

Linear Regression-1

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Characteristic Function

The characteristic function of a random variable is

$$\phi_x(t) \equiv E(e^{itx}) = \int_{-\infty}^{\infty} e^{itx} f_x(x) dx$$

$$e^{itx} = \frac{1}{\mathcal{L}_0} + \frac{itx}{\mathcal{L}_1} + \frac{(it)^2 x^2}{\mathcal{L}_2} + \frac{(it)^3 x^3}{\mathcal{L}_3} + \dots$$

$$\therefore \phi_x(t) = E \left[\frac{1}{\mathcal{L}_0} + \frac{itx}{\mathcal{L}_1} + \frac{(it)^2 x^2}{\mathcal{L}_2} + \frac{(it)^3 x^3}{\mathcal{L}_3} + \dots \right]$$

$$\Rightarrow \phi_x(t) = 1 + it E(x) + \frac{(it)^2}{\mathcal{L}_2} E(x^2) + \frac{(it)^3}{\mathcal{L}_3} E(x^3) + \dots$$

Characteristic Function

$$\therefore \phi_X(t) = \frac{1}{L^0} + \frac{it m_1}{L^1} + \frac{(it)^2 m_2}{L^2} + \frac{(it)^3 m_3}{L^3} + \dots$$

where m_n is the n^{th} moment of the r.v.

i.e. $m_n = E(X^n)$

$$\therefore \phi_X(t) \Big|_{t=0} = 1 ; \quad \frac{d}{dt} \phi_X(t) \Big|_{t=0} = i m_1 ; \quad \frac{d^n}{dt^n} \phi_X(t) \Big|_{t=0} = (i)^n m_n$$

Moment generating Function

The moment generation function of a r.v. is

$$M_X(t) = \phi_X(-it) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$$

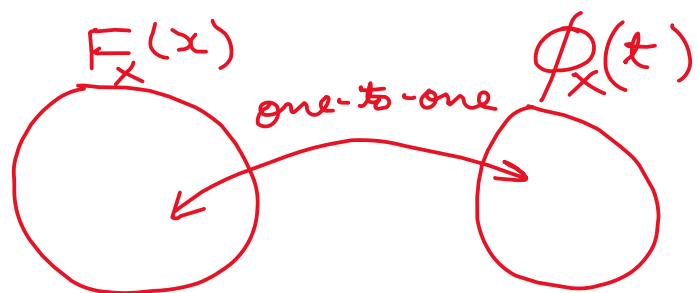
\therefore

$$\left. \frac{d^n}{dt^n} \phi_X(t) \right|_{t=0} = (i)^n m_n \Rightarrow \frac{d^n}{dt^n} M_X(t) = m_n = n^{\text{th}} \text{ moment}$$

Characteristic Function

$$\therefore \phi_X(t) = \frac{1}{\mathcal{L}^0} + \frac{it m_1}{\mathcal{L}^1} + \frac{(it)^2 m_2}{\mathcal{L}^2} + \frac{(it)^3 m_3}{\mathcal{L}^3} + \dots$$

There is a one-to-one correspondence between the cumulative distribution function and the characteristic function.



If the r.v. has a probability density function $f_X(x)$ then

$$f_X(x) = F'_X(x) = \frac{1}{2\pi} \int e^{-itx} \phi_X(t) dt$$

Characteristic Function

$$\therefore \phi_X(t) = \frac{1}{L^0} + \frac{it m_1}{L^1} + \frac{(it)^2 m_2}{L^2} + \frac{(it)^3 m_3}{L^3} + \dots$$

If a r.v. X has $\mu=0$ and $\sigma^2=1$ i.e. $X \sim (0, 1)$

then $\phi_X(t) = 1 + 0 - \frac{t^2}{2} + o(t^2)$

For normal distribution $N(\mu, \sigma^2)$

$$\phi_X(t) = e^{it\mu - \frac{1}{2}\sigma^2 t^2}$$

and for $N(0, 1)$, $\phi_X(t) = e^{-\frac{t^2}{2}}$

$$\left[\begin{array}{l} \text{For } N(\mu, \sigma^2) \\ f_X = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \\ \text{For } N(0, 1) \\ f_X = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \end{array} \right]$$

Central limit theorem

$$X_i \stackrel{iid}{\sim} (\mu, \sigma^2)$$

Sample mean $\bar{X}_n \equiv \frac{1}{n} \sum X_i$

\therefore Expected value of sample mean is

$$E(\bar{X}_n) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} n \mu = \mu$$

