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1 Inverse transform method.

$$f(x) \Rightarrow F(x) = u \sim Unif(0,1) \Rightarrow F^{-1}(u) = x$$

2 Accept reject method

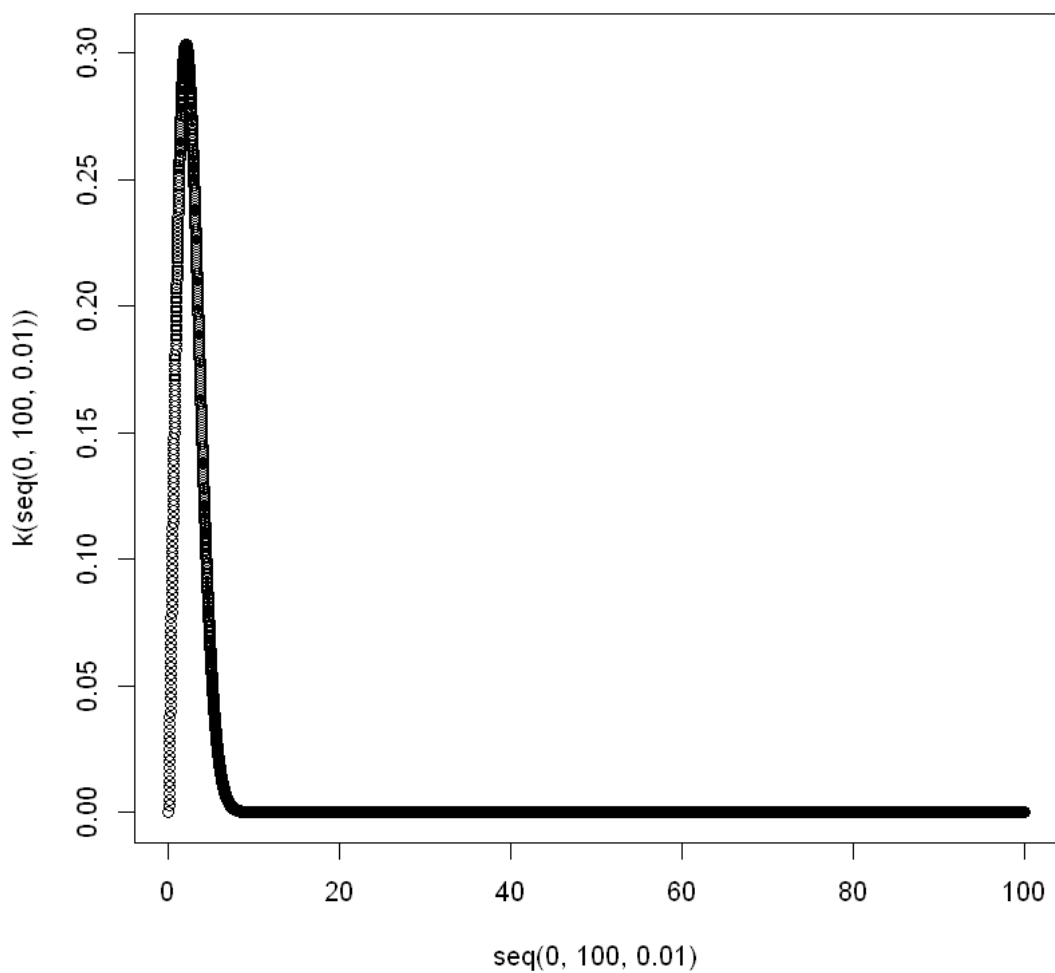
- $f(x) \Rightarrow$ we want to generate
- $g(x) \Rightarrow$ we can generate
- find $c \Rightarrow \max(f(x)/g(x))$
- generate u from $unif(0,1)$, x from g and see if $c \leq f(x)/(g(x)*c)$
- if yes x is added else discarded try again

2.1 properties

- Do not change the configuration of u NEVER
- Range of x defines the range of sample

```
[88]: f <- \(x) x/4 * exp(-(x**2)/8)
      g <- \(x) 1
```

