# NOCIONES BASICAS DE PROBABILIDAD

- Leyes de Morgan I:  $\left(\bigcup_{i=1}^n A_i\right)^c = \bigcap_{i=1}^n A_i^c$  II:  $\left(\bigcap_{i=1}^n A_i\right)^c = \bigcup_{i=1}^n A_i^c$
- Principio Inclusión Exclusión  $P(\bigcup_{i=1}^n A_i) = \sum_{i=1}^n P(A_i) \sum_{1 \le i < j \le n} P(A_i \cap A_j) + \dots + (-1)^{n-1} P(A_1 \cap \dots \cap A_n)$
- Definición axiomática probabilidad I:  $P(\Omega) = 1$  II:  $P(A) \ge 0$  III:  $A \subset B \Rightarrow P(A) \le P(B)$  IV:  $P(\bigcup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} P(A_i)$
- Permutaciones y combinaciones I:  $\binom{n}{r} = \binom{n}{n-r}$  II:  $k\binom{k-1}{r-1} = r\binom{k}{r}$
- Distribución multinomial  $(x_1 + \cdots x_r)^n = \sum \binom{n}{n_1 \cdots n_r} x_1^{n_1} \cdots x_r^{n_r}$ , donde la suma es sobre los  $n_i$  de manera que  $\sum_{i=1}^r n_i = n$
- Condicional  $\bigcup_{i=1}^n B_i = \Omega$ ,  $B_i \cap B_j = \emptyset \ \forall_{i \neq j}$  LPT:  $P(A) = \sum_{i=1}^n P(A/B_i)P(B_i)$  Bayes:  $P(B_i/A) = \frac{P(A/B_i)P(B_i)}{\sum_{i=1}^n P(A/B_i)P(B_i)}$  dato:  $P(\bigcup_{i=1}^\infty A_i|B) = \sum_{i=1}^\infty P(A_i|B)$
- Independencia  $P(A/B) = P(A) \longrightarrow P(A \cap B) = P(A)P(B)$

### DISTRIBUCIONES DISCRETAS

- Binomial  $X \sim B(n,p)$  **PMF:**  $P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$  **FGM:**  $M(t) = (pe^t + 1 p)^n$  **E(X)** = np **V(X)** = np(1-p)
- Geométrica  $X \sim G(p)$  **PMF:**  $P(X = k) = (1 p)^{k-1} p, k \ge 1$  **FGM:**  $M(t) = \frac{pe^t}{1 (1 p)e^t}$  **E(X)**  $= \frac{1}{p}$  **V(X)**  $= \frac{1}{p^2} \frac{1}{p}$
- BinomNeg  $X \sim BN(r,p)$  **PMF:**  $P(X = k) = \binom{k-1}{r-1} p^r (1-p)^{k-r}, k \ge r$  **FGM:**  $M(t) = \left(\frac{pe^t}{1-(1-p)e^t}\right)^r$  **E(X):**  $\frac{r}{p}$  **V(X):**  $\frac{r(1-p)}{p^2}$
- Poisson  $X \sim P(\lambda)$  PMF:  $P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}, k \ge 0$  FGM:  $M(t) = e^{\lambda(e^t 1)} \mathbf{E}(\mathbf{X}) = \lambda$  (recordar reescalar  $\lambda$ )
- Hipergeométrica  $X \sim H(n,r,m)$  PMF:  $P(X=k) = \frac{\binom{r}{k}\binom{n-r}{m-k}}{\binom{n}{m}}$  FGM:  $E(X) = \frac{mr}{n}$   $Var(X) = \frac{mr}{n^2(n-1)}(m-n)(r-n)$ ,

