

ECE 286: Bayesian Machine Perception

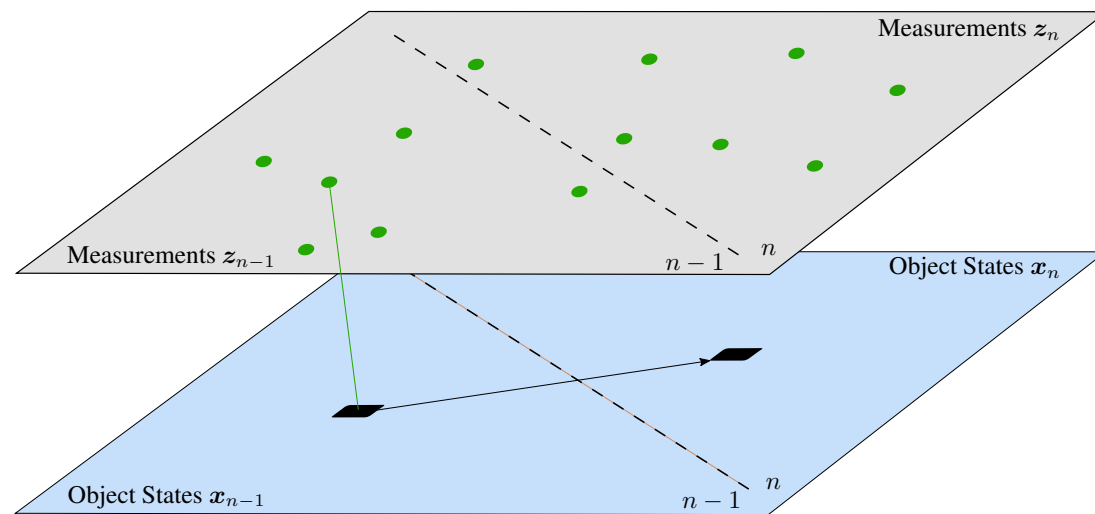
Class 8: Probabilistic Data Association

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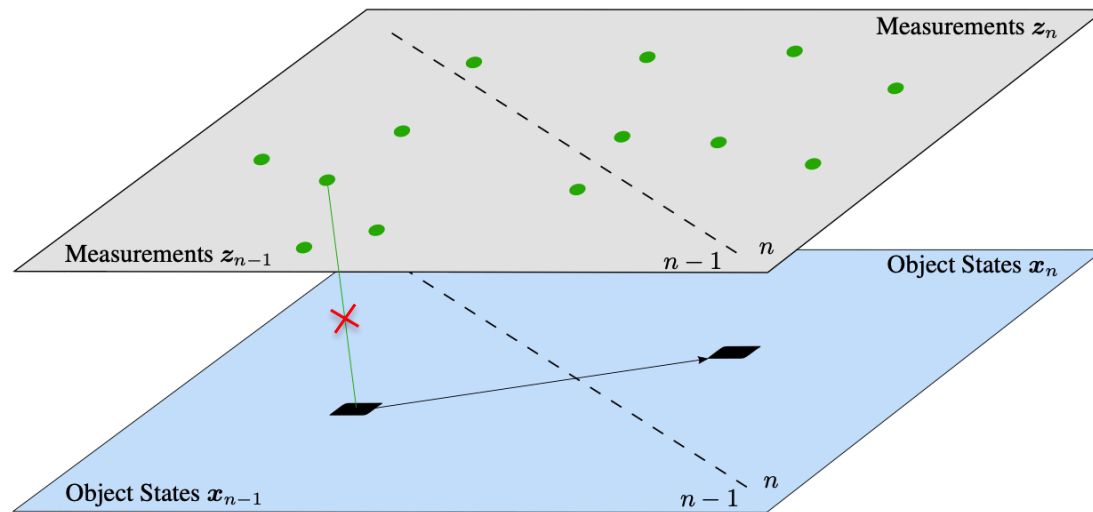
Basic Setting

