

## Quiz-4 answers and solutions

Coursera. Stochastic Processes

December 30, 2020

1. Choose the functions  $K(t, s)$ , which can be covariance functions of some stochastic processes defined for  $t \in [0, 1]$ .

**Options:**

- a)  $K(t, s) = 1 - \max\{t, s\}$       b)  $K(t, s) = \max\{t, s\} - \min\{t, s\}$   
c)  $K(t, s) = 1 + (t - s)^3$       d)  $c > 0$

**Answer:** a)  $K(t, s) = 1 - \max\{t, s\}$ , d)  $c > 0$

**Solution:**

a) Clearly, this function is symmetric, since

$$K(t, s) = 1 - \max\{t, s\} = 1 - \max\{s, t\} = K(s, t).$$

Let us show that it is also positive semi-definite. Indeed, define  $g_t(x) := \mathbb{I}\{x \in [t, 1]\}$ . Then

$$1 - \max\{t, s\} = \int_0^\infty g_s(x) g_t(x) dx,$$

from which

$$\begin{aligned} \sum_{j=1}^n \sum_{k=1}^n u_j u_k K(t_j, t_k) &= \sum_{j=1}^n \sum_{k=1}^n u_j u_k \int_0^\infty g_{t_j}(x) g_{t_k}(x) dx \\ &= \int_0^\infty \sum_{j=1}^n \sum_{k=1}^n u_j u_k g_{t_j}(x) g_{t_k}(x) dx \\ &= \int_0^\infty \left( \sum_{j=1}^n u_j g_{t_j}(x) \right) \left( \sum_{k=1}^n u_k g_{t_k}(x) \right) dx \\ &= \int_0^\infty \left( \sum_{j=1}^n u_j g_{t_j}(x) \right)^2 dx \geq 0. \end{aligned}$$

Thus,  $K(t, s) = 1 - \max\{t, s\}$  can be a covariance function of some process defined for  $t \in [0, 1]$ .

b) This function cannot be a covariance function of any process. Assume the converse: let there exist some process  $X_t$  such that its covariance function  $K(t, s) = \max\{t, s\} - \min\{t, s\}$ . Then

$$\text{Var } X_t = K(t, t) = \max\{t, t\} - \min\{t, t\} = t - t = 0,$$

i.e.,  $X_t = \text{const}$  for all  $t \in [0, 1]$ . However,

$$\text{cov}(X_t, X_s) = \max\{t, s\} - \min\{t, s\} \neq \text{const} \quad \forall s \neq t, \quad (s, t) \in [0, 1]^2,$$

which leads to the contradiction.

c) Since

$$K(t, s) = 1 + (t - s)^3 \neq 1 + (s - t)^3 = K(s, t),$$

the symmetry property is violated, and this function cannot be a covariance function of any process.

d) It can be seen that

$$K(t, s) = c = K(s, t)$$

and

$$\sum_{j=1}^n \sum_{k=1}^n u_j u_k K(t_j, t_k) = \sum_{j=1}^n \sum_{k=1}^n u_j u_k c = c \left( \sum_{j=1}^n u_j \right)^2 \geq 0,$$

meaning that this function is symmetric and positive semi-definite. Thus, it can be a covariance function of some process. A corresponding example is given by  $X_t = \xi$ , where  $\xi$  is a random variable such that  $\text{Var } \xi = 1$ .

2. Which of the following processes are Gaussian? (In the answers below  $W_t$  is a Brownian motion).

**Options:**

- a)  $X_t = W_t^2$   
b)  $X_t = a\xi_{t-1} + \xi_t, \quad \xi_i \sim \text{i.i.d. } \mathcal{N}(0, 1) \quad \forall i = 0, 1, \dots, t, \quad t = 1, 2, \dots$   
c)  $X_t = W_{1-t} - W_1, \quad t \in [0, 1]$   
d)  $X_t = W_t + \eta\xi, \quad \xi \sim \mathcal{N}(0, 1), \eta = \begin{cases} -1, & 1/2 \\ 1, & 1/2 \end{cases}, \quad W_t, \xi, \eta \text{ are independent for all } t \geq 0$

