Math Review Summer 2017

Topic 7

7. Probability

We go through the basics of probability building from the foundations and notations we need to work with. This topic is very definition and information heavy – I have tried to add as many examples as I could, but still the lecture outweighs the exercises. It might be the nature of the coverage of this topic – we'll try to keep things interesting and cruise through the material. The main goal is to nail down the basic notions of probability and build tools to use in statistical inferences and modelling (eventually).

7.1 Random variables and distributions

An experiment is any procedure that can be infinitely repeated and has a well-defined set of outcomes.

<u>Definition</u>. **Experiment:**

We are not always interested in an experiment itself, but rather in some consequence of its random outcome. Such consequences, when real valued, may be thought of as random variables.

Definition. Random variable:

7.1.1. Discrete random variable

A **discrete random variable** is a random variable that takes on only a finite number of values. The simplest type of discrete random variable is called a Bernoulli random variable (or "binary") where $X = \{0, 1\}$. Typically an outcome of "1" denotes a success and an outcome of "0" denotes a failure.

<u>Definition</u> .	Sample	space:
---------------------	--------	--------

<u>Definition</u>. Event:

Example. If I plant ten bean seeds and count the number that germinate, the sample space is $S = \{0,1,2,3,4,5,6,7,8,9,10\}$

Q: If you toss a 'fair' coin three times and record the result, what is the sample space?

 \mathcal{A} :

<u>Definition</u>. An event, E, is a designated collection of sample outcomes. An event is, therefore, a subset of the sample space.

Example. As we saw, by definition, an event is a subset of the sample space. We can specify an event by listing all the outcomes that make it up. In the coin toss example, let A be the event 'more heads than tails.'

How can we write out *A*?

Q: Imagine you toss a 'fair' coin three times and record the result. What is the event 'heads on last throw'? Denote the event as B and write it out.

The events E_1, E_2, \ldots, E_m are **mutually exclusive** if for each pair of events E_i and E_j in the collection we have $E_i \cap E_j = \emptyset$; i.e. the two events have no outcomes in common. The complement of E, E^c , denotes all possible outcomes not in the event E.

7.1.2. What is probability?

No, we will not get into a philosophical discussion! But broadly speaking, it is hard to concretely define what a probability is. Some view it as a 'limiting frequency.' For example, if I toss a coin 1000 times, I am likely to get tails about ½ of the time. Other may view it as a more subjective matter. For our purposes, we will work with a mathematical construct centered around an experiment which satisfies some key axioms.

Probability:

Q: What is $E \cap E^c$?

 \mathcal{A} :

Q: What is $P(E) + P(E^c) = ?$

Discrete random variable:

Notation use: Often we use capital letters at the end of the alphabet, X, Y, Z, to define the random variable. The corresponding lower case letters, x, y, z, to represent the actual values. Thus, the expression P(X = x) symbolizes the probability that the random variable X takes on the particular value x. Often, this is written simply as P(x). Likewise, $P(X \le x)$ is the probability that the random variable X is less than or equal to the specific value x; $P(a \le X \le b)$ is the probability that X lies between values A and A.

In context of a binary variable, these conditions imply that we only need to state one of the two probabilities, i.e. $P(X = 1) = \alpha$ implies that $P(X = 0) = 1 - \alpha$.

We can use all possible probabilities to create the **probability density function (pdf).** The pdf is just a function where the input is any possible outcome, x_i , of the random variable X and the outcome is the probability of that outcome, p_i . The function is denoted as:

$$pdf = f(x_i) = p_i \text{ for } i = 1, 2, ..., k$$

$$f(x) = 0$$
 if $x \neq x_i$ for some $i = 1, 2, ..., k$

Example. The probability density function for a standard die is given by:

Let's try a more involved example as a group.

Group Example. Consider the simple experiment of tossing a coin three times. Let X = number of times the coin comes up heads. Write out the pdf for this experiment.

The **cumulative distribution function** is then just the probability that the outcome is below some number:

For a discrete random variable this function reduces to:

Recall that outcomes of an experiment are individual outcomes that are **mutually** exclusive.

Example. Consider a baseball player who has a 68% chance of striking out, a 15% chance of getting to first base, a 10% chance of getting to second base, a 5% chance of getting to third base, and a 2% chance of hitting a home run. Then, the baseball players pdf and cdf are given by:

7.1.2. Continuous random variable

A continuous random variable is a random variable that can take on every point in some interval of the real numbers. A continuous random variable differs from a discrete random variable in that it takes on an uncountable infinite number of possible outcomes.

