

#6 Let $Y = X\beta + \varepsilon$ where $\varepsilon \sim N_n(0, \sigma^2 I)$ and let $H = X(X'X)^{-1}X'$. In addition, let $d_i = \begin{bmatrix} 0 \\ \vdots \\ 1 \\ \vdots \\ 0 \end{bmatrix}$ $\leftarrow i^{\text{th}}$ spot has one in it

a) Let $B_i = \frac{(n-p)\hat{\sigma}^2}{\sigma^2} - \frac{\hat{\varepsilon}_i^2/(1-h_{ii})}{\sigma^2}$.

Note that $(n-p)\hat{\sigma}^2 = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 = Y'(I-H)Y$.

In addition, $1-h_{ii} = d_i'(I-H)d_i$ and $\hat{\varepsilon}_i = d_i'\hat{\varepsilon} = d_i'(I-H)Y$.

Hence, $\hat{\varepsilon}_i^2/(1-h_{ii}) = Y'(I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H)Y$.

Thus, $B_i = \left(\frac{Y}{\sigma}\right)' \left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \left(\frac{Y}{\sigma}\right)$

Note that $Y/\sigma \sim N_n(0, I)$.

Moreover, $\left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right]$
 $= (I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H)$

Since $(I-H)^2 = I-H$ and $d_i'(I-H)d_i = 1-h_{ii}$.

Also, $(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H)$ is symmetric and so we've got a projection matrix.

Therefore, $B_i \sim \chi_r^2$ where $r = \text{rank} \left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right]$.

$$\begin{aligned}
\text{Now, rank} & \left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \\
&= \text{tr} \left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \\
&= n-p - \frac{1}{1-h_{ii}} \text{tr} \left[(I-H) d_i d_i' (I-H) \right] \\
&= n-p - \frac{1}{1-h_{ii}} \text{tr} \left[d_i' (I-H) d_i \right] \\
&= n-p - \frac{1-h_{ii}}{1-h_{ii}} \\
&= n-p-1.
\end{aligned}$$

And so $B_i \sim \chi^2_{n-p-1}$.

b) Note that $A_i = \frac{1}{\sigma^2} Y' (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) Y$.

We look at "A" & "B" (the matrices between Y' & Y)

$$\begin{aligned}
\Rightarrow BA &= \frac{1}{\sigma^2} \left[(I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \cdot \frac{1}{\sigma^2} \left[(I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \\
&= \frac{1}{(\sigma^2)^2} \left[(I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right] \\
&= \frac{1}{(\sigma^2)^2} \left[(I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) - (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \right]
\end{aligned}$$

$$\text{since } d_i' (I-H) d_i = 1-h_{ii}$$

$$= 0.$$

Thus, $A_i \perp B_i$.

c) For $r_i = \frac{\hat{\varepsilon}_i}{\sqrt{1-h_{ii}}}$, note that

$$A_i = \frac{\hat{\sigma}^2}{\sigma^2} r_i^2.$$

symmetric, idempotent, rank = 1.

$$\text{In addition, } A_i = \left(\frac{Y}{\sigma}\right)' (I-H) \frac{d_i d_i'}{1-h_{ii}} (I-H) \left(\frac{Y}{\sigma}\right)$$

$\sim \chi^2_1$ (similar to part a) process).

And so since $A_i \perp B_i$: $\frac{A_i}{A_i+B_i} \sim \text{Beta}\left(\frac{1}{2}, \frac{n-p-1}{2}\right).$

$$\text{Thus, } \frac{A_i}{A_i+B_i} = \frac{\frac{\hat{\sigma}^2}{\sigma^2} r_i^2}{\frac{\hat{\sigma}^2}{\sigma^2} (n-p)} = \frac{r_i^2}{n-p} \sim \text{Beta}\left(\frac{1}{2}, \frac{n-p-1}{2}\right). \checkmark$$

d) The notes used a "trick" where we consider $(I-H)d_i \phi$ instead (to "orthogonalize" things).

$$\begin{aligned} \text{From there } \hat{\phi} &= [d_i' (I-H) d_i]^{-1} d_i' (I-H) Y \\ &= \frac{\hat{\varepsilon}_i}{1-h_{ii}} \quad (\text{as seen earlier, above}). \end{aligned}$$