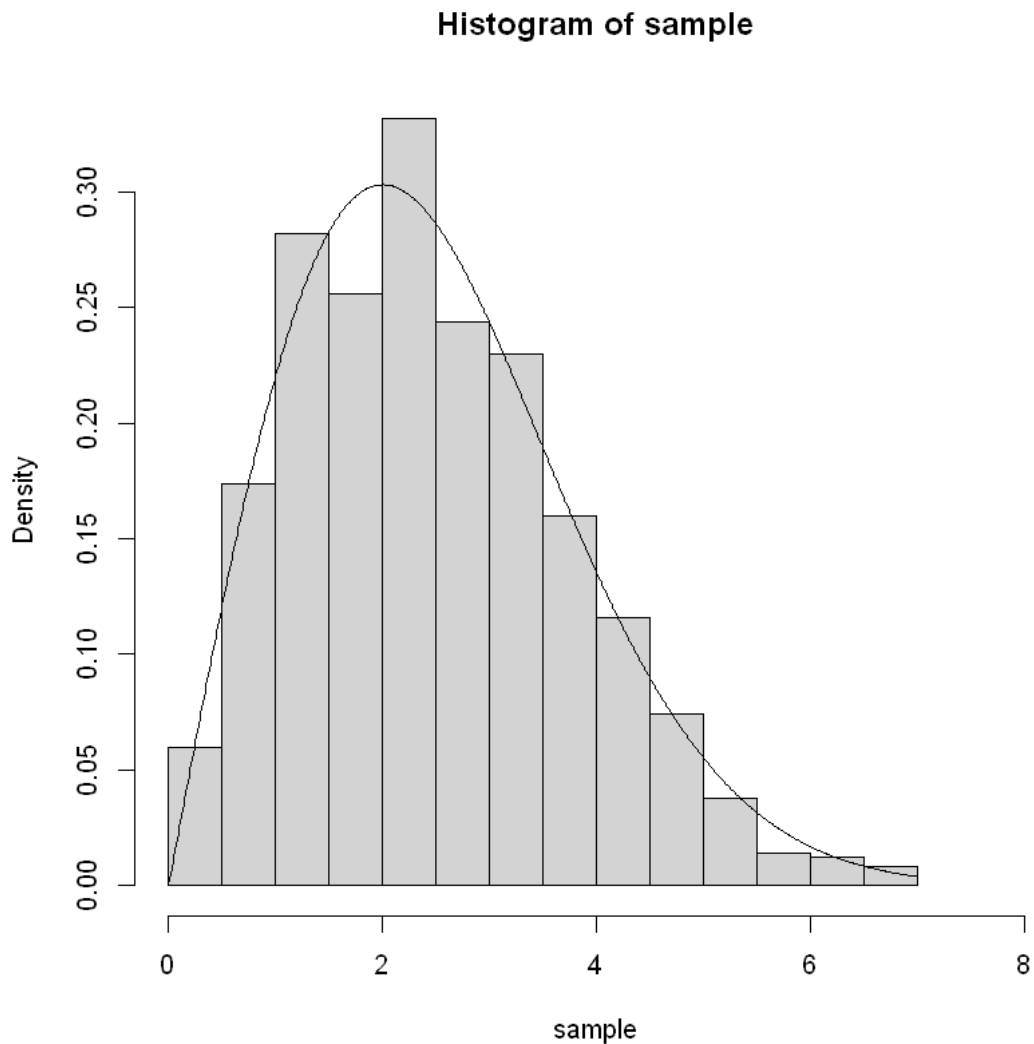
```
[89]: k <- \(x) f(x)/g(x)
      plot(seq(0,100,0.01), k(seq(0,100,0.01)))
```



```
[90]: c <- optimize(\(x) f(x)/g(x), c(0, 100), maximum=1)$objective #find max value of f(x)/g(x)
```

```
[111]: sample <- numeric(1000)
for(i in 1:1000){
  repeat{
    u <- runif(1)
    x <- runif(1, 0, 7)
    if(u <= f(x)/(c*g(x))) break
  }
  sample[i] <- x
}
```

```
[112]: hist(sample, freq = 0, xlim=c(0,8))
lines(seq(0,7,0.01), dgamma(seq(0,7,0.01),2))
```



3 sums and mixtures

- The same formulas binding distributions together can be used with the samples from these distributions.
- Summing the lines of matrix that have the same distribution items will result in sample from $\sum f(x)$

