

# 2009 Theory II #3

$$\begin{aligned}
 3a) L_n(\theta) &= \prod_{k=1}^2 \prod_{i=1}^n \frac{\frac{y_{ik}}{\lambda_{ik}} e^{-\lambda_{ik}}}{y_{ik}!} (2\pi\sigma_k^2)^{-\gamma_k/2} \exp\left\{-\frac{1}{2\sigma_k^2} x_{ik}^2\right\} \\
 &= \prod_{k=1}^2 \prod_{i=1}^n \frac{(2\pi\sigma_k^2)^{-\gamma_k}}{y_{ik}!} \exp\left\{y_{ik}(\alpha_k + \beta x_{ik}) - e^{\alpha_k + \beta x_{ik}} - \frac{1}{2\sigma_k^2} x_{ik}^2\right\} \\
 l_n(\theta) &= \sum_{k=1}^2 \sum_{i=1}^n \left\{ -\frac{1}{2} \log(2\pi) - \log(y_{ik}) - \frac{1}{2} \log(\sigma_k^2) \right. \\
 &\quad \left. + y_{ik}(\alpha_k + \beta x_{ik}) - e^{\alpha_k + \beta x_{ik}} - \frac{1}{2\sigma_k^2} x_{ik}^2 \right\}
 \end{aligned}$$

Need to find  $I_n(\theta)$

$$\begin{aligned}
 \frac{\partial}{\partial \alpha_1} l_n(\theta) &= \frac{\partial}{\partial \alpha_1} \sum_{i=1}^n \left\{ y_{i1}(\alpha_1 + \beta x_{i1}) - e^{\alpha_1 + \beta x_{i1}} \right\} \\
 &= \sum_{i=1}^n \left\{ y_{i1} - e^{\alpha_1 + \beta x_{i1}} \right\}
 \end{aligned}$$

$$\frac{\partial}{\partial \alpha_2} l_n(\theta) = \sum_{i=1}^n \left\{ y_{i2} - e^{\alpha_2 + \beta x_{i2}} \right\}$$

$$\begin{aligned}
 \frac{\partial}{\partial \beta} l_n(\theta) &= \frac{\partial}{\partial \beta} \sum_{k=1}^2 \sum_{i=1}^n \left\{ y_{ik}(\alpha_k + \beta x_{ik}) - e^{\alpha_k + \beta x_{ik}} \right\} \\
 &= \sum_{k=1}^2 \sum_{i=1}^n \left\{ x_{ik} y_{ik} - x_{ik} e^{\alpha_k + \beta x_{ik}} \right\}
 \end{aligned}$$

$$\begin{aligned}
 \frac{\partial}{\partial \sigma_1^2} l_n(\theta) &= \frac{\partial}{\partial \sigma_1^2} \sum_{i=1}^n \left\{ -\frac{1}{2} \log(\sigma_1^2) - \frac{1}{2\sigma_1^2} x_{ik}^2 \right\} \\
 &= \sum_{i=1}^n \left\{ -\frac{1}{2\sigma_1^2} + \frac{1}{2(\sigma_1^2)^2} x_{ik}^2 \right\} = -\frac{n}{2\sigma_1^2} + \frac{1}{2(\sigma_1^2)^2} \sum_{i=1}^n x_{ik}^2
 \end{aligned}$$

$$\frac{\partial}{\partial \sigma_2^2} l_n(\theta) = -\frac{n}{2\sigma_2^2} + \frac{1}{2(\sigma_2^2)^2} \sum_{i=1}^n x_{ik}^2$$

Now the second-order partial derivatives:

$$\frac{\partial^2}{\partial \alpha_2} l_n(\theta) = -\sum_{i=1}^n e^{\alpha_2 + \beta x_{iz}}$$

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$$\frac{\partial^2}{\partial \beta^2} l_n(\theta) = -\sum_{k=1}^2 \sum_{i=1}^n x_{ik}^2 e^{\alpha_k + \beta x_{ik}}$$

$$\frac{\partial^2}{\partial (\sigma_2^2)^2} l_n(\theta) = \frac{n}{2(\sigma_2^2)^2} - \frac{1}{(\sigma_2^2)^3} \sum_{i=1}^n x_{ik}^2$$

$$\frac{\partial^2}{\partial (\sigma_2^2)^2} l_n(\theta) = \frac{n}{2(\sigma_2^2)^2} - \frac{1}{(\sigma_2^2)^3} \sum_{i=1}^n x_{ik}^2$$

$$\frac{\partial^2}{\partial \alpha_2 \partial \beta} l_n(\theta) = -\sum_{i=1}^n x_{iz} e^{\alpha_2 + \beta x_{iz}}$$