**Answer:** b), c), d)

**Options:** The process  $X_t$  is Gaussian if and only if for any  $t_1, \dots, t_n$ :  $(X_{t_1}, \dots, X_{t_n})$  is a Gaussian vector, that is,  $\sum_{k=1}^n \lambda_k X_{t_k} \sim \mathcal{N}, \forall \vec{\lambda}$ .

a) Since  $\mathbb{P}\{W_t^2 < 0\} = 0$ ,  $W_t^2$  is not normally distributed, and this process is not Gaussian.

b) A finite dimensional distribution for this process has the form

$$(X_{t_1}, X_{t_2}, \dots, X_{t_n}) = (a\xi_{t_1-1} + \xi_{t_1}, a\xi_{t_2-1} + \xi_{t_2}, \dots, a\xi_{t_n-1} + \xi_{t_n}).$$

Clearly, any linear combination of finite dimensional distributions of  $X_t$  can be represented as a sum of i.i.d. normally distributed random variables: for instance, if  $t_i$  are such that  $t_i - 1 = t_{i-1} \quad \forall i = 1, 2, \dots, n$

$$\begin{aligned} \sum_{k=1}^n \lambda_k X_{t_k} &= \lambda_1(a\xi_{t_1-1} + \xi_{t_1}) + \lambda_2(a\xi_{t_2-1} + \xi_{t_2}) + \dots + \lambda_n(a\xi_{t_n-1} + \xi_{t_n}) \\ &= \lambda_1(a\xi_{t_1-1} + \xi_{t_1}) + \lambda_2(a\xi_{t_1} + \xi_{t_2}) + \dots + \lambda_n(a\xi_{t_{n-1}} + \xi_{t_n}) \\ &= \lambda_1 a\xi_{t_1-1} + (\lambda_1 + a\lambda_2)\xi_{t_1} + (\lambda_2 + a\lambda_3)\xi_{t_1} + \dots + \lambda_n \xi_{t_n} \sim \mathcal{N}. \end{aligned}$$

c) This process is Gaussian as

$$\begin{aligned}\sum_{k=1}^n \lambda_k X_{t_k} &= \lambda_1 W_{1-t_1} - \lambda_1 W_1 + \dots + \lambda_n W_{1-t_n} - \lambda_n W_1 \\ &= \sum_{k=1}^n \lambda_k W_{1-t_k} - W_1 \sum_{k=1}^n \lambda_k \sim \mathcal{N}\end{aligned}$$

as a linear combination of the components of the finite dimensional distribution  $(W_{1-t_1}, W_{1-t_2}, \dots, W_{1-t_n}, W_1)$  of a Brownian motion which is a Gaussian process.

d) Let us show that  $\eta\xi \sim \mathcal{N}(0, 1)$ . Indeed,

$$\begin{aligned}\mathbb{P}\{\eta\xi \leq x\} &= \mathbb{P}\{-\xi \leq x|\eta = -1\}\mathbb{P}\{\eta = -1\} + \mathbb{P}\{\xi \leq x|\eta = 1\}\mathbb{P}\{\eta = 1\} \\ &= \frac{1}{2}(1 - \mathbb{P}\{\xi > x\} + \mathbb{P}\{\xi \leq x\}) = \frac{1}{2}(1 - 1 + \mathbb{P}\{\xi \leq x\} + \mathbb{P}\{\xi \leq x\}) \\ &= \mathbb{P}\{\xi \leq x\}.\end{aligned}$$

Thus

$$\sum_{k=1}^n \lambda_k X_{t_k} = \sum_{k=1}^n \lambda_k (W_{t_k} + \eta\xi) = \sum_{k=1}^n \tilde{\lambda}_k (W_{t_k} - W_{t_{k-1}}) + \eta\xi \sum_{k=1}^n \lambda_k \sim \mathcal{N}$$

as a sum of i.i.d. normally distributed random variables.

