

See also 18-Apr-2018 and 16-May-2018 at
<http://phyloseminar.org/recorded.html>

Likelihood in Phylogenetics

Workshop on Molecular Evolution
Woods Hole, Massachusetts

28 May 2022

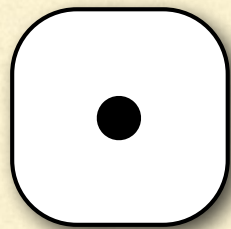
Paul O. Lewis
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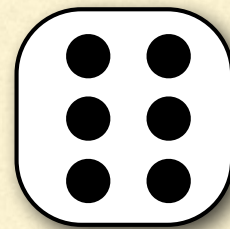
Probability

Probabilities: the AND rule

Rolling 2 dice, what is the probability of seeing (simultaneously) a 1 on the first die and a 6 on the second die?



AND



$(1/6)$

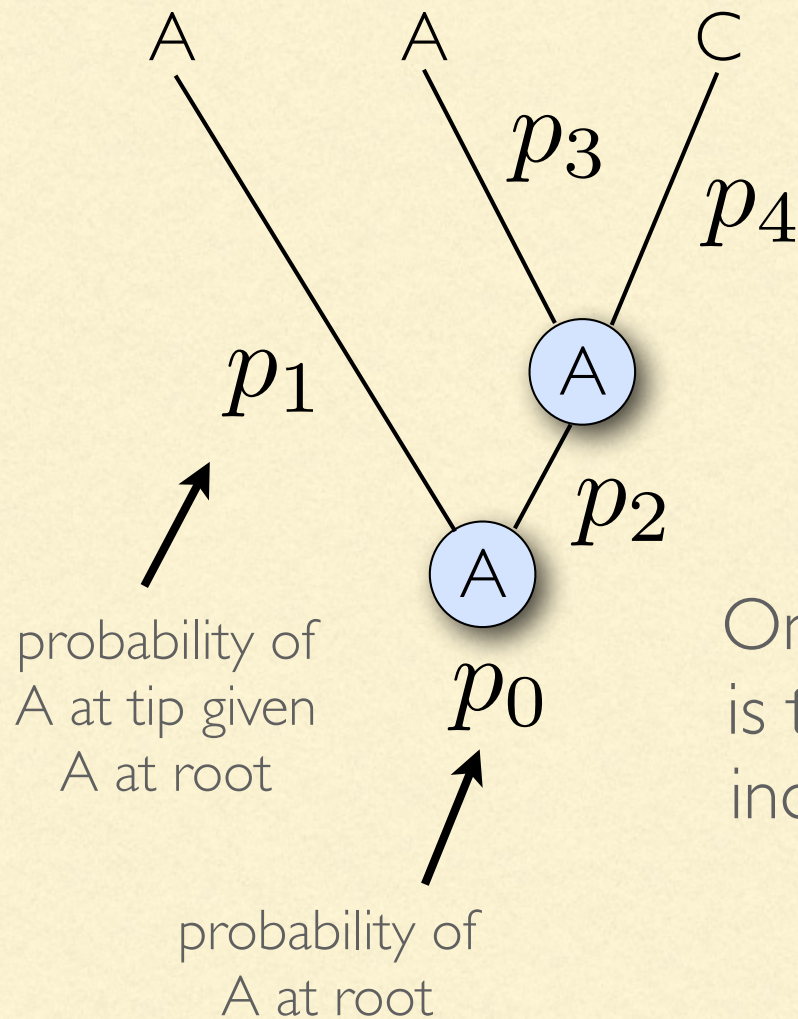
\times

$(1/6)$

$=$

$1/36$

AND rule in phylogenetics



we have observed these states

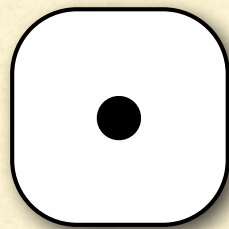
pretend (for now) we also know these ancestral states

$$\Pr(A, A, C, A, A) = p_0 p_1 p_2 p_3 p_4$$

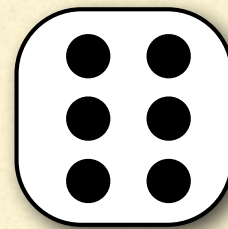
One use of the AND rule in phylogenetics is to combine probabilities associated with individual branches to produce the overall probability of the data for one site.

Probabilities: the OR rule

Rolling 1 die, what is the probability of seeing either a 1 or a 6?



OR



$(1/6)$

+

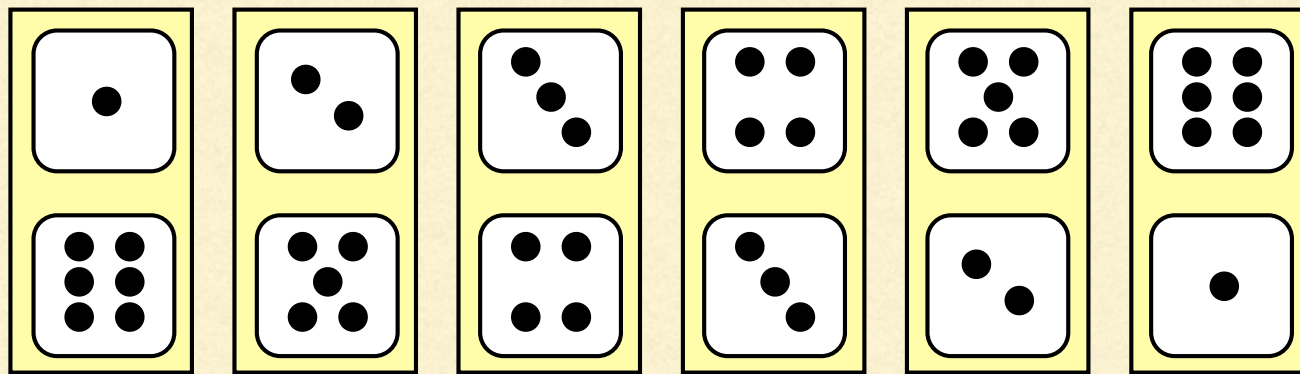
$(1/6)$

=

$1/3$

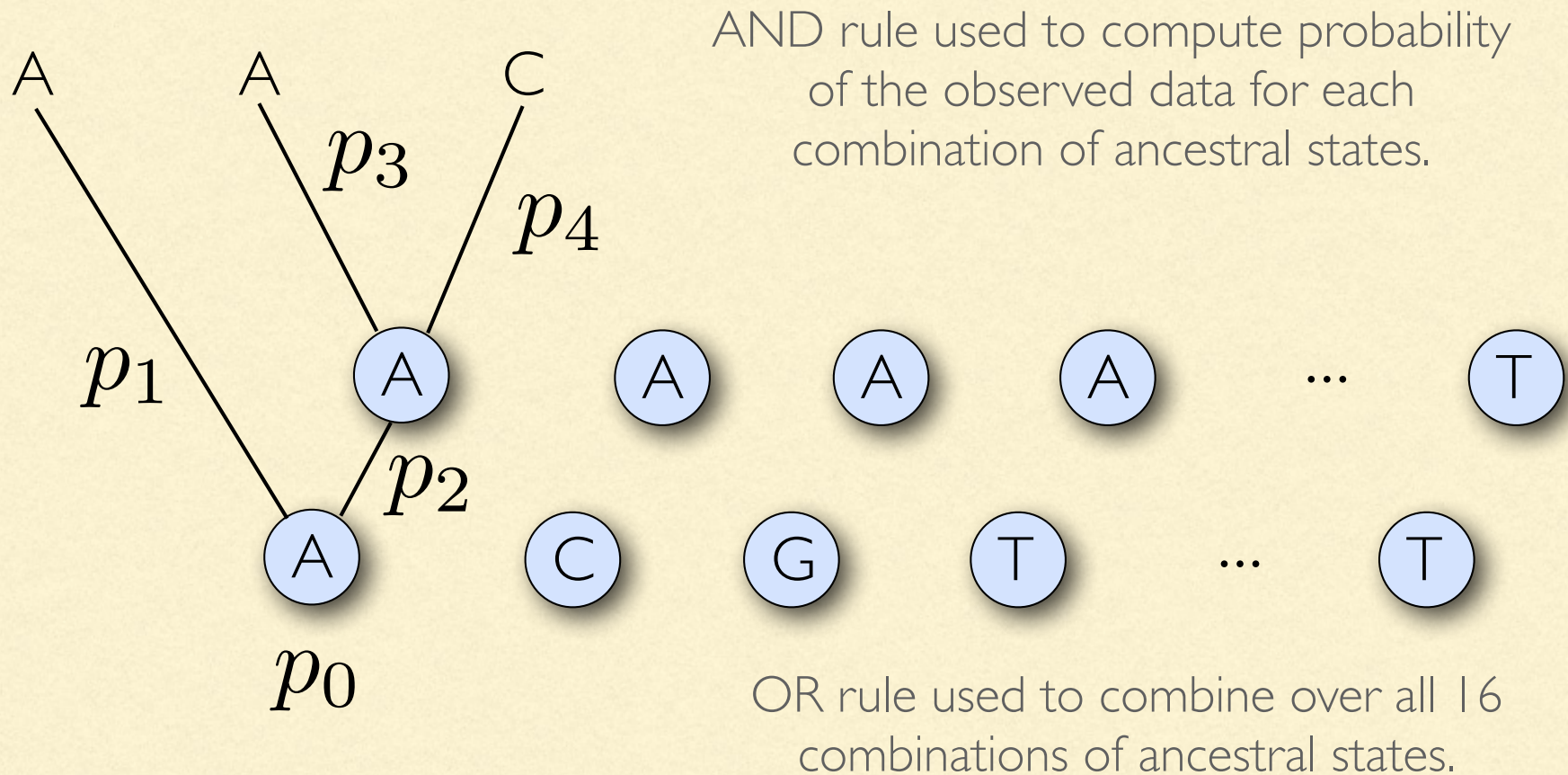
Combining AND and OR

What is the probability that the sum of two dice is 7?



$$(1/36) + (1/36) + (1/36) + (1/36) + (1/36) + (1/36) = 1/6$$

Using both AND and OR in phylogenetics



$$\text{Pr}(\mathbf{A}, \mathbf{A}, \mathbf{C}) = \text{Pr}(\mathbf{A}, \mathbf{A}, \mathbf{C}, \text{A}, \text{A}) + \text{Pr}(\mathbf{A}, \mathbf{A}, \mathbf{C}, \text{A}, \text{C}) + \dots + \text{Pr}(\mathbf{A}, \mathbf{A}, \mathbf{C}, \text{T}, \text{T})$$

Independence

$$\Pr(A, B) = \Pr(A) \Pr(B)$$

Probability of flipping a coin twice and
getting heads both times:

$$\Pr(H, H) = \Pr(H) \Pr(H)$$

Non-independence

$$\Pr(A, B) = \Pr(A) \Pr(B|A)$$

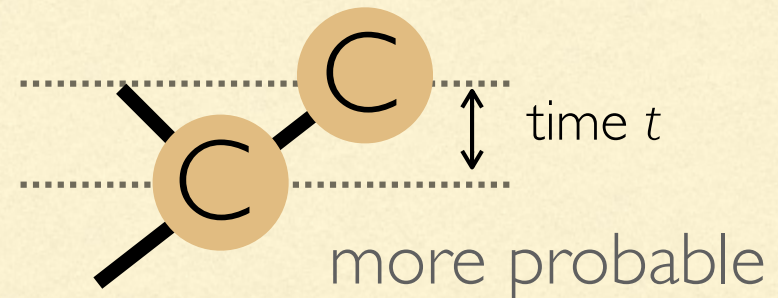
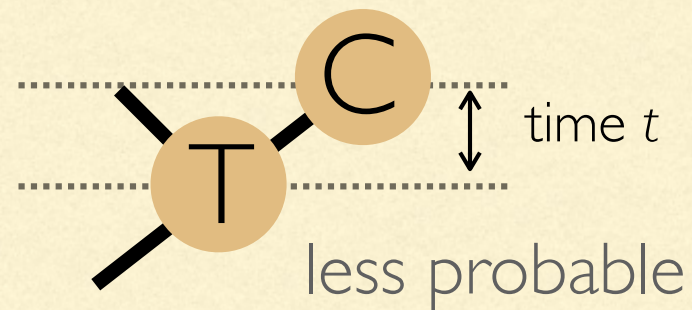
↑
↑
 joint probability of A and B conditional probability of B given A

$$\begin{aligned}\Pr(\text{walk to work}|\text{sunny}) &= 0.99 \\ \Pr(\text{walk to work}|\text{raining}) &= 0.50\end{aligned}$$

