The Axioms of Subjective Probability

Peter C. Fishburn

Abstract. This survey recounts contributions to the axiomatic foundations of subjective probability from the pioneering era of Ramsey, de Finetti, Savage, and Koopman to the mid-1980's. It is designed to be accessible to readers who have little prior acquaintance with axiomatics. At the same time, it provides a fairly complete picture of the present state of the measurement-theoretic foundations of subjective probability.

Key words and phrases: Axioms for comparative probability, numerical representations, additivity, countable additivity.

The probability of any event is the ratio between the value at which an expectation depending on the happening of the event ought to be computed, and the value of the thing expected upon its happening. (Bayes, 1763, page 376)

We are driven therefore to the second supposition that the degree of a belief is a causal property of it, which we can express vaguely as the extent to which we are prepared to act on it. (Ramsey, 1931, page 170)

... the degree of probability attributed by an individual to a given event is revealed by the conditions under which he would be disposed to bet on that event. (de Finetti, 1937; from Kyburg and Smokler, 1964, page 101)

The intuitive thesis in probability holds that ... probability derives directly from the intuition, and is prior to objective experience ... (Koopman, 1940a, page 269)

Personalistic views hold that probability measures the confidence that a particular individual has in the truth of a particular proposition, for example, the proposition that it will rain tomorrow. (Savage, 1954, page 3)

1. INTRODUCTION

The theory of subjective probability attempts to make precise the connection between coherent dispositions toward uncertainty and quantitative probability as axiomatized by Kolmogorov (1933) and others. It accommodates the classical interpretations of probability in Bayes (1763) and Laplace (1812), the intui-

Peter C. Fishburn is a member of the Mathematical Studies Department in the Mathematical Sciences Research Center, AT&T Bell Laboratories, 600 Mountain Avenue, Murray Hill, New Jersey 07974. tive views of Koopman (1940a, 1940b) and Good (1950), and the decision-oriented approach of Ramsey (1931), de Finetti (1931, 1937), and Savage (1954). The aim of this essay is to recount the axiomatic development of subjective or personal (Savage) probability from 1926, when Frank P. Ramsey wrote Truth and Probability, to the present. My hope is that this will not only provide a useful current perspective on subjective probability per se but that it will also promote appreciation of a vital part of the Bayesian approach to statistical decision theory pioneered by Good (1950) and Savage (1954) and further developed by Schlaifer (1959), Raiffa and Schlaifer (1961), DeGroot (1970), and Hartigan (1983) among others. A brief but very informative introduction to problems and perspectives of the Bayesian approach is given by Sudderth (1985).

The axioms of subjective probability refer to assumed properties of a binary relation is more probable than, or its nonstrict companion is at least as probable as, on a set of propositions or events. This relation, often referred to as a qualitative or comparative probability relation, can be taken either as an undefined primitive (intuitive views) or as a relation derived from a preference relation (decision-oriented approach). In the latter case, to say that you regard rain as more probable tomorrow than shine, or that you believe the pound sterling is more likely to fall than rise against the dollar next year means roughly that you would rather bet on the first-named event for a valuable prize that you will receive if your chosen event obtains.

Comparative probability axioms are usually thought of as criteria of consistency and coherence for a person's attitudes toward uncertainty. The prevailing view has been that these criteria do not purport to describe actual behavioral attitudes so much as they characterize the partial beliefs of a rational, idealized individual. Lately, however, some traditional axioms

have been weakened to yield theories that accommodate observed departures from those axioms, and the lines between the ideal and the actual have become blurred. I shall say more about this later.

Except when noted otherwise, I adopt the strict comparative probability relation > as basic. Read A > B as "event A is (regarded by the individual as) more probable than event B." When > is basic, its symmetric complement \sim and nonstrict companion \geq are defined by

$$A \sim B$$
 if neither $A > B$ nor $B > A$,
 $A > B$ if either $A > B$ or $A \sim B$.

We usually view ~ as a comparative equiprobability relation in either a precise sense or an approximate "no significant difference" sense.

On the other hand, when \geq ("is at least as probable as") is taken as basic, then \sim and \succ are defined from it by

$$A \sim B$$
 if $A \succeq B$ and $B \succeq A$,
 $A \succ B$ if $A \succeq B$ and not $(B \succeq A)$.

This has the advantage of identifying A and B as noncomparable (Keynes, 1921) when neither $A \geq B$ nor $B \geq A$. In the \succ basic case, noncomparability is embedded in \sim and there is no obvious way to distinguish between it and comparative equiprobability when such a distinction is desired.

It is easily checked that, when \succ is basic, the definitions in the preceding paragraphs agree with each other if and only if \succ is asymmetric, i.e., whenever $A \succ B$ then not $(B \succ A)$. Alternatively, if \succsim is basic, the definitions agree if and only if \succsim is complete, i.e., $A \succsim B$ or $B \succsim A$ for every pair of events. We shall generally assume that \succ is asymmetric when it is taken as basic, but will consider cases where \succsim is not complete when it is taken as basic (see Section 3).

The set on which \succ or \succeq is defined is assumed to be a Boolean algebra $\mathscr A$ of subsets A, B, \cdots of a universal set S. We refer to each A in $\mathscr A$ as an event. The empty event is $\mathscr O$, the universal event is S, and $\mathscr O \subseteq A \subseteq S$ for every A in $\mathscr A$. We recall that $\mathscr A$ is a Boolean algebra if $S \in \mathscr A$, $A \in \mathscr A \Rightarrow S \setminus A \in \mathscr A$, and A, $B \in \mathscr A \Rightarrow A \cup B \in \mathscr A$. The complement $S \setminus A$ of A in S will also be written as A^c .

A probability measure on $\mathscr A$ is a real valued function P on $\mathscr A$ such that P(S)=1 and, for all A and B in $\mathscr A$, $P(A)\geq 0$, and $A\cap B=\varnothing \Rightarrow P(A\cup B)=P(A)+P(B)$. This last property implies that P is finitely additive: if A_1, \dots, A_n are mutually disjoint events in $\mathscr A$, then $P(\cup_j A_j)=\sum_j P(A_j)$. Countable additivity and σ -algebras, which are taken for granted in some standard works on probability (Loève, 1960; Feller, 1966), will be discussed later.

We say that P on \mathscr{A}

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partially agrees with \succ if A \succ B \Rightarrow P(A) \gt P(B), almost agrees with \succeq if A \succeq B \Rightarrow P(A) \succeq P(B), agrees with \succ (or \succeq) if A \succ B \Leftrightarrow P(A) \gt P(B),
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for all A and B in \mathscr{A} . Partial agreement requires asymmetry for \succ , but almost agreement does not require completeness for \succeq , so \succeq is the appropriate basis for consideration of almost agreement.

The most demanding axiomatizations of subjective probability specify conditions for \succ on $\mathscr A$ that are necessary and sufficient, or perhaps only sufficient, for the existence of a probability measure that agrees with \succ . Less demanding theories seek only partial agreement or almost agreement. We also mention cases intermediate between partial and full agreement and note others in which additivity of P must be replaced by a weaker concept to obtain a suitable numerical representation.

