

April 4th, 2025.

Moment Generating Function.

For a random variable Y , and for an independent argument denoted by t , we define the **moment generating function (mgf)** of Y as this function of t :

$$\underline{M_Y(t)} := \underline{\mathbb{E} \left[e^{Y \cdot t} \right]} \quad \text{for all } t \text{ such that the expectation exists}$$

- Note:
- $M_Y(0) = \underline{1} \Rightarrow$ @ least $t=0$ is in the domain
 - We say that the mgf exists if it's finite for t such that $|t| < b$ for some $b > 0$.

Goal: To understand e^X w/ $X \sim \text{Normal}(\text{mean}=m, \text{var}=\sigma^2)$

Recall: In terms of $Z \sim N(0,1)$,

$$X = m + \sigma \cdot Z$$

Fact: $M_Z(t) = e^{\frac{t^2}{2}}$ for all $t \in \mathbb{R}$

\Rightarrow For any normal X :

$$\begin{aligned} \underline{M_X(t)} &= \mathbb{E} \left[e^{X \cdot t} \right] = \mathbb{E} \left[e^{(m + \sigma \cdot Z) \cdot t} \right] \\ &= \mathbb{E} \left[e^{mt} \cdot e^{\sigma t \cdot Z} \right] \\ &= e^{mt} \mathbb{E} \left[e^{\sigma t \cdot Z} \right] \\ &= e^{mt} \cdot M_Z(\sigma t) \\ &= e^{mt} \cdot e^{\frac{\sigma^2 t^2}{2}} = \underline{e^{mt + \frac{\sigma^2 t^2}{2}}} \end{aligned}$$

The lognormal distribution.

Definition 1.1. Let $X \sim \text{Normal}(\text{mean} = m, \text{variance} = \nu^2)$. Define the random variable $Y = e^X$. We say that the random variable Y is *lognormally distributed*.

1.1. First properties.

- The expected value of the lognormally distributed random variable Y can be obtained as follows:

$$\mathbb{E}[Y] = \mathbb{E}[e^X] = M_X(1) = e^{m + \frac{\nu^2}{2}}.$$

- Let Y be a lognormal and let $a \neq 0$. Then, the random variable Y^a is also lognormal. *Note:* For $a = 0$, we get a degenerate random variable at 1 which can, technically, be interpreted as lognormal, but is not fun.
- Let Y_1 and Y_2 be independent and lognormally distributed. Then, $Y_1 Y_2$ is also lognormal.

1.2. Quantiles.

Definition 1.2. For p such that $0 < p < 1$, we define the $100p^{\text{th}}$ quantile of a random variable X as any value π_p such that

$$F_X(\pi_p -) \leq p \leq F_X(\pi_p).$$

In particular, the 50^{th} quantile is referred to as the *median*.

Note: When the random variable X is continuous, we can obtain the $100p^{\text{th}}$ quantile by simply solving for π_p in

$$F_X(\pi_p) = p.$$

Consider a probability p . Let z_p be the $100p^{\text{th}}$ quantile of the standard normal distribution. Let Y be lognormally distributed as above. My claim is that the value

$$y_p = e^{m + \nu z_p}$$

is the $100p^{\text{th}}$ quantile of Y . Let us simply verify this claim by calculating $F_Y(y_p)$. We have, with $Z \sim N(0, 1)$,

$$F_Y(y_p) = \mathbb{P}[Y \leq y_p] = \mathbb{P}[e^X \leq y_p] = \mathbb{P}[e^{m + \nu Z} \leq e^{m + \nu z_p}].$$

Since the logarithmic function is increasing, we have that the above equals

$$F_Y(y_p) = \mathbb{P}[m + \nu Z \leq m + \nu z_p] = \mathbb{P}[Z \leq z_p] = p.$$

The above concludes our proof.

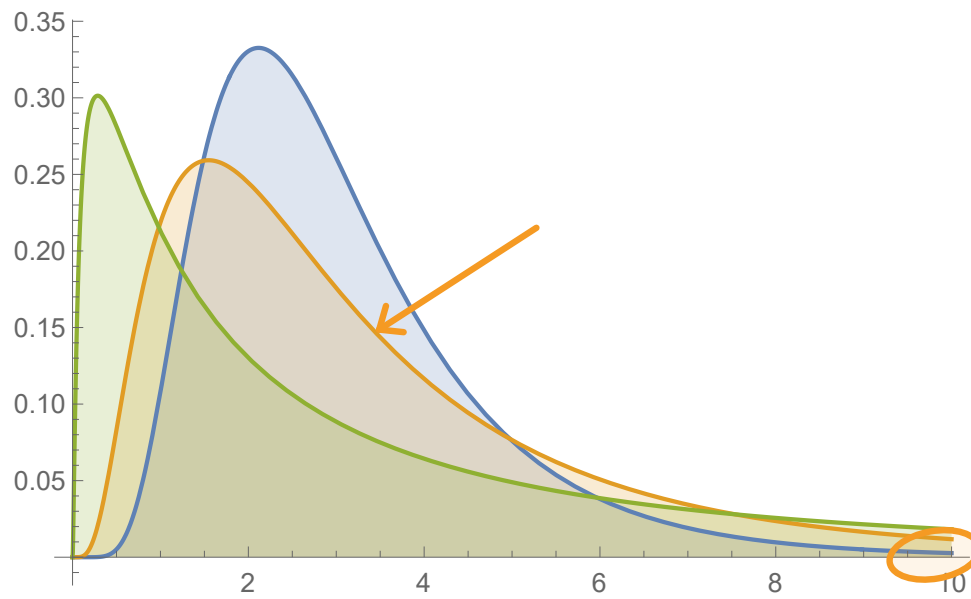
In particular, since the median of the standard normal distribution equals 0, the median of the lognormal distribution will be e^m .