Non-independence in phylogenies

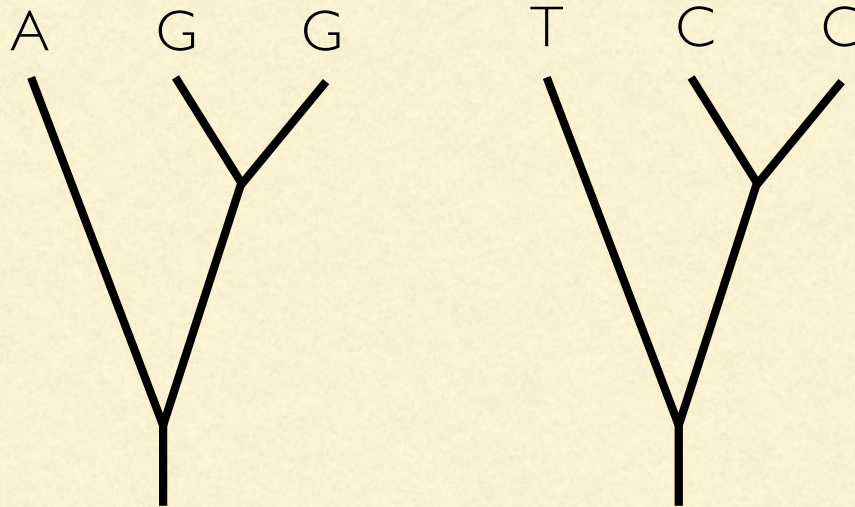
Normally, for a given rate of substitution and time, the probability of the end state is *dependent* on the starting state

$$\Pr(C|C) > \Pr(C|T)$$



Conditional Independence

$$\Pr(A,B|C) = \Pr(A|C) \Pr(B|C)$$

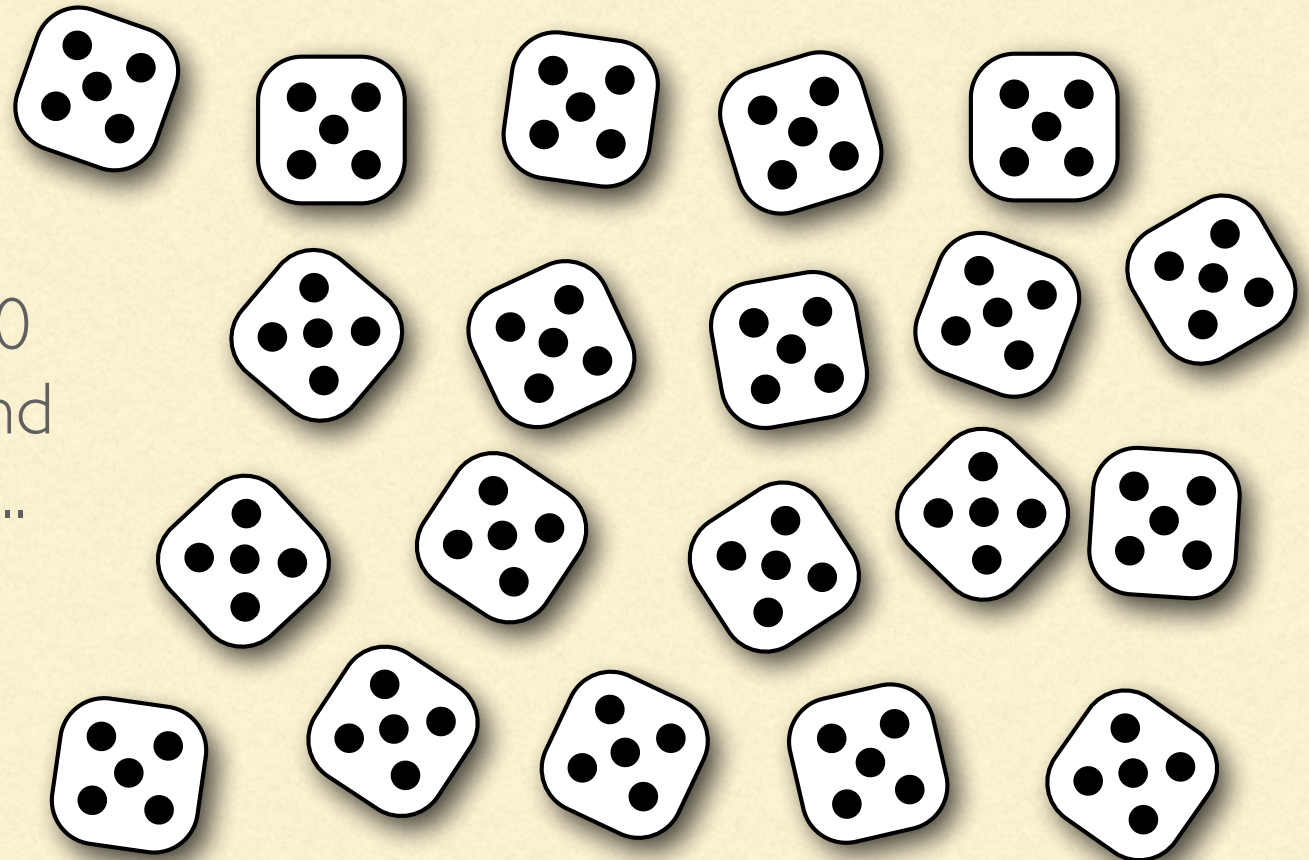


$$\Pr(\text{AGG}, \text{TCC} | \text{tree}) = \Pr(\text{AGG} | \text{tree}) \Pr(\text{TCC} | \text{tree})$$

Likelihood

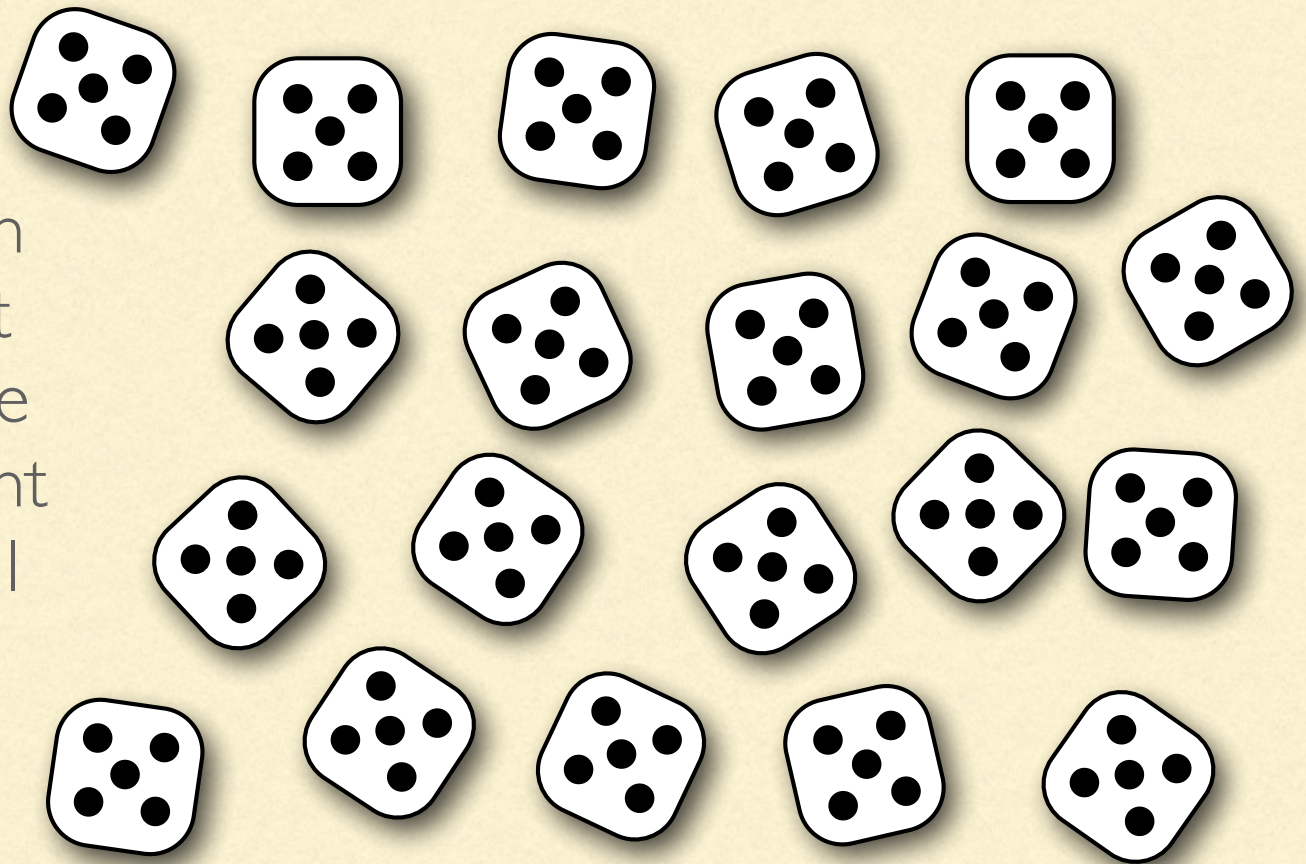
The Likelihood Criterion

Suppose I threw 20
dice on the table and
this was the result...



The Fair Dice model

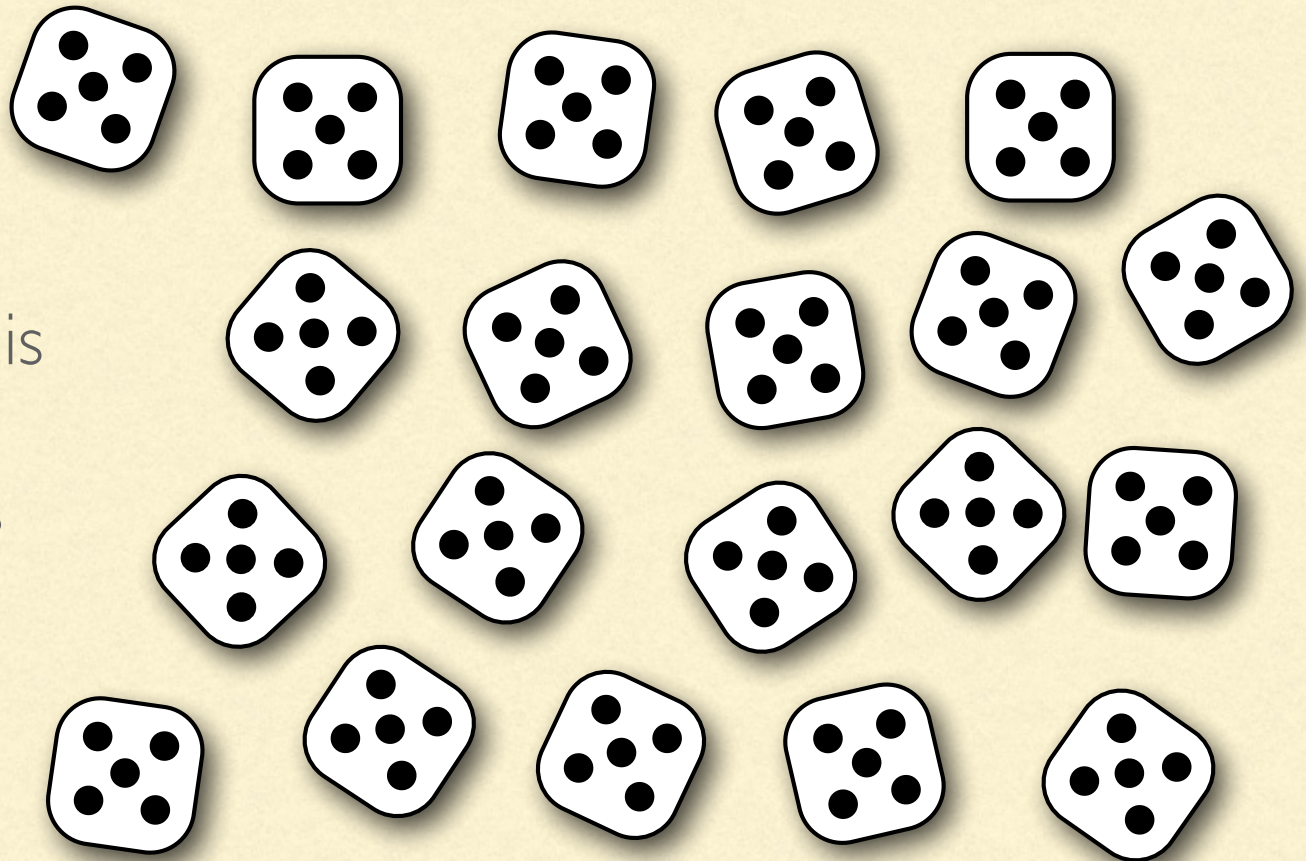
You should have been
very surprised at
this result because the
probability of this event
is **very small**: only 1
in 3.6 quadrillion!



The Trick Dice model


(assumes dice each have 5 on every side)

You should be **not surprised** at all at this result because the observed outcome is **certain** under this model.



What's changed? (the model)

The **winning model** makes us **least surprised** at the data we've observed



Model	Likelihood	Surprise level
Fair Dice	$\frac{1}{3,656,158,440,062,976}$	Very, very, very surprised
→ Trick Dice	1	Not surprised at all

Why do we need the term **likelihood**?

