

Study 2

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This file reproduces the preprocessing and analysis steps of Study 2. The data are automatically imported from Github and necessary packages will be downloaded and installed if they are not yet available.

Create directory to save plots:

```
if (!dir.exists('final_plots')) {dir.create('final_plots')}
```

```
rm(list=ls())
if (!require("pacman"))
  install.packages("pacman")
required_packages = c(
  "tidyverse",
  "afex",
  'brms',
  'BayesFactor',
  'emmeans',
  'gridExtra'
)
pacman::p_load(required_packages,
               update = F,
               character.only = T
)
```

```
set.seed(42)
```

```
sessionInfo()
```

```
## R version 4.0.3 (2020-10-10)
## Platform: x86_64-apple-darwin17.0 (64-bit)
## Running under: macOS Big Sur 10.16
##
## Matrix products: default
## BLAS:   /Library/Frameworks/R.framework/Versions/4.0/Resources/lib/libRblas.dylib
## LAPACK: /Library/Frameworks/R.framework/Versions/4.0/Resources/lib/libRlapack.dylib
```

```
##
## locale:
## [1] en_US.UTF-8/en_US.UTF-8/en_US.UTF-8/C/en_US.UTF-8/en_US.UTF-8
##
## attached base packages:
## [1] stats      graphics  grDevices  utils      datasets  methods   base
##
## other attached packages:
## [1] gridExtra_2.3      emmeans_1.5.3      BayesFactor_0.9.12-4.2
## [4] coda_0.19-4        brms_2.14.4         Rcpp_1.0.5
## [7] afex_0.28-0         lme4_1.1-26         Matrix_1.2-18
## [10] forcats_0.5.0      stringr_1.4.0       dplyr_1.0.2
## [13] purrr_0.3.4        readr_1.4.0         tidyr_1.1.2
## [16] tibble_3.0.4       ggplot2_3.3.2       tidyverse_1.3.0
## [19] pacman_0.5.1
##
## loaded via a namespace (and not attached):
## [1] readxl_1.3.1        backports_1.2.1     plyr_1.8.6
## [4] igraph_1.2.6        splines_4.0.3       crosstalk_1.1.0.1
## [7] TH.data_1.0-10      rstantools_2.1.1    inline_0.3.17
## [10] digest_0.6.27       htmltools_0.5.0     rsconnect_0.8.16
## [13] lmerTest_3.1-3      fansi_0.4.1         magrittr_2.0.1
## [16] openxlsx_4.2.3      modelr_0.1.8        RcppParallel_5.0.2
## [19] matrixStats_0.57.0  xts_0.12.1          sandwich_3.0-0
## [22] prettyunits_1.1.1   colorspace_2.0-0    rvest_0.3.6
## [25] haven_2.3.1         xfun_0.19           callr_3.5.1
## [28] crayon_1.3.4        jsonlite_1.7.2      survival_3.2-7
## [31] zoo_1.8-8           glue_1.4.2          gtable_0.3.0
## [34] MatrixModels_0.4-1  V8_3.4.0            car_3.0-10
## [37] pkgbuild_1.1.0      rstan_2.21.3        abind_1.4-5
## [40] scales_1.1.1        mvtnorm_1.1-1       DBI_1.1.0
## [43] miniUI_0.1.1.1      xtable_1.8-4        foreign_0.8-80
## [46] StanHeaders_2.21.0-6 stats4_4.0.3        DT_0.16
## [49] htmlwidgets_1.5.3   httr_1.4.2          threejs_0.3.3
## [52] ellipsis_0.3.1      pkgconfig_2.0.3     loo_2.4.1
## [55] dbplyr_2.0.0        tidyselect_1.1.0    rlang_0.4.9
## [58] reshape2_1.4.4      later_1.1.0.1       munsell_0.5.0
## [61] cellranger_1.1.0    tools_4.0.3         cli_2.2.0
## [64] generics_0.1.0      broom_0.7.2         ggribes_0.5.2
## [67] evaluate_0.14       fastmap_1.0.1       yaml_2.2.1
## [70] processx_3.4.5      knitr_1.30          fs_1.5.0
## [73] zip_2.1.1           pbapply_1.4-3       nlme_3.1-149
## [76] mime_0.9            projpred_2.0.2       xml2_1.3.2
## [79] compiler_4.0.3      bayesplot_1.7.2     shinythemes_1.1.2
## [82] rstudioapi_0.13     curl_4.3            gamm4_0.2-6
## [85] reprex_0.3.0        statmod_1.4.35      stringi_1.5.3
## [88] ps_1.5.0            Brodningnag_1.2-6    lattice_0.20-41
## [91] nloptr_1.2.2.2      markdown_1.1         shinyjs_2.0.0
## [94] vctrs_0.3.5         pillar_1.4.7        lifecycle_0.2.0
## [97] bridgesampling_1.0-0 estimability_1.3     data.table_1.13.4
## [100] httpuv_1.5.4        R6_2.5.0            promises_1.1.1
## [103] rio_0.5.16          codetools_0.2-16    boot_1.3-25
## [106] colourpicker_1.1.0  MASS_7.3-53         gtools_3.8.2
## [109] assertthat_0.2.1    withr_2.3.0         shinystan_2.5.0
```

```
## [112] multcomp_1.4-15      mgcv_1.8-33      parallel_4.0.3
## [115] hms_0.5.3            grid_4.0.3      minqa_1.2.4
## [118] rmarkdown_2.6        carData_3.0-4   numDeriv_2016.8-1.1
## [121] shiny_1.5.0         lubridate_1.7.9.2 base64enc_0.1-3
## [124] dygraphs_1.1.1.6
```

Import data

```
github_link = 'https://raw.githubusercontent.com/mertensu/thinking-in-ratios/master/'
file_name = 'data_total_study2.csv'
df = read.csv(paste0(github_link, file_name))
```

Preprocessing

```
# replace zero ratings with 0.001
df[df$brightness_rating == 0, 'brightness_rating'] = 0.001
# df = df[df$brightness_rating!=0, ]

# compute correct binary choice
df = df %>% mutate(brightness_dichotom_correct = ifelse(cd < 10, 0, 1))

df$correct_response = NA
df[df$condition == 'unidirectional', 'correct_response'] = df[df$condition == 'unidirectional', 'brightness_rating']

# remove wrong binary choice(s) in unidirectional condition
df = df %>% filter(!(correct_response == F &
                     condition == 'unidirectional'))

df[, "brightness_rating_stevens"] = ifelse(
  (df$condition == "unidirectional") &
  (df$brightness_dichotom_correct == 0),
  100 / df$brightness_rating,
  df$brightness_rating
)

df$log_brightness_rating = log(df$brightness_rating_stevens)

df$cd_factor = factor(df$cd, levels = c(1, 1.8, 3.2, 5.7, 17.9, 32.0, 57.2, 100.0))

df$brightness_rating_stevens = df$brightness_rating_stevens / 100
df$cd = df$cd / 100
```

