

Relatedness analysis (allele sharing)

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Motivation

The detection of (closely) related individuals in genetic studies is of interest in various contexts.

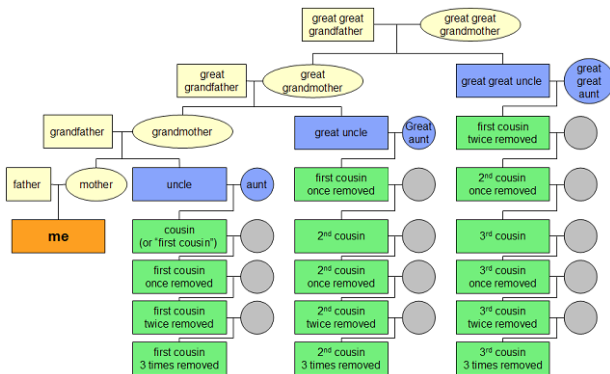
- In association studies, many methods assume independent individuals. Closely related individuals will not be independent.
- In conservation genetics, breeding programs are set up for preferably unrelated individuals.
- In quality control of the data, samples can be accidentally duplicated, and it is of interest to detect it.
- In paternity testing.
- In forensic genetics, e.g. identification of remains.
- To verify documented family relationships.
- To uncover cryptic relatedness.
- ...

Close and remote relatedness

- A distinction is generally drawn between
 - **Close or recent relatedness**: family relationships (MZ, PO, FS, HS, AV, FC, ...)
 - **Distant or remote relatedness**: population substructure (non-homogeneous genetic data)
- Here we mostly address **recent relatedness**
- Focus is on 1° and 2° relationships:

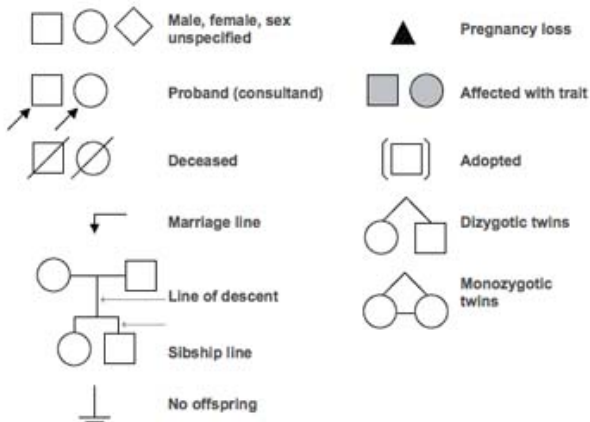
1°	2°
MonoZygotic twins (MZ)	Half Sibs (HS)
Full Sibs (FS)	Avuncular (AV)
Parent-Offspring (PO)	Grandparent-Grandchild (GG)

Close relatedness: family relationships



Close relatedness: family data and pedigrees

Standard Pedigree Nomenclature

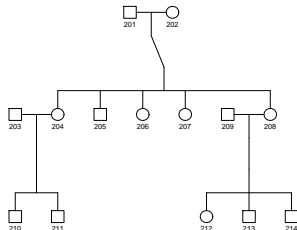


Coding of family data

- A database of related individuals is typically coded in **.ped file format**.
- Besides the genotype information, Family ID, Sample ID, Paternal ID, Maternal ID, Sex (1=male; 2=female; other=unknown) and Affection status (1=affected; 0=unaffected) are registered.

Example:

Family id	Sample id	Father	Mother	Sex	Affected
2	201	0	0	1	1
2	202	0	0	2	NA
2	203	0	0	1	1
2	204	201	202	2	0
2	205	201	202	1	NA
2	206	201	202	2	1
2	207	201	202	2	1
2	208	201	202	2	0
2	209	0	0	1	0
2	210	203	204	1	0
2	211	203	204	1	0
2	212	209	208	2	0
2	213	209	208	1	0
2	214	209	208	1	1



Allele Sharing

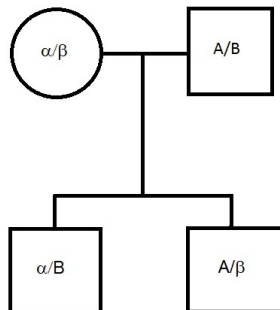
Much of relatedness research is based on the principle of [allele sharing](#)

- For diploid individuals, a pair of individuals can share 0, 1 or 2 alleles for a certain locus.
- The degree to which individuals share alleles indicates the extent to which they are related.

