

ECE 286: Bayesian Machine Perception

Class 4: The Particle Filter

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Estimation of Expectation and Covariance

- Samples $\mathbf{x}^{(j)} \sim f(\mathbf{x})$, $j = 1, \dots, J$ of continuous random vector \mathbf{x}
- **Expectation** of a random vector \mathbf{x}

$$\mathbb{E}\{\mathbf{x}\} = \int \mathbf{x} f(\mathbf{x}) d\mathbf{x} \approx \frac{1}{J} \sum_{j=1}^J \mathbf{x}^{(j)} = \tilde{\boldsymbol{\mu}}_{\mathbf{x}}$$

- Expectation of transformed random vector $g(\mathbf{x})$

$$\mathbb{E}\{g(\mathbf{x})\} = \int g(\mathbf{x}) f(\mathbf{x}) d\mathbf{x} \approx \frac{1}{J} \sum_{j=1}^J g(\mathbf{x}^{(j)})$$

- **Covariance** of a random vector \mathbf{x}

$$\mathbb{C}\{\mathbf{x}\} = \mathbb{E}\{\mathbf{x}\mathbf{x}^T\} - \mathbb{E}\{\mathbf{x}\}\mathbb{E}\{\mathbf{x}\}^T \approx \frac{1}{J} \sum_{j=1}^J \mathbf{x}^{(j)}\mathbf{x}^{(j)T} - \tilde{\boldsymbol{\mu}}_{\mathbf{x}}\tilde{\boldsymbol{\mu}}_{\mathbf{x}}^T$$

- Approximations are arbitrarily good for J sufficiently large

Importance Sampling (i)

- Recap expectation of transformed random vector

$$\mathbb{E}\{g(\boldsymbol{x})\} = \int g(\boldsymbol{x}) f_t(\boldsymbol{x}) d\boldsymbol{x} = \int g(\boldsymbol{x}) \frac{f_t(\boldsymbol{x})}{f_p(\boldsymbol{x})} f_p(\boldsymbol{x}) d\boldsymbol{x}$$

- **Draw** samples proposal pdf $\boldsymbol{x}^{(j)} \sim f_p(\boldsymbol{x}), j = 1, \dots, J$
- **Calculate** approximate expectation

Note that
 $\sum_{j=1}^J w^{(j)} = 1$
for $J \rightarrow \infty$

$$\mathbb{E}\{g(\boldsymbol{x})\} \approx \frac{1}{J} \sum_{j=1}^J g(\boldsymbol{x}^{(j)}) \frac{f_t(\boldsymbol{x}^{(j)})}{f_p(\boldsymbol{x}^{(j)})} = \sum_{j=1}^J g(\boldsymbol{x}^{(j)}) w^{(j)}$$

where $w^{(j)} = \frac{1}{J} \frac{f_t(\boldsymbol{x}^{(j)})}{f_p(\boldsymbol{x}^{(j)})}$

- This motivates a particle representation $\{\boldsymbol{x}^{(j)}, w^{(j)}\}_{j=1}^J \simeq f_t(\boldsymbol{x})$

Importance Sampling (ii)