### DISTRIBUCIONES CONTINUAS

- Definición I:  $f: \mathbb{R} \to \mathbb{R}$   $P(X \in B) = \int_B f(x) dx$  II:  $P(X \in (-\infty, \infty)) = \int_{-\infty}^{+\infty} f(x) dx = 1$  III:  $P(a \le X \le b) = \int_a^b f(x) dx$
- Distribución Uniforme  $X \sim U(a,b)$  PDF:  $a \leq x \leq b \rightarrow f(x) = 1$  FGM:  $M(t) = \frac{e^{tb} e^{ta}}{t(b-a)}$
- Exponencial  $X \sim Exp(\lambda)$  **PDF:**  $x \geq 0 \rightarrow f(x) = \lambda e^{-\lambda x}$  **FGM:**  $M(t) = \frac{\lambda}{\lambda t}$ , **prop olvido**: P(T > t + s | T > s) = P(T > t)
- Gamma  $X \sim G(\alpha, \lambda)$  **PDF:**  $x \ge 0 \to f(x) = \frac{\lambda^{\alpha}}{\Gamma(\alpha)} x^{\alpha 1} e^{-\lambda x}$   $\Gamma(\alpha) = \int_{0}^{\infty} u^{\alpha 1} e^{-u} du$  **FGM:**  $M(t) = \left(\frac{\lambda}{\lambda t}\right)^{\alpha}$
- $\blacksquare \ \ \textit{Notas de Color sobre la Gamma} \ \Gamma(n) = (n-1)! \ \text{si } n \ \text{entero} \ \Gamma(\frac{1}{2}) = \sqrt{\pi}, \ \Gamma(\alpha) = (\alpha-1)\Gamma(\alpha-1), \ \Gamma(\frac{n}{2}) = \frac{(n-2)!!\sqrt{\pi}}{2^{\frac{n-1}{2}}}$
- Normal  $X \sim N(\mu, \sigma^2)$  **PDF:**  $-\infty \le x \le \infty \to f(x) = \frac{1}{\sigma\sqrt{2\pi}}exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$  **FGM:**  $M(t) = exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right)$
- Functiones de variables aleatorias  $I Y = g(X) F_Y(y) = P(Y \le y) = P(g(X) \le y) = F_x(g^{-1}(y)) \rightarrow f_y(y) = f_x(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|$
- Funciones de v.a. IIX con dist F, si  $Z = F(X) \Rightarrow Z \sim U[0,1]$   $P(Z \le z) = P(F(X) \le z) = P(X \le F^{-1}(z)) = F(F^{-1}(z)) = Z$
- Funciones de v.a. III  $U \sim U[0,1] \, X = F^{-1}(U) \Rightarrow X \sim F(U)$ .  $P(X \le x) = P(F^{-1}(U) \le x) = p(U \le F(x)) = F(x)$
- $Exp(\lambda) \equiv G(1,\lambda), \chi^2_g \equiv G(\frac{g}{2},\frac{1}{2}) \ X_i \sim P(\lambda_i) \ \text{indep} \Rightarrow X_1 + X_2 \sim P(\lambda_1 + \lambda_2), \ X_i \sim Exp(\lambda) \ \text{indep} \Rightarrow \sum X_i \sim G(n,\lambda), \ Z \sim P(\lambda), \ X|_{Z=z} \sim Bi(z,p) \Rightarrow X \sim P(p\lambda), \ \text{si} \ n \to \infty, p \to 0, np \to \lambda \Rightarrow B(n,p) \approx P(\lambda)$
- $\blacksquare X \sim U[a,b], E(X) = \tfrac{a+b}{2}, V(X) = \tfrac{(b-a)^2}{12}, \ X \sim Exp(\lambda), E(X) = \tfrac{1}{\lambda}, V(X) = \tfrac{1}{\lambda^2}, X \sim \Gamma(\alpha,\lambda)E(X) = \tfrac{\alpha}{\lambda}, V(X) = \tfrac{\alpha}{\lambda^2}$

# DISTRIBUCION CONJUNTA

- $P(x_1 \le X \le x_2, y_1 \le Y \le y_2) = F(x_2, y_2) F(x_1, y_2) F(x_2, y_1) + F(x_1, y_1)$
- Variables continuas  $F(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(u,v) du dv \ f(x,y) = \frac{\partial^{2}}{\partial x \partial y} F(x,y)$
- $\bullet \ (X,Y) \sim f(x,y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}(\frac{X-\mu_x}{\sigma_x})^2 + (\frac{y-\mu_y}{\sigma_y})^2 \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}} \\ \Rightarrow Y|_{X=x} \sim N(\mu_y \rho\frac{\sigma_x}{\sigma_y}(x-\mu_x), \sigma_y^2(1-\rho^2))$
- Incrementos  $P(x \le X \le X + \delta_x, y \le Y \le y + \delta_y) \approx f(x, y)\delta_x\delta_y$
- Distribuciones marginales  $f_x(x) = F_X'(x) = \int_{-\infty}^{\infty} f(x,y) dy$  (Si quiero  $P(X \ge Y)$  hago  $\int_{-\infty}^{x} f(x,y) dy$ )
- Independencia  $\vec{X} \to F(\vec{X}) = \prod F_{x_i}$  Estadístico k-ésimo de orden  $X_{(k)}$  con  $f_k(x) = \frac{n!}{(k-1)!(n-k)!} f(x) F^{k-1}(x) (1 F(x))^{n-k}$

