ECE 286: Bayesian Machine Perception Class 3: The Kalman Filter

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Expectation and Covariance

• Expectation of a random vector $oldsymbol{x}$

discrete case

continuous case

$$\mathbb{E}\{\boldsymbol{x}\} = \sum_{\boldsymbol{x} \in \mathcal{X}} \boldsymbol{x} \, p(\boldsymbol{x})$$

$$\mathbb{E}\{oldsymbol{x}\} = \int oldsymbol{x} f(oldsymbol{x}) \mathrm{d}oldsymbol{x}$$

ullet Expectation of transformed random vector $\,g(oldsymbol{x})\,$

$$\mathbb{E}\{g(\boldsymbol{x})\} = \sum_{\boldsymbol{x} \in \mathcal{X}} g(\boldsymbol{x}) p(\boldsymbol{x}) \qquad \mathbb{E}\{g(\boldsymbol{x})\} = \int g(\boldsymbol{x}) f(\boldsymbol{x}) d\boldsymbol{x}$$

• Covariance of a random vector x

$$\mathbb{C}\{\boldsymbol{x}\} = \mathbb{E}\big\{(\boldsymbol{x} - \mathbb{E}\{\boldsymbol{x}\})(\boldsymbol{x} - \mathbb{E}\{\boldsymbol{x}\})^{\mathrm{T}}\big\} = \mathbb{E}\big\{\boldsymbol{x}\boldsymbol{x}^{\mathrm{T}}\big\} - \mathbb{E}\{\boldsymbol{x}\}\mathbb{E}\{\boldsymbol{x}\}^{\mathrm{T}}$$

The Gaussian Distribution

• Gaussian distribution of continuous random vector $\ensuremath{m{x}} = [x_1, x_2, \dots, x_I]^{\mathrm{T}}$

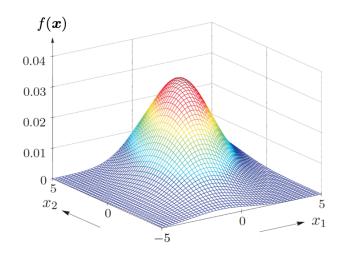
$$f(\boldsymbol{x}) = \det(2\pi \boldsymbol{\Sigma}_{\boldsymbol{x}})^{-\frac{1}{2}} \exp\left(-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu}_{\boldsymbol{x}})^{\mathrm{T}} \boldsymbol{\Sigma}_{\boldsymbol{x}}^{-1}(\boldsymbol{x} - \boldsymbol{\mu}_{\boldsymbol{x}})\right)$$

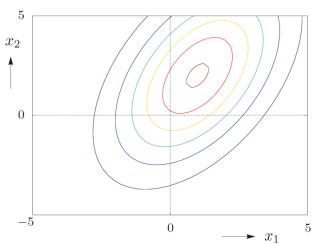
where $oldsymbol{\mu_x} = \mathbb{E}\{oldsymbol{x}\}$ and $oldsymbol{\Sigma_x} = \mathbb{C}\{oldsymbol{x}\}$

• Example (I = 2):

$$oldsymbol{\mu_x} = egin{bmatrix} 1 \\ 2 \end{bmatrix}$$

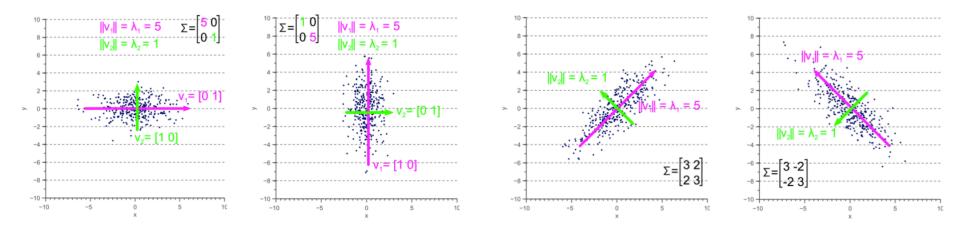
$$\boldsymbol{\Sigma_x} = \begin{bmatrix} 4 & 3 \\ 3 & 9 \end{bmatrix}$$





Geometric Interpretation of the Covariance Matrix

- Variances $[\sigma_{x_1}^2 \, \sigma_{x_2}^2 \dots \sigma_{x_I}^2]^{\mathrm{T}} = \mathrm{diag} \, \boldsymbol{\varSigma_x}$ represent the spread of \boldsymbol{x} along the axes
- The eigenvalues of Σ_x represent the variance of x along eigenvector directions
- Example ($\mu = [0\ 0]^{\mathrm{T}}$):