Central limit theorem

$$\left[\begin{array}{l} X_i \stackrel{iid}{\sim} (\mu, \sigma^2) \\ \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \end{array} \right]$$

Variance of sample mean $\text{Var}(\bar{X}_n)$

$$\text{Var}(\bar{X}_n) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{1}{n^2} n \sigma^2 = \frac{\sigma^2}{n}$$

\therefore for $X_i \stackrel{iid}{\sim} (\mu, \sigma^2)$

Mean of Sample mean $E(\bar{X}_n) = \mu$

and variance of sample mean $\text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$

Central limit theorem

Now we define $Z_n = \frac{n \bar{X}_n - n\mu}{\sigma \sqrt{n}}$

$$\Rightarrow Z_n = \frac{n \frac{1}{n} \sum_{i=1}^n X_i - n\mu}{\sigma \sqrt{n}} = \frac{\sum_{i=1}^n X_i - n\mu}{\sigma \sqrt{n}}$$

$$\Rightarrow Z_n = \frac{\sum_{i=1}^n (X_i - \mu)}{\sigma \sqrt{n}} = \frac{\sqrt{n} (\bar{X}_n - \mu)}{\sigma}$$

If now we define

$$Y_i = \frac{X_i - \mu}{\sigma}, \text{ then}$$

$$Z_n = \sum_{i=1}^n \frac{Y_i}{\sqrt{n}}$$

Central limit theorem

$$\therefore Y_i = \frac{X_i - \mu}{\sigma} ; E(Y_i) = 0 \text{ and}$$

$$\text{Var}(Y_i) = \frac{\text{Var}(X_i)}{\sigma^2} = \frac{\sigma^2}{\sigma^2} = 1$$

$$\therefore \phi_Y(t) = 1 - \frac{t^2}{2} + o(t^2)$$

$$\therefore \phi_{Z_n}(t) = E\left(e^{it \frac{(Y_1 + Y_2 + \dots + Y_n)}{\sqrt{n}}}\right)$$

$$= E\left(\prod_{k=1}^n e^{i \frac{t}{\sqrt{n}} Y_k}\right)$$

$$= \prod_{k=1}^n E\left(e^{i \frac{t}{\sqrt{n}} Y_k}\right) = \left[\phi_Y\left(t/\sqrt{n}\right)\right]^n$$

$$Z_n = \sum_{i=1}^n \frac{Y_i}{\sqrt{n}}$$

Central limit theorem

$$\therefore \phi_{Z_n}(t) = \left[\phi_Y(t/\sqrt{n}) \right]^n = \left[1 - \frac{t^2}{2n} + o\left(\frac{t^2}{n}\right) \right]^n$$

As we increase the sample size n , we get the limit

$$\begin{aligned} \lim_{n \rightarrow \infty} \phi_{Z_n}(t) &= \lim_{n \rightarrow \infty} \left[1 - \frac{t^2}{2n} + o\left(\frac{t^2}{n}\right) \right]^n \\ &= e^{-t^2/2} \end{aligned}$$

This is same as the characteristic function for $N(0,1)$.

Hence, $\lim_{n \rightarrow \infty} Z_n = N(0,1)$

Central limit theorem

$$\therefore \lim_{n \rightarrow \infty} Z_n = N(0, 1)$$

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{\sqrt{n} (\bar{X}_n - \mu)}{\sigma} = N(0, 1)$$

$$\Rightarrow \lim_{n \rightarrow \infty} \sqrt{n} (\bar{X}_n - \mu) = N(0, \sigma^2)$$

$$\Rightarrow \lim_{n \rightarrow \infty} (\bar{X}_n - \mu) = N\left(0, \frac{\sigma^2}{n}\right)$$

$$\Rightarrow \lim_{n \rightarrow \infty} \bar{X}_n = \mu + N\left(0, \frac{\sigma^2}{n}\right)$$

