#### Generalized additive models in R

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Material available at:

https://github.com/mfasiolo/workshop\_ENBIS18

# Today's plan

#### First session

- 1 Intro to Generalized Additive Models (GAMs) in R
- Hands-on session

Coffee break

#### Second session

- Beyond mean modelling: GAMLSS and quantile GAMs
- Hands-on session

# Intro to Generalized Additive Models (GAMs)

#### Structure:

- What is an additive model?
- Introducing smooth effects
- 3 Big Data GAM methods
- Diagnostics and model selection tools
- GAM modelling using mgcv and mgcViz

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#### Regression setting:

- y is response or dependent variable
- x is vector of covariates or independent variables

In **distributional regression** we want a good model for Dist(y|x).

Model is  $\mathsf{Dist}_m\{y|\theta_1(\mathbf{x}),\dots,\theta_q(\mathbf{x})\}$ , where  $\theta_1(\mathbf{x}),\dots,\theta_q(\mathbf{x})$  are param.

In a Gaussian model, the mean depends on the covariates

$$y|\mathbf{x} \sim N\{y|\mu = \theta(\mathbf{x}), \sigma^2\},$$

where  $\mu = \mathbb{E}(y|\mathbf{x})$  and  $\sigma^2 = \text{Var}(y)$ .

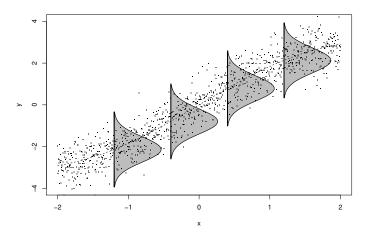


Figure: Gaussian model with variable mean.
In mgcv: gam(y~s(x), family=gaussian).

Generalized additive model (GAM) (Hastie and Tibshirani, 1990):

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\theta_1 = \mu(\mathbf{x}), \theta_2, \dots, \theta_p\},$$

where

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = g^{-1} \Big\{ \sum_{i=1}^{m} f_i(\mathbf{x}) \Big\},\,$$

and g is the link function.

 $f_i$ 's can be fixed, random or smooth effects with coefficients  $\beta$ .

#### Poisson GAM:

- $y|\mathbf{x} \sim \mathsf{Pois}\{y|\mu(\mathbf{x})\}$
- $\mathbb{E}(y|\mathbf{x}) = \mathsf{Var}(y|\mathbf{x}) = \exp\left\{\sum_{j=1}^{m} f_j(\mathbf{x})\right\}$
- $g = \log \text{ assures } \mu(\mathbf{x}) > 0$

Here relation between  $\mathbb{E}(y|\mathbf{x})$  and  $\text{Var}(y|\mathbf{x})$  is implied by model...

... or we can have extra parameters for scale and shape.

Scaled Student's t GAM:

- $y|\mathbf{x} \sim \mathsf{ScaledStud}\{y|\mu(\mathbf{x}), \sigma, \nu\}$
- $\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = \sum_{j=1}^{m} f_j(\mathbf{x})$
- $\bullet$   $\sigma$  is scale parameter
- $\nu$  is shape parameter (degrees of freedom)
- $Var(y|\mathbf{x}) = \sigma^2 \frac{\nu}{\nu 2}$

Later we'll see models with multiple linear predictors, eg:

•  $y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma(\mathbf{x})\}$ 

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### Introducing smooth effects

Consider additive model

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = g^{-1}\Big\{f_1(\mathbf{x}) + f_2(\mathbf{x}) + \cdots\Big\},\,$$

where

- $f_1(\mathbf{x}) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2$
- $f_2(\mathbf{x}) = f_2(x_2)$  is a non-linear smooth function

Smooth effects built using spline bases

$$f_2(x_2) = \sum_{k=1}^r \beta_k b_k(x_2)$$

where  $\beta_k$  are unknown coeff and  $b_k(x_2)$  are known spline basis functions.

NB: we call  $\sum_{i=1}^{m} f_i(\mathbf{x})$  linear predictor because it is linear in  $\beta$ .

