

$$\textcircled{1} \quad f(y) = \begin{cases} e^{-(y-\theta)} & y > \theta, \quad -\infty < \theta < \infty \\ 0 & \text{otherwise} \end{cases}$$

$$\Pr(y[k] \geq y) = \int_y^{\infty} e^{-(y-\theta)}$$

$$\text{Let } e^{\theta} = \beta$$

$$\Rightarrow \Pr(y[k] \geq y) = \begin{cases} \beta(e^{-y}) & \text{where } y > \theta \\ 1 & y \leq \theta \end{cases}$$

$$\text{For } \min(y_N) \geq y,$$

$$y[k] \geq y \quad \forall k.$$

$$\Rightarrow \Pr(\min(y_N) \geq y) = (\beta e^{-y})^N \quad y > \theta$$

$$\Rightarrow \Pr(\min(y_N) \leq y) = 1 - \beta^N e^{-Ny} \quad \text{--- } \textcircled{1}$$

$$\Rightarrow f(\min(y_N) = y) = \frac{d}{dy} (1 - \beta^N e^{-Ny})$$

$$= \begin{cases} N\beta^N e^{-Ny} & y > \theta \\ 0 & y \leq \theta \end{cases} \quad \text{--- } \textcircled{2}$$

$$\textcircled{1} \Rightarrow \Pr(2 \min(Y_N) \leq 2\theta) = 1 - \beta^N e^{-N\theta}$$

$$\Rightarrow \Pr(2 \min(Y_N) \leq t) = 1 - \beta^N e^{-Nt/2}$$

where  $t = 2\theta$

$$\rightarrow \Pr(T \leq t) = 1 - \beta^N e^{-Nt/2}$$

$$\Rightarrow f(T_N = t) = \frac{N \beta^N e^{-Nt/2}}{2}$$

$$f(T_N = t) = \begin{cases} N \exp(N(\theta - t/2)) & t > 2\theta \\ \frac{1}{2} & t \geq 2\theta \\ 0 & \text{otherwise } (t \leq 2\theta) \end{cases}$$

$$b) E(T_N) = \int_{-\infty}^{\infty} t f(t) dt$$

$$= \frac{N}{2} \int_{2\theta}^{\infty} t e^{N\theta - Nt/2} dt$$

$$= \frac{N \beta^N}{2} \int_{2\theta}^{\infty} t e^{-Nt/2} dt \quad (\beta = e^{N\theta})$$

$$= \frac{N \beta^N}{2} \left[ \frac{2t e^{-Nt/2}}{-1/2} \Big|_{2\theta}^{\infty} + \int_{2\theta}^{\infty} \frac{2 e^{-Nt/2}}{N} dt \right]$$

$$= \frac{N \beta^N}{2} \left[ \frac{2 \cdot 2\theta e^{-N\theta}}{N} + \frac{4 e^{-N\theta}}{N^2} \right]$$

$$= 2\theta + \frac{2}{N}$$

$E(T_N) \neq \theta \Rightarrow T_N$  is a biased estimator.

Correcting it:  $\frac{T_N}{2} - \frac{1}{N} = T_N'$

$$\begin{aligned} E(T_N') &= E\left(\frac{T_N}{2} - \frac{1}{N}\right) = \frac{E(T_N)}{2} - \frac{1}{N} \\ &= \theta + \frac{1}{N} - \frac{1}{N} \end{aligned}$$

$$\Rightarrow E(T_N') = \theta \quad \checkmark \text{ unbiased}$$

$\therefore$  Corrected statistic:

$$T_N' = \frac{T_N}{2} - \frac{1}{N}$$

c)  $T_N' \xrightarrow{p} \theta$  [To prove]

To prove:  $\lim_{N \rightarrow \infty} \Pr(|X_n - X| \geq \varepsilon) = 0$  where  $\varepsilon > 0$

$$\Rightarrow \lim_{N \rightarrow \infty} \Pr(|T_N' - \theta| \geq \varepsilon) = 0$$

$$\Rightarrow \lim_{N \rightarrow \infty} \left[ \Pr(T_N' - \theta \geq \varepsilon) + \Pr(T_N' - \theta \leq -\varepsilon) \right] = 0$$

(Because for the condition to satisfy either  $T_N' - \theta \geq \varepsilon$  or  $T_N' - \theta \leq -\varepsilon$ .)

