Intro to generalized additive models in R (with mgcv)

Matteo Fasiolo (University of Bristol, UK)

matteo.fasiolo@bristol.ac.uk

Material available at:

 $\verb|https://github.com/mfasiolo/GAM_Workshop_Dortmund_25|$

Workshop content

Today's sessions will cover:

- Intro to standard GAMs
- Smooth effects and penalties
- Multi-parameter GAMs, including GAMLSS
- Quantile GAMs and Big Data methods

Focus on **GAM modelling**, not fitting/inferential/computational aspects.

On Github you can find:

- slides
- html files for R demos
- exercises and solutions

We firstly cover:

- What is an additive model?
- Introducing smooth effects
- Oiagnostics and model selection tools
- GAM modelling with mgcv

What is an additive model?

Regression setting:

- y is our response or dependent variable
- x is a vector of covariates or independent variables

In **distributional regression** we want a good model for $Dist(y|\mathbf{x})$.

Model is $\mathrm{Dist}_m\{y|\theta_1(\mathbf{x}),\theta_2,\ldots,\theta_q\}$, where θ_1,\ldots,θ_q are param.

We assume that $\theta_2, \ldots, \theta_q$ do not depend on **x**.

Gaussian additive model:

$$y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma^2\},$$

where

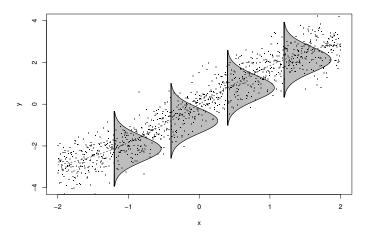
$$\mu(\mathbf{x}) = \mathbb{E}(y|\mathbf{x}) = \sum_{j=1}^{m} f_j(\mathbf{x}),$$

and

$$\sigma^2 = \mathsf{Var}(y|\mathbf{x}).$$

 f_i 's can be fixed, random or smooth effects.

NB: we call $\sum_{i=1}^{m} f_i(\mathbf{x})$ linear predictor because it is linear in β .



Gaussian model with variable mean.
In mgcv: gam(y~s(x), family=gaussian).

Generalized additive model (GAM) (Hastie and Tibshirani, 1990):

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\theta_1 = \mu(\mathbf{x}), \theta_2, \dots, \theta_p\},\$$

where

$$g\{\mu(\mathbf{x})\} = \sum_{j=1}^{m} f_j(\mathbf{x}),$$

and g is a one-to-one link function.

Poisson GAM:

- $y|\mathbf{x} \sim \mathsf{Pois}\{y|\mu(\mathbf{x})\}$
- $\mu(\mathbf{x}) = \exp\left\{\sum_{j=1}^{m} f_j(\mathbf{x})\right\}$
- $g = \log$ assures $\mu(\mathbf{x}) > 0$

Here relation between $\mathbb{E}(y|\mathbf{x})$ and $Var(y|\mathbf{x})$ is implied by model...

... or we can have extra parameters for scale and shape.

Scaled Student's t GAM:

- $y|\mathbf{x} \sim \mathsf{ScaledStud}\{y|\mu(\mathbf{x}), \sigma, \nu\}$ (i.e. $(y-\mu)/\sigma \sim \mathsf{Stud}(y|\nu)$)
- $\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = \sum_{j=1}^{m} f_j(\mathbf{x})$
- ullet σ is scale parameter
- ullet u is shape parameter (degrees of freedom)
- $Var(y|\mathbf{x}) = \sigma^2 \frac{\nu}{\nu 2}$

Later we'll let all parameters be functions of x, eg:

• $y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma(\mathbf{x})\}$

These slides cover:

- What is an additive model?
- Introducing smooth effects
- Oiagnostics and model selection tools
- GAM modelling with mgcv

Introducing smooth effects

Consider additive model

$$g\{\mu(\mathbf{x})\} = f_1(\mathbf{x}) + f_2(\mathbf{x}) + f_3(\mathbf{x}),$$

where

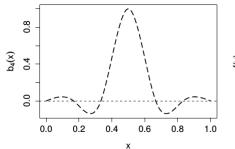
- $f_1(\mathbf{x}) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2$
- $f_2(\mathbf{x}) = \begin{cases} 0 & \text{if } x_2 = \text{FALSE} \\ \beta_4 & \text{if } x_2 = \text{TRUE} \end{cases}$
- $f_3(\mathbf{x}) = f_3(x_3)$ is a non-linear smooth function.

