

# ECE 275A: Parameter Estimation I

## The Kalman Filter

Florian Meyer

Electrical and Computer Engineering Department  
University of California San Diego

UCSanDiego  
JACOBS SCHOOL OF ENGINEERING

0

## Recap: State-Space Model

- Consider a sequence of states  $\mathbf{x}_n$  and a sequence of measurements  $\mathbf{y}_n$

### State-Transition Model:

State  $\mathbf{x}_n$  evolves according to

$$\mathbf{x}_n = g_n(\mathbf{x}_{n-1}) + \underbrace{\mathbf{u}_n}_{\text{driving noise (white)}}$$

This determines the joint prior

$$f(\mathbf{x}_{0:n}) = f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1})$$

### Measurement Model:

Measurement  $\mathbf{y}_n$  is generated as

$$\mathbf{y}_n = h_n(\mathbf{x}_n) + \underbrace{\mathbf{v}_n}_{\text{measurement noise (white)}}$$

This determines the joint likelihood

$$f(\mathbf{y}_{1:n} | \mathbf{x}_{1:n}) = \prod_{n'=1}^n f(\mathbf{y}_{n'} | \mathbf{x}_{n'})$$

- By using Bayes' rule we obtain the marginal posterior pdf (for  $\mathbf{y}_{1:n}$  fixed)

$$f(\mathbf{x}_{0:n} | \mathbf{y}_{1:n}) \propto f(\mathbf{x}_{0:n}) f(\mathbf{y}_{1:n} | \mathbf{x}_{1:n}) = f(\mathbf{x}_0) \prod_{n'=1}^n f(\mathbf{x}_{n'} | \mathbf{x}_{n'-1}) f(\mathbf{y}_{n'} | \mathbf{x}_{n'})$$

1

1

## Linear-Gaussian State-Space Model

- Consider a sequence of states  $\mathbf{x}_n$  and a sequence of measurements  $\mathbf{y}_n$

### State-Transition Model:

State  $\mathbf{x}_n$  evolves according to

$$\mathbf{x}_n = \mathbf{G}_n \mathbf{x}_{n-1} + \underbrace{\mathbf{u}_n}_{\text{driving noise (white)}}$$

with Gaussian driving noise

$$\mathbf{u}_n \sim \mathcal{N}(\mathbf{0}, \Sigma_{\mathbf{u}_n})$$

### Measurement Model:

Measurement  $\mathbf{y}_n$  is generated as

$$\mathbf{y}_n = \mathbf{H}_n \mathbf{x}_n + \underbrace{\mathbf{v}_n}_{\text{measurement noise (white)}}$$

with Gaussian measurement noise

$$\mathbf{v}_n \sim \mathcal{N}(\mathbf{0}, \Sigma_{\mathbf{v}_n})$$

- Prior PDF at  $n = 0$ ,  $\mathbf{x}_0 \sim \mathcal{N}(\boldsymbol{\mu}_{\mathbf{x}_0}, \boldsymbol{\Sigma}_{\mathbf{x}_0})$

2

2

## Kalman Prediction Step

- Recall prediction step of sequential Bayesian estimation

$$\underbrace{f(\mathbf{x}_n | \mathbf{y}_{1:n-1})}_{\text{Predicted posterior pdf}} = \int \underbrace{f(\mathbf{x}_n | \mathbf{x}_{n-1})}_{\text{State-transition pdf}} \underbrace{f(\mathbf{x}_{n-1} | \mathbf{y}_{1:n-1})}_{\text{Previous posterior pdf}} d\mathbf{x}_{n-1}$$

- If  $f(\mathbf{x}_{n-1} | \mathbf{y}_{1:n-1}) = \mathcal{N}(\boldsymbol{\mu}_{\mathbf{x}_{n-1}}, \boldsymbol{\Sigma}_{\mathbf{x}_{n-1}})$  and  $f(\mathbf{x}_n | \mathbf{x}_{n-1}) = \mathcal{N}(\mathbf{G}_n \mathbf{x}_{n-1}; \boldsymbol{\Sigma}_{\mathbf{u}_n})$