# Introducing smooth effects

#### **B-splines**:

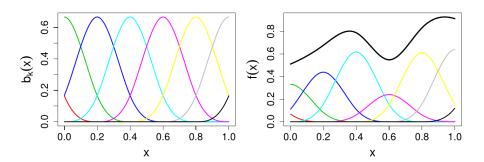
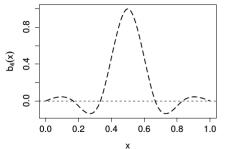
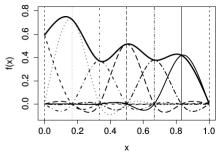


Figure: B-spline basis (left) and smooth (right).

$$s(x, bs = "cr", k = 20)$$

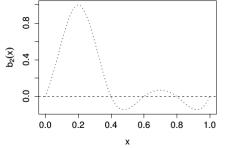


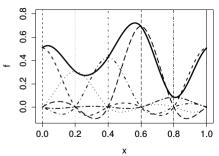


Cubic regression splines are related to the optimal solution to

$$\sum_{i=1}^{n} \{y_i - f(x_i)\}^2 + \gamma \int f''(x)^2 dx.$$

$$s(x, bs = "cc")$$

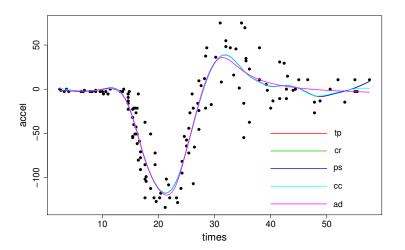




Cyclic cubic regression splines make so that

- $f(x_{min}) = f(x_{max})$
- $f'(x_{min}) = f'(x_{max})$

$$s(x, bs = "ad")$$



The wiggliness or smoothness of f(x) depends on x.

Based on thin plate regression splines basis (Wood, 2006).

Related to optimal solution to:

$$\sum_{i} \{y_i - f(x_i, z_i)\}^2 + \gamma \int f_{xx}^2 + 2f_{xz}^2 + f_{zz}^2 dx dz$$

A single smoothing parameter  $\gamma$ .

Isotropic: same smoothness along  $x_1, x_2, ...$ 

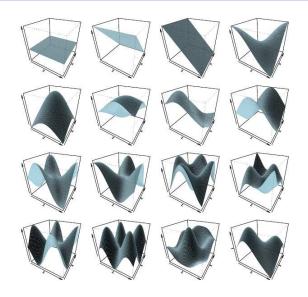
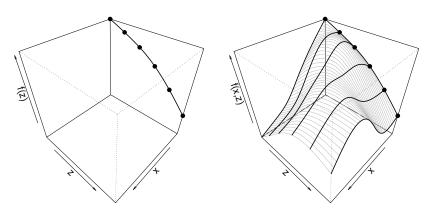


Figure: Rank 15 2D TPRS basis. Courtesy of Simon Wood.

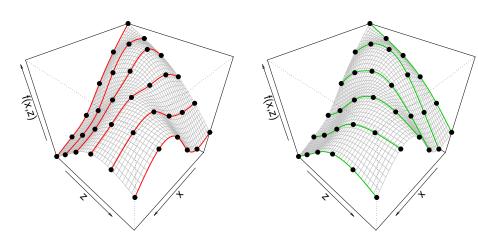
Isotropic effects make sense if  $x_1$ ,  $x_2$  are in same unit (e.g. Km).

If different units better use tensor product smooths te(x1, x2).

Construction: make a spline  $f_z(z)$  a function of x by letting its coefficients vary smoothly with x



- x-penalty: average wiggliness of red curves
- z-penalty: average wiggliness of green curves



Can use (almost) any kind of marginal:

- te(x1, x2, x3) product of 3 cubic regression splines bases
- te(x1, x2, bs = c("cc", "cr"), k = c(10, 6))
- te(LO, LA, t, d=c(2,1), k=c(20,10), bs=c("tp","cc"))

Basis of te contains functions of the form  $f(x_1)$  and  $f(x_2)$ .

To fit  $f(x_1) + f(x_2) + f(x_1, x_2)$  separately use:

$$y \sim ti(x1) + ti(x2) + ti(x1, x2)$$

#### By-factor smooths

Approach (1) is s(x, by = subject), which means

- $\mu(x) = f_1(x) + ...$  if subject = 1
- $\mu(x) = f_2(x) + ...$  if subject = 2
- ...

