

Food Inspections: Modeling Risk with Ordinal Logistic Regression

MSCA 31010 Linear & Nonlinear Models

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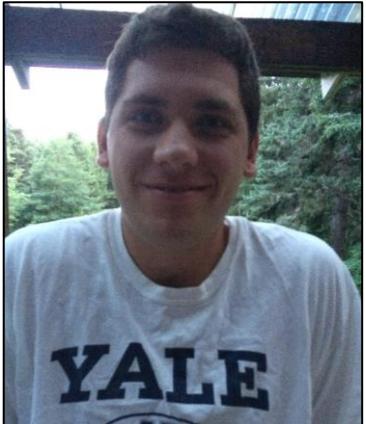
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Our Team



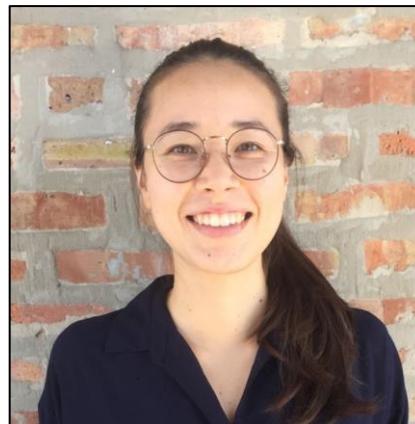
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Agenda

1. Executive Summary
2. Initial Exploratory Analysis
3. Ordinal Logistic Model
4. Our Model
5. Findings & Insights
6. Lessons Learned
7. Questions

Executive Summary

The Chicago Department of Public Health regularly inspects businesses serving food to ensure restaurants and other food retail outlets are following safe food handling procedures. Inspections determine that safeguards are in place to protect food from contamination by food handlers, cross-contamination, and contamination from other sources in the restaurant.

We created an **Ordinal Logistic Regression Model** to explore our business case research question.

| Our Business Case | Application | Data | Our Approach |
|---|--|--|---|
| What factors influence risk assignment for restaurants which pass their inspections? | <p>Stakeholders: Restaurant owners and managers</p> <p>Use Case: Determine which violations to prioritize to minimize the severity of their risk assignment during inspections</p> | <ul style="list-style-type: none">• 43,573 Rows, 17 Attributes• July 1, 2018 to February 19, 2021• Data Types: Categorical and Ordinal | <ul style="list-style-type: none">• Initial Exploratory Analysis• Design & Fit the Model• Extract Insights• Document Lessons Learned |

I. Initial Exploratory Analysis

Initial exploratory analysis revealed an establishment could pass an inspection but still be categorized as high risk. The dataset recorded Risk as an ordered factor which could be affected by multiple categorical variables.

Categorical Data Types

Ordinal (High > Med > Low)

| Inspection ID | DBA Name | AKA Name | License # | Facility Type | Risk | Address | City | State | Zip | Inspection Date | Inspection Type | Results | Violations |
|---------------|---------------------|---------------------|-----------|---------------|-----------------|-------------------------|---------|-------|---------|-----------------|-----------------------|--------------------|---|
| 2184275 | MCCB | MCCB | 2600357.0 | Restaurant | Risk 1 (High) | 2138 S ARCHER AVE | CHICAGO | IL | 60616.0 | 7/3/2018 | License | Pass w/ Conditions | NaN |
| 2184211 | JAMES KITCHEN + BAR | JAMES HOTEL CHICAGO | 1884681.0 | Restaurant | Risk 1 (High) | 616 N RUSH ST | CHICAGO | IL | 60611.0 | 7/3/2018 | Canvass | Out of Business | NaN |
| 2182184 | CAFE EL TAPATIO | CAFE EL TAPATIO | 2432567.0 | Restaurant | Risk 1 (High) | 3400-3402 N ASHLAND AVE | CHICAGO | IL | 60657.0 | 7/3/2018 | Canvass Re-Inspection | Pass | NaN |
| 2170218 | ROYS LUNCH BAG | ROYS LUNCH BAG | 46653.0 | Restaurant | Risk 1 (High) | 403 E 71ST ST | CHICAGO | IL | 60619.0 | 7/3/2018 | Canvass Re-Inspection | Pass | NaN |
| 2182196 | MCDONALD'S | MCDONALD'S | 1840645.0 | Restaurant | Risk 2 (Medium) | 6740 N CLARK ST | CHICAGO | IL | 60626.0 | 7/3/2018 | Canvass | Pass w/ Conditions | 5. PROCEDURES FOR RESPONDING TO VOMITING AND D... |

I. Initial Exploratory Analysis

We transformed the dataset so each violation would be listed with an inspection was recorded as its own categorical variable (“0” indicated a pass and “1” indicating a hit or violation in that inspection area). This new format would then be compatible with building an ordinal logistic regression model.