IBD and IBS

- A pair of alleles can be **identical by state (IBS)** or **identical by descent (IBD)**
- IBS alleles simply match irrespective of their provenance
- IBD alleles match because of a **common ancestor**.
- IBD implies IBS but not the reverse.

IBD and IBS



2 alleles IBS but 0 alleles IBD

IBS alleles

- For any locus, we can record for a pair of individuals how many alleles are IBS (how many alleles "match") and this can be 0, 1 or 2.
- E.g., for an A/T single nucleotide polymorphism (SNP):

	AA	AT	TT
AA	2	1	0
AT	1	1	1
TT	0	1	2

- The number of IBS alleles can be recorded for many loci, and averaged over loci.
- An average of 2 would mean that the two individuals are identical (monozygotic twins) or that a sample has been accidentally duplicated.
- This principle can be used to uncover [closely related individuals](#), or to detect [sample heterogeneity](#) (individuals from different populations).

Allele sharing

Allele sharing statistics are often graphed in one of the following ways:

- By plotting means (m) and standard deviation (s) of IBS statistics: (m, s) plot
- By plotting percentages of markers with 0, 1 or 2 IBS alleles: (p_0, p_2) plot
- By plotting estimates of IBD probabilities with 0, 1 or 2 IBS alleles: (k_0, k_1) plot

Notes:

- The (p_0, p_2) plot and (k_0, k_1) plot leave out one of the three proportions. The three proportions can be explicitly visualized simultaneously in a [ternary diagram](#)
- Variants with multiple alleles (e.g. microsatellites) are more informative for discriminating relationship categories than bi-allelic variants (SNP data).
- High MAF variants are more informative for discriminating relationship categories.

(m, s) plot (Abecasis et al, 2001)

- Let

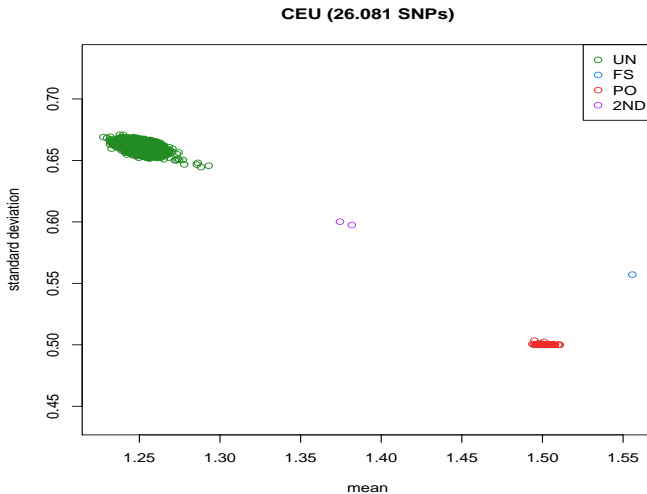
x_{ijk} = number of shared alleles between individual i and j for marker k (0,1,2)

- Compute:

$$m_{ij} = \frac{1}{K} \sum_{k=1}^K x_{ijk} \text{ and } s_{ij}^2 = \frac{1}{K-1} \sum_{k=1}^K (x_{ijk} - m_{ij})^2$$

- Plot m_{ij} against s_{ij} .
- This plot reveals characteristic **clusters** that correspond to the different family relationships.
- Precise position of the different clusters depends on the distribution of the allele frequencies.

