Statistics 203: Introduction to Regression and Analysis of Variance Course review

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Today

■ Review / overview of what we learned.



General themes in regression models

- Specifying regression models.
 - What is the joint (conditional) distribution of all outcomes given all covariates?
 - Are outcomes independent (conditional on covariates)? If not, what is an appropriate model?
- Fitting the models.
 - Once a model is specified how are we going to estimate the parameters?
 - Is there an algorithm or some existing software to fit the model?
- Comparing regression models.
 - ◆ Inference for coefficients in the model: are some zero (i.e. is a smaller model better?)
 - What if there are two competing models for the data? Why would one be preferable to the other?
 - What if there are many models for the data? How do we compare models for the data?



Simple linear regression model

- Only one covariate

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i, \qquad 1 \le i \le n$$

■ Errors ε are independent $N(0, \sigma^2)$.



Multiple linear regression model

$$Y_i = \beta_0 + \beta_1 X_{i1} + \dots + \beta_{p-1} X_{i,p-1} + \varepsilon_i, \qquad 1 \le i \le n$$

- Errors ε are independent $N(0, \sigma^2)$.
- \blacksquare β_i 's: (partial) regression coefficients.
- Special cases: polynomial / spline regression models where extra columns are functions of one covariate.



ANOVA & categorical variables

- Generalization of two-sample tests
- One-way (fixed)

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}, \qquad 1 \le i \le r, 1 \le j \le n$$

 α 's are constants to be estimated. Errors ε_{ij} are independent $N(0, \sigma^2)$.

■ Two-way (fixed):

$$Y_{ijk} = \mu + \alpha_i + \beta_j + (\alpha \beta)_{ij} + \varepsilon_{ijk}, \qquad 1 \le i \le r, 1 \le j \le m, 1 \le k \le n$$

 α 's, β 's, $(\alpha\beta)$'s are constants to be estimated. Errors ε_{ijk} are independent $N(0,\sigma^2)$.

■ Experimental design: when balanced layouts are impossible, which is the best design?



Generalized linear models

- Non-Gaussian errors.
- Binary outcomes: logistic regression (or probit).
- Count outcomes: Poisson regression.
- A "link" and "variance" function determine a GLM.
- Link:

$$g(\mathbb{E}(Y_i)) = g(\mu_i) = \eta_i = x_i \beta = X_{i0} \beta_0 + \dots + X_{i,p-1} \beta_{p-1}.$$

Variance function:

$$Var(Y_i) = V(\mu_i).$$



Nonlinear regression models

- Regression function depends on parameters in a nonlinear fashion.

$$Y_i = f(X_{i1}, \dots, X_{ip}; \theta_1, \dots, \theta_q) + \varepsilon_i, \qquad 1 \le i \le n$$

■ Errors ε_i are independent $N(0, \sigma^2)$.



Robust regression

Suppose that we have additive noise, but not Gaussian. Likelihood of the form

$$L(\beta|Y, X_0, \dots, X_{p-1}) \propto \exp\left(-\rho\left(\frac{Y - \sum_{j=0}^{p-1} \beta_j X_j}{s}\right)\right)$$

Leads to robust regression

$$\sum_{i=1}^{n} \rho \left(\frac{Y_i - \sum_{j=0}^{p-1} X_{ij} \beta_j}{s} \right).$$

Can downweight residuals with bigger tails than normal random variables.



Random & mixed effects ANOVA

- When the levels of the categorical variables in an ANOVA are a sample from a population, effects should be treated as random.
- One-way (random):

$$Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}, \qquad 1 \le i \le r, 1 \le j \le n$$

where $\alpha_i \sim N(0, \sigma_{\alpha}^2)$ are random, independent of the errors ε_{ij} which are independent $N(0, \sigma^2)$.

■ Introduces correlation in the *Y*'s:

$$Cov(Y_{ij}, Y_{i'j'}) = \delta_{ii'} \left(\sigma_{\alpha}^2 + \delta_{jj'}\sigma^2\right).$$



Mixed linear models

- Essentially a model of covariance between observations based on "subject" effects.
- General form:

$$Y_{n\times 1} = X_{n\times p}\beta_{p\times 1} + Z_{n\times q}\gamma_{q\times 1} + \varepsilon_{n\times 1}$$

where

- $\varepsilon \sim N(0, \sigma^2 I)$;
- $\gamma \sim N(0, D)$ for some covariance D.
- In this model

$$Y \sim N(X\beta, ZDZ' + \sigma^2 I).$$

Covariance is modelled through "random effect" design matrix Z and covariance D.



