

$$= 6 + 4 \cdot 4 + 9 \cdot 2 = 40$$

Negative Binomial Dist'n.

Binomial Coefficients.

For $n, k \in \mathbb{N}_0$ w/ $n \geq k$:

$$\binom{n}{k} = \frac{n!}{k! (n-k)!}$$

We can generalize to $x \in \mathbb{R}_+$ and $k \in \mathbb{N}_0$.

If $x > k-1$, then

$$\binom{x}{k} = \frac{\Gamma(x+1)}{\Gamma(k+1) \cdot \Gamma(x-k+1)}$$

A Modelling Problem.

Imagine a sequence of Bernoulli trials such that:

- each trial results in success or failure;
- the probability of success in every trial is p ;
- the trials are independent

The experiment continues until a fixed number r of successes is achieved.

N ... random variable denoting the total # of failures before the r^{th} success.

Q: What is the pmf of N ?

→: for $k = 0, 1, 2, \dots$

$$p_N(k) := \mathbb{P}[N=k] = \binom{r+k-1}{k} p^r (1-p)^k$$

What's the probab. of seeing k failures before reaching the r^{th} success?

The very last trial must be a success.

In this class, and in the STAM tables, we use the parametrisation:

$$p = \frac{1}{1+\beta}$$

for some $\beta > 0$

With this parametrisation:

$$\checkmark \quad p_N(k) = \binom{r+k-1}{k} \left(\frac{1}{1+\beta} \right)^r \left(\frac{\beta}{1+\beta} \right)^k \quad k = 0, 1, 2, \dots$$

Returning to the generalization of the binomial coefficients, we generalize the pmf to be well-defined for any pair of parameters $r > 0$ and $\beta > 0$

We write $N \sim \text{NegBinomial}(r, \beta)$.

$$\checkmark \bullet \mathbb{E}[N] = r \cdot \beta$$

$$\checkmark \bullet \text{Var}[N] = r \cdot \beta (1+\beta) > \mathbb{E}[N]$$

$$\checkmark \bullet P_N(z) = \left(1 - \beta(z-1) \right)^{-r}$$

Note: The geometric dist'n is a special case: $r=1$.

We write $N \sim \text{Geometric}(\beta)$.

Do you remember (or can you find) one nifty property of the geometric dist'n?