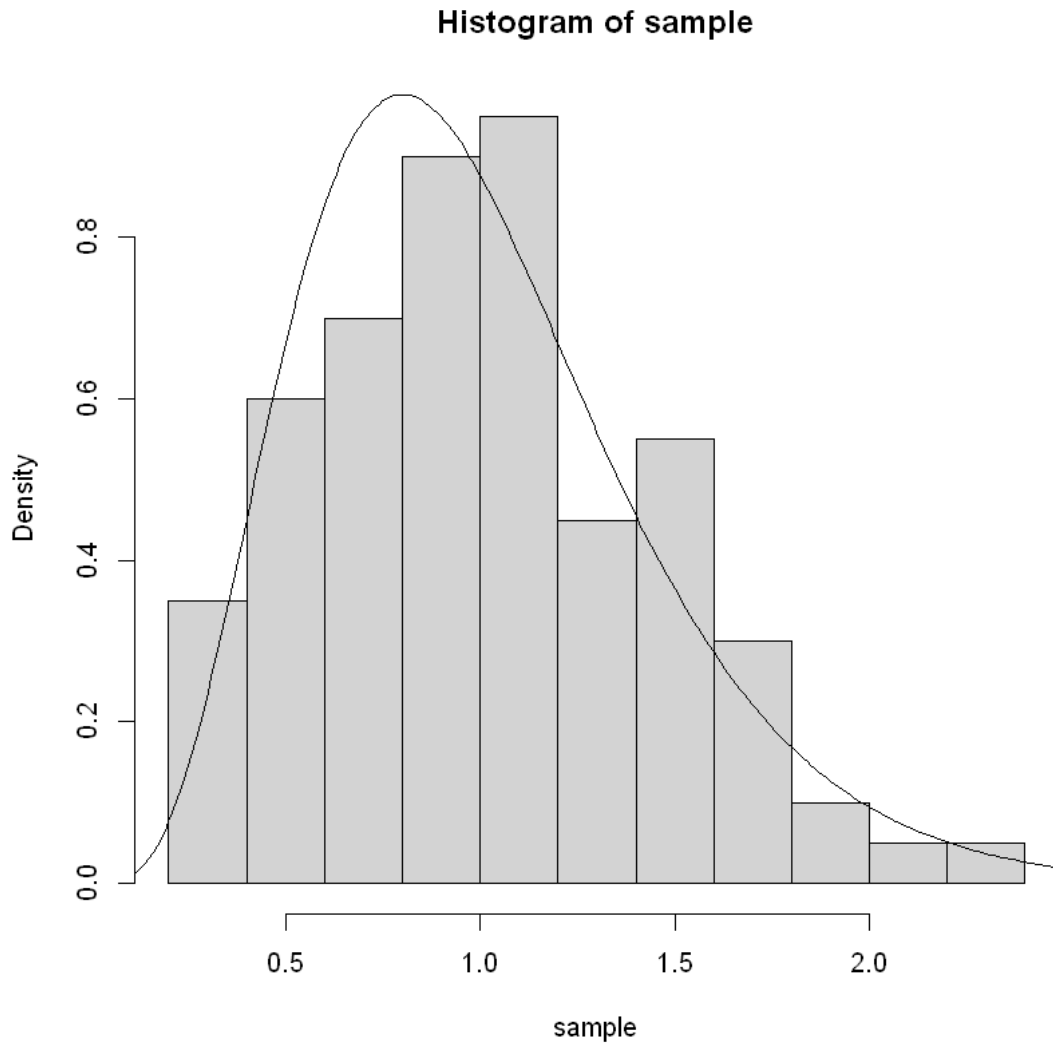
```
[68]: sample <- rgamma(100, 1, 5)
for(i in 1:4){
```

```

    sample = sample + rgamma(100, 1, 5)
}

hist(sample, freq=0)
lines(seq(0,3,0.01), dgamma(seq(0,3,0.01), 5, 5), ylim=c(0,1))

```



4 Monte Carlo

- Basically its integral, but with simulations.
- Get n random points, see how many of those points are acceptable by the region that is covered by the $f(x)$, get the ratio between accepted/all = result.
- If wanted to get $(-\infty, \infty)$ range use `rnorm` instead of `runif`

