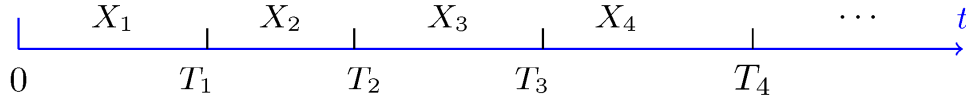


# 1 Poisson Process



- $N(t)$ : The number of arrivals in  $[0, t]$ .
- $N(I)$ : The number of arrivals in  $I \subset [0, \infty)$ .

Characterization

- iid
- $N(b) - N(a) \sim \text{Poisson}((b - a))$   
 $N(s_1), N(s_2) - N(s_1), \dots, N(s_n) - N(s_{n-1})$  are independent

## 1.1 Thinning Property

$N(t)$  is a Poisson process with rate  $\lambda$ . We assign a type  $Y_j \in \{1, 2, \dots, l\}$  for the  $j^{\text{th}}$  arrival. We assume that  $Y_1, Y_2, \dots$ , are iid.

Let  $N_j(t)$  be the number of arrivals with type  $J$  in  $[0, t]$ . Then  $N_j$  is a Poisson process of rate  $P_j \cdot \lambda$ .  $N_1, N_2, \dots$  are independent.

### 1.1.1 Example

Vehicles arrive at a toll booth with a rate of  $2/\text{min}$ . The probability that a given vehicle is a truck is  $2/3$ .

(a)  $P(\text{exactly 2 cars and 3 trucks arriving in the next 5 minutes})$

- Arriving vehicles form a Poisson process with rate  $\lambda = 2$ .
- Let  $N_1$  be the number of arrivals of trucks and  $N_2$  be that of arrivals of cars.
- Then, for trucks, the rate of Poisson process is  $\frac{2}{3}\lambda$  and the rate of Poisson process for cars is  $\frac{1}{3}\lambda$ .
- So  $P(N_1(5) = 3, N_2(5) = 2) = P(N_1(5) = 3)P(N_2(5) = 2) = \frac{(\frac{20}{3})^3}{3!}e^{-\frac{20}{3}} \cdot \frac{(\frac{10}{3})^2}{2!}e^{-\frac{10}{3}}$

(b)  $P(\text{First arrival is a truck and the second one is a car}) = \frac{2}{3} \cdot \frac{1}{3}$

(c) Given that 20 trucks arrived in an hour what is expected the number of cars within the same hour?  $\frac{1}{3} \cdot 60 = 20$

The given 20 trucks actually does not affect the expected number of cars within the same time period.

## 1.2 Superposition Property

Suppose that  $N_1, N_2, \dots, N_j$  are independent Poisson processes with the rate of  $N_j$  is  $\lambda_j$ . We look at the union of all arrivals. This process is also a Poisson process, the rate  $\lambda_1 + \lambda_2 + \dots + \lambda_j$ . The types of the arrivals will form an iid sequence where the probability of type  $j = \frac{\lambda_j}{\lambda_1 + \dots + \lambda_j}$ .

*Proof.* Let  $N_j$  be a counting function of the  $j^{\text{th}}$  Poisson process. Denote by  $N$  the counting function of arrivals  $N(t) = N_1(t) + N_2(t) + \dots + N_j(t)$

$$N(b) - N(a) = \sum_{j=1}^j \underbrace{(N_j(b) - N_j(a))}_{\text{Poisson}(\lambda_j(b-a))} \sim \text{Poisson} \left( (b-a) \sum_{j=1}^k \lambda_j \right)$$

### 1.2.1 Example

Customers arrive at a ticket counter. 30 girls arrive per an hour and 20 boys arrive per an hour.

(a) What is the expected waiting time between the first and third customer?

