

Notes: This is a closed book and closed notes exam. The maximal score on the real exam will be 100 points.

There are many ways in which any single problem can be solved. The solutions herein are just one possible way to tackle the given problems.

Time: 50 minutes

All written work handed in by the student is considered to be
their own work, prepared without unauthorized assistance.

The University Code of Conduct

"The core values of The University of Texas at Austin are learning, discovery, freedom, leadership, individual opportunity, and responsibility. Each member of the university is expected to uphold these values through integrity, honesty, trust, fairness, and respect toward peers and community. As a student of The University of Texas at Austin, I shall abide by the core values of the University and uphold academic integrity."

"I agree that I have complied with the UT Honor Code during my completion of this exam."

Signature:

1.1. Formulas. If Y has the binomial distribution with parameters n and p , then $p_Y(k) = \mathbb{P}[Y = k] = \binom{n}{k} p^k (1-p)^{n-k}$, for $k = 0, \dots, n$, $\mathbb{E}[Y] = np$, $\text{Var}[Y] = np(1-p)$. The binomial coefficients are defined as follows for integers $0 \leq k \leq n$, $\binom{n}{k} = \frac{n!}{k!(n-k)!}$. The moment generating function of Y is given by $m_Y(t) = (pe^t + q)^n$.

If Y has a geometric distribution with parameter p , then $p_Y(k) = p(1-p)^k$ for $k = 0, 1, \dots$, $\mathbb{E}[Y] = \frac{1-p}{p}$, $\text{Var}[Y] = \frac{1-p}{p^2}$. Its mgf is $m_Y(t) = \frac{p}{1-qe^t}$ for t such that $qe^t < 1$.

If Y has a Poisson distribution with parameter λ , then $p_Y(k) = e^{-\lambda} \frac{\lambda^k}{k!}$ for $k = 0, 1, \dots$, $\mathbb{E}[Y] = \text{Var}[Y] = \lambda$. Its mgf is $m_Y(t) = e^{\lambda(e^t-1)}$.

If Y has a uniform distribution on $[l, r]$, its density is

$$f_Y(y) = \frac{1}{r-l} \mathbf{1}_{(l,r)}(y),$$

its mean is $\frac{l+r}{2}$, and its variance is $\frac{(r-l)^2}{12}$. Let $U \sim U(0, 1)$. The mgf of U is $m_U(t) = \frac{1}{t}(e^t - 1)$.

If Y has the standard normal distribution, then its mean is zero, its variance is one, and its density equals

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \quad x \in \mathbb{R}.$$

Its mgf is $m_Y(t) = e^{\frac{t^2}{2}}$.

If Y has the exponential distribution with parameter τ , then its cumulative distribution function is $F_Y(y) = 1 - e^{-\frac{y}{\tau}}$ for $y \geq 0$, its probability density function is $f_Y(y) = \frac{1}{\tau} e^{-y/\tau}$ for $y \geq 0$. Also, $\mathbb{E}[Y] = \text{SD}[Y] = \tau$. Its mgf is $m_Y(t) = \frac{1}{1-\tau t}$.

The mgf of $Y \sim \Gamma(k, \tau)$ is

$$m_Y(t) = \frac{1}{(1-\tau t)^k} \text{ for } t < 1/\tau.$$

Its expectation is $k\tau$ and its variance is $k\tau^2$. The χ^2 -distribution with n degrees of freedom is the special case $\Gamma(\frac{n}{2}, 2)$

1.2. DEFINITIONS.

Problem 1.1. Write down the definition of the **independence** of two events E and F .

Solution: Two events E and F are said to be independent if

$$\mathbb{P}[E \cap F] = \mathbb{P}[E]\mathbb{P}[F].$$

Problem 1.2. Write down the definition of the **cumulative distribution function** of a random variable Y .

Solution:

$$F_Y(x) = \mathbb{P}[Y \leq x] \quad \text{for } x \in \mathbb{R}.$$

Problem 1.3. Let Y be a continuous random variable with the probability density function denoted by f_Y . Let g be a function taking real values such that $g(Y)$ is well defined. How is $\mathbb{E}[g(Y)]$ evaluated using f_Y , if it exists?

Solution: We have that

$$\mathbb{E}[g(Y)] = \int_{-\infty}^{\infty} g(y)f_Y(y) dy$$

if the above integral is absolutely convergent.