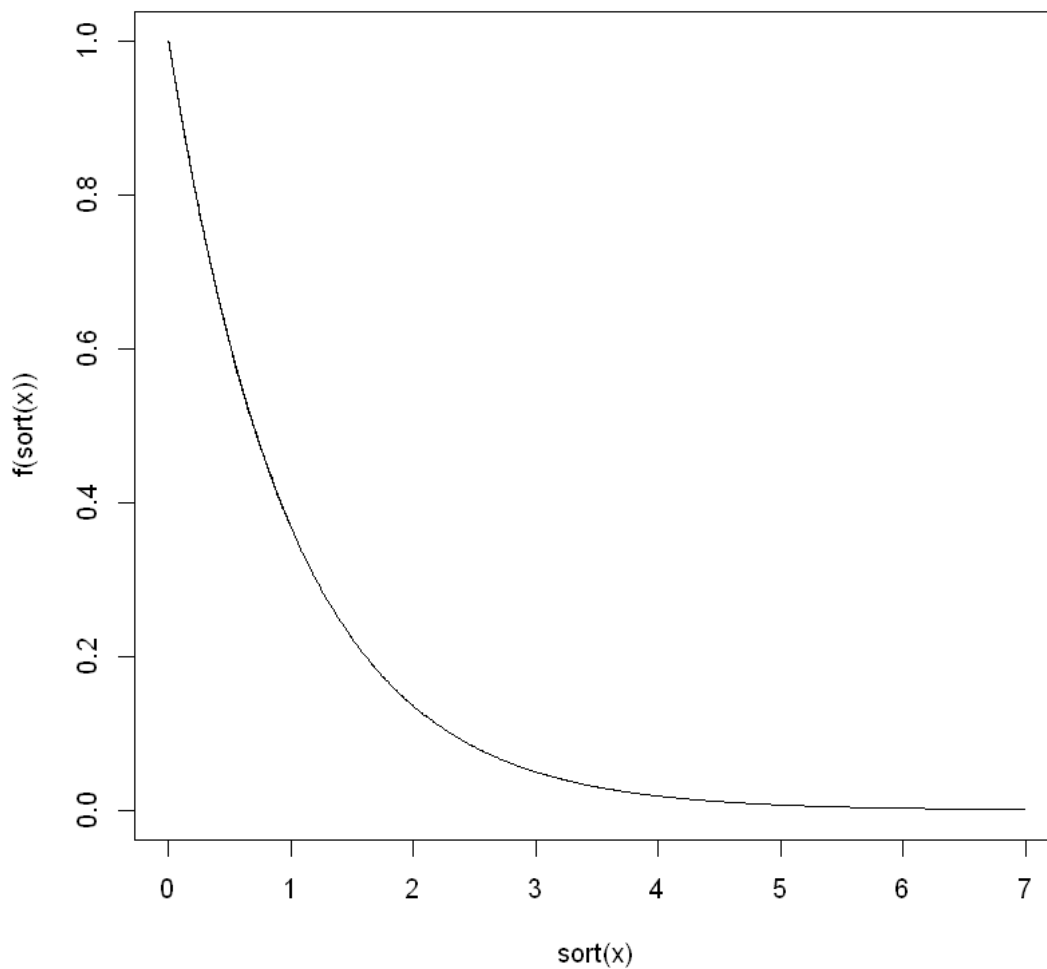
4.0.1 Example

Say; $f(x) = e^{-x}$

```
[36]: f <- \(x) exp(-x) #function
x <- runif(1000, 0, 5) #if from 0 to 5
plot(sort(x), f(sort(x)), type="l")
lines(seq(0,1,0.001), dexp(seq(0,1,0.001), 1))
res <- mean(f(x))*(5-0)
res
integrate(f, 0, 5)
```

0.907036446982713

0.9990881 with absolute error < 1.1e-14



```
[43]: k <- function(f, p1=0, p2=0){
      g <- \(x) if(x > p1){return(1)}else{return(0)}
      k <- mean(sapply(f(rnorm(1000)), g))
      return(k)
    }
    mean(k(f, 0, 1000))
    integrate(f, 0, Inf)
```

1

1 with absolute error < 5.7e-05

4.1 Variance Reduction

4.1.1 PROPERTIES

- Antithetic & Control variate methods cannot be used with ∞ boundaries

4.1.2 Antithetic Variables

- Using $u \sim Uni(a, b)$ and their counterpart $(b-u) \sim Uni(a, b)$ You can reduce variance of $f(u)$ because the newly added Covariances

```
[58]: f <- \(x) exp(-x)
      u <- runif((10**4)/2, 0, 5)
      x <- f(u)
      y <- f(5-u)
      ant <- (x+y)/2
      mean(ant)*5
      var(ant)
```

1.00100756672662

0.0141440714701547

```
[48]: f <- \(x) exp(-x)
      n <- (10**4)/2
      x <- runif(n, 0, 5)
      mean(f(x))*5
      var(f(x))
```

1.02074890175252

0.063463067682122

4.1.3 Control variate

```
[59]: #define parameters
n <- 10**4
x <- runif(n) # from 0 to 1

#define f(X) and g(X)

f.cont <- function(x) 1/(1+x)
g.cont <- function(x) 1+x
```

$$c^* = -\frac{\text{Cov}(f(x), g(x))}{\text{Var}(g(x))}$$

$$MC = \text{mean}(f(x) + c^*(g(x) - \frac{1}{b-a} \int_a^b g(x) dx)) * (b-a)$$

$$\text{Var}(MC) = \text{Var}(\theta^{c^*}) = \text{Var}(f(X)) - \frac{[\text{Cov}(f(X), g(X))]^2}{\text{Var}(g(X))}$$

```
[67]: c_star <- -cov(f.cont(x), g.cont(x)) / var(g.cont(x))
c_star
mean(f.cont(x) + c_star * (g.cont(x) - (1/(1-0))*integrate(g.cont, 0, 1)$value)) * (1-0)
integrate(f.cont, 0, 1)
```

0.477653726862992

0.693316188201991

0.6931472 with absolute error < 7.7e-15

5 Excercises from book

5.1 Monte Carlo

```
[4]: f <- \(x) sin(x)
x <- runif(10000, 0, 60)
mean(f(x))*60
```

1.74592437282211

```
[5]: integrate(\(x) sin(x), 0, 60)
```

1.952413 with absolute error < 3.4e-06

```
[12]: f <- \(x) 1/sqrt(2*pi) * exp(-1/2 * x^2)
norm <- function(x){
```

```

k <- runif(10000, 0, x)
mc <- mean(f(k))*x
var <- var(f(k))
conf_up <- mc + qnorm(0.95, lower.tail = 1)*sqrt(var/10000)
conf_down <- mc - qnorm(0.95, lower.tail = 1)*sqrt(var/10000)
cat("Estimation from mc =", mc, ", Var of mc =", var,
    "\nConfidence interval = (", conf_down, ",", conf_up, ")",
    "\nReal value =", 0.5-(1-pnorm(x)))
}
norm(2)

```

Estimation from mc = 0.476589 , Var of mc = 0.01331552
 Confidence interval = (0.474691 , 0.4784871)
 Real value = 0.4772499

```

[1]: n <- c(5, 30, 100)
M <- 10^4
mu <- 3; variance <- 5

```

```

[16]: mu_ <- matrix(0, ncol=3, nrow=M)
variance_ <- matrix(0, ncol=3, nrow=M)
for(i in 1:M){
  for(j in 1:3){
    mu_[i, j] <- mean(rnorm(n[j], mu, sqrt(variance)))
    variance_[i, j] <- var(rnorm(n[j], mu, sqrt(variance)))
  }
}
apply(mu_, 2, mean)-3
apply(variance_, 2, mean)-5

```

1. 0.0239984741423527 2. -0.00282657038733714 3. 0.000814203544638126
 1. 0.0240282628787876 2. -0.0110375322940222 3. -0.00716311065229025

```

[4]: library(tidyverse)