- “Single object tracking in clutter” problem



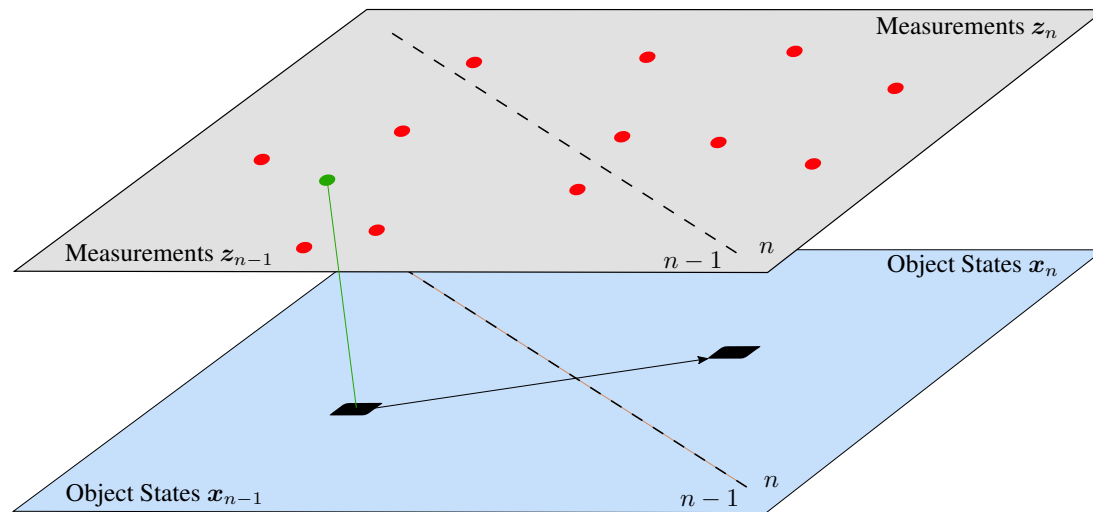
Basic Setting

- “Single object tracking in clutter” problem
- **Measurement-origin uncertainty (MOU)**, false clutter measurements and missed detections



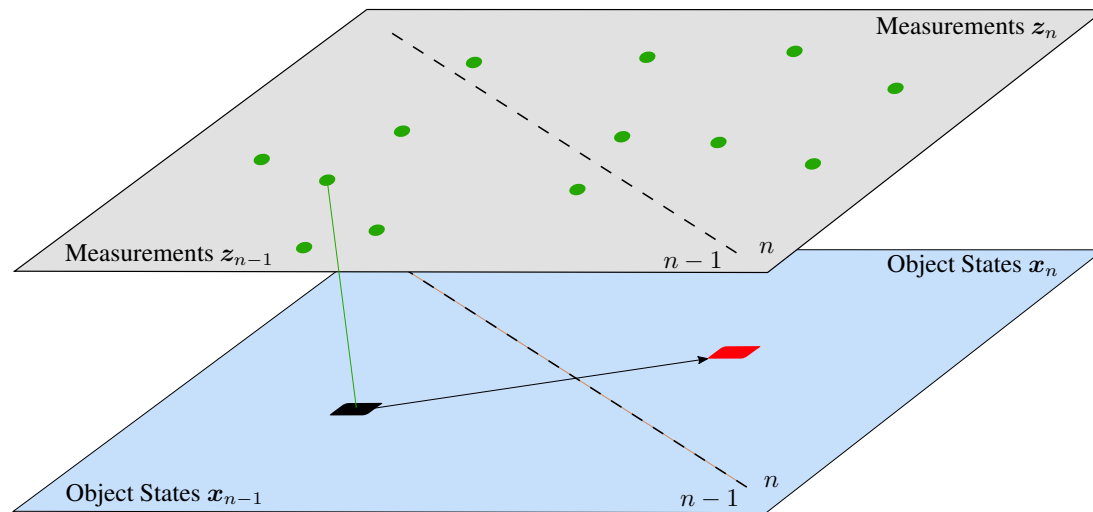
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Basic Setting