For example, if we let X denote the height (in meters) of a randomly selected maple tree, then X is a continuous random variable. In other words, the probability of any given number occurring is zero. This zero probability is because the continuous random variable can take on so many possible outcomes that we cannot count them (infinitely many). There is, however, a nonzero probability that the outcome falls in a certain range or interval. To compute this probability we generally use the pdf.

Note that each individual point and image on a pdf does not give us the probability of the event occurring as it does in the discrete case, i.e.:

$$f(x_i) \neq p_i$$

Instead, in the continuous case, we have:

The cdf is then given by:

One of the important continuous distribution is the Standard Normal Distribution. The Random Variable if often denoted by Z. The graph for Z is a symmetrical bell-shaped curve:

Example. Find P(0 < z < 0.45). You can use the standard normal distribution table (for now you can believe me), that this is equal to:

Q: What is the P(Z < -0.9)? How about P(-1.40 < z < -1.20)?

z	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	80.0	0.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7703	0.7734	0.7764	0.7794	0.7823	0.7852
8.0	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
9.0	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.90147
1.3	0.90320	0.90490	0.90658	0.90824	0.90988	0.91149	0.91309	0.91466	0.91621	0.91774
1.4	0.91924	0.92073	0.92220	0.92364	0.92507	0.92647	0.92785	0.92922	0.93056	0.93189
1.5	0.93319	0.93448	0.93574	0.93699	0.93822	0.93943	0.94062	0.94179	0.94295	0.9440
1.6	0.94520	0.94630	0.94738	0.94845	0.94950	0.95053	0.95154	0.95254	0.95352	0.95449
1.7	0.95543	0.95637	0.95728	0.95818	0.95907	0.95994	0.96080	0.96164	0.96246	0.96327
1.8	0.96407	0.96485	0.96562	0.96638	0.96712	0.96784	0.96856	0.96926	0.96995	0.97062
1.9	0.97128	0.97193	0.97257	0.97320	0.97381	0.97441	0.97500	0.97558	0.97615	0.97670
2.0	0.97725	0.97778	0.97831	0.97882	0.97932	0.97982	0.98030	0.98077	0.98124	0.98169
2.1	0.98214	0.98257	0.98300	0.98341	0.98382	0.98422	0.98461	0.98500	0.98537	0.98574
2.2	0.98610	0.98645	0.98679	0.98713	0.98745	0.98778	0.98809	0.98840	0.98870	0.98899

Properties of the CDF for a continuous random variable

7.1.3. Multiple random variables - Joint distributions

We have thus far considered discrete and continuous distributions but we only focused on single random variables. Probability distributions can, however, be applied to grouped random variables which gives rise to joint probability distributions. We'll mostly focus on 2-dimensional distributions (i.e. only two random variables) but higher dimensions (more than two variables) are also possible.

Essentially, joint probability distributions describe situations where by both outcomes represented by random variables occur. While we only X to represent the random variable, we now have X and Y as the pair of random variables.

<u>Definition</u>. **Joint distribution**:

Example. The table below represents the joint probability distribution obtained for the outcomes when a die is flipped and a coin is tossed.

f(x,y)	1	2	3	4	5	16	Row Totals
Heads	a	b	С	d	e	f	α
Tails	g	h	i	j	k	1	β
Column Totals	γ	δ	ε	ζ	θ	ψ	ω

In the table above, x = 1, 2, 3, 4, 5, 6 as outcomes when the die is tossed while y = Heads, Tails are outcomes when the coin is flipped. The letters a through l represent the joint probabilities of the different events formed from the combinations of x and y while the Greek letters represent the totals and ω should equal to 1. The row sums and column sums are referred to as the marginal probability distribution functions (PDF).

\mathcal{Q} : In the above table, assume that flipping a coin and tossing a die are independent. What are the letters a through l equal to? How about α and δ for instance?						
\mathcal{A} :						

We will not spend too much time on continuous joint distributions, but as one can expect, the results can be summarized as:

For a given interval A, we have:

$$P[(X,Y) \in A] = \iint_A f(x,y) dx \, dy$$

This tells us that in order to determine the probability of an event A, we must integrate the function f(x, y) over the space defined by the event A.

