

# ECE 275A: Parameter Estimation I

## Bayesian Estimation – Part II

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# Optimality of $E(\theta|y)$

$$\begin{aligned}\text{BMSE}(\hat{\theta}) &= E_{y,\theta}(\tilde{\theta}\tilde{\theta}^T) \\&= E_{y,\theta}((\hat{\theta} - \theta)(\hat{\theta} - \theta)^T) \\&= E_{y,\theta}(\underbrace{(\hat{\theta} - \hat{\theta}_{\text{MMSE}})}_{\Delta\hat{\theta}} + \underbrace{\hat{\theta}_{\text{MMSE}} - \theta}_{\tilde{\theta}_{\text{MMSE}}})(\hat{\theta} - \theta)^T) \\&= E_{y,\theta}((\Delta\hat{\theta} + \tilde{\theta}_{\text{MMSE}})(\Delta\hat{\theta} + \tilde{\theta}_{\text{MMSE}})^T) \\&= E_{y,\theta}(\Delta\hat{\theta}\Delta\hat{\theta}^T) + 2E_{y,\theta}(\Delta\hat{\theta}\tilde{\theta}_{\text{MMSE}}^T) + E_{y,\theta}(\tilde{\theta}_{\text{MMSE}}\tilde{\theta}_{\text{MMSE}}^T) \\&= E_{y,\theta}(\Delta\hat{\theta}\Delta\hat{\theta}^T) + 2E_y\left(\Delta\hat{\theta}\underbrace{E_{\theta|y}(\tilde{\theta}_{\text{MMSE}}^T|y)}_{=0}\right) + E_{y,\theta}(\tilde{\theta}_{\text{MMSE}}\tilde{\theta}_{\text{MMSE}}^T) \\&= \underbrace{E_{y,\theta}(\Delta\hat{\theta}\Delta\hat{\theta}^T)}_{\geq 0} + \underbrace{E_{y,\theta}(\tilde{\theta}_{\text{MMSE}}\tilde{\theta}_{\text{MMSE}}^T)}_{\text{BMSE}(\hat{\theta}_{\text{MMSE}})} \geq \text{BMSE}(\hat{\theta}_{\text{MMSE}})\end{aligned}$$

- The equality holds iff  $\hat{\theta} = \hat{\theta}_{\text{MMSE}}$

## Example: Binary Symbols in Noise

- Suppose we have a scalar observation  $y = x + n$  where  $n \sim \mathcal{N}(0, 1)$  and the prior distribution of the unknown parameter  $x \in \{-1, 1\}$  is

$$p(x) = \frac{1}{2}\delta(x + 1) + \frac{1}{2}\delta(x - 1)$$

where  $\delta(\cdot)$  is the Kronecker delta and  $x \perp\!\!\!\perp n$

- The marginal distribution of  $y$  is given by

$$p(y) = \frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y+1)^2}{2}\right) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y-1)^2}{2}\right)$$

## Example: Binary Symbols in Noise

- The posterior distribution  $p(x|y)$  can be calculated as

$$\begin{aligned} p(x|y) &= \frac{p(y|x)p(x)}{p(y)} \\ &= \frac{\frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y-x)^2}{2}\right)}{\frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y+1)^2}{2}\right) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y-1)^2}{2}\right)} \delta(x+1) \\ &\quad + \frac{\frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y-x)^2}{2}\right)}{\frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y+1)^2}{2}\right) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(y-1)^2}{2}\right)} \delta(x-1) \\ &= \frac{\exp(xy)}{\exp(-y) + \exp(y)} \delta(x+1) + \frac{\exp(xy)}{\exp(-y) + \exp(y)} \delta(x-1) \end{aligned}$$

- Thus, we obtain

$$\hat{x}_{\text{MMSE}} = E(x|y) = \frac{\exp(y) - \exp(-y)}{\exp(y) + \exp(-y)} = \tanh(y)$$

- A MAP estimate can be obtained as

$$\hat{x}_{\text{MAP}} = \begin{cases} 1 & \text{if } y \geq 0 \\ -1 & \text{if } y < 0 \end{cases}$$