3. Which of the following processes is a Brownian motion? (In the answers below  $W_t$  is a Brownian motion).

**Options:**

- a)  $X_t = W_t^2$
- b)  $X_t = a\xi_{t-1} + \xi_t, \quad \xi_i \sim i.i.d. \mathcal{N}(0, 1) \quad \forall i = 0, 1, \dots, t, \quad t = 1, 2, \dots$
- c)  $X_t = W_{1-t} - W_1, \quad t \in [0, 1]$
- d)  $X_t = W_t + \eta\xi, \quad \xi \sim \mathcal{N}(0, 1), \eta = \begin{cases} -1, & 1/2 \\ 1, & 1/2 \end{cases}, \quad W_t, \xi, \eta \text{ are independent for all } t \geq 0$

**Answer:** c)  $X_t = W_{1-t} - W_1, \quad t \in [0, 1]$

**Solution:** Since Brownian motion is a Gaussian process  $W_t$  such that  $\mathbb{E}W_t = 0$  and  $\text{cov}(W_t, W_s) = \min\{t, s\}$ , we should only check the last two properties for the processes b)-d).

For b)

$$\mathbb{E}[X_t] = \mathbb{E}[a\xi_{t-1} + \xi_t] = a\mathbb{E}[\xi_{t-1}] + \mathbb{E}[\xi_t] = 0,$$

but

$$\text{cov}(X_t, X_s) = \text{cov}(a\xi_{t-1} + \xi_t, a\xi_{s-1} + \xi_s) = (1+a^2)\mathbb{I}\{t = s\} + a\mathbb{I}\{|t-s| = 1\},$$

thus, it is not a Brownian motion.

For c)

$$\mathbb{E}[W_{1-t} - W_1] = \mathbb{E}[W_{1-t}] - \mathbb{E}[W_1] = 0$$

and

$$\begin{aligned}
\text{cov}(W_{1-t} - W_1, W_{1-s} - W_1) &= \text{cov}(W_{1-t}, W_{1-s}) - \text{cov}(W_1, W_{1-t}) \\
&\quad - \text{cov}(W_1, W_{1-s}) + \text{cov}(W_1, W_1) \\
&= \min\{1-t, 1-s\} - (1-t) - (1-s) + 1 \\
&= t + s - \max\{t, s\} = \min\{t, s\}.
\end{aligned}$$

Therefore,  $X_t$  is Brownian motion.

For d)

$$\mathbb{E}[W_t + \eta\xi] = \mathbb{E}[W_t] + E[\eta]E[\xi] = 0,$$

but

$$\begin{aligned}
\text{cov}(W_t + \eta\xi, W_s + \eta\xi) &= \text{cov}(W_t, W_s) + \text{cov}(W_t, \eta\xi) + \text{cov}(W_s, \eta\xi) + \text{cov}(\eta\xi, \eta\xi) \\
&= \min\{t, s\} + 0 + 0 + \mathbb{E}[(\eta\xi)^2] - (\mathbb{E}[\eta\xi])^2 \\
&\stackrel{\eta\xi \sim \mathcal{N}(0,1)}{=} \min\{t, s\} + 1 - 0 \\
&= \min\{t, s\} + 1 \neq \min\{t, s\}.
\end{aligned}$$

So,  $X_t$  is not a Brownian motion.

4. Find the covariance function of the process  $X_t = W_t - tW_1$ ,  $t \in [0, 1]$ , where  $W_t$  is a Brownian motion.

**Answer:**  $\min\{t, s\} - ts$

**Solution:**

$$\begin{aligned}
\text{cov}(X_t, X_s) &= \text{cov}(W_t - tW_1, W_s - sW_1) \\
&= \text{cov}(W_t, W_s) - t \text{cov}(W_1, W_s) - s \text{cov}(W_t, W_1) + ts \text{cov}(W_1, W_1) \\
&= \min\{t, s\} - ts - st + ts = \min\{t, s\} - ts.
\end{aligned}$$

5. Find the covariance function of the process  $X_t = \xi W_t$ ,  $\xi \sim \mathcal{N}(1, 1)$  is independent of  $W_t \quad \forall t, t \geq 0$ .

