Beyond mean modelling: GAMLSS and quantile GAMs

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GAMLSS

Structure:

- Intro to GAMs for Location Scale and Shape
- Intro to quantile GAMs
- 3 GAMLSS and QGAM modelling in mgcv and ggam

Beyond mean modelling: GAMLSS models

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- **1** Intro to GAMs for Location Scale and Shape
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Recall GAM model structure:

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\theta_1 = \mu(\mathbf{x}), \theta_2, \dots, \theta_p\},$$

where

$$\mu(\mathbf{x}) = g^{-1} \Big\{ \sum_{j=1}^m f_j(\mathbf{x}) \Big\},\,$$

and g is the link function.

Example, Scaled Student-t distribution:

- location $\mu(\mathbf{x}) = \mathbb{E}(y|\mathbf{x})$
- scale $\theta_2 = \sigma$
- shape $\theta_3 = \nu$

In Generalized Additive Models for Location Scale and Shape (GAMLSS) (Rigby and Stasinopoulos, 2005) we let scale and shape change with the covariates \mathbf{x} .

GAMLSS model structure:

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\theta_1 = \mu_1(\mathbf{x}), \theta_2 = \mu_2(\mathbf{x}), \dots, \theta_p = \mu_p(\mathbf{x})\},$$

where

$$\mu_1(\mathbf{x}) = g_1^{-1} \Big\{ \sum_{j=1}^m f_j^1(\mathbf{x}) \Big\},$$

$$\mu_p(\mathbf{x}) = g_p^{-1} \Big\{ \sum_{i=1}^m f_j^p(\mathbf{x}) \Big\},\,$$

and g_1, \ldots, g_p are link function.

Example: Gaussian model for location and scale

Model is

$$y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma(\mathbf{x})\}$$

where

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = \sum_{j=1}^{m} f_j^1(\mathbf{x})$$

$$\operatorname{var}(y|\mathbf{x})^{1/2} = \sigma(\mathbf{x}) = \exp\Big\{\sum_{j=1}^m f_j^2(\mathbf{x})\Big\}$$

that is $g_2 = \log$ to guarantee $\sigma > 0$.

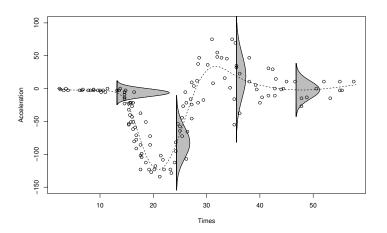
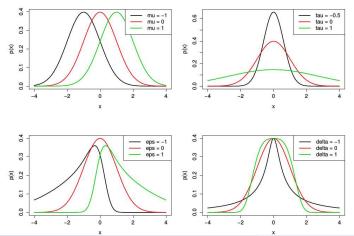


Figure: Gaussian model with variable mean and variance. In mgcv: gam(list(y~s(x), ~s(x)), family=gaulss).

Example: Sinh-arcsinh (shash) distribution

Four parameter distribution where location, scale, skewness (asymmetry) and kurtosis (tail behaviour) can depend on \mathbf{x} (Jones and Pewsey, 2009).



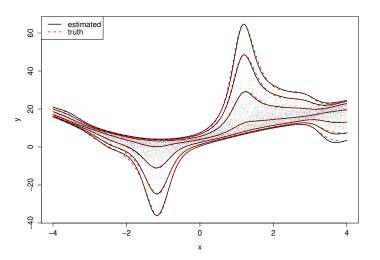
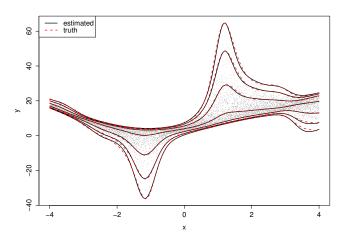


Figure: gam(list(y s(x), s(x), s(x), s(x), s(x)), family=shash).

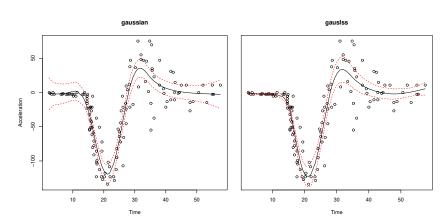
Why is this useful?

R1: you might be interested in whole distribution $y|\mathbf{x}$ not just $\mathbb{E}(y|\mathbf{x})$.



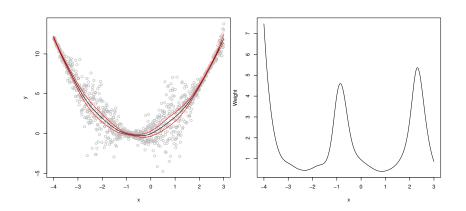
Why is this useful?