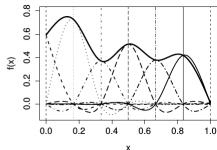
Smooth effects built using spline bases

$$f_3(x_3) = \sum_{k=1}^r \beta_k b_k(x_3)$$

where β_k are unknown coeff and $b_k(x_3)$ are known spline basis functions.

$$s(x, bs = "cr", k = 20)$$

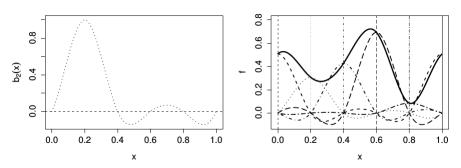




Cubic regression splines are related to the optimal solution to

$$\sum_{i=1}^{n} \{y_i - f(x_i)\}^2 + \lambda \int f''(x)^2 dx.$$

mgcv offers many smooths (see ?smooth.terms). s(x, bs = "cc")



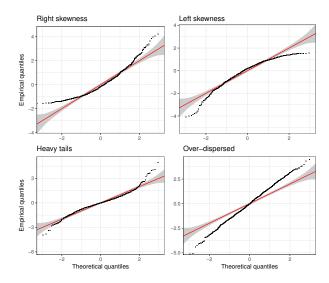
Cyclic cubic regression splines make so that

- $f(x_{min}) = f(x_{max})$
- $f'(x_{min}) = f'(x_{max})$
- $f''(x_{min}) = f''(x_{max})$

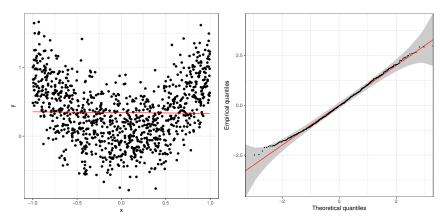
These slides cover:

- What is an additive model?
- Introducing smooth effects
- Oiagnostics and model selection tools
- 4 GAM modelling with mgcv

In the first hands-on session we'll use few basic diagnostics. **QQ-plots**



Useful for choosing model $\operatorname{Dist}_m(y|\mathbf{x})$ (e.g. Poisson vs Neg. Binom.) Less useful for finding omitted variables and non-linearities.



Recall structure of smooth effects:

$$f(\mathbf{x}) = \sum_{j=1}^k \beta_j b_j(\mathbf{x}).$$

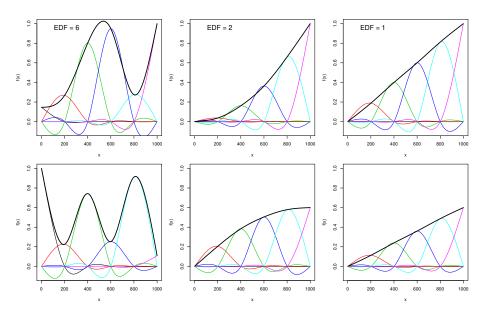
where β estimated by Maximum a Posterior (MAP)

$$\hat{\boldsymbol{\beta}} = \operatorname*{argmax}_{\boldsymbol{\beta}} \ \log p(\boldsymbol{\beta}|\mathbf{y},\boldsymbol{\lambda}) = \operatorname*{argmax}_{\boldsymbol{\beta}} \big\{ \underbrace{\log p(\mathbf{y}|\boldsymbol{\beta})}_{\text{prior penalising complexity}} + \underbrace{\log p(\boldsymbol{\beta}|\boldsymbol{\lambda})}_{\text{prior penalising complexity}} \big]$$

Exact k is unimportant, we choose it large enough and let penalty work.

Effective number of parameters we are using is $\leq k$.

Approximation is **Effective Degrees of Freedom** (EDF) $\leq k$.