■ Estadísticos de orden y extremos  $\vec{X}$  iid F,f.  $U = max\{X_i\}$   $V = min\{X_i\}$ .  $F_U(u) = [F(u)]^n$   $F_V(v) = 1 - [1 - F(v)]^n$ 

#### PROBABILIDAD CONDICIONAL

- Caso Discreto:  $P(X = x_i/Y = y_j) = \frac{P(X = x_i, Y = y_j)}{P(Y = y_i)}$
- Independencia Discreto: X e Y independientes,  $P(X = x_i/Y = y_j) = p_x(x:i)$
- Caso Continuo:  $f_{x/y}(x/y) = \frac{f_{xy}(x,y)}{f_y(y_j)} f_{xy}(x,y) = f_{x/y}(x/y) f_y(y_j) = f_{y/x}(y/x) f_x$
- Independencia Continuo: X e Y independientes,  $P(X = x_i/Y = y_i) = f_x(x_i)$

# FUNCIONES DE VARIABLES ALEATORIAS CONTINUAS

- Suma  $X \in Y$  Z = X + Y  $F_z(z) = \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f_{xy}(x,y) dy dx = \int_{-\infty}^{\infty} \int_{-\infty}^{z} f_{xy}(x,v-x) dv dx$
- Convolución:  $f_z(z) = \int_{-\infty}^{\infty} f_{xy}(x, z x) dx$ . Si X e Y independientes  $f_z(z) = \int_{-\infty}^{\infty} f_x(x) f_y(z x) dx$ .
- Cocientes: Z = Y/X.  $x > 0 \rightarrow y \le xz$   $x < 0 \rightarrow y \ge xz$   $F_z(z) = \int_{-\infty}^0 \int_{xz}^\infty f_{xy}(x,y) dy dx + \int_0^\infty \int_{-\infty}^{xz} f_{xy}(x,y) dy dx = \int_{-\infty}^\infty z \int_{-\infty}^\infty |x| f_{xy}(x,xv) dx dv$ , Z con dist. de Cauchy  $f_Z(z) = \frac{1}{\pi(z^2+1)}$
- densidad del cociente  $f_z(z) = \int_{-\infty}^{\infty} |x| f_{xy}(x,xz) dx$  Si X e Y son indep  $\Rightarrow f_z(z) = \int_{-\infty}^{\infty} |x| f_X(x) f_Y(xz) dx$
- Función Invertible: X e Y con  $f_{xy}(x,y)$   $u = g_1(x,y)$   $v = g_2(x,y)$  invertibles  $x = h_1(u,v)$   $y = h_2(u,v)$ .  $f_{uv}(u,v) = f_{xy}(h_1(u,v),h_2(u,v)) \cdot |J(h_1(u,v),h_2(u,v))|$

# ESPERANZA MATEMATICA

- $\bullet \ E(g(X)) = \sum_{i:p(x_i)>0} g(x_i)p(x_i), \text{ en part: } M_X(t) = E(e^{tX}) = \sum_{x:p(x)>0} e^{tx}p(x), \ E(X^n) = \sum_{x:p(x)>0} x^np(x), \ E(X) = \int_{-\infty}^{\infty} xf(x)dx$
- FGM: I: Si  $0 \in X \Rightarrow \exists ! F(x)$  II:  $M^{(r)}(0) = E(X^r)$  III:  $Y = aX + b, M_Y(t) = e^{bt}M_x(at)$  IV: X, Y ind Z = X + Y  $M_Z(t) = E(e^{t(X+Y)}) = E(e^{tX}e^{tY}) = E(e^{tX})E(e^{tY}) = M_x(t)M_y(t)$ .
- Functiones de Variables:  $Y = g(X_1, \dots, X_n)$   $E(Y) = \int \dots \int_{x_1, \dots, x_n} g(x_1, \dots, x_n) p(x_1, \dots, x_n) dx_1 \dots dx_n$
- Independencia: X, Y independences. E(g(X)h(Y)) = E(g(X))E(h(Y))
- Suma:  $Y = a + \sum_{i=1}^{n} b_i X_i \ E(Y) = a + \sum_{i=1}^{n} b_i E(X_i)$
- Designaldad Jensen: X y  $\phi$  convexa.  $\phi(E(X)) \leq E(\phi(X))$ , Designaldad Markov: X no negativa.  $P(X \geq a) \leq \frac{E(X)}{a}$
- Designal and Chebyshev:  $X \sim (\mu, \sigma^2) P(|X \mu| > t) \leq \frac{\sigma^2}{t^2} P(|X \mu| > k\sigma) \leq \frac{1}{k^2}$
- Covariancia: Cov(X,Y) = E(XY) E(X)E(Y)  $Cov(aX,bY) = a \cdot b \cdot Cov(X,Y)$  Cov(X+Y,Z) = Cov(X,Z) + Cov(Y,Z)  $Var\left(a + \sum_{i=1}^{n} b_i X_i\right) = \sum_{j=1}^{n} \sum_{i=1}^{n} b_i b_j Cov(X_i, X_j)$ , Coeficiente de Correlación:  $\rho = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}} 1 \le \rho \le 1$
- Esperanza Condicional:  $E(Y/X=x)=\int_{-\infty}^{\infty}yf_{y/x}(y/x)dy,\ E(h(Y)/X=x)=\int_{-\infty}^{\infty}h(y)f_{y/x}dy$
- $L.E.I.: \mathbf{I}: E(Y) = E(E(Y/X)) \mathbf{II}: Var(Y) = Var(E(Y/X)) + E(Var(Y/X)), V(Y|X) = E[(Y E(Y|X))^2|X]$