Demographics

```
psych::describe(df$age)
```

```
##      vars      n mean   sd median trimmed  mad min max range skew kurtosis   se
## X1      1 1359 23.77 5.83      23   22.67 2.97  18  46    28 2.57      6.88 0.16
```

```
df %>% distinct(File, .keep_all = T) %>% group_by(gender) %>% summarise(
  N = n(),
  Min =
    min(age),
  Max =
    max(age),
  Mean =
    mean(age),
  Sd =
    sd(age)
)
```

```
## 'summarise()' ungrouping output (override with '.groups' argument)
```

```
## # A tibble: 2 x 6
##   gender      N   Min   Max  Mean    Sd
##   <chr> <int> <int> <int> <dbl> <dbl>
## 1 m         5    21    46   28   10.2
## 2 w        29    18    43  23.0   4.75
```

```
df %>% distinct(File, .keep_all = T) %>% count(student)
```

```
##   student  n
## 1         0 3
## 2         1 31
```

```
df %>% distinct(File, .keep_all = T) %>% filter(student == 1) %>% count(psycho)
```

```
##   psycho  n
## 1         0 21
## 2         1 10
```

```
df %>% distinct(File, .keep_all = T) %>% count(condition)
```

```
##           condition  n
## 1           standard 16
## 2  unidirectional 18
```

Analysis

ANOVA I (mixed 2(method) x 8(luminance))

```
(
  fit = aov_ez(
```

```

    dv = "log_brightness_rating",
    within = 'cd_factor',
    between = "condition",
    id = "File",
    data = df
  )
)

```

frequentist fit

```

## Converting to factor: condition

## Warning: More than one observation per cell, aggregating the data using mean
## (i.e, fun_aggregate = mean)!

## Contrasts set to contr.sum for the following variables: condition

## Anova Table (Type 3 tests)
##
## Response: log_brightness_rating
##           Effect             df  MSE          F ges p.value
## 1           condition             1, 32 0.14   21.73 *** .090  <.001
## 2           cd_factor 2.02, 64.70 0.39 176.75 *** .825  <.001
## 3 condition:cd_factor 2.02, 64.70 0.39    4.71 * .112   .012
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Sphericity correction method: GG

```

```

df$File = factor(df$File)
df$condition = factor(df$condition)
bfs = anovaBF(
  log_brightness_rating ~ cd_factor * condition + File,
  whichRandom = 'File',
  whichModels = 'top',
  data = df
)

# BF cd_factor
bf_1 = lmBF(log_brightness_rating ~ condition + File,
            whichRandom = 'File',
            data = df)
bf_2 = lmBF(
  log_brightness_rating ~ cd_factor + condition + File,
  whichRandom = 'File',
  data = df
)
(bf_cd_factor = bf_2 / bf_1)

```

bayesian fit

```
## Bayes factor analysis
## -----
## [1] cd_factor + condition + File : 1.136327e+326 ±3.12%
##
## Against denominator:
##   log_brightness_rating ~ condition + File
## ---
## Bayes factor type: BFlinearModel, JZS
```

```
print(paste0('logBF ', bf_cd_factor@bayesFactor$bf))
```

```
## [1] "logBF 750.770541046337"
```

```
# BF condition
bf_1 = lmBF(log_brightness_rating ~ cd_factor + File,
            whichRandom = 'File',
            data = df)
bf_2 = lmBF(
  log_brightness_rating ~ cd_factor + condition + File,
  whichRandom = 'File',
  data = df
)
(bf_condition = bf_2 / bf_1)
```

```
## Bayes factor analysis
## -----
## [1] cd_factor + condition + File : 34.16728 ±3.8%
##
## Against denominator:
##   log_brightness_rating ~ cd_factor + File
## ---
## Bayes factor type: BFlinearModel, JZS
```

```
print(paste0('logBF ', bf_condition@bayesFactor$bf))
```

```
## [1] "logBF 3.53126845930751"
```

```
# BF interaction
bf_1 = lmBF(
  log_brightness_rating ~ cd_factor + condition + File,
  whichRandom = 'File',
  data = df
)
bf_2 = lmBF(
  log_brightness_rating ~ cd_factor * condition + File,
  whichRandom = 'File',
  data = df
)
(bf_interaction = bf_2 / bf_1)
```

```
## Bayes factor analysis
```

```
## -----
## [1] cd_factor * condition + File : 371867928683 ±2.36%
##
## Against denominator:
##   log_brightness_rating ~ cd_factor + condition + File
## ---
## Bayes factor type: BFlinearModel, JZS
```

```
print(paste0('logBF ', bf_interaction@bayesFactor$bf))
```

```
## [1] "logBF 26.6418045977596"
```

```
scaleFUN <- function(x)
  sprintf("%.1f", x)

grid = data.frame(emmeans(fit, ~ cd_factor + condition))

ggplot(grid, aes(
  x = cd_factor,
  y = exp(emmean),
  group = condition
)) +
  geom_pointrange(aes(
    ymin = exp(lower.CL),
    ymax = exp(upper.CL),
    color = condition
  ), size =
    0.3) +
  scale_y_continuous(
    trans = 'log2',
    breaks = c(1.0, 10.0, 100.0),
    limits = c(1.0, 100.0),
    labels = scaleFUN
  ) +
  geom_line(linetype = 'dashed') +
  labs(y = 'Brightness judgement') +
  xlab(expression(paste("cd/", m ^ 2, sep = ""))) +
  scale_color_manual(
    values = c("darkgrey", "black"),
    name = "method",
    labels = c("unidirectional", "standard")
  ) +
  scale_x_discrete(labels = substring(grid$cd_factor, 2)) +
  theme_classic() +
  theme(
    legend.position = c(0.2, 0.8),
    legend.background = element_rect(color = "black")
  )
```