V. Spruyt, A geometric interpretation of the covariance matrix, 2014.

Recap: State-Space Model

• Consider a sequence of states $oldsymbol{x}_n$ and a sequence of measurements $oldsymbol{y}_n$

State-Transition Model:

State x_n evolves according to

$$oldsymbol{x}_n = g_n(oldsymbol{x}_{n-1}, oldsymbol{\underline{u}}_n)$$
 driving noise (white)

This determines the joint prior

$$f(\boldsymbol{x}_{0:n}) = f(\boldsymbol{x}_0) \prod_{n'=1}^{n} f(\boldsymbol{x}_{n'} | \boldsymbol{x}_{n'-1})$$

Measurement Model:

Measurement y_n is generated as

$$y_n = h_n(x_n, v_n)$$
measurement noise (white)

This determines the joint likelihood

$$f(y_{1:n}|x_{1:n}) = \prod_{n'=1}^{n} f(y_{n'}|x_{n'})$$

• By using Bayes' rule we obtain the marginal posterior pdf (for $oldsymbol{z}_{1:n}$ fixed)

$$f(m{x}_{0:n}|m{z}_{1:n}) \propto f(m{x}_{0:n})f(m{z}_{1:n}|m{x}_{1:n}) = f(m{x}_0)\prod_{n'=1}^n f(m{x}_{n'}|m{x}_{n'-1})f(m{z}_{n'}|m{x}_{n'})$$

Linear-Gaussian State-Space Model

• Consider a sequence of states $oldsymbol{x}_n$ and a sequence of measurements $oldsymbol{y}_n$

State-Transition Model:

State \boldsymbol{x}_n evolves according to

$$oldsymbol{x}_n = oldsymbol{G}_n \, oldsymbol{x}_{n-1} + oldsymbol{u}_n$$
 driving noise (white)

with Gaussian driving noise

$$oldsymbol{u}_n \sim \mathcal{N}(oldsymbol{0}, oldsymbol{arSigma}_{oldsymbol{u}_n})$$

Measurement Model:

Measurement y_n is generated as

$$oldsymbol{y}_n = oldsymbol{H}_n oldsymbol{x}_n + oldsymbol{v}_n$$
measurement noise (white)

with Gaussian measurement noise

$$oldsymbol{v}_n \sim \mathcal{N}(oldsymbol{0}, oldsymbol{arSigma}_{oldsymbol{v}_n})$$

ullet Prior PDF at n=0 , $oldsymbol{x}_0 \sim \mathcal{N}(oldsymbol{\mu}_{oldsymbol{x}_0}, oldsymbol{\Sigma}_{oldsymbol{x}_0})$

Kalman Prediction Step

Recall prediction step of sequential Bayesian estimation

$$\underbrace{f(\boldsymbol{x}_{n}|\boldsymbol{y}_{1:n-1})}_{\text{Predicted posterior pdf}} = \underbrace{\int \underbrace{f(\boldsymbol{x}_{n}|\boldsymbol{x}_{n-1})}_{\text{State-transition}} \underbrace{f(\boldsymbol{x}_{n-1}|\boldsymbol{y}_{1:n-1})}_{\text{Previous posterior pdf}} d\boldsymbol{x}_{n-1}$$

- $f(m{x}_{n-1}|m{y}_{1:n-1})$ is Gaussian with mean $m{\mu}_{m{x}_{n-1}}$ and covariance $m{\Sigma}_{m{x}_{n-1}}$
- $f(m{x}_n|m{y}_{1:n-1})$ is Gaussian with mean $m{\mu}_{m{x}_n}^-$ and covariance $m{\Sigma}_{m{x}_n}^-$ given as

$$\boldsymbol{\mu}_{\boldsymbol{x}_n}^- = G_n \boldsymbol{\mu}_{\boldsymbol{x}_{n-1}}$$

$$oldsymbol{\Sigma}_{oldsymbol{x}_n}^- = oldsymbol{G}_n oldsymbol{\Sigma}_{oldsymbol{x}_{n-1}} oldsymbol{G}_n^{ ext{T}} + oldsymbol{\Sigma}_{oldsymbol{u}_n}$$