Central limit theorem

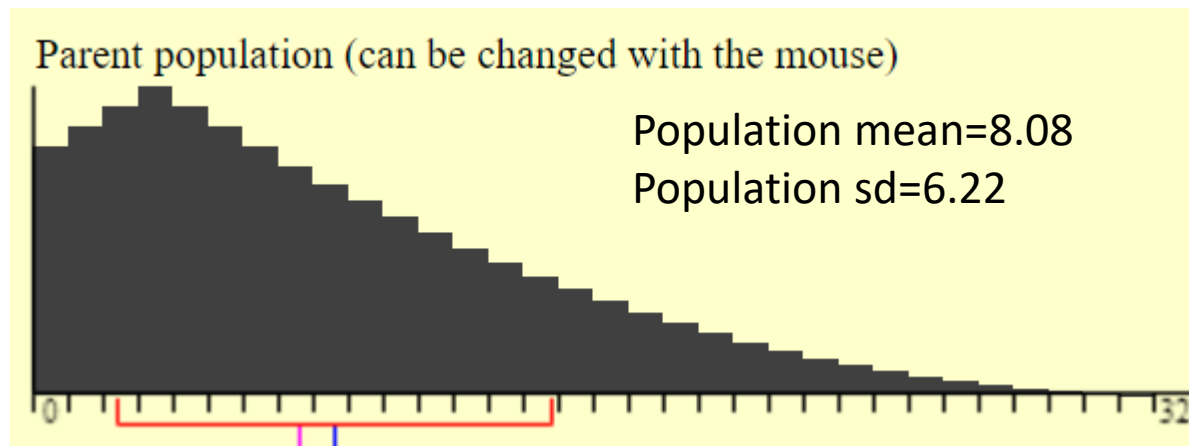
For $X_i \stackrel{iid}{\sim} (\mu, \sigma^2)$ if we define sample mean $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$

Then mean of sample mean $E(\bar{X}_n) = \mu$
and variance of sample mean $\text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$

Now by central limit theorem we get that

$$\therefore \lim_{n \rightarrow \infty} \bar{X}_n = N\left(\mu, \frac{\sigma^2}{n}\right)$$

Central limit theorem

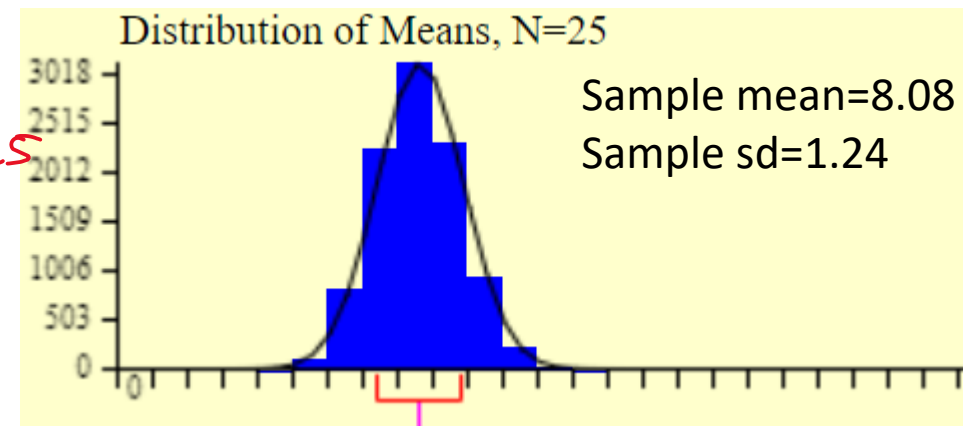
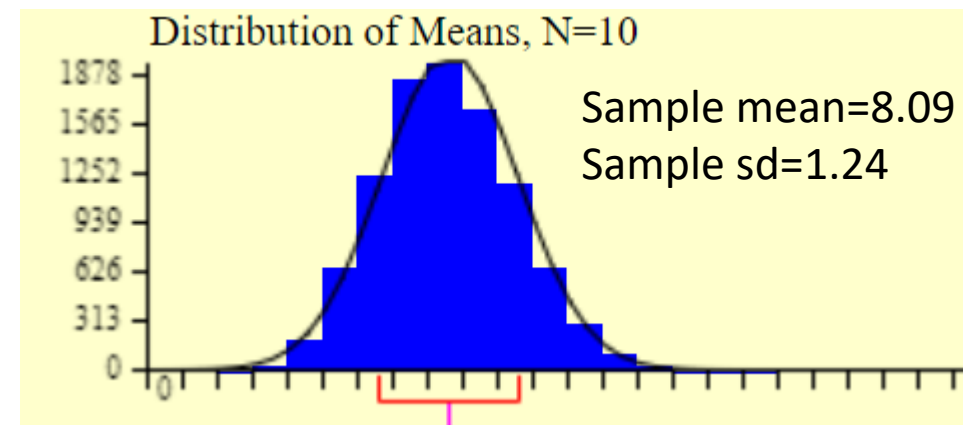
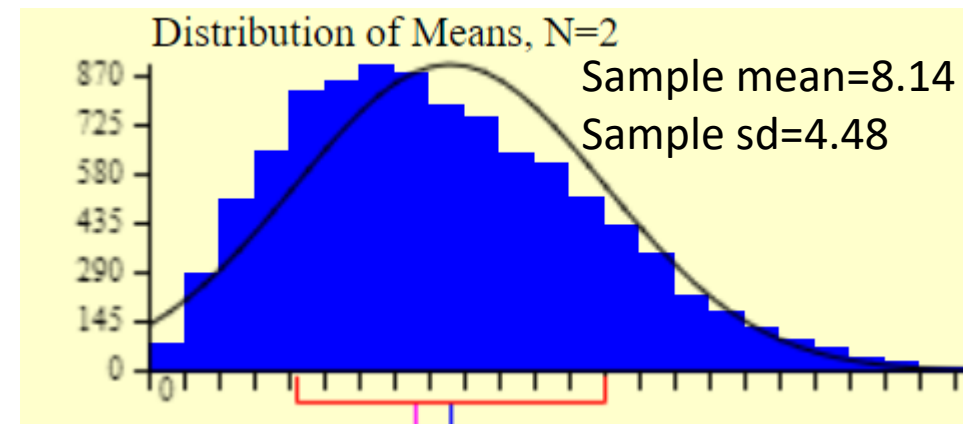