- **Goal:** Calculate

$$\{\mathbf{x}^{(j)}, w^{(j)}\}_{j=1}^J \simeq f_t(\mathbf{x})$$

- **Step 1:** Sample

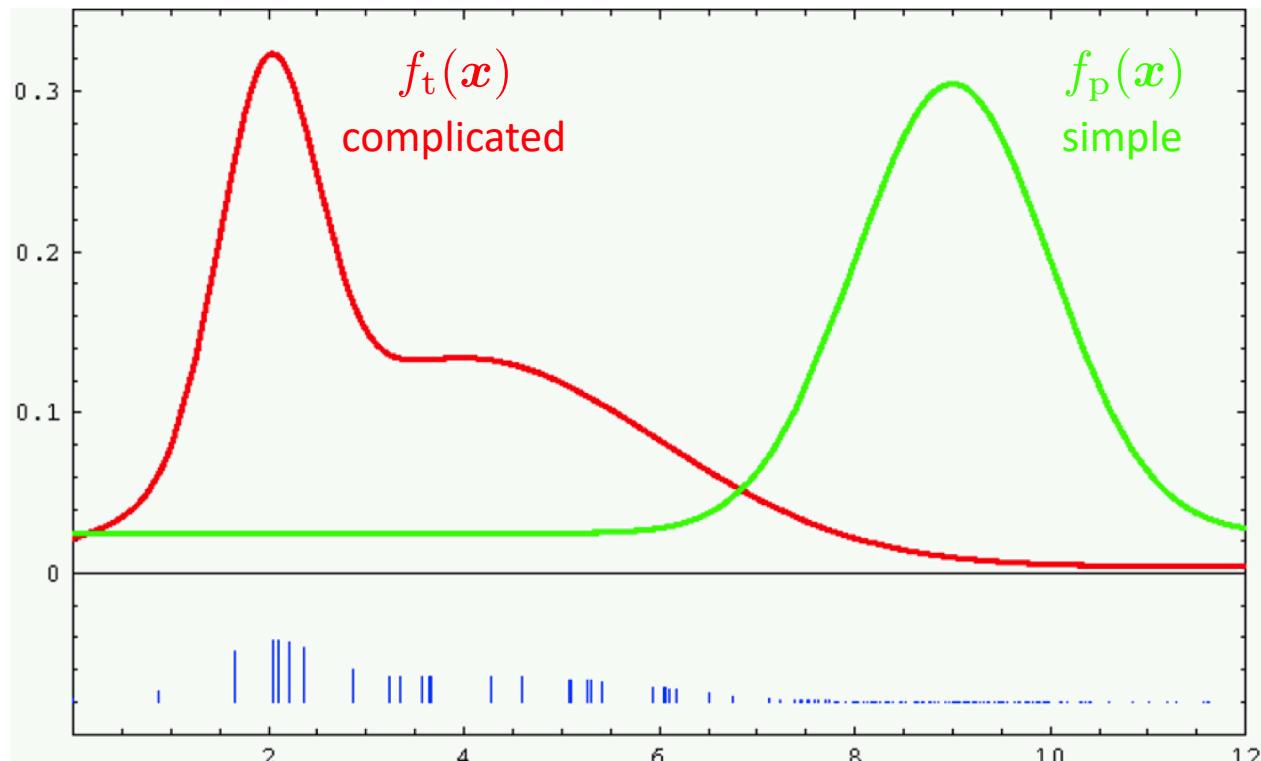
$$\mathbf{x}^{(j)} \sim f_p(\mathbf{x})$$

$$j \in \{1 \dots J\}$$

- **Step 2:** Calculate weights

$$w^{(j)} = \frac{1}{J} \frac{f_t(\mathbf{x}^{(j)})}{f_p(\mathbf{x}^{(j)})}$$

$$j \in \{1 \dots J\}$$



Estimation of Expectation and Covariance

- Particles and weights $\{(\mathbf{x}^{(j)}, w^{(j)})\}_{j=1}^J \simeq f(\mathbf{x})$ representing the PDF of random vector \mathbf{x}
- Expectation of a random vector \mathbf{x}

$$\mathbb{E}\{\mathbf{x}\} \approx \sum_{j=1}^J w^{(j)} \mathbf{x}^{(j)} = \tilde{\boldsymbol{\mu}}_{\mathbf{x}}$$

- Expectation of transformed random vector $g(\mathbf{x})$

$$\mathbb{E}\{g(\mathbf{x})\} \approx \sum_{j=1}^J g(\mathbf{x}^{(j)}) w^{(j)}$$

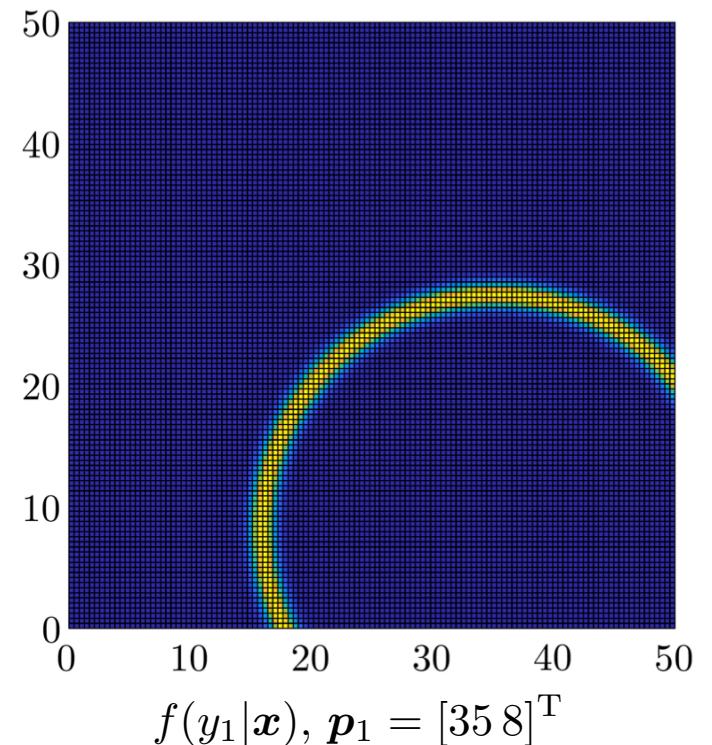
- Covariance of a random vector \mathbf{x}