- The arrival process of the customers is a Poisson process with rate 50/hour:  
Girls:  $\frac{3}{5}$  Boys:  $\frac{2}{5}$
- $\tau_1, \tau_2, \dots \sim \exp(50)$   
 $E[\tau_2 + \tau_3] = E[\tau_2] + E[\tau_3] = \frac{1}{50} + \frac{1}{50} = \frac{1}{25}$

(b)  $P(\text{The first 3 customers are all girls}) = \left(\frac{3}{5}\right)^3$

## 1.3 Conditioning the Poisson Process

We want to consider on  $\{N(t) = k\}$ . What is the (conditional) distribution of the  $k$  points in  $[0, t)$ .

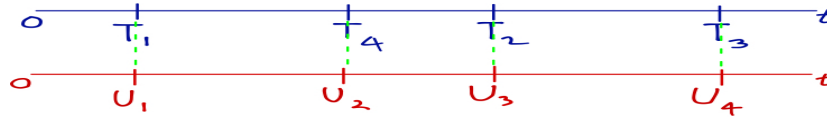
- $k = 1$

$$\begin{aligned} P(T_1 \leq s \mid N(t) = 1) &= \frac{P(T_1 \leq s)(N(t) = 1)}{P(N(t) = 1)} = \frac{(P(N(s) = 1)P(N(t) - N(s) = 0))}{P(N(t) = 1)} \\ &= \frac{P(N(s) = 1)P(N(t) - N(s) = 0)}{P(N(t) = 1)} = \frac{(\lambda s)e^{-\lambda s} \cdot e^{-\lambda(t-s)}}{(\lambda t)e^{-\lambda t}} = \frac{s}{t} \end{aligned}$$

$$P(T_1 \leq s \mid N(t) = 1) = \begin{cases} 1 & \text{if } s \geq t \\ \frac{s}{t} & \text{if } 0 < s < t \\ 0 & \text{if } s \leq 0 \end{cases}$$

**Theorem** Let  $k \geq 1$ . Then the conditional distribution of  $(T_1, T_2, \dots, T_k)$  given that  $N(t) = k$

**Theorem 2.14** If we condition on  $N(t) = n$ , then the vector  $(T_1, T_2, \dots, T_n)$  has the same distribution as  $(V_1, V_2, \dots, V_n)$  and hence the set of arrival times  $(T_1, T_2, \dots, T_n)$  has the same distribution as  $\{U_1, U_2, \dots, U_n\}$



To construct the distribution  $(T_1, T_2, \dots, T_n)$  given  $N(t) = n$

- (i) Put on  $[0, t]$   $n$  i.i.d uniform random points  $(U_1, U_2, \dots, U_n)$
- (ii) Let  $V_1 < V_2 < \dots < V_n$  be the ordered set of values  $(U_1, \dots, U_n)$
- (iii) Then  $(V_1, \dots, V_n)$  has the same joint distribution as  $(T_1, T_2, \dots, T_n)$  conditioned on  $N(t) = n$

Now  $(T_1, T_2, \dots, T_n)$  is the new distribution. Consequently, the joint PDF of  $(T_1, T_2, \dots, T_n)$  conditional on  $N(t) = n$  is  $f(x_1, \dots, x_n) = \frac{n!}{t^n}$  on the set  $\underbrace{\{0 < x_1 < x_2 < \dots < x_n < t\}}_{\text{These random variables are ordered}}$

Aside

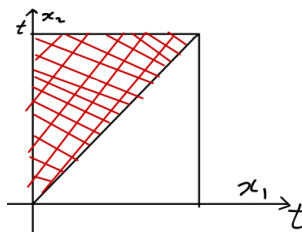
- If  $X$  is uniform  $X \sim \text{Unif}[a, b]$ , then  $X$  has PDF

$$f(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & x \notin [a, b] \end{cases}$$

- $\vec{x} = (x_1, \dots, x_n)$  is uniform on a set  $H \subset \mathbb{R}^n$  if  $\vec{x}$  has joint PDF

$$f(x_1, \dots, x_n) = \begin{cases} \frac{1}{\text{VOL}(H)} & (x_1, \dots, x_n) \in H \\ 0 & (x_1, \dots, x_n) \notin H \end{cases}$$

The volume of uniform density on the set  $\{0 < x_1 < x_2 < \dots < x_n < t\} = \frac{t^n}{n!}$



Closer look at the case  $n=2$

If  $X$  has CDF of  $F$  and PDF of  $f$ , then  $f(x) = F'(x)$ .