Outcome	Fair coin model	Two-heads model
H	0.5	1
T	0.5	0
	1	1

Likelihoods of models given one particular data set are not expected to sum to 1.0

Probabilities of data outcomes given one particular model sum to 1.0

Probability of the **data**
given the model
Likelihood of the **model**
given the data

Likelihood of a single vertex

First 32 nucleotides of the $\psi\eta$ -globin gene of gorilla:

● **GAAGTCCTTGAGAAATAAACTGCACACACTGG**

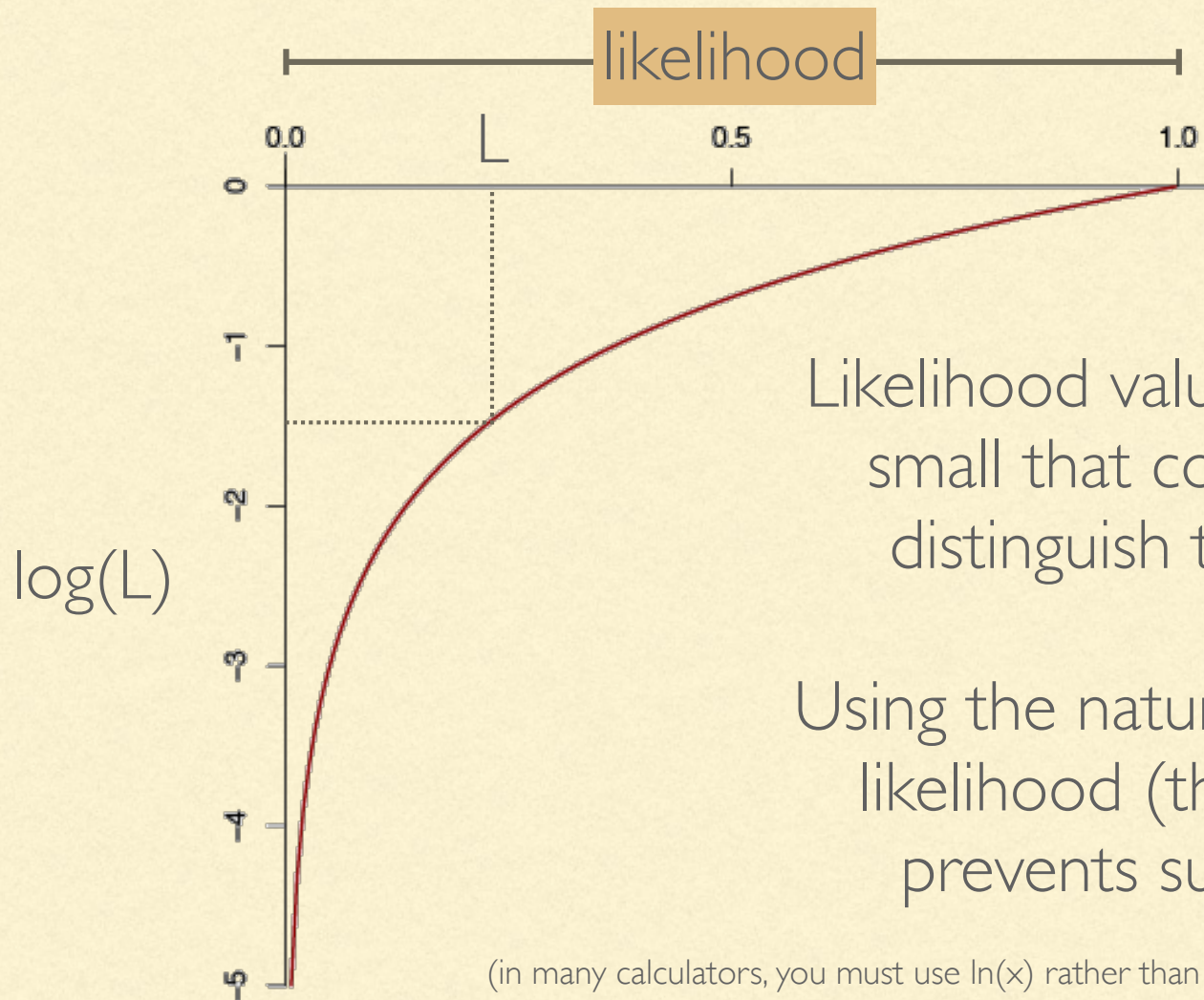
$$L = \Pr(G) \Pr(A) \Pr(A) \Pr(G) \Pr(T) \cdots \Pr(G)$$

$$L = \pi_G \pi_A \pi_A \pi_G \pi_T \cdots \pi_G$$

$$L = \pi_A^{12} \pi_C^7 \pi_G^7 \pi_T^6$$

$$\log L = 12 \log(\pi_A) + 7 \log(\pi_C) + 7 \log(\pi_G) + 6 \log(\pi_T)$$

Natural logarithm



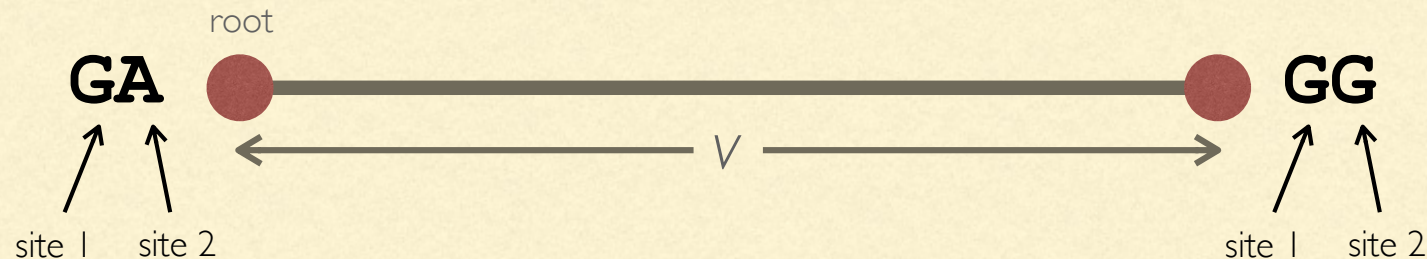
Likelihood values can become so small that computers cannot distinguish them from zero.

Using the natural logarithm of the likelihood (the log-likelihood) prevents such "underflow"

(in many calculators, you must use $\ln(x)$ rather than $\log(x)$ to take the natural log of the value x)

Likelihood of a single-edge tree

Two nodes have sequence data (but only for two sites)



$$L = \left[\left(\frac{1}{4} \right) \left(\frac{1}{4} + \frac{3}{4} e^{-4v/3} \right) \right] \left[\left(\frac{1}{4} \right) \left(\frac{1}{4} - \frac{1}{4} e^{-4v/3} \right) \right]$$

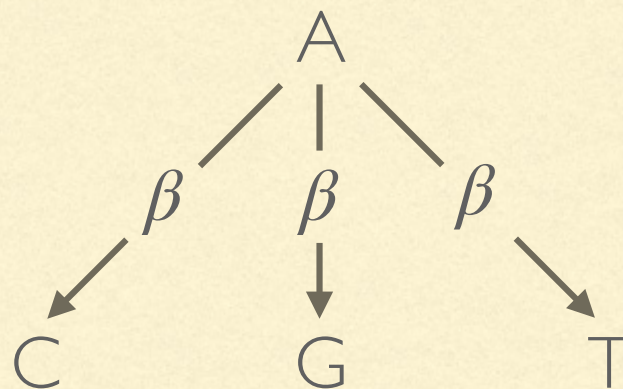
Below the equation, two horizontal double-headed arrows represent the components. The first arrow is labeled 'site 1' and is positioned under the first bracketed term. The second arrow is labeled 'site 2' and is positioned under the second bracketed term.

Each **site likelihood** is the probability of the **starting state** at the root ($1/4$) times the **transition probability** (probability of the end state given the starting state)

What is the edge length v ?

number of
substitutions
per site = substitution
rate per site \times **time**

$$v = 3\beta \times t$$



3 possible substitutions, each of which happens with rate β

Jukes and Cantor (1969)

JC69 model

to:

Parameters: β

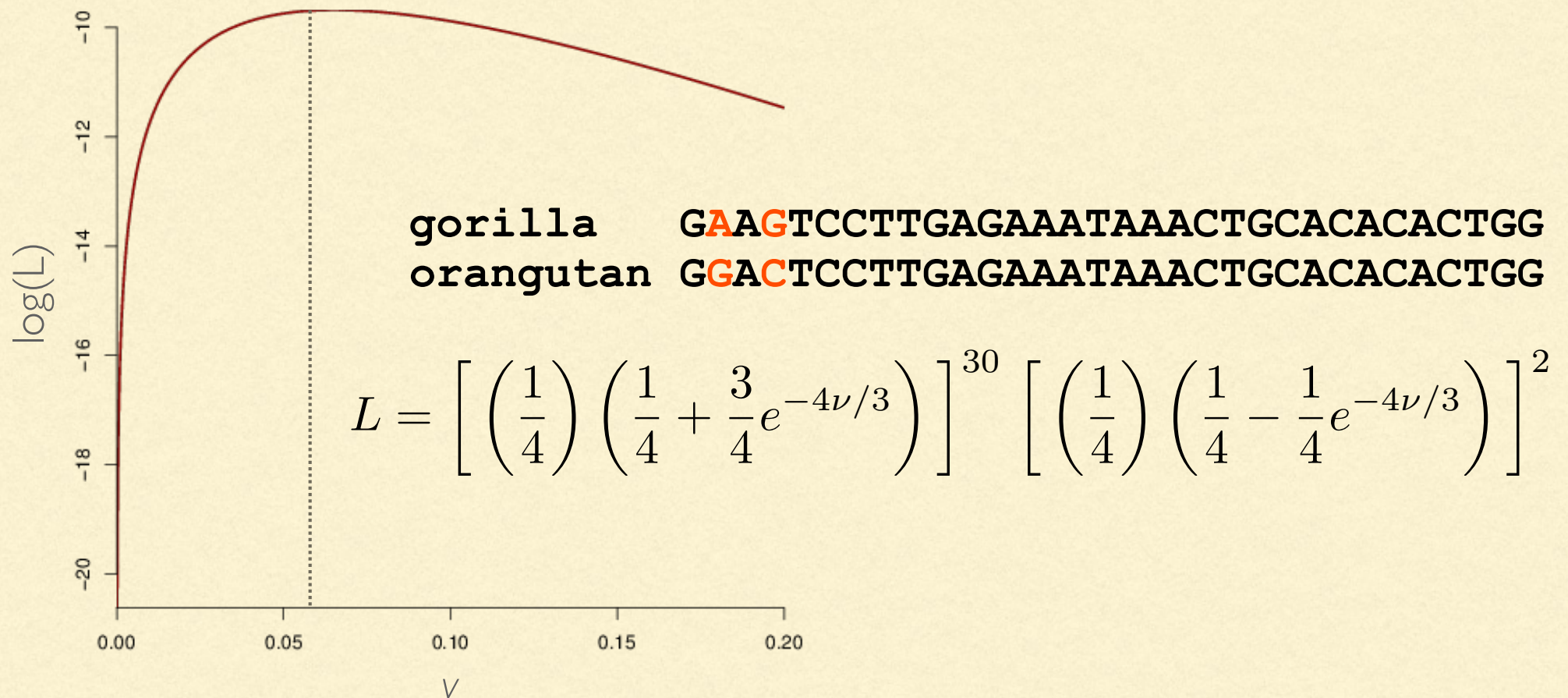
from:

	A	C	G	T
A	-3β	β	β	β
C	β	-3β	β	β
G	β	β	-3β	β
T	β	β	β	-3β

Jukes, T. H., and C. R. Cantor. 1969. Evolution of protein molecules. Pages 21-132 in H. N. Munro (ed.), Mammalian Protein Metabolism. Academic Press, New York.

Maximum likelihood estimation

0.065 is the maximum likelihood estimate (MLE) of ν



Transition probabilities

$$\frac{1}{4} + \frac{3}{4}e^{-4\nu/3}$$

same state

$$\frac{1}{4} - \frac{1}{4}e^{-4\nu/3}$$

different states

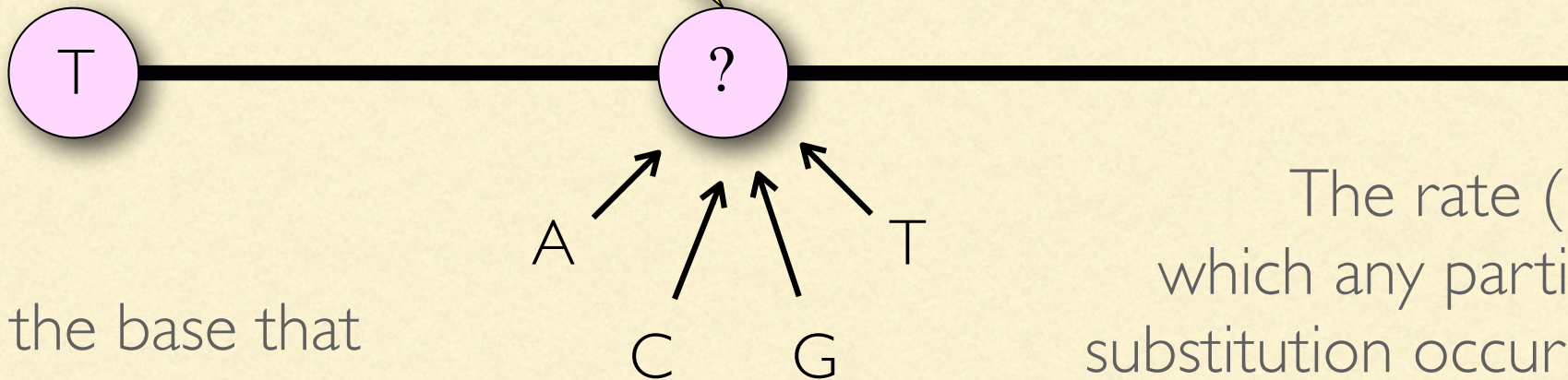
Conditional probability of end state given
starting state and edge length

How do these formulas arise?