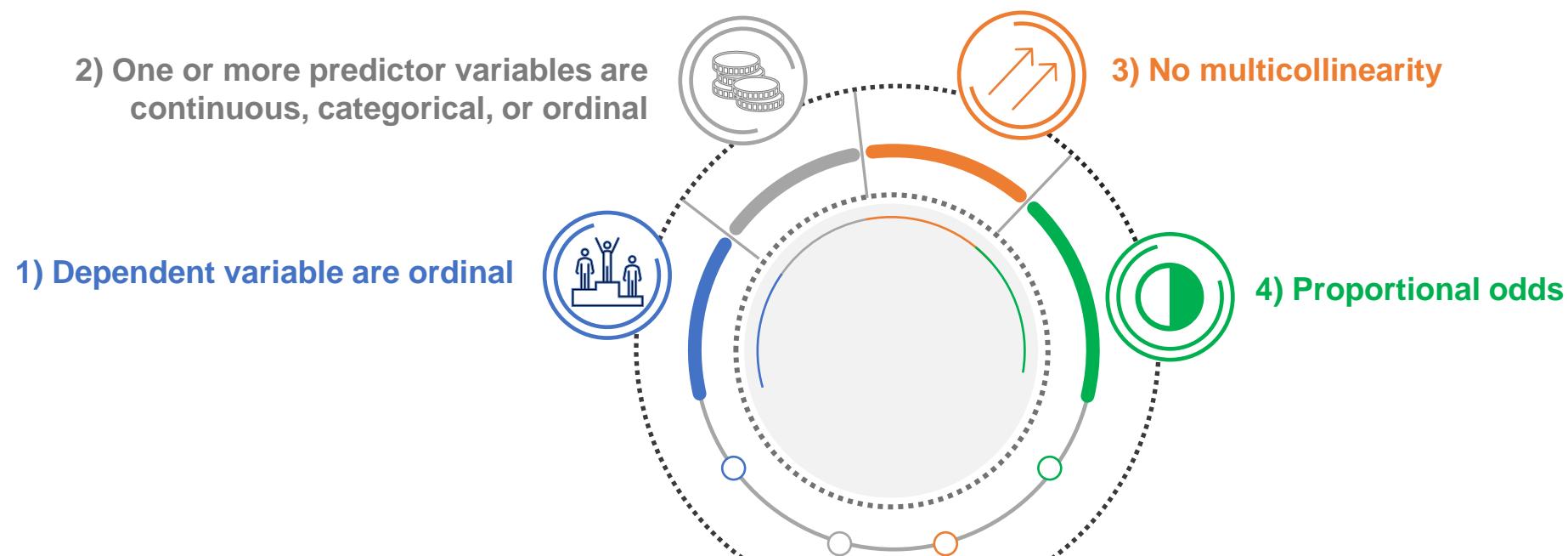
| Risk | Facility Type | Inspection Date | Inspection Type | Violations |
|------|-----------------|-----------------|-----------------|------------|
| 0 | Risk 2 (Medium) | Restaurant | 7/3/2018 | Canvass |
| 1 | Risk 2 (Medium) | Restaurant | 7/6/2018 | License |
| 2 | Risk 1 (High) | Restaurant | 7/9/2018 | Complaint |
| 3 | Risk 1 (High) | Restaurant | 7/9/2018 | Canvass |
| 4 | Risk 1 (High) | Restaurant | 7/10/2018 | Canvass |

| Risk | Facility Type | Inspection Date | Inspection Type | Supervision | EmployeeHealth | HygenicPractices | HandContamination | ApprovedSource | Contamination |
|------|---------------|-----------------|-----------------|-------------|----------------|------------------|-------------------|----------------|---------------|
| 0 | Med | Restaurant | Summer | Canvass | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |
| 1 | High | Restaurant | Summer | Canvass | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |
| 2 | High | Restaurant | Summer | Canvass | 0.0 | 1.0 | 0.0 | 0.0 | 0.0 |
| 3 | High | Restaurant | Summer | Canvass | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |
| 4 | High | Restaurant | Summer | Canvass | 0.0 | 0.0 | 0.0 | 1.0 | 0.0 |

Ordinal Logistic Regression (OLR)

Ordinal Logistic Regression is used to predict an ordinal dependent variable given one or more independent variables.

The Ordinal Logistic Model must meet four assumptions:



Ordinal Logistic Regression (OLR)

- **Multinomial (MLR) vs. Ordinal Logistic Regression Models**

- Key difference: the response variable is nominal vs ordinal
- mlogit(y ~., data=learn)
- In a MLR model the coefficients represent intercepts for one response variable against each other.
- In an OLR model the coefficients represent "cut points", which represent where one variable switches to the next in a hierarchical order.