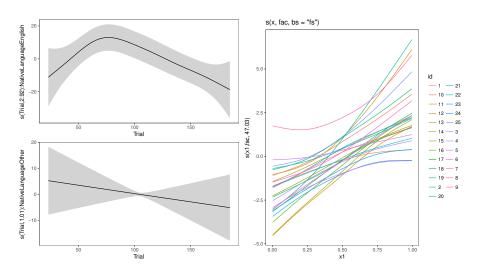
Approach (2) is s(x, subject, bs = "fs"), which means

- $\mu(x) = b_1 + f_1(x) + ...$  if subject = 1
- $\mu(x) = b_2 + f_2(x) + ...$  if subject = 2
- ...

where  $b_1, b_2, \dots \sim N(0, \gamma_{\mathbf{b}} \mathbf{I})$  are random effects.

In (1) each  $f_j$  has its own smoothing parameter.

In (2) all  $f_i$ 's have the same smoothing parameter.



# GAM model fitting

In general

$$f(\mathbf{x}) = \sum_{k=1}^{r} \beta_k b_k(\mathbf{x}).$$

To determine complexity of  $f(\mathbf{x})$ :

- the basis rank r is large enough for sufficient flexibility
- ullet a complexity penalty on eta controls the wiggliness of the effects

# GAM model fitting

 $\hat{oldsymbol{eta}}$  is the maximizer of **penalized** log-likelihood

$$\hat{\beta} = \operatorname*{argmax}_{\beta} \operatorname{PenLogLik}(\beta|\gamma) = \operatorname*{argmax}_{\beta} \big\{ \underbrace{\widetilde{L_{y}(\beta)}}_{\text{penalize complexity}} - \underbrace{\underbrace{\operatorname{Pen}(\beta|\gamma)}_{\text{penalize complexity}}} \big\}$$

where:

- $L_{y}(\beta) = \sum_{i} \log p(y_{i}|\beta)$  is log-likelihood
- ullet Pen $(eta|\gamma)$  penalizes the complexity of the  $f_j$ 's
- $\gamma > 0$  smoothing parameters ( $\uparrow \gamma \uparrow$ smoothness)

Concrete example  $\mu(\mathbf{x}) = f(x_1) + g(x_2, x_3)$ :

$$\mathsf{Pen}(oldsymbol{eta}|oldsymbol{\gamma}) pprox \gamma_1 \int f_{\mathsf{x}_1\mathsf{x}_1}^2 d\mathsf{x}_1 + \gamma_2 \int g_{\mathsf{x}_2\mathsf{x}_2}^2 + 2g_{\mathsf{x}_2\mathsf{x}_3}^2 + g_{\mathsf{x}_3\mathsf{x}_3}^2 d\mathsf{x}_2 d\mathsf{x}_3$$

# GAM model fitting

We use a hierarchical fitting framework:

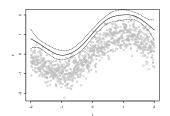
 $\ \, \textbf{ 9} \ \, \textbf{ Select } \gamma \ \, \textbf{ determine smoothness} \\$ 

$$\hat{\gamma} = \mathop{\mathsf{argmax}}_{\gamma} \mathsf{LAML}(\gamma)$$

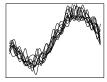
where LAML $(\gamma) \approx p(y|\gamma) = \int p(y,\beta|\gamma)d\beta$ .

**2** For fixed  $\gamma$ , estimate  $\beta$  to determine actual fit

$$\hat{eta} = \mathop{\mathsf{argmax}}_{eta} \mathop{\mathsf{PenLogLik}}(eta|\gamma).$$







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### Big Data methods

Recall the GAM model structure

$$\mu(\mathbf{x}) = \mathbb{E}(y|\mathbf{x}) = g^{-1}\Big\{\sum_{j=1}^m f_j(\mathbf{x})\Big\}$$

Here  $\mu(\mathbf{x}_i)$  can be written as  $g^{-1}(\mathbf{X}_i\beta)$ , where  $\mathbf{X}_i$  row of

$$\mathbf{X} = \begin{bmatrix} A_{11} & A_{12} & \cdots & b_{11}(x_{11}) & b_{12}(x_{11}) & \cdots & b_{21}(x_{21}) & \cdots \\ A_{21} & A_{22} & \cdots & b_{11}(x_{12}) & b_{12}(x_{12}) & \cdots & b_{21}(x_{22}) & \cdots \\ \vdots & \vdots & \ddots & \vdots & \ddots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \ddots & \vdots & \vdots & \ddots & \vdots & \ddots & \end{bmatrix}$$

with n rows and

$$d = p + k_1 + \cdots + k_j + \cdots + k_m$$

columns.