"ACHNyons" vs. substitutions

I made up this term
(Anything Can
Happen Now)

When an **achnyon** occurs, any
base can appear in a sequence.



If the base that
appears is different
from the base that
was already there, then a
substitution event has occurred.

The rate (β) at
which any particular
substitution occurs will
be 1/4 the achnyon rate (μ).
That is, $\beta = \mu/4$ (or $\mu = 4\beta$)

Deriving a transition probability

Calculate the probability that a site currently T will change to G over time t when the rate of this particular substitution is β :

$$\text{Pr}(0 \text{ achnyons}) = e^{-\mu t} \quad (\text{Poisson probability of zero events})$$

$$\text{Pr}(\text{at least 1 achnyon}) = 1 - e^{-\mu t}$$

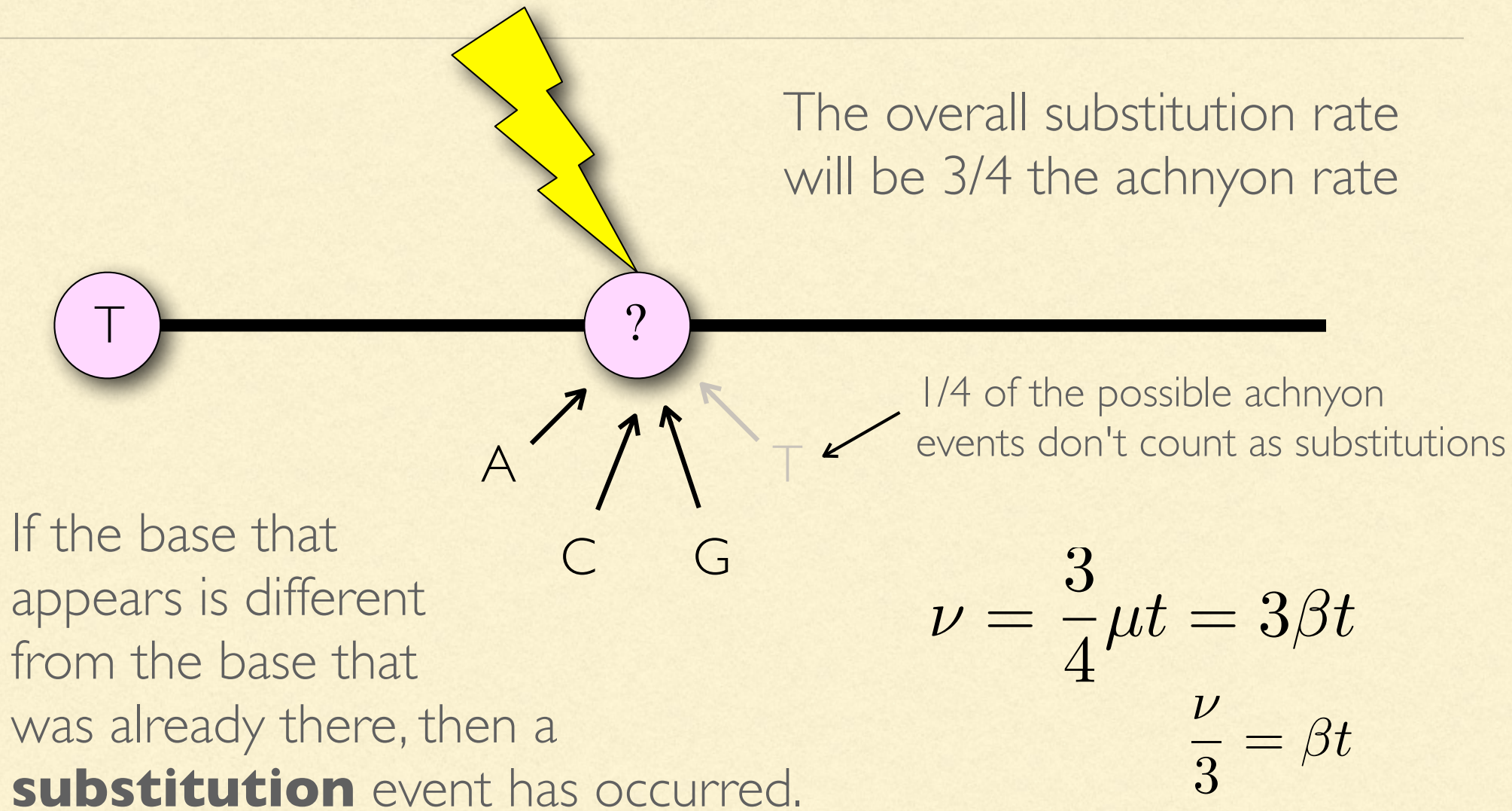
$$\text{Pr}(\text{last achnyon results in base G}) = \frac{1}{4}$$

$$\text{Pr}(\text{end in G} \mid \text{start in T}) = \frac{1}{4} (1 - e^{-\mu t})$$

Remember that the achnyon rate (μ) is 4 times the rate (β) of any *particular* substitution:

$$P_{TG}(t) = \frac{1}{4} (1 - e^{-4\beta t})$$

Expected number of substitutions



Deriving a transition probability

Calculate the probability that a site currently T will change to G over time t when the rate of this particular substitution is β :

⋮

Remember that the achnyon rate (μ) is 4 times the rate (β) of any *particular* substitution:

$$P_{TG}(t) = \frac{1}{4} (1 - e^{-4\beta t})$$

Substitute $v/3$ for βt :

$$P_{TG}(t) = \frac{1}{4} \left(1 - e^{-4v/3} \right)$$

Transition Probabilities: Remarks

$$P_{TA}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

$$P_{TC}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

$$P_{TG}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

$$P_{TT}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

These should add to 1.0
because *T must* change
to something!

$$1 - e^{-4\nu/3}$$

Doh! Something must be
wrong here...

Transition Probabilities: Remarks

$$P_{TA}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

$$P_{TC}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

$$P_{TG}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right)$$

$$P_{TT}(t) = \frac{1}{4} \left(1 - e^{-4\nu/3} \right) + e^{-4\nu/3}$$

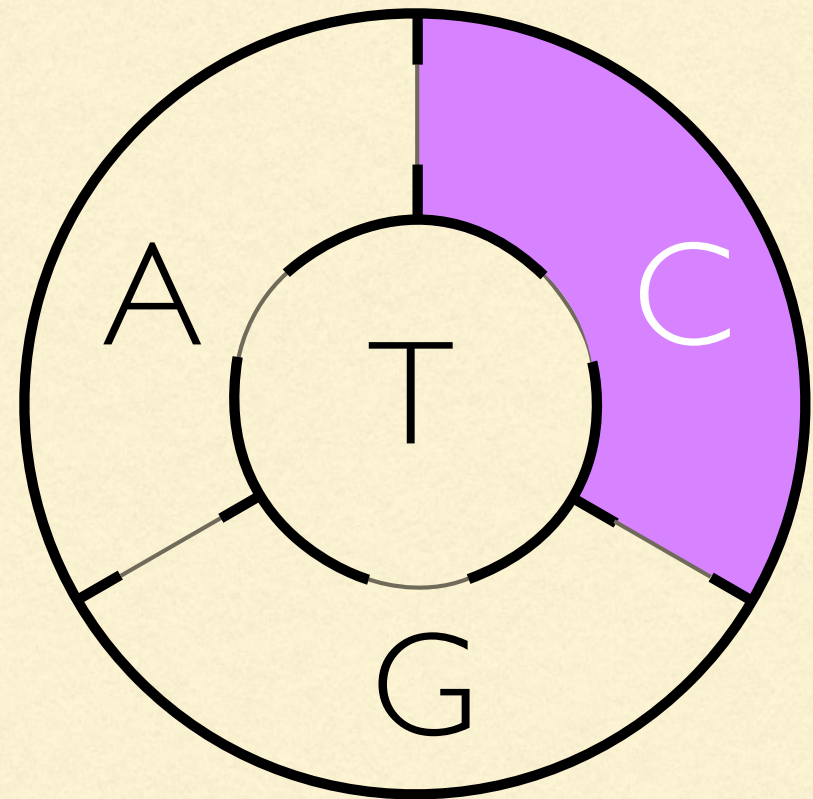
I forgot to account for the possibility of *no* achnyons over time t

Equilibrium Frequencies

Imagine a bottle of perfume has been spilled in room C.

The doors to the other rooms are closed, so the perfume has, thus far, not been able to spread.

What would happen if we opened all the doors?



Architect: Joe Bielawski

Equilibrium Frequencies

At the **instant the doors open**, perfume molecules...

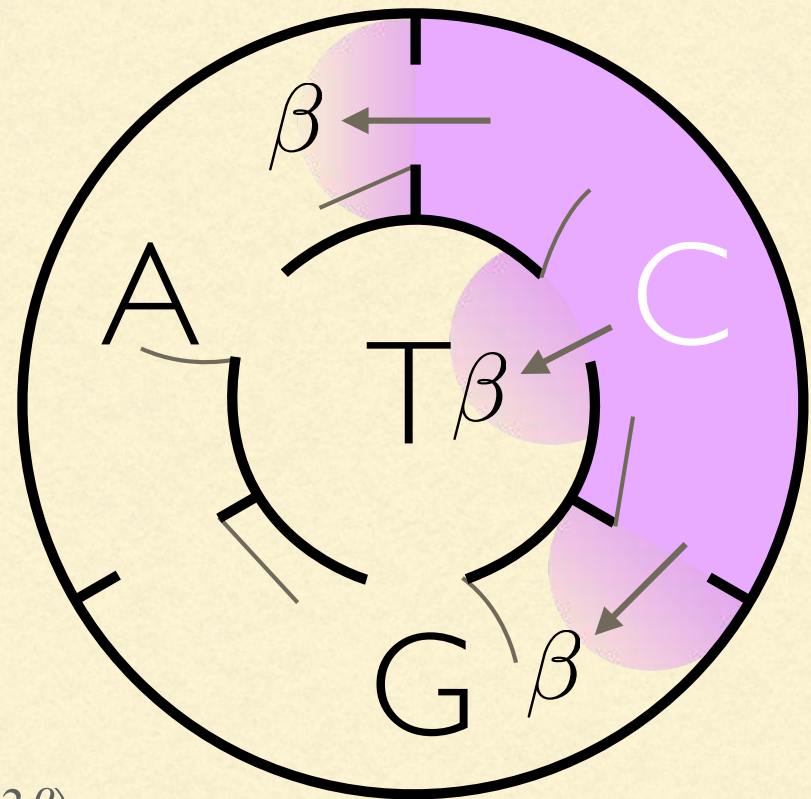
enter room A at rate β

enter room T at rate β

enter room G at rate β

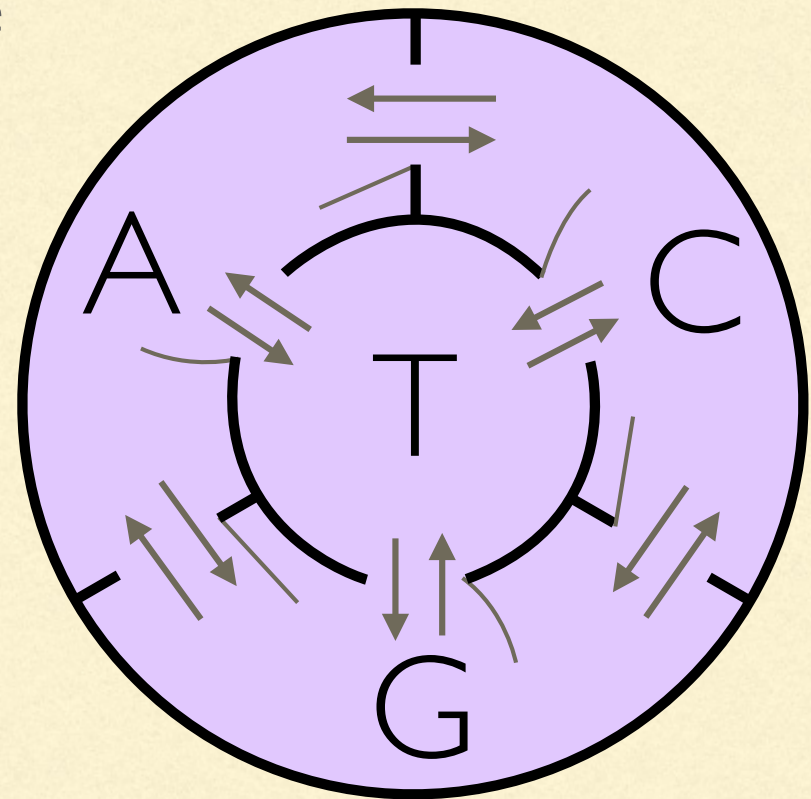
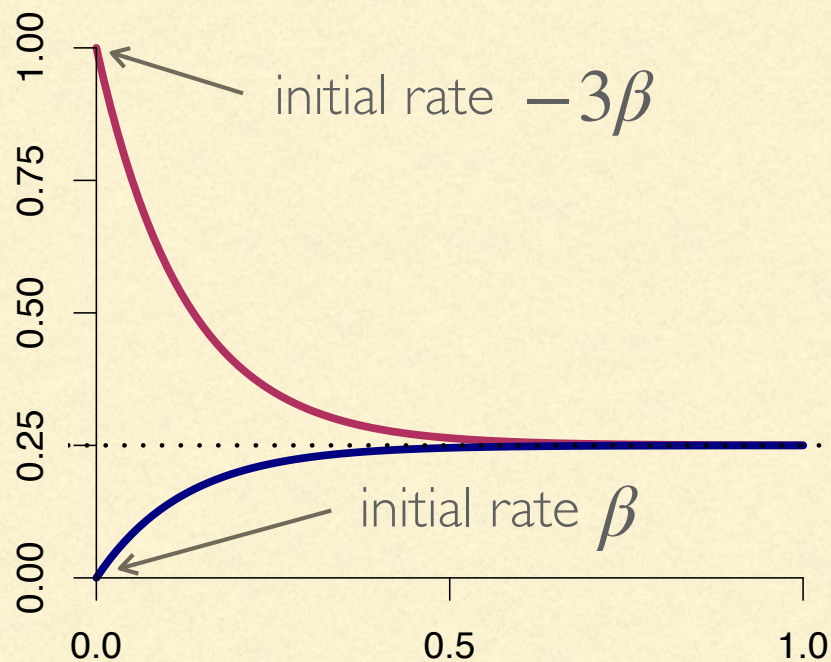
enter room C at rate -3β