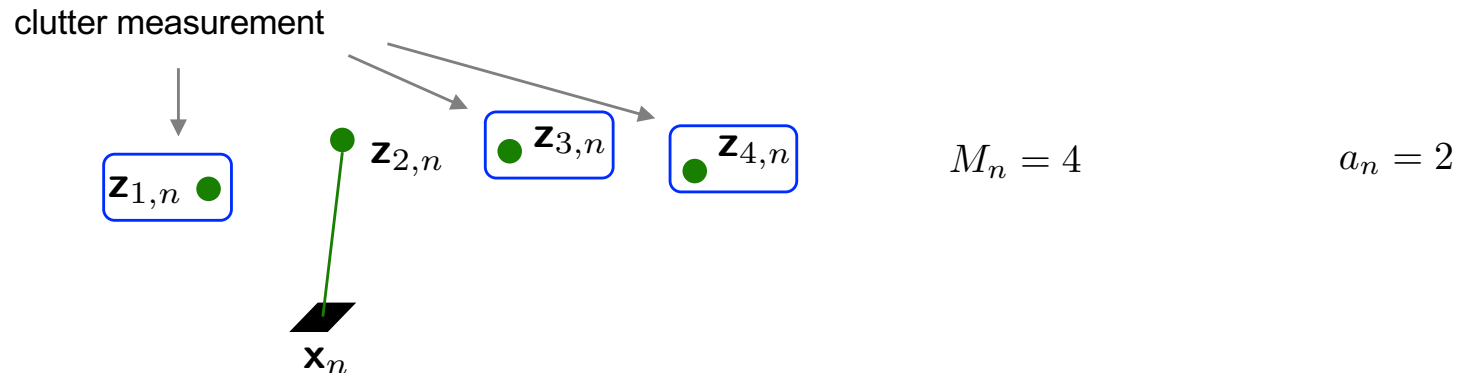
- “Single object tracking in clutter” problem
- Measurement-origin uncertainty (MOU), false clutter measurements and missed detections



The Association Variable

- Object-oriented association variable $a_n \in \{0, 1, \dots, M_n\}$ ← number of measurements
 - $a_n = m > 0$: at time n , the object generates the measurement with index m

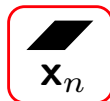
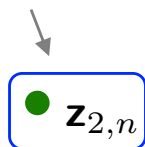
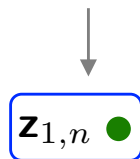
- Example 1:



The Association Variable

- Object-oriented association variable $a_n \in \{0, 1, \dots, M_n\}$
 - $a_n = m > 0$: at time n , the object generates the measurement with index m
 - $a_n = 0$: at time n , the object did not generate a measurement
- Example 2:

clutter measurement



missed detection

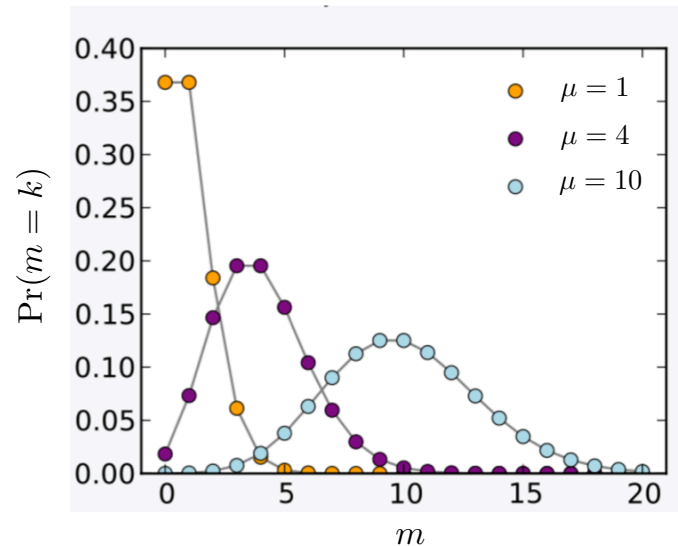
$$M_n = 2$$

$$a_n = 0$$

The Poisson Distribution

- A discrete random variable m is said to have a Poisson distribution with parameter $\mu > 0$, if for $m = 0, 1, 2, \dots$ the probability mass function is given by