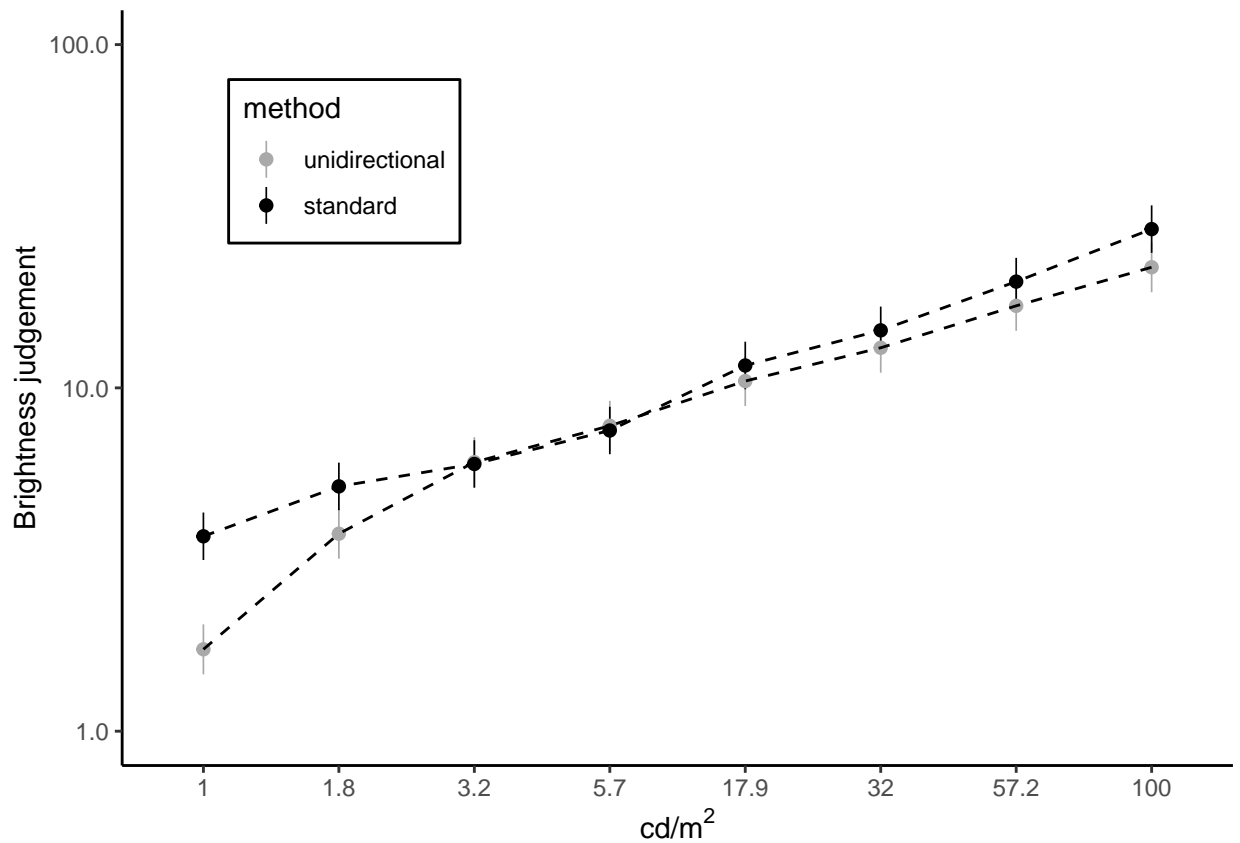


Figure 2

```
ggsave(
  paste0("final_plots/study2_figure2.png"),
  dpi = 600,
  height = 4,
  width = 5,
  units = "in"
)
```

Bayesian mixed effects models (standard condition)

```
prior_settings <- c(set_prior('normal(0, 5)', nlpar = "a"),
  set_prior('normal(1.0, 0.5)', nlpar = "b", lb = 0))
```

```
(formula = bf(brightness_rating_stevens ~ a * cd ^ b, a ~ (1 | File), b ~ (1 | File), nl = TRUE))
```

Fit power law

```
## brightness_rating_stevens ~ a * cd^b
## a ~ (1 | File)
## b ~ (1 | File)
```



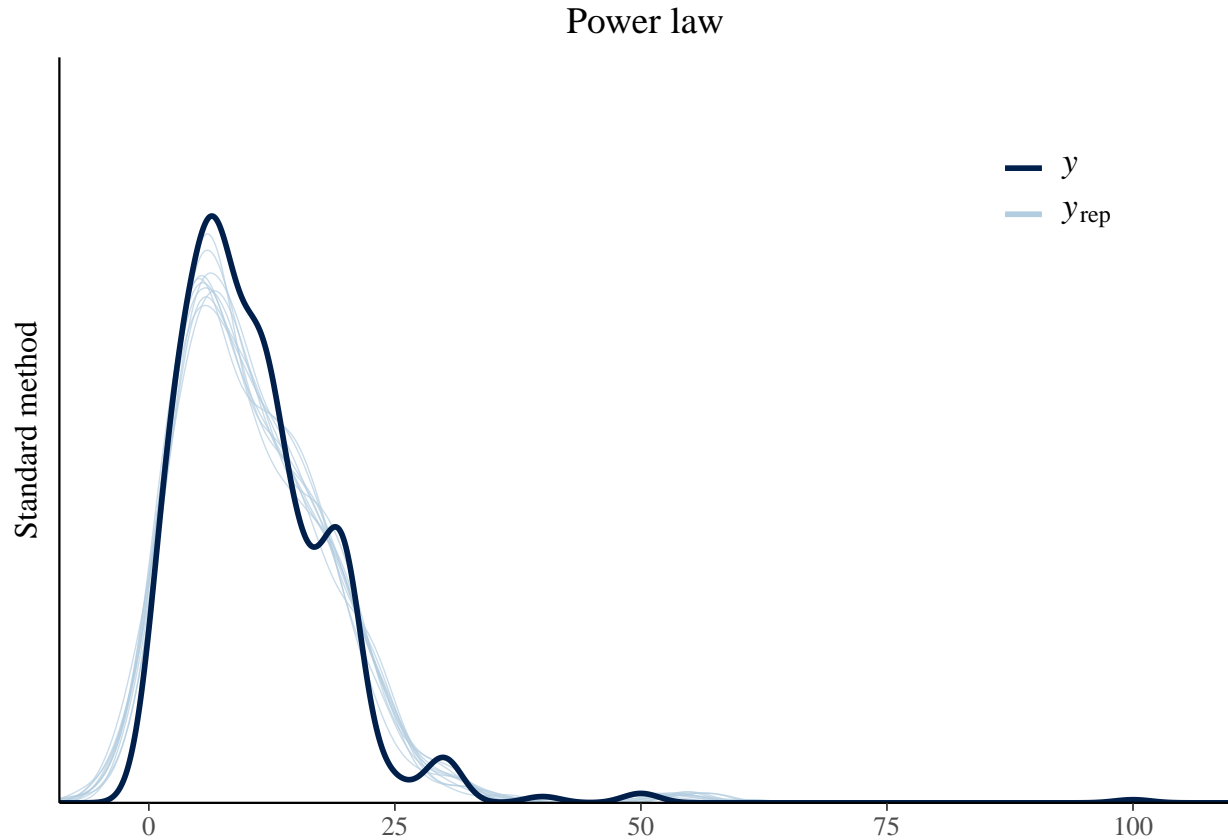
```
# make_stancode(formula, data = df[df$condition == 'standard', ], prior = prior_settings)
```

```
(power_law_standard <-
  brm(formula,
    data = df[df$condition == 'standard', ],
    save_all_pars = T,
    sample_prior = 'yes',
    warmup = 2000,
    iter = 7000,
    cores = 4,
    control = list(adapt_delta = 0.96, max_treedepth = 20),
    prior = prior_settings,
    file = 'study2_power_standard_x'))
```

```
## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: brightness_rating_stevens ~ a * cd^b
##      a ~ (1 | File)
##      b ~ (1 | File)
## Data: df[df$condition == "standard", ] (Number of observations: 640)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##      total post-warmup samples = 20000
##
## Group-Level Effects:
## ~File (Number of levels: 16)
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(a_Intercept)    0.09     0.02    0.06    0.14 1.00     5851     9960
## sd(b_Intercept)    0.18     0.04    0.12    0.27 1.00     5892     8201
##
## Population-Level Effects:
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## a_Intercept    0.23     0.02    0.19    0.28 1.00     2723     5081
## b_Intercept    0.44     0.05    0.35    0.53 1.00     3523     6853
##
## Family Specific Parameters:
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma    0.03     0.00    0.03    0.03 1.00     22498     13684
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
(
  pow_standard_plot = pp_check(power_law_standard) +
    labs(title = 'Power law', y = 'Standard method') +
    scale_x_continuous(labels = c(0, 25, 50, 75, 100), limits=c(NA,1.1)) +
    scale_y_continuous(limits=c(0,8)) +
    theme(
      plot.title = element_text(hjust = 0.5),
      legend.position = c(0.85, 0.85),
    )
)
```

```
## Using 10 posterior samples for ppc type 'dens_overlay' by default.
```



```
brms::loo(power_law_standard)
```