(you could also say they *leave* C at rate 3β)



Equilibrium Frequencies

At **equilibrium**, the relative concentration of perfume is **equal** in all rooms



$$\pi_A = \pi_C = \pi_G = \pi_T = \frac{1}{4}$$

Transition probability demo

<https://plewis.github.io/applets/jc-transition-probabilities/>

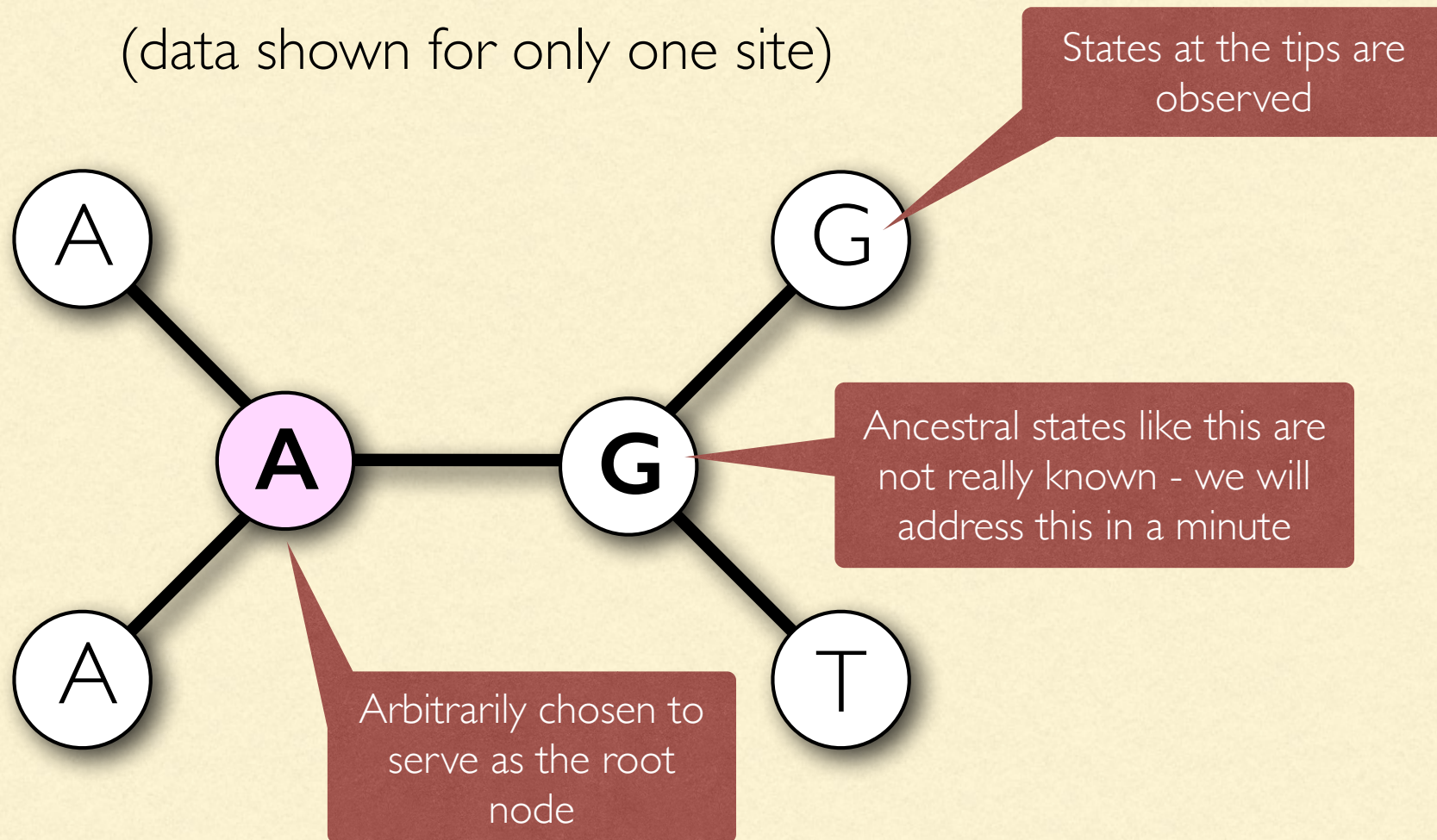
Sequence data for four taxa

one site

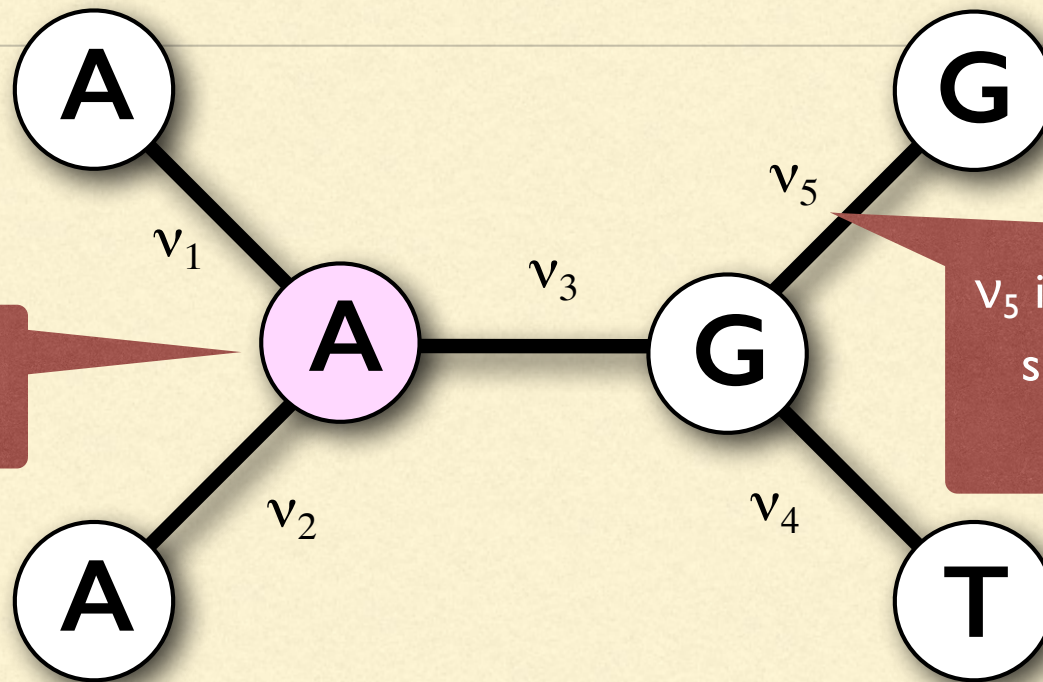
Sphagnum	GGCAGCATTTCGAATGACTCCTCAACCTGGAGT	A	TCACCCG...
Asplenium	GGCAGCTTTCCGGATGACCCCAACAACCCGGAGT	A	TCAGCTG...
Picea	GGCAGCATTCCGAGTAACTCCTCAACCAGGGGT	G	TCGCCCG...
Avena	GGCAGCATTCCGAGTAACTCCTCAACCTGGGGT	T	TCGCCGG...

Likelihood for tree (one site)

(data shown for only one site)

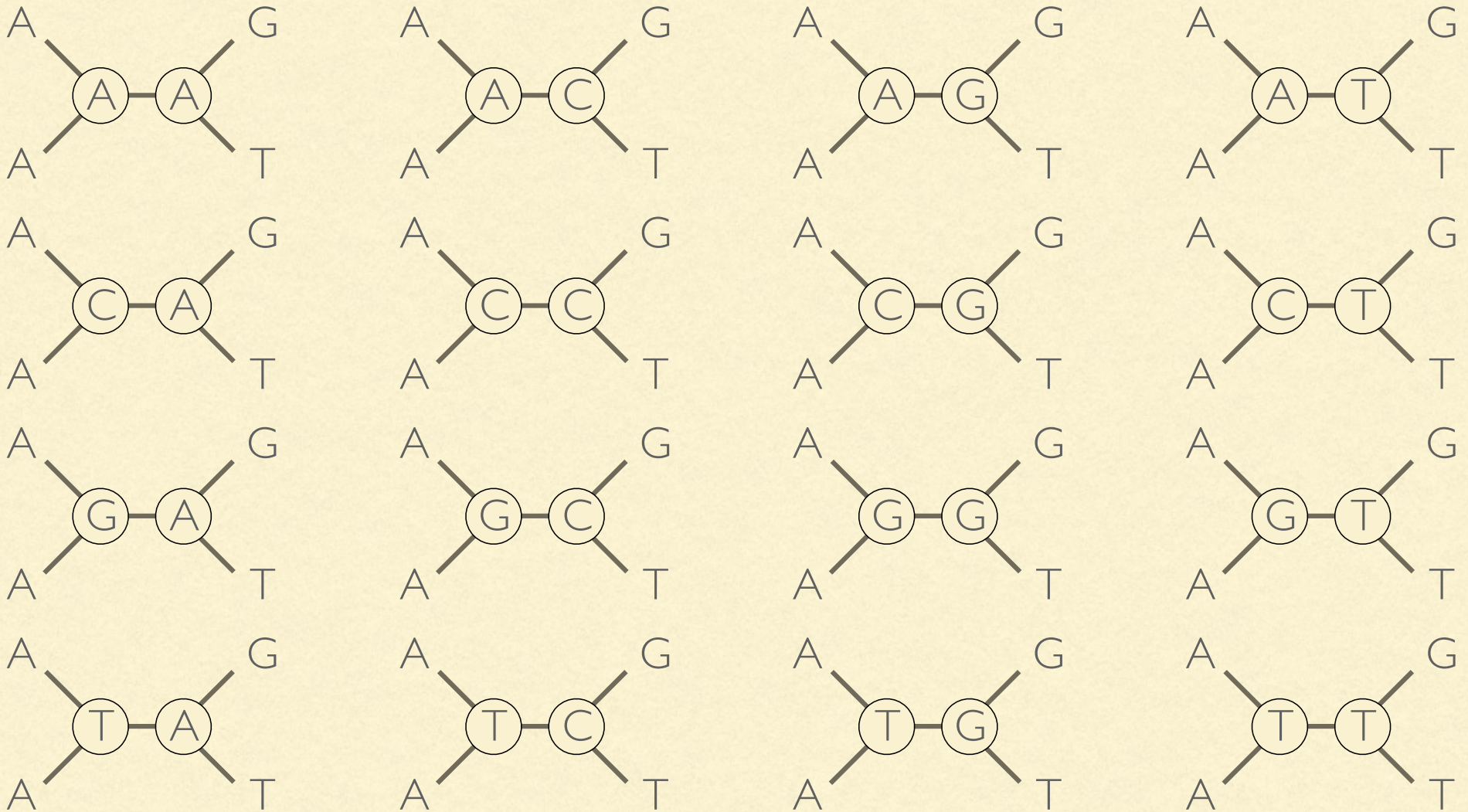


Likelihood for tree (one site)