$$p(m) = \frac{\mu^m e^{-\mu}}{m!}$$



The parameter μ is the mean
as well as the variance

Single Object Tracking in Clutter

- The **state of the object** is denoted $\mathbf{x}_n \in \mathbb{R}^w$ and the **joint measurement** is given by $\mathbf{z}_n \triangleq [\mathbf{z}_{1,n}^T, \mathbf{z}_{2,n}^T, \dots, \mathbf{z}_{M_n,n}^T]^T$ with entries $\mathbf{z}_{m,n} \in \mathbb{R}^d$
- The **association variable** a_n is given by
 - $m \in \{1, 2, \dots, M_n\}$, if measurement $\mathbf{z}_{m,n}$ was generated by the object
 - 0, if no measurement was generated by the object
- Association variable a_n and the number of measurements M_n are random variables

Prior Distribution

- Key Assumptions I:

- Object detection $\theta_n \in \{0, 1\}$ is a Bernoulli trial with success probability $0 < p_d \leq 1$
- The number of clutter measurements L_n is Poisson distributed with mean μ_c
- At most one measurement $z_{m,n}$ is generated by the object

- Joint prior probability mass function (pmf):

$$p(a_n, \theta_n, L_n) = p(a_n | \theta_n, L_n) p(\theta_n) p(L_n)$$

$$p(\theta_n) = \begin{cases} p_d & \theta_n = 1 \\ 1 - p_d & \theta_n = 0 \end{cases}$$

$$p(a_n | \theta_n = 1, L_n) = \begin{cases} \frac{1}{L_n + 1} & a_n \in \{1, \dots, M_n\} \\ 0 & a_n = 0 \end{cases}$$

$$p(L_n) = \frac{\mu_c^{L_n}}{L_n!} e^{-\mu_c}$$

$$p(a_n | \theta_n = 0, L_n) = \begin{cases} 0 & a_n \in \{1, \dots, M_n\} \\ 1 & a_n = 0 \end{cases}$$

Prior Distribution

- After variable transform $L_n + \theta_n \rightarrow M_n$ we obtain

$$p(a_n, M_n) = \begin{cases} p_d \frac{\mu_c^{M_n-1}}{M_n!} e^{-\mu_c} & a_n \in \{1, \dots, M_n\} \\ (1 - p_d) \frac{\mu_c^{M_n}}{M_n!} e^{-\mu_c} & a_n = 0 \end{cases}$$

- **Properties:**

- For all arbitrary $L_n + \theta_n = M_n$ we have $p(a_n, \theta_n, L_n) = p(a_n, M_n)$
- $p(a_n, M_n)$ is a valid pmf in the sense that $\sum_{M_n=0}^{\infty} \sum_{a_n=0}^{M_n} p(a_n, M_n) = 1$

Likelihood Function

- Key Assumption II:

- Clutter measurements are independent and identically distributed (iid) according to $f_c(\mathbf{z}_{m,n})$
- Condition on \mathbf{x}_n , the object-generated measurement $\mathbf{z}_{a_n,n}$ is conditionally independent of all the other measurements

- Likelihood function:

- for $\mathbf{z}_n \in \mathbb{R}^{M_n d}$,

measurement model $\mathbf{z}_{a_n,n} = h_n(\mathbf{x}_n, \mathbf{v}_n)$ with noise \mathbf{v}_n

$$f(\mathbf{z}_n | \mathbf{x}_n, a_n, M_n) = \begin{cases} \prod_{m=1}^{M_n} f_c(\mathbf{z}_{m,n}) & a_n = 0 \\ \frac{f(\mathbf{z}_{a_n,n} | \mathbf{x}_n)}{f_c(\mathbf{z}_{a_n,n})} \prod_{m=1}^{M_n} f_c(\mathbf{z}_{m,n}) & a_n \in \{1, \dots, M_n\} \end{cases}$$

- For $\mathbf{z}_n \notin \mathbb{R}^{M_n d}$,

$$f(\mathbf{z}_n | \mathbf{x}_n, a_n, M_n) = 0$$

Joint Distributions

- Joint prior for $\mathbf{x}_{0:n}$

State transition model $\mathbf{x}_n = g_n(\mathbf{x}_{n-1}, \mathbf{u}_n)$ with noise \mathbf{u}_n

$$f(\mathbf{x}_{0:n}) = f(\mathbf{x}_0) \prod_{n'=1}^n \boxed{f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1})}$$