Time series regression models

Another model of covariance between observations, based on dependence in time.

$$Y_{n\times 1} = X_{n\times p}\beta_{p\times 1} + \varepsilon_{n\times 1}$$

- In these models, $\varepsilon \sim N(0, \Sigma)$ where the covariance Σ depends on what kind of time series model is used (i.e. which ARMA(p,q) model?).
- **Example**, if ε is AR(1) with parameter ρ then

$$\Sigma_{ij} = \sigma^2 \rho^{|i-j|}.$$



Functional linear model

- We talked about a functional two-sample *t*-test.
- General form

$$Y_{i,t} = \beta_{0,t} + \sum_{j=1}^{p-1} X_{ij}\beta_{j,t} + \varepsilon_{i,t}$$

where the noise $\varepsilon_{i,t}$ is a random function, independent across "observation" (curves) $Y_{i,\cdot\cdot}$.

■ Parameter estimates are curves: leads to nice inference problems for smooth random curves.



Least squares

■ Multiple linear regression – OLS

$$\widehat{\beta} = (X^t X)^{-1} X^t Y.$$

■ Non-constant variance but independent – WLS

$$\widehat{\beta} = (X^t W X)^{-1} X^t W Y, \qquad W_i = 1/\mathsf{Var}(Y_i)$$

■ General correlation – GLS

$$\widehat{\beta} = (X^t \Sigma^{-1} X)^{-1} X^t \Sigma^{-1} Y, \quad \text{Cov}(Y) = \sigma^2 \Sigma.$$



Maximum likelihood

- In the Gaussian setting, with Σ known least squares is MLE.
- In other cases, we needed iterative techniques to solve MLE:
 - nonlinear regression: iterative projections onto the tangent space;
 - ullet robust regression: IRLS with weights determined by $\psi=\rho'$
 - generalized linear models: IRLS, Fisher scoring
 - time series regression models: two-stage procedure approximates MLE (can iterate further, though)
 - mixed models: similar techniques (though we skipped the details)



Diagnostics: influence and outliers

- Diagnostic plots:
 - Added variable plots.
 - QQplot.
 - Residuals vs. fitted.
 - Standardized residuals vs. fitted.
- Measures of influence
- Cook's distance.
- \blacksquare DFFITS.
- \blacksquare DFBETAS.
- Outlier test with Bonferroni correction.
- Techniques are most developed for multiple linear regression model, but some can be generalized (using "whitened" residuals).



Penalized regression

- We looked at ridge regression, too.
- A generic example of the "bias-variance" tradeoff in statistics.
- Minimize

$$SSE_{\lambda}(\beta) = \sum_{i=1}^{n} \left(Y_i - \sum_{j=1}^{p-1} X_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^{p-1} \beta_j^2.$$

- Other penalties possible: basic idea is that the penalty is a measure of "complexity" of the model.
- Smoothing spline: ridge regression for scatterplot smoothers.



Hypothesis tests: multiple linear regression

■ Multiple linear regression: model R has j less coefficients than model F — equivalently there are j linear constraints on β 's.

$$F = rac{\frac{SSE(R) - SSE(F)}{j}}{\frac{SSE(F)}{n-p}}$$
 $\sim F_{j,n-p}(ext{if } H_0 ext{ is true})$

■ Reject $H_0: R$ is true at level α if $F > F_{1-\alpha,j,n-p}$.



Hypothesis tests: general case

- Other models: $DEV(\mathcal{M}) = -2 \log L(\mathcal{M})$ replaces $SSE(\mathcal{M})$.
- Difference $D = DEV(R) DEV(F) \sim \chi_j^2$ (asymptotically).
- Denominator in the F statistic is usually either known or based on something like Pearson's X^2 :

$$\widehat{\phi} = \frac{1}{n-p} \sum_{i=1}^{n} r_i(F)^2.$$

In general, residuals are "whitened"

$$r(\mathcal{M}) = \Sigma(\mathcal{M})^{-1/2} (Y - X\widehat{\beta}(\mathcal{M})).$$

■ Reject $H_0: R$ is true at level α if $D > \chi^2_{1-\alpha,n-p}$.



Model selection: AIC, BIC, stepwise

- Best subsets regression (leaps): adjusted R^2 , C_p .
- Akaike Information Criterion (AIC)

$$AIC(\mathcal{M}) = -2\log L(\mathcal{M}) + 2 \cdot p(\mathcal{M}).$$

in model \mathcal{M} evaluated at the MLE (Maximum Likelihood Estimators).

Schwarz's Bayesian Information Criterion (BIC)

$$BIC(\mathcal{M}) = -2\log L(\mathcal{M}) + p(\mathcal{M}) \cdot \log n$$

■ Penalized regression can be thought of as model selection as well: choosing the "best" penalty parameter on the basis of

$$GCV(\mathcal{M}) = \frac{1}{\mathsf{Tr}(S(\mathcal{M}))} \sum_{i=1}^{n} \left(Y_i - \widehat{Y}_i(\mathcal{M}) \right)^2$$

where

$$\widehat{Y}(\mathcal{M}) = S(\mathcal{M})Y.$$



That's it!

Thanks!