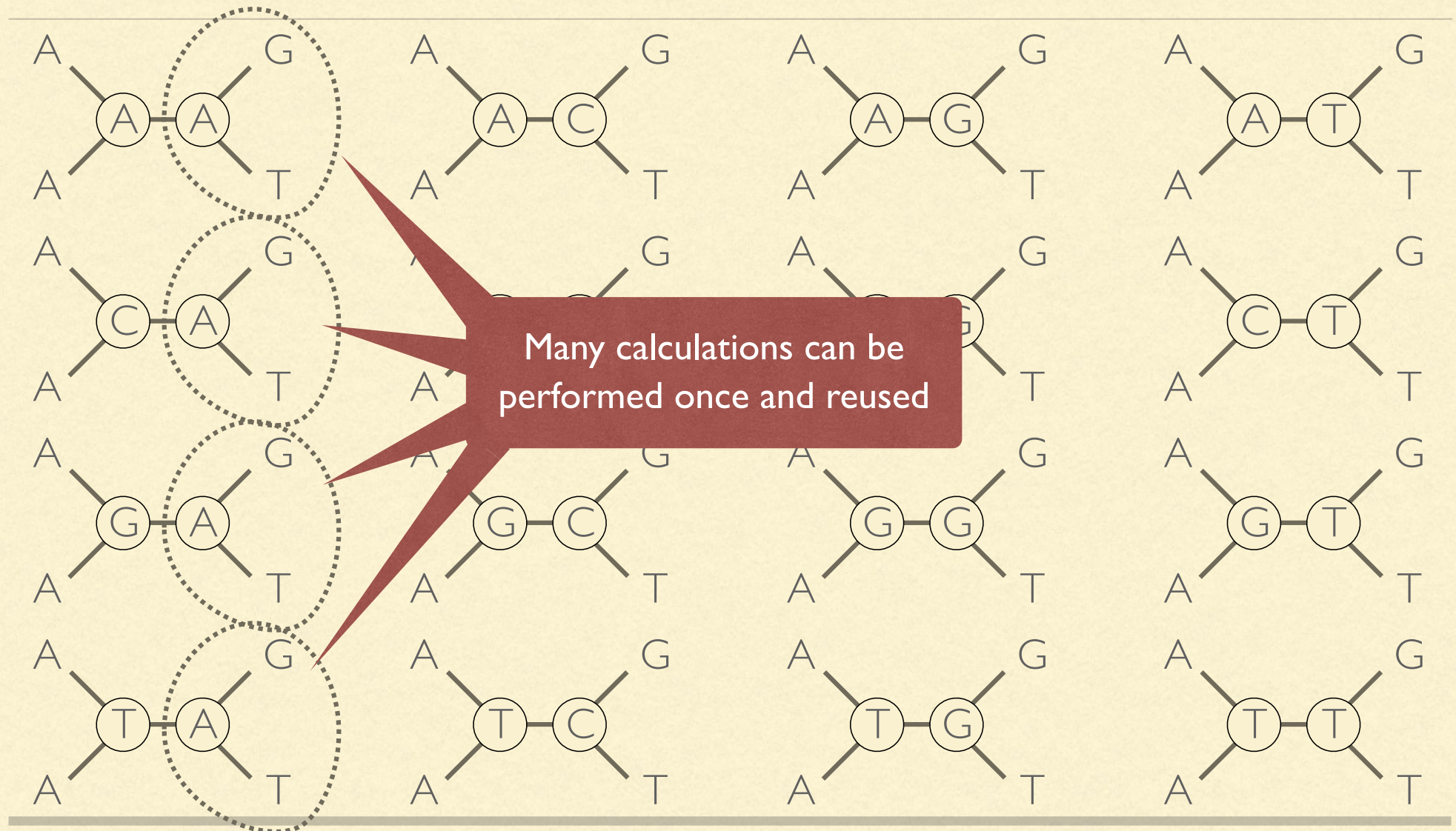
$$L = \frac{1}{4} \left[\frac{1}{4} + \frac{3}{4} e^{-4\nu_1/3} \right] \left[\frac{1}{4} + \frac{3}{4} e^{-4\nu_2/3} \right] \left[\frac{1}{4} - \frac{1}{4} e^{-4\nu_3/3} \right] \left[\frac{1}{4} - \frac{1}{4} e^{-4\nu_4/3} \right] \left[\frac{1}{4} + \frac{3}{4} e^{-4\nu_5/3} \right]$$

Brute force approach

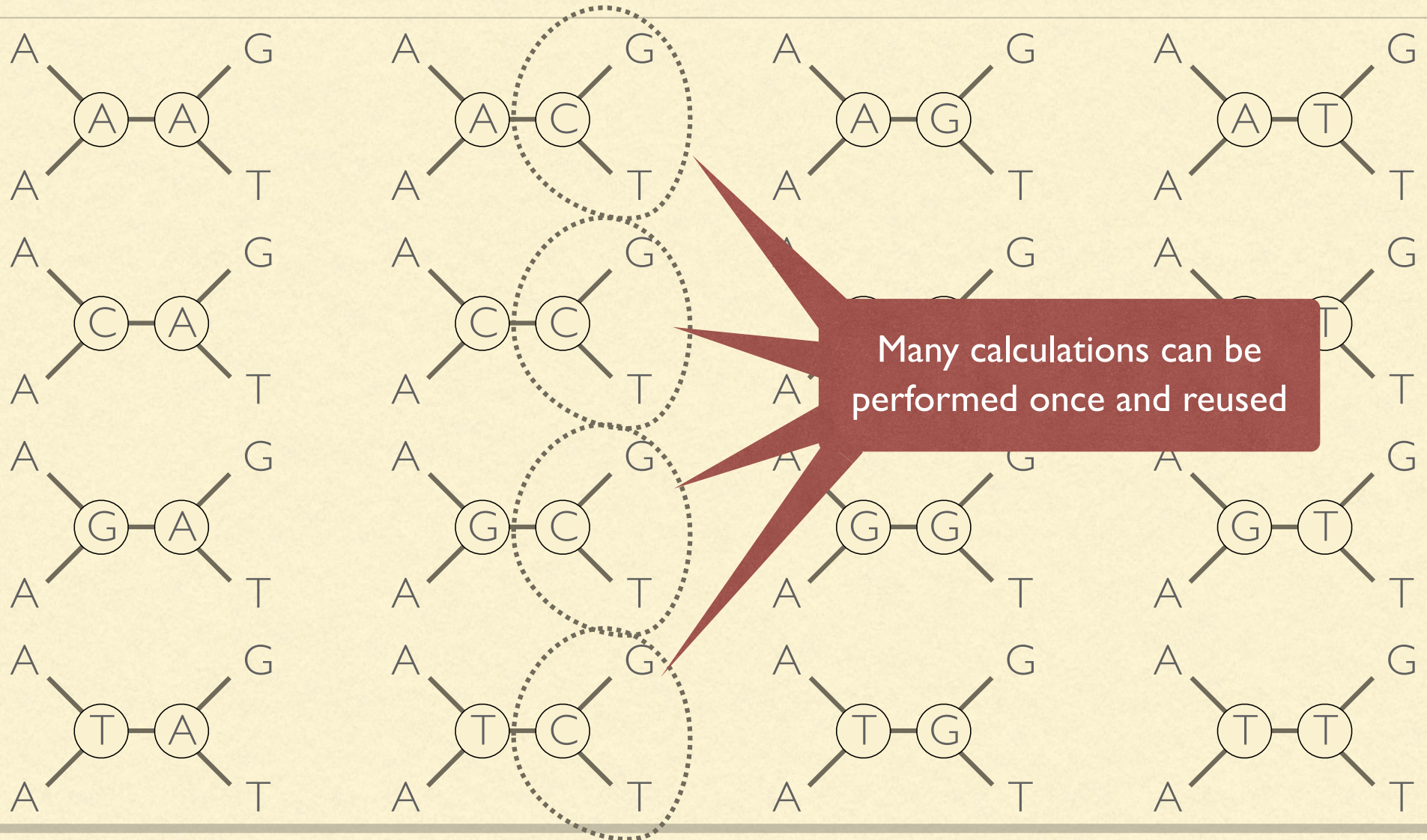


Pruning algorithm

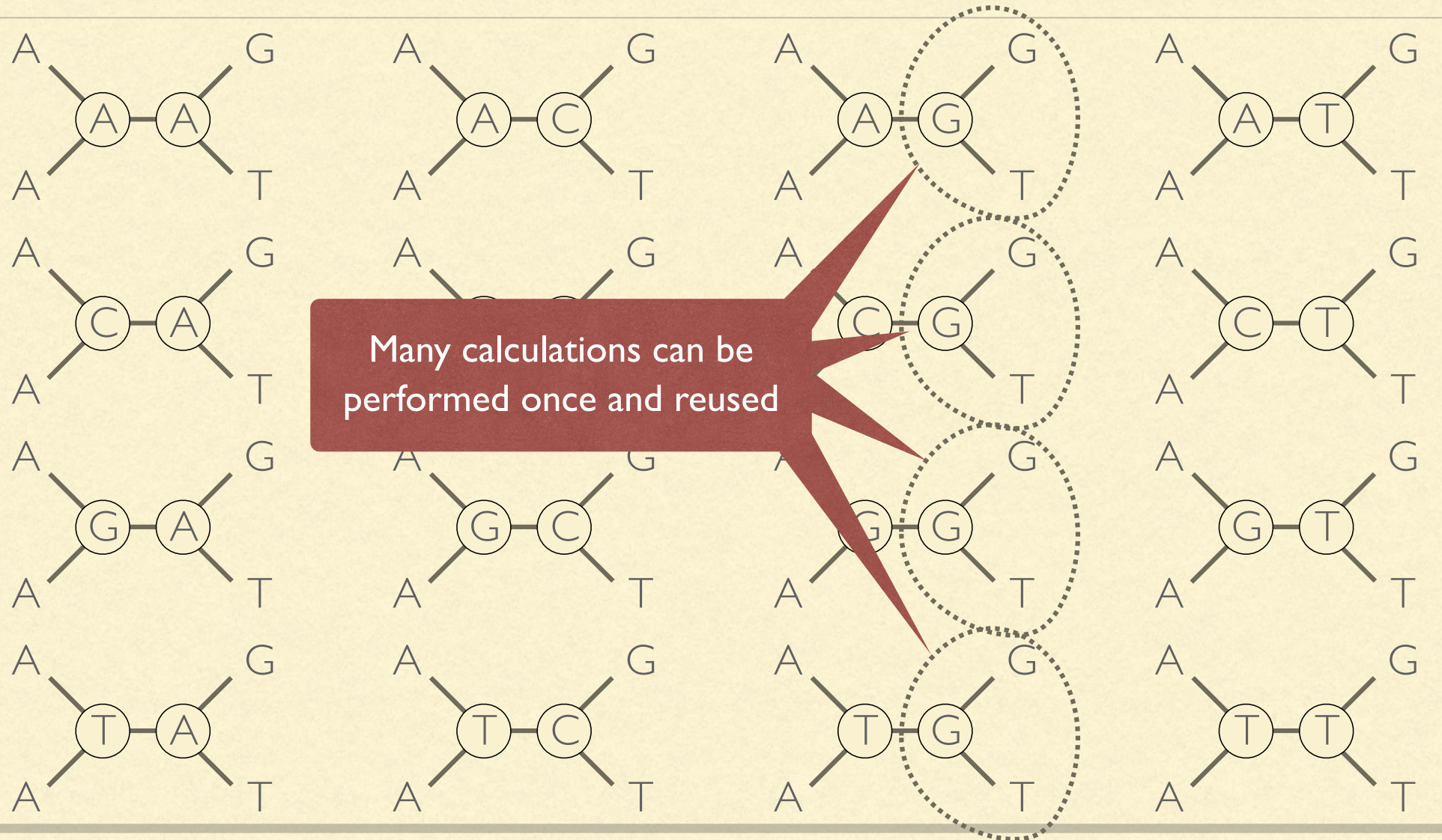
Felsenstein, J. 1981. Evolutionary trees from DNA sequences: a maximum likelihood approach. *Journal of Molecular Evolution* 17:368-376.