$$\mathbb{C}\{\mathbf{x}\} \approx \sum_{j=1}^J w^{(j)} \mathbf{x}^{(j)} \mathbf{x}^{(j)\top} - \tilde{\boldsymbol{\mu}}_{\mathbf{x}} \tilde{\boldsymbol{\mu}}_{\mathbf{x}}^{\top}$$

- Approximations are arbitrarily good for J sufficiently large

Example: Fusion of Range Measurements (i)

- \mathbf{x} is the unknown position of an agent
- Measurement $\mathbf{y} = [y_1 \ y_2 \ y_3]^T$ is modeled as
$$y_i = \|\mathbf{x} - \mathbf{p}_i\| + v_i$$
where \mathbf{p}_i is the position of a known reference and $v_i \sim \mathcal{N}(0; \sigma_v)$ is noise (independent across i)
- Each likelihood function $f(y_i | \mathbf{x})$ describes potential agent positions on a ring

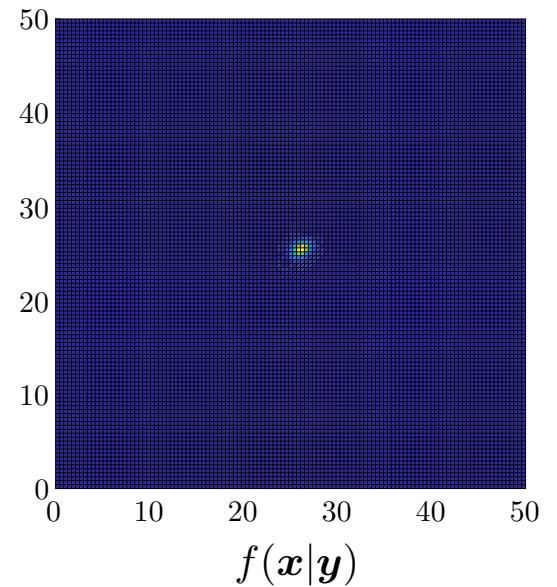
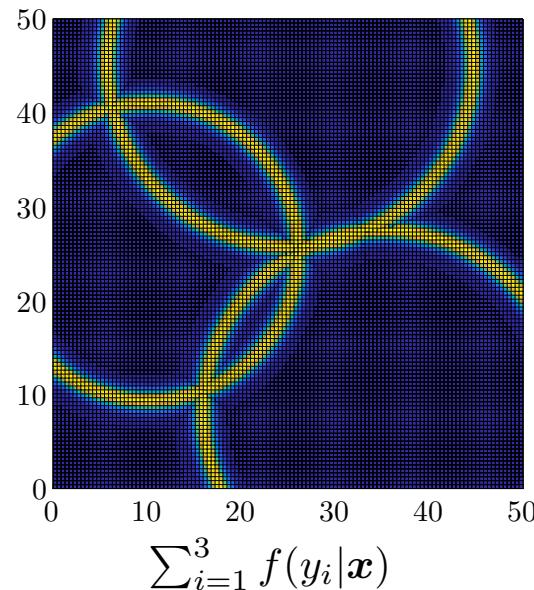
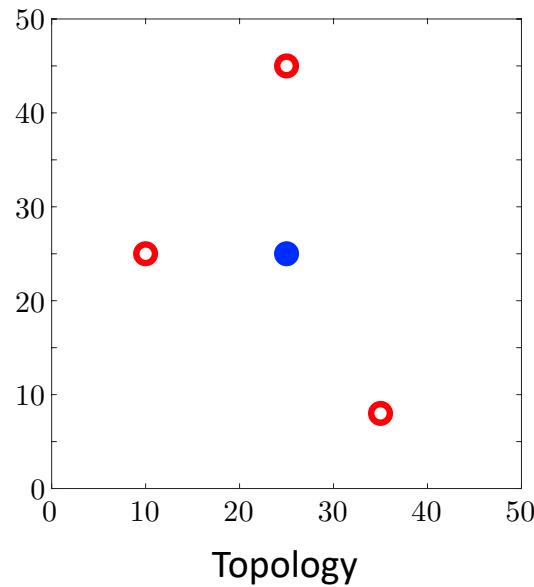