Driving noise independent
across time n and
independent of \mathbf{x}_0

- Joint prior for $\mathbf{a}_{1:n}$ and $\mathbf{M}_{1:n}$

$$p(\mathbf{a}_{1:n}, \mathbf{M}_{1:n}) = \prod_{n'=1}^n p(a_{n'}, M_{n'})$$

Measurement generation
independent across time n

- Joint likelihood function

$$f(\mathbf{z}_{1:n} | \mathbf{x}_{1:n}, \mathbf{a}_{1:n}, \mathbf{M}_{1:n}) = \prod_{n'=1}^n f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, a_{n'}, M_{n'})$$

Measurement noise and
clutter independent across
time n

The Joint Posterior Distribution

- The joint posterior distribution ($M_{1:n}$ and $z_{1:n}$ are observed and thus fixed)

$$f(\mathbf{x}_{0:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) = f(\mathbf{x}_{0:n}, \mathbf{a}_{1:n}, \mathbf{M}_{1:n} | \mathbf{z}_{1:n}) \quad \longleftarrow \mathbf{M}_{1:n} \text{ fixed}$$

$$\text{Bayes rule} \quad \longrightarrow \quad \propto f(\mathbf{z}_{1:n} | \mathbf{x}_{1:n}, \mathbf{a}_{1:n}, \mathbf{M}_{1:n}) f(\mathbf{x}_{0:n}, \mathbf{a}_{1:n}, \mathbf{M}_{1:n})$$

$$\mathbf{x}_{0:n} \perp\!\!\!\perp \mathbf{a}_{1:n}, \mathbf{M}_{1:n} \quad \longrightarrow \quad = f(\mathbf{z}_{1:n} | \mathbf{x}_{1:n}, \mathbf{a}_{1:n}, \mathbf{M}_{1:n}) f(\mathbf{x}_{0:n}) p(\mathbf{a}_{1:n}, \mathbf{M}_{1:n})$$

$$\text{Expressions for joint distributions} \quad \longrightarrow \quad = f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, \mathbf{a}_{n'}, M_{n'}) p(\mathbf{a}_{n'}, M_{n'})$$

Problem Formulation

- Input at time n :
 - All observations up to time $\mathbf{z}_{1:n}$
 - “Markovian” statistical model

$$f(\mathbf{x}_{1:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) \propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, \mathbf{a}_{n'}, M_{n'}) p(\mathbf{a}_{n'}, M_{n'})$$


- Output at time n :
 - Estimate of $\hat{\mathbf{x}}_n$
- Calculation of an estimate $\hat{\mathbf{x}}_n$ is based on the **marginal posterior pdf** $f(\mathbf{x}_n | \mathbf{z}_{1:n})$

The Factor Graph

- Recall factorization of the joint posterior distribution:


$$f(\mathbf{x}_{1:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) \propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, a_{n'}, M_{n'}) p(a_{n'}, M_{n'})$$

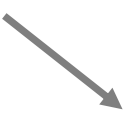
observed and fixed



The Factor Graph

- Recall factorization of the joint posterior distribution:

$$\begin{aligned}
 f(\mathbf{x}_{1:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) &\propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, a_{n'}, M_{n'}) p(a_{n'}, M_{n'}) \\
 &\propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) g_1(\mathbf{x}_{n'}, a_{n'}) g_2(a_{n'})
 \end{aligned}$$



$$f(\mathbf{z}_n | \mathbf{x}_n, a_n, M_n) = \begin{cases} \frac{f(\mathbf{z}_{a_n, n} | \mathbf{x}_n)}{f_c(\mathbf{z}_{a_n, n})} \prod_{m=1}^M f_c(\mathbf{z}_{m, n}) & a_n \in \{1, \dots, M_n\} \\ \prod_{m=1}^M f_c(\mathbf{z}_{m, n}) & a_n = 0 \end{cases}$$


↑
constant

$$g_1(\mathbf{x}_n, a_n) = \begin{cases} \frac{f(\mathbf{z}_{a_n, n} | \mathbf{x}_n)}{f_c(\mathbf{z}_{a_n, n})} & a_n \in \{1, \dots, M_n\} \\ 1 & a_n = 0 \end{cases}$$