Kalman Update Step

• Recall measurement update step of sequential Bayesian estimation

$$rac{f(oldsymbol{x}_n|oldsymbol{y}_{1:n})}{ ext{Posterior pdf}} \propto \underbrace{f(oldsymbol{y}_n|oldsymbol{x}_n)}_{ ext{Likelihood function}} \underbrace{f(oldsymbol{x}_n|oldsymbol{y}_{1:n-1})}_{ ext{Predicted posterior pdf}}$$

- $f(m{x}_n|m{y}_{1:n-1})$ is Gaussian with mean $m{\mu}_{m{x}_n}^-$ and covariance $m{\Sigma}_{m{x}_n}^-$
- $f(x_n|y_{1:n})$ is Gaussian with mean μ_{x_n} and covariance Σ_{x_n} (both can be calculated in closed form)

Kalman Update Step

• Kalman gain

$$oldsymbol{K}_n = oldsymbol{\Sigma}_{oldsymbol{x}_n}^- oldsymbol{H}_n^{\mathrm{T}} ig(oldsymbol{H}_n oldsymbol{\Sigma}_{oldsymbol{x}_n}^- oldsymbol{H}_n^{\mathrm{T}} + oldsymbol{\Sigma}_{oldsymbol{v}_n} ig)^{-1}$$

• Mean and covariance update

$$egin{aligned} oldsymbol{\mu_{x_n}} &= oldsymbol{\mu_{x_n}}^- + K_nig(y_n - oldsymbol{H_noldsymbol{\mu_{x_n}}}^-ig) \ oldsymbol{\Sigma_{x_n}} &= oldsymbol{\Sigma_{x_n}}^- - K_noldsymbol{H_noldsymbol{\Sigma_{x_n}}}^- \end{aligned}$$

Kalman Filter Properties (i)

- Kalman filter provides a procedure for calculating the entire posterior PDF $f(x_n|y_{1:n})$ of linear-Gaussian sequential Bayesian estimation problems
- The covariance matrix Σ_{x_n} does not depend on the measurements $y_{1:n}$ and can thus be calculated offline
- Since for Gaussian distributions the mean is equal to the maximum (or mode), μ_{x_n} is the optimum MMSE estimate and MAP estimate, i.e.,

$$\hat{oldsymbol{x}}_n^{ ext{MMSE}} = \hat{oldsymbol{x}}_n^{ ext{MAP}} = oldsymbol{\mu}_{oldsymbol{x}_n}$$

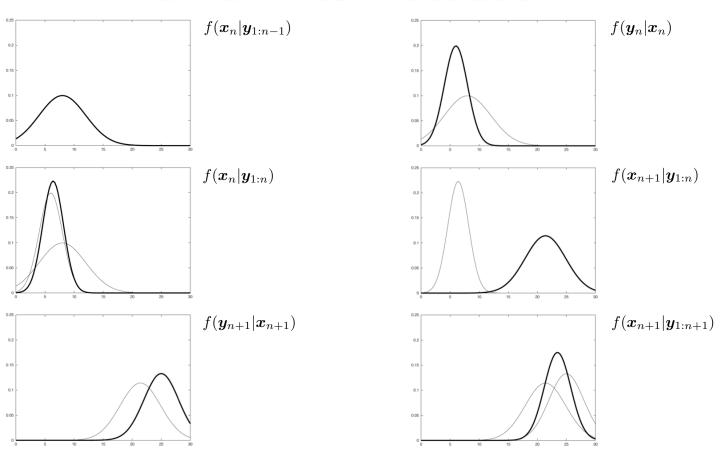
and Σ_{x_n} is the error covariance matrix of the estimate

Kalman Filter Properties (ii)

- In case $f(u_{n'})$ and $f(v_{n'}), n' = 0, 1, ..., n$ or $f(x_0)$ are not Gaussian distributions, it can be shown that the Kalman filter is the best linear MMSE estimator for x_n
- ullet The complexity of the Kalman filter is quite moderate, for I the dimension of the state space and L the dimension of the measurement space
 - it scales as $\mathcal{O}(I^2)$ due matrix multiplication in $\Sigma_{x_n} = (I K_n H_n) \Sigma_{x_n}^-$
 - it scales as $\mathcal{O}(L^{2.4})$ due matrix inversion in $m{K}_n = m{\Sigma}_{m{x}_n}^- m{H}_n^{\mathrm{T}} m{(\Sigma_{m{z}_n})}^{-1}$

J. Anderson and B. Moore, Optimal Filtering, Prentice-Hall, 1979.

Kalman Filter Illustration



S. Thrun, W. Burgard, and D. Fox, *Probabilistic Robotics*, MIT Press, 2006.