What axioms for > or \geq on $\mathscr A$ are so obvious and uncontroversial as to occasion no serious criticism? A few that might qualify for this distinction are

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asymmetry: If A > B then not (B > A);
nontriviality: S > \emptyset;
nonnegativity: A \geq \emptyset;
monotonicity: If A \supseteq B then A \geq B;
inclusion monotonicity:
If (A \supseteq B, B > C) or (A > B, B \supseteq C)
then A > C.
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One might also nominate

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transitivity: If A > B and B > C then A > C;
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additivity: If
$$A \cap C = \emptyset = B \cap C$$

then $A > B \Leftrightarrow A \cup C > B \cup C$;

complementarity: If A > B then not $(A^c > B^c)$,

but caution is advised here since these assumptions have not gone unchallenged. Examples appear in Section 3.

In any event, axioms like these form the foundations of theories of subjective probability and of our remarks about those theories. We begin with agreement when the event set is finite, then consider weaker finite representations, finitely and countably additive representations with infinite algebras, and conclude with comments on conditional probability.

Excellent sources for articles of historical interest for subjective probability are Kyburg and Smokler (1964) and Savage, Hacking, and Shimony (1967). Technical surveys are included in Luce and Suppes (1965), Krantz et al. (1971), Fine (1973), and Roberts (1979). Other interpretations of probability are discussed in Savage (1954), Good (1959), Fishburn (1964), Fine (1973), de Finetti (1978), and Walley and Fine (1979).

2. FINITE AGREEMENT

This section and the next assume that \mathscr{A} is finite. With no loss in generality, let \mathscr{A} be the family of all subsets of $S = \{1, 2, \dots, n\}$. Each i in S is a *state* and S is the set of states or the set of states of the world (Savage, 1954).

We take > as basic in the present section. For convenience, let

$$p_i = P(\{i\}), \qquad i = 1, \dots, n,$$

so that P agrees with \succ if, for all A and B in \mathscr{A} ,

(1)
$$A > B \Leftrightarrow \sum_{i \in A} p_i > \sum_{i \in B} p_i$$
.

This requires \succ to be asymmetric and transitive and, since $A \sim B \Leftrightarrow P(A) = P(B)$, it also requires \sim to be transitive. We refer to \succ as a *weak order* when it has these properties.

What else besides weak order is needed for agreement? Some time ago de Finetti (1931) noted that (1) also entails nontriviality $(S > \emptyset)$, nonnegativity $(A \geq \emptyset)$ for each (A), and additivity $(A \cap C = \emptyset) = (A \cap C) = (A \cap C)$. The question of whether these axioms are sufficient for agreement remained open until it was settled in the negative by Kraft, Pratt, and Seidenberg (1959). They constructed an example with $(A \cap C) = (A \cap C)$ that satisfies de Finetti's basic axioms and includes the comparisons

$$\{4\} > \{1, 3\},$$

 $\{2, 3\} > \{1, 4\},$
 $\{1, 5\} > \{3, 4\},$
 $\{1, 3, 4\} > \{2, 5\}.$

If (1) holds then

$$p_4 > p_1 + p_3,$$
 $p_2 + p_3 > p_1 + p_4,$
 $p_1 + p_5 > p_3 + p_4,$
 $p_1 + p_3 + p_4 > p_2 + p_5.$

But these are inconsistent since addition and cancellation leaves us with 0 > 0.

What Kraft, Pratt, and Seidenberg discovered is that a much stronger additivity condition is needed for (1). To motivate it, let $(A_1, \dots, A_m) =_0 (B_1, \dots, B_m)$ mean that the A_j and B_j are in $\mathscr A$ and, for each

 $1 \le i \le n$, the number of A_j that contain i equals the number of B_j that contain i. In other words, the sums of the indicator functions over the two event sequences are identical. Then, for any real numbers p_1, \dots, p_n ,

$$(A_1, \dots, A_m) =_0 (B_1, \dots, B_m)$$

$$\Rightarrow \sum_{j=1}^m \sum_{i \in A_j} p_i = \sum_{j=1}^m \sum_{i \in B_i} p_i.$$

Consequently, if (1) holds and $(A_1, \dots, A_m) =_0 (B_1, \dots, B_m)$, then we cannot have $A_j \geq B_j$ for every j along with $A_j > B_j$ for at least one j. In the preceding example,

$$(A_1, A_2, A_3, A_4) = (\{4\}, \{2, 3\}, \{1, 5\}, \{1, 3, 4\}),$$

 $(B_1, B_2, B_3, B_4), = (\{1, 3\}, \{1, 4\}, \{3, 4\}, \{2, 5\}).$

We shall say that \succ on \mathscr{A} is strongly additive if, for all $m \ge 2$ and all A_i and B_i ,

$$[(A_1, \dots, A_m) =_0 (B_1, \dots, B_m), A_j \gtrsim B_j$$

$$(2) \qquad \text{for each } j < m]$$

$$\Rightarrow \text{not } (A_m > B_m).$$

As just noted, this is necessary for (1). Its strength is suggested by the fact that it implies weak order as well as additivity. For example,

$$(A, B) =_0 (B, A)$$
, so $A > B \Rightarrow \text{not} (B > A)$;
 $(A, B, C) =_0 (B, C, A)$,
so $(A \ge B, B \ge C) \Rightarrow A \ge C$.

But strong additivity implies much more. Namely (Fishburn, 1970, Chapter 4), \succ on $\mathscr A$ is strongly additive if and only if there are real numbers p_1, \dots, p_n that satisfy (1). In the special case of subjective probability with $p_i \ge 0$ and $\sum p_i = 1$, it is enough to assume that \succ is nontrivial and nonnegative as well as strongly additive. Proofs that these or equivalent axioms are necessary and sufficient for agreement appear in Kraft, Pratt, and Seidenberg (1959), Scott (1964), and Krantz et al. (1971).

Those proofs, and ones for theorems in the next section that use conditions similar to strong additivity, are based on solution theory for systems of linear inequalities (Kuhn and Tucker, 1956). Relevant theorems are referred to Farkas's lemma, the theorem of the alternative, Fan's theorem, and separating-hyperplane lemmas. The basic algebraic results of linear solution theory are also used by Heath and Sudderth (1972) and Buehler (1976) among others to expand on de Finetti's ideas on coherent systems of bets.

A crude but instructive comparison between betting and strong additivity can be made as follows. Suppose

strong additivity fails with $(A_1, \dots, A_m) =_0 (B_1, \dots, B_m)$, $A_j \geq B_j$ for all j < m, and $A_m > B_m$. Because of $A_m > B_m$, an individual with these comparisons would presumably be willing to pay some positive amount to engage in the following game. For each A_j that obtains he receives \$1, and for each B_j that obtains he pays out \$1. (Event A obtains if the true state is in A.) However, this would be foolish, for regardless of which state obtains, $(A_1, \dots, A_m) =_0 (B_1, \dots, B_m)$ ensures that his net take, exclusive of stake, will be precisely zero. I shall say more about decision models shortly.