The Factor Graph

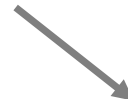
- Recall factorization of the joint posterior distribution:

$$\begin{aligned}
 f(\mathbf{x}_{1:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) &\propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, a_{n'}, M_{n'}) p(a_{n'}, M_{n'}) \\
 &\propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) g_1(\mathbf{x}_{n'}, a_{n'}) g_2(a_{n'})
 \end{aligned}$$


$$p(a_n, M_n) = \begin{cases} p_d \frac{\mu_c^{M_n-1}}{M_n!} e^{-\mu_c} & a_n = \{1, \dots, M_n\} \\ (1 - p_d) \frac{\mu_c^{M_n}}{M_n!} e^{-\mu_c} & a_n = 0 \end{cases}$$

constant



$$g_2(a_n) = \begin{cases} \frac{p_d}{\mu_c} & a_n \in \{1, \dots, M_n\} \\ (1 - p_d) & a_n = 0 \end{cases}$$


The Factor Graph

- Recall factorization of the joint posterior distribution:

$$\begin{aligned} f(\mathbf{x}_{1:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) &\propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{z}_{n'} | \mathbf{x}_{n'}, a_{n'}, M_{n'}) p(a_{n'}, M_{n'}) \\ &\propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) g_1(\mathbf{x}_{n'}, a_{n'}) g_2(a_{n'}) \\ &= f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) g_{\mathbf{z}_n}(\mathbf{x}_{n'}, a_{n'}) \end{aligned}$$

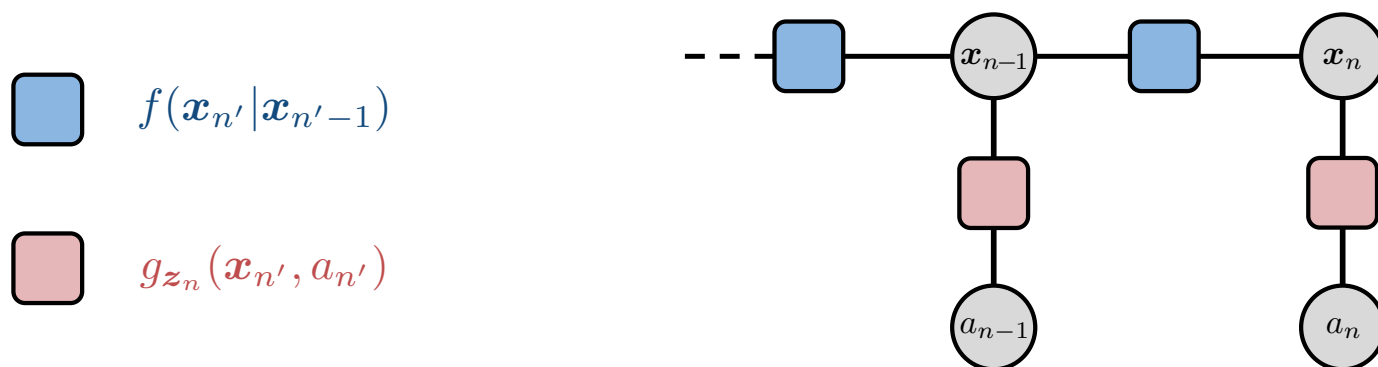
$$g_{\mathbf{z}_n}(\mathbf{x}_n, a_n) = g_1(\mathbf{x}_n, a_n) g_2(a_n) = \begin{cases} \frac{p_d f(\mathbf{z}_{a_n, n} | \mathbf{x}_n)}{\mu_c f_c(\mathbf{z}_{a_n, n})} & a_n \in \{1, \dots, M_n\} \\ (1 - p_d) & a_n = 0 \end{cases}$$

The Factor Graph

- Recall factorization of the joint posterior distribution:

$$f(\mathbf{x}_{1:n}, \mathbf{a}_{1:n} | \mathbf{z}_{1:n}) \propto f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) g_{\mathbf{z}_n}(\mathbf{x}_{n'}, a_{n'})$$