Sample size = 2

Sample size = 10

Sample size = 25

As $n \uparrow$ the distribution of sample mean approaches a normal distribution.



Ref: http://onlinestatbook.com/stat_sim/sampling_dist/index.html

Application of Central limit theorem

From a population of $X_i \overset{iid}{\sim} (\mu, \sigma^2)$ we draw a sample of size n

$$\text{Population mean } \mu = E(\text{sample mean}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right)$$

$$\text{and population variance } \sigma^2 = E(\text{sample variance}) = E\left(\frac{1}{n-1} \sum_{i=1}^n (\bar{X}_n - X_i)^2\right)$$

Now by central limit theorem

$$\lim_{n \rightarrow \infty} \bar{X}_n = N\left(\mu, \frac{\sigma^2}{n}\right)$$

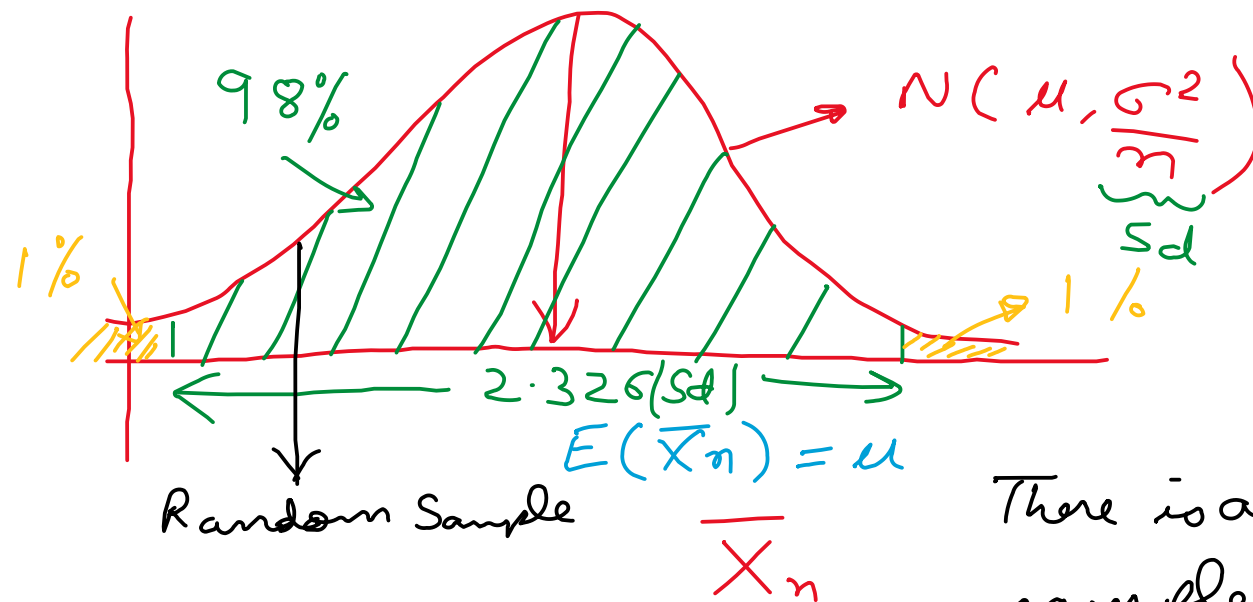
→ An unbiased estimate of σ^2 is s^2 i.e. sample variance

