A toolbox of smooth effects and penalties

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Material available at:

 $\verb|https://github.com/mfasiolo/GAM_Workshop_Dortmund_25|$

These slides cover:

More on smoothing penalties

A toolbox of smooth effects

Practical modelling with mgcViz

Recall the GAM model structure:

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\mu(\mathbf{x}), \theta_1, \dots, \theta_p\}$$

where
$$\mu(\mathbf{x}) = \mathbb{E}(y|\mathbf{x}) = g^{-1}\{\sum_{j=1}^m f_j(\mathbf{x})\}.$$

The f_i 's can be

- parametric e.g. $f_j(\mathbf{x}) = \beta_1 x_j + \beta_2 x_j^2$
- random effects
- spline-based smooths such as

$$f_j(x_j) = \sum_{i=1}^r \beta_{ji} b_{ji}(x_j)$$

where β_{ji} are coefficients and $b_{ji}(x_j)$ are known spline basis functions.

NB: we call $\sum_{i=1}^{m} f_i(\mathbf{x})$ linear predictor because it is linear in β .

The $oldsymbol{eta}$ vector is shrunk by penalty

$$\hat{eta} = \operatorname*{argmax}_{eta} \operatorname{PenLogLik}(eta|oldsymbol{\lambda}) = \operatorname*{argmax}_{eta} ig\{ \overbrace{\log p(\mathbf{y}|eta)}^{\operatorname{goodness of fit}} - \underbrace{\operatorname{Pen}(eta|oldsymbol{\lambda})}_{\operatorname{penalize complexity}} ig\}.$$

In mgcv the penalty takes the form

$$\mathsf{Pen}(\boldsymbol{\beta}|\boldsymbol{\lambda}) = \boldsymbol{\beta}^\mathsf{T} \mathbf{S}^{\boldsymbol{\lambda}} \boldsymbol{\beta},$$

where

$$\mathbf{S}^{\lambda} = \sum_{j} \lambda_{j} \mathbf{S}_{j}.$$

and the S_i are positive semi-definite matrices.

NOTE: the penalty is linear w.r.t. λ .

But penalty $\lambda \beta^{\mathsf{T}} \mathbf{S} \beta$ is not very interpretable.

How do we relate it to something like $\lambda \int f''(x)^2 dx$?

Consider a spline-based effect

$$f(x) = \sum_{k} b_{k}(x)\beta_{k} = \mathbf{b}(x)^{\mathsf{T}}\boldsymbol{\beta}.$$

We have that

$$\int f''(x)^2 dx = \int \{\mathbf{b}''(x)^\mathsf{T} \boldsymbol{\beta}\}^2 dx$$

$$= \int (\mathbf{b}''(x)^\mathsf{T} \boldsymbol{\beta}) \mathbf{b}''(x)^\mathsf{T} \boldsymbol{\beta} dx$$

$$= \boldsymbol{\beta}^\mathsf{T} \left\{ \int \mathbf{b}''(x) \mathbf{b}''(x)^\mathsf{T} dx \right\} \boldsymbol{\beta}$$

$$= \boldsymbol{\beta}^\mathsf{T} \mathbf{S} \boldsymbol{\beta}.$$

So, interpretable penalties can be "translated" to $\lambda \beta^T \mathbf{S} \beta$.

Note that $\lambda \int f''(x)^2 dx$ is just one type of penalty.

The type of penalty and type of bases function are often inter-related.

E.g., cubic regression splines (c.r.s.) are related to the optimal solution to

$$\sum_{i=1}^{n} \{y_i - f(x_i)\}^2 + \lambda \int f''(x)^2 dx.$$

So, if you want f(x) to be smooth in terms of f''(x), use c.r.s..

Conversely, if you use c.r.s. it makes sense to penalise f''(x).

But there are cases where you can "mix-and-match" bases and penalties.

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mgcv offers a wide variety of smooths (see ?smooth.terms).

Univariate types:

- s(x, bs = "bs") B-splines regression spline
- s(x, bs = "ad") adaptive smooth
- s(x) = s(x, bs = "tp") thin-plate-splines

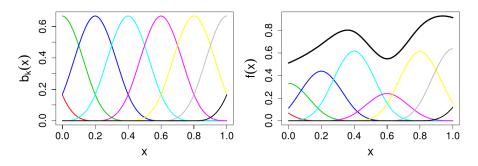
Multivariate type:

- s(x1, x2) = s(x1, x2, bs = "tp") thin-plate-splines
- te(x1, x2) tensor-product-smooth

They can depend on factors:

- s(Subject, bs = "re")
- s(x, by = Subject)
- s(x, Subject, bs = "fs")

B-splines: s(x, bs = "bs", m = c(4, 2))



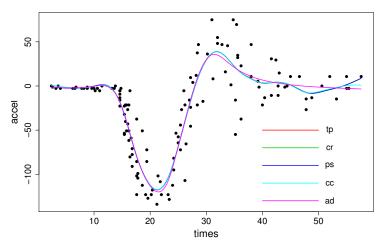
m[1] and m[2] are orders of the spline basis and penalty.