Assuming agreement as in (1) with nonnegative p_i that sum to 1, it is trivial to note that the p_i are unique up to transformations that preserve nonnegativity, the sum of 1, and the weak order of partial sums induced by \succ . Of more interest is the question of when the p_i in (1) are unique. A partial answer, sufficient for uniqueness (Luce, 1967; Krantz et al., 1971) is that whenever events A, B, C, and D are such that $A \cap B = \emptyset$, $A \succ B$, and $B \succeq D$, then there are events C', D', and E such that

(3)
$$E \sim A \cup B,$$

$$C' \cap D' = \emptyset,$$

$$C' \cup D' \subseteq E,$$

$$C' \sim C, \quad \text{and} \quad D' \sim D.$$

This is a difficult condition to unravel and it is not necessary even in simple cases. An example not covered by the condition is, for n = 5,

$$\{1, 2\} \sim \{4\},\$$
 $\{1, 4\} \sim \{2, 3\},\$
 $\{2, 4\} \sim \{5\},\$
 $\{2, 5\} \sim \{1, 3, 4\}.$

Under (1) these yield $p_1 + p_2 = p_4$, $p_1 + p_4 = p_2 + p_3$, $p_2 + p_4 = p_5$, and $p_2 + p_5 = p_1 + p_3 + p_4$. The unique solution with nonnegative p_i that sum to 1 is

$$(p_1, p_2, p_3, p_4, p_5) = (2/22, 3/22, 4/22, 5/22, 8/22).$$

In general, given $p_i > 0$ for all i, the p_i are unique if and only if there are n-1 event pairs (A_j, B_j) with $A_j \cap B_j = \emptyset$ and $A_j \sim B_j$ such that the corresponding n-1 linear equations

$$\sum_{i \in A_j} p_i - \sum_{i \in B_j} p_i = 0, \quad j = 1, \dots, n - 1,$$

are linearly independent.

Other avenues to uniqueness are available in the finite states setting if \mathscr{A} is embedded in or extended to a richer structure and something more than (1) is required of P that will ensure its uniqueness. For example, Suppes and Zanotti (1976) and Luce and

Narens (1978) extend $\mathscr A$ to infinite structures with corresponding extensions of \succ such that the representation required of the probability measure for the extended system implies that it is unique. If P for (1) is the restriction to $\mathscr A$ of this measure, then P is unique, but only because of the conditions imposed on its parent. In particular, the linear independence condition of the preceding paragraph need not hold for \sim on $\mathscr A$. Extendability is treated in detail in Kaplan and Fine (1977).

A more traditional route to uniqueness arises when A is embedded in a decision structure and extraneous scaling probabilities (Anscombe and Aumann, 1963; Fishburn, 1967, 1970; Myerson, 1979) or canonical lotteries (Pratt, Raiffa, and Schlaifer, 1964) are used to derive a subjective expected utility model in which the state probabilities are unique. One might, for example, use probabilities associated with events for a random device to construct lotteries on a set of monetary outcomes. We then consider an individual's preference relation on the set $\{(l_1, \dots, l_n): \text{ each } l_i \text{ is a } \}$ lottery}, where (l_1, \dots, l_n) is a "lottery act" that selects lottery l_i if state i obtains or is the "true state." The selected lottery is then played to determine the final outcome. If the individual's preference relation on the set of lottery acts satisfies the expected utility axioms of von Neumann and Morgenstern (1944), then for each i in S there is a real valued utility function u_i on the outcome set M such that, for all pairs of lottery

$$(l_1, \dots, l_n)$$
 is preferred to (l'_1, \dots, l'_n)
 $\Leftrightarrow \sum_{i=1}^n \sum_{a \in M} l_i(a)u_i(a) > \sum_{i=1}^n \sum_{a \in M} l'_i(a)u_i(a),$

where l(a) is the probability that lottery l yields outcome a.

Subjective probabilities emerge from this utility representation (Fishburn, 1967) if we adopt an independence axiom which says that preferences over lotteries within each nonnull state are essentially identical. [State i is null if (l_1, \dots, l_n) is indifferent to (l'_1, \dots, l'_n) whenever $l_i = l'_j$ for all $j \neq i$.] We can then replace u_i by $p_i u$ under suitable rescaling of the u_i to obtain

$$(l_1, \dots, l_n)$$
 is preferred to (l'_1, \dots, l'_n)
 $\Leftrightarrow \sum_{i=1}^n p_i \sum_{a \in M} l_i(a)u(a) > \sum_{i=1}^n p_i \sum_{a \in M} l'_i(a)u(a),$

with the p_i unique and $p_i = 0$ if and only if state i is null.

Additional details of this and other lottery-based theories are given in Fishburn (1981, Section 8). Other decision-oriented theories that derive P on finite or

infinite \mathscr{A} as part of the utility representation for preferences are described there also.

3. OTHER FINITE MODELS

Finite models weaker than (1) arise when we relax strong additivity, including perhaps its implication of weak order. Plausible reasons for considering weaker models will be illustrated by examples before we get into specific representations. I omit asymmetry, nontriviality, and monotonicity from the examples.

3.1 Noncomparability (Incompleteness of ≿)

- A = Mexico City's population will exceed 20,000,000 by 1994;
- B = The first card drawn from this old and probably incomplete bridge deck will be a heart.

As Keynes (1921) might have argued, A and B are sufficiently disparate to discourage a reasonable person from judging either $A \geq B$ or $B \geq A$.

3.2 Nontransitive ~

For a slightly bent coin,

A = The next 101 flips will give at least 40 heads;

B =The next 100 flips will give at least 40 heads;

C =The next 1000 flips will give at least 460 heads.

The judgments $A \sim C$, $C \sim B$, and A > B do not seem unreasonable.

3.3 Nontransitive > (May, 1954; Tversky, 1969; Fishburn, 1983a)

Cyclic patterns could arise from comparisons between multidimensional events. Sue will meet Mike X. Smith at a party next Saturday. She knows nothing about him except that he has just published a hot new book on winning strategies in real estate. Her judgments on separate attributes for Mr. Smith include

Height: 6'-0'' > 6'-1'' > 6'-2'',

Age: 40 > 50 > 60,

Hair color: brunette > red > blonde.

Three composites for Mr. Smith are

 $A = 6' \cdot 0'' 60$ -year-old redhead;

B = 6'-1'' 40-year-old blonde;

C = 6'-2'' 50-year-old brunette.

Sue considers one of these more believable than another if the first is more probable on two of the three attributes. Hence A > B, B > C, and C > A.

3.4 Nonadditivity (Ellsberg, 1961)

An urn contains 90 colored balls, 30 of which are ecru. The other 60 are red and navy in unknown

proportion. One ball will be chosen at random from the urn.

E =chosen ball is ecru;

R =chosen ball is red;

N =chosen ball is navy.

An individual can earn a valuable prize by guessing the color of the chosen ball. Two comparisons are made: E versus R, and $E \cup N$ versus $R \cup N$. Ellsberg's experiments indicate that $E \succ R$ and $R \cup N \succ E \cup N$ are likely. In each case the "more probable" event has greater specificity: exactly 30 balls correspond to E, and exactly 60 correspond to $R \cup N$. The noted judgments clearly violate additivity. They also violate complementarity.

Two relatives of $=_0$ in the preceding section will be used to express necessary and sufficient conditions for several weaker models. First,

$$(A_1, \dots, A_m) \geq_0 (B_1, \dots, B_m)$$

means that, for each $1 \le i \le n$, the number of A_j that contain i equals or exceeds the number of B_j that contain i. Second,

$$(A_1, \dots, A_m) >_0 (B_1, \dots, B_m)$$

means that for every i the number of A_j that contain i exceeds the number of B_j that contain i. In both cases, the A_i and B_i are understood to be in \mathscr{A} .

We first consider almost agreement with \gtrsim as basic. In this case (Kraft, Pratt, and Seidenberg, 1959) there is a probability measure P on $\mathscr A$ that almost agrees with \gtrsim if and only if, for all $m \ge 1$ and all A_i and B_i ,

$$[(A_1, \dots, A_m) >_0 (B_1, \dots, B_m), B_j \gtrsim A_j$$

$$(4) \qquad \text{for each } j < m]$$

$$\Rightarrow \text{not} \quad (B_m \gtrsim A_m).$$

This is a very weak condition. Its only implication for m=1 is not $(\emptyset \geq S)$. For m=2, since $(A, A^c) >_0 (\emptyset, \emptyset)$, it implies that we cannot have both $\emptyset \geq A$ and $\emptyset \geq A^c$. Indeed, the weakness of almost agreement is seen by the fact that it does not forbid cyclic \succ . For example, if $S=\{1, 2, 3\}$ with $\{1\} > \{2\} > \{3\} > \{1\}$, almost agreement could hold with $p_1=p_2=p_3=\frac{1}{3}$.