Example: CEU sample from the 1000G project ($n = 165$, $p = 26.081$ pruned highly variable SNPs)



(p_0, p_2) plot (Rosenberg et al, 2001)

- Compute for each pair ij

$$p_0 = \frac{1}{K} \sum_{k=1}^K l_{(x_{ijk}=0)}$$

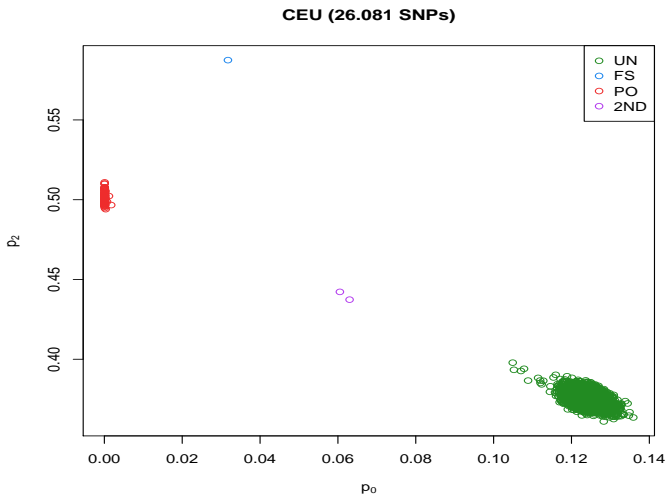
$$p_1 = \frac{1}{K} \sum_{k=1}^K l_{(x_{ijk}=1)}$$

$$p_2 = \frac{1}{K} \sum_{k=1}^K l_{(x_{ijk}=2)}$$

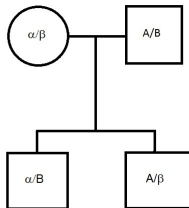
- Plot p_0 against p_2 (or other combinations).
- This plot also reveals **clusters** that correspond to the different family relationships.
- (p_0, p_2) and (m, s) are mathematically related:

$$m = 1 - p_0 + p_2 \quad \text{and} \quad s = \sqrt{p_0(1 - p_0) + p_2(1 - p_2) + 2p_0p_2}$$

Example: CEU sample from the 1000G project ($n = 165$, $p = 26.081$ pruned highly variable SNPs)



IBD probabilities for a given relationship



	α/A	α/B	β/A	β/B
α/A	2	1	1	0
α/B	1	2	0	1
β/A	1	0	2	1
β/B	0	1	1	2

Cotterman coefficients:

$$k_0 = P(\#IBD = 0|FS) = 0.25$$

$$k_1 = P(\#IBD = 1|FS) = 0.50$$

$$k_2 = P(\#IBD = 2|FS) = 0.25$$

Cotterman coefficients

Identity-by-descent probabilities for some standard relationships:

Relationship	k_0	k_1	k_2	θ
MZ	0	0	1	$\frac{1}{2}$
PO	0	1	0	$\frac{1}{4}$
FS	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$
HS	$\frac{1}{2}$	$\frac{1}{2}$	0	$\frac{1}{8}$
AV	$\frac{1}{2}$	$\frac{1}{2}$	0	$\frac{1}{8}$
GG	$\frac{1}{2}$	$\frac{1}{2}$	0	$\frac{1}{8}$
UN	1	0	0	0

- **Kinship** or **coancestry coefficient**: $\theta = \frac{1}{4}k_1 + \frac{1}{2}k_2$
- Probability that two alleles at a locus, one taken at random from two individuals, are identical-by-descent.

Procedure

- IBD probabilities can be estimated from the genotype data by maximum likelihood (Thompson, 1975)
- If the estimated probabilities are "close" to one of the standard relationships, then we infer that particular relationship.
- The **inferred relationship** may (or not) differ from the **putative relationship**.

ML approach

- Let G_1 and G_2 be the pair of genotypes observed at a locus for two individuals.
- Let m (0, 1 or 2) represent the number of IBD alleles.
- By the law of total probability:

$$P(G_1 \cap G_2 | k_0, k_1, k_2) = P(G_1 \cap G_2 | m = 0) k_0 + P(G_1 \cap G_2 | m = 1) k_1 + P(G_1 \cap G_2 | m = 2) k_2$$

- The probabilities $P(G_1 \cap G_2 | m = 0)$ depend on the genotypes of the individuals and are calculated from the allele frequencies in the population.