Model metrics (LOOCV, WAIC, marginal likelihood)

```
## Warning: Found 4 observations with a pareto_k > 0.7 in model
## 'power_law_standard'. It is recommended to set 'moment_match = TRUE' in order to
## perform moment matching for problematic observations.
```

```
##
## Computed from 20000 by 640 log-likelihood matrix
##
##      Estimate    SE
## elpd_loo  1254.8 142.1
## p_loo      89.2  51.9
## looic     -2509.6 284.2
## -----
## Monte Carlo SE of elpd_loo is NA.
##
## Pareto k diagnostic values:
##              Count Pct.    Min. n_eff
## (-Inf, 0.5] (good)   636  99.4%   3143
```

```
## (0.5, 0.7] (ok) 0 0.0% <NA>
## (0.7, 1] (bad) 2 0.3% 57
## (1, Inf) (very bad) 2 0.3% 2
## See help('pareto-k-diagnostic') for details.
```

```
brms::waic(power_law_standard)
```

```
## Warning:
## 15 (2.3%) p_waic estimates greater than 0.4. We recommend trying loo instead.

##
## Computed from 20000 by 640 log-likelihood matrix
##
##           Estimate      SE
## elpd_waic  1228.0 168.9
## p_waic      116.0  79.0
## waic       -2456.0 337.8
##
## 15 (2.3%) p_waic estimates greater than 0.4. We recommend trying loo instead.
```

```
brms::bridge_sampler(power_law_standard)
```

```
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6

## Bridge sampling estimate of the log marginal likelihood: 1223.834
## Estimate obtained in 6 iteration(s) via method "normal".
```

```
prior_settings <-
  c(
    set_prior('normal(0,1)', class = 'b', coef = 'cd'),
    set_prior('normal(0,5)', class = 'b', coef = 'Intercept')
  )
```

```
(formula = brightness_rating_stevens ~ 0 + Intercept + cd + (1 + cd | File))
```

Fit linear model

```
## brightness_rating_stevens ~ 0 + Intercept + cd + (1 + cd | File)
```

```
# make_stancode(formula, data = df[df$condition == 'standard', ], prior = prior_settings)
```

```

(linear_standard <-
  brm(
    formula,
    data = df[df$condition == 'standard', ],
    save_all_pars = T,
    sample_prior = 'yes',
    warmup = 2000,
    cores = 4,
    prior = prior_settings,
    iter = 7000,
    file = 'study2_linear_standard_x'))

## Warning: There were 9 divergent transitions after warmup. Increasing adapt_delta
## above 0.8 may help. See http://mc-stan.org/misc/warnings.html#divergent-
## transitions-after-warmup

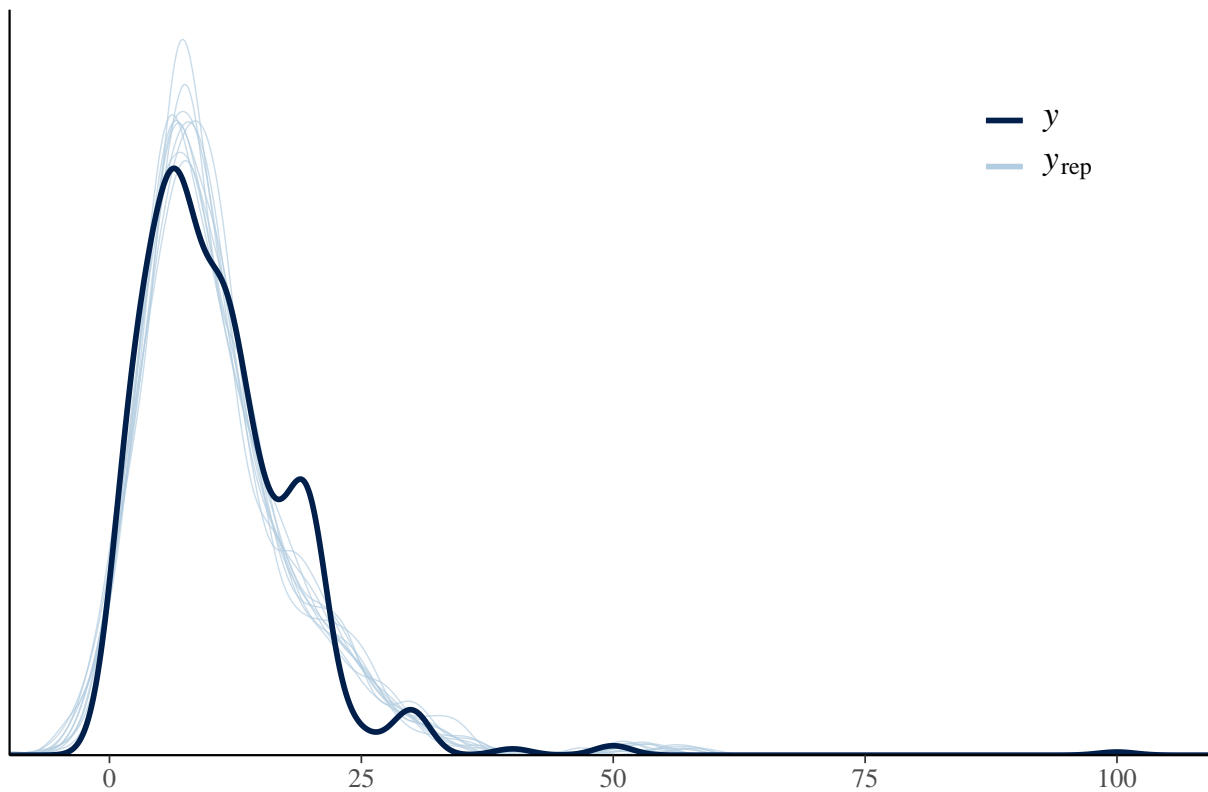
## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: brightness_rating_stevens ~ 0 + Intercept + cd + (1 + cd | File)
## Data: df[df$condition == "standard", ] (Number of observations: 640)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##           total post-warmup samples = 20000
##
## Group-Level Effects:
## ~File (Number of levels: 16)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.01      0.00    0.01    0.02 1.00     9842    11547
## sd(cd)             0.10      0.02    0.07    0.14 1.00     8200    10097
## cor(Intercept,cd)  -0.91      0.09   -1.00   -0.66 1.00     8033    10322
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept      0.05      0.00    0.05    0.06 1.00     7984     9466
## cd             0.20      0.03    0.15    0.25 1.00     7436     8469
##
## Family Specific Parameters:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma      0.04      0.00    0.03    0.04 1.00    26295    14071
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).

(
  lin_standard_plot = pp_check(linear_standard) + labs(title = 'Linear model') +
    scale_x_continuous(labels = c(0, 25, 50, 75, 100), limits=c(NA,1.1)) +
    scale_y_continuous(limits=c(0,8)) +
    theme(
      plot.title = element_text(hjust = 0.5),
      legend.position = c(0.85, 0.85),
    )
)

```

```
## Using 10 posterior samples for ppc type 'dens_overlay' by default.
```

Linear model



```
brms::loo(linear_standard)
```

Model metrics (LOOCV, WAIC, marginal likelihood)

```
## Warning: Found 2 observations with a pareto_k > 0.7 in model 'linear_standard'.
## It is recommended to set 'moment_match = TRUE' in order to perform moment
## matching for problematic observations.
```

```
##
## Computed from 20000 by 640 log-likelihood matrix
##
##           Estimate      SE
## elpd_loo  1189.7 115.2
## p_loo      69.8  40.1
## looic     -2379.4 230.5
## -----
## Monte Carlo SE of elpd_loo is NA.
##
## Pareto k diagnostic values:
##               Count Pct.    Min. n_eff
## (-Inf, 0.5] (good)   636  99.4%   2172
```

```
## (0.5, 0.7] (ok) 2 0.3% 1016
## (0.7, 1] (bad) 1 0.2% 32
## (1, Inf) (very bad) 1 0.2% 1
## See help('pareto-k-diagnostic') for details.
```

```
brms::waic(linear_standard)
```

```
## Warning:
## 13 (2.0%) p_waic estimates greater than 0.4. We recommend trying loo instead.

##
## Computed from 20000 by 640 log-likelihood matrix
##
##           Estimate      SE
## elpd_waic  1179.1 126.3
## p_waic      80.3  51.3
## waic       -2358.3 252.5
##
## 13 (2.0%) p_waic estimates greater than 0.4. We recommend trying loo instead.
```

```
brms::bridge_sampler(linear_standard)
```

```
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6

## Bridge sampling estimate of the log marginal likelihood: 1169.17
## Estimate obtained in 6 iteration(s) via method "normal".
```