Partial agreement with \succ as basic forbids cyclic \succ and seems somewhat more demanding than almost agreement. In this case (Adams, 1965; Fishburn, 1969) there is a probability measure P on $\mathscr A$ that partially agrees with \succ if and only if, for all $m \ge 1$ and all A_j and B_j ,

[
$$(A_1, \dots, A_m) \ge_0 (B_1, \dots, B_m), B_j > A_j$$

(5) for all $j < m$]
 $\Rightarrow \text{not} \quad (B_m > A_m).$

This entails asymmetry, nonnegativity, monotonicity, and acyclicity, but not nontriviality, transitivity, or additivity. However, it does forbid additivity reversals like E > R and $R \cup N > E \cup N$ in our fourth example.

Stronger versions of partial agreement have been developed to explicitly include discriminatory thresholds in "if and only if" representations. They were motivated by Luce's (1956) semiorder model and Fishburn's (1970, 1985) interval order model in preference theory. We call > an interval order if it is asymmetric and satisfies $(A > B \text{ and } C > D) \Rightarrow (A > D \text{ or } C > B)$, for all A, B, C, and D. If, in addition, $(A > B > C) \Rightarrow$ (A > D or D > C), for all A, B, C, and D, then > is a semiorder. Examples like those given earlier show that these conditions can be violated by reasonable judgments in some circumstances. For instance, if $A \supset B$ and $C \supset D$ with A and B of a substantially different nature than C and D (c.f. A versus B in our first example), then A > B, C > D, $A \sim C \sim B$, and $A \sim D \sim B$ might obtain.

The semiorder partial agreement representation is a constant threshold model with

$$A > B \Leftrightarrow P(A) > P(B) + \delta, \quad 0 \le \delta < 1,$$

and the interval order partial agreement representation is a variable threshold model with

$$A > B \Leftrightarrow P(A) > P(B) + \sigma(B), \quad \sigma \ge 0.$$

In both cases, P is a probability measure on \mathscr{A} . Necessary and sufficient axioms (Fishburn, 1969; Domotor and Stelzer, 1971) for the semiorder representation are asymmetry, nontriviality, inclusion monotonicity, and

$$[(A_1, \dots, A_{2m}) =_0 (B_1, \dots, B_{2m}),$$

$$B_j \gtrsim A_j \text{ for } j = 1, \dots, m, \text{ and}$$

$$A_j > B_j \text{ for } j = m + 1, \dots, 2m - 1]$$

$$\Rightarrow \text{ not } (A_{2m} > B_{2m}).$$

Sufficient axioms (Fishburn, 1969) for the interval order representation are asymmetry, $(A > B \text{ and } C > D) \Rightarrow (A > D \text{ or } C > B)$, inclusion monotonicity (not strictly necessary), and

$$[(A_1, \dots, A_m) =_0 (B_1, \dots, B_m), A_j >^* B_j$$
for all $j < m$]
$$\Rightarrow \text{not} \quad (A_m >^* B_m),$$

where $A >^* B$ means that, for all C, A > C if B > C, and for some C, A > C and not (B > C). The displayed conditions essentially reflect the difference between the constant and variable threshold notions.

The interval order representation bears comparison to models of upper and lower probabilities that arise in Koopman (1940a), Smith (1961), Good (1962),

Dempster (1967, 1968), Suppes (1974), Williams (1976), Walley and Fine (1979), and Kumar (1982) among others. See also Shafer (1976) on the concept of belief functions. One model of this type has

$$A > B \Leftrightarrow P_*(A) > P^*(B),$$

where P_* and P^* are monotonic but not necessarily additive measures on $\mathscr A$ with $P^*(S) = P_*(S) = 1$, $P^*(\varnothing) = P_*(\varnothing) = 0$, and $P^*(A) \ge P_*(A)$ for each A. If we let $\sigma(A) = P^*(A) - P_*(A)$, then

$$A > B \Leftrightarrow P_{\star}(A) > P_{\star}(B) + \sigma(B),$$

which is similar to the interval order model. Axioms for this case consist of the interval order conditions plus suitable nonnegativity and monotonicity assumptions. It requires a special additivity condition only if there is to be a probability measure P for which $P_* \leq P \leq P^*$.

Davidson and Suppes (1956) and Schmeidler (1984) axiomatize utility models in which P is monotonic but not necessarily additive. Davidson and Suppes adopt Ramsey's (1931) measurement approach. They first use an event with subjective probability $\frac{1}{2}$ to scale outcome utilities and then use these to scale probabilities of other events. Their outcome utilities are evenly spaced and P(A) is defined as

$$P(A) = \frac{u(d^*) - u(d)}{u(d^*) - u(d) + u(c) - u(c^*)}$$

when $u(c) \neq u(c^*)$, $u(d) \neq u(d^*)$, and the act that yields c if A and d if A^c is indifferent to the act that yields c^* if A and d^* if A^c . This gives $P(A) + P(A^c) = 1$, and their P is unique if $n \geq 5$.

Schmeidler (1984) weakens the lottery act axioms for the model at the end of the preceding section to obtain a modified expected utility representation in which P is monotonic and unique. Because P need not be additive, his definition of expectation (integration) is special. Schmeidler's model is designed to accommodate Ellsberg-type violations of additivity.

Another nonadditive model, axiomatized in both the intuitive and decision modes (Fishburn, 1983a, 1983b), allows intransitivities. Unlike preceding representations, it uses a two-argument function ρ on $\mathscr{A} \times \mathscr{A}$ with

$$A > B \Leftrightarrow \rho(A, B) > 0$$

along with $\rho(\Omega, \emptyset) = 1$, $\rho(B, A) = -\rho(A, B)$, $A \supseteq B \Rightarrow \rho(A, B) \ge 0$, and

$$A \cap B = \emptyset$$

$$\Rightarrow \rho(A \cup B, C) + \rho(\emptyset, C) = \rho(A, C) + \rho(B, C).$$

The last property is a nonseparable generalization of additivity. Necessary and sufficient conditions are nontriviality, monotonicity, and a generalization of strong additivity.

4. INFINITE ALGEBRAS

This and the next two sections assume that $\mathscr A$ is an infinite Boolean algebra of subsets of S. The present section introduces terminology and special axioms developed for this case. It also notes a partial agreement axiomatization that does not fit into the next two sections, which discuss agreement and almost agreement when \succ is a weak order. The next section focuses on finite additivity; Section 6 looks at countably additive measures. We take \succ as basic throughout.

The Boolean algebra \mathscr{A} is a σ -algebra if A is in \mathscr{A} whenever A is the union of a countable number of events in \mathscr{A} . A probability measure P is countably additive if

$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i)$$

whenever the A_i are mutually disjoint events whose union is also in \mathscr{A} . Note that \mathscr{A} does not have to be a σ -algebra for P to be countably additive.