Multinomial

$$P(Y \leq j|x_i) = \sum_{k \leq j} P(Y = k|x_i)$$

$$\log \left(\frac{P(Y_i \leq j|\mathbf{x}_i)}{1 - P(Y_i \leq j|\mathbf{x}_i)} \right) = \theta_j - \mathbf{x}_i \beta$$

Ordinal

$P(Y \leq j)$ is the cumulative probability of Y less than or equal to a specific category j .

$$\log \frac{P(Y \leq j)}{P(Y > j)} = \text{logit}(P(Y \leq j))$$

$$\text{logit}(P(Y \leq j)) = \beta_0 j + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p$$

```
> summary(m1)
Call: nnet::multinom(formula = V1 ~ GarbageInfo)

Coefficients:
            (Intercept) GarbageInfo
red          -0.6011338 -0.1331142
yellow       -0.3221203 -0.5995860

Std. Errors:
            (Intercept) GarbageInfo
red          0.5164521  0.8448432
yellow       0.4932972  0.8380932
...

> summary(m2)
...
Call: MASS::polr(formula = V2 ~ GarbageInfo)

Coefficients:
              Value Std. Error t value
GarbageInfo -0.2181    0.6431 -0.3391

Intercepts:
              Value Std. Error t value
green|yellow -0.1541   0.3910  -0.3941
yellow|red    0.9856   0.4045   2.4364
...
```

Image from StackExchange

P = probability of a response

j = response level

β = coefficient of explanatory variable

x = explanatory (independent) variable

p = number of predictor

Our Model

Our ordered response variable can be modeled using the cumulative logistic model:

$$\text{logit}(P(Y \leq j)) = \beta_0 j + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p$$

Where the response Y (Risk) has j-ordered levels (low, med, high) described by a combination of explanatory variables (inspection season and types of health inspection focus areas).

The model will be described by two equations:

$$\log \frac{P(Y \leq L)}{P(Y > M)} = \text{logit}(P(Y \leq L)) = \beta_{0L|M} + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p$$

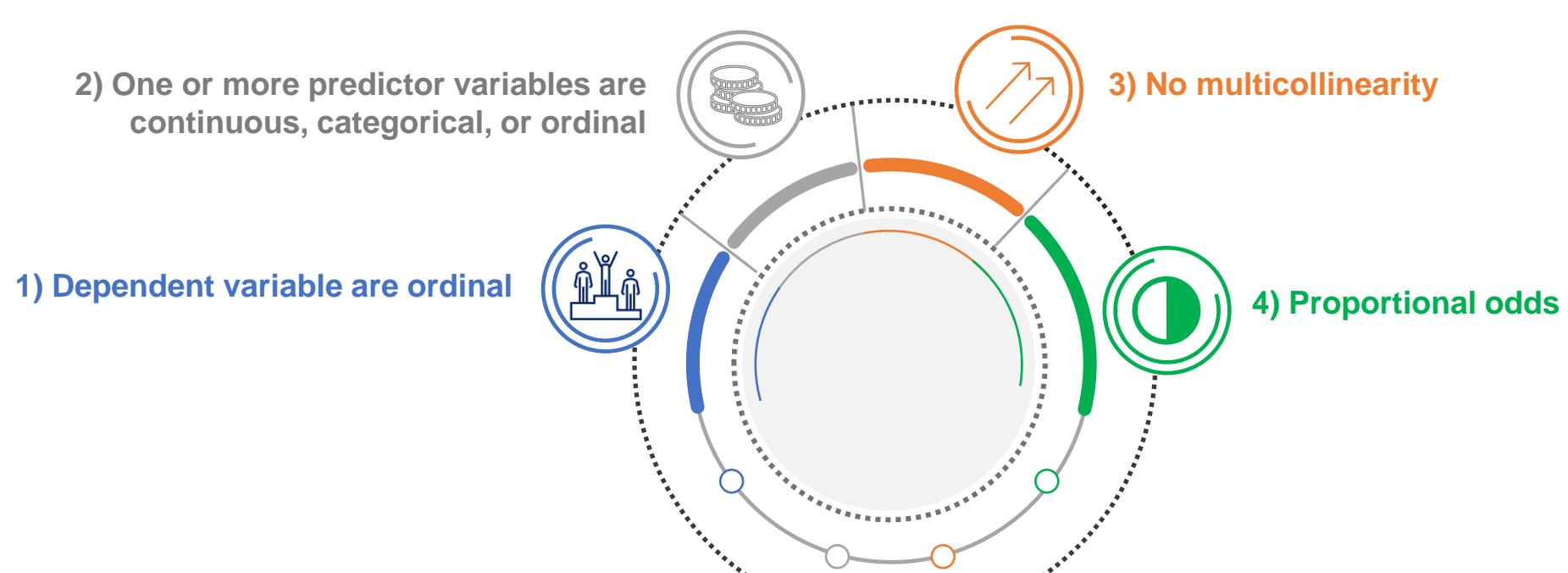
$$\log \frac{P(Y \leq M)}{P(Y > H)} = \text{logit}(P(Y \leq M)) = \beta_{0M|H} + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p$$

