

$W_n$  es un estimador de  $\theta$  ( $X \sim f(x; \theta)$ )

$$\tau(W_n; \theta) = \text{distancia entre } W_n \text{ y } \theta$$

$$\stackrel{\text{ej}}{=} \underbrace{(W_n - \theta)^2}$$

$$R_\tau(W_n; \theta) = E\{\tau(W_n; \theta)\}$$

$$\stackrel{\text{ej}}{=} E\{(W_n - \theta)^2\}$$

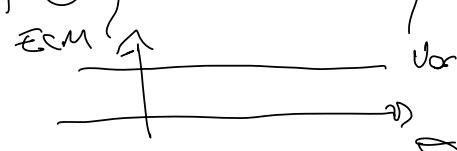
$$= \underbrace{[E\{W_n\} - \theta]^2}_{\text{Sesgo}^2(W_n)} + \underbrace{E\{(W_n - E(W_n))^2\}}_{\text{Var}(W_n)}$$

$W_n$  insesgado p/  $\theta$  si

$$E\{W_n\} = \theta \quad \forall n \geq 1; \theta \in \Theta$$

$$\text{Sesgo}(W_n) = 0 \quad \forall n \geq 1; \theta \in \Theta.$$

$$\underline{\text{Ej 1}}: \quad \underline{\theta = E(X)} \quad \bar{X}_n \Rightarrow E\{\bar{X}_n\} = E(X) = \theta$$

$$ECM(\bar{X}_n; \theta) = \text{Var}(\bar{X}_n) = \frac{\text{Var}(X)}{n}$$


Ex 2:  $X \sim N(\mu; \sigma^2)$   $\sigma^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$

$$= \frac{1}{n-1} \sum_{i=1}^n X_i^2 - \frac{n}{n-1} (\bar{X}_n)^2$$

$$ECM(\hat{\sigma}_n^2, \sigma^2) = \text{Bias}(\hat{\sigma}_n^2) + \text{Var}(\hat{\sigma}_n^2)$$

$$E\{\hat{\sigma}_n^2\} = \frac{1}{n-1} \sum_{i=1}^n E(X_i^2) - \frac{n}{n-1} E(\bar{X}_n^2)$$

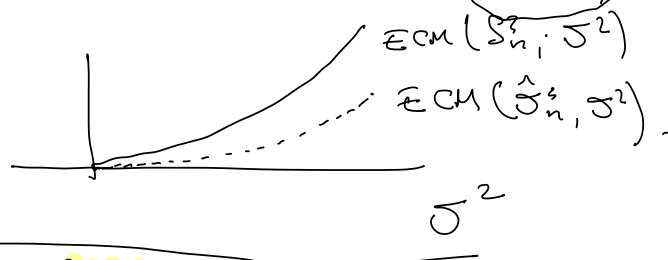
$$\begin{aligned} \text{Var}(Y) &= E(Y^2) - [E(Y)]^2 \\ E(Y^2) &= \text{Var}(Y) + [E(Y)]^2 \end{aligned}$$

$$= \frac{n}{n-1} (\sigma^2 + \mu^2) - \frac{n}{n-1} [\text{Var}(\bar{X}_n) + (E(\bar{X}_n))^2]$$

$$= \frac{n}{n-1} (\sigma^2 + \mu^2) - \frac{n}{n-1} \left[ \frac{\sigma^2}{n} + \mu^2 \right]$$

$$E\{\hat{\sigma}_n^2\} = \sigma^2 \quad (\text{unbiased})$$

$$ECM(\hat{\sigma}_n^2; \sigma^2) = \text{Var}(\hat{\sigma}_n^2) = \frac{2\sigma^4}{n-1}$$



$$\hat{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 = \frac{n-1}{n} S_n^2$$

$$ECM(\hat{\sigma}_n^2; \sigma^2) = [E(\hat{\sigma}_n^2) - \sigma^2]^2 + \text{Var}(\hat{\sigma}_n^2)$$

$$E\{\hat{\sigma}_n^2\} = \frac{n-1}{n} E(S_n^2) = \frac{n-1}{n} \sigma^2 \neq \sigma^2$$