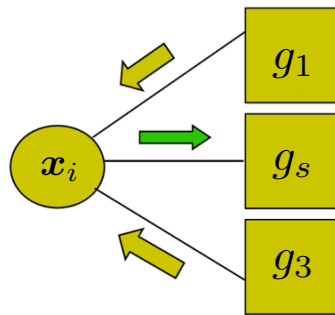
- Factor graph for two time steps $n' \in \{n-1, n\}$



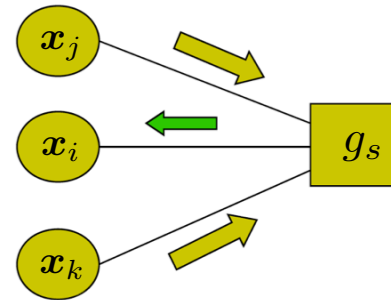
- The factor graph is cycle free \Rightarrow message passing can provide exact marginals

Recall Message Passing Rules

- **Message passing protocol:** A message to a neighboring node can only be send when it has received messages from all its other neighbors
- Marginal distribution can be calculated as $b(\mathbf{x}_i) \propto \prod_{t \in \mathcal{N}(i)} \phi_{ti}(\mathbf{x}_i) = \phi_{ti}(\mathbf{x}_i) \nu_{it}(\mathbf{x}_i)$



$$\nu_{is}(\mathbf{x}_i) = \prod_{t \in \mathcal{N}(i) \setminus s} \phi_{ti}(\mathbf{x}_i)$$



$$\phi_{si}(\mathbf{x}_i) = \int \left(g_s(\mathbf{x}_{\mathcal{N}(s)}) \prod_{j \in \mathcal{N}(s) \setminus i} \nu_{js}(\mathbf{x}_j) \right) d\mathbf{x}_{\mathcal{N}(s) \setminus i}$$

Prediction

- Prediction step:

$$f(\mathbf{x}_n | \mathbf{y}_{1:n-1}) = \int f(\mathbf{x}_n | \mathbf{x}_{n-1}) f(\mathbf{x}_{n-1} | \mathbf{y}_{1:n-1}) d\mathbf{x}_{n-1}$$

$$\phi_{\rightarrow n}(\mathbf{x}_n) = \int f(\mathbf{x}_n | \mathbf{x}_{n-1}) \nu_{\rightarrow n}(\mathbf{x}_{n-1}) d\mathbf{x}_{n-1}$$

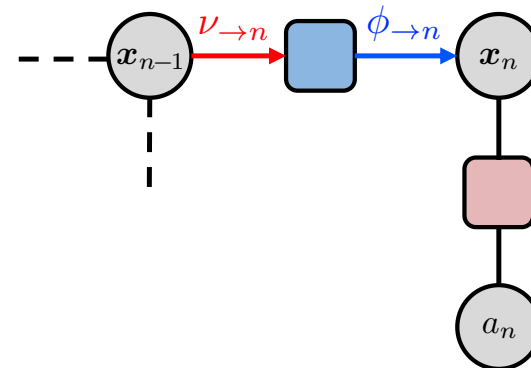
- Factor graph for two time steps $n' \in \{n-1, n\}$



$$f(\mathbf{x}_n | \mathbf{x}_{n-1})$$



$$g_{z_n}(\mathbf{x}_{n'}, a_{n'})$$



Data Association

- Data association step:

$$\phi_n(\mathbf{x}_n) = \sum_{m=0}^{M_n} g_{z_n}(\mathbf{x}_n, a_n = m) \nu_n(a_n = m) \quad \nu_n(a_n) = 1$$

(no other neighbors)