Example: Fusion of Range Measurements (ii)

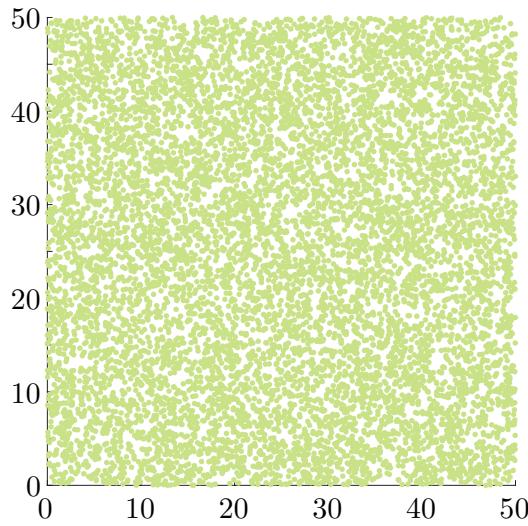
- Marginal posterior PDF (y is observed)

$$f(\mathbf{x}|\mathbf{y}) \propto f(\mathbf{y}|\mathbf{x})f(\mathbf{x}) = \prod_{i=1}^3 f(y_i|\mathbf{x})f(\mathbf{x})$$

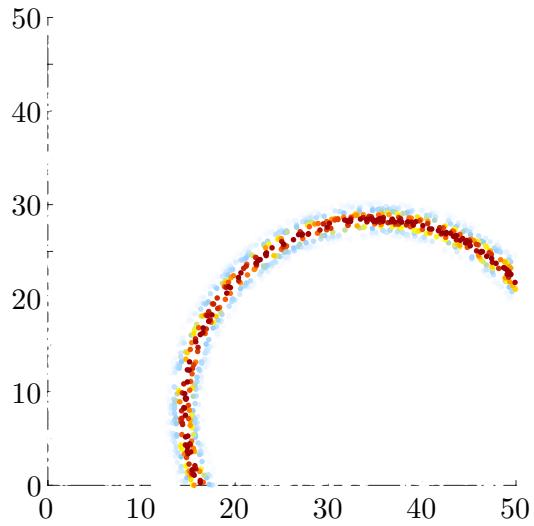
- Example scenario with uniform prior $f(\mathbf{x})$



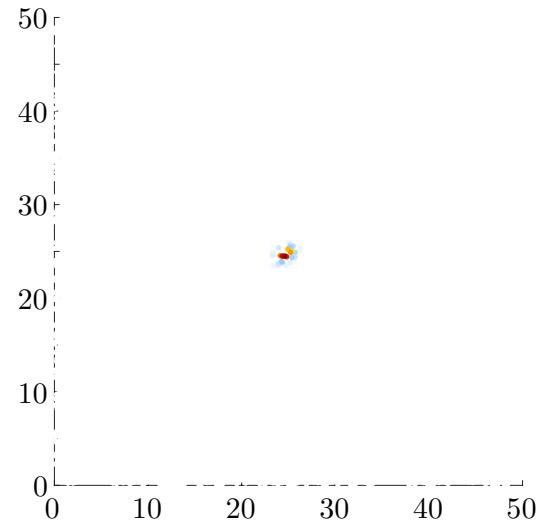
Example: Fusion of Range Measurements (iii)



$$\{\mathbf{x}^j\}_{j=1}^J \sim f(\mathbf{x})$$



$$\{\mathbf{x}_1^j, w_1^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}|y_1)$$



$$\{\mathbf{x}^j, w^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}|\mathbf{y})$$

