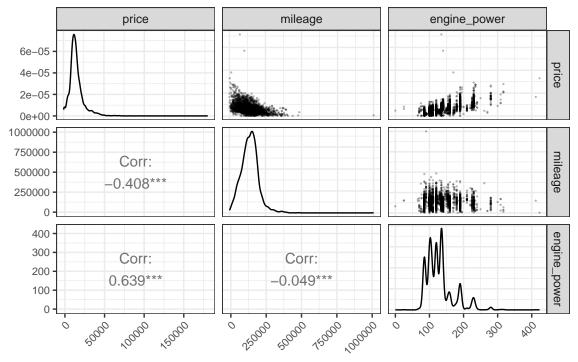
```
mylibrary <- c("dplyr", "ggplot2", "cowplot", "GGally", "MASS")
invisible(lapply(mylibrary, library, character.only = TRUE))
### load the data
dat_bmw <- read.csv("BMWpricing_updated.csv")
### Assigning NA to data points with negative mileage
dat_bmw[dat_bmw$mileage < 0, "mileage"] <- NA
### Converting mileage, engine power, and price into numeric
dat_bmw[,c(3,4,17)] <- apply(dat_bmw[,c(3,4,17)], 2, as.numeric)
### Create a data frame only has the variable of interests
dat_bmw2 <- dat_bmw[,c("price", "mileage", "engine_power")]
### Drop the data point with NA
dat bmw2 <- dat bmw2[complete.cases(dat bmw2),]</pre>
```

scatterplot matrix

Distribution and Correlation between Price, Mileage, and Engine Power



scatter plot of price and mileage We first set the negative mileage (in one sample) as the missing data and excluded it in our following analysis.

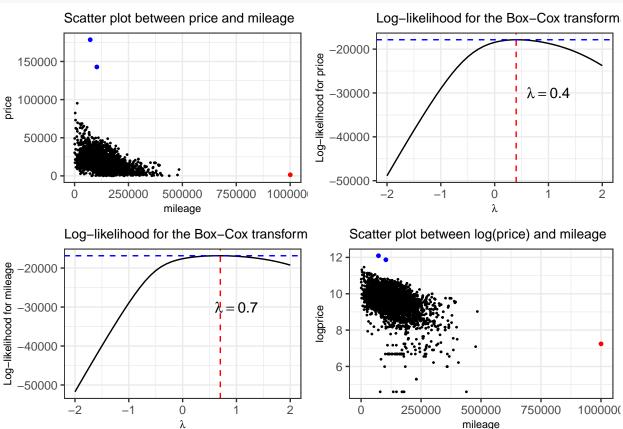
Our scatter plot shows a negative correlation between price and mileage, indicating that price tends to decrease as mileage increases. However, this relationship is not strongly linear and a straight-line regression model may not be the best option. We find some problems in our data: 1) Both price and mileage follow highly right-skewed distributions; 2) High-leverage points appear on the right side of the scatter plot, particularly the red point with approximately 1,000,000 miles; 3) Outliers with extreme high sold prices are observed on the top-left of the plot, especially these blue points with price over 100000.

To improve normality, we applied the Box-Cox transformation to the price variable. Though the λ with the greatest log-likelihood ($\lambda_{price} = 0.4, \lambda_{mileage} = 0.7$) and its 95% CI do not include 0, 0.5, and 1, there are slight differences in log-likelihood for price when $\lambda \in [0,0.5]$ and for mileage when $\lambda \in [0.5,1]$. For

interpretability, we finally apply log-transformation on the price and keep the mileage in the data. In the scatter plot of log(price) and mileage, we observe some potential outliers with low prices and short mileage.

```
# Scatter plot of price vs. mileage with highlighted outliers
scatterPlotPriceMileage <- ggplot(data = dat bmw2, mapping = aes(x = mileage, y = price)) +</pre>
  geom_point(pch=19, cex=0.3) + # Default points
  geom_point(data = subset(dat_bmw2, mileage > 500000), mapping = aes(x = mileage, y = price), pch=19,
  geom_point(data = subset(dat_bmw2, price > 100000), mapping = aes(x = mileage, y = price), pch=19, ce.
 labs(title = "Scatter plot between price and mileage") +
  theme_bw() + theme(title = element_text(size=8))
# Box-Cox transformation for the price variable
boxcox_price <- boxcox(lm(dat_bmw2$price ~ 1), plotit = F)</pre>
boxcox_priceDF <- data.frame("lambda" = boxcox_price$x, "ll" = boxcox_price$y)</pre>
maxlabmda_price <- boxcox_price$x[which.max(boxcox_price$y)]</pre>
# Box-cox Plot for the price variable
boxcoxPlotPrice <- ggplot(data = boxcox_priceDF, mapping = aes(x = lambda, y = 11)) +</pre>
  geom_line(color = "black", linewidth = 0.5) +
  geom_vline(xintercept = maxlabmda_price, linetype = "dashed", color = "red") +
  geom_hline(yintercept = max(boxcox_price$y) - qchisq(0.95, df = 1)/2, linetype = "dashed", color = "b
  labs(title = "Log-likelihood for the Box-Cox transformation",
       x = expression(lambda),
       y = "Log-likelihood for price") +
  annotate("text", label = paste0("lambda ==", maxlabmda_price), x = 1, y = -30000, parse = T) +
  theme bw() + theme(title = element text(size=8))
# Box-cox transformation for the mile variable
boxcox_mile <- boxcox(lm(dat_bmw2$mileage ~ 1), plotit = F)</pre>
boxcox_mileDF <- data.frame("lambda" = boxcox_mile$x, "ll" = boxcox_mile$y)</pre>
maxlabmda_mile <- boxcox_mile$x[which.max(boxcox_mile$y)]</pre>
# Box-cox Plot for the mile variable
boxcoxPlotMile <- ggplot(data = boxcox_mileDF, mapping = aes(x = lambda, y = 11)) +
  geom_line(color = "black", linewidth = 0.5) +
  geom_vline(xintercept = maxlabmda_mile, linetype = "dashed", color = "red") +
  geom_hline(yintercept = max(boxcox_mile$y) - qchisq(0.95, df = 1)/2, linetype = "dashed", color = "bl
  labs(title = "Log-likelihood for the Box-Cox transformation",
       x = expression(lambda),
       y = "Log-likelihood for mileage") +
  annotate("text", label = paste0("lambda ==", maxlabmda_mile), x = 1, y = -30000, parse = T) +
  theme_bw() + theme(title = element_text(size=8))
# Create a column of logprice in dat_bmw2
dat_bmw2 <- dat_bmw2 %>% mutate(logprice = log(price))
# Scatter plot of logprice vs. mileage with highlighted outliers
scatterPlotLogPriceMileage <- ggplot(data = dat_bmw2, mapping = aes(x = mileage, y = logprice)) +</pre>
  geom_point(pch=19, cex=0.3) +
  geom_point(data = subset(dat_bmw2, mileage > 500000), mapping = aes(x = mileage, y = logprice), pch=1
  geom_point(data = subset(dat_bmw2, logprice > log(100000)), mapping = aes(x = mileage, y = logprice),
  labs(title = "Scatter plot between log(price) and mileage") +
  theme_bw() + theme(title = element_text(size=8))
```