Model comparison (Bayes factors) Compute logBF in favor of power law. Takes a few minutes to run, hence commented out.

```
bfs_standard = vector('numeric', length = 5)
for (i in 1:5) {
  bfs_standard[i] = log(1 / bayes_factor(linear_standard, power_law_standard)$bf)
}
```

```
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
```

```
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 7
## Iteration: 8
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 7
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 7
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
```

```
print(paste0('Mean logBF:', mean(bfs_standard), 'Std logBF:', sd(bfs_standard)))
```

```
## [1] "Mean logBF:54.6536442422723Std logBF:0.00576493641493995"
```

Bayesian mixed-effects models (unidirectional condition)

```
prior_settings <- c(set_prior('normal(0, 5)', nlpar = "a"),
  set_prior('normal(1.0, 0.5)', nlpar = "b", lb = 0)
)
```

```
(formula = bf(brightness_rating_stevens ~ a * cd ^ b, a ~ (1 | File), b ~ (1 | File), nl = TRUE))
```

Fit power law

```
## brightness_rating_stevens ~ a * cd^b
## a ~ (1 | File)
## b ~ (1 | File)
```

```
# make_stancode(formula, data = df[df$condition == 'unidirectional', ], prior = prior_settings)
```

```
(power_law_unidirectional <-
  brm(formula,
    data = df[df$condition == 'unidirectional', ],
    save_all_pars = T,
    sample_prior = 'yes',
    warmup = 2000,
    iter = 7000,
    cores = 4,
    control = list(adapt_delta = 0.96, max_treedepth = 20),
    prior = prior_settings,
    file = 'study2_power_unidirectional_x'))
```

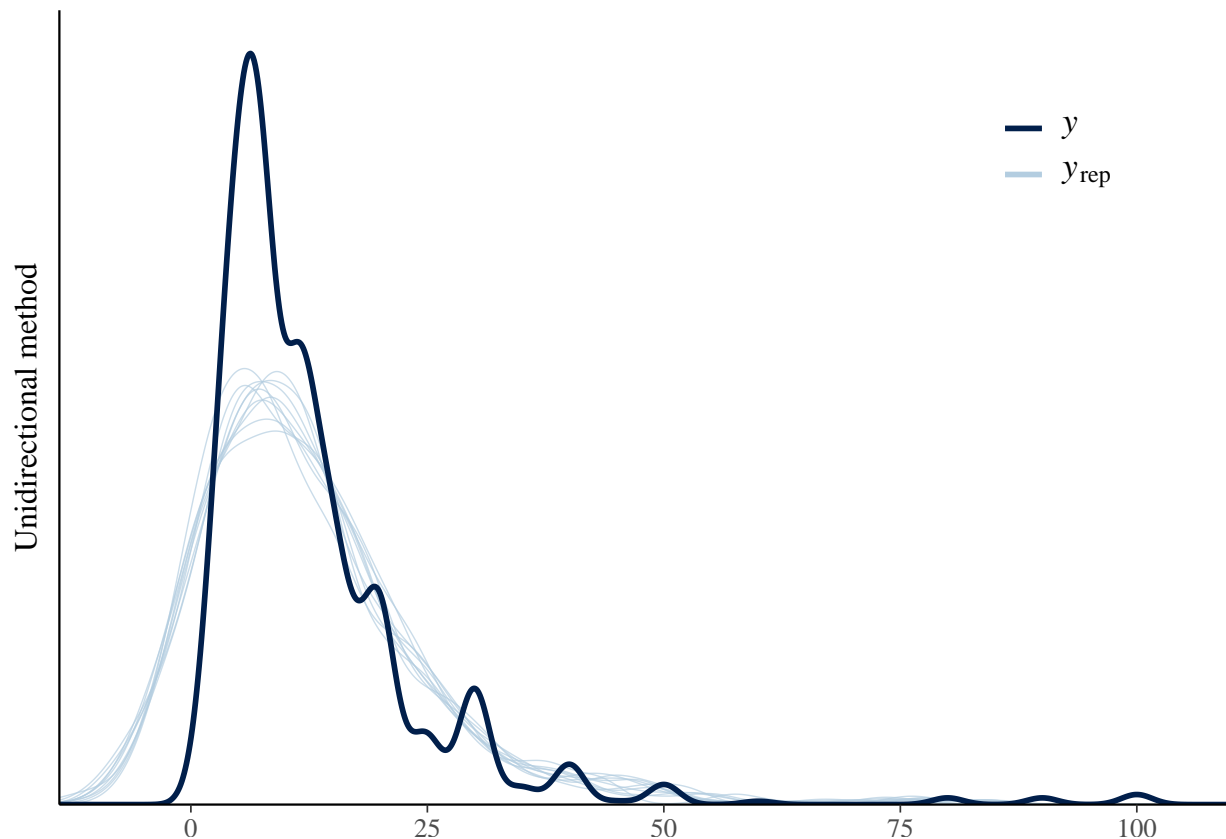
```
## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: brightness_rating_stevens ~ a * cd^b
##      a ~ (1 | File)
##      b ~ (1 | File)
## Data: df[df$condition == "unidirectional", ] (Number of observations: 719)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##      total post-warmup samples = 20000
##
## Group-Level Effects:
## ~File (Number of levels: 18)
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(a_Intercept)    0.17    0.03    0.12    0.24 1.00    3967    7473
## sd(b_Intercept)    0.30    0.06    0.21    0.45 1.00    4096    7561
##
```



```
## Population-Level Effects:
##           Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## a_Intercept    0.32     0.04    0.24    0.40 1.00    2575    4882
## b_Intercept    0.55     0.07    0.41    0.70 1.00    3468    5553
##
## Family Specific Parameters:
##           Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma         0.05     0.00    0.05    0.05 1.00   19667   15132
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
(
  pow_unidirectional_plot = pp_check(power_law_unidirectional) +
    labs(y = 'Unidirectional method') +
    scale_x_continuous(
      labels = c(0, 25, 50, 75, 100),
      breaks = c(0, 0.25, 0.5, 0.75, 1),
      limits = c(NA, 1.1)
    ) +
    scale_y_continuous(limits = c(0, 8)) +
    theme(legend.position = c(0.85, 0.85))
)
```

```
## Using 10 posterior samples for ppc type 'dens_overlay' by default.
```



```
brms::loo(power_law_unidirectional)
```