Note: Since

$$e^m < e^{m + \frac{\nu^2}{2}},$$

(1.1)

i.e., since the mean of a lognormal distribution always exceeds the median, we say that it's *right-skewed*. In fact, this is what its probability density function looks like.



Jensen's Inequality.

Caveat:

$$\mathbb{E}[e^x] \geq e^{\mathbb{E}[x]}$$

Theorem. Let X be a random variable,
and let g be a convex function
such that $g(x)$ is well-defined
and $\mathbb{E}[g(x)]$ exists.

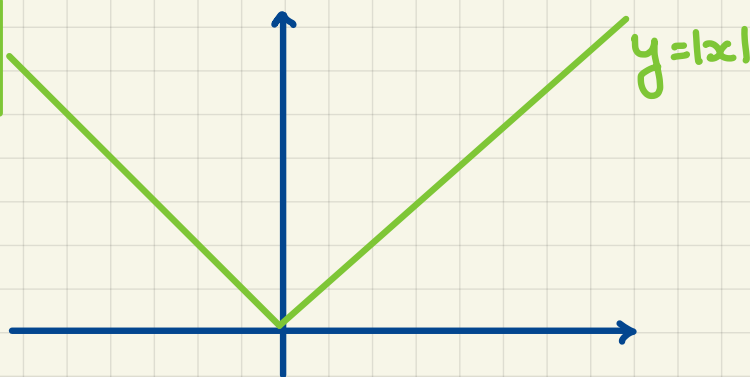
Then,

$$\mathbb{E}[g(x)] \geq g(\mathbb{E}[x])$$

Examples. i.

$$g(x) = |x|$$

$$\mathbb{E}[|x|] \geq |\mathbb{E}[x]|$$



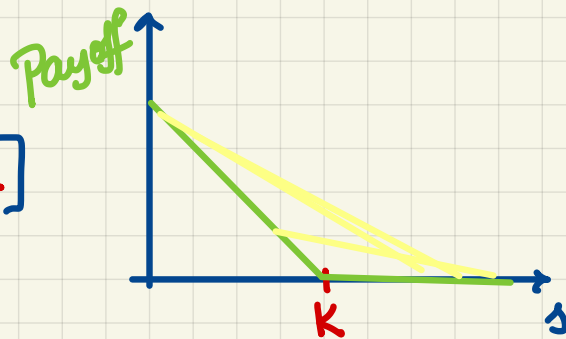
ii. Consider a European put w/ strike K .

Its payoff function: $v_p(s) = (K - s)_+$

The expected payoff

$$\mathbb{E}[v_p(S(T))] = \mathbb{E}[(K - S(T))_+]$$

By Jensen's inequality
it's $\geq (K - \mathbb{E}[S(T)])_+$



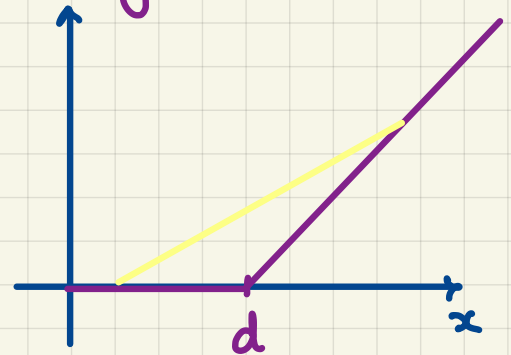
iii. In classical insurance:

$\begin{cases} X \dots \text{(ground-up) loss, i.e., severity r.v.} \\ d \dots \text{deductible} \end{cases}$

The insurer pays $(X-d)_+$, i.e., $g(x) = (x-d)_+$

By Jensen's inequality:

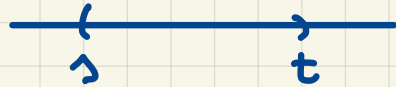
$$\mathbb{E}[(X-d)_+] \geq (\mathbb{E}[X] - d)_+$$



Log-Normal Stock Prices.

Temporarily fix a time horizon T .

$S(t)$, $t \in [0, T]$... time- t stock price



Recall:

$$R(s, t) = \ln\left(\frac{S(t)}{S(s)}\right) \Leftrightarrow S(t) = S(s) \cdot e^{R(s, t)}$$

In particular:

$R(0, T)$... realized return for $[0, T]$

We model realized returns as normal.

$$R(0, T) \sim \text{Normal}(\text{mean} = \textcircled{\mu}, \text{var} = \textcircled{\sigma^2})$$

$\Rightarrow S(T)$ is **lognormal**

and

$\mathbb{E}^*[S(T)] = S(0)e^{\mu + \frac{\sigma^2}{2}}$

Market model:

- Riskless asset w/ ccrf r
- Risky asset: a non-dividend paying stock w/ **volatility** σ