1.3. TRUE/FALSE QUESTIONS.

Problem 1.4. (5 points) The pdf (probability density function) of the random variable Y is $f_Y(y) = c \exp(-2y)$ for $y > 0$ and $f_Y(y) = 0$ for $y \leq 0$. The constant c is 2. *True or false? Why?*

Solution: TRUE

We can recognize Y as exponential with mean $\tau = \frac{1}{2}$. Also, we have $1 = \int_0^{\infty} ce^{-2y} dy = c \times \frac{1}{2}$.

Problem 1.5. Let $Y \sim b(n, p)$. Then, $\mathbb{E}[Y] \geq \text{Var}[Y]$. *True or false? Why?*

Solution: TRUE

We have

$$\mathbb{E}[Y] = np \geq np(1 - p) = \text{Var}[Y]$$

Problem 1.6. (5 points) Let Y be a random variable with mean $\mu = 1$ and standard deviation equal to $\sigma = 4$. Then, $\mathbb{E}[Y^2] = 5$. *True or false? Why?*

Solution: FALSE

$$\mathbb{E}[X^2] = \text{Var}[X] + (\mathbb{E}[X])^2 = 16 + 1^2 = 17.$$

Problem 1.7. (5 points) Let Y be a continuous random variable. Then, $\mathbb{P}[Y = y] = 0$ for every $y \in \mathbb{R}$. *True or false? Why?*

Solution: **TRUE**

For every y , we have that, in our usual notation,

$$\mathbb{P}[Y = y] = \mathbb{P}[y \leq Y \leq y] = \int_y^y f_Y(u) du = 0.$$

1.4. Free-response problems.

Please, explain carefully all your statements and assumptions. Numerical results or single-word answers without an explanation (even if they're correct) are worth 0 points.

Problem 1.8. (10 points) Assume that the time T until the arrival of the bus at the bus stop is exponential with mean 5. You have been waiting at the bus stop for 3 minutes. What is the probability that your **total waiting time** will **exceed** 7 minutes? *Note: You can leave your response in the form that uses the exponential function, but you must **simplify** it as much as possible!*

Solution: By the memoryless property, the probability equals

$$\mathbb{P}[T > 4] = e^{-\frac{4}{5}}.$$

Problem 1.9. (15 points) Let $Y \sim b(n, p)$ such that its mean equals 8 and its variance equals 1.6. What is the probability of exactly 3 successes? *Note: Leave your answer in the form of a fraction containing only integers **without** any binomial coefficients.*

Solution: We are given that

$$\mathbb{E}[Y] = np = 8 \quad \text{and} \quad \text{Var}[Y] = np(1-p) = 1.6 \quad \Rightarrow \quad 1-p = 0.2 \quad \Rightarrow \quad p = 0.8 \quad \Rightarrow \quad n = 10.$$

So,

$$\mathbb{P}[Y = 3] = \binom{10}{3} (0.8)^3 (0.2)^7 = \frac{10 \cdot 9 \cdot 8}{3!} \cdot \frac{4^3}{5^{10}} = \frac{120 \cdot 4^3}{5^{10}} = \frac{24 \cdot 4^3}{5^9}.$$

Problem 1.10. A die is rolled 5 times; let the obtained numbers be given by Y_1, \dots, Y_5 . Use counting to compute the probability that

- (1) all Y_1, \dots, Y_5 are even?
- (2) at most 4 of Y_1, \dots, Y_5 are odd?
- (3) the values of Y_1, \dots, Y_5 are all different from each other?

Solution: There are 6^5 different (equally likely) outcomes of 5 rolls of a die. We need to find the number of those 5-tuples of rolls that correspond to the situation described in each question, and simply divide by 6^5 .

- (1) The number of 5-tuples of rolls where each outcome is even is 3^5 , because each roll can come up an even number in three ways, namely 2, 4 or 6. Therefore, The answer is

$$\frac{3^5}{6^5} = \left(\frac{1}{2}\right)^5 = \frac{1}{32}.$$

- (2) We have seen above that the number of 5-tuples of rolls where all values are even is 3^5 . Therefore, the number of 5-tuples where at most 4 are even is $6^5 - 3^5$, and the required probability is

$$\frac{6^5 - 3^5}{6^5} = \frac{31}{32}.$$

- (3) Exactly $6 \times 5 \times 4 \times 3 \times 2 = 6!$ 5-tuples have all numbers different. Therefore, the required probability is $\frac{6!}{6^5} = \frac{5}{54}$.

Problem 1.11. *Source: "Probability" by Jim Pitman.*

Consider a set-up in which a transmitter is transmitting either a 0 or a 1 and the receiver indicates that it received either a 0 or a 1. Denote the events that $i = 0, 1$ was transmitted by T_i , and the events that $i = 0, 1$ was indicated as received by R_i .