R2: standard GAM inference (e.g. p-value & confidence interval) is valid if the model for $y|\mathbf{x}$ is correct



Why is this useful?

R3: the accuracy of the fit is improved if the weight of each observation is inversely proportional to $Var(y|\mathbf{x})$.



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Regression setting:

- y is our response or dependent variable
- x is a vector of covariates or independent variables

In **distributional regression** we want a good model for $p(y|\mathbf{x})$.

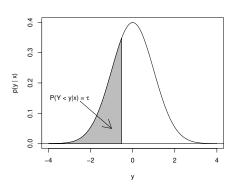
Model is $p_m\{y|\theta_1(\mathbf{x}),\ldots,\theta_q(\mathbf{x})\}$, where $\theta_1(\mathbf{x}),\ldots,\theta_q(\mathbf{x})$ are parameters.

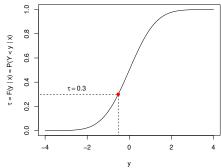
Lots of options for $p_m(y|\mathbf{x})$: binomial, gamma, Poisson, Tweedie...

We consider continuous (not discrete) y.

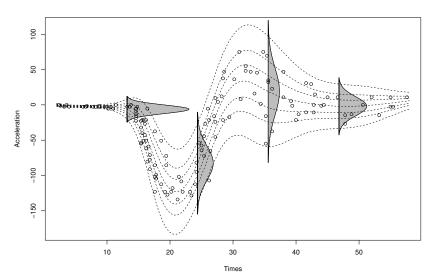
Define $F(y|\mathbf{x}) = \text{Prob}(Y \leq y|\mathbf{x})$.

The τ -th $(\tau \in (0,1))$ quantile is $\mu_{\tau}(\mathbf{x}) = F^{-1}(\tau|\mathbf{x})$.



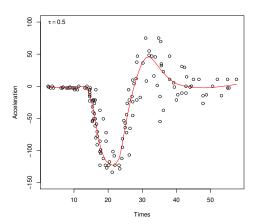


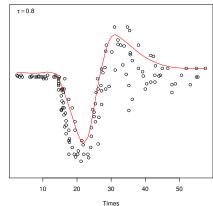
Given $p_m(y|\mathbf{x})$ we can get the conditional quantiles $\mu_{\tau}(\mathbf{x})$.



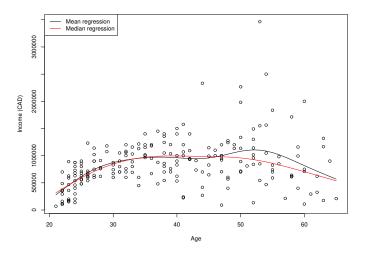
Quantile regression estimates conditional quantiles $\mu_{\tau}(\mathbf{x})$ directly.

No model for $p(y|\mathbf{x})$.

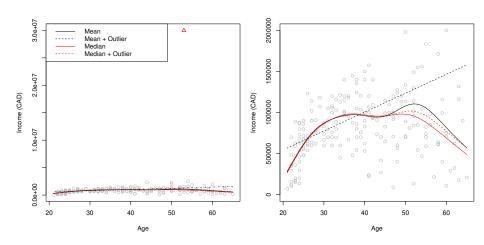




Median income is a better indicator of how the "average" person is doing, relative to mean income.

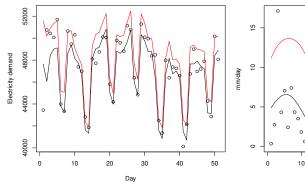


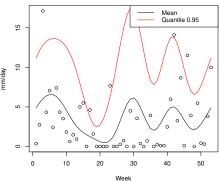
The median is also more **resistant to outliers**.



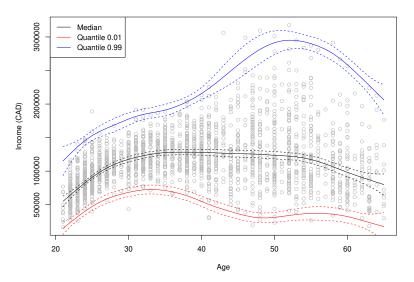
Some quantiles are more important than others:

- electricity producers need to satisfy top electricity demand
- urban planners need estimates of extreme rainfall