P-splines are B-splines with penalty such as $\sum_{k} (\beta_{k+1} - \beta_k)^2$.

Order of basis and penalty can be different, e.g.:

$$s(x, bs = "ps", m = c(3, 1))$$

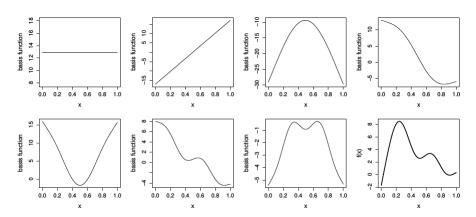
Adaptive P-splines: s(x, bs = "ad")



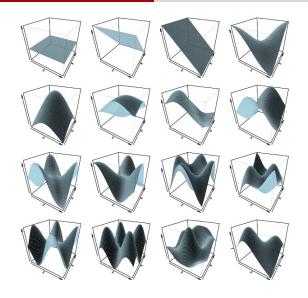
The wiggliness or smoothness of f(x) depends on x.

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Thin plate regression splines (TPRS): s(x)



Rank 7 TPRS basis. Image from Wood (2006).



Rank 17 2D TPRS basis. Courtesy of Simon Wood.

Based on thin plate regression splines basis.

Related to optimal solution to:

$$\sum_{i} \{y_{i} - f(x_{i}, z_{i})\}^{2} + \lambda \int f_{xx}^{2} + 2f_{xz}^{2} + f_{zz}^{2} dx dz$$

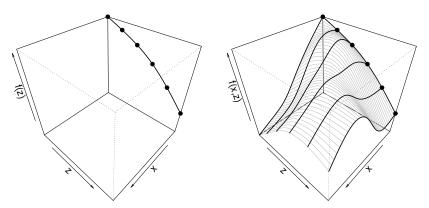
A single smoothing parameter λ .

Isotropic: same smoothness along $x_1, x_2, ...$

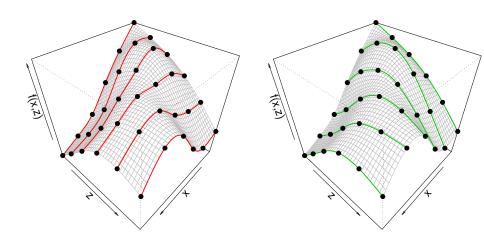
Isotropic effect of x_1 , x_2 are in same unit (e.g. Km).

If different units better use tensor product smooths te(x1, x2).

Construction: make a spline $f_z(z)$ a function of x by letting its coefficients vary smoothly with x



- x-penalty: average wiggliness of red curves
- z-penalty: average wiggliness of green curves



Can use (almost) any kind of marginal:

- te(x1, x2, x3) product of 3 cubic regression splines bases
- te(x1, x2, bs = c("cc", "cr"), k = c(10, 6))
- te(LO, LA, t, d=c(2,1), k=c(20,10), bs=c("tp","cc"))

Basis of te contains functions of the form $f(x_1)$ and $f(x_2)$.

To fit $f(x_1) + f(x_2) + f(x_1, x_2)$ separately use:

$$y \sim ti(x1) + ti(x2) + ti(x1, x2)$$

Random effects

Suppose we have data on bone mineral density (y) as a function of x.

We have m subjects and n data pairs per subject

- subj 1: $\{y_{11}, x_{11}\}, \dots, \{y_{n1}, x_{n1}\}$
- subj j: $\{y_{1j}, x_{1j}\}, \dots, \{y_{nj}, x_{nj}\}$
- subj m: $\{y_{1m}, x_{1m}\}, \dots, \{y_{nm}, x_{nm}\}$

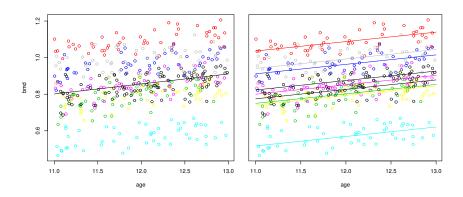
Standard linear model ignores individual differences

$$\mathbb{E}(y|x_{ij}) = \mu(x_{ij}) = \alpha + \beta x_{ij}.$$

We can include random intercept per subject

$$\mu(x_{ij}) = \alpha + \beta x_{ij} + a_j,$$

with penalty $\lambda_a \sum_j a_j^2 = \lambda_a \mathbf{a}^\mathsf{T} \mathbf{a}$ via s(subject, bs = "re")



We can also include random slopes

$$\mu(x_{ij}) = \alpha + (\beta + b_j)x_{ij} + a_j,$$

with penalty $\lambda_b \mathbf{b}^\mathsf{T} \mathbf{b}$ via s(x, subject, bs = "re").