0 for all others

$$\frac{\partial^2}{\partial \alpha_2 \partial \beta} l_n(\theta) = -\sum_{i=1}^n x_{iz} e^{\alpha_2 + \beta x_{iz}}$$

Next,

$$E[X_{ik}^2] = \sigma_k^2$$

$$\begin{aligned} E[X_{ik}^2 e^{\alpha_k + \beta X_{ik}}] &= \int_{-\infty}^{\infty} x_{ik}^2 e^{\alpha_k + \beta x} \exp\left\{-\frac{1}{2\sigma_k^2} x^2\right\} dx \\ &= e^{\alpha_k} \int_{-\infty}^{\infty} x^2 \exp\left\{-\frac{1}{2\sigma_k^2} x^2 + \beta x\right\} dx = e^{\alpha_k} \int_{-\infty}^{\infty} x^2 \exp\left\{-\frac{1}{2\sigma_k^2} \left[x^2 - 2\sigma_k^2 \beta x + \sigma_k^4 \beta^2 - \sigma_k^4 \beta^2\right]\right\} dx \end{aligned}$$

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$$= \exp\left\{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2\right\} \int_{-\infty}^{\infty} x^2 \exp\left\{-\frac{1}{2\sigma_k^2}(x - \sigma_k^2\beta)^2\right\} dx$$

$\sim N(\beta\sigma_k^2, \sigma_k^2)$  kernel

$$= \cancel{\exp\left\{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2\right\}} \left[ \frac{\sigma_k^2 + (\sigma_k^2\beta)^2}{\beta} \right]$$

$\sigma_k^2$        $\sigma_k^2$

Then  $\alpha_1 \left[ \begin{array}{c} ne^{\alpha_2 + \frac{1}{2}\sigma_2^2\beta^2} \\ ne^{\alpha_2 + \frac{1}{2}\sigma_2^2\beta^2} \\ ne^{\alpha_2 + \frac{1}{2}\sigma_2^2\beta^2} \\ n \sum_{k=1}^2 [\sigma_k^2 + (\sigma_k^2\beta)^2] \exp\{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2\} \\ \sigma_1^2 \\ \sigma_2^2 \end{array} \right]$

$$\ln(\alpha_1, \alpha_2, \beta, \sigma_1^2, \sigma_2^2) = \beta \frac{\partial \alpha_1^2 e^{\alpha_2 + \frac{1}{2}\sigma_2^2\beta^2}}{\partial \sigma_1^2}$$

$\frac{n}{2(\sigma_2^2)^2}$        $\frac{n}{2(\sigma_2^2)^2}$

$$E[e^{\alpha_k + \beta X_{ik}}] = \int_{-\infty}^{\infty} e^{\alpha_k + \beta x} \exp\left\{-\frac{1}{2\sigma_k^2}x^2\right\} dx = e^{\alpha_k} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma_k^2}} \exp\left\{-\frac{1}{2\sigma_k^2}x^2 + \beta x\right\} dx$$

$$= e^{\alpha_k} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma_k^2}} \exp\left\{-\frac{1}{2\sigma_k^2}\left[x^2 - 2\sigma_k^2\beta x + \sigma_k^4\beta^2 - \sigma_k^4\beta^2\right]\right\} dx$$

$$= e^{\alpha_k + \sigma_k^2\beta^2/2} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma_k^2}} \exp\left\{-\frac{1}{2\sigma_k^2}(x - \sigma_k^2\beta)^2\right\} dx = e^{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2}$$

$$E[X_{ik} e^{\alpha_k + \beta X_{ik}}] = \dots = \cancel{\beta \sigma_k^2} e^{\alpha_k + \frac{1}{2}\beta^2\sigma_k^2}$$

Use next page!

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$$\begin{array}{c|ccccc|c}
 & \sigma_1 & \sigma_2 & \alpha_1 & \alpha_2 & & \beta \\
 \hline
 \sigma_1 & \frac{n}{2(\sigma_1^2)^2} & 0 & 0 & 0 & A & 0 \\
 \sigma_2 & 0 & \frac{n}{2(\sigma_2^2)^2} & 0 & 0 & & 0 \\
 \hline
 \alpha_1 & 0 & 0 & ne^{\alpha_1 + \frac{1}{2}\sigma_1^2\beta} & 0 & & n\beta\sigma_1^2 e^{\alpha_1 + \frac{1}{2}\beta^2\sigma_1^2} \\
 \alpha_2 & 0 & 0 & 0 & ne^{\alpha_2 + \frac{1}{2}\sigma_2^2\beta} & & n\beta\sigma_2^2 e^{\alpha_2 + \frac{1}{2}\beta^2\sigma_2^2} \\
 \hline
 \beta & 0 & 0 & n\beta\sigma_1^2 e^{\alpha_1 + \frac{1}{2}\beta^2\sigma_1^2} & n\beta\sigma_2^2 e^{\alpha_2 + \frac{1}{2}\beta^2\sigma_2^2} & n \sum_{k=1}^2 [\sigma_k^2 + (\sigma_k^2\beta)^2] \exp\left\{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2\right\} & C \\
 \end{array}$$