Following Savage (1954), \succ is *fine* if, whenever $A \succ \emptyset$, S can be partitioned into events B_1, \dots, B_m with $A \succeq B_j$ for all j. Events A and B are almost equivalent if, for all C and D in \mathscr{A} ,

$$[C \succ \varnothing, D \succ \varnothing, A \cap C = \varnothing = B \cap D]$$

 $\Rightarrow A \cup C \succeq B \text{ and } B \cup D \succeq A;$

and \succ is tight if $A \sim B$ whenever A and B are almost equivalent. Savage (1954, page 37) subsequently strengthened his definition of fine by replacing $A \gtrsim B_j$ by $A \succ B_j$ for all j. We refer to \succ as superfine in the latter case.

Following Villegas (1964), event A is an atom if $A > \emptyset$ and $A > B > \emptyset$ for no event $B \subseteq A$. The algebra $\mathscr A$ is atomless if it has no atom. It should be noted that these definitions depend on the behavior of >.

Three of the earliest axioms for infinite algebras are due to de Finetti (1931, 1937) and Savage (1954). The first of the following is from de Finetti, and also Bernstein (1917) and Koopman (1940a). The other two are Savage's. All three suggest the possibility of partitioning $\mathscr A$ into a very large number of events with very small probabilities.

Partition Axiom 1. For each $n \ge 2$, S can be partitioned into n equally likely (\sim) events.

Partition Axiom 2. For each $n \ge 2$, S can be partitioned into n events such that the union of no r events is more probable than the union of any r + 1.

Partition Axiom 3. For all A and B in \mathscr{A} , if A > B then S can be partitioned into a finite number of events C_1, \dots, C_m such that $A > B \cup C_i$ for all j.

Partitions for Axiom 1 are uniform partitions, and those for Axiom 2 are almost uniform partitions. Given de Finetti's basic axioms (Section 2), Savage notes that Axiom 3 implies that \succ is both superfine and tight and that it implies Axioms 1 and 2.

Partition Axiom 3 is used by Fishburn (1975) to axiomatize partial agreement with a finitely additive measure. Sufficient axioms for partial agreement are inclusion monotonicity, Axiom 5 with $=_0$ in place of \ge_0 , Partition Axiom 3, and the following part of additivity:

$$[A \cap C = \emptyset = B \cap C, A \cup C > B \cup C] \Rightarrow A > B.$$

Partial additivity is defended by the contention that if $A \cup C > B \cup C$ when $(A \cup B) \cap C = \emptyset$, then the difference between the two events will be even more apparent when the common part C is removed. The fourth example in the preceding section questions its general acceptability.

Because of the possibility of atoms, none of the partition axioms is necessary for agreement. We therefore note three more conditions that are necessary for the existence of a probability measure that agrees with >. These are referred to as Archimedean axioms because they lead to real representations with no infinitesimal nonstandard numbers or infinitely large numbers. A few preliminary definitions are needed.

A subset $\mathscr C$ of $\mathscr A$ is order dense in $\mathscr A$ if, whenever $A \succ B$ and neither A nor B is in $\mathscr C$, then $A \succsim C \succsim B$ for some C in $\mathscr C$. A countable (finite or denumerable) sequence of events A_1, A_2, \cdots is a standard sequence relative to event A if for each A_i there are disjoint events B_i and C_i such that $A_1 = B_1, A_{i+1} = B_i \cup C_i$ (when i+1 is present), $B_1 \sim A$, $C_i \sim A$, and $B_i \sim A_i$. The upshot of this is that if P agrees with \succ then $P(A_i) = iP(A)$ for each i.

Archimedean Axiom 1. $\mathscr A$ includes a countable order dense subset.

Archimedean Axiom 2. For each event $A > \emptyset$, every standard sequence relative to A is finite.

Archimedean Axiom 3. For all $m \ge 2$ and all events $A_1, \dots, A_m, B_1, \dots, B_m$: if $A_1 > B_1$ then there is a positive integer N such that, whenever $(k\emptyset, nA_1, A_2, \dots, A_m) =_0 (kS, nB_1, B_2, \dots, B_m)$ with $k \ge 0, n \ge 1$, and $A_j \ge B_j$ for all $j \ge 2$, then k/n > 1/N.

The first Archimedean axiom, which dates from Cantor's (1895) work on transfinite numbers, is necessary and sufficient in conjunction with weak order for the existence of a real valued function f on $\mathscr A$ such that, for all A and B in $\mathscr A$, $A > B \Leftrightarrow f(A) > f(B)$ (Birkhoff, 1967; Fishburn, 1970; Krantz et al., 1971).

Archimedean Axiom 2, from Luce (1967), is one of several standard sequence conditions used in Krantz et al. (1971). To be applicable, it requires structural conditions that allow the construction of such sequences. Neither of the first two axioms relates to additivity of probability, which must enter through another axiom.

Strong additivity is built into Archimedean Axiom 3 via k = 0 and n = 1. In its statement, due to Chateauneuf and Jaffray (1984), kA denotes A repeated k times. Its necessity for agreement follows from the fact that agreement and its hypotheses imply

$$k = n[P(A_1) - P(B_1)] + \sum_{j \ge 2} [P(A_j) - P(B_j)]$$

$$\ge n[P(A_1) - P(B_1)],$$

so $k/n \ge P(A_1) - P(B_1)$. Since $P(A_1) - P(B_1) > 0$ by $A_1 > B_1$, k/n cannot be arbitrarily small without violating agreement.

5. FINITE ADDITIVITY

The basic axioms of de Finetti, i.e., weak order, nontriviality, nonnegativity, and additivity, have become so standard that they are used more or less directly in all infinite $\mathscr A$ axiomatizations for agreement, and in most axiomatizations for almost agreement. Savage (1954) defines \succeq (or \succ) as a qualitative probability when it satisfies de Finetti's axioms, and this name has persisted in more or less this form.

Since all axiomatizations in this section and the next use de Finetti's basic axioms, they will be taken for granted until the final section of the paper. The present section is concerned solely with the existence of finitely additive probability measures that agree, or almost agree, with \succ when $\mathscr A$ is infinite. We consider almost agreement first.

Savage (1954) proves that there exists a P that almost agrees with > when Partition Axiom 2 holds (almost uniform) and $\mathscr A$ is the σ -algebra of all subsets of S. Moreover, P is unique, and if P(A) > 0 and $0 < \lambda < 1$, then $P(B) = \lambda P(A)$ for some $B \subseteq A$. The same conclusions follow if > is superfine and, in this case, P(A) > 0 if $A > \emptyset$ and P(A) = P(B) if and only if A and B are almost equivalent. See also Wakker (1981, Lemmas 3–5).

Related results for almost agreement have been obtained by others when \mathscr{A} is merely assumed to be a Boolean algebra. Niiniluoto (1972) and Wakker (1981) note that there is a unique almost agreeing P when \succ is fine. Narens (1974) and Wakker (1981) show that almost agreement (not necessarily unique) follows from (4).

The earliest axiomatizations for agreement (Bernstein, 1917; de Finetti, 1931, 1937; Koopman, 1940a) used Partition Axiom 1 (uniform partitions). Savage (1954), again with \mathscr{A} as the family of all subsets of S,

obtains a unique agreeing P with Partition Axiom 3. Since the $P(B) = \lambda P(A)$ property holds here, \mathscr{A} must be atomless. A related result is obtained in French (1982) by adjoining an auxiliary experiment (extraneous scaling probabilities) to S. This approach, of enriching a possibly finite S with events generated from a random device, was used earlier by DeGroot (1970), and indeed did not escape notice by Savage (1954, pages 33 and 38).