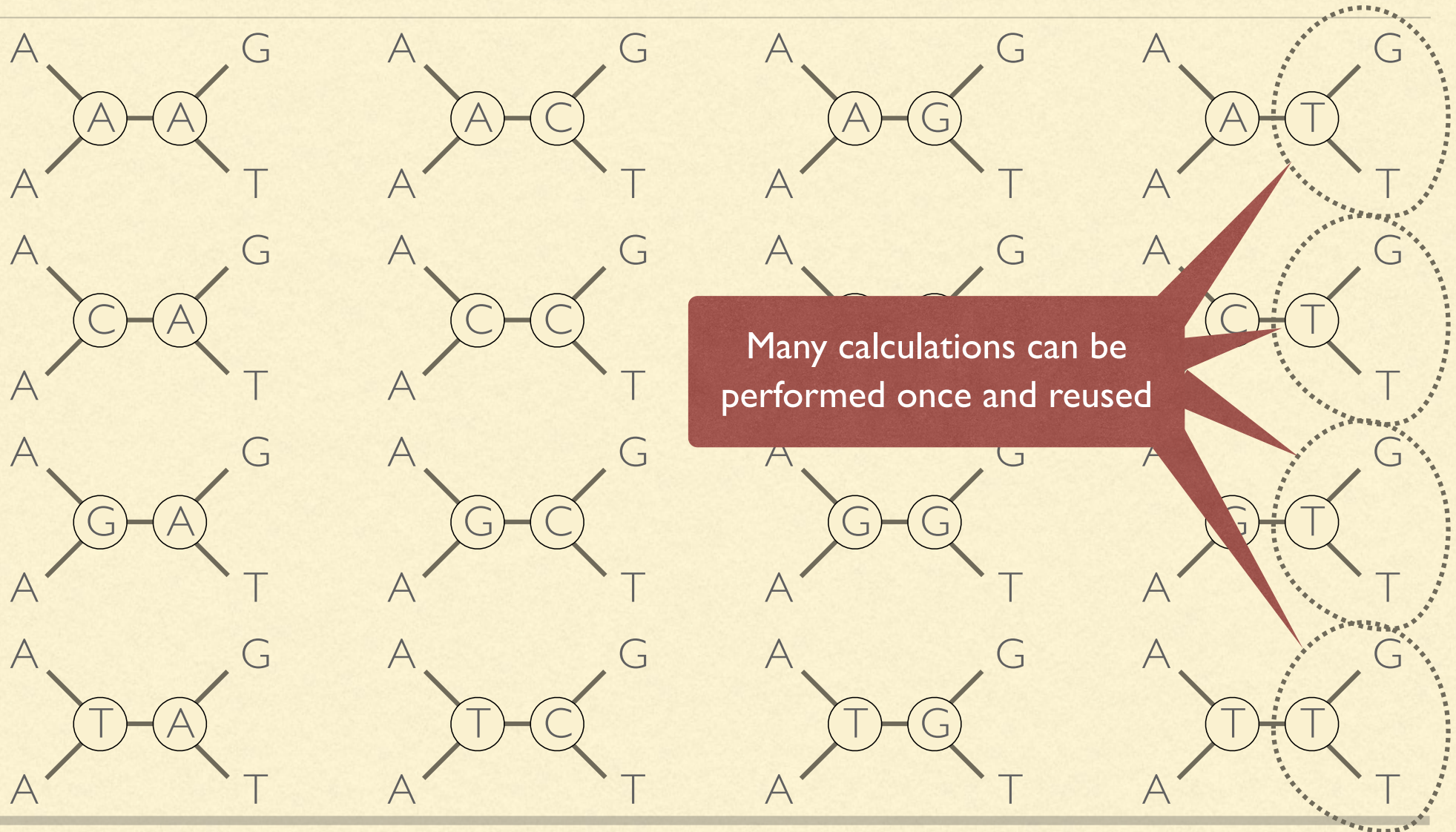
Pruning algorithm



Pruning algorithm

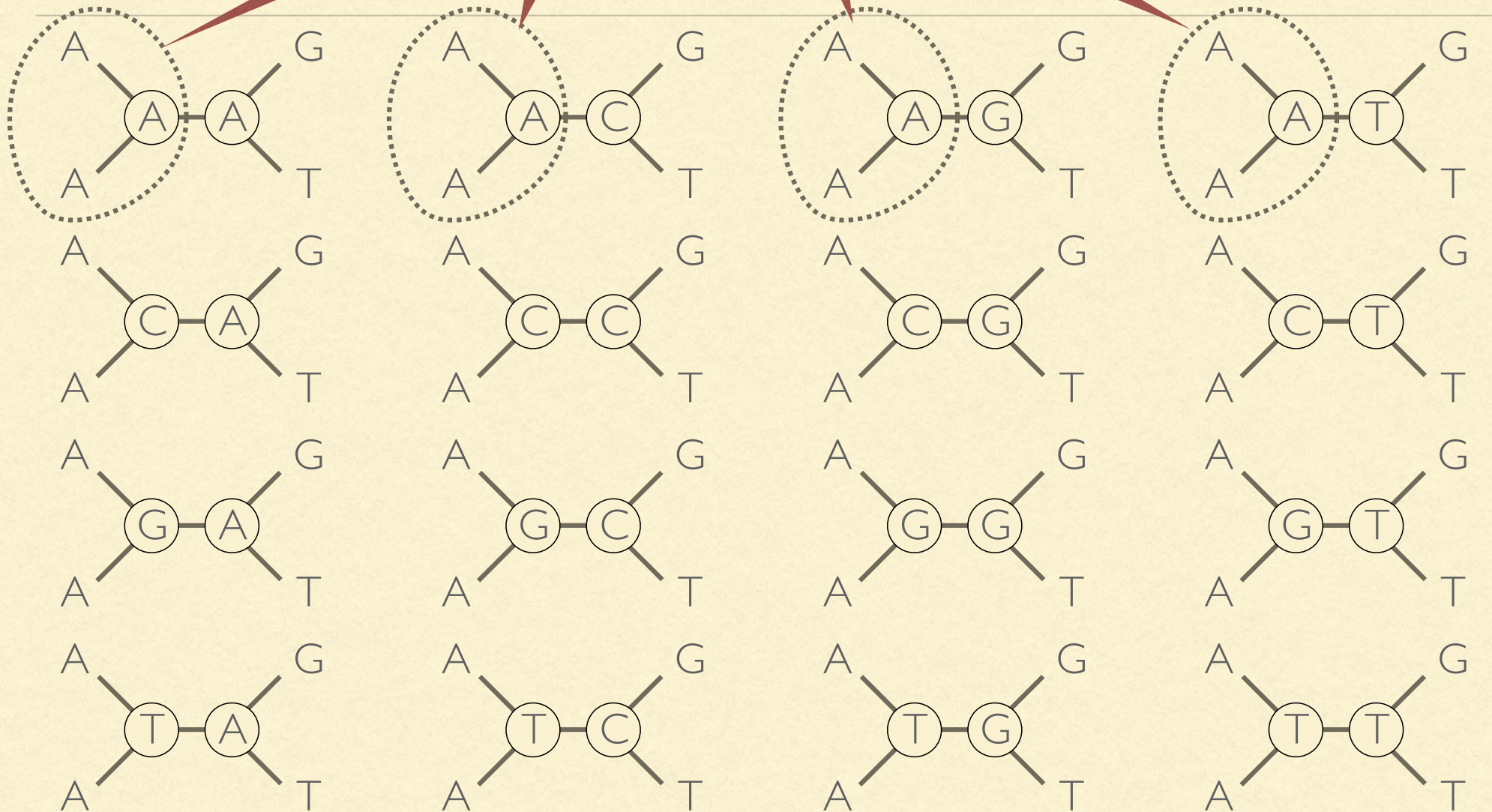


Pruning algorithm

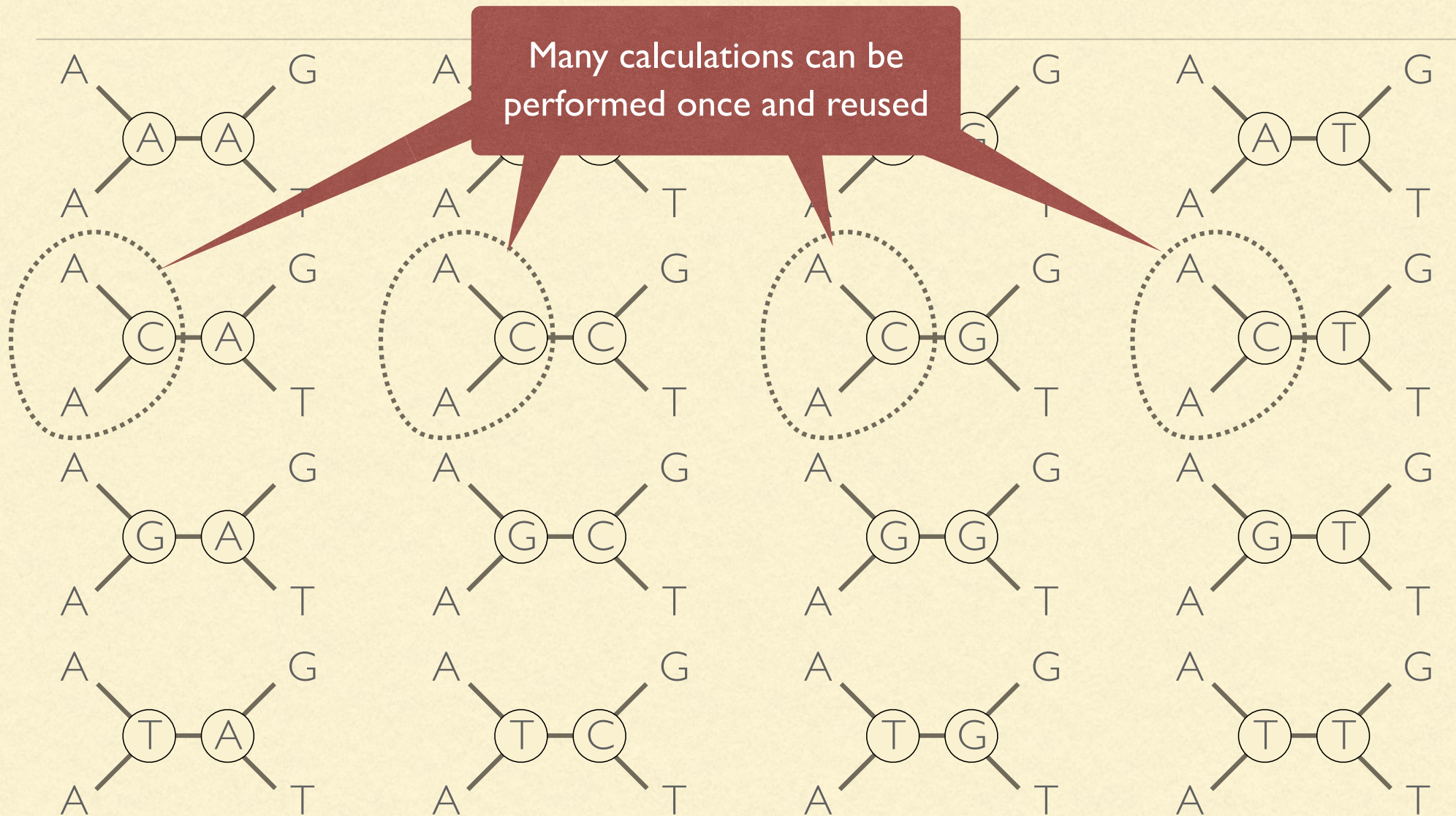


Pruning algo

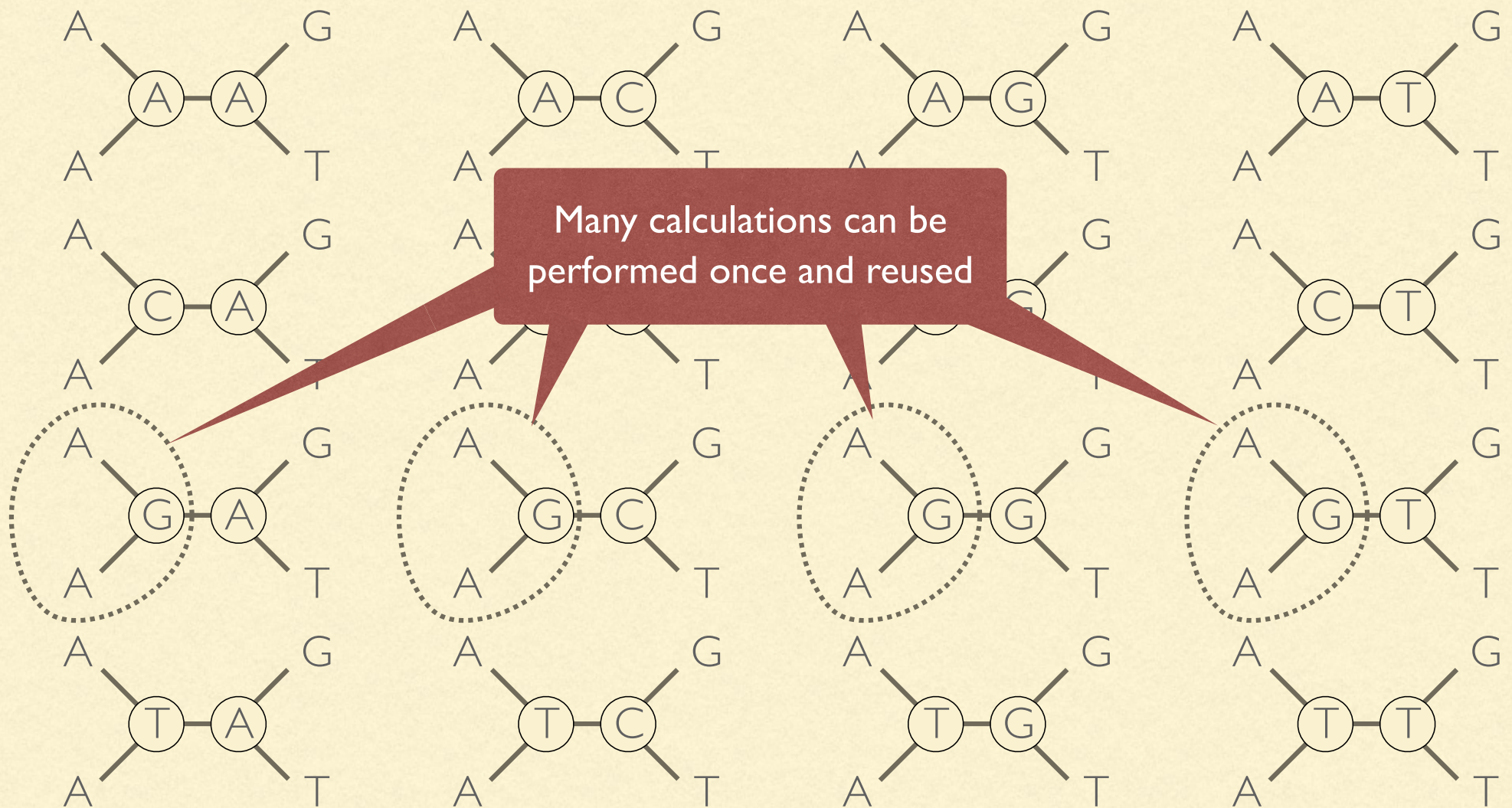
Many calculations can be performed once and reused