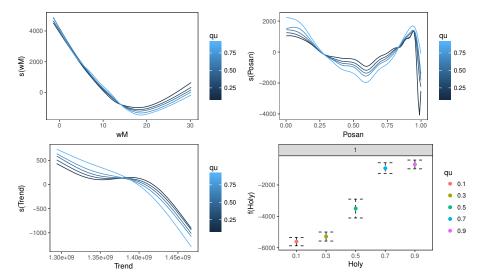




Effect of explanatory variables may depend on quantile

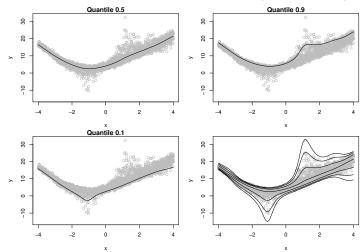


$$q_{\tau}(\mathsf{Demand}) = f_1(\mathsf{Temp}) + f_2(\mathsf{TimeOfYear}) + f_3(\mathsf{Trend}) + f_4(\mathsf{Holiday}) + \cdots$$



No assumptions on $p(y|\mathbf{x})$:

- no need to find good model for $p(y|\mathbf{x})$;
- no need to find normalizing transformations (e.g. Box-Cox);

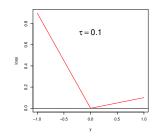


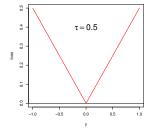
Recall definition $\mu_{\tau}(\mathbf{x}) = F^{-1}(\tau | \mathbf{x})$.

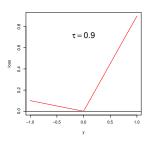
Key fact: $\mu_{\tau}(\mathbf{x})$ is the minimizer of

$$L(\mu|\mathbf{x}) = \mathbb{E}\{ \rho_{\tau}(y-\mu) | \mathbf{x} \},$$

where ρ_{τ} is the "pinball" loss (Koenker, 2005):







In additive modelling context $\mu_{\tau}(\mathbf{x}) = \mu_{\tau}(\beta) = \sum_{i=1}^{m} f_{i}(\mathbf{x})$.

Problem: how to perform Bayesian update $p(\beta|y) \propto p(y|\beta)p(\beta)$?

A coherent approach is the Gibbs posterior (Bissiri et al., 2016)

$$p(eta|y) \propto \underbrace{\mathrm{e}^{-rac{1}{\sigma}
ho_{ au}\{y-\mu(eta)\}}}_{ ext{pseudo}\ p(y|eta)} p(eta),$$

where $1/\sigma > 0$ is the "learning rate".

Recall that $p(\beta) = p(\beta|\lambda)$, hence we need to:

- select learning rate $1/\sigma$
- ullet select smoothing parameters $oldsymbol{\lambda}$
- ullet estimate regression coefficients eta

We use a hierarchical fitting framework:

 $lack {\odot}$ Select σ to optimise coverage of credible intervals

$$\hat{\sigma} = \underset{\sigma}{\operatorname{argmax}} \mathsf{IKL}(\sigma).$$

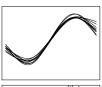
2 Select γ determine smoothness

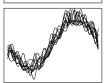
$$\hat{\gamma} = \mathop{\mathsf{argmax}}_{\gamma} \, \mathsf{LAML}(\gamma)$$

where LAML
$$(\gamma) \approx p(y|\gamma) = \int p(y,\beta|\gamma)d\beta$$
.

3 For fixed λ and σ , estimate β

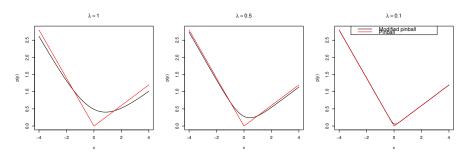
$$\hat{\boldsymbol{\beta}} = \underset{\boldsymbol{\beta}}{\operatorname{argmax}} \log p(\boldsymbol{\beta}|y).$$





ggam uses a modified loss which we call Extended log-F (ELF) loss.

This is smooth and convex and, as $\lambda \to 0$, we have recover pinball loss.



Since qgam 1.3.0, λ (err parameter) is selected automatically.

Learning rate can depend on covariates $\sigma = \sigma(\mathbf{x})$ (see examples).

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Demonstration in R

For more details on methodology, see:

Fasiolo, M., Goude, Y., Nedellec, R. and Wood, S.N., 2017. Fast calibrated additive quantile regression. arXiv preprint arXiv:1707.03307. and the file "intro_to_qgam.pdf".

For more software training material see

http://mfasiolo.github.io/qgam/articles/qgam.html

https://mfasiolo.github.io/mgcViz/articles/qgam_mgcViz.html

Now we move to "gamlss_qgam.html"

References I

- Bissiri, P. G., C. Holmes, and S. G. Walker (2016). A general framework for updating belief distributions. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*.
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