Now we will compute the 2 probabilities individually and add them up

1st term:  $\Pr\left(\frac{T_N}{2} - \frac{1}{N} - \theta \geq \varepsilon\right)$

$\Rightarrow \Pr\left(\frac{T_N}{2} \geq \varepsilon + \frac{1}{N} + \theta\right) = \Pr\left(T_N \geq 2\left(\varepsilon + \frac{1}{N} + \theta\right)\right)$

We now find out the CDF from PDF of  $T$ ,

$$\Pr(T_N \leq t) = \int_{2\theta}^t \frac{N}{2} \exp\left(N\left(\theta - \frac{t}{2}\right)\right) dt$$

$$= 1 - \beta^N e^{-\frac{Nt}{2}} \quad \text{--- ①}$$

$$\Pr(T_N \geq t) = 1 - \Pr(T_N \leq t) \quad \left[ \beta^N = e^{N\theta} \text{ as defined in part (a)} \right]$$

$$= \beta^N e^{-Nt/2} \quad \text{--- ②}$$

$$\therefore \Pr\left(T_N \geq 2\left(\varepsilon + \frac{1}{N} + \theta\right)\right) = \beta^N e^{-N\left(\varepsilon + \frac{1}{N} + \theta\right)}$$

$$= e^{-N\varepsilon - 1} \quad \text{--- ③}$$

2nd term:  $\Pr\left(\frac{T_N}{2} - \frac{1}{N} - \theta \leq -\varepsilon\right)$

using eqn ①,

$$\Rightarrow \Pr\left(\frac{T_N}{2} \leq 2\left(-\varepsilon + \frac{1}{N} + \theta\right)\right) = 1 - \beta^N e^{-N\left(-\varepsilon + \frac{1}{N} + \theta\right)}$$

$$= 1 - e^{N\varepsilon - 1}$$

However this eqn is valid ~~for~~ only if,

$$2\left(-\varepsilon + \frac{1}{N} + \theta\right) \geq 2\theta \Rightarrow \frac{1}{N} - \varepsilon \geq 0$$

In the event  $N \rightarrow \infty$ ,  $-\varepsilon \geq 0$  But  $\varepsilon > 0$   
 $\Rightarrow -\varepsilon < 0$

$\Rightarrow$  We can't use  $1 - e^{-N\varepsilon - 1}$  expression in the given limit

$$\therefore \text{ since } 2\left(\theta - \frac{1}{N} - \varepsilon\right) \leq 2\theta,$$

$$F(T_N \leq t) \text{ at such } t = 0 \quad (\because t_{\min} = 2\theta)$$

$$\Rightarrow \lim_{N \rightarrow \infty} \Pr\left(T_N \leq 2\left(-\varepsilon + \frac{1}{N} + \theta\right)\right) = 0 \quad \text{④}$$

Using ④ & ③ we have,

$$\begin{aligned} \lim_{N \rightarrow \infty} \left[ \Pr(T_N' - \theta \geq \varepsilon) + \Pr(T_N' - \theta \leq -\varepsilon) \right] \\ = \lim_{N \rightarrow \infty} \left[ e^{-N\varepsilon - 1} + 0 \right] \end{aligned}$$

$$= 0$$

$$\Rightarrow \lim_{N \rightarrow \infty} \left[ \Pr(T_N' - \theta \geq \varepsilon) + \Pr(T_N' - \theta \leq -\varepsilon) \right] = 0$$

$$\therefore T_N' \xrightarrow{P} \theta$$