**Answer:**  $2 \min\{t, s\}$ .

**Solution:**

$$\begin{aligned}
\text{cov}(\xi W_t, \xi W_s) &= \mathbb{E}[\xi^2 W_t W_s] - \mathbb{E}[\xi W_t] \mathbb{E}[\xi W_s] \\
&= \mathbb{E}[\xi^2 \mathbb{E}[W_t W_s]] - (\mathbb{E}[\xi])^2 \mathbb{E}W_t \mathbb{E}W_s \\
&\stackrel{t \geq s}{=} (1 + 1^2) \mathbb{E}[(W_t - W_s + W_s)W_s] - 0^2 \cdot 0 \cdot 0 \\
&= 2(\mathbb{E}[(W_t - W_s)(W_s - W_0)] + \mathbb{E}[W_s^2]) \\
&= 2s = 2 \min\{t, s\}.
\end{aligned}$$

6. Harry is playing some game. At each given time moment  $t$  the number of points he gets is determined by the value of the process  $X_t = \sigma W_t$ , that is, he earns points if  $X_t > 0$  and loses if  $X_t < 0$ . Every hour (at time moments  $t = 1, 2, \dots$ ) the results are automatically recorded, and if the sum of 3 results in a row is greater than 10, Harry gets a special prize. Find the probability that Harry gets this prize after 3 hours of playing (in

the answers below  $\Phi$  is the distribution function of the standard normal distribution).

**Answer:**  $1 - \Phi\left(\frac{10}{\sigma\sqrt{14}}\right)$

**Solution:** We need to find the probability that

$$\begin{aligned}\mathbb{P}\{X_1 + X_2 + X_3 \geq 10\} &= \mathbb{P}\{\sigma(W_1 + W_2 + W_3) \geq 10\} \\ &= 1 - \mathbb{P}\{W_1 + W_2 + W_3 \leq 10/\sigma\}.\end{aligned}$$

As

$$\mathbb{E}[W_1 + W_2 + W_3] = \mathbb{E}[W_1] + \mathbb{E}[W_2] + \mathbb{E}[W_3] = 0$$

and

$$\begin{aligned}\text{Var}(W_1 + W_2 + W_3) &= \text{Var } W_1 + \text{Var } W_2 + \text{Var } W_3 + 2\text{cov}(W_1, W_2) \\ &\quad + 2\text{cov}(W_2, W_3) + 2\text{cov}(W_1, W_3) \\ &= 1 + 2 + 3 + 1 \cdot 2 + 2 \cdot 2 + 1 \cdot 2 = 14,\end{aligned}$$

$$1 - \mathbb{P}\{W_1 + W_2 + W_3 \leq 10/\sigma\} = 1 - \Phi\left(\frac{10}{\sigma\sqrt{14}}\right).$$

7. Choose the correct statements about the process  $X_t = \sigma W_t$ ,  $\sigma > 0$ :

**Options:**

- a)  $\text{cov}(X_t, X_s) = \min\{t, s\}$
- b) the Kolmogorov continuity theorem  $\mathbb{E}[|X_t - X_s|^\alpha] \leq c|t - s|^{1+\beta}$  holds with  $c = 3, \alpha = 4, \beta = 1$
- c)  $X_t$  is not stationary
- d)  $X_t$  is a Gaussian process

**Answer:**  $X_t$  is not stationary and is a Gaussian process.

**Solution:** Since  $X_t$  is a Brownian motion multiplied by a positive constant, it is also a Gaussian process. Now,

$$\text{cov}(X_t, X_s) = \sigma^2 \text{cov}(W_t, W_s) = \sigma^2 \min\{t, s\}$$

and since the covariance function depends not on the difference, but on the time moments  $t$  and  $s$ ,  $X_t$  is not stationary. Finally,

$$\mathbb{E}[|X_t - X_s|^\alpha] = \mathbb{E}[|\sigma(W_t - W_s)|^\alpha] = \sigma^\alpha \mathbb{E}[|W_t - W_s|^\alpha],$$

therefore, the condition of the Kolmogorov continuity theorem

$$\mathbb{E}[|X_t - X_s|^\alpha] \leq c|t - s|^{1+\beta}$$

holds with  $\alpha = 4, \beta = 1$  and  $c = 3\sigma^4$ .