Other axiomatizations for agreement do not assume that \mathscr{A} is a σ -algebra. Luce (1967), whose axioms apply also to finite S, shows that a unique agreeing P follows from (3) and Archimedean Axiom 2. Roberts (1973) obtains unique agreement from Partition Axiom 2 and Archimedean Axiom 1, after a similar result in Fine (1971). Wakker (1981) notes that, when \succ is fine, unique agreement holds if and only if either \succ is tight or $\mathscr A$ has an atom.

Necessary and sufficient conditions for agreement are discussed by Domotor (1969) and Chateauneuf and Jaffray (1984). Domotor's complex conditions will not be recalled here. Chateauneuf and Jaffray show that if \mathscr{A} is countable, e.g., the family of finite subsets of $\{1, 2, \cdots\}$ and their complements, then some P agrees with \succ if and only if Archimedean Axiom 3 holds. Because their conditions are necessary as well as sufficient, they do not imply uniqueness. Indeed, a price payed for uniqueness is nonnecessity since conditions that guarantee unique agreement cannot be wholly necessary.

Further references on mathematical details of finitely additive measures include Bochner (1939), Sobczyk and Hammer (1944), Yosida and Hewitt (1952), Hewitt and Savage (1955), Dubins (1974), Purves and Sudderth (1976), Bhashara Rao and Bhaskara Rao (1983), and Schervish, Seidenfeld, and Kadane (1984). Contributions of special interest in decision theory include Dubins and Savage (1965), de Finetti (1975), Heath and Sudderth (1978), Lane and Sudderth (1978), and Seidenfeld and Schervish (1983).

6. COUNTABLE ADDITIVITY

We continue to assume de Finetti's four basic axioms.

Villegas (1964) identifies the key assumption on > needed for countable additivity. Modified slightly to accommodate Boolean algebras that are not also σ -algebras, his axiom is

Monotone continuity: For all A, B, A_1, A_2, \cdots in \mathscr{A} , if $A_1 \subseteq A_2 \subseteq \cdots$, $A = \bigcup_i A_i$ and $B \gtrsim A_i$ for all i, then $B \gtrsim A$.

Thus, if the nondecreasing A_i converge to a limit event A, then the judgment that B is at least as

probable as A_i for all i cannot be reversed in the limit by a jump to A > B. Monotone continuity is quite appealing and is necessary for agreement with a countably additive probability measure. In fact, Villegas (1964) proves that if $\mathscr A$ is a σ -algebra and if P is a finitely additive probability measure that agrees with >, then P is countably additive if and only if > is monotonely continuous. Chateauneuf and Jaffray (1984) remark that this remains true even when $\mathscr A$ is not a σ -algebra.

Consequently, for every axiomatization in the preceding section that implies the existence of P that agrees with \succ , such a P is countably additive if and only if \succ is monotonely continuous.

Other axiomatizations have been developed specifically for countable additivity when $\mathscr A$ is a σ -algebra. All assume monotone continuity. Villegas (1964) proves that if $\mathscr A$ is atomless then there is a unique countably additive P that agrees with \succ . DeGroot (1970) and French (1982) obtain the same conclusion with the explicit use of an auxiliary experiment that itself is atomless but allows atoms in the original state set. Similarly, Chuaqui and Malitz (1983) adopt strong additivity to obtain unique agreement whenever \mathscr{A} has a nontrivial atomless subalgebra. They also consider the purely atomic case, where P is a discrete measure. and prove an almost agreement theorem under strong additivity. Finally, Chateauneuf and Jaffray (1984) show that Archimedean Axiom 3 is necessary and sufficient for agreement. Their complete set of axioms for countably additive agreement when \mathscr{A} is a σ -algebra is equivalent to de Finetti's axioms, monotone continuity, and Archimedean Axiom 3. Since these are necessary as well as sufficient, their P is not necessarily unique.

7. CONDITIONAL PROBABILITY

A typical version of Bayes' formula for conditional probability when $S=\Theta\times Z$ and the underlying state or parameter space Θ and sample space Z are finite is

$$P_1(\boldsymbol{\theta} \mid \boldsymbol{z}) = \frac{P_2(\boldsymbol{z} \mid \boldsymbol{\theta}) P_0(\boldsymbol{\theta})}{\sum_{\boldsymbol{\theta}'} P_2(\boldsymbol{z} \mid \boldsymbol{\theta}') P_0(\boldsymbol{\theta}')}.$$

Here $P_1(\cdot \mid z)$ and $P_2(\cdot \mid \theta)$ are conditional probability mass functions, and P_0 is a (marginal) probability mass function on Θ . In the usual subjectivistic interpretation, P_0 is the prior distribution and P_1 is the posterior on Θ given observation z. Each of P_0 , P_1 , and P_2 may be based on a joint distribution over S as described in the next paragraph.

If there is a *unique* probability measure P on \mathscr{A} that agrees or almost agrees with \succ or \succeq , then subjective conditional probabilities can be defined unambiguously in the usual way as

(6)
$$P(A | B) = P(A \cap B)/P(B)$$
 when $P(B) > 0$,

thus putting Bayes' theorem and other mathematical machinery of conditional probability at our disposal. The theoretical foundations of Bayesian decision theory can be understood in this way when all uncertainties, including those associated with experiment's outcomes, are embedded in \mathscr{A} . At least in theory, if not in common practice, priors and other special probabilities can be extracted from the global P measure by the usual operations. For the setting of the preceding paragraph, $P_0(\theta) = P(\{\theta\} \times Z)$, $P_1(\theta \mid z) = P(\{(\theta, z)\})/P(\Theta \times \{z\})$, and $P_2(z \mid \theta) = P(\{(\theta, z)\})/P(\{\theta\} \times Z)$ when the denominators do not vanish.

A potential problem with (6) arises when P is not unique, as for example in Section 2, since conditional probabilities might not be uniquely ordered. For example, one agreeing P may give $P(A \mid B) > P(C \mid D)$ by (6), while another gives $P(C \mid D) > P(A \mid B)$.

Partly for this reason, and partly because some authors, including Koopman (1940a, 1940b), believe that conditional judgments are basic to subjective probability, there has been interest in axiomatizing conditional probability directly. This is usually done by means of a comparative conditional probability relation \geq_0 on $\mathscr{A} \times \mathscr{A}_0$, where \mathscr{A}_0 is the set of nonnull events in \mathscr{A} and $A \mid B \geq_0 C \mid D$ is interpreted as "A given B is at least as probable as C given D." When \geq_0 is complete and \sim_0 is defined by

$$A \mid B \sim_0 C \mid D$$

if $A \mid B \succeq_0 C \mid D$ and $C \mid D \succeq_0 A \mid B$,

a null event is an A in \mathscr{A} for which $A \mid S \sim_0 \varnothing \mid S$. A typical representation for comparative conditional probability is, for all $A \mid B$ and $C \mid D$ in $\mathscr{A} \times \mathscr{A}_0$,

(7)
$$A \mid B \succeq_0 C \mid D \\ \Leftrightarrow P(A \cap B)/P(B) \ge P(C \cap D)/P(D),$$

where P is a probability measure on \mathscr{A} .