# Linear Minimum Mean Square Error Estimation

- Now we restrict the estimator to be a linear function with respect to the observation, i.e.,

$$\hat{\boldsymbol{\theta}} = \mathbf{C}\mathbf{y} + \mathbf{d}$$

where  $\mathbf{C} \in \mathbb{R}^{p \times m}$  and  $\mathbf{d} \in \mathbb{R}^p$

- We aim to find the linear estimator  $\hat{\boldsymbol{\theta}}_{\text{LMMSE}} = \mathbf{A}\mathbf{y} + \mathbf{b}$  that minimizes the Bayesian mean square error
- By using the orthogonality principle, we have

$$\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}_{\text{LMMSE}} \perp \text{LF}(\mathbf{y}) \subset \mathcal{L}(\Omega, \Theta)$$

where

$$\text{LF}(\mathbf{y}) = \{ \mathbf{C}\mathbf{y} + \mathbf{d} \mid \forall \mathbf{C} \in \mathbb{R}^{p \times m}, \mathbf{d} \in \mathbb{R}^p \}$$

# Linear Minimum Mean Square Error Estimation

$$\begin{aligned}\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}_{\text{LMMSE}} \perp \text{LF}(\mathbf{y}) &\Leftrightarrow \langle \hat{\boldsymbol{\theta}}, \boldsymbol{\theta} - \hat{\boldsymbol{\theta}}_{\text{LMMSE}} \rangle = 0, \forall \hat{\boldsymbol{\theta}} \in \text{LF}(\mathbf{y}) \\ &\Rightarrow \langle \mathbf{C}\mathbf{y} + \mathbf{d}, \boldsymbol{\theta} - \mathbf{A}\mathbf{y} - \mathbf{b} \rangle = 0, \forall \mathbf{C}, \mathbf{d} \\ &\Rightarrow E_{\boldsymbol{\theta}, \mathbf{y}}((\mathbf{C}\mathbf{y} + \mathbf{d})^T (\boldsymbol{\theta} - \mathbf{A}\mathbf{y} - \mathbf{b})) = 0, \forall \mathbf{C}, \mathbf{d}\end{aligned}$$

- Let  $\mathbf{C} = \mathbf{0}$ ,  $\mathbf{m}_{\boldsymbol{\theta}} = E(\boldsymbol{\theta})$ , and  $\mathbf{m}_{\mathbf{y}} = E(\mathbf{y})$

$$\begin{aligned}E_{\boldsymbol{\theta}, \mathbf{y}}(\mathbf{d}^T (\boldsymbol{\theta} - \mathbf{A}\mathbf{y} - \mathbf{b})) &= 0 \\ \Rightarrow \mathbf{d}^T (\mathbf{m}_{\boldsymbol{\theta}} - \mathbf{A}\mathbf{m}_{\mathbf{y}} - \mathbf{b}) &= 0\end{aligned}$$

- Now we set  $\mathbf{d} = \mathbf{m}_{\boldsymbol{\theta}} - \mathbf{A}\mathbf{m}_{\mathbf{y}} - \mathbf{b}$

$$\begin{aligned}\Rightarrow \|\mathbf{m}_{\boldsymbol{\theta}} - \mathbf{A}\mathbf{m}_{\mathbf{y}} - \mathbf{b}\|^2 &= 0 \\ \Rightarrow \mathbf{b} &= \mathbf{m}_{\boldsymbol{\theta}} - \mathbf{A}\mathbf{m}_{\mathbf{y}}\end{aligned}$$

- Hence

$$\hat{\boldsymbol{\theta}}_{\text{LMMSE}} = \mathbf{m}_{\boldsymbol{\theta}} + \mathbf{A}(\mathbf{y} - \mathbf{m}_{\mathbf{y}}) \tag{1}$$