- Then  $f(\mathbf{x}_n | \mathbf{y}_{1:n-1})$  is Gaussian with mean  $\boldsymbol{\mu}_{\mathbf{x}_n}^-$  and covariance  $\boldsymbol{\Sigma}_{\mathbf{x}_n}^-$  given as

$$\boldsymbol{\mu}_{\mathbf{x}_n}^- = \mathbf{G}_n \boldsymbol{\mu}_{\mathbf{x}_{n-1}} \quad \boldsymbol{\Sigma}_{\mathbf{x}_n}^- = \mathbf{G}_n \boldsymbol{\Sigma}_{\mathbf{x}_{n-1}} \mathbf{G}_n^T + \boldsymbol{\Sigma}_{\mathbf{u}_n}$$

3

3

## Kalman Update Step

- Recall measurement update step of sequential Bayesian estimation

$$\underbrace{f(\mathbf{x}_n | \mathbf{y}_{1:n})}_{\text{Posterior pdf}} \propto \underbrace{f(\mathbf{y}_n | \mathbf{x}_n)}_{\text{Likelihood function}} \underbrace{f(\mathbf{x}_n | \mathbf{y}_{1:n-1})}_{\text{Predicted posterior pdf}}$$

- If  $f(\mathbf{x}_n | \mathbf{y}_{1:n-1}) = \mathcal{N}(\boldsymbol{\mu}_{\mathbf{x}_n}^-, \boldsymbol{\Sigma}_{\mathbf{x}_n}^-)$  and  $f(z_n | \mathbf{x}_n) = \mathcal{N}(\mathbf{H}_n \mathbf{x}_n; \boldsymbol{\Sigma}_{\mathbf{v}_n})$
- Then  $f(\mathbf{x}_n | \mathbf{y}_{1:n})$  is Gaussian with mean  $\boldsymbol{\mu}_{\mathbf{x}_n}$  and covariance  $\boldsymbol{\Sigma}_{\mathbf{x}_n}$  (both can be calculated in closed form)

4

4

## Kalman Update Step

- Kalman gain

$$\mathbf{K}_n = \boldsymbol{\Sigma}_{\mathbf{x}_n}^- \mathbf{H}_n^T (\mathbf{H}_n \boldsymbol{\Sigma}_{\mathbf{x}_n}^- \mathbf{H}_n^T + \boldsymbol{\Sigma}_{\mathbf{v}_n})^{-1}$$

- Mean and covariance update

$$\begin{aligned}\boldsymbol{\mu}_{\mathbf{x}_n} &= \boldsymbol{\mu}_{\mathbf{x}_n}^- + \mathbf{K}_n (\mathbf{y}_n - \mathbf{H}_n \boldsymbol{\mu}_{\mathbf{x}_n}^-) \\ \boldsymbol{\Sigma}_{\mathbf{x}_n} &= \boldsymbol{\Sigma}_{\mathbf{x}_n}^- - \mathbf{K}_n \mathbf{H}_n \boldsymbol{\Sigma}_{\mathbf{x}_n}^-\end{aligned}$$

5

5

## Kalman Filter Properties (i)

- Kalman filter provides a procedure for calculating the entire **posterior PDF**  $f(\mathbf{x}_n | \mathbf{y}_{1:n})$  of linear-Gaussian sequential Bayesian estimation problems
- The covariance matrix  $\Sigma_{\mathbf{x}_n}$  does not depend on the measurements  $\mathbf{y}_{1:n}$  and can thus be calculated offline
- Since for Gaussian distributions the mean is equal to the maximum (or mode),  $\mu_{\mathbf{x}_n}$  is the optimum **MMSE estimate** and **MAP estimate**, i.e.,

$$\hat{\mathbf{x}}_n^{\text{MMSE}} = \hat{\mathbf{x}}_n^{\text{MAP}} = \boldsymbol{\mu}_{\mathbf{x}_n}$$