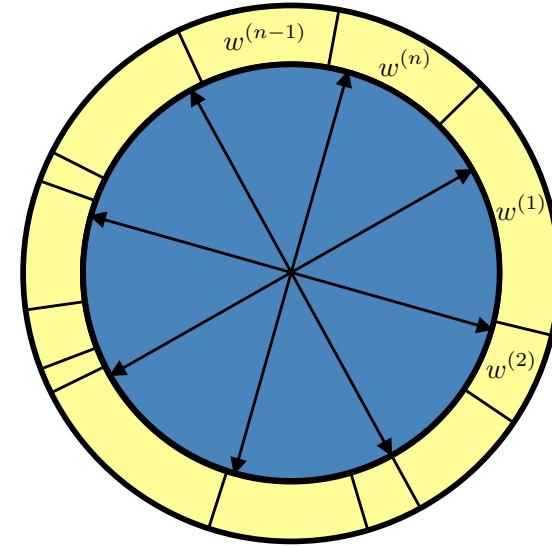
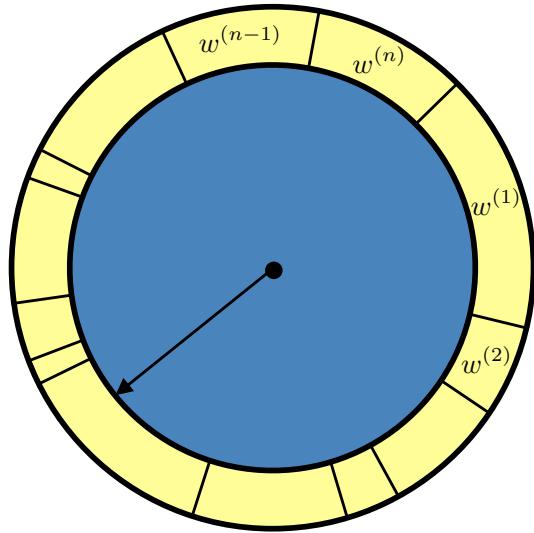
- Approximate MMSE estimate: $\hat{\mathbf{x}}^{\text{MMSE}} = \sum_{j=1}^J \mathbf{x}^{(j)} w^{(j)}$

Particle Resampling

- We can consider a particle representation $\{\boldsymbol{x}^{(j)}, w^{(j)}\}_{j=1}^J \simeq f(\boldsymbol{x})$ as probability mass function and draw J' samples from it
- In this way, we get a new particle representation $\{\bar{\boldsymbol{x}}^{(j)}\}_{j=1}^{J'} \simeq f(\boldsymbol{x})$
- This is motivated by the approximation ($\delta(\cdot)$ is the Dirac delta function)

$$\begin{aligned} f(\boldsymbol{x}) &\approx \sum_{j=1}^J w^{(j)} \delta(\boldsymbol{x} - \boldsymbol{x}^{(j)}) & \{\bar{\boldsymbol{x}}^{(j)}\}_{j=1}^{J'} &\sim \tilde{f}(\boldsymbol{x}) \\ &= \tilde{f}(\boldsymbol{x}) \end{aligned}$$

Particle Resampling



- ``Roulette wheel''
- Binary search, complexity $\mathcal{O}(J \log J)$
- Systematic resampling
- Complexity $\mathcal{O}(J)$
- Easy to implement, low variance

Particle Marginalization

- If we have a particle representation of a PDF $f(\mathbf{x}) = f(x_1, x_2)$, i.e.,

$$\{(\mathbf{x}^{(j)}, w^{(j)})\}_{j=1}^J = \{(x_1^{(j)}, x_2^{(j)}, w^{(j)})\}_{j=1}^J \simeq f(\mathbf{x})$$

we immediately have particle representations of the marginals $f(x_1)$ and $f(x_2)$, i.e.,

$$\{(x_1^{(j)}, w^{(j)})\}_{j=1}^J \simeq f(x_1) \quad \{(x_2^{(j)}, w^{(j)})\}_{j=1}^J \simeq f(x_2)$$

- This is motivated by

$$\begin{aligned} f(x_1) &\approx \int \tilde{f}(\mathbf{x}) dx_2 = \sum_{j=1}^J w^{(j)} \int \delta(\mathbf{x} - \mathbf{x}^{(j)}) dx_2^{(j)} \\ &= \sum_{j=1}^J w^{(j)} \delta(x_1 - x_1^{(j)}) \end{aligned}$$

The Particle Filter

- Recap sequential Bayesian estimation

Prediction Step:

$$\underbrace{f(\mathbf{x}_n | \mathbf{y}_{1:n-1})}_{\text{Predicted posterior pdf}} = \int \underbrace{f(\mathbf{x}_n | \mathbf{x}_{n-1})}_{\text{State-transition pdf}} \underbrace{f(\mathbf{x}_{n-1} | \mathbf{y}_{1:n-1})}_{\text{Previous posterior pdf}} d\mathbf{x}_{n-1}$$