Case  $n = 2$  : The joint CDF of  $(X, Y)$  is  $F(x, y) = p(X \leq x, Y \leq y)$ .

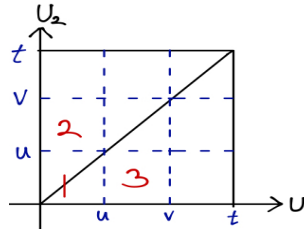
If  $(X, Y)$  has joint PDF  $f$ , then connection between

$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(u, v) dv du$$

If  $f$  is continuous at  $(x, y)$ ,

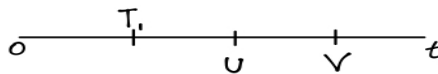
$$\frac{\partial^2}{\partial x \partial y} F(x, y) = f(x, y)$$

Let  $U_1, U_2$  be independent uniform random variables on  $[0, t]$ . Let  $V_1, V_2$  be  $V_1 = \min(U_1, U_2)$ ,  $V_2 = \max(U_1, U_2)$ . Find joint distribution function CDF  $F_{V_1, V_2}$  of  $(V_1, V_2)$ . Let  $u < v$ .



$$\begin{aligned} F(u, v) &= P(V_1 \leq u, V_2 \leq v) \\ &= P(U_1 \leq u, U_2 \leq v) + P(U_1 \leq u, u < U_2 < v) + P(u < U_1 \leq v, U_2 \leq u) \\ &= \left(\frac{u}{t}\right)^2 + \left(\frac{u}{t}\right) \left(\frac{v-u}{t}\right) + \left(\frac{v-u}{t}\right) \left(\frac{u}{t}\right) = \frac{u^2 + 2uv - 2u^2}{t^2} \\ F_{V_1, V_2}(u, v) &= \left(\frac{u}{v}\right)^2 + \frac{2u(v-u)}{t^2} \quad \text{for } 0 < u < v < t \end{aligned}$$

Let  $N(\cdot)$  be a rate  $\lambda$  Poisson Process condition on  $N(t) = 2$ . Find the conditional joint PDF. ( $0 \leq u < v \leq t$ )



$$\begin{aligned} F_{T_1, T_2}(u, v \mid N(t) = 2) &= P(T_1 \leq u, T_2 \leq v \mid N(t) = 2) \\ &= P(N(u) \geq 1, N(v) = 2 \mid N(t) = 2) \\ &= P(N(u) = 1, N(v) = 2 \mid N(t) = 2) \\ &\quad + P(N(u) = 2, N(v) = 2 \mid N(t) = 2) \end{aligned}$$

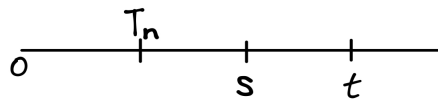
$$\begin{aligned}
&= \frac{P(N(u) = 1, N(v) = 2, N(t) = 2)}{P(N(t) = 2)} + \frac{P(N(u) = 2, N(v) = 2, N(t) = 2)}{P(N(t) = 2)} \\
&= \frac{P(N(u) = 1, N(v) - N(u) = 1, N(t) - N(v) = 0)}{P(N(t) = 2)} + \frac{P(N(u) = 2, N(v) - N(t) = 2)}{P(N(t) = 2)} \\
&= \frac{e^{-u\lambda} u \lambda e^{-\lambda(v-u)} \lambda (v-u) e^{-\lambda(t-v)} + e^{-u\lambda} \frac{(u\lambda)^2}{2} e^{-\lambda(t-u)}}{e^{-t\lambda} \frac{(-t\lambda)^2}{2}} = \frac{2u(v-u)}{t^2} + \frac{u^2}{t^2} = \frac{2uv - u^2}{t^2}
\end{aligned}$$

Conditional Joint Density Function PDF

$$f_{T_1, T_2}(u, v \mid N(t) = 2) \frac{\partial^2}{\partial u \partial v} F_{T_1, T_2}(u, v \mid N(t) = 2) = \frac{2}{t^2}$$