Example and exercise:

A certain farm produces two kinds of eggs on any given day; organic and non-organic. Let these two kinds of eggs be represented by the random variables X and Y respectively. Given that the joint probability density function of these variables is given by

$$f(x,y) = \begin{cases} \frac{2}{3}(x+2y), 0 \le x \le 1, 0 \le y \le 1, \\ 0 \text{ elsewhere} \end{cases}$$

Find the $P(X \le \frac{1}{2}, Y \le \frac{1}{2})$

7.1.4. Conditional distributions

<u>Definition</u>: **Conditional distribution**:

If X and Y are independent random variables then what do we have?

 \mathcal{Q} : Say you have two binary random variables given by X and Y with the following conditional probabilities. Are X and Y independent?

$$f_{Y|X}(1|1) = 0.85$$

 $f_{Y|X}(1|0) = 0.70$
 $f_{Y|X}(0|1) = 0.15$
 $f_{Y|X}(0|0) = 0.30$

7.2 Moments and properties

<u>Definition</u>. Expected value:

Q: A fair six-sided die is tossed. You win \$2 if the result is a "1," you win \$1 if the result is a "6," but otherwise you lose \$1. Find E(X).

Hint: Let X = Amount Won or Lost.

 \mathcal{A} :

Here are some key properties of the expected value:

Definition. Let X be a random variable and let $k \geq 1$. If $E[X^k] < \infty$, then $E[X^k]$ is called the k - th moment of X.

For the random variable X the **first moment** is given by:

The **second moment** is:

Some may be familiar with the formula for the variance.

The variance of the random variable X is a measure of the expected distance of X from its mean, given by: $var[X] = \sigma^2 \equiv E[(X - \mu)^2]$. By definition, the variance becomes a random variable where $g(X) = (X - \mu)$. Thus, the variance can be described as the second moment of (X - E[X]).

Properties of variance

The standard deviation of a random variable, denoted sd[X], is the positive square root of the variance:

$$sd[X] = \sigma = +\sqrt{var[X]}$$

Properties of standard deviation

- (i) For any constant c, sd[c] = 0
- (ii) For any constants a and b, $sd[aX + b] = |a| \cdot sd[X]$

Here are couple of definitions that one can keep handy which we will not further expand on. In many statistical analyses the task is to really characterize the location and variability of a data set. A further characterization of the data includes skewness and kurtosis.

Skewness is a measure of symmetry, or more precisely, the lack of symmetry. A distribution, or data set, is symmetric if it looks the same to the left and right of the center point.

Kurtosis is a measure of whether the data are heavy-tailed or light-tailed relative to a normal distribution. That is, data sets with high kurtosis tend to have heavy tails, or outliers. Data sets with low kurtosis tend to have light tails, or lack of outliers.

Accordingly, the third and fourth moments of the the random variable $(X - \mu)$ can also be used to describe the distribution. We have:

Skewness:

Kurtosis:

These coefficients are the respective descriptive statistics for these moments.

Example. If you have the following data:

X	3	4	5
f(x)	0.3	0.4	0.3

We want to find the variance and standard deviation of X.

x	1	2	6	8
f(x)	0.4	0.1	0.3	0.2

Q:	What i	is the	variance	and	stand	ard d	leviati	on of	X5

<u>Definition</u>. Covariance:

Properties of covariance

Another statistic you may be interested in is the correlation coefficient. Both covariance and correlation indicate whether variables are positively or inversely related.

<u>Definition</u>. **Correlation**:

Properties of correlation:

Example. Suppose X and Y have the following joint probability function, let's find the covariance and correlation between X and Y.

			y				
	f(x,y)	1	2	3	$f_X(x)$		
X	1	0.25	0.25	0	0.5		
	2	0	0.25	0.25	0.5		
$f_{Y}(y)$		0.25	0.5	0.25	1		

7.4 Conditional expectation

The conditional expectation of Y given X computes the *expected value* of the random variable Y, given that we know X = x:

$$E[Y|X = x] = \begin{cases} \sum_{j=1}^{m} y_j f_{Y|X}(y_j|x) \\ \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy \end{cases}$$

The top is for discrete and bottom for continuous Y. This conditional expectation tells us how the expected value of Y varies with X.