→ An unbiased estimate of this is \bar{X}_n i.e. sample mean

Application of Central limit theorem

$X_i \text{ iid } (\mu, \sigma^2)$ If we take a sample of size n then $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$

Then for large n , CLT $\Rightarrow \bar{X}_n \rightarrow N\left(\mu, \frac{\sigma^2}{n}\right)$



| C | Z^x |
|-----|-------|
| 99% | 2.576 |
| 98% | 2.326 |
| 95% | 1.96 |

There is a 98% probability that a random sample is within $\pm 2.326(Sd)$ of μ

\therefore we can say with 98% confidence that $\mu = \bar{X}_n \pm 2.326 Sd$

Application of Central limit theorem

$X_i \text{ iid } (\mu, \sigma^2)$ If we take a sample of size n then $\bar{X}_n = \sum_{i=1}^n \frac{x_i}{n}$

Then for large n , CLT $\Rightarrow \bar{X}_n \rightarrow N(\mu, \frac{\sigma^2}{n})$

| C | Z* |
|-----|-------|
| 99% | 2.576 |
| 98% | 2.326 |
| 95% | 1.96 |

Therefore we can say with C confidence

that $\mu = \bar{X}_n \pm Z^*(C) \underbrace{\left(\frac{\sigma}{\sqrt{n}} \right)}_{\substack{\text{Sample} \\ \text{variance} \\ \downarrow \\ = \frac{E(S)}{\sqrt{n}}}}$

$$\therefore \mu = \frac{1}{n} \sum_{i=1}^n x_i \pm Z^*(C) \frac{\frac{1}{n-1} \sum (\bar{X}_n - x_i)^2}{\sqrt{n}}$$

t-distribution

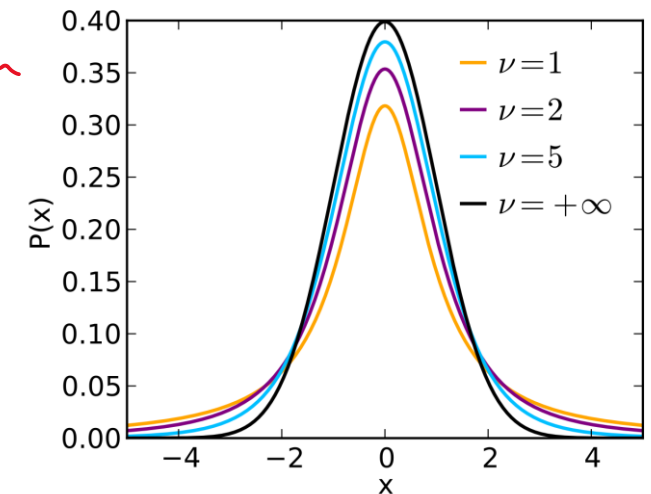
If $X_i \stackrel{iid}{\sim} N(\mu, \sigma^2)$ and $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$, then we can define

$$Z_n \equiv \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \text{ and by CLT as } n \rightarrow \infty \quad Z_n \rightarrow N(0, 1)$$

Similarly, it can be shown that if $t \equiv \frac{\bar{X}_n - \mu}{S/\sqrt{n}}$ P.S: As $n \rightarrow \infty$
t-dist $\rightarrow N(0, 1)$

then the r.v. t follows t-distribution
with $\nu = n - 1$ degrees of freedom

$$\text{pdf for } t\text{-distribution} = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\nu\pi} \Gamma\left(\frac{\nu}{2}\right)} \left(1 + \frac{x^2}{\nu}\right)^{-\frac{\nu+1}{2}}$$



Confidence Interval

CLT based: If $X_i \overset{iid}{\sim} (\mu, \sigma^2)$

then
$$\mu = \bar{X} \pm Z^*(C) \frac{\sigma}{\sqrt{n}}$$

Sample variance S is an unbiased estimate of σ

| C | Z^* |
|-----|-------|
| 99% | 2.676 |
| 98% | 2.326 |
| 95% | 1.96 |

t-distribution: If $X_i \overset{iid}{\sim} N(\mu, \sigma^2)$

then
$$\mu = \bar{X} \pm t^*(C) \frac{S}{\sqrt{n}}$$

For 95% C

| n | n-1 | t^* |
|----------|----------|-------|
| 6 | 5 | 2.571 |
| 11 | 10 | 2.228 |
| 31 | 30 | 2.042 |
| ∞ | ∞ | 1.960 |

If $n > 30$ use CLT, else use t-distribution