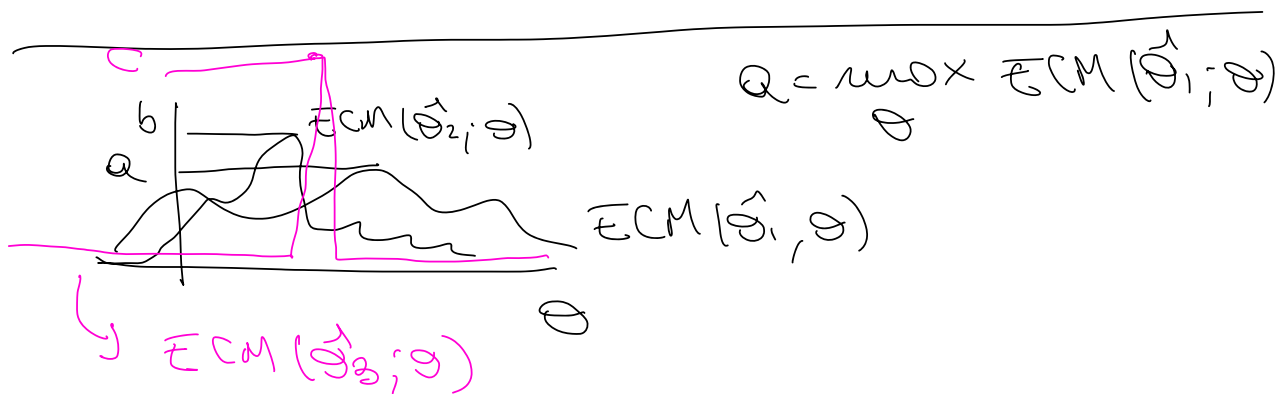
$$\text{Bias}(\hat{\sigma}_n^2) = \frac{n-1}{n} \sigma^2 - \sigma^2 = -\frac{1}{n} \sigma^2$$

$$\text{Var}(\hat{\sigma}_n^2) = \text{Var}\left(\frac{n-1}{n} S_n^2\right) = \frac{(n-1)^2}{n^2} \frac{2\sigma^4}{(n-1)}$$

$$\text{ECM}(\hat{\sigma}_n^2; \sigma^2) = \frac{(2n-1)}{n^2} \sigma^2$$

$$\frac{\text{ECM}(\hat{\sigma}_n^2; \sigma^2)}{\text{ECM}(S_n^2; \sigma^2)} = 1 - \underbrace{\frac{2n-1}{2n^2}}_{\geq 0 \text{ if } n \geq 1} < 1$$

$$\text{ECM}(\hat{\sigma}_n^2; \sigma^2) < \text{ECM}(S_n^2; \sigma^2) \text{ if } n \geq 1$$



$$X \sim \text{Pois}(\lambda) \Rightarrow E(X) = V(X) = \lambda$$

Have 2 est. integrals of  $\lambda$

$$\textcircled{1} \bar{X}_n \Rightarrow E(\bar{X}_n^2) = E(X) = \lambda$$

$$\textcircled{2} S_n^2 \Rightarrow E(S_n^2) = \text{Var}(X) = \lambda$$

$$\text{Var}(\bar{Y}_n) = \frac{\lambda}{n} \quad \text{Var}(S_n^2) = ?$$

$$\underline{W_{\alpha,n}} = \alpha \bar{X}_n + (1-\alpha) S_n^2 \quad \alpha \in (0,1)$$

$$\underline{W_n} \in C_0 \Rightarrow \text{Var}(W_n) \geq \underbrace{C_0 \alpha}_{\downarrow \text{CR}}$$

$$I_n(\lambda) = n \hat{I}(\lambda) = n \mathbb{E} \left\{ \left[ \frac{\partial}{\partial \lambda} \ln f(x; \lambda) \right]^2 \right\}$$

$$\ln f(x; \lambda) = \ln \frac{e^{-\lambda} \lambda^x}{x!} = -\lambda + x \ln(\lambda) - \ln(x!) \quad \hat{I}(\lambda)$$

$$\frac{\partial}{\partial \lambda} \ln f(x; \lambda) = -1 + \frac{x}{\lambda} \Rightarrow \frac{\partial^2}{\partial \lambda^2} \ln f(x; \lambda) = -\frac{x}{\lambda^2}$$

$$\begin{aligned} \hat{I}(\lambda) &= \mathbb{E} \left\{ \left[ -1 + \frac{x}{\lambda} \right]^2 \right\} \\ &= \frac{1}{\lambda^2} \underbrace{\mathbb{E} \{ (x - \lambda)^2 \}}_{\text{Var}(X)} = \frac{\lambda}{\lambda^2} \Rightarrow \hat{I}(\lambda) = \frac{1}{\lambda} \end{aligned}$$

$$\underline{I_n(\lambda) = n \hat{I}(\lambda)}$$

Si  $W_n \in C_0$  on est. insup. de  $\lambda$   
 $\Rightarrow \text{Var}(W_n) \geq I_n^{-1}(\lambda) = \lambda/n$

$$\text{Var}(\bar{X}_n) = \frac{\text{Var}(X)}{n} = \frac{\lambda}{n} \rightarrow \text{Alcance la cota}$$

⇓

UMVUE.