Domotor (1969) includes necessary and sufficient conditions on \succeq_0 for (7) when $\mathscr A$ is finite. His axioms extend the approach of Kraft, Pratt, and Seidenberg (1959) for finite agreement. Suppes and Zanotti (1982) give necessary and sufficient conditions for a model like (7) for general & by using extensions of indicator functions in a manner similar to Suppes and Zanotti (1976). Other sufficient axiomatizations for (7) and related representations appear in Koopman (1940a, 1940b), Aczel (1961, 1966), Luce (1968), and Krantz et al. (1971). The last of these provides a nice description of previous work (pages 221-228). Their main axiomatization is based on Luce (1968) and consists of seven conditions necessary for (7) plus a nonnecessary solvability condition which says that if $A \mid B \geq_0 C \mid D$ then there is a C' in $\mathscr A$ such that $C \cap D \subseteq C'$ and $A \mid B \sim_0 C' \mid D$.

Conditional probabilities also arise directly from the preference-based axiomatizations of Ramsey (1931), Pfanzagl (1967, 1968), Luce and Krantz (1971), Fishburn (1973), and Balch and Fishburn (1974). Reviews of their theories are included in Fishburn (1981).

REFERENCES

- ACZEL, J. (1961). Über die Begründung der Additions- und Multiplikationsformeln von bedingten Wahrscheinlichkeiten. Magyar Tud. Akad. Mat. Kutató Int. Közl. 6 110-122.
- ACZEL, J. (1966). Lectures on Functional Equations and Their Applications. Academic, New York.
- Adams, E. W. (1965). Elements of a theory of inexact measurement. *Philos. Sci.* **32** 205–228.
- Anscombe, F. J. and Aumann, R. J. (1963). A definition of subjective probability. *Ann. Math. Statist.* **34** 199-205.
- Balch, M. and Fishburn, P. C. (1974). Subjective expected utility for conditional primitives. In *Essays on Economic Behavior* under Uncertainty 57-69. North-Holland, Amsterdam.
- BAYES, T. (1763). An essay towards solving a problem in the doctrine of chances. *Philos. Trans. Roy. Soc.* **53** 370-418. Reprinted in *Facsimiles of Two Papers of Bayes* (W. E. Deming, ed.). Department of Agriculture, Washington, D. C., 1940.
- BERNSTEIN, S. N. (1917). On the axiomatic foundation of probability theory (in Russian). Soob. Khar'k. ob-va 15 209-274.
- BHASHARA RAO, K. P. S. and BHASKARA RAO, M. (1983). Theory of Charges. Academic, New York.
- BIRKHOFF, G. (1967). Lattice Theory. 3rd ed. Amer. Math. Soc., Providence, R. I.
- BOCHNER, S. (1939). Additive set functions on groups. *Ann. Math.* **40** 769–799.
- BUEHLER, R. J. (1976). Coherent preferences. Ann. Statist. 4 1051-1064.
- CANTOR, G. (1895). Beiträge zur Begründung der transfinite Mengenlehre. Math. Ann. 46 481-512; 49 (1897) 207-246. (English translation in Contributions to the Founding of the Theory of Transfinite Numbers. Dover, New York, n.d.)
- CHATEAUNEUF, A. and JAFFRAY, J.-Y. (1984). Archimedean qualitative probabilities. J. Math. Psychology 28 191–204.
- CHUAQUI, R. and MALITZ, J. (1983). Preorderings compatible with probability measures. Trans. Amer. Math. Soc. 279 811–824.
- DAVIDSON, D. and SUPPES, P. (1956). A finitistic axiomatization of subjective probability and utility. *Econometrica* **24** 264–275.
- DE FINETTI, B. (1931). Sul significato soggettivo della probabilità. Fund. Math. 17 298–329.
- DE FINETTI, B. (1937). La prévision ses lois logiques, ses sources subjectives. Ann. Inst. H. Poincaré 7 1-68. (Translated by H. E. Kyburg in Kyburg and Smokler (1964).)
- DE FINETTI, B. (1975). Theory of Probability. Wiley, New York.
- DE FINETTI, B. (1978). Probability: interpretations (II). International Encyclopedia of Statistics 744-754. Free Press, New York.
- Degroot, M. H. (1970). Optimal Statistical Decisions. McGraw-Hill, New York.
- DEMPSTER, A. P. (1967). Upper and lower probabilities induced by a multivalued mapping. *Ann. Math. Statist.* **38** 325-339.
- DEMPSTER, A. P. (1968). A generalization of Bayesian inference. J. Roy. Statist. Soc. Ser. B 30 205-247.
- DOMOTOR, Z. (1969). Probabilistic relational structures and their applications. Technical Report 144, Institute for Mathematical Studies in the Social Sciences, Stanford Univ.
- DOMOTOR, Z. and STELZER, J. (1971). Representation of finitely additive semiordered qualitative probability structures. J. $Math.\ Psychology\ 8\ 145-158$.
- DUBINS, L. E. (1974). On Lebesgue-like extensions of finitely additive measures. Ann. Probab. 2 456-463.

- DUBINS, L. E. and SAVAGE, L. J. (1965). How to Gamble if You Must. McGraw-Hill, New York. Dover edition, 1976.
- ELLSBERG, D. (1961). Risk, ambiguity, and the Savage axioms. *Quart. J. Econ.* **75** 643-669.
- FELLER, W. (1966). An Introduction to Probability Theory and Its Applications 2. Wiley, New York.
- FINE, T. (1971). A note on the existence of quantitative probability. Ann. Math. Statist. 42 1182-1186.
- FINE, T. (1973). Theories of Probability. Academic, New York.
- FISHBURN, P. C. (1964). Decision and Value Theory. Wiley, New York.
- FISHBURN, P. C. (1967). Preference-based definitions of subjective probability. *Ann. Math. Statist.* **38** 1605–1617.
- FISHBURN, P. C. (1969). Weak qualitative probability on finite sets. Ann. Math. Statist. 40 2118–2126.
- FISHBURN, P. C. (1970). *Utility Theory for Decision Making*. Wiley, New York. Krieger edition, 1979.
- FISHBURN, P. C. (1973). A mixture-set axiomatization of conditional subjective expected utility. *Econometrica* 41 1-25.
- FISHBURN, P. C. (1975). Weak comparative probability on infinite sets. *Ann. Probab.* **3** 889–893.
- FISHBURN, P. C. (1981). Subjective expected utility: a review of normative theories. *Theory and Decision* 13 139-199.
- FISHBURN, P. C. (1983a). A generalization of comparative probability on finite sets. J. Math. Psychology 27 298-310.
- FISHBURN, P. C. (1983b). Ellsberg revisited: a new look at comparative probability. *Ann. Statist.* 11 1047-1059.
- FISHBURN, P. C. (1985). Interval Orders and Interval Graphs. Wiley, New York.
- FRENCH, S. (1982). On the axiomatisation of subjective probabilities. *Theory and Decision* 14 19-33.
- Good, I. J. (1950). Probability and the Weighing of Evidence. Griffin, London.
- GOOD, I. J. (1959). Kinds of probability. Science 129 443-447.
- Good, I. J. (1962). Subjective probability as the measure of a non-measurable set. Logic, Methodology and Philosophy of Science 319-329. Stanford Univ. Press, Stanford, Calif.
- HARTIGAN, J. A. (1983). Bayes Theory. Springer, New York.
- HEATH, D. C. and SUDDERTH, W. D. (1972). On a theorem of de Finetti, oddsmaking, and game theory. *Ann. Math. Statist.* 43 2072–2077.
- HEATH, D. C. and SUDDERTH, W. D. (1978). On finitely additive priors, coherence, and extended admissibility. Ann. Statist. 6 333-345.
- HEWITT, E. and SAVAGE, L. J. (1955). Symmetric measures on Cartesian products. *Trans. Amer. Math. Soc.* **80** 470-501.
- KAPLAN, M. and FINE, T. (1977). Joint orders in comparative probability. Ann. Probab. 5 161-179.
- KEYNES, J. M. (1921). A Treatise on Probability. Macmillan, New York. Torchbook edition, 1962.
- KOLMOGOROV, A. N. (1933). Grundbegriffe der Wahrscheinlichkeitsrechnung. Springer, Berlin. Chelsea edition (English), 1950.
- KOOPMAN, B. O. (1940a). The axioms and algebra of intuitive probability. *Ann. Math.* 41 269-292.
- KOOPMAN, B. O. (1940b). The bases of probability. *Bull. Amer. Math. Soc.* **46** 763-774. Reprinted in Kyburg and Smokler (1964).
- Kraft, C. H., Pratt, J. W. and Seidenberg, A. (1959). Intuitive probability on finite sets. *Ann. Math. Statist.* **30** 408-419.
- KRANTZ, D. H., LUCE, R. D., SUPPES, P. and TVERSKY A. (1971).
 Foundations of Measurement 1. Academic, New York.
- KUHN, H. W. and TUCKER, A. W., eds. (1956). Linear Inequalities and Related Systems, Ann. Math. Study 38. Princeton Univ. Press, Princeton, N. J.
- KUMAR, A. (1982). Lower probabilities on infinite spaces and instability of stationary sequences. Ph.D. dissertation, Cornell Univ.
- KYBURG, H. E., Jr. and SMOKLER, H. E., eds. (1964). Studies in Subjective Probability. Wiley, New York.