Properties of conditional expectations

The **conditional variance** is given by:

$$var[Y | X = x] = E[Y^2 | x] - [E[Y | x]]^2$$

If X and Y are independent, however, var[Y | X] = var[Y]

Definition. A statistic:

Covariance:

7.5 Sampling distribution

A sample of n observations, x_1, x_2, \ldots, x_n , is a random sample if the n observations are drawn independently from the same population or probability density function. The sample is called univariate if x_i is a single random variable and is multivariate if x_i contains several variables.

The most common descriptive statistics are:	
Mean:	
Median:	
Sample midrange:	
Standard deviation:	
Standard error:	
Skewness:	
Kurtosis:	
Covariance:	

One philosophical foundation of our approach in thinking about estimation and hypothesis testing is that sample data, say, X_1, X_2, \ldots, X_n , of a sample of size n, are thought of as a (subset of an infinite) collection of independent, identically distributed (i.i.d.) random variables, following the probability distribution in question (for example, say a normal distribution). This does not always have to be true, sometimes things are not that nice. A good example is a succession of throws of a fair coin: The coin has no memory, so all the throws are "independent". And every throw is 50:50 (heads, tails), so the coin is and stays fair - the distribution from which every throw is drawn, so to speak, is and stays the same: "identically distributed". This a good example, but it

should be not confused to mean equi-probable, i.e., all outcomes are equally likely. This is more related to this example.

Let θ be some unknown parameter. An estimator is then a rule or strategy that we use to estimate the parameter θ given our random sample of data. Let's denote our estimate as $\hat{\theta}$.

<u>Definition</u>. **Unbiasedness**:

<u>Definition</u>. **Efficient**:

Example: Show that \bar{x} is an unbiased estimators for μ .

We need to show that $E[\bar{x}] = \mu$

7.5.1. Brief notes on common probability distributions:

Normal distribution:

Although I won't do a proof here, when standardized to variable $z = \frac{X-\mu}{\sigma}$, we have E[Z] = 0 and var(Z) = 1. As we saw earlier, the normal distribution is always a symmetric distribution. Then $Z \sim N[0,1]$

We will often work with normal distributions and we further define the Chi-squared, t and F-distribution.

Let Z_i for i = 1, 2, ..., n be independent random variables each distributed as standard normal. A chi-square distribution with n degrees of freedom is given as:

$$X = \sum_{i=1}^{n} Z_i^2 \sim \mathcal{X}_n^2$$

Let $Z \sim N[0,1]$ and $X \sim X_n^2$ be independent random variables, then the ratio:

 $T = \frac{Z}{\sqrt{\frac{X}{n}}} \sim t_n$. has the t-distribution with n degrees of freedom.

Let $X_1 \sim X_{n1}^2$ and $X_2 \sim X_{n2}^2$ be independent random variables, then the ratio:

$$F = \frac{X_1/n_1}{X_2/n_2} \sim F_{n_1,n_2}$$

has the F-distribution with n_1 and n_2 degrees of freedom. We say n_1 is the numerator degrees of freedom and n_2 is the denominator degrees of freedom.

You may also come across the following distributions but going through those is outside the scope of this class.

Lognormal distribution

The lognormal distribution is convenient for modeling size distributions, such as the distribution of firm sizes or distribution of income. It is given as:

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma x} e^{-\frac{1}{2}(\frac{(\ln x - \mu)^2}{\sigma^2})}$$
, for $x > 0$

Gamma distribution

The Gamma distribution can also be used for the study of income distribution, we have:

$$f(x) = \frac{\lambda^P}{\Gamma(P)} e^{-\lambda x} x^{P-1}, \text{ for } x \ge 0, \lambda > 0, P > 0$$

 $\Gamma(P)$ is the gamma function which is equal to $\int_0^\infty t^{P-1}e^{-t}dt$

Logistic distribution

The logistic distribution is similar to the normal distribution but allows for more "thickness" in the tails of the distribution. It is given by:

Discrete random variables

All the distributions above apply to continuous random variables. For discrete random variables, we have two main (most common) distributions. These distributions are based on binary random variables. The **binomial distribution** is given by:

If the number of trails n becomes large and the mean is stable, then the binomial distribution converges to the **Poisson distribution**:

7.6. Hypothesis testing – very basic

Why do we care about hypothesis testing?

So far, we assumed to know the parameter values and investigated the properties of the distribution (i.e. we knew that μ and σ pretty straightforwardly). However, knowing the true values of the parameters is not a straightforward task. We can use the data to estimate the parameters, the second is to guess a value for the parameters and ask the data whether this value is true – this is close to a proof by contradiction in essence. The former approach is **estimation** and the latter is **hypothesis testing.**

Hypothesis testing will let us make decisions about specific values of parameters or relationships between parameters.