```

-- Attaching packages

tidyverse 1.3.1 --

v ggplot2	3.3.5	v purrr	0.3.4
v tibble	3.1.6	v dplyr	1.0.8
v tidyr	1.2.0	v stringr	1.4.0
v readr	2.1.2	v forcats	0.5.1

-- Conflicts -----

----- tidyverse_conflicts() -----

x dplyr::filter() masks stats::filter()


```
x dplyr::lag()      masks stats::lag()
```

```
[7]: apply(mu_, 2, mean)
```

```
1. -0.0202090534564321 2. -0.00939024148553381 3. -0.000799617315271596
```

```
[27]: n <- 10000
f <- function(x) exp(x)/(1+x)

start1 <- Sys.time()
u <- runif(n, 0, 1)
estimate_n <- mean(f(u))
variance_n <- var(f(u))
end1 <- Sys.time()

start2 <- Sys.time()
u <- runif(n/2, 0, 1)
x <- f(u)
y <- f(1-u)
ant <- (x+y)/2
estimate_ant <- mean(ant)
variance_ant <- var(ant)
end2 <- Sys.time()

cat("Estimate of normal mc =", estimate_n, ", variance = ", variance_n, ",\n",
    "taken time = ", as.numeric(end1-start1), "ms",
    "\nEstimate of antithetic mc =", estimate_ant, ", variance = ",\n",
    "variance_ant, ", taken time = ", as.numeric(end2-start2), "ms")
```

```
Estimate of normal mc = 1.126077 , variance = 0.01151673 , taken time =
0.007977009 ms
Estimate of antithetic mc = 1.124603 , variance = 0.0005687255 , taken time =
0.007979155 ms
```

```
[32]: f <- \x 1/(2*sqrt(2*pi)) * exp(-1/2 * ((x-8)/2)**2)
z_calculated = (7.5-8)/(2/sqrt(150))
u <- runif(n, -10, z_calculated)
mean(f(u))*(z_calculated+10)
```

```
0.0000000147264651838957
```

as this is lower than 0.05 we can reject the null hypothesis. So yes they indeed drink less water.

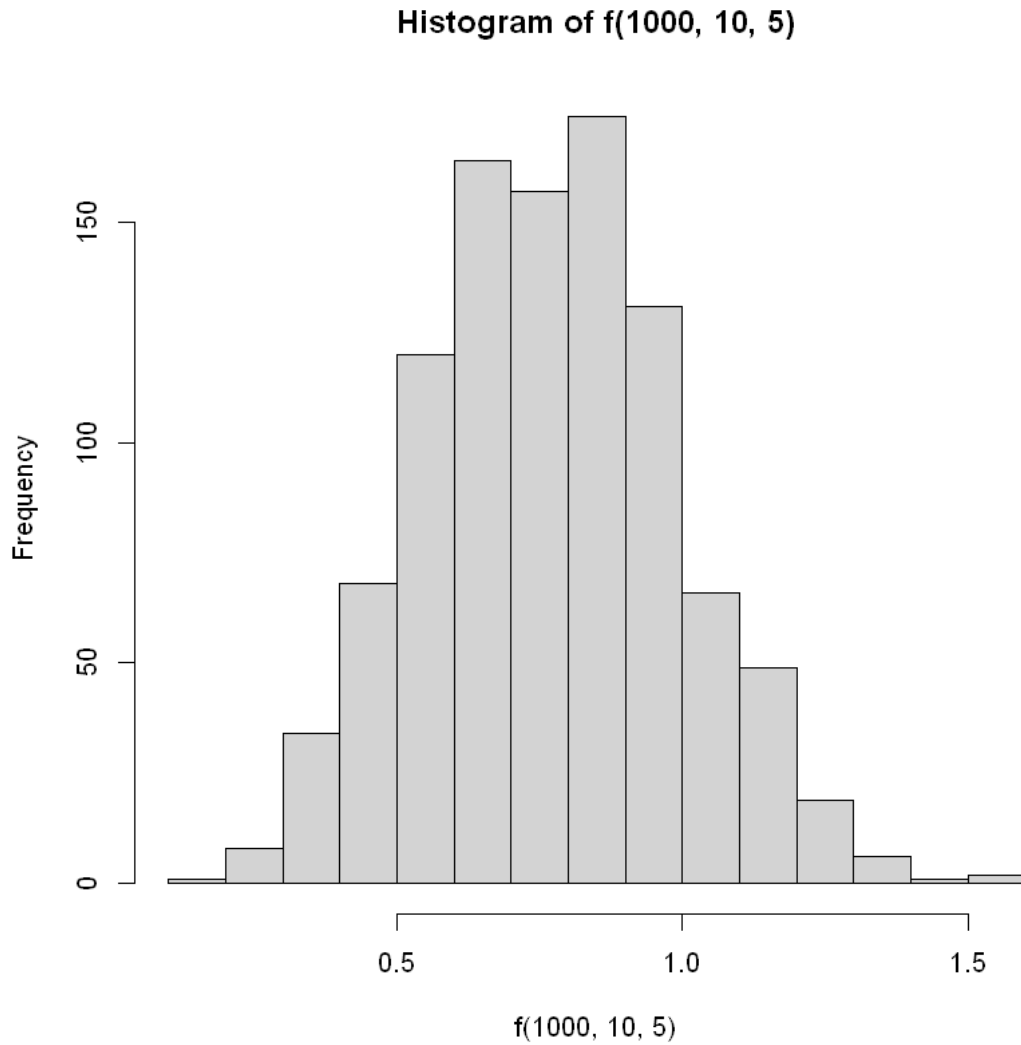
```
[29]: options(scipen=99999999)
```

```
[30]: dnorm(10)
```