And the coefficients of the model can be interpreted as an odds ratio holding all other variables constant, $e^{x_p \beta_j}$.

| Call: | | | |
|---|----------|------------|----------|
| polr(formula = Risk ~ ., data = dta_mini, Hess = T) | | | |
| Coefficients: | | | |
| | Value | Std. Error | t value |
| Inspection.DateWinter | 0.38477 | 0.15640 | 2.4602 |
| Inspection.DateSpring | -0.40910 | 0.15307 | -2.6727 |
| Inspection.DateSummer | -0.61075 | 0.15436 | -3.9567 |
| HandContamination.L | 0.27371 | 0.12982 | 2.1084 |
| Contamination.L | 0.20936 | 0.17876 | 1.1712 |
| FoodTemperature.L | -0.42486 | 0.14030 | -3.0283 |
| FoodIdentification.L | 0.28734 | 0.13189 | 2.1787 |
| FoodContamination.L | 0.10263 | 0.09566 | 1.0728 |
| ProperUtensils.L | 0.52630 | 0.16659 | 3.1593 |
| Equipment.L | 0.10667 | 0.07631 | 1.3978 |
| PhysicalFacilities.L | 0.06233 | 0.09528 | 0.6542 |
| EmployeeTraining.L | 0.74643 | 0.14920 | 5.0030 |
| Intercepts: | | | |
| | Value | Std. Error | t value |
| Low Med | -7.3898 | 0.5110 | -14.4608 |
| Med High | -2.6670 | 0.2511 | -10.6206 |
| Residual Deviance: 2312.469 | | | |
| AIC: 2340.469 | | | |

II. Model Building – Assumptions

The Ordinal Logistic Model must meet four assumptions:



II. Model Building – Assumptions 1 & 2

| Assumption | Description | Pass |
|--|---|------|
| 1) The dependent variable are ordered | Risk is ordered Low < Med < High | ✓ |
| 2) One or more predictor variables are continuous, categorical, or ordinal | All predictor variables are categorical (Date: Seasons, Violations: Pass and Hit) | ✓ |

| ▲ Risk | Inspection.Date | HandContamination | Contamination | FoodTemperature | FoodIdentification | FoodContamination |
|--------|-----------------|-------------------|---------------|-----------------|--------------------|-------------------|
| 1 Med | Summer | Pass | Pass | Pass | Pass | Pass |
| 2 High | Summer | Pass | Pass | Pass | Pass | Pass |
| 3 High | Summer | Pass | Pass | Pass | Pass | Pass |
| 4 High | Summer | Pass | Pass | Pass | Pass | Pass |
| 5 High | Summer | Hit | Pass | Pass | Pass | Pass |
| 6 High | Summer | Pass | Pass | Pass | Pass | Pass |
| 7 Med | Summer | Pass | Pass | Pass | Pass | Pass |
| 8 High | Summer | Pass | Pass | Pass | Pass | Pass |

II. Model Building – Assumption 3

| Assumption | Description | Pass |
|-------------------------|---|------|
| 3) No multicollinearity | Using the VIFfit() test, we can conclude that no variables have a VIF score > 5, therefore there is no multicollinearity. | ✓ |

Variance Inflation Factor

The VIF estimates how much the variance of a regression coefficient is inflated due to multicollinearity in the model.