Calculating the joint genotype probabilities

Let $G_1 = i/i$ and $G_2 = i/i$, and let p_i be the i th allele frequency.

$$P(G_1 = i/i \cap G_2 = i/i | m = 0) = P(G_1 = i/i) P(G_2 = i/i) = p_i^2 p_i^2 = p_i^4$$

$$P(G_1 = i/i \cap G_2 = i/i | m = 2) = P(G_1 = i/i) = P(G_2 = i/i) = p_i^2$$

$$P(G_1 = i/i \cap G_2 = i/i | m = 1) = P(G_1 = i/i) P(G_2 = i/i | G_1 = i/i | m = 1) = p_i^2 p_i = p_i^3$$

For all possible genotype pairs

Pair	Shared alleles	$m = 0$	$m = 1$	$m = 2$
$(A_i/A_i, A_i/A_i)$	2	p_i^4	p_i^3	p_i^2
$(A_i/A_i, A_j/A_j)$	0	$p_i^2 p_j^2$		
$(A_i/A_i, A_i/A_j)$	1	$2p_i^3 p_j$	$p_i^2 p_j$	
$(A_i/A_i, A_j/A_m)$	0	$2p_i^2 p_j p_m$		
$(A_i/A_j, A_i/A_j)$	2	$4p_i^2 p_j^2$	$p_i p_j (p_i + p_j)$	$2p_i p_j$
$(A_i/A_j, A_i/A_m)$	1	$4p_i^2 p_j p_m$	$p_i p_j p_m$	
$(A_i/A_j, A_m/A_l)$	0	$4p_i p_j p_m p_l$		

$$P(G_1 \cap G_2 | k_0, k_1, k_2) = d_0 k_0 + d_1 k_1 + d_2 k_2$$

$$L(k_0, k_1, k_2 | G) = \prod_{i=1}^n (d_{0i} k_0 + d_{1i} k_1 + d_{2i} k_2)$$

Assumptions:

- Hardy-Weinberg equilibrium
- Known population allele frequencies
- Independent variants

Example: HapMap Phase III, Mexican population ($n = 86$)

It.	l	\hat{k}_0	\hat{k}_1	\hat{k}_2
1	-9483.1290	0.41422	0.48104	0.10474
2	-9368.1777	0.18452	0.56753	0.24796
3	-9366.4621	0.21746	0.52776	0.25478
4	-9366.4615	0.21697	0.52798	0.25505
5	-9366.4615	0.21697	0.52798	0.25505

ML estimation of IBD probabilities of a FS pair, using 5.000 SNPs, with initial point (0.575,0.400,0.025). Iteration history for the maximization of the log-likelihood (l)

Software for relatedness research

- R-package SNPRelate
- R-package GWASTools
- GRR
- Relpair
- PLINK
- ...

Estimation of IBD probabilities with PLINK

```
#
# convert .ped to .bed and .fam files
#
runstring01 <- paste("plink -file hapmap3_r3_b36_fwd.consensus.qc.poly",
                    " -make-bed -out hapmap",sep="")
system(runstring01)

#
# Select the CEU individuals
#

runstring02 <- "plink -bfile hapmap -keep CEUsubset.txt -make-bed -out CEU"
system(runstring02)

#
# exclude the X chromosome
#

runstring03 <- "plink --bfile CEU --chr 1-22 --make-bed -out CEU2"
system(runstring03)

#
# Selecting complete SNPs with MAF > 0.40 only
#

runstring05 <- "plink --bfile CEU2 --geno 0 --maf 0.40 -make-bed -out CEU3"
system(runstring05)

#
# HWE filter
#

runstring06 <- "plink --bfile CEU3 --hwe 0.05 midp -make-bed -out CEU4"
system(runstring06)
```