# NOCIONES DE CONVERGENCIA

- Convergencia en probabilidad:  $\lim_{x\to\infty} Prob(|x_n-c|>\epsilon)=0 \ \forall \epsilon>0 \longrightarrow plim\ x_n=c$
- Convergencia en media cuadrática:  $x_n \sim (\mu_n, \sigma_n^2) \lim_{n \to \infty} \mu_n = c \lim_{n \to \infty} \sigma_n^2 = 0 \Rightarrow plim x_n = c$
- Ley débil GN:  $X_1, \ldots, X_n$ i.i.d.  $E(X_i) = \mu \ Var(X_i) = \sigma^2 \ \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \longrightarrow plim \ \bar{X} = \mu$
- Convergencia A.S.:  $\lim_{n\to\infty} Prob(|x_i-c|>\epsilon, \exists i\geq n)=0 \longrightarrow x \stackrel{a.s.}{\to} c$
- Ley fuerte GN:  $X_1, \ldots, X_n$  iid  $E(X_i) = \mu_i < \infty \ Var(X_i) = \sigma_i^2 < \infty \ \sum_{i=1}^{\infty} \sigma_i^2/i^2 < \infty \longrightarrow \bar{x}_n \mu_n \stackrel{a.s.}{\to} 0$
- Teorema Slutsky:  $plim g(x_n) = g(plim x_n) g(x_n)$  es continua y no es función de n
- Reglas plim: plim  $(x_n, y_n) = (c, d)$  I:plim  $(x_n + y_n) = c + d$  II:plim  $(x_n \cdot y_n) = c \cdot d$  III:plim  $(x_n/y_n) = c/d$
- Convergencia en distribución:  $\lim_{n\to\infty} F(x_n) = F(x) \ (M_n(t) \to M(t)) \longrightarrow (F(x_n) \to F(x))$  si contiene a 0
- Reglas distribuciones límite I:  $x_n \stackrel{d}{\to} x$  plim  $y_n = c$  I:  $x_n \cdot y_n \stackrel{d}{\to} c \cdot x$  II:  $x_n + y_n \stackrel{d}{\to} x + c$  III:  $x_n/y_n \stackrel{d}{\to} x/c$
- Reglas distribuciones límite II:  $x_n \stackrel{d}{\to} x$   $g(x_n)$  continua  $\Rightarrow g(x_n) \stackrel{d}{\to} g(x)$
- TCL Lindberg-Levy:  $X_1, \ldots, X_n \ \mu < \infty, \sigma^2 < \infty \ \bar{x}_n = \frac{1}{n} \sum_{i=1}^n X_i \Rightarrow \sqrt{n}(\bar{x}_n \mu) \stackrel{d}{\to} N(0, \sigma^2)$
- Método Delta Si  $\sqrt{n}(Z_n \mu) \stackrel{d}{\to} N(0, \sigma^2)$  y  $g(Z_n)$  derivable y que no depende de  $n \Rightarrow \sqrt{n}(g(Z_n) g(\mu)) \stackrel{d}{\to} N(0, (g'(\mu))^2 \sigma^2)$