Model metrics (LOOCV, WAIC, marginal likelihood)

```
## Warning: Found 7 observations with a pareto_k > 0.7 in model
## 'power_law_unidirectional'. It is recommended to set 'moment_match = TRUE' in
## order to perform moment matching for problematic observations.
```

```
##
## Computed from 20000 by 719 log-likelihood matrix
##
##           Estimate      SE
## elpd_loo    1074.0 104.1
## p_loo        119.3  37.4
## looic       -2148.0 208.2
## -----
## Monte Carlo SE of elpd_loo is NA.
##
## Pareto k diagnostic values:
##           Count Pct.    Min. n_eff
## (-Inf, 0.5] (good)    709  98.6%  1665
## (0.5, 0.7] (ok)         3   0.4%   305
## (0.7, 1] (bad)         2   0.3%    88
## (1, Inf) (very bad)    5   0.7%     3
## See help('pareto-k-diagnostic') for details.
```

```
brms::waic(power_law_unidirectional)
```

```
## Warning:
## 23 (3.2%) p_waic estimates greater than 0.4. We recommend trying loo instead.
```

```
##
## Computed from 20000 by 719 log-likelihood matrix
##
##           Estimate      SE
## elpd_waic    1069.0 106.7
## p_waic        124.2  40.1
## waic         -2138.0 213.4
##
## 23 (3.2%) p_waic estimates greater than 0.4. We recommend trying loo instead.
```

```
brms::bridge_sampler(power_law_unidirectional)
```

```
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
```

```
## Bridge sampling estimate of the log marginal likelihood: 1052.961
## Estimate obtained in 6 iteration(s) via method "normal".
```

```
prior_settings <-
  c(
    set_prior('normal(0,1)', class = 'b', coef = 'cd'),
    set_prior('normal(0,5)', class = 'b', coef = 'Intercept')
  )
```

```
(formula = brightness_rating_stevens ~ 0 + Intercept + cd + (1 + cd | File))
```

Fit linear model

```
## brightness_rating_stevens ~ 0 + Intercept + cd + (1 + cd | File)
```

```
# make_stancode(formula, data = df[df$condition == 'unidirectional', ], prior = prior_settings)
```

```
(linear_unidirectional <-
  brm(
    formula,
    data = df[df$condition == 'unidirectional', ],
    save_all_pars = T,
    sample_prior = 'yes',
    warmup = 2000,
    cores = 4,
    prior = prior_settings,
    iter = 7000,
    file = 'study2_linear_unidirectional_x'))
```

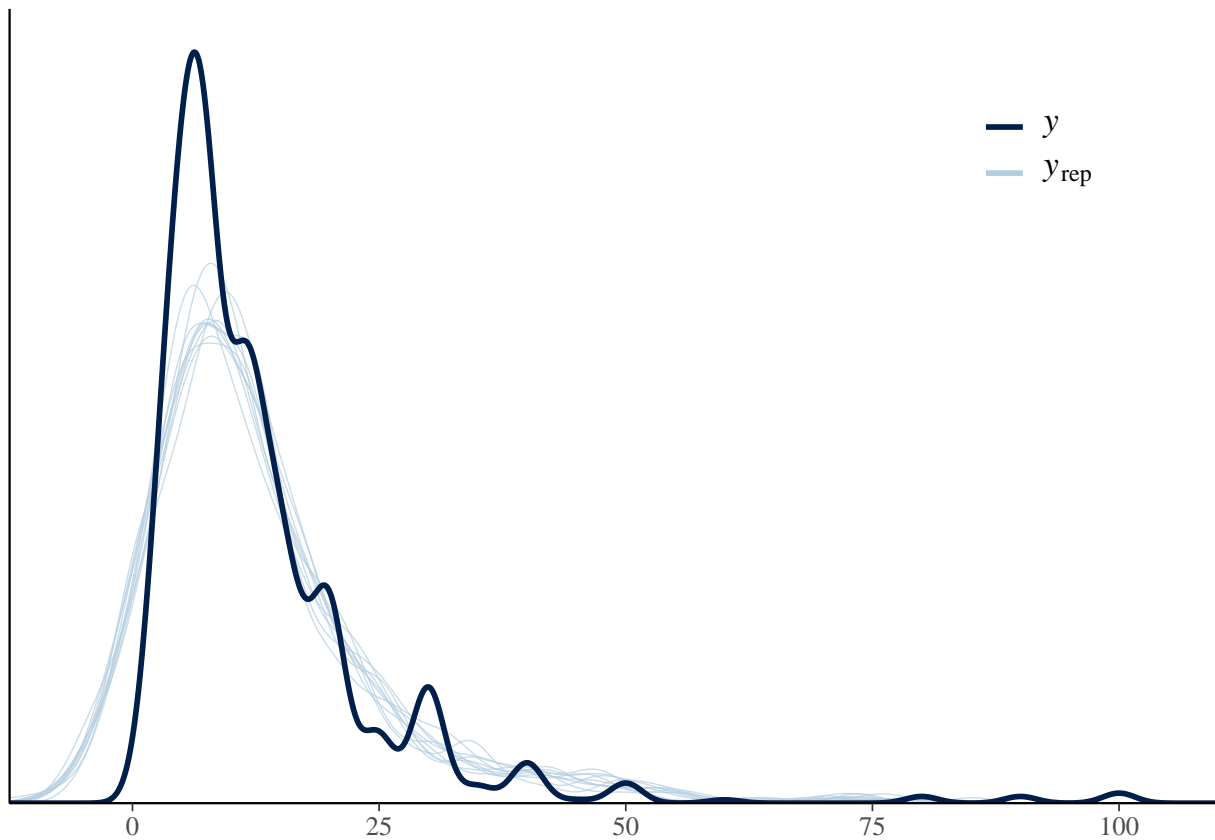
```
## Warning: There were 1 divergent transitions after warmup. Increasing adapt_delta
## above 0.8 may help. See http://mc-stan.org/misc/warnings.html#divergent-
## transitions-after-warmup
```

```
## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: brightness_rating_stevens ~ 0 + Intercept + cd + (1 + cd | File)
## Data: df[df$condition == "unidirectional", ] (Number of observations: 719)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##           total post-warmup samples = 20000
##
## Group-Level Effects:
## ~File (Number of levels: 18)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.02      0.00   0.01   0.03 1.00    8165    11658
## sd(cd)              0.17      0.03   0.12   0.25 1.00    6883    10056
## cor(Intercept,cd)  -0.98      0.02  -1.00  -0.92 1.00   11423    12818
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept      0.05      0.01   0.04   0.06 1.00    5306     8291
## cd             0.28      0.04   0.20   0.36 1.00    4713     8208
```

```
##
## Family Specific Parameters:
##      Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma      0.05      0.00    0.05    0.05 1.00    41479    13515
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
(
  lin_unidirectional_plot = pp_check(linear_unidirectional) +
    scale_x_continuous(
      labels = c(0, 25, 50, 75, 100),
      breaks = c(0, 0.25, 0.5, 0.75, 1),
      limits = c(NA, 1.1)
    ) +
    scale_y_continuous(limits = c(0, 8)) +
    theme(legend.position = c(0.85, 0.85))
)
```

```
## Using 10 posterior samples for ppc type 'dens_overlay' by default.
```



```
brms::loo(linear_unidirectional)
```

Model metrics (LOOCV, WAIC, marginal likelihood)

```
## Warning: Found 6 observations with a pareto_k > 0.7 in model
## 'linear_unidirectional'. It is recommended to set 'moment_match = TRUE' in order
## to perform moment matching for problematic observations.
```

```
##
## Computed from 20000 by 719 log-likelihood matrix
##
##           Estimate      SE
## elpd_loo    1077.9   99.4
## p_loo        96.1   30.3
## looic       -2155.8 198.9
## -----
## Monte Carlo SE of elpd_loo is NA.
##
## Pareto k diagnostic values:
##           Count Pct.    Min. n_eff
## (-Inf, 0.5] (good)    708   98.5%   2491
## (0.5, 0.7] (ok)         5    0.7%    561
## (0.7, 1] (bad)          1    0.1%     36
## (1, Inf) (very bad)    5    0.7%     5
## See help('pareto-k-diagnostic') for details.
```

```
brms::waic(linear_unidirectional)
```

```
## Warning:
## 18 (2.5%) p_waic estimates greater than 0.4. We recommend trying loo instead.
```

```
##
## Computed from 20000 by 719 log-likelihood matrix
##
##           Estimate      SE
## elpd_waic    1075.6 101.0
## p_waic        98.4  32.0
## waic         -2151.3 202.0
##
## 18 (2.5%) p_waic estimates greater than 0.4. We recommend trying loo instead.
```

```
brms::bridge_sampler(linear_unidirectional)
```

```
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
```

```
## Bridge sampling estimate of the log marginal likelihood: 1066.358
## Estimate obtained in 6 iteration(s) via method "normal".
```

