

M339G: October 25th, 2024.

Any Bivariate Normal.

Random variables U and V are said to be bivariate normal w/ parameters $\mu_U, \mu_V, \sigma_U, \sigma_V$, and ρ if

$$\left(X := \frac{U - \mu_U}{\sigma_U}, Y := \frac{V - \mu_V}{\sigma_V} \right)$$

has the standard bivariate normal dist'n w/ correlation ρ .

Note: • $\rho(U, V) = ?$

By def'n:

$$\rho(U, V) = \frac{\text{Cov}[U, V]}{\text{SD}[U] \cdot \text{SD}[V]}$$

$$\begin{aligned} U &= \mu_U + \sigma_U \cdot X \\ V &= \mu_V + \sigma_V \cdot Y \end{aligned}$$

$$= \frac{\text{Cov}[\cancel{\mu_U + \sigma_U \cdot X}, \cancel{\mu_V + \sigma_V \cdot Y}]}{\text{SD}[\cancel{\mu_U + \sigma_U \cdot X}] \cdot \text{SD}[\cancel{\mu_V + \sigma_V \cdot Y}]}$$

$$= \frac{\text{Cov}[\sigma_U \cdot X, \sigma_V \cdot Y]}{\text{SD}[\sigma_U \cdot X] \cdot \text{SD}[\sigma_V \cdot Y]}$$

$$= \frac{\sigma_U \cdot \sigma_V \cdot \text{Cov}[X, Y]}{\sigma_U \cdot \text{SD}[X] \cdot \sigma_V \cdot \text{SD}[Y]} = \rho(X, Y) = \rho$$

μ_U and μ_V
are
deterministic

- U and V are independent

\Leftrightarrow

$$\rho = 0$$

Example. Midterm and Final.

Midterm and final scores in a large class have an (approximately) bivariate normal dist'n w/ parameters:

	<u>mean</u>	<u>sd</u>	
<u>midterm scores:</u>	65	18	
<u>final scores:</u>	60	20	correlation: 0.76

Q: What is the "estimated" mean final score of the students who were above the mean on the midterm?

→ Let U be the midterm score, and let V be the final score.

Let X and Y be U and V in std units, resp.

Our task is to find:

$$\mathbb{E}[V \mid U > \mu_U] = ?$$

Our ancillary task is to find:

$$\mathbb{E}[Y \mid X > 0] = \int_{-\infty}^{\infty} \mathbb{E}[Y \mid X = x] \cdot f_X(x \mid X > 0) dx$$

The Law of Total Probability

and for $x > 0$:

$$f_X(x \mid X > 0) dx = \frac{\mathbb{P}[X \in (x, x+dx) \mid X > 0]}{\mathbb{P}[X > 0]} =$$

$$= \frac{\mathbb{P}[X \in (x, x+dx)]}{\frac{1}{2}} = 2 \cdot f_X(x) dx$$

$$= 2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

$Y \mid X=x \sim \text{Normal}(px)$
 $\text{var} = 1-p^2$

$$\begin{aligned} \mathbb{E}[Y | X > 0] &= \int_0^{\infty} \underline{f_X(x)} \cdot 2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \frac{2g}{\sqrt{2\pi}} \int_0^{\infty} x \cdot e^{-\frac{x^2}{2}} dx = \boxed{u = \frac{x^2}{2}, du = xdx} \\ &= \frac{2g}{\sqrt{2\pi}} \int_0^{\infty} e^{-u} du = \frac{2g}{\sqrt{2\pi}} \left[-e^{-u} \right]_{u=0}^{\infty} = \frac{2g}{\sqrt{2\pi}} \end{aligned}$$

In this problem: $\frac{1.52}{\sqrt{2\pi}} = \underline{0.6063923}$

\Rightarrow Our answer is $60 + 20(0.6063923) = \underline{72.12785}$



Matrix Notation.

In two dimensions, we can place the means in a vector

$\begin{pmatrix} \mu_x \\ \mu_y \end{pmatrix}$ and the variances/covariance in a matrix

$$\Sigma = \begin{bmatrix} \sigma_x^2 & \sigma_{xy} \\ \sigma_{yx} & \sigma_y^2 \end{bmatrix}$$

(positive definite)

In 1D:
 $\text{UNN}(\mu, \sigma)$

$$f_U(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$