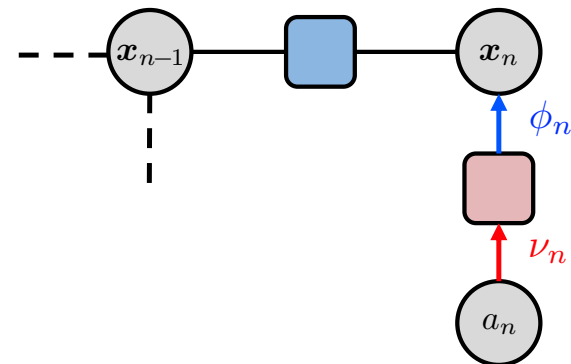
- Factor graph for two time steps $n' \in \{n-1, n\}$



$$f(\mathbf{x}_n | \mathbf{x}_{n-1})$$



$$g_{z_n}(\mathbf{x}_{n'}, a_{n'})$$



Update

- Update step:

$$f(\mathbf{x}_n | \mathbf{z}_{1:n}) \propto \phi_n(\mathbf{x}_n) \phi_{\rightarrow n}(\mathbf{x}_n)$$

$$f(\mathbf{x}_n | \mathbf{z}_{1:n}) \propto \phi_n(\mathbf{x}_n) f(\mathbf{x}_n | \mathbf{y}_{1:n-1})$$

$$\nu_{\rightarrow n+1}(\mathbf{x}_n) = \phi_n(\mathbf{x}_n) \phi_{\rightarrow n}(\mathbf{x}_n)$$

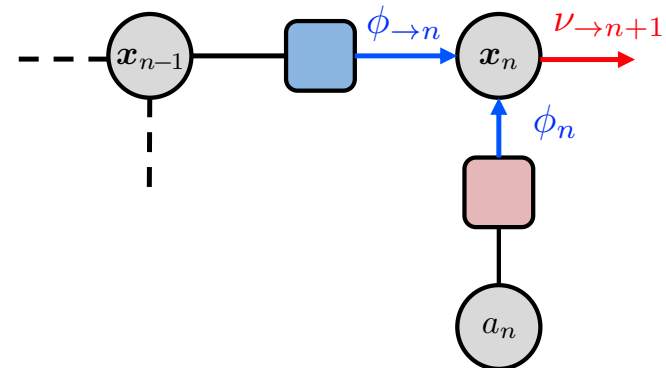
- Factor graph for two time steps $n' \in \{n-1, n\}$



$$f(\mathbf{x}_n | \mathbf{x}_{n-1})$$



$$g_{z_n}(\mathbf{x}_{n'}, a_{n'})$$



Particle-Based Update Step (cf. Class 4)

- **Given:** Particles $\{(\mathbf{x}_n^{(j)})\}_{j=1}^J \simeq f(\mathbf{x}_n|\mathbf{y}_{1:n-1})$ representing the **predicted posterior PDF**
- **Wanted:** Particles $\{(\bar{\mathbf{x}}_n^{(j)})\}_{j=1}^J \simeq f(\mathbf{x}_n|\mathbf{y}_{1:n})$ representing the **posterior PDF**
- Perform importance sampling with proposal distribution $f_p(\mathbf{x}_n) = f(\mathbf{x}_n|\mathbf{y}_{1:n-1})$ and target distribution $f_t(\mathbf{x}_n) \propto \phi_n(\mathbf{x}_n)f(\mathbf{x}_n|\mathbf{y}_{1:n-1})$
 - calculate unnormalized weights $\tilde{w}_n^{(j)} = \sum_{m=0}^{M_n} g_{z_n}(\mathbf{x}_n^{(j)}, a_n = m) \propto f_t(\mathbf{x}_n^{(j)})/f_p(\mathbf{x}_n^{(j)})$
 - normalize weights $w_n^{(j)} = \tilde{w}_n^{(j)} / \sum_{j'=1}^J \tilde{w}_n^{(j')}$, $j = 1, \dots, J$
- Perform resampling to get $\{(\bar{\mathbf{x}}_n^{(j)})\}_{j=1}^J \simeq f(\mathbf{x}_n|\mathbf{y}_{1:n})$ from $\{(\mathbf{x}_n^{(j)}, w_n^{(j)})\}_{j=1}^J \simeq f(\mathbf{x}_n|\mathbf{y}_{1:n})$

Summary

- Single object tracking in clutter
 - possible association events are modelled by discrete random variable
 - data association is performed by summing over all possible association events
 - the sequential estimation problem that can be represented by a cycle free factor graph
 - a particle-based implementation can provide exact estimation results as the number of particles goes to infinity