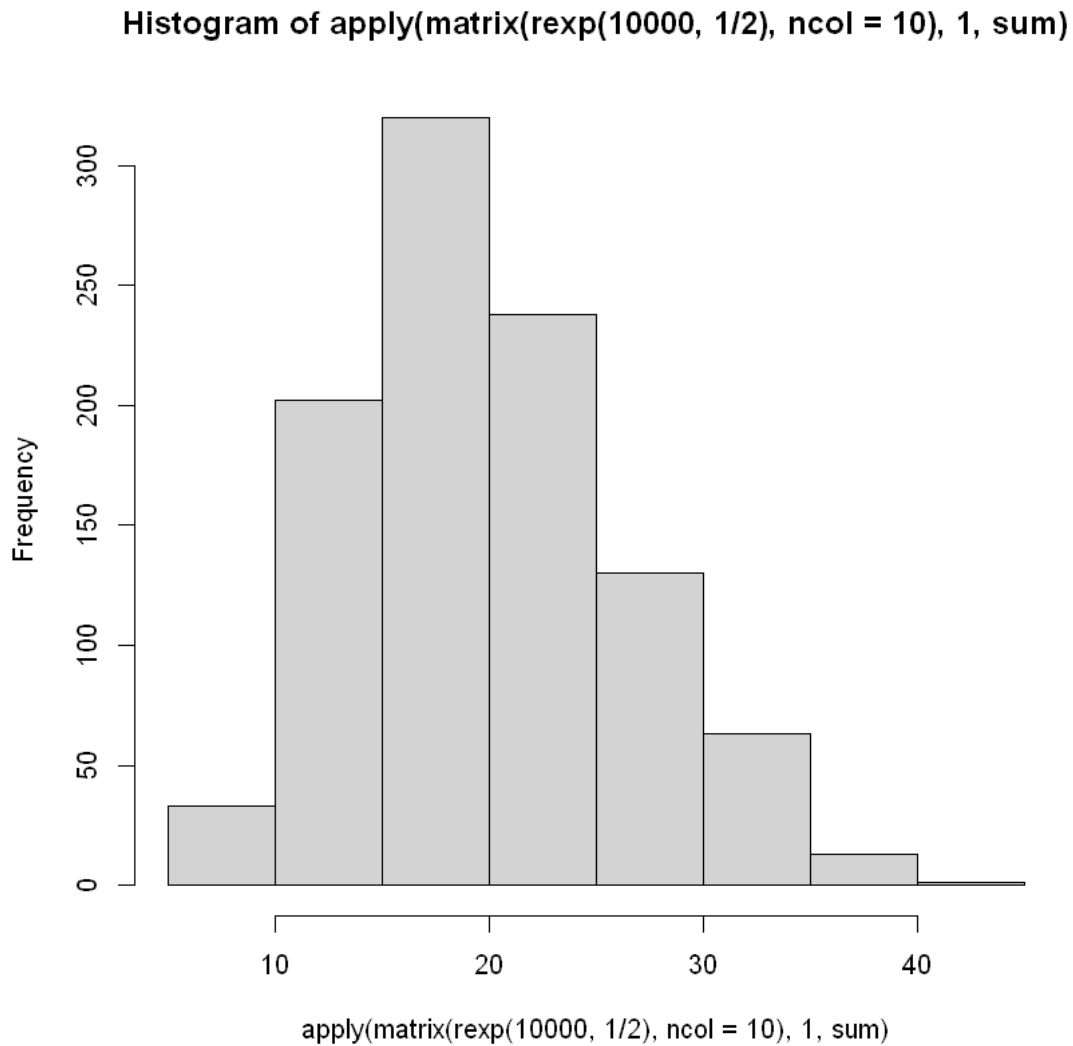
```
0.000000000000000000000000769459862670642
```

```
[48]: f <- function(n, b, m){
  g <- \(x) -m/b * log(1-(m*x)/b)
  x <- matrix(g(runif(n*m)), ncol=m, nrow=n, byrow=1)
  return(apply(x, 1, sum))
}
```

```
[50]: hist(f(1000, 10, 5))
```



```
[40]: hist(apply(matrix(rexp(10000, 1/2), ncol=10), 1, sum))
```



6 Monte Carlo Methods In Statistical Inference

6.1 Statistic estimation using Monte Carlo

- Be sure of the formula you are trying to apply.

