

$$Q: \mathbb{E}[(Y-a)^2] \xrightarrow{a} \min$$

$$a^* = \mathbb{E}[Y]$$

November 17<sup>th</sup>, 2025.

## Maximum Likelihood Estimation.

### Likelihood.

Def'n. Given a random sample  $Y_1, Y_2, \dots, Y_n$  from a discrete dist'n  $D$  w/ an unknown parameter  $\theta$ , the likelihood f'n is defined as

$$\begin{aligned} L(\theta; y_1, y_2, \dots, y_n) &= p_{Y_1, \dots, Y_n}^\theta(y_1, \dots, y_n) \\ &= p_{Y_1}^\theta(y_1) \cdot p_{Y_2}^\theta(y_2) \cdots p_{Y_n}^\theta(y_n) \\ &= p^\theta(y_1) \cdot p^\theta(y_2) \cdots p^\theta(y_n) \end{aligned}$$

where  $p^\theta$  is the pmf of  $D$ .

### Example. Bernoulli.

$$Y_1, Y_2, \dots, Y_n \sim B(p)$$

$$p \leftrightarrow \theta$$

$$\text{pmf of } B(p): p(y) = \begin{cases} p & \text{for } y=1 \\ 1-p & \text{for } y=0 \end{cases}$$

$$p(y) = p^y (1-p)^{1-y} \quad \text{for } y=0,1$$

$$\begin{aligned} L(p; y_1, y_2, \dots, y_n) &= p^{y_1} (1-p)^{1-y_1} p^{y_2} (1-p)^{1-y_2} \cdots p^{y_n} (1-p)^{1-y_n} \\ &= \prod_{i=1}^n p^{y_i} \cdot \prod_{i=1}^n (1-p)^{1-y_i} \\ &= p^{\sum y_i} \cdot (1-p)^{n - \sum y_i} \end{aligned}$$

For computational reasons, take the  $\ln(\cdot)$  increasing and get the log-likelihood,

$$\text{i.e., } l(p; y_1, \dots, y_n) = \sum_{i=1}^n y_i \cdot \ln(p) + (n - \sum_{i=1}^n y_i) \ln(1-p) \xrightarrow{p} \max$$

Next, we differentiate with respect to  $p$

$$l'(p; y_1, \dots, y_n) = \left( \sum_{i=1}^n y_i \right) \frac{1}{p} + \left( n - \sum_{i=1}^n y_i \right) (-1) \cdot \frac{1}{1-p}$$

We equate this derivative to zero and solve for  $p$ .

$$\left( \sum_i y_i \right) \cdot \frac{1}{p} - \left( n - \sum_i y_i \right) \frac{1}{1-p} = 0 \quad / \cdot p(1-p)$$

$$\left( \sum_i y_i \right) (1-p) - \left( n - \sum_i y_i \right) p = 0$$

$$\sum_i y_i - p \cdot \sum_i y_i - np + p \cdot \sum_i y_i = 0$$

$$p = \frac{\sum y_i}{n} = \bar{y}$$

$$\hat{p}_{MLE} = \bar{y}$$



Def'n. If  $Y_1, \dots, Y_n$  come from a continuous dist'n  $D$  w/ pdf  $f^\theta$ , then the likelihood f'n is

$$L(\theta; y_1, \dots, y_n) = f_{Y_1, \dots, Y_n}^\theta(y_1, \dots, y_n) = f^\theta(y_1) \cdot f^\theta(y_2) \cdots f^\theta(y_n)$$

Example. Exponential.

$$Y_1, \dots, Y_n \sim E(\theta)$$

$$\text{the pdf is } f^\theta(y) = \frac{1}{\theta} e^{-\frac{y}{\theta}} \quad y > 0$$

$$\begin{aligned} \Rightarrow L(\theta; y_1, \dots, y_n) &= \frac{1}{\theta} e^{-\frac{y_1}{\theta}} \cdot \frac{1}{\theta} e^{-\frac{y_2}{\theta}} \cdots \frac{1}{\theta} e^{-\frac{y_n}{\theta}} \\ &= \left( \frac{1}{\theta} \right)^n e^{-\frac{1}{\theta} \sum y_i} \end{aligned}$$

$$\Rightarrow l(\theta; y_1, \dots, y_n) = -n \cdot \ln(\theta) + \left( -\frac{1}{\theta} \sum_{i=1}^n y_i \right)$$

$$\Rightarrow l'(\theta; y_1, \dots, y_n) = -n \cdot \frac{1}{\theta} + (-1) \cdot \frac{1}{\theta^2} \sum_{i=1}^n y_i = 0 \quad / \cdot \theta^2$$

$$-n \cdot \theta + \sum_{i=1}^n y_i = 0$$

$$\theta = \bar{y} \Rightarrow \hat{\theta}_{MLE} = \bar{y}$$

### Example. Normal

$$Y_1, \dots, Y_n \sim N(\mu, \sigma)$$

the pdf:  $f(y) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)^2}$  for all  $y \in \mathbb{R}$

$$L(\mu, \sigma; y_1, \dots, y_n) = \prod_{i=1}^n \left( \frac{1}{\sigma\sqrt{2\pi}} \cdot e^{-\frac{1}{2} \frac{(y_i - \mu)^2}{\sigma^2}} \right)$$
$$= (2\pi)^{-\frac{n}{2}} \cdot \sigma^{-n} \cdot e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2}$$
$$\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2}$$

unknown

$$l(\mu; y_1, \dots, y_n) = \text{const} - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2$$

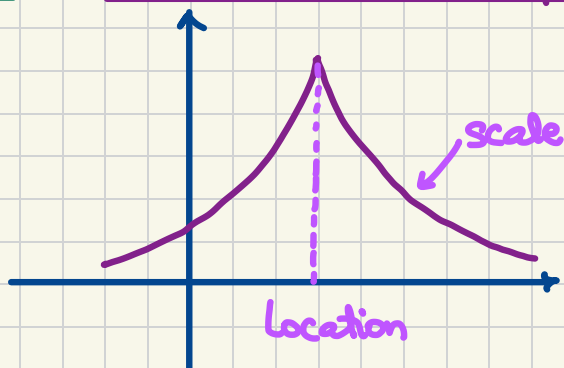
$$l'(\mu; y_1, \dots, y_n) = -\frac{1}{\cancel{2\sigma^2}} \sum_{i=1}^n \cancel{2}(y_i - \mu)(-1) = 0$$

$$\sum_{i=1}^n (y_i - \mu) = 0$$

$$\sum_{i=1}^n y_i - n \cdot \mu = 0$$

$$\hat{\mu}_{MLE} = \bar{Y}$$

### Example. Laplace / Double exponential



Look up on  
Wikipedia