$$i(\lambda) = -E\left\{\frac{\partial^2}{\partial \lambda^2} \ln f(X; \lambda)\right\} = -E\left\{-\frac{X}{\lambda^2}\right\} = \frac{E(X)}{\lambda^2} = \underline{\frac{1}{\lambda}}$$

$$I_n(\lambda) = n / \lambda.$$

$$X \sim \text{Exp}(\theta) \quad \theta = E(X) \quad \bar{X}_n$$

$$E\{\bar{X}_n\} = E(X) = \theta \quad \gamma \left[ \text{Var}(\bar{X}_n) = \frac{\text{Var}(X)}{n} = \frac{\theta^2}{n} \right]$$

$$I_n(\theta) = -n E\left\{\frac{\partial^2}{\partial \theta^2} \ln f(X; \theta)\right\}$$

$$\ln f(X; \theta) = \ln \frac{1}{\theta} e^{-X/\theta} = -\ln(\theta) - \frac{X}{\theta}$$

$$\frac{\partial^2}{\partial \theta^2} \ln f(X; \theta) = +\frac{1}{\theta^2} - \frac{2X}{\theta^3}$$

$$\underline{I_n(\theta)} = -n E\left\{\frac{1}{\theta^2} - \frac{2X}{\theta^3}\right\} = \frac{-n}{\theta^2} + \frac{2n}{\theta^2} = \frac{n}{\theta^2}$$

$$\text{Si } W_n \in \underline{C_\theta} \Rightarrow \left[ \text{Var}(W_n) \geq \frac{\theta^2}{n} \right]$$

$\Rightarrow \bar{X}_n$  es el UMVU de  $\theta$ .

$X \sim \text{Bern}(\theta)$   $\hat{\theta}_n = \bar{X}_n \Rightarrow E\{\bar{X}_n\} = E\{X\} = \theta$   
 $C_\theta$   $\boxed{\text{Var}(\bar{X}_n) = \frac{\theta(1-\theta)}{n}}$

$$l(\theta) = \ln(\theta) \sum_{i=1}^n x_i + \ln(1-\theta) \left(n - \sum_{i=1}^n x_i\right)$$

$$l'_\theta: \frac{\sum_{i=1}^n x_i}{\theta} = \frac{(n - \sum_{i=1}^n x_i)}{1-\theta} = 0 \Rightarrow \hat{\theta}_n = \bar{X}_n$$

$$l''_\theta: -\frac{\sum_{i=1}^n x_i}{\theta^2} - \frac{(n - \sum_{i=1}^n x_i)}{(1-\theta)^2} < 0 \quad \forall \theta \in (0,1)$$

$$\begin{aligned}
 I_n(\theta) &= -E\{l''_\theta\} \\
 &= -E\left\{ \frac{\sum_{i=1}^n x_i}{\theta^2} - \frac{(n - \sum_{i=1}^n x_i)}{(1-\theta)^2} \right\} \\
 &= \frac{\sum_{i=1}^n E(x_i)}{\theta^2} + \frac{n - \sum_{i=1}^n E(x_i)}{(1-\theta)^2}
 \end{aligned}$$

$$= \frac{n\cancel{\theta}}{\theta^2} + \frac{n(1-\cancel{\theta})}{(1-\theta)^2}$$

$$I_n(\theta) = \frac{n}{\theta} + \frac{n}{(1-\theta)} = \frac{n}{\theta(1-\theta)}$$

$$\text{Si } W_n \in C_\theta \Rightarrow \text{Var}(W_n) \geq \frac{\theta(1-\theta)}{n}$$

$$\bar{X}_n \text{ es insesgado y } \text{var}(\bar{X}_n) = \frac{\theta(1-\theta)}{n}$$



$\bar{X}_n$  es el MME

Consistencia y  $\text{ECM}(W_n; \theta) = \text{Bias}^2(W_n) + \text{Var}(W_n)$

$$W_n \xrightarrow{p} \theta \Rightarrow \lim_{n \rightarrow \infty} P(|W_n - \theta| \geq \varepsilon) = 0$$

Cheby: Si  $g(x)$  es una función no-neg.

$$P(g(x) \geq \varepsilon) \leq \frac{\text{ECM}(W_n; \theta)}{\varepsilon} \quad \text{y } \varepsilon > 0$$

$$P(|W_n - \theta| \geq \varepsilon) = P(\underbrace{|W_n - \theta|^2}_{g(W_n)} \geq \varepsilon^2) \leq \frac{\text{ECM}(W_n; \theta)}{\varepsilon^2}$$

$\uparrow$   
 $E[(W_n - \theta)^2]$

$$\lim_{n \rightarrow \infty} P(|W_n - \theta| \geq \varepsilon) \leq \frac{E S_{\varepsilon^2}(W_n) + \text{Var}(W_n)}{\varepsilon^2}$$

$$ECM(W_n; \theta) \rightarrow 0 \quad n \rightarrow \infty$$

$$\underline{Ej}: \quad \underline{ECM}(S_n^2; \sigma^2) = \frac{2\sigma^4}{(n-1)} \rightarrow 0 \quad n \rightarrow \infty$$

$X \sim N(\mu, \sigma^2)$

$$S_n^2 \rightarrow \text{consistent } \sigma^2$$

$$S_n^2 \xrightarrow{P} \sigma^2$$

$$ECM(\hat{\sigma}_n^2; \sigma^2) = \frac{(2n-1)\sigma^4}{n^2} \rightarrow 0 \quad n \rightarrow \infty$$

$\downarrow$

$$E_j \text{ consistent } \hat{\sigma}_n^2 \text{ and } \hat{\sigma}_n^2 \xrightarrow{P} \sigma^2$$