# Linear Minimum Mean Square Error Estimation

- Let  $\mathbf{d} = \mathbf{0}$

$$\begin{aligned}E_{\theta,y}(\mathbf{y}^T \mathbf{C}^T (\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}_{\text{LMMSE}})) &= 0 \\ \Rightarrow E_{\theta,y}(\mathbf{y}^T \mathbf{C}^T (\underbrace{\boldsymbol{\theta} - \mathbf{m}_{\theta}}_{\Delta \boldsymbol{\theta}} - \mathbf{A}(\underbrace{\mathbf{y} - \mathbf{m}_y}_{\Delta \mathbf{y}}))) &= 0 \\ \Rightarrow E_{\theta,y}(\mathbf{y}^T \mathbf{C}^T (\Delta \boldsymbol{\theta} - \mathbf{A} \Delta \mathbf{y})) &= 0 \\ \Rightarrow E_{\theta,y}(\mathbf{y}^T \mathbf{C}^T (\Delta \boldsymbol{\theta} - \mathbf{A} \Delta \mathbf{y})) - \mathbf{m}_y^T \mathbf{C}^T \underbrace{E_{\theta,y}(\Delta \boldsymbol{\theta} - \mathbf{A} \Delta \mathbf{y})}_{=0} &= 0 \\ \Rightarrow E_{\theta,y}(\Delta \mathbf{y}^T \mathbf{C}^T (\Delta \boldsymbol{\theta} - \mathbf{A} \Delta \mathbf{y})) &= 0 \\ \Rightarrow \text{tr}(\mathbf{C}^T E_{\theta,y}((\Delta \boldsymbol{\theta} - \mathbf{A} \Delta \mathbf{y}) \Delta \mathbf{y}^T)) &= 0 \\ \Rightarrow \text{tr}(\mathbf{C}^T (\mathbf{C}_{\theta y} - \mathbf{A} \mathbf{C}_{yy})) &= 0 \quad \forall \mathbf{C}\end{aligned}$$

- For  $\mathbf{C} = \mathbf{C}_{\theta y} - \mathbf{A} \mathbf{C}_{yy}$ , we have  $\|\mathbf{C}_{\theta y} - \mathbf{A} \mathbf{C}_{yy}\|_F^2 = 0$ , i.e.,

$$\mathbf{A} = \mathbf{C}_{\theta y} \mathbf{C}_{yy}^{-1} \quad (2)$$

# Linear Minimum Mean Square Error Estimation

- Combining equation (1) and (2), we have

$$\hat{\boldsymbol{\theta}}_{\text{LMMSE}} = \mathbf{m}_{\boldsymbol{\theta}} + \mathbf{C}_{\boldsymbol{\theta}\mathbf{y}}\mathbf{C}_{\mathbf{y}\mathbf{y}}^{-1}(\mathbf{y} - \mathbf{m}_{\mathbf{y}})$$

- The estimation error can be expressed as

$$\tilde{\boldsymbol{\theta}}_{\text{LMMSE}} = \hat{\boldsymbol{\theta}}_{\text{LMMSE}} - \boldsymbol{\theta} = -\Delta\boldsymbol{\theta} + \mathbf{A}\Delta\mathbf{y}$$

- The covariance of the estimator is obtained as

$$\begin{aligned}\text{Cov}(\hat{\boldsymbol{\theta}}_{\text{LMMSE}}) &= E_{\boldsymbol{\theta},\mathbf{y}}(\tilde{\boldsymbol{\theta}}_{\text{LMMSE}}\tilde{\boldsymbol{\theta}}_{\text{LMMSE}}^T) \\ &= E_{\boldsymbol{\theta},\mathbf{y}}((\mathbf{A}\Delta\mathbf{y} - \Delta\boldsymbol{\theta})(\mathbf{A}\Delta\mathbf{y} - \Delta\boldsymbol{\theta})^T) \\ &= \mathbf{C}_{\boldsymbol{\theta}\boldsymbol{\theta}} - \mathbf{C}_{\boldsymbol{\theta}\mathbf{y}}\mathbf{C}_{\mathbf{y}\mathbf{y}}^{-1}\mathbf{C}_{\mathbf{y}\boldsymbol{\theta}}\end{aligned}$$

- Note that if  $\boldsymbol{\theta}$  and  $\mathbf{y}$  are jointly Gaussian, we get  $\hat{\boldsymbol{\theta}}_{\text{LMMSE}} = \hat{\boldsymbol{\theta}}_{\text{MMSE}}$