It is possible to have transmission errors. In fact, you are given the following data on accuracy and the frequency of transmitted signals:

$$\mathbb{P}[R_0 | T_0] = 0.99, \quad \mathbb{P}[R_1 | T_1] = 0.98,$$

and

$$\mathbb{P}[T_0] = 0.75.$$

- (a) Given that the receiver indicated 1, what is the probability that there was an error in the transmission?
- (b) What is the overall probability that there was an error in transmission?

Solution:

- (1) We need $\mathbb{P}[T_0 | R_1]$. By the Bayes formula,

$$\begin{aligned} \mathbb{P}[T_0 | R_1] &= \frac{\mathbb{P}[R_1 | T_0] \mathbb{P}[T_0]}{\mathbb{P}[R_1 | T_0] \mathbb{P}[T_0] + \mathbb{P}[R_1 | T_1] \mathbb{P}[T_1]} \\ &= \frac{(1 - 0.99) \times 0.75}{(1 - 0.99) \times 0.75 + 0.98 \times 0.25} \\ &= \frac{3}{101}. \end{aligned}$$

(2) An error will happen if $T_0 \cap R_1$ or $T_1 \cap R_0$ occur, i.e.,

$$\begin{aligned}
 \mathbb{P}[\text{error}] &= \mathbb{P}[T_0 \cap R_1] + \mathbb{P}[T_1 \cap R_0] \\
 &= \mathbb{P}[R_1|T_0] \times \mathbb{P}[T_0] + \mathbb{P}[R_0|T_1] \times \mathbb{P}[T_1] \\
 &= (1 - \mathbb{P}[R_0|T_0]) \times \mathbb{P}[T_0] \\
 &\quad + (1 - \mathbb{P}[R_1|T_1]) \times (1 - \mathbb{P}[T_0]) \\
 &= 0.01 \times 0.75 + 0.02 \times 0.25 \\
 &= \frac{1}{80}.
 \end{aligned}$$

Problem 1.12. *Source: Sample P Exam, Problem #483.*

A doctor tests 100 patients for two diseases, **A** and **B**. Each patient has probability p of having disease **A** and probability p of having disease **B**, with $0 \leq p \leq \frac{1}{2}$.

For each patient, the event of having disease **A** and the event of having disease **B** are independent. The test outcomes for different patients are mutually independent.

The variance of the number of patients who have disease **A** is 9.

Calculate the variance of the number of patients who have at least one of the two diseases.

Solution: Let Y_A be the number of people who have disease **A** and let Y_B be the number of people who have disease **B**. We are given that they are both $b(100, p)$.

For disease **A**, we are also given

$$\text{Var}[Y_A] = 100p(1 - p) = 9 \quad \Rightarrow \quad p = 0.1$$

since we know that $0 \leq p \leq 0.5$. So, the probability that a single patient has neither of the two diseases is $0.9 \cdot 0.9 = 0.81$. Hence, the probability that a single patient has at least one of the two diseases is $1 - 0.81 = 0.19$. The variance of the number of people who have at least one disease is

$$100(0.81)(0.19) = 15.39.$$

Problem 1.13. *Source: Sample P exam, Problem #29.*

An actuary has discovered that policyholders are three times as likely to file two claims as to file four claims. The number of claims filed has a Poisson distribution.

Calculate the variance of the number of claims filed.

Solution: Let N denote the number of claims. We are given that $N \sim P(\lambda)$. Also,

$$p_2 = 3p_4 \quad \Rightarrow \quad e^{-\lambda} \frac{\lambda^2}{2!} = 3e^{-\lambda} \frac{\lambda^4}{4!} \quad \Rightarrow \quad \lambda^2 = 4 \quad \Rightarrow \quad \lambda = 2.$$

We now know that $\text{Var}[N] = \lambda = 2$.

Problem 1.14. *Source: Sample P exam, Problem #442.*

Let Y be a random variable that is uniform on $[a, b]$. The probability that Y is greater than 8 is 0.60. The probability that Y is greater than 11 is 0.20.

Calculate the variance of Y .

Solution: We are given the two probabilities that imply that the probability that Y lands between 8 and 11 is 0.4. So,

$$\frac{11 - 8}{b - a} = \frac{2}{5} \quad \Rightarrow \quad b - a = \frac{15}{2}.$$

Finally, the variance is

$$\frac{\left(\frac{15}{2}\right)^2}{12} = \frac{225}{48}.$$

Problem 1.15. A random variable Y has the normal distribution with mean 6 Its 0.8-quantile is 8. What is its standard deviation?