Find the conditional expectation:  $E[T_n \mid N(t) = n]$

The conditional joint density function PDF  $f_{T_n}(s \mid N(t) = n) = \frac{d}{ds} \left( \frac{s}{t} \right)^n = \frac{ns^{n-1}}{t^n}$  on the set  $\{(x_1, \dots, x_n) \mid 0 < x_1 < \dots < x_n < t\}$



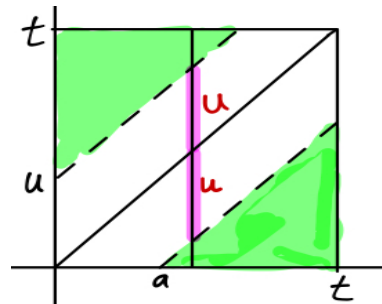
To find PDF of  $T_n$ , let's first find  $P(T_n \leq s \mid N(t) = n) = \left( \frac{s}{t} \right)^n$ .

$$f_{T_n}(s \mid N(t) = n) = \frac{d}{ds} \left( \frac{s}{t} \right)^n = \frac{ns^{n-1}}{t^n} \quad (0 \leq s \leq t)$$

$$E[T_n \mid N(t) = n] = \int_0^t \frac{s \cdot ns^{n-1}}{t^n} ds = \frac{n}{n+1} t$$

$$P(\text{arrivals at least } u \text{ time units apart} \mid N(t)=2) = \int_0^{t-u} \int_{x+u}^t \frac{2}{t^2} dy dx = \frac{(t-u)^2}{t^2}$$

$$\text{Probability} : \frac{\text{shaded area}}{t^2} = \frac{(t-u)^2 \cdot \frac{1}{2} \cdot 2}{t^2}$$



## 1.4 Practice Exam

1.d Let  $N$  be the process of ice cream cones sold with rate 3.

$\lambda_c$  : Rate for chocolate =  $0.7 \cdot 3$     $\lambda_s$  : Rate for strawberry =  $0.2 \cdot 3$     $\lambda_v$  : Rate for vanilla =  $0.1 \cdot 3$

$$\begin{aligned}
 P(N_s(6, 9] = 2 \mid N(7, 9] = 5) &= \frac{P(N_5(6, 9] = 2, N(7, 9] = 5)}{P(N(7, 9] = 5)} \\
 &= \sum_{j=0}^2 \frac{P(N_s(6, 7] = j, N_s(7, 9] = 2-j, N(7, 9] = 5)}{P(N(7, 9] = 5)} \\
 &= \sum_{j=0}^2 \frac{P(N_s(6, 7] = j) \overbrace{P(N_s(7, 9] = 2-j, N(7, 9] = 5)}^{P(N_s(7, 9]=2-j, N_{cv}(7, 9]=3+j): \text{Independence by the thinning property}}}{P(N(7, 9] = 5)} \\
 &= \sum_{j=0}^2 \frac{P(N_s(6, 7] = j) \overbrace{P(N_s(7, 9] = 2-j, N(7, 9] = 5)}^{\text{Independence of arrivals in } (6, 7] \text{ \& } (7, 9]}}{P(N(7, 9] = 5)} \\
 &= \sum_{j=0}^2 \frac{e^{\frac{3}{5}} \left(\frac{3}{5}\right)^j \cdot \frac{1}{j!} \cdot e^{-\frac{6}{5}} \left(\frac{6}{5}\right)^{2-j} \frac{1}{(2-j)!} \cdot e^{-\frac{24}{5}} \left(\frac{24}{5}\right)^{3+j} \frac{1}{(3+j)!}}{e^{-6} \frac{6^5}{5!}} \\
 &= \sum_{j=0}^2 \underbrace{e^{-\frac{3}{5}} \left(\frac{3}{5}\right)^j \frac{1}{j!}}_{\text{Poisson probability for strawberries in } (6, 7]} \underbrace{\binom{5}{2-j} \left(\frac{1}{5}\right)^{2-j} \left(\frac{4}{5}\right)^{3+j}}_{\text{Binomial probability of strawberries in } (7, 9], \text{ comes from conditioning on the total number of sales}}
 \end{aligned}$$

