

Statistics 203: Introduction to Regression and Analysis of Variance

Multiple Linear Regression + Multivariate Normal

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Today

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- Multiple linear regression
- Model
- Design matrix
- Fitting the model: SSE
- Solving for $\hat{\beta}$
- Multivariate normal
- Multivariate normal
- Projections
- Projections
- Identity covariance, projections & χ^2
- Properties of multiple regression estimates

- Multiple linear regression
- Some proofs: multivariate normal distribution.



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- Specifying the model.
- Fitting the model: least squares.
- Interpretation of the coefficients.



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- Basically, rather than one predictor, we more than one predictor, say $p - 1$.



$$Y_i = \beta_0 + \beta_1 X_{i1} + \cdots + \beta_{p-1} X_{i,p-1} + \varepsilon_i$$

- Errors $(\varepsilon_i)_{1 \leq i \leq n}$ are assumed independent $N(0, \sigma^2)$, as in simple linear regression.
- Coefficients are called (partial) regression coefficients because they “allow” for the (partial) effect of other variables.



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- Define the $n \times p$ matrix

$$X = \begin{pmatrix} 1 & X_{11} & X_{12} & \dots & X_{1,p-1} \\ \vdots & \vdots & \ddots & \vdots & \\ 1 & X_{n1} & X_{n2} & \dots & X_{n,p-1} \end{pmatrix}$$

and the column vectors $X_j = (X_{1j}, \dots, X_{nj})$.

- Model can be expressed as

$$Y = X\beta + \varepsilon.$$



Fitting the model: SSE

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- Just as in simple linear regression, model is fit by minimizing

$$SSE(\beta_0, \dots, \beta_p) = \sum_{i=1}^n (Y_i - (\beta_0 + \sum_{j=1}^p \beta_j X_{ij}))^2.$$

- Minimizers: $\hat{\beta} = (\hat{\beta}_0, \dots, \hat{\beta}_p)$ are the “least squares estimates” and are also normally distributed as in simple linear regression.
- Explicit expression when X is full rank (next slide)

$$\hat{\beta} = (X^t X)^{-1} X^t Y.$$



Solving for $\hat{\beta}$

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■ Normal equations

$$\left. \frac{\partial}{\partial \beta_j} SSE \right|_{\hat{\beta}} = -2 \left(Y - X\hat{\beta} \right)^t X_j = 0, \quad 0 \leq j \leq p-1.$$

■ Equivalent to

$$(Y - X\hat{\beta})^t X = 0$$

$$Y^t X = \hat{\beta}^t (X^t X)$$

$$X^t Y = (X^t X) \hat{\beta}$$

$$\hat{\beta} = (X^t X)^{-1} X^t Y$$

■ Properties: after some facts about multivariate normal random vectors.



Multivariate normal

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- $Z = (Z_1, \dots, Z_n) \in \mathbb{R}^n$ is *multivariate Gaussian* if, for every $\alpha = (\alpha_1, \dots, \alpha_n) \in \mathbb{R}^n$, $\langle \alpha, Z \rangle = \sum_{i=1}^n \alpha_i Z_i$ is Gaussian.

- Mean vector: $\mu \in \mathbb{R}^n$ has components

$$\mu_i = \mathbb{E}(Z_i).$$

- Covariance matrix: Σ a non-negative definite $n \times n$ matrix

$$\Sigma_{ij} = \text{Cov}(Z_i, Z_j).$$

- Non-negative (positive) definite: for any $\alpha \in \mathbb{R}^n$

$$\alpha^t \Sigma \alpha \geq (>) 0.$$

- We write $Z \sim N(\mu, \Sigma)$.



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- For any $m \times n$ matrix A

$$AZ \sim N(A\mu, A\Sigma A^t).$$

- If Σ is positive definite then the density of Z is

$$f_Z(z) = (2\pi)^{-n/2} |\Sigma|^{-1/2} e^{-(z-\mu)^t \Sigma^{-1} (z-\mu)/2}.$$

- If Σ is only non-negative definite (i.e. rank of $\Sigma < n$) then Z lives on a lower dimensional space and has no density on \mathbb{R}^n .



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- If an $n \times n$ matrix P satisfies
 - ◆ $P^2 = P$ (idempotent)
 - ◆ $P = P^t$ (symmetric)then P is a projection matrix.
- That is, there exists a subspace $L \subset \mathbb{R}^n$ of dimension $r \leq n$ such that for any $z \in \mathbb{R}^n$ Pz is the projection of z onto L . We write P_L to denote the subspace L projects onto.
- Given any orthonormal basis $\{w_1, \dots, w_r\}$ of L

$$P_L z = \sum_{j=1}^r \langle z, w_j \rangle w_j.$$

- If P_L is a projection matrix then

$$I - P_L = P_{L^\perp}$$

is also a projection matrix which projects onto L^\perp , the orthogonal complement of L in \mathbb{R}^n .



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- Let $\{X_1, \dots, X_r\}$ be a set of linearly independent vectors in \mathbb{R}^n and

$$X = \begin{pmatrix} X_1 & X_2 & \dots & X_r \end{pmatrix}$$

is the $n \times r$ matrix made by “concatenating” the X_i ’s.

- If

$$L = \text{span}(X_1, \dots, X_r)$$

is the subspace of \mathbb{R}^n spanned by $\{X_1, \dots, X_r\}$ then

$$P_L = X(X^t X)^{-1} X^t.$$



Identity covariance, projections & χ^2

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- If $\Sigma = \sigma^2 I$ and L is a subspace of \mathbb{R}^n then

$$P_L Z \sim N(P_L \mu, \sigma^2 P_L)$$

where P_L is the projection matrix onto L .

- If $P_L \mu = 0$ then

$$\|P_L Z\|^2 \sim \chi_{\dim(L)}^2$$

and

$$\dim(L) = \text{Tr}(P_L).$$

- If $P_L \mu \neq 0$ then

$$\|P_L Z\|^2 \sim \chi_{\dim(L), \|P_L \mu\|^2}^2$$

has a *non-central* χ^2 distribution.



Properties of multiple regression estimates

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■ $\hat{\beta} \sim N(\beta, \sigma^2(X^t X)^{-1}).$

■ As in simple regression

$$\hat{\sigma}^2 = MSE = \frac{SSE}{n - p} \sim \sigma^2 \cdot \frac{\chi_{n-p}^2}{n - p}$$

independent of $\hat{\beta}$.

■ The least squares estimates are “minimum variance” linear unbiased estimators. (Gauss-Markov theorem)