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By-factor smooths

Approach (1) is s(x, by = subject), which means

- $\mu(x) = f_1(x) + ...$ if subject = 1
- $\mu(x) = f_2(x) + ...$ if subject = 2
- ...

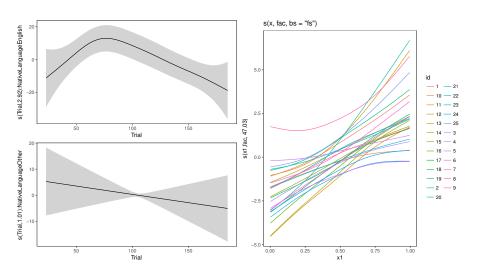
Approach (2) is s(x, subject, bs = "fs"), which means

- $\mu(x) = b_1 + f_1(x) + ...$ if subject = 1
- $\mu(x) = b_2 + f_2(x) + ...$ if subject = 2
- ...

where b_1, b_2, \ldots are random intercepts.

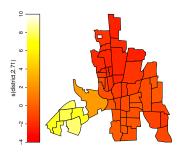
In (1) each f_i has its own smoothing parameter.

In (2) all f_i 's have the same smoothing parameter.



Markov random field effects

Sometimes data come allocated to irregular partitions of space.



- Markov random fields area a popular way of smoothing such data.
- The smooth has a coefficient, β_i , for each region.
- N_i is the set of indices of the neighbours of region i, then penalty is

$$\lambda \sum_{i} (\sum_{j \in N_i} (\beta_i - \beta_j))^2.$$

Some extra smooths available in mgcv:

- smooth on sphere
- soap film smooths
- functional smooths (see ?linear.functional.terms)

There are "standard", i.e. linear w.r.t. β : $f(\mathbf{x}) = \sum_k b_k(\mathbf{x})\beta_k$.

scam package provides shape-constrained effects.

These are non-linear w.r.t. β .

gamFactory provides nested smooth effect

$$f(\mathbf{x}) = s(\tilde{s}(\mathbf{x})),$$

where

- \tilde{s} is a scalar-valued transformation
- s is a spline-based smooth

Simple example is single index effect

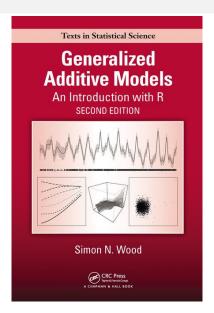
$$f(\mathbf{x}) = s(\alpha^{\mathsf{T}}\mathbf{x}).$$

The effect is non linear w.r.t. α .

If interested try

library(devtools)
install_github("mfasiolo/gamFactory")

Further reading



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Now we'll look at $R_demos/s2_effects_and_mgcViz.html$

References I

- Brezger, A., T. Kneib, and S. Lang (2003). Bayesx: Analysing bayesian structured additive regression models. Technical report, Discussion paper//Sonderforschungsbereich 386 der Ludwig-Maximilians
- Bürkner, P. C. et al. (2017). brms: An r package for bayesian multilevel models using stan. *Journal of Statistical Software 80*(1), 1–28.
- Eilers, P. H. and B. D. Marx (1996). Flexible smoothing with b-splines and penalties. *Statistical science* 11(2), 89–121.
- Fasiolo, M., R. Nedellec, Y. Goude, and S. N. Wood (2020). Scalable visualization methods for modern generalized additive models. *Journal of computational and Graphical Statistics* 29(1), 78–86.
- Hothorn, T., P. Bühlmann, T. Kneib, M. Schmid, and B. Hofner (2010). Model-based boosting 2.0. *The Journal of Machine Learning Research* 11, 2109–2113.
- Pya, N. and S. N. Wood (2015). Shape constrained additive models. *Statistics and computing* 25, 543–559.
- Rigby, R. A. and D. M. Stasinopoulos (2005). Generalized additive models for location, scale and shape. *Journal of the Royal Statistical Society: Series C (Applied Statistics)* 54(3), 507–554.

References II

Rue, H., S. Martino, and N. Chopin (2009). Approximate bayesian inference for latent gaussian models by using integrated nested laplace approximations. *Journal of the royal statistical society: Series b (statistical methodology)* 71(2), 319–392.

Wood, S. (2006). Generalized additive models: an introduction with R. CRC press.