In hypothesis testing, we use sample data to choose between two competing hypotheses. There are 2 choices, the null hypothesis and the alternative hypothesis. We assume the null hypothesis is true until the alternative is shown beyond "chance".

Definition. Type 1 Error:

<u>Definition</u>. Type II Error:

More simply stated, a type I error is detecting an effect that is not present, while a type II error is failing to detect an effect that is present.

Power of a test: power = $1 - \beta = 1 - P$ [type II error], β is simply the probability of making a type II error.

In any hypothesis test, you need to balance α and β . You can never get them both to be 0 with a sample. In practice, we choose the largest α that is tolerable - usually .01, .05 or .10. The precise relationship between α and β cannot be easily determined. That is to say, when α increases, β decreases (and power would increase), and vice versa. It hard to say by how much. Power really gets at the probability of correctly rejecting the null hypothesis in favor of the alternative when the alternative is true. Ideally, you want power to be as large (close to 1) as possible. There are 2 quantities that you have control over that can help you increase power.

To increase power, you can:

- 1. Increase sample size, n
- 2. Increase α , the significance level. (Note you still don't want to set this too high).

To begin, a hypothesis is a claim about a population characteristic (parameter). A hypothesis test is set up as follows:

- (1) State the null hypothesis
- (2) State the alternative hypothesis
- (3) Decide on the significance level α of the test.

Obtain the critical value for the test.

- (4) Calculate the test statistic.
- (5) If the test statistic falls in the rejection region then reject the null hypothesis. If the test statistic does not fall in the rejection region then we fail to reject the null hypothesis.

Assume that the sampled data follows a normal distribution and use a t-statistic to approximate this data. The cases for hypothesis testing can be summarized as follows:

H_0	H_1	Test	Rejection area
$\mu = \mu_0$	µ ≠ μ ₀	Two-tailed	$ t = \left \frac{\bar{\theta} - \mu_0}{\frac{S_{\theta}}{\sqrt{n}}} \right > C_{1 - \frac{\alpha}{2}}$
$\mu = \mu_0$	$\mu > \mu_0$	One-tailed	$t = \frac{\bar{\theta} - \mu_0}{s_\theta / \sqrt{n}} > C_{1-\alpha}$
$\mu = \mu_0$	$\mu < \mu_0$	One-tailed	$t = \frac{\bar{\theta} - \mu_0}{s_\theta / \sqrt{n}} < -C_{1-\alpha}$

 θ is a parameter and s_{θ} is the standard deviation of that parameter. Note that $C_{1-\frac{\alpha}{2}}$ denotes the critical value for the corresponding distribution and confidence level.

We will do one quick example and get to the next section.

A principal at a certain school claims that the students in his school are above average intelligence. A random sample of thirty students IQ scores have a mean score of 112. Is there sufficient evidence to support the principal's claim? The mean population IQ is 100 with a standard deviation of 15.

State the null hypothesis:

State the alternate hypothesis:

Is this a one tailed or two tailed test?

What would be an appropriate rejection area if you are given an $\alpha = 0.05$.

What is the t test statistic?

If the test statistic is greater than the z-score, we reject the null.

7.1 Brief review of large sample distribution theory

After some thought, I decided to do only a brief review of this section. I've been convinced that this part is nicely and thoroughly covered by Prof. McCullough, so I will let him take you on this ride. You will learn much more about this in class — so do not worry too much about this being only a brief introduction.

To check larger sample properties we use probability limits.

Definition: PLIM

7.6.1 Large numbers

(Khinchine's Weak Law of Large Numbers) Theorem.

This law of large numbers states that the mean of a random sample is always a consistent estimate of the true population average. In simple terms, we can get arbitrarily close to estimating the true mean by using a large sample.

Properties of probability limits

You will use these rules to show that Ordinary Least Squares (OLS) estimate is consistent for example.

Example: Let X be a binary random variable that follows a binomial distribution where x=1 denotes a success and x=0 denotes a failure. Assume that $P[X=1]=\alpha$. Let x_1,x_2,\ldots,x_n be a random sample of n observations. Is the estimate $G=\frac{\bar{x}}{(1-\bar{x})}$ an unbiased estimator for $\lambda=\frac{\alpha}{1-\alpha}$? Is G consistent?