Solution: Since $Y \sim N(\mu = 6, \sigma)$, we know that Y can be expressed as

$$Y = \mu + \sigma Z$$

where Z is standard normal. We are also given that

$$\mathbb{P}[Y \leq 8] = 0.8.$$

So,

$$\mathbb{P}[6 + Z\sigma \leq 8] = 0.8 \Rightarrow \mathbb{P}[Z\sigma \leq 2] = 0.8 \Rightarrow \mathbb{P}\left[Z \leq \frac{2}{\sigma}\right] = 0.8.$$

Using the standard normal tables, we see that

$$\frac{2}{\sigma} = 0.8416 \Rightarrow \frac{1}{\sigma} = 0.4208 \Rightarrow \sigma = \frac{1}{0.4208}.$$

Problem 1.16. *Source: Sample P exam, Problem #385.*

A computer manufacturer collects data on how long it takes before its computers fail. The time to fail, in years, follows an exponential distribution. Twenty percent of its computers fail within two years.

The probability a randomly selected computer fails before time t^* , in years, is 0.80.

Calculate t^* .

Solution: Let T be the exponential failure time, i.e., $T \sim E(\tau)$. We are given that

$$\mathbb{P}[T \leq 2] = 0.2 \Rightarrow 1 - e^{-\frac{2}{\tau}} = 0.2 \Rightarrow e^{-\frac{2}{\tau}} = 0.8 \Rightarrow \frac{1}{\tau} = -\frac{1}{2} \ln(0.8) \Rightarrow \tau = -\frac{2}{\ln(0.8)}.$$

We are looking for t^* such that

$$\mathbb{P}[T \leq t^*] = 0.8 \Rightarrow 1 - e^{-\frac{t^*}{\tau}} = 0.8 \Rightarrow e^{-\frac{t^*}{\tau}} = 0.2 \Rightarrow t^* = \frac{2 \ln(0.2)}{\ln(0.8)}.$$

1.5. MULTIPLE CHOICE QUESTIONS.

Problem 1.17. (5 pts) A biased coin for which *Heads* is twice as likely as *Tails* is tossed twice, independently. If the outcomes are two consecutive *Heads*, Bertie gets a \$15 reward; if the outcomes are different, Bertie gets a \$10 reward; if the outcomes are two consecutive *Tails*, Bertie has to pay \$5. What is the expected value of Bertie's winnings?

- (a) 75/9
- (b) 80/9
- (c) 85/9
- (d) 95/9

(e) None of the above.

Solution: The correct answer is (d).

Since *Heads* is twice as likely as *Tails*, *Heads* appears with probability $2/3$, while *Tails* appears with probability $1/3$.

Let X denote the amount Bertie wins. Then, X has the following distribution:

$$X \sim \begin{cases} 15, & \text{with probability } 4/9, \\ 10, & \text{with probability } 4/9, \\ -5, & \text{with probability } 1/9. \end{cases}$$

$$\mathbb{E}[X] = \frac{4}{9}(15) + \frac{4}{9}(10) + \frac{1}{9}(-5) = \frac{95}{9}.$$

Problem 1.18. (5 points) There are 10 green balls and 20 red balls in a box, well mixed together. We pick a ball at random, and then roll a die. If the ball was green, we write down the outcome of the die, but if the ball is red, we write down the number 3.

If the number 3 was written down, the probability that the ball was green is:

- (a) $1/3$
- (b) $1/2$
- (c) $5/6$
- (d) 1
- (e) none of the above

Solution: The correct answer is (e).

Let R denote the event when the ball drawn was red, and $G = R^c$ the event corresponding to drawing a green ball, so that $\mathbb{P}[R] = 2/3$ and $\mathbb{P}[G] = 1/3$. If X denotes the number written down, we have

$$\mathbb{P}[X = 3|G] = 1/6 \text{ and } \mathbb{P}[X = 3|R] = 1.$$

Using Bayes formula,

$$\begin{aligned} \mathbb{P}[G|X = 3] &= \frac{\mathbb{P}[X = 3|G] \times \mathbb{P}[G]}{\mathbb{P}[X = 3|G] \times \mathbb{P}[G] + \mathbb{P}[X = 3|R] \times \mathbb{P}[R]} \\ &= \frac{1/6 \times 1/3}{1/6 \times 1/3 + 1 \times 2/3} = \frac{1}{13}. \end{aligned}$$