Update Step:

$$\underbrace{f(\mathbf{x}_n | \mathbf{y}_{1:n})}_{\text{Posterior pdf}} \propto \underbrace{f(\mathbf{y}_n | \mathbf{x}_n)}_{\text{Likelihood function}} \underbrace{f(\mathbf{x}_n | \mathbf{y}_{1:n-1})}_{\text{Predicted posterior pdf}}$$

- Approximate implementation by means of importance sampling, particle marginalization, and resampling
- Particle representation of posterior PDFs

$$\{\mathbf{x}_n^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}_n | \mathbf{y}_{1:n-1})$$

$$\{\bar{\mathbf{x}}_n^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}_n | \mathbf{y}_{1:n})$$

B. Ristic, S. Arulampalam, and N. Gordon, *Beyond the Kalman Filter: Particle Filters for Tracking Applications*, Artech House, 2004.

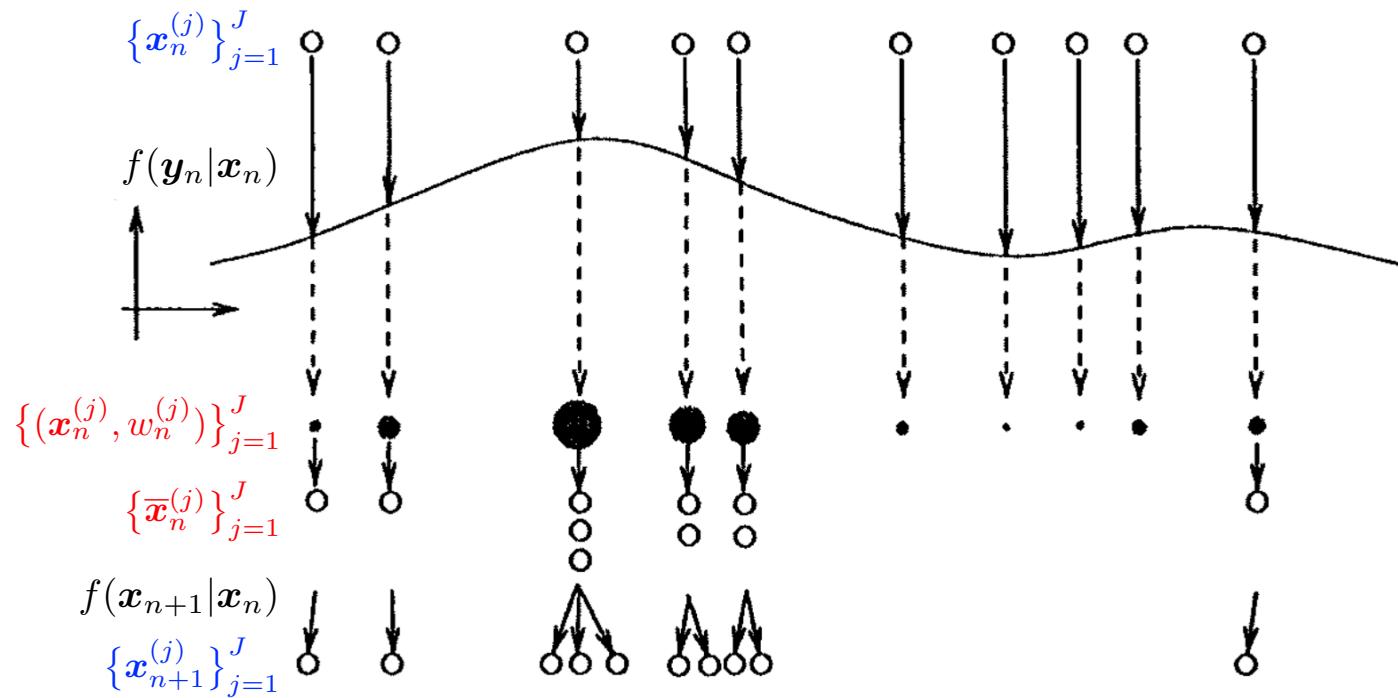
Prediction Step

- **Given:** Particles $\{\bar{x}_{n-1}^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}_{n-1}|\mathbf{y}_{1:n-1})$ representing the **previous posterior PDF**
- **Wanted:** Particles $\{x_n^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}_n|\mathbf{y}_{1:n-1})$ representing the **predicted posterior PDF**
- Calculate a particle representation of $f(\mathbf{x}_n, \mathbf{x}_{n-1}|\mathbf{y}_{1:n-1}) = f(\mathbf{x}_n|\mathbf{x}_{n-1})f(\mathbf{x}_{n-1}|\mathbf{y}_{1:n-1})$ by repeating J times
 - draw sample from driving noise PDF $\mathbf{u}_n^{(j)} \sim f(\mathbf{u}_n)$
 - draw sample from $x_n^{(j)} \sim f(\mathbf{x}_n|\bar{x}_{n-1}^{(j)})$ by applying state transition model $\mathbf{x}_n^{(j)} = g_n(\bar{x}_{n-1}^{(j)}, \mathbf{u}_n^{(j)})$
- Perform particle marginalization to get $\{x_n^{(j)}\}_{j=1}^J \simeq f(\mathbf{x}_n|\mathbf{y}_{1:n-1})$ from $\{(x_n^{(j)}, \bar{x}_{n-1}^{(j)})\}_{j=1}^J \simeq f(\mathbf{x}_n, \mathbf{x}_{n-1}|\mathbf{y}_{1:n-1})$