So to test $\phi=0$:

consider $\frac{A_i}{\sqrt{B_i}} \sim t(n-p-1)$. Then

$$t^* = \frac{A_i}{\sqrt{B_i}} = \frac{\hat{\phi} - 0}{\left[(n-p)\hat{\sigma}^2 - \hat{\phi} \hat{\varepsilon}_i \right]^{1/2}} \sim t(n-p-1); \quad \text{and compare to } t\left(\frac{2-\alpha}{2}, n-p-1\right).$$

$$\text{Let } P_I = X_I(X_I'X_I)^{-1}X_I'$$

$$e) \text{ For a single } i, \hat{\beta}_{(i)} = \hat{\beta} - \frac{(X'X)^{-1}X_i'\hat{\varepsilon}_i}{1-h_{ii}}$$

$$= \hat{\beta} - (X'X)^{-1}(X'd_i)[d_i'(I-H)d_i]^{-1}d_i'\hat{\varepsilon}$$

$$\begin{aligned} \text{So, in general, } \hat{\beta}_I &= \hat{\beta} - (X'X)^{-1}(X'd_I)[d_I'(I-H)d_I]^{-1}d_I'\hat{\varepsilon} \\ &= \hat{\beta} - (X'X)^{-1}X_I[d_I'd_I - d_I'X(X'X)^{-1}X'd_I]^{-1}\hat{\varepsilon}_I \\ &= \hat{\beta} - (X'X)^{-1}X_I[I_m - X_I(X'X)^{-1}X_I']\hat{\varepsilon}_I \\ &= \hat{\beta} - (X'X)^{-1}X_I[I - P_I]^{-1}\hat{\varepsilon}_I. \end{aligned}$$

We need to show that equivalently

$$(\hat{\beta} - \hat{\beta}_I)'(X'X)(\hat{\beta} - \hat{\beta}_I) = \sum_{i=1}^m \frac{(d_i'\hat{\varepsilon}_I)^2}{1-h_{ii}} \left(\frac{\lambda_i}{1-\lambda_i} \right).$$

$$\text{Note that } (\hat{\beta} - \hat{\beta}_I)'(X'X)(\hat{\beta} - \hat{\beta}_I)$$

$$= \hat{\varepsilon}_I'(I-P_I)^{-1}X_I'(X'X)^{-1}(X'X)(X'X)^{-1}X_I(I-P_I)^{-1}\hat{\varepsilon}_I$$

Now, $I - P_I = QQ' - Q\Lambda Q'$
 $= Q[I - \Lambda]Q'$

And $(I - P)^{-1} = Q[I - \Lambda]^{-1}Q'$ [Note: $Q[I - \Lambda]Q'Q[I - \Lambda]^{-1}Q' = Q[I - \Lambda][I - \Lambda]^{-1}Q' = QQ' = I$]
 $= Q \text{diag}\left(\frac{1}{1 - \lambda_i}\right)Q'$

So $(\hat{\beta} - \hat{\beta}_I)'(X'X)(\hat{\beta} - \hat{\beta}_I)$

$= \hat{\varepsilon}_I' \left(Q \text{diag}\left(\frac{1}{1 - \lambda_i}\right)Q' \right) (Q\Lambda Q') \left(Q \text{diag}\left(\frac{1}{1 - \lambda_i}\right)Q' \right) \hat{\varepsilon}_I$

$= \hat{\varepsilon}_I' Q \text{diag}\left(\frac{\lambda_i}{(1 - \lambda_i)^2}\right) Q' \hat{\varepsilon}_I$

$= \begin{bmatrix} y_1' \hat{\varepsilon}_I \\ \vdots \\ y_m' \hat{\varepsilon}_I \end{bmatrix}' \text{diag}\left(\frac{\lambda_i}{(1 - \lambda_i)^2}\right) \begin{bmatrix} y_1' \hat{\varepsilon}_I \\ \vdots \\ y_m' \hat{\varepsilon}_I \end{bmatrix}$

$= \sum_{i=1}^m \frac{\lambda_i}{(1 - \lambda_i)^2} (y_i' \hat{\varepsilon}_I)^2$

$= \sum_{i=1}^m \frac{(y_i' \hat{\varepsilon}_I)^2}{1 - \lambda_i} \left(\frac{\lambda_i}{1 - \lambda_i} \right), \text{ as needed.}$

By equivalence,

$\frac{(\hat{\beta} - \hat{\beta}_I)'(X'X)(\hat{\beta} - \hat{\beta}_I)}{p\hat{\sigma}^2} = \frac{1}{p} \sum_{i=1}^m h_i^2 \frac{\lambda_i}{1 - \lambda_i}$

where $h_i^2 = \frac{(y_i' \hat{\varepsilon}_I)^2}{\hat{\sigma}^2(1 - \lambda_i)}$

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