**Example:**  $P(T_3 \leq s \mid N(t) = 4)$

$$\begin{aligned}
 P(T_3 \leq s \mid N(t) = 4) &= P(N(s) \geq 3 \mid N(t) = 4) = P(N(s) = 3 \mid N(t) = 4) + P(N(s) = 4 \mid N(t) = 4) \\
 &= P(\text{out of 4 independent uniforms, 3 land in } [0, s]) \\
 &\quad + P(\text{out of 4 independent uniforms, all 4 land in } [0, s]) \\
 &= 4 \binom{s}{t}^3 \left(1 - \frac{s}{t}\right) + \left(\frac{s}{t}\right)^4
 \end{aligned}$$

Longer way:  $P(T_3 \leq s \mid N(t) = 4)$

$$\begin{aligned}
 &= \frac{P(N(s) = 3, N(t) = 4)}{P(N(t) = 4)} + \frac{P(N(s) = 4, N(t) = 4)}{P(N(t) = 4)} \\
 &= \frac{P(N(s) = 3, N(s, t] = 1)}{P(N(t) = 4)} + \frac{P(N(s) = 4, N(s, t] = 0)}{P(N(t) = 4)} \\
 &= \frac{e^{-s\lambda} (s\lambda)^3 \frac{1}{3!} \cdot e^{-(t-s)\lambda} (t-s)\lambda}{e^{-t\lambda} (t\lambda)^4 \frac{1}{4!}} + \frac{e^{-s\lambda} (s\lambda)^4 \frac{1}{4!} \cdot e^{-(t-s)\lambda} (t-s)\lambda}{e^{-t\lambda} (t\lambda)^4 \frac{1}{4!}} \\
 &= \frac{4! s^3 (t-s)}{3! t^4} + \frac{s^4}{t^4}
 \end{aligned}$$

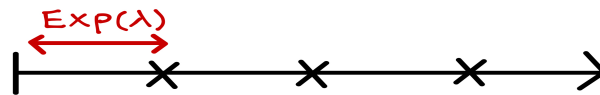
**Fact**

Conditional on  $N(t) = 4$ ,  $(T_1, T_2, T_3, T_4)$  has conditional PDF  $f(x_1, x_2, x_3, x_4) = \frac{4!}{t^4}$  on the set  $\{(x_1, x_2, x_3, x_4) : 0 < x_1 < x_2 < x_3 < x_4 < t\}$   
 whose volume is  $\frac{t^4}{4!}$

$$P(T_3 \leq s \mid N(t) = 4) = \int \dots \int \frac{4!}{t^4} dx_4 dx_3 dx_2 dx_1 = \int_0^s dx_1 \int_{x_1}^s dx_2 \int_{x_2}^s dx_3 \int_{x_3}^t dx_4 \frac{4!}{t^4}$$

**2.27** Person waits for a bus. Time till arrival of bus is  $\text{Unif}(0,1)$ . Cars go by at rate 6. Each car gives this person a ride with probability 0.5. What is the probability that this person rides the bus.

*Remark:* Arrivals come as a rate  $\lambda$  Poisson process then interarrival times are  $\text{Exp}(\lambda)$ .



$$f_S(s) = \begin{cases} 3e^{-3s} & s > 0 \\ 0 & s \leq 0 \end{cases} \quad f_U(s) = \begin{cases} 1 & s \in (0, 1) \\ 0 & s \notin (0, 1) \end{cases}$$

$$\begin{aligned} P(U < S) &= \iint_{u < s} \underbrace{f_U(u)f_S(s)}_{\text{Joint PDF of (U,S)}} du ds \\ &= \int_0^1 1 \, du \int_u^\infty 3e^{-3s} ds \\ &= \int_0^1 e^{-3u} du = \frac{1}{3} \int_0^1 3e^{-3u} du \\ &= \frac{1 - e^{-3}}{3} \end{aligned}$$