- 1 = Not correlated
- 1 - 5 = Moderately correlated
- > 5 = Highly correlated

```
numRisk <- as.numeric(dta$Risk)
VIFfit <- lm(scale(numRisk) ~ ., data = dta_mini)
VIF(VIFfit)
```

| | GVIF | Df | GVIF^(1/(2*Df)) |
|--------------------|----------|----|-----------------|
| Risk | 1.048278 | 2 | 1.011857 |
| Inspection.Date | 1.050227 | 3 | 1.008201 |
| HandContamination | 1.012984 | 1 | 1.006471 |
| Contamination | 1.005303 | 1 | 1.002648 |
| FoodTemperature | 1.013250 | 1 | 1.006603 |
| FoodIdentification | 1.023206 | 1 | 1.011536 |
| FoodContamination | 1.019121 | 1 | 1.009515 |
| ProperUtensils | 1.029838 | 1 | 1.014809 |
| Equipment | 1.008589 | 1 | 1.004285 |
| PhysicalFacilities | 1.020507 | 1 | 1.010202 |
| EmployeeTraining | 1.024369 | 1 | 1.012111 |

II. Model Building – Assumption 4

| Assumption | Description | Pass |
|----------------------|---|------|
| 4) Proportional odds | Brant Test & PoTest indicates that the probabilities are > than our alpha, 0.05 | ✓ |

Brant Test

| Test for | x2 | df | probability |
|-----------------------|------|----|-------------|
| <hr/> | | | |
| Omnibus | 1.49 | 12 | 1 |
| Inspection.DateWinter | 0 | 1 | 1 |
| Inspection.DateSpring | 0 | 1 | 1 |
| Inspection.DateSummer | 0 | 1 | 0.97 |
| HandContamination.L | 0 | 1 | 1 |
| Contamination.L | 0 | 1 | 1 |
| FoodTemperature.L | 1.04 | 1 | 0.31 |
| FoodIdentification.L | 0 | 1 | 1 |
| FoodContamination.L | 0.11 | 1 | 0.74 |
| ProperUtensils.L | 0 | 1 | 1 |
| Equipment.L | 0.64 | 1 | 0.42 |
| PhysicalFacilities.L | 0 | 1 | 0.96 |
| EmployeeTraining.L | 0 | 1 | 1 |
| <hr/> | | | |

Proportional Odds Test

| Tests for Proportional Odds polr(formula = Risk ~ ., data = dta_mini, Hess = T) | | | | | | |
|--|-----------|-----------|-----------|-----------|------|------------|
| | b[polr] | b[>Low] | b[>Med] | chisquare | df | Pr(>Chisq) |
| Overall | | | | | 1.49 | 12 |
| Inspection.DateWinter | 3.85e-01 | 1.86e+01 | 3.81e-01 | 0.00 | 1 | 1.00 |
| Inspection.DateSpring | -4.09e-01 | 1.87e+01 | -4.15e-01 | 0.00 | 1 | 1.00 |
| Inspection.DateSummer | -6.11e-01 | -6.38e-01 | -6.07e-01 | 0.00 | 1 | 0.97 |
| HandContamination.L | 2.74e-01 | 1.27e+01 | 2.71e-01 | 0.00 | 1 | 1.00 |
| Contamination.L | 2.09e-01 | 1.30e+01 | 2.06e-01 | 0.00 | 1 | 1.00 |
| FoodTemperature.L | -4.25e-01 | -1.26e+00 | -4.17e-01 | 1.04 | 1 | 0.31 |
| FoodIdentification.L | 2.87e-01 | 1.26e+01 | 2.84e-01 | 0.00 | 1 | 1.00 |
| FoodContamination.L | 1.03e-01 | -1.67e-01 | 1.02e-01 | 0.11 | 1 | 0.74 |
| ProperUtensils.L | 5.26e-01 | 1.27e+01 | 5.24e-01 | 0.00 | 1 | 1.00 |
| Equipment.L | 1.07e-01 | -4.26e-01 | 1.07e-01 | 0.64 | 1 | 0.42 |
| PhysicalFacilities.L | 6.23e-02 | 2.02e-02 | 6.02e-02 | 0.00 | 1 | 0.96 |
| EmployeeTraining.L | 7.46e-01 | 1.27e+01 | 7.45e-01 | 0.00 | 1 | 1.00 |

III. Model Testing (Train-Test Split)

Now that we know our model passes the Ordinal Logistic Regression model assumptions, we tested our model.

- Training and testing **60:40**
- While an **82% accuracy** is relatively high, we can see the model is only predicting “High” because of the underlying imbalance in the dataset.

Training set: 83.4% accuracy

```
#view frequency table  
table(datatrain$Risk, pred1)  
  
##      pred1  
##      Low Med High  
##  Low     0   0    3  
##  Med     0   0  257  
##  High    0   0 1308  
  
#model test accuracy  
sum(diag(table(datatrain$Risk, pred1)))/sum(table(datatrain$Risk, pred1))  
  
## [1] 0.8341837
```

Testing set: 82% accuracy

```
#view frequency table  
table(datatest$Risk, pred_test)  
  
##      pred_test  
##      Low Med High  
##  Low     0   0    2  
##  Med     0   0  181  
##  High    0   0  863  
  
#model test accuracy  
sum(diag(table(datatest$Risk, pred_test)))/sum(table(datatest$Risk, pred_test))  
  
## [1] 0.8250478
```