plot_grid(scatterPlotPriceMileage, boxcoxPlotPrice, boxcoxPlotMile, scatterPlotLogPriceMileage, align =



Outliers

```
modelDF <- dat_bmw2
### Assigning NA to car with mileage > 50,0000 miles
modelDF[modelDF$mileage > 500000, "mileage"] <- NA
### Assigning NA to car with price > 100000 usd
modelDF[modelDF$price > 100000, "price"] <- NA
modelDF <- modelDF[complete.cases(modelDF),]
modelDF$logprice <- log(modelDF$price)</pre>
```

model fitting We applied an Ordinary Least Square regression model to predict the natural logrithm of BMW car prices from mileage. The model may be represented as:

$$log(price) = \beta_0 + \beta_1 \times mileage + \epsilon$$

The residuals represent the difference between the model-predicted values and the actual $\log(\text{price})$ values. Ideally, they should be symmetrically distributed around zero for a well-behaved model. But the minimum of residual is -5.1907. However, the minimum of residual is -5.1907. That means that, for at least one observation, the model-predicted $\log(\text{price})$ is much larger than the actual value.

The "Coefficients" section provides us with detailed statistics for the intercept and the mileage coefficient. The intercept is estimated to be 10.17 with very small standard error (0.02175) and thus having an extremely large t-value of 467.7 and an extremely small p-value of less than 2e-16. This is very strong evidence against

the null hypothesis that the intercept is zero. Similarly, mileage coefficient is also estimated as -4.746e-06, indicating that for one unit increase in mileage, logprice decreases by about 4.746e-06 units. Its standard error is 1.425e-07, which yields a t-value of -33.3 and a p-value similarly less than 2e-16. These allow us to reject the null hypothesis for both coefficients, meaning mileage is a statistically significant logprice predictor.

Besides the individual coefficients, the fit of the entire model is also tested by means of the F-test. The F-statistic of 1109 with degrees of freedom 1 and 4837, and p-value smaller than 2.2×10^{-16} tests whether the model that includes mileage as a predictor explains significantly more variation in log(price) than does a model with no predictors. This result confirms that our model as a whole is statistically significant although mileage by itself explains only about 18.65% of the variation in log(price) (as indicated by the Multiple R-squared of 0.1865). The relatively low R-squared value suggests that while mileage is a significant predictor, other variables likely contribute to the variation in car prices.

```
mod1 <- lm(logprice ~ mileage, data = modelDF)
summary(mod1)</pre>
```

```
##
## Call:
## lm(formula = logprice ~ mileage, data = modelDF)
##
##
  Residuals:
##
       Min
                1Q
                                3Q
                   Median
                                       Max
##
   -5.1907 -0.2404
                   0.0496
                            0.3229
                                    1.3558
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                1.017e+01
                           2.175e-02
                                       467.7
                                                <2e-16 ***
               -4.746e-06 1.425e-07
                                        -33.3
                                                <2e-16 ***
## mileage
                 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 0.5838 on 4837 degrees of freedom
## Multiple R-squared: 0.1865, Adjusted R-squared: 0.1864
## F-statistic: 1109 on 1 and 4837 DF, p-value: < 2.2e-16
```

In the plot below, we overlay the scatter plot of mileage vs price with the linear regression line. We can see that as mileage increases, log-price will decrease overall. But there is plenty of scatter around that best-fit line. Particularly, there are a few points at the low end of log-price (close to 5–6 on the y-axis) that are rather far away from the rest of the data—these points are potential high-leverage points for the regression because they are far out in mileage and have exceptionally low log-prices. They can pull the regression line down more than if they were not present. Despite those outliers, the linear fit still shows the general inverse relation (higher mileage \rightarrow lower log-price), but it does not explain all of the price variability, given the broad scatter of points. So we can conclude that simple linear regression is not a good fit here. A multiple regression model would be a better choice as it would account for the variation in price that is not accounted for by mileage alone.