and  $\Sigma_{\mathbf{x}_n}$  is the **error covariance matrix** of the estimate

7

7

## Kalman Filter Properties (ii)

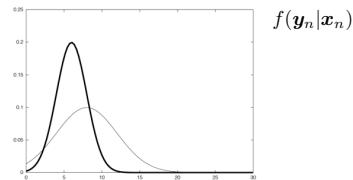
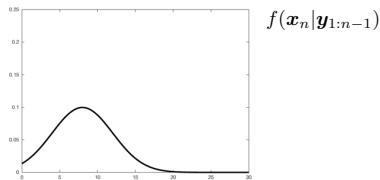
- In case  $f(\mathbf{u}_{n'})$  and  $f(\mathbf{v}_{n'}), n' = 0, 1, \dots, n$  or  $f(\mathbf{x}_0)$  are not Gaussian distributions, it can be shown that the Kalman filter is the **best linear MMSE estimator** for  $\mathbf{x}_n$
- The complexity of the Kalman filter is quite moderate, for  $I$  the dimension of the state space and  $L$  the dimension of the measurement space
  - it scales as  $\mathcal{O}(I^2)$  due matrix multiplication in  $\Sigma_{\mathbf{x}_n} = (\mathbf{I} - \mathbf{K}_n \mathbf{H}_n) \Sigma_{\mathbf{x}_n}^-$
  - it scales as  $\mathcal{O}(L^{2.4})$  due matrix inversion in  $\mathbf{K}_n = \Sigma_{\mathbf{x}_n}^- \mathbf{H}_n^T (\Sigma_{\mathbf{z}_n})^{-1}$

J. Anderson and B. Moore, *Optimal Filtering*, Prentice-Hall, 1979.

8

8

## Kalman Filter Illustration

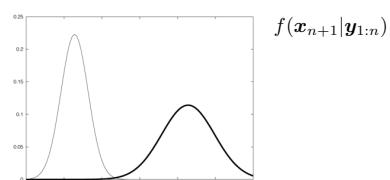
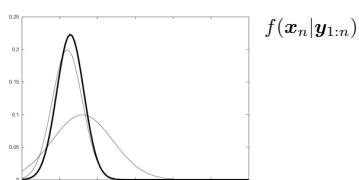
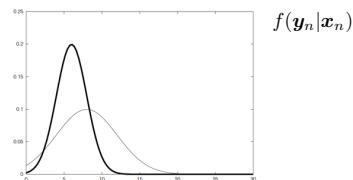
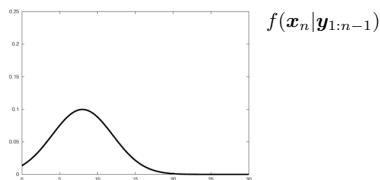


S. Thrun, W. Burgard, and D. Fox, *Probabilistic Robotics*, MIT Press, 2006.

9

9

## Kalman Filter Illustration

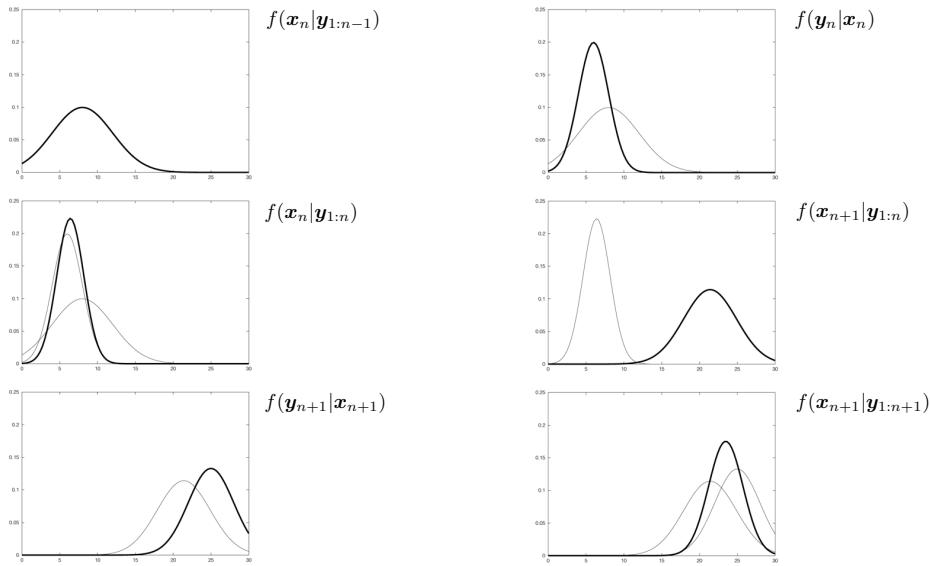


S. Thrun, W. Burgard, and D. Fox, *Probabilistic Robotics*, MIT Press, 2006.

10

10

## Kalman Filter Illustration



S. Thrun, W. Burgard, and D. Fox, *Probabilistic Robotics*, MIT Press, 2006.

11

11