Full dta_mini dataset: 83% accuracy

```
#view frequency table  
table(dta_mini$Risk, pred_dta_mini)  
  
##      pred_dta_mini  
##      Low Med High  
##  Low     0   0    5  
##  Med     0   0  438  
##  High    0   0 2171  
  
#model test accuracy  
sum(diag(table(dta_mini$Risk, pred_dta_mini)))/sum(table(dta_mini$Risk, pred_dta_mini))  
  
## [1] 0.8305279
```

III. Model Testing (Oversampling)

We can examine **random oversampling** where duplicates are randomly selected from the minority class (any Risk factor other than “High”).

The overall model accuracy **decreases to 49%** so our OLR model works best by just predicting “High” Risk factors for all inspections because so “Med” and “Low” assignments are less frequent.

Given the drop in accuracy with resampling, we continued using our “m_mini” model in upcoming analysis. The “m_mini” OLR model had an **82%** accuracy.

```
#creating predictions based on test set  
random.oversample.pred <- predict(m_resample, datatest, type = "class")
```

```
#recall the original confusion matrix using the dta_mini_train model and test set  
table(datatest$Risk, pred_test)
```

```
##      pred_test  
##        Low Med High  
##  Low     0   0    2  
##  Med     0   0  181  
##  High    0   0  863
```

```
#confusion matrix after random oversampling  
oversample.table = table(datatest$Risk, random.oversample.pred)  
oversample.table
```

```
##      random.oversample.pred  
##        Low Med High  
##  Low     1   1    0  
##  Med    58  56   67  
##  High   181 228  454
```

```
#viewing accuracy and confusion matrix summary after random oversampling  
sum(diag(oversample.table))/sum(oversample.table)
```

```
## [1] 0.4885277
```

IV. Model Interpretation & Results

The log odd coefficients are not easily interpretable from the model summary. Converting the coefficients into an odds ratio lets us describe the affect of a variable unit change on the response.

- For restaurants with **Inspection Dates in Winter**, the odds of getting a higher risk inspection rating (ex. Med or High vs Low) **is 1.47 times** that of restaurants who get Inspections in other seasons, holding constant all other variables.
- For restaurants with **Inspection Dates in Spring**, the odds of getting a higher risk inspection rating (ex. Med or High vs Low) **is 0.66 times** that of restaurants who get Inspections in other seasons, holding constant all other variables.

Model Summary

```
## polr(formula = Risk ~ ., data = dta_mini, Hess = T)
##
## Coefficients:
##              Value Std. Error t value
## Inspection.DateWinter  0.38477  0.15640  2.4602
## Inspection.DateSpring -0.40910  0.15307 -2.6727
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## FoodIdentification.L  0.28734  0.13189  2.1787
## FoodContamination.L   0.10263  0.09566  1.0728
## ProperUtensils.L     0.52630  0.16659  3.1593
## Equipment.L           0.10667  0.07631  1.3978
## PhysicalFacilities.L  0.06233  0.09528  0.6542
## EmployeeTraining.L    0.74643  0.14920  5.0030
##
## Intercepts:
##             Value Std. Error t value
## Low|Med   -7.3898  0.5110 -14.4608
## Med|High  -2.6670  0.2511 -10.6206
##
## Residual Deviance: 2312.469
## AIC: 2340.469
```

Odds Ratio and Confidence Intervals

| | OR | 2.5 % | 97.5 % |
|--------------------------|--------|--------|--------|
| ## Inspection.DateWinter | 1.4693 | 1.0806 | 1.9967 |
| ## Inspection.DateSpring | 0.6642 | 0.4911 | 0.8954 |
| ## Inspection.DateSummer | 0.5429 | 0.4002 | 0.7334 |
| ## HandContamination.L | 1.3148 | 1.0276 | 1.7116 |
| ## Contamination.L | 1.2329 | 0.8821 | 1.7843 |
| ## FoodTemperature.L | 0.6538 | 0.4994 | 0.8667 |
| ## FoodIdentification.L | 1.3328 | 1.0377 | 1.7425 |
| ## FoodContamination.L | 1.1080 | 0.9214 | 1.3411 |
| ## ProperUtensils.L | 1.6927 | 1.2400 | 2.3898 |
| ## Equipment.L | 1.1126 | 0.9583 | 1.2927 |
| ## PhysicalFacilities.L | 1.0643 | 0.8807 | 1.2801 |
| ## EmployeeTraining.L | 2.1094 | 1.5957 | 2.8705 |

IV. Model Interpretation & Results

Key Takeaways:

- Variables that Increased the Odds of Higher Risk Assignment
 - Inspections in Winter
 - Violations in Hand Contamination, Food Identification, Proper Utensils, and Employee Training
- Inspections in the Spring or Summer reduced the odds of higher risk assignment

Interpretation:

- Inspection date impacts the odds of risk assignment.
 - Odds of risk assignment might increase in Winter because restaurants are more likely to lose power (impacting the facility's ability to maintain health code standards)
 - Restaurants may also be more prepared for inspections in Spring and Summer as those are traditional busy seasons
- The log odds of food temperature violations is < 0 , but this might be tied to less frequent power outages outside of winter months (freezers maintain temperatures more reliably)

Odds Ratio and Confidence Intervals

| | OR | 2.5 % | 97.5 % |
|--------------------------|--------|--------|--------|
| ## Inspection.DateWinter | 1.4693 | 1.0806 | 1.9967 |
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IV. Model Interpretation & Results

Business and stakeholders can't control the date of a random inspection, but could reduce their odds of higher risk assignment by passing all checks in hand contamination, food identification, proper utensils, and employee training:

Preventing Contamination by Hands (HandContamination)

- Hands clean & properly washed
- No bare hand contact with RTE food or pre-approved alternative procedure properly allowed
- Adequate handwashing sinks properly supplied and accessible

Food Identification

- Food properly labeled; original container

Proper Use of Utensils (ProperUtensils)

- In-use utensils: properly stored
- Utensils, equipment & linens: properly stored, dried, & handled
- Single-use/single-service articles: properly stored & used
- Gloves used properly

Employee Training

- All food employees have food handler training
- Allergen training as required

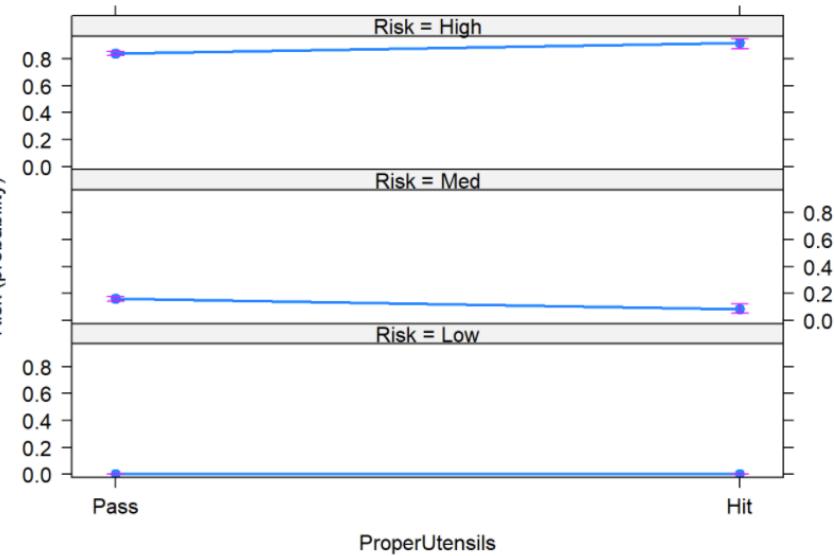
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| ## PhysicalFacilities.L | 1.0643 | 0.8807 | 1.2801 |
| ## EmployeeTraining.L | 2.1094 | 1.5957 | 2.8705 |

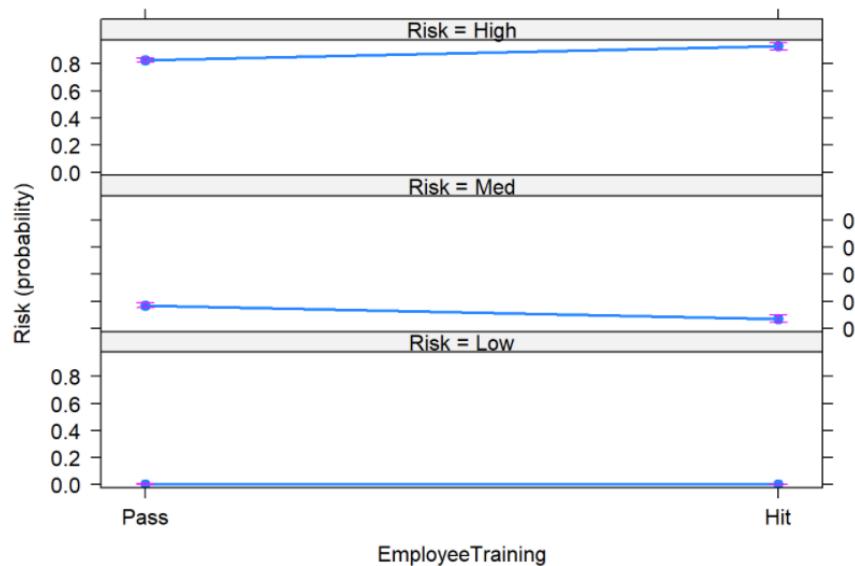
IV. Model Interpretation & Results

Another way to visualize the change in risk assignment is to inspect a plot of the probability changes across each explanatory variable. The plots below illustrate how the probability of “high” risk assignment increases (holding all other variable constant), if a restaurant receives a hit on an inspection area.