Update Step

- **Given:** Particles $\{(\boldsymbol{x}_n^{(j)})\}_{j=1}^J \simeq f(\boldsymbol{x}_n | \boldsymbol{y}_{1:n-1})$ representing the **predicted posterior PDF**
- **Wanted:** Particles $\{(\bar{\boldsymbol{x}}_n^{(j)})\}_{j=1}^J \simeq f(\boldsymbol{x}_n | \boldsymbol{y}_{1:n})$ representing the **posterior PDF**
- Perform importance sampling with proposal distribution $f_p(\boldsymbol{x}_n) = f(\boldsymbol{x}_n | \boldsymbol{y}_{1:n-1})$ and target distribution $f_t(\boldsymbol{x}_n) \propto f(\boldsymbol{y}_n | \boldsymbol{x}_n) f(\boldsymbol{x}_n | \boldsymbol{y}_{1:n-1})$
 - calculate unnormalized weights $\tilde{w}_n^{(j)} = f(\boldsymbol{y}_n | \boldsymbol{x}_n^{(j)}) \propto f_t(\boldsymbol{x}_n^{(j)}) / f_p(\boldsymbol{x}_n^{(j)})$
 - normalize weights $w_n^{(j)} = \tilde{w}_n^{(j)} / \sum_{j'=1}^J \tilde{w}_n^{(j')}, \quad j = 1, \dots, J$
- Perform resampling to get $\{(\bar{\boldsymbol{x}}_n^{(j)})\}_{j=1}^J \simeq f(\boldsymbol{x}_n | \boldsymbol{y}_{1:n})$ from $\{(\boldsymbol{x}_n^{(j)}, w_n^{(j)})\}_{j=1}^J \simeq f(\boldsymbol{x}_n | \boldsymbol{y}_{1:n})$

Particle Filter Illustration



B. Ristic, S. Arulampalam, and N. Gordon, *Beyond the Kalman Filter: Particle Filters for Tracking Applications*, Artech House, 2004.

Particle Filter Simulation (i)

- We consider **single object tracking**
 - One mobile object, four static references at known locations
 - The state of the mobile object consists of location and velocity, i.e.,

$$\boldsymbol{x}_n = [x_{1,n} \ x_{2,n} \ \dot{x}_{1,n} \ \dot{x}_{2,n}]^T$$

- **State-transition model** for the mobile object:

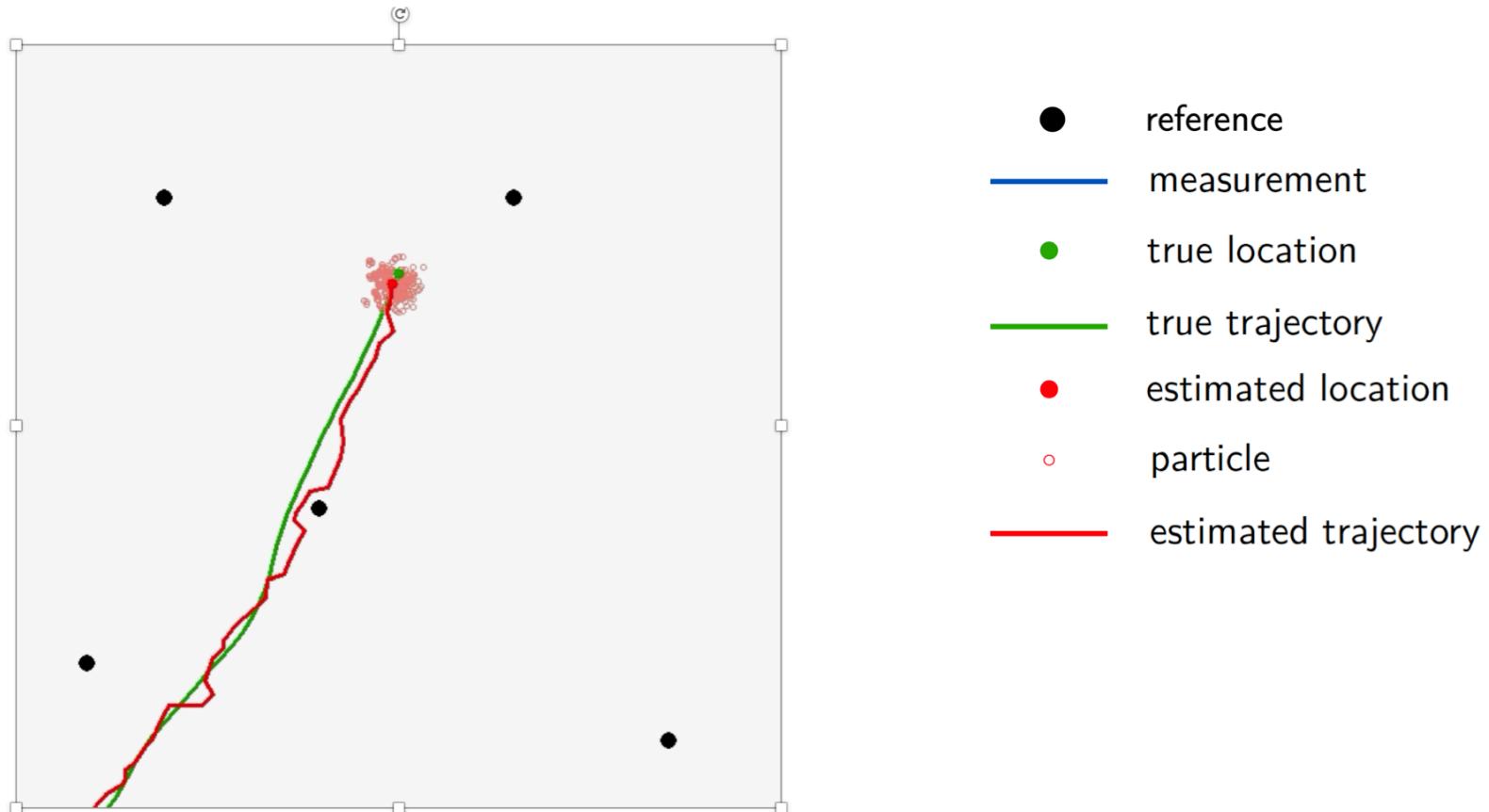
$$\boldsymbol{x}_n = \begin{bmatrix} 1 & 0 & 1 & 0 \\ 0 & 1 & 0 & 1 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \boldsymbol{x}_{n-1} + \begin{bmatrix} 0.5 & 0 \\ 0 & 0.5 \\ 1 & 0 \\ 0 & 1 \end{bmatrix} \boldsymbol{u}_n \quad \text{with } \boldsymbol{u}_n \sim \mathcal{N}(\mathbf{0}, \sigma_u^2 \mathbf{I})$$

- **Measurement model:**

$$y_{n,i} = \|[\boldsymbol{x}_{1,n} \ \boldsymbol{x}_{2,n}]^T - \boldsymbol{p}_k\| + v_{n,i}, \quad i = 1, 2, 3, 4$$

with $v_i \sim \mathcal{N}(0; \sigma_v)$ and known location \boldsymbol{p}_k

Particle Filter Simulation (ii)



Summary

- The Unscented Kalman Filter
 - extension of the Kalman filter to moderately nonlinear systems
 - uses sigma points for linearization
 - is suitable for high-dimensional sequential Bayesian estimation problems
- The Particle Filter
 - represents posterior PDF by random particles
 - provides exact estimation results as the number of particles goes to infinity
 - is suitable for highly nonlinear sequential Bayesian estimation problem
 - is limited to low-dimensional problems ('`curse of dimensionality'')