Example $x \sim \text{Gamma}(\alpha = 4, \beta = 7)$, take sample size as $n = 100$, $M = 10000$, find **MSE** of both α and β

$$x \sim \text{Gam}(\alpha, \beta) \Rightarrow E(x) = \alpha\beta \Rightarrow \text{Var}(x) = \alpha\beta^2$$

using these

$$E(x) = \alpha\beta, \text{Var}(x) = \alpha\beta^2 \Rightarrow \alpha = \frac{E^2(x)}{\text{Var}(x)}, \beta = \frac{\text{Var}(x)}{E(x)}$$

and for the MSE and Bias

$$MSE = \text{Var}(\hat{\theta}) + \text{Bias}^2(\theta) \text{Bias}(\theta) = \frac{\sum \hat{\theta}_i}{N} - \theta$$

```
[15]: bias_mse <- matrix(0,ncol = 2 , nrow = 2,
                        dimnames = list(c("Bias","MSE"),
                                         c("alpha","beta"))) #for result

MSE <- \(n,alpha,beta) {
  M <- 10^4 # 10000 replicates
  alpha.hat <- numeric(M)
  beta.hat <- numeric(M)
  for(i in 1:M){
    x <- rgamma(n,alpha,beta) #generate n gamma distributed random numbers

    beta.hat[i] <- mean(x)/var(x)
    alpha.hat[i] <- (mean(x))^2 / var(x)
  }

  #bias estimation
  bias_mse[1,1] <- mean(alpha.hat) - alpha
  bias_mse[1,2] <- mean(beta.hat) - beta

  #MSE estimation
  bias_mse[2,1] <- var(alpha.hat) + (mean(alpha.hat) - alpha)^2
  bias_mse[2,2] <- var(beta.hat) + (mean(beta.hat) - beta)^2

  print(bias_mse)
}

MSE(n=100, alpha=4, beta=7)
```

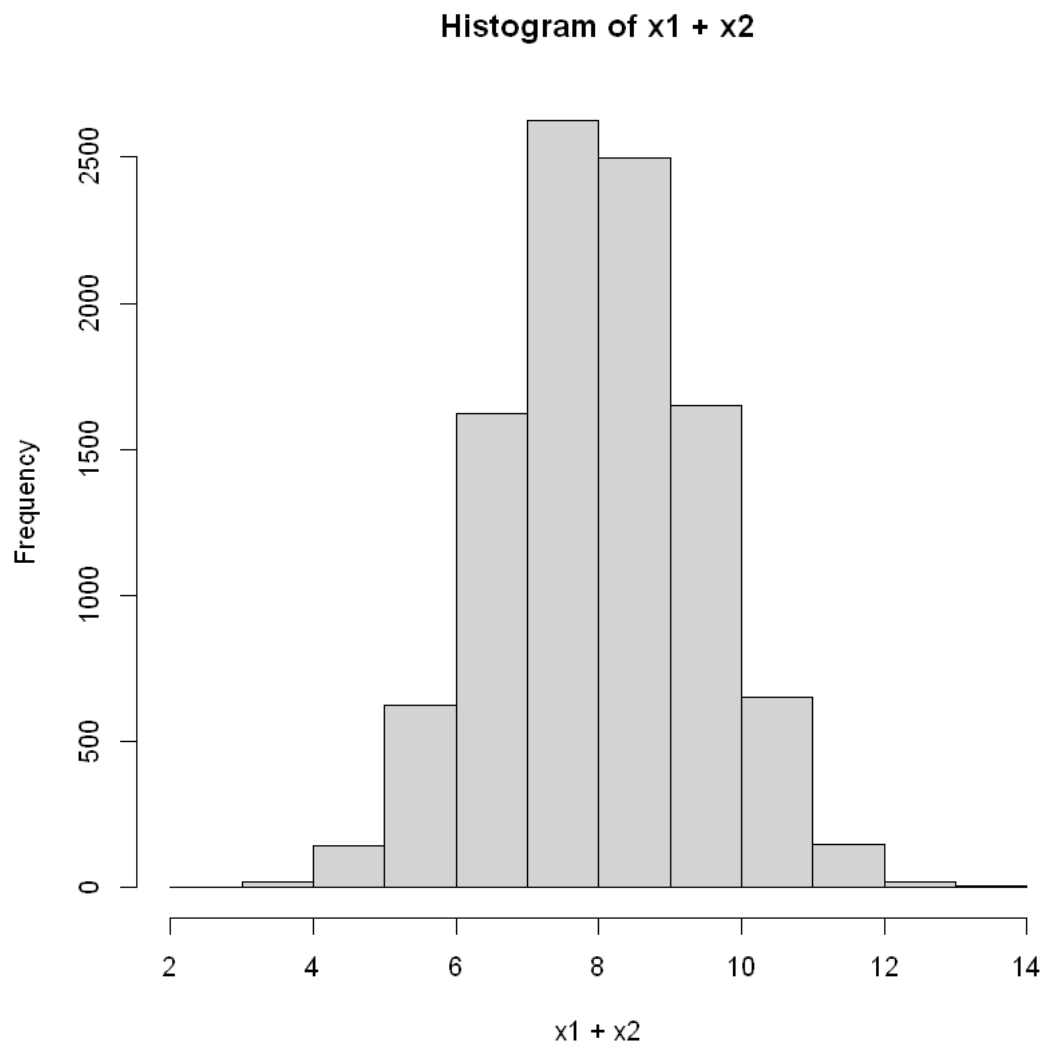
	alpha	beta
Bias	0.1188180	0.2301708
MSE	0.4419034	1.5046290