Estimation of IBD probabilities with PLINK

```
#
# LD pruning
#

runstring07 <- "plink --bfile CEU4 --indep-pairwise 50 5 0.2 -make-bed -out CEU5"
system(runstring07)

runstring08 <- "plink --bfile CEU5 --extract CEU5.prune.in --make-bed --out CEU6"
system(runstring08)

#
# Calculate IBD probabilities
#

runstring09 <- "plink --bfile CEU6 --genome --genome-full --out CEU7"
system(runstring09)

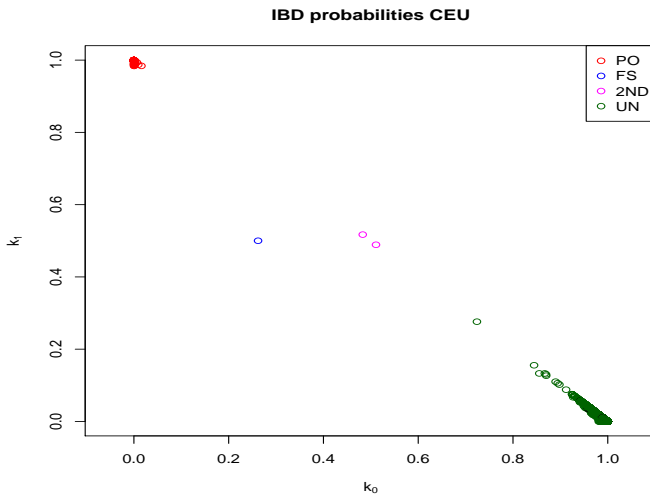
#
# Read the IBD probabilities in R
#

X <- read.table("CEU7.genome",header=TRUE)

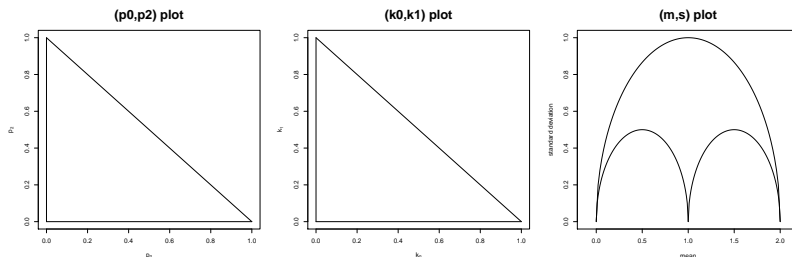
#
# Make a k_0 versus k_1 plot
#

plot(X$Z0,X$Z1,asp=1,xlab=expression(k[0]),ylab=expression(k[1]),main="IBD probabilities CEU")
```

IBD probabilities CEU sample



Restrictions on estimators



All estimators live in a **constrained space**

(p_0, p_1, p_2) is a composition with $p_0 + p_1 + p_2 = 1$

(k_0, k_1, k_2) is a composition with $k_0 + k_1 + k_2 = 1$

Some references

- Abecasis, G.R., Cherny, S.S., Cookson W.O.C. and Cardon, L. R. (2001) GRR: graphical representation of relationship errors. *Bioinformatics*, 17(8) pp. 742–743.
- Graffelman, J., Galván-Femenía, I., De Cid, R., and Barceló-Vidal, C. (2019) A log-ratio biplot approach for exploring genetic relatedness based on identity by state. *Frontiers in Genetics* doi: 10.3389/fgene.2019.00341
- Rosenberg, N. A. (2006). Standardized subsets of the HGDP-CEPH Human Genome Diversity Cell Line Panel, accounting for atypical and duplicated samples and pairs of close relatives. *Annals of Human Genetics*, 70: 841-847.
- Thompson, E.A. (1975). The estimation of pairwise relationships. *Annals of Human Genetics*, 39(2): 173-188.
- Weir, B.S., Anderson, A.D., Hepler, A.B. (2006) Genetic relatedness analysis: modern data and new challenges. *Nature Review Genetics* 7(10) pp. 771–780.

Computer exercise

The file `YRI.raw` contains SNPs of a Yoruba population consisting of parent-offspring trios (2 parents and 1 child). We wish to investigate if the genetic data is consistent with the specified relationships. The PLINK files `YRI.fam`, `YRI.bed` and `YRI.bim` are also available.

- Load the data in `YRI.raw`
- Compute the mean m of the number of alleles shared for each pair of individuals.
- Compute the standard deviation s of the number of alleles shared for each pair of individuals.
- Plot all pairs in a scatterplot of s against m .
- Plot the fraction of variants for which the individuals share 0 alleles against the fraction of variants for which the individuals share 2 alleles, and try to interpret the results.
- Use PLINK to estimate the IBD probabilities, and plot the probabilities of sharing 0 and 1 IBD alleles (k_0 and k_1) for all pairs of individuals.
- Do you think all relationships between all individuals were correctly specified?