

⑥ a) Jeffery's prior,  $\pi(\theta) \sim \{I(\theta)\}^{1/2}$

$$f(y[k]|\theta) = \begin{cases} \frac{\theta^{y[k]} e^{-\theta}}{(y[k])!} & \text{if } y[k] \text{ is a non-negative integer} \\ 0 & \text{otherwise} \end{cases}$$

$$L = \log \left( \prod_{k=1}^N \frac{\theta^{y[k]} e^{-\theta}}{(y[k])!} \right)$$

$$= \log(e^{-\theta})^N + \log \theta^{\sum y[k]} + C$$

$$\Rightarrow L = \sum_{k=1}^N y[k] \log \theta - N\theta + C$$

where  $C$  is a constant  
(term independent of  $\theta$ )

$$S = \frac{\partial L}{\partial \theta} = \frac{\sum y[k]}{\theta} - N$$

$$\Rightarrow \frac{\partial^2 L}{\partial \theta^2} = \frac{-\sum y[k]}{\theta^2}$$

$$E\left(-\frac{\partial^2 L}{\partial \theta^2}\right) = \frac{\sum E(y[k])}{\theta^2} = \frac{N\theta}{\theta^2}$$

( $\because E(y[k]) = \theta$ )

$$\Rightarrow I(\theta) = \frac{1}{\theta}$$

$$\therefore \text{Jeffrey's prior, } \pi(\theta) \sim \left(\frac{1}{\theta}\right)^{1/2} \\ \sim \boxed{\theta^{-1/2}}$$

So a Data Raja's prior is in the class of  
Jeffrey's priors.

$$\begin{aligned} b) \quad f(\theta | y_N) &= \frac{(f(y_N | \theta)) f(\theta)}{f(y_N)} \\ &= c \times \theta \times \frac{\prod_{k=1}^N \theta^{y[k]} e^{-\theta}}{(y[k])!} \times \theta^{-1/2} \\ &= c \times \theta^{\sum y[k]} e^{-N\theta} \times \theta^{-1/2} \\ &\quad \frac{\prod_{k=1}^N (y[k])!}{\prod_{k=1}^N (y[k])!} \end{aligned}$$

For given  $y_N$ ,  $\frac{\prod_{k=1}^N (y[k])!}{\prod_{k=1}^N (y[k])!}$  is a constant

$$\Rightarrow f(\theta | y_N) = c' \times \theta^{\sum y[k]} \times e^{-N\theta} \times \theta^{-1/2}$$

Jeffrey's prior doesn't change upon transforming  
the RV  $\Rightarrow f(2N\theta) \sim f(\theta) \sim \theta^{-1/2}$

$$\begin{aligned} f(2N\theta | y_N) &= \frac{f(y_N | 2N\theta) f(2N\theta)}{f(y_N)} \\ &= c' \times \left(\frac{2N\theta}{2N}\right)^{\sum y[k]} e^{-\frac{2N\theta}{2}} \times \left(\frac{2N\theta}{2N}\right)^{-1/2} \end{aligned}$$

Let  $x = 2N\theta$ ,

$$f(x|y_N) = c'' x \left( x \right)^{\frac{(2\sum y(k))}{2}} e^{-x/2} (x)^{-1/2}$$

$$\bar{y} = \frac{\sum y(k)}{N} \Rightarrow \sum y(k) = N\bar{y}$$

$$\Rightarrow f(x|y_N) = c'' \left( x \right)^{\left( \frac{2N\bar{y} + 1 - 1}{2} \right)} e^{-x/2} (x)^{-1/2}$$

$$= c'' \left( x \right)^{\left( \frac{2N\bar{y} + 1}{2} - 1 \right)} e^{-x/2}$$

This is of the  $\chi^2$  distribution form.

def =  $(2N\bar{y} + 1)$ . So  $c'' = \frac{2^{-\frac{(2N\bar{y} + 1)}{2}}}{\Gamma\left(\frac{2N\bar{y} + 1}{2}\right)}$

$$\Rightarrow f(x|y_N) = \frac{1}{2^{\frac{(2N\bar{y} + 1)}{2}} \Gamma\left(\frac{2N\bar{y} + 1}{2}\right)} x^{\left(\frac{2N\bar{y} + 1}{2} - 1\right)} e^{-x/2}$$

$$= \frac{1}{2^{k/2} \Gamma(k/2)} x^{k/2 - 1} e^{-x/2}$$

where  $k = \boxed{2N\bar{y} + 1}$

Mean of this distribution,  $E(x) = E(2N\theta)$  is the def  $k$ .

$$\Rightarrow E(2N\theta) = (2N\bar{y} + 1)$$

$$\therefore \boxed{f(2N\theta|y_N) \sim \chi^2(2N\bar{y} + 1)}$$

c)

$$\hat{\theta}_{MMSE} = \min_{\theta} E((\theta - \hat{\theta})^2)$$

$$= \min_{\theta} E((2N\theta - 2N\hat{\theta})^2) \times \frac{1}{(2N)^2}$$

$$\Rightarrow \hat{\theta}_{MMSE} = \min_{\theta} (E(2N\theta - 2N\hat{\theta})^2) \times \frac{1}{(2N)^2}$$

$$\hat{\theta}_{MMSE} = \frac{1}{2N} \times \min_{\theta} \left( \frac{E((x - 2N\hat{\theta})^2)}{(2N)^2} \right)$$

∵ 2N is a constant for given data

$$\hat{\theta}_{MMSE} = \frac{\hat{x}_{MMSE}}{2N} \Rightarrow \hat{\theta}_{MMSE} = \frac{\hat{x}_{MMSE}}{2N}$$

E  $\hat{x}_{MMSE}$  is the mean of the posterior  $X^2(k)$   
 $= k$

$$\Rightarrow \hat{x}_{MMSE} = 2N\bar{y} + 1$$

$$\therefore \hat{\theta}_{MMSE} = \bar{y} + \frac{1}{2N}$$

d) Let  $a$  be the lower end and  $b$  be the higher end of the credible interval. Let  $F$  be the CDF of  $X^2$

i) If a symmetric credible interval over the MMSE estimate is required:

$$F(b) - F(a) = 1 - \alpha \quad \text{--- ①}$$



$$\frac{a+b}{2} = (2N\bar{y} + 1) \quad \text{--- (2)}$$

Solving eqns (1) & (2) for  $a$  and  $b$  gives the symmetric credible interval around mean estimate of  $2N\theta$

ii) Unsymmetric credible interval ~~(when  $\hat{\theta} = \text{MMSE}$ )~~  
 → This interval will be unsymmetric when the ~~Median estimate~~ MMSE estimate is used  
 → If Median estimate is used the interval is symmetric.

→ Solve the eqns (3) & (4)

$$F(a) = 0.025 \quad \text{--- (3)} \quad F(b) = 0.975 \quad \text{--- (4)}$$

to get  $a$  &  $b$  which will represent the end points of the interval for  $2N\theta$  estimate

$$\text{CDF, } F(x = 2N\theta) = \frac{\gamma\left(\frac{k}{2}, \frac{x}{2}\right)}{\Gamma\left(\frac{k}{2}\right)}; \quad k = 2N\bar{y} + 1$$

$\gamma(\cdot) \rightarrow$  lower incomplete gamma fn.;  $\Gamma \rightarrow$  upper incomplete gamma function

$$\therefore \text{Interval for } \theta \text{ is } \left[ \frac{a}{2N}, \frac{b}{2N} \right]$$

where  $a$  &  $b$  are obtained by solving (1) & (2)  
 or (3) & (4) as needed.