ProperUtensils effect plot



EmployeeTraining effect plot



Odds Ratio and Confidence Intervals

| | OR | 2.5 % | 97.5 % |
|--------------------------|--------|--------|--------|
| ## Inspection.DateWinter | 1.4693 | 1.0806 | 1.9967 |
| ## Inspection.DateSpring | 0.6642 | 0.4911 | 0.8954 |
| ## Inspection.DateSummer | 0.5429 | 0.4002 | 0.7334 |
| ## HandContamination.L | 1.3148 | 1.0276 | 1.7116 |
| ## Contamination.L | 1.2329 | 0.8821 | 1.7843 |
| ## FoodTemperature.L | 0.6538 | 0.4994 | 0.8667 |
| ## FoodIdentification.L | 1.3328 | 1.0377 | 1.7425 |
| ## FoodContamination.L | 1.1080 | 0.9214 | 1.3411 |
| ## ProperUtensils.L | 1.6927 | 1.2400 | 2.3898 |
| ## Equipment.L | 1.1126 | 0.9583 | 1.2927 |
| ## PhysicalFacilities.L | 1.0643 | 0.8807 | 1.2801 |
| ## EmployeeTraining.L | 2.1094 | 1.5957 | 2.8705 |

IV. Comparing Methods

- Fit the models using probit or complementary log-log (cloglog) methods
- The probit and cloglog model coefficients are not in terms of log odds and therefore, more difficult to interpret relative to the default logistic method that our model uses

Anova of Default and Probit Methods

```
#update and refit our model with the probit method  
m_mini_probit <- update(m_mini, method = "probit", Hess = TRUE)  
  
#compare models  
anova(m_mini, m_mini_probit)  
  
## Likelihood ratio tests of ordinal regression models  
##  
## Response: Risk  
##  
Model  
## 1 Inspection.Date + HandContamination + Contamination + Food  
Temperature + FoodIdentification + FoodContamination + ProperUt  
ensils + Equipment + PhysicalFacilities + EmployeeTraining  
## 2 Inspection.Date + HandContamination + Contamination + Food  
Temperature + FoodIdentification + FoodContamination + ProperUt  
ensils + Equipment + PhysicalFacilities + EmployeeTraining  
## Resid. df Resid. Dev Test Df LR stat. Pr(Chi)  
## 1 2600 2312.469  
## 2 2600 2309.677 1 vs 2 0 2.79238 0
```

Anova of Default and Cloglog Methods

```
#update and refit model using complementary log log  
m_mini_clog <- update(m_mini, method = "cloglog", Hess = TRUE)  
  
#compare models  
anova(m_mini, m_mini_clog)  
  
## Likelihood ratio tests of ordinal regression models  
##  
## Response: Risk  
##  
Model  
## 1 Inspection.Date + HandContamination + Contamination + Food  
Temperature + FoodIdentification + FoodContamination + ProperUt  
ensils + Equipment + PhysicalFacilities + EmployeeTraining  
## 2 Inspection.Date + HandContamination + Contamination + Food  
Temperature + FoodIdentification + FoodContamination + ProperUt  
ensils + Equipment + PhysicalFacilities + EmployeeTraining  
## Resid. df Resid. Dev Test Df LR stat. Pr(Chi)  
## 1 2600 2312.469  
## 2 2600 2313.683 1 vs 2 0 -1.213399 1
```

Given the residual deviance and AIC values for the probit, cloglog, and logistic (the default OLR model which we used to create our m_mini model) are similar, we recommend using our original m_mini model for this analysis.

Lessons Learned

- OLS requires a large sample size to avoid pitfalls of imbalanced datasets; this problem is compounded during train-test splits.
- If the data is imbalanced, it is more common for OLS variables to fail the Proportional Odds Assumptions.
- Accuracy is not always be the best performance metric to evaluate a model.
- Dropping certain datapoints due to scoping limitations is acceptable and does not adversely affect the model and its outcome.
- The logit model is easier to interpret due to the odds ratio compared to probit and complimentary log-log models if the performance is similar.

Questions

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END