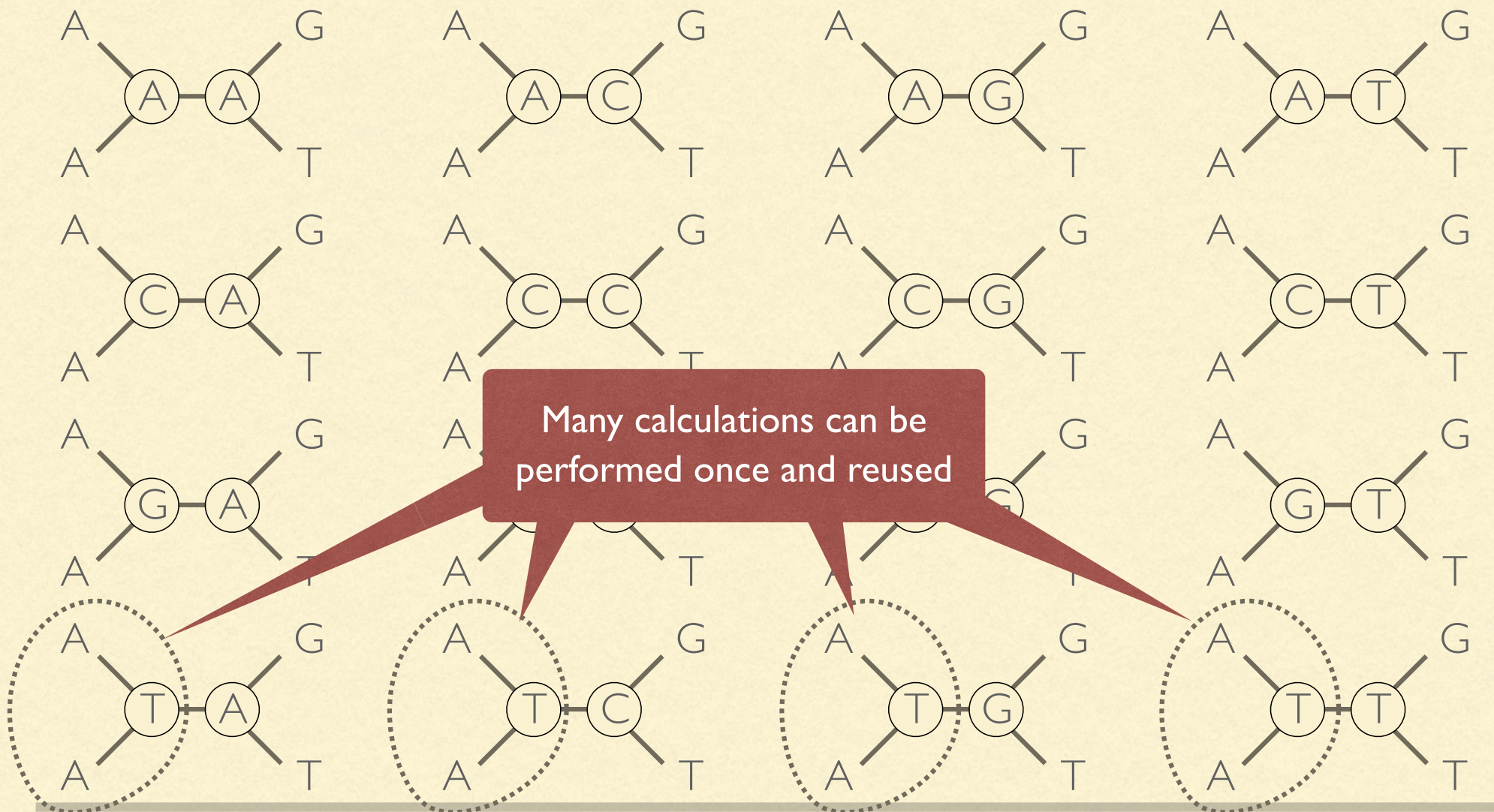
Pruning algorithm



Pruning algorithm



Pruning algorithm



Total likelihood

$$L = L_1 L_2 \cdots L_n$$

↑ ↑ ↑
site 1 site 2 site n


$$\log L = \log L_1 + \log L_2 + \cdots + \log L_n$$

Jukes and Cantor (1969)

JC69 model

to:

Parameters: β

from:

$$\begin{array}{c} \text{A} \\ \text{C} \\ \text{G} \\ \text{T} \end{array} \begin{bmatrix} -3\beta & \beta & \beta & \beta \\ \beta & -3\beta & \beta & \beta \\ \beta & \beta & -3\beta & \beta \\ \beta & \beta & \beta & -3\beta \end{bmatrix}$$

Kimura (1980)

K80 (or K2P) model

Parameters: α, β

$$\begin{array}{c} \text{A} \\ \text{C} \\ \text{G} \\ \text{T} \end{array} \begin{bmatrix} \text{A} & \text{C} & \text{G} & \text{T} \\ -\alpha - 2\beta & \beta & \alpha & \beta \\ \beta & -\alpha - 2\beta & \beta & \alpha \\ \alpha & \beta & -\alpha - 2\beta & \beta \\ \beta & \alpha & \beta & -\alpha - 2\beta \end{bmatrix}$$

Kimura (1980)

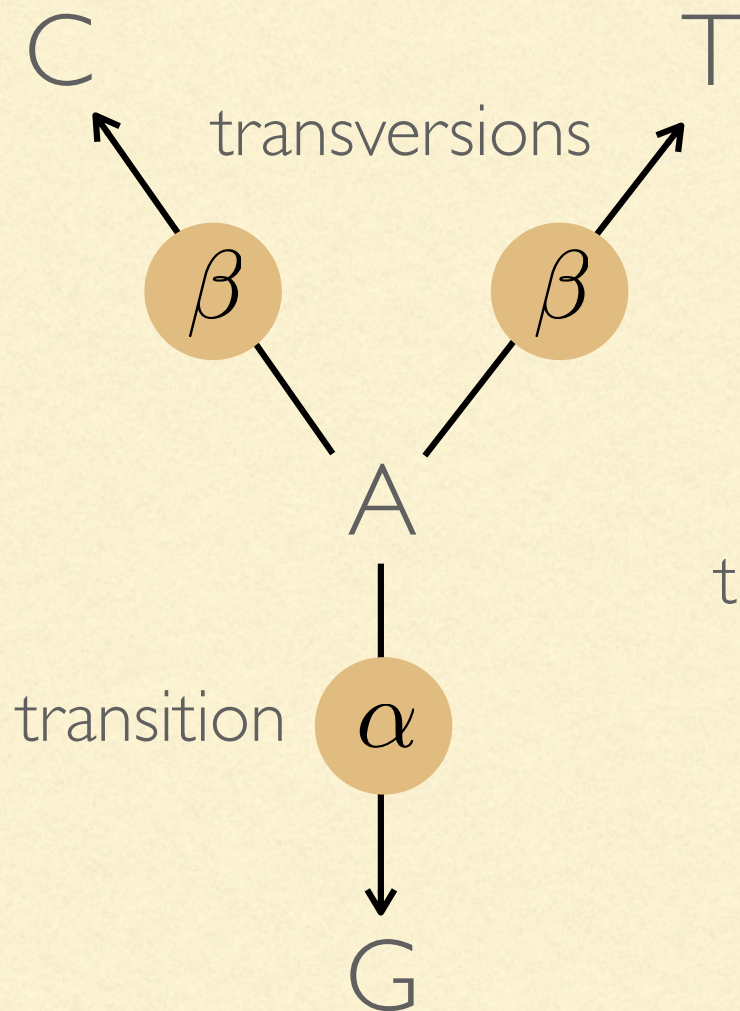
K80 (or K2P) model

$$\kappa = \alpha/\beta$$

Parameters: κ, β

	A	C	G	T
A	$-\beta(\kappa + 2)$	β	$\kappa\beta$	β
C	β	$-\beta(\kappa + 2)$	β	$\kappa\beta$
G	$\kappa\beta$	β	$-\beta(\kappa + 2)$	β
T	β	$\kappa\beta$	β	$-\beta(\kappa + 2)$

Transition-transversion (rate) ratio



transition rate = α

transversion rate = β

assume $\alpha = \beta$

transition-transversion rate ratio = 1.0

transition-transversion ratio = 0.5

Felsenstein (1981)

F81 model

Parameters: μ, π_A, π_C, π_G

	A	C	G	T
A	$-\mu(1 - \pi_A)$	$\pi_C \mu$	$\pi_G \mu$	$\pi_T \mu$
C	$\pi_A \mu$	$-\mu(1 - \pi_C)$	$\pi_G \mu$	$\pi_T \mu$
G	$\pi_A \mu$	$\pi_C \mu$	$-\mu(1 - \pi_G)$	$\pi_T \mu$
T	$\pi_A \mu$	$\pi_C \mu$	$\pi_G \mu$	$-\mu(1 - \pi_T)$

JC69 is a special case of F81

	A	C	G	T		A	C	G	T
A	$-\frac{3}{4}\mu$	$\frac{1}{4}\mu$	$\frac{1}{4}\mu$	$\frac{1}{4}\mu$	$\left[\begin{array}{cccc} -3\beta & \beta & \beta & \beta \\ \beta & -3\beta & \beta & \beta \\ \beta & \beta & -3\beta & \beta \\ \beta & \beta & \beta & -3\beta \end{array} \right]$	-3β	β	β	β
C	$\frac{1}{4}\mu$	$-\frac{3}{4}\mu$	$\frac{1}{4}\mu$	$\frac{1}{4}\mu$		β	-3β	β	β
G	$\frac{1}{4}\mu$	$\frac{1}{4}\mu$	$-\frac{3}{4}\mu$	$\frac{1}{4}\mu$		β	β	-3β	β
T	$\frac{1}{4}\mu$	$\frac{1}{4}\mu$	$\frac{1}{4}\mu$	$-\frac{3}{4}\mu$		β	β	β	-3β

$$\beta = \frac{1}{4}\mu$$

Hasegawa, Kishino, and Yano (1985)

HKY85 model

Parameters: $\mu, \kappa, \pi_A, \pi_C, \pi_G$

these are global
parameters
(apply to all
edge lengths)

one parameter in each model is
associated with the length of an edge

	A	C	G	T
A	$-\mu(\pi_C + \pi_G\kappa + \pi_T)$	$\pi_C\mu$	$\pi_G\mu\kappa$	$\pi_T\mu$
C	$\pi_A\mu$	$-\mu(\pi_A + \pi_G + \pi_T\kappa)$	$\pi_G\mu$	$\pi_T\mu\kappa$
G	$\pi_A\mu\kappa$	$\pi_C\mu$	$-\mu(\pi_A\kappa + \pi_C + \pi_T)$	$\pi_T\mu$
T	$\pi_A\mu$	$\pi_C\mu\kappa$	$\pi_G\mu$	$-\mu(\pi_A + \pi_C\kappa + \pi_G)$

Hasegawa, Kishino, and Yano (1985)

HKY85 model

Parameters: μ , κ , π_A , π_C , π_G

sum of the circled rates equals the total rate given that we start with an A

	A	C	G	T
A	$-\mu(\pi_C + \pi_G\kappa + \pi_T)$	$\pi_C\mu$	$\pi_G\mu\kappa$	$\pi_T\mu$
C	$\pi_A\mu$	$-\mu(\pi_A + \pi_G + \pi_T\kappa)$	$\pi_G\mu$	$\pi_T\mu\kappa$
G	$\pi_A\mu\kappa$	$\pi_C\mu$	$-\mu(\pi_A\kappa + \pi_C + \pi_T)$	$\pi_T\mu$
T	$\pi_A\mu$	$\pi_C\mu\kappa$	$\pi_G\mu$	$-\mu(\pi_A + \pi_C\kappa + \pi_G)$

Hasegawa, Kishino, and Yano (1985)

HKY85 model

Parameters: $\mu, \kappa, \pi_A, \pi_C, \pi_G$

The diagonal element conveniently equals the negative of the total rate away from A

$$\begin{array}{c} \text{A} \\ \text{C} \\ \text{G} \\ \text{T} \end{array} \begin{bmatrix} \text{A} & \text{C} & \text{G} & \text{T} \\ \begin{array}{c} -\mu(\pi_C + \pi_G\kappa + \pi_T) \\ \pi_A\mu \\ \pi_A\mu\kappa \\ \pi_A\mu \end{array} & \begin{array}{c} \pi_C\mu \\ -\mu(\pi_A + \pi_G + \pi_T\kappa) \\ \pi_C\mu \\ \pi_C\mu\kappa \end{array} & \begin{array}{c} \pi_G\mu\kappa \\ \pi_G\mu \\ -\mu(\pi_A\kappa + \pi_C + \pi_T) \\ \pi_G\mu \end{array} & \begin{array}{c} \pi_T\mu \\ \pi_T\mu\kappa \\ \pi_T\mu \\ -\mu(\pi_A + \pi_C\kappa + \pi_G) \end{array} \end{array} \right]$$