- LANE, D. A. and SUDDERTH, W. D. (1978). Diffuse models for sampling and predictive inference. Ann. Statist. 6 1318– 1336.
- LAPLACE, P. S. (1812). Théorie Analytique des Probabilités. Paris. Reprinted in Oeuvres Complétes 7 1947.
- LOÈVE, M. (1960). Probability Theory. 2nd ed. Van Nostrand, Princeton, N. J.
- LUCE, R. D. (1956). Semiorders and a theory of utility discrimination. *Econometrica* 24 178-191.
- LUCE, R. D. (1967). Sufficient conditions for the existence of a finitely additive probability measure. Ann. Math. Statist. 38 780-786.
- Luce, R. D. (1968). On the numerical representation of qualitative conditional probability. *Ann. Math. Statist.* **39** 481-491.
- LUCE, R. D. and Krantz, D. H. (1971). Conditional expected utility. *Econometrica* **39** 253–271.
- LUCE, R. D. and NARENS, L. (1978). Qualitative independence in probability theory. Theory and Decision 9 225-239.
- LUCE, R. D. and SUPPES, P. (1965). Preference, utility, and subjective probability. *Handbook of Mathematical Psychology* 3, 249-410. Wiley, New York.
- MAY, K. O. (1954). Intransitivity, utility, and the aggregation of preference patterns. *Econometrica* 22 1-13.
- MYERSON, R. B. (1979). An axiomatic derivation of subjective probability, utility and evaluation functions. *Theory and Deci*sion 11 339-352.
- NARENS, L. (1974). Minimal conditions for additive conjoint measurement and qualitative probability. J. Math. Psychology 11 404-430.
- NIINLUOTO, I. (1972). A note on fine and tight qualitative probabilities. Ann. Math. Statist. 43 1581-1591.
- PFANZAGL, J. (1967). Subjective probability derived from the Morgenstern-von Neumann utility concept. In Essays in Mathematical Economics 237-251. Princeton Univ. Press, Princeton, N. J.
- PFANZAGL, J. (1968). Theory of Measurement. Wiley, New York.
- Pratt, J. W., Raiffa, H. and Schlaifer, R. (1964). The foundations of decision under uncertainty: an elementary exposition. J. Amer. Statist. Assoc. **59** 353-375.
- Purves, R. A. and Sudderth, W. D. (1976). Some finitely additive probability. *Ann. Probab.* 4 259-276.
- RAIFFA, H. and SCHLAIFER, H. (1961). Applied Statistical Decision Theory. Harvard Univ., Boston.
- RAMSEY, F. P. (1931). Truth and probability. In *The Foundations* of Mathematics and Other Logical Essays. Humanities Press, New York. Reprinted in Kyburg and Smokler (1964).
- ROBERTS, F. S. (1973). A note on Fine's axioms for qualitative probability. *Ann. Probab.* **1** 484-487.
- ROBERTS, F. S. (1979). Measurement Theory. Addison-Wesley, Reading, Mass.

- SAVAGE, L. J. (1954). The Foundations of Statistics. Wiley, New York, Dover edition, 1972.
- SAVAGE, L. J., HACKING, I. and SHIMONY, A. (1967). A panel discussion of personal probability. *Philos. Sci.* **34** 305–332.
- SCHERVISH, M. J., SEIDENFELD, T. and KADANE, J. B. (1984). The extent of non-conglomerability of finitely additive probabilities. Z. Wahrsch. Verw. Gebiete 66 205–226.
- Schlaffer, R. (1959). Probability and Statistics for Business Decisions. McGraw-Hill, New York.
- SCHMEIDLER, D. (1984). Subjective probability and expected utility without additivity. Preprint 84, Institute for Mathematics and Its Applications, Univ. Minnesota.
- Scott, D. (1964). Measurement structures and linear inequalities. J. Math. Psychology 1 233-247.
- SEIDENFELD, T. and SCHERVISH, M. J. (1983). A conflict between finite additivity and avoiding Dutch book. *Philos. Sci.* **50** 398-412.
- SHAFER, G. (1976). A Mathematical Theory of Evidence. Princeton Univ. Press, Princeton, N. J.
- SMITH, C. A. B. (1961). Consistency in statistical inference and decision. J. Roy. Statist. Soc. Ser. B 23 1-37.
- SOBCZYK, A. and HAMMER, P. C. (1944). A decomposition of additive set functions. *Duke Math. J.* 11 839–846.
- SUDDERTH, W. D. (1985). Review of Hartigan (1983). Bull. Amer. Math. Soc. 12 294-297.
- SUPPES, P. (1974). The measurement of belief. J. Roy. Statist. Soc. Ser. B 36 160-175.
- SUPPES, P. and ZANOTTI, M. (1976). Necessary and sufficient conditions for existence of a unique measure strictly agreeing with a qualitative probability ordering. J. Philos. Logic 5 431-438.
- Suppes, P. and Zanotti, M. (1982). Necessary and sufficient qualitative axioms for conditional probability. Z. Wahrsch. Verw. Gebiete 60 163–169.
- TVERSKY, A. (1969). Intransitivity of preferences. Psychol. Rev. 76
- VILLEGAS, C. (1964). On qualitative probability σ-algebras. Ann. Math. Statist. 35 1787–1796.
- VON NEUMANN, J. and MORGENSTERN, O. (1944). Theory of Games and Economic Behavior. Princeton Univ. Press, Princeton, N. J.
- WAKKER, P. (1981). Agreeing probability measures for comparative probability structures. *Ann. Statist.* **9** 658-662.
- WALLEY, P. and FINE, T. L. (1979). Varieties of modal (classificatory) and comparative probability. Synthese 41 321–374.
- WILLIAMS, P. M. (1976). Indeterminate probabilities. Formal Methods in the Methodology of Empirical Sciences 229-246. Ossolineum and Reidel, Dordrecht, Netherlands.
- YOSIDA, K. and HEWITT, E. (1952). Finitely additive measures. Trans. Amer. Math. Soc. 72 46-66.