\* Take  $\lim_{n \rightarrow \infty} \frac{1}{n}$  first next time

$$([I_n(\theta)]^{-1})_{5,5} = \{C - B' A^{-1} B\}^{-1}$$

$$= \left\{ n \sum_{k=1}^2 [\sigma_k^2 + (\sigma_k^2\beta)^2] \exp\left\{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2\right\} - \frac{(n\beta\sigma_1^2 e^{\alpha_1 + \frac{1}{2}\beta^2\sigma_1^2})^2}{ne^{\alpha_1 + \frac{1}{2}\sigma_1^2\beta}} - \frac{(n\beta\sigma_2^2 e^{\alpha_2 + \frac{1}{2}\beta^2\sigma_2^2})^2}{ne^{\alpha_2 + \frac{1}{2}\sigma_2^2\beta}} \right\}^{-1}$$

$$= \left\{ n \sum_{k=1}^2 \sigma_k^2 e^{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2} \right\}^{-1}$$

Then

$$\sqrt{n}(\hat{\beta} - \beta) \xrightarrow{D} N(0, \gamma) \text{ where}$$

$$\gamma = \left[ \lim_{n \rightarrow \infty} \frac{1}{n} ([I_n(\theta)]^{-1})_{5,5} \right] = \left\{ \frac{1}{n} \sum_{k=1}^2 \sigma_k^2 e^{\alpha_k + \frac{1}{2}\sigma_k^2\beta^2} \right\}^{-1}$$

$k=1, 2$ 

2b.i) Same process as before, now using only the data from one center.  
 See Joe's answer. For an estimator, use  $\left[ \lim_{n \rightarrow \infty} \frac{1}{n} I_n(\theta) \right]^{-1} \Big|_{\theta=\hat{\theta}}$   
 When is this consistent?

2b.ii)

Under the MLE theory regularity conditions and since the data from the two centers is independent we obtain

$$\sqrt{n} \left( \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} - \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} \right) \xrightarrow{\text{D}} N \left( 0, \begin{bmatrix} V_1 & 0 \\ 0 & V_2 \end{bmatrix} \right) \quad \text{where } V_1, V_2 \text{ as found in (2b.i)}$$

so that by the delta method for  $g$  as specified we have

$$\sqrt{n} (g(\hat{\beta}_1, \hat{\beta}_2) - \beta) \xrightarrow{\text{D}} N \left( 0, \left( \frac{\partial g_1(\beta, \beta)}{\partial \beta} \right)^2 V_1 + \left( \frac{\partial g_2(\beta, \beta)}{\partial \beta} \right)^2 V_2 \right)$$

↓  
derivative is continuous

2b.iii)  $g^*$  is continuously differentiable, scalar, and  
 $g^*(\beta, \beta) = \beta$ . How to show it is optimum?

minimize  $a^2 V_1 + b^2 V$  subject to

$$g(\beta) = \beta$$

$$\dot{g}_1(\beta) = a$$

$$\dot{g}_2(\beta) = b$$

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$$3b. iv) \frac{\partial}{\partial(x,y)} g_{opt}^*(x,y) = \begin{pmatrix} \frac{V_2}{V_1+V_2}, & \frac{V_1}{V_1+V_2} \end{pmatrix}$$

$\downarrow V_1 V_2 (V_1 + V_2)$

$$\text{Then } \dot{g}_{opt}^*(\beta, \beta) \begin{bmatrix} V_1 & 0 \\ 0 & V_2 \end{bmatrix} [\dot{g}_{opt}^*(\beta, \beta)]' = \underbrace{\frac{V_1^2 V_2 + V_2^2 V_1}{(V_1+V_2)^2}}_{\text{so}} = \frac{V_1 V_2}{V_1+V_2}$$

$\sqrt{2n}(g_{opt}^*(\hat{\beta}_1, \hat{\beta}_2) - \beta) \xrightarrow{L} N\left(0, \frac{2V_1 V_2}{V_1