Estimation & Decisions

STA721 Linear Models Duke University

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Bayesian Estimation

Model

$$\mathbf{Y} \sim \mathsf{N}(\mathbf{X}\boldsymbol{\beta}, \mathbf{I}_n/\phi)$$

with precision $\phi = 1/\sigma^2$.

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More Prior Choices:

- More on g-priors
- Zellner-Siow Cauchy Prior
- Utility and choice of Estimators

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Note

$$(\mathbf{X}^{T}(\mathbf{I}_{n}-\mathbf{P}_{1})\mathbf{X}) = (\mathbf{X}^{T}(\mathbf{I}_{n}-\mathbf{P}_{1})^{T}(\mathbf{I}_{n}-\mathbf{P}_{1})\mathbf{X}) = (\mathbf{X}-\mathbf{1}_{n}\bar{\mathbf{X}}^{T})^{T}(\mathbf{X}-\mathbf{1}_{n}\bar{\mathbf{X}})$$

$$\begin{array}{rcl} \mathbf{Y} & = & \mathbf{1}\beta_0 + \mathbf{X}_1\boldsymbol{\beta} + \boldsymbol{\epsilon} \\ p(\beta_0, \phi) & \propto & 1 \\ \boldsymbol{\beta} \mid \phi & \sim & \mathsf{N}(\mathbf{0}, \frac{\mathbf{g}}{\phi}(\mathbf{X}^T(\mathbf{I}_n - \mathsf{P}_1)\mathbf{X})^{-1}) \end{array}$$

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Let $(\mathbf{X} - \mathbf{1}_n \bar{\mathbf{X}}^T)^T (\mathbf{X} - \mathbf{1}\bar{\mathbf{X}}) = SS_{\mathbf{X}} = \mathbf{U}^T \mathbf{U}$ Contribution quadratic to the log likelihood from prior after integrating out β_0

$$(\mathbf{Y}_c - \mathbf{X}_c \boldsymbol{\beta})^T (\mathbf{Y}_c - \mathbf{X}_c \boldsymbol{\beta}) + (\boldsymbol{\beta}^T \frac{\mathbf{U}^T \mathbf{U}}{g} \boldsymbol{\beta})$$



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Example

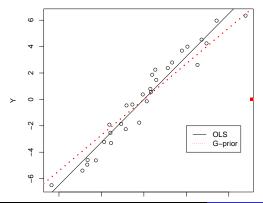
In SLR it is like an extra $Y_0=0$ at $\mathbf{X}_o=\sqrt{\frac{\mathrm{SS}_x}{g}}$:

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- Mixtures of Conjugate Priors

Theorem (Diaconis & Ylivisaker 1985)

Given a sampling model $p(y \mid \theta)$ from an exponential family, any prior distribution can be expressed as a mixture of conjugate prior distributions

• Prior $p(\theta) = \int p(\theta \mid \omega) p(\omega) d\omega$

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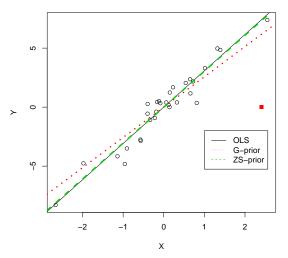
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- Choice of g?
- $\frac{g}{1+g}$ weight given to the data
- Let $\tau = 1/g$ assign $\tau \sim G(1/2, 1/2)$
- Find prior distribution
- Can expres posterior as a mixture of g-priors

Example Again

From JAGS:



JAGS Code: library(R2jags), library(R2WinBUGS)

```
model = function(){
  for (i in 1:n) {
      Y[i] ~ dnorm(X[i]*beta, phi)
  }
  beta ~ dnorm(0, SSX/(n*phi*lambda))
  phi ~ dgamma(.05, .05)
  lambda ~ dgamma(.5, .5)
write.model(model, "ZSmodel")
model.file="ZSmodel"
data = list(Y=Y, X=X, n =length(Y), SSX=sum(X^2) )
ZSout = jags(data,inits=NULL,
             parameters.to.save=c("beta", "lambda", "phi")
             model=model.file, n.iter=10000)
```

Quadratic loss for estimating β using estimator a

$$L(\boldsymbol{\beta}, \mathbf{a}) = (\boldsymbol{\beta} - \mathbf{a})^{\mathsf{T}} (\boldsymbol{\beta} - \mathbf{a})$$

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where λ_j are eigenvalues of $\mathbf{X}^T\mathbf{X}$.



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where λ_i are eigenvalues of $\mathbf{X}^T \mathbf{X}$.

- If smallest $\lambda_i \to 0$ then MSE $\to \infty$
- Note: estimate is unbiased!
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Is the *g*-prior better?

Explore Frequentist properties of using a Bayesian estimator

$$\mathsf{E}_{\mathbf{Y}}[(\beta-\hat{\boldsymbol{\beta}}_{g})^{T}(\beta-\hat{\boldsymbol{\beta}}_{g})$$

but now $\hat{\boldsymbol{\beta}}_g = g/(1+g)\hat{\boldsymbol{\beta}}$ for g prior.

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when is the g prior better than the Reference prior of OLS?

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- Solutions:
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 - other shrinkage estimators