## Big Data methods

Bottom line: X can get very big, which causes problems:

- storing X takes too much memory
- $\bullet$  computing things involving **X** (e.g.  $\mathbf{X}^T\mathbf{X}$ ) takes time

Solution implemented in mgcv::bam function:

• do not create **X** but only sub-blocks:

$$\mathbf{X} = \left[ \begin{array}{cc} \mathbf{X}_{11} & \mathbf{X}_{12} \\ \mathbf{X}_{21} & \mathbf{X}_{22} \end{array} \right]$$

do not store them either, but create them when needed;

- any computation involving X is based on the blocks;
- use parallelization when possible;

## Big Data methods

Further acceleration and memory savings by discretization.

Instead of having n unique rows of X discretize to b << n rows.

#### In mgcv:

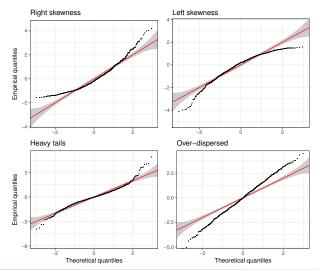
#### Structure of the talk

#### Structure:

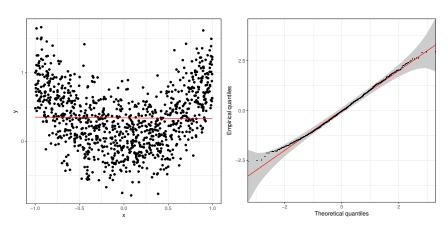
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In first hands-on session we'll use few basic diagnostics.

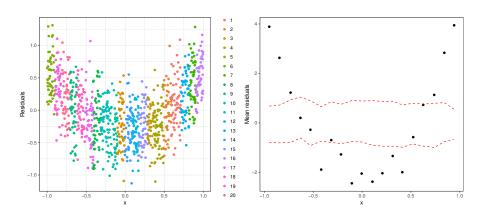
### QQ-plots



Useful for choosing model  $\operatorname{Dist}_m(y|\mathbf{x})$  (e.g. Poisson vs Neg. Binom.) Less useful for finding omitted variables and non-linearities.



#### Conditional residuals checks are more helpful here.



Recall structure of smooth effects:

$$f(\mathbf{x}) = \sum_{j=1}^k \beta_j b_j(\mathbf{x}).$$

where  $\beta$  shrunk toward zero by smoothness penalty.

Effective number of parameters we are using is < k.

Approximation is **Effective Degrees of Freedom** (EDF) < k.

By default k = 10 but this is arbitrary.

Exact choice of k not important, but it must not be too low.

#### Checking whether *k* is too low:

- look at conditional residuals checks
- ② look at output of check(fit):

```
## k' edf k-index p-value

## s(wM) 9.00 8.60 0.91 <2e-16 ***

## s(wM_s95) 9.00 8.13 1.02 0.76

## s(Posan) 8.00 2.66 1.04 0.97
```

increase k and see if a model selection criterion improves

#### Model selection

General criterion is approximate Akaike Information Criterion (AIC):

$$AIC = \underbrace{-2 \log p(\mathbf{y}|\hat{\beta})}_{\text{goodness of fit}} + \underbrace{2\tau}_{\text{model complexity}}$$

where  $\tau$  is EDF.

If  $AIC_{m1} < AIC_{m2}$  choose model 1.

To select which effects to include we can also look at p-values:

```
summary(fit)
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 267.2004 75.4197 3.543 0.000405 ***
## Fl 6.2854 1.0457 6.010 2.20e-09 ***
## loc2 79.8459 80.4130 0.993 0.320858
## loc3 -71.2728 86.1725 -0.827 0.408284
```

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# GAMs in mgcv and mgcViz

 ${\tt mgcv}$  is the recommended R package for fitting GAMs.

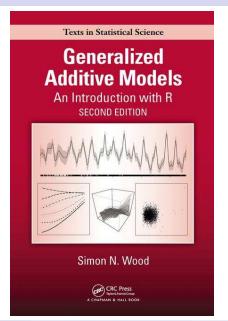
Today we'll work with mgcViz's interface.

mgcViz extends mgcv's tools for:

- plotting the estimated effects
- doing visual model checking

But most of the computation is done by mgcv.

## Further reading



#### References I

Hastie, T. and R. Tibshirani (1990). *Generalized Additive Models*, Volume 43. CRC Press.

Wood, S. (2006). Generalized additive models: an introduction with R. CRC press.