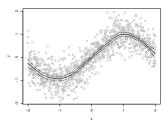
mgcv uses a hierarchical fitting framework:

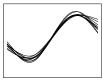
1 Select λ to determine smoothness

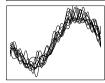
$$\hat{oldsymbol{\lambda}} = \mathop{\mathsf{argmax}}_{oldsymbol{\lambda}} \mathsf{LAML}(oldsymbol{\lambda}).$$

2 For fixed λ , estimate β to determine actual fit

$$\hat{oldsymbol{eta}} = \mathop{\mathsf{argmax}}_{oldsymbol{eta}} \, \log p(oldsymbol{eta} | \mathbf{y}, oldsymbol{\lambda}).$$







By default k = 10 but this is arbitrary.

Exact choice of k not important, but it must not be too low.

Checking whether *k* is too low:

- look at conditional residuals checks
- ② look at output of gam.check(fit):

```
## k' edf k-index p-value

## s(x1) 9.00 8.60 0.91 <2e-16 ***

## s(x2) 9.00 8.13 1.02 0.76

## s(x3) 8.00 2.66 1.04 0.97
```

increase k and see if a model selection criterion improves

Popular criterion is approximate Akaike Information Criterion (AIC):

$$AIC = \underbrace{-2 \log p(\mathbf{y}|\hat{\beta})}_{\text{goodness of fit}} + \underbrace{2\tau}_{\text{model complexity}}$$

where τ is EDF.

If $AIC_{m_1} < AIC_{m_2}$ choose model 1.

To select which effects to include we can also look at p-values:

```
summary(fit)
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 267.2004 75.4197 3.543 0.000405 ***
## x1 6.2854 1.0457 6.010 2.20e-09 ***
## x2 79.8459 80.4130 0.993 0.320858
## x3 -71.2728 86.1725 -0.827 0.408284
```

The exercises will be based on the mgcv package for GAM modelling.

mgcv is a recommended R package, included in R by default.

It contains methods for:

- creating GAM models
- fitting them
- visualizing and summarizing model output

The mgcv ecosystem:

- mgcViz visualising GAMs
- qgam quantile GAMs
- SCM multivariate Gaussian GAMs
- gamFactory nested smooth effects in GAMs
- and many others gamm4, refund, scam, vgam, GJRM, itsadug, ...

There are alternatives to mgcv, such as:

- mboost (Hothorn et al., 2010)
- gamlss (Rigby and Stasinopoulos, 2005)
- brms (Bürkner et al., 2017)
- BayesX (Brezger et al., 2003)
- INLA (Rue et al., 2009)
- bamlss (Umlauf et al., 2018)

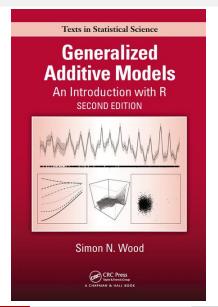
Each offers much flexibility (e.g. smooth effects types and distributions).

Strong points of mgcv's methods:

- little tuning needed (automatic smoothing parameters selection)
- 2 fast and stable numerical implementation

Now we'll be looking at R_demos/1_Intro_Gefcom14.html

Further reading



References I

- Brezger, A., T. Kneib, and S. Lang (2003). Bayesx: Analysing bayesian structured additive regression models. Technical report, Discussion paper//Sonderforschungsbereich 386 der Ludwig-Maximilians
- Bürkner, P. C. et al. (2017). brms: An r package for bayesian multilevel models using stan. *Journal of Statistical Software 80*(1), 1–28.
- Hastie, T. and R. Tibshirani (1990). Generalized Additive Models, Volume 43. CRC Press.
- Hothorn, T., P. Bühlmann, T. Kneib, M. Schmid, and B. Hofner (2010). Model-based boosting 2.0. *The Journal of Machine Learning Research* 11, 2109–2113.
- Rigby, R. A. and D. M. Stasinopoulos (2005). Generalized additive models for location, scale and shape. *Journal of the Royal Statistical Society: Series C (Applied Statistics)* 54(3), 507–554.
- Rue, H., S. Martino, and N. Chopin (2009). Approximate bayesian inference for latent gaussian models by using integrated nested laplace approximations. *Journal of the royal statistical society: Series b (statistical methodology)* 71(2), 319–392.
- Umlauf, N., N. Klein, and A. Zeileis (2018). BAMLSS: Bayesian additive models for location, scale, and shape (and beyond). *Journal of Computational and Graphical Statistics* 27(3), 612–627.