Model comparison (Bayes factors) Compute logBF in favor of power law. Takes a few minutes to run, hence commented out.

```
bfs_unidirectional = vector('numeric', length = 5)
for (i in 1:5) {
  bfs_unidirectional[i] = log(1 / bayes_factor(linear_unidirectional, power_law_unidirectional)$bf)
}
```

```
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6
## Iteration: 7
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
```

```

## Iteration: 5
## Iteration: 6
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 1
## Iteration: 2
## Iteration: 3
## Iteration: 4
## Iteration: 5
## Iteration: 6

print(paste0('Mean logBF:', mean(bfs_unidirectional), 'Std logBF:', sd(bfs_unidirectional)))

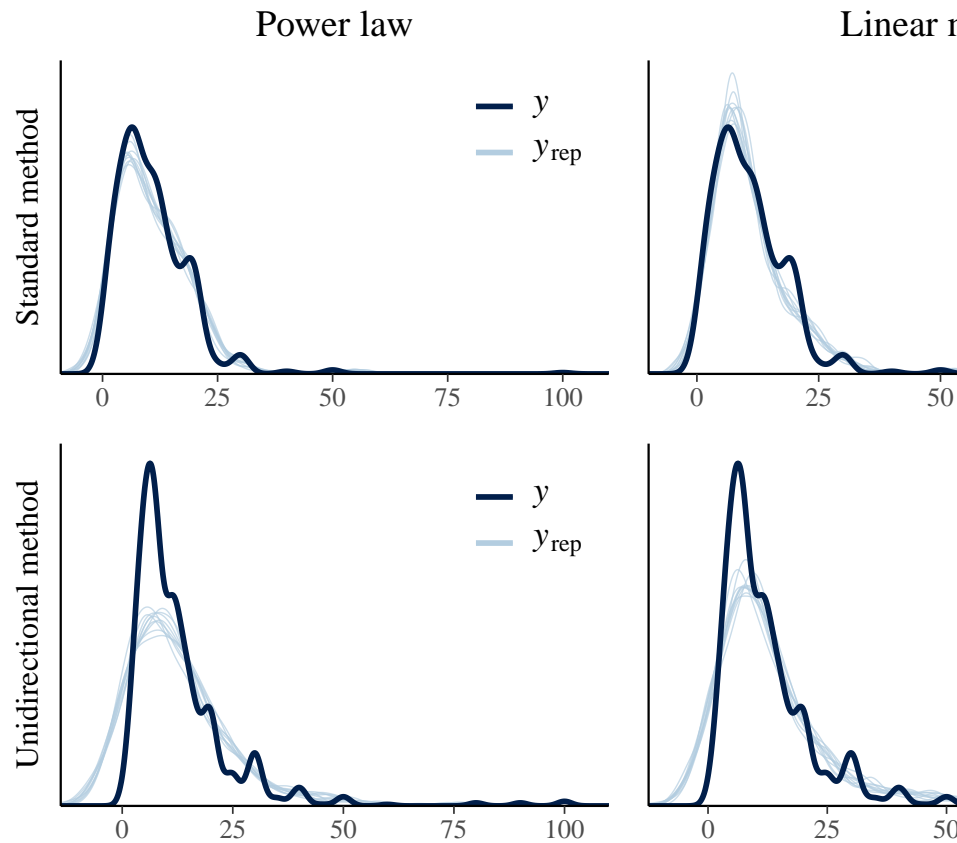
## [1] "Mean logBF:-13.3779609597161Std logBF:0.0405369791611522"

```

```

pp_grid = (
  grid.arrange(
    pow_standard_plot,
    lin_standard_plot,
    pow_unidirectional_plot,
    lin_unidirectional_plot
  )
)

```



Posterior predictive plot (Figure 3)

```
ggsave(
  paste0("final_plots/study2_pp_plot.png"),
  plot = pp_grid,
  dpi = 600,
  height = 6,
  width = 7,
  units = "in"
)
```

Bayesian mixed-effects models (aggregated data)

```
df_agg = df %>% group_by(cd, condition) %>% summarise(median_rating = median(brightness_rating_stevens))
```

Fit power law (standard condition)

```
## 'summarise()' regrouping output by 'cd' (override with '.groups' argument)
```

```
prior_settings <- c(set_prior('normal(0, 5)', nlpar = "a"),
  set_prior('normal(1.0, 0.5)', nlpar = "b", lb = 0))
```

```
(formula = bf(median_rating ~ a * cd ^ b , a ~ 1, b ~ 1 , nl = TRUE))
```



```
## median_rating ~ a * cd^b
## a ~ 1
## b ~ 1
```

```
# make_stancode(formula, data = df_agg[df_agg$condition == 'standard', ], prior = prior_settings)
```

```
(
  power_law_standard_agg <-
    brm(
      formula,
      data = df_agg[df_agg$condition == 'standard',],
      save_all_pars = T,
      sample_prior = 'yes',
      warmup = 2000,
      iter = 7000,
      cores = 4,
      control = list(adapt_delta = 0.96, max_treedepth = 20),
      prior = prior_settings,
      file = 'study2_power_standard_agg_x'
    )
)
```

```
## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: median_rating ~ a * cd^b
##      a ~ 1
##      b ~ 1
## Data: df_agg[df_agg$condition == "standard", ] (Number of observations: 8)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##      total post-warmup samples = 20000
##
## Population-Level Effects:
##      Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## a_Intercept    0.20     0.01   0.18    0.23 1.00    6778    5687
## b_Intercept    0.38     0.04   0.31    0.46 1.00    6929    6449
##
## Family Specific Parameters:
##      Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma    0.01     0.01   0.01    0.03 1.00    5258    5513
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
prior_settings <- c(set_prior('normal(0, 5)', nlpar = "a"),
  set_prior('normal(1.0, 0.5)', nlpar = "b", lb = 0))
```

```
(formula = bf(median_rating ~ a * cd ^ b , a ~ 1, b ~ 1 , nl = TRUE))
```

Fit power law on aggregated data (unidirectional condition)

```
## median_rating ~ a * cd^b
## a ~ 1
## b ~ 1

# make_stancode(formula, data = df_agg[df_agg$condition == 'unidirectional', ], prior = prior_settings)

(
  power_law_unidirectional_agg <-
    brm(
      formula,
      data = df_agg[df_agg$condition == 'unidirectional',],
      save_all_pars = T,
      sample_prior = 'yes',
      warmup = 2000,
      iter = 7000,
      cores = 4,
      control = list(adapt_delta = 0.96, max_treedepth = 20),
      prior = prior_settings,
      file = 'study2_power_unidirectional_agg_x'
    )
)

## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: median_rating ~ a * cd^b
##      a ~ 1
##      b ~ 1
## Data: df_agg[df_agg$condition == "unidirectional", ] (Number of observations: 8)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##      total post-warmup samples = 20000
##
## Population-Level Effects:
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## a_Intercept    0.28     0.02   0.25   0.33 1.00    6197    5744
## b_Intercept    0.52     0.09   0.39   0.73 1.00    6073    5480
##
## Family Specific Parameters:
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma    0.02     0.01   0.01   0.05 1.00    5536    5761
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

Supplemental analysis