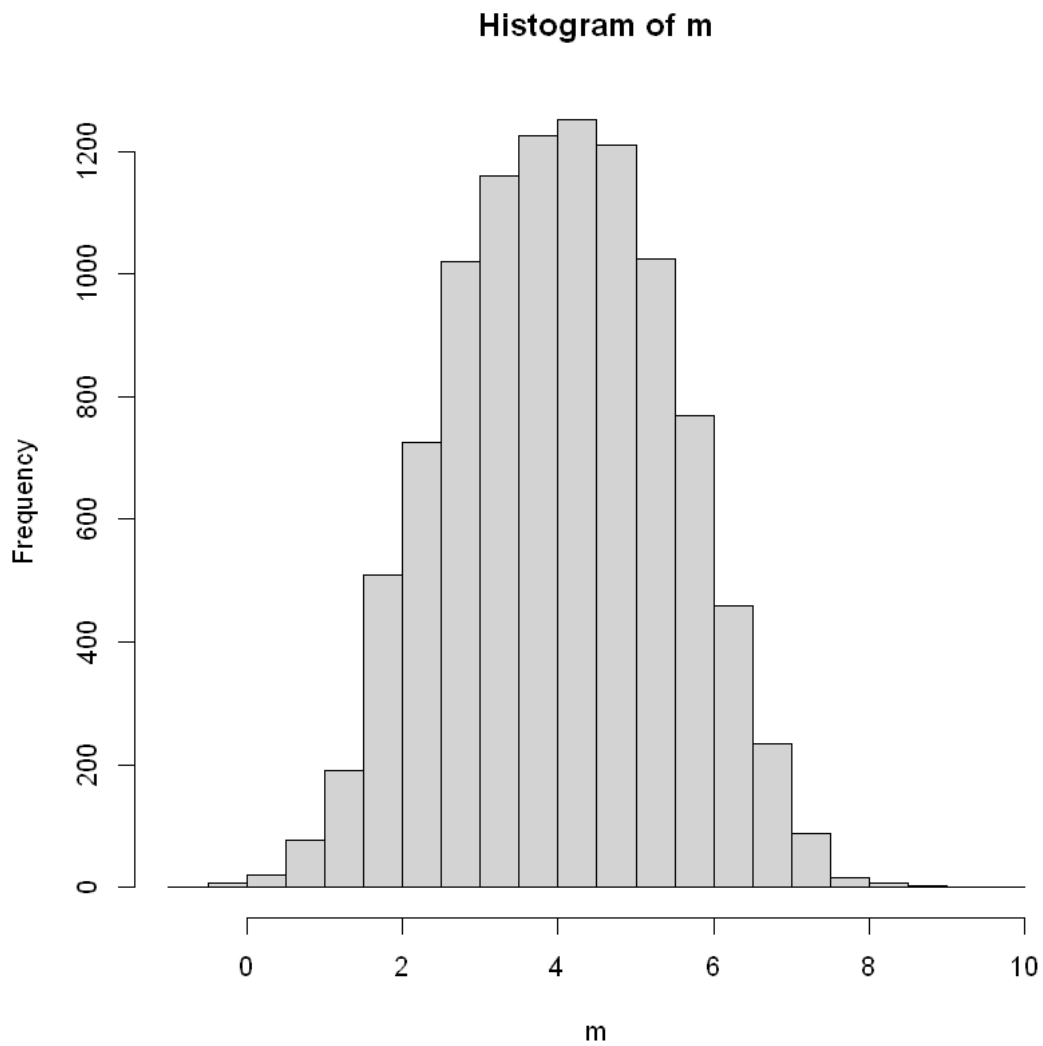
7 OLD FINAL

```
[38]: # convolution
# x1 + x2
x1 <- rnorm(10000, 3, 1)
x2 <- rnorm(10000, 5, 1)
```

```
hist(x1 + x2)

# mixture
m <- c(sample(x1, length(x1)/2), sample(x2, length(x2)/2))
hist(m, breaks=20)
```





```
[39]: density(m)
```

Call:

```
density.default(x = m)
```

Data: m (10000 obs.); Bandwidth 'bw' = 0.2001

x	y
Min. : -1.173	Min. : 2.280e-06
1st Qu.: 1.685	1st Qu.: 1.385e-03
Median : 4.543	Median : 3.442e-02
Mean : 4.543	Mean : 8.739e-02

```
3rd Qu.: 7.401    3rd Qu.:1.878e-01
Max.    :10.259   Max.    :2.496e-01
```

```
[45]: M <- 10^4
      asuc <- 0
      bsuc <- 0

      for(i in 1:M){
        a <- runif(100, 0, 2)
        b <- rexp(100, 1)
        if(t.test(a, alternative = "two.sided", mu=1)$p.value <= 0.05)(asuc =
        ↪asuc+1)
        if(t.test(b, alternative = "two.sided", mu=1)$p.value <= 0.05)(bsuc =
        ↪bsuc+1)
      }

      asuc/M
      bsuc/M
```

0.0525

0.0563