```
github_link = 'https://raw.githubusercontent.com/mertensu/thinking-in-ratios/master/'
file_name = 'data_total_supp.csv'
df_supp = read.csv(paste0(github_link, file_name))
```

```

df_supp[df_supp$brightness_rating == 0, 'brightness_rating'] = 0.001

df_supp = df_supp %>% mutate(brightness_dichotom_correct = ifelse(cd < 10, 0, 1))

df_supp$correct_response = NA
df_supp[, 'correct_response'] = df_supp['brightness_dichotom_correct'] == df_supp[, 'brightness_dichotom']

df_supp$cd_factor = factor(df_supp$cd, levels = c(1, 1.8, 3.2, 5.7, 17.9, 32.0, 57.2, 100.0))

# use same column name as in Study 2
df_supp$brightness_rating_stevens = df_supp$brightness_rating
df_supp$log_brightness_rating = log(df_supp$brightness_rating_stevens)

df_supp$brightness_rating_stevens = df_supp$brightness_rating_stevens/100
df_supp$cd = df_supp$cd/100

```

```

psych::describe(df_supp$age)

```

```

##      vars      n mean      sd median trimmed  mad min max range skew kurtosis   se
## X1      1 760 21.37 2.37      21   21.14 1.48  18  27      9 0.82    -0.24 0.09

```

```

df_supp %>% distinct(File, .keep_all = T) %>% group_by(gender) %>% summarise(
  N = n(),
  Min =
    min(age),
  Max =
    max(age),
  Mean =
    mean(age),
  Sd =
    sd(age)
)

```

```

## 'summarise()' ungrouping output (override with '.groups' argument)

```

```

## # A tibble: 2 x 6
##   gender      N    Min    Max  Mean    Sd
##   <chr> <int> <int> <int> <dbl> <dbl>
## 1 m         4     18     25  21.8  2.99
## 2 w        15     19     27  21.3  2.37

```

```

df_supp %>% distinct(File, .keep_all = T) %>% count(student)

```

```

##   student  n
## 1         1 19

```

```

df_supp %>% distinct(File, .keep_all = T) %>% filter(student == 1) %>% count(psycho)

```

```
## psycho n
## 1 0 12
## 2 1 7
```

```
df_supp %>% distinct(File, .keep_all = T) %>% count(condition)
```

```
## condition n
## 1 binary_standard 19
```

```
df_s2 = df
df_s2$File = paste0(df_s2$File, 'study2')
df_tot = rbind(df_supp, df_s2)
df_tot = df_tot %>% filter(condition != 'unidirectional')
```

```
(
  fit = aov_ez(
    dv = "log_brightness_rating",
    within = 'cd_factor',
    between = "condition",
    id = "File",
    data = df_tot
  )
)
```

```
## Converting to factor: condition
```

```
## Warning: More than one observation per cell, aggregating the data using mean
## (i.e, fun_aggregate = mean)!
```

```
## Contrasts set to contr.sum for the following variables: condition
```

```
## Anova Table (Type 3 tests)
```

```
##
## Response: log_brightness_rating
##      Effect      df  MSE      F    ges p.value
## 1      condition      1, 33 0.40      0.13 <.001   .716
## 2      cd_factor 1.60, 52.89 0.90 111.11 ***   .724   <.001
## 3 condition:cd_factor 1.60, 52.89 0.90      0.24   .006   .739
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Sphericity correction method: GG
```

```
df_tot$File = factor(df_tot$File)
df_tot$condition = factor(df_tot$condition)
bfs = anovaBF(log_brightness_rating ~ cd_factor * condition + File,
              whichRandom = 'File', whichModels = 'top', data=df_tot)
```

```
# BF cd_factor
```

```
bf_1 = lmBF(log_brightness_rating ~ condition + File, whichRandom = 'File', data=df_tot)
bf_2 = lmBF(log_brightness_rating ~ cd_factor + condition + File, whichRandom = 'File', data=df_tot)
(bf_cd_factor = bf_2 / bf_1)
```

```

## Bayes factor analysis
## -----
## [1] cd_factor + condition + File : 1.594378e+279 ±7.31%
##
## Against denominator:
##   log_brightness_rating ~ condition + File
## ---
## Bayes factor type: BFlinearModel, JZS

print(paste0('logBF ', bf_cd_factor@bayesFactor$bf))

## [1] "logBF 642.887724511602"

# BF condition
bf_1 = lmBF(log_brightness_rating ~ cd_factor + File, whichRandom = 'File', data=df_tot)
bf_2 = lmBF(log_brightness_rating ~ cd_factor + condition + File, whichRandom = 'File', data=df_tot)
(bf_condition = bf_2 / bf_1)

## Bayes factor analysis
## -----
## [1] cd_factor + condition + File : 0.1640309 ±1.67%
##
## Against denominator:
##   log_brightness_rating ~ cd_factor + File
## ---
## Bayes factor type: BFlinearModel, JZS

print(paste0('logBF ', bf_condition@bayesFactor$bf))

## [1] "logBF -1.80770074176746"

# BF interaction
bf_1 = lmBF(log_brightness_rating ~ cd_factor + condition + File, whichRandom = 'File', data=df_tot)
bf_2 = lmBF(log_brightness_rating ~ cd_factor * condition + File, whichRandom = 'File', data=df_tot)
(bf_int = bf_2 / bf_1)

## Bayes factor analysis
## -----
## [1] cd_factor * condition + File : 0.002013232 ±2.23%
##
## Against denominator:
##   log_brightness_rating ~ cd_factor + condition + File
## ---
## Bayes factor type: BFlinearModel, JZS

print(paste0('logBF ', bf_int@bayesFactor$bf))

## [1] "logBF -6.20801400047162"

```

```
prior_settings <-
  c(set_prior('normal(0, 5)', nlpar = "a"),
    set_prior('normal(1.0, 0.5)', nlpar = "b", lb = 0))
```

```
(formula = bf(brightness_rating_stevens ~ a * cd ^ b , a ~ (1 | File), b ~ (1 | File), nl = TRUE))
```

Fit power law

```
## brightness_rating_stevens ~ a * cd^b
## a ~ (1 | File)
## b ~ (1 | File)
```

```
# make_stancode(formula, data = df_supp, prior = prior_settings)
```

```
(
  power_law_supp <-
    brm(
      formula,
      data = df_supp,
      save_all_pars = T,
      sample_prior = 'yes',
      warmup = 2000,
      iter = 7000,
      cores = 4,
      control = list(adapt_delta = 0.96, max_treedepth = 20),
      prior = prior_settings,
      file = 'power_binary_standard'
    )
)
```

```
## Family: gaussian
## Links: mu = identity; sigma = identity
## Formula: brightness_rating_stevens ~ a * cd^b
##      a ~ (1 | File)
##      b ~ (1 | File)
## Data: df_supp (Number of observations: 760)
## Samples: 4 chains, each with iter = 7000; warmup = 2000; thin = 1;
##          total post-warmup samples = 20000
##
## Group-Level Effects:
## ~File (Number of levels: 19)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(a_Intercept)    0.09     0.02   0.06   0.13 1.00   3549    6546
## sd(b_Intercept)    0.18     0.03   0.12   0.26 1.00   4340    7621
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## a_Intercept    0.24     0.02   0.19   0.28 1.00   2495    4228
## b_Intercept    0.42     0.04   0.34   0.51 1.00   3223    5098
```

```

##
## Family Specific Parameters:
##      Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sigma      0.03      0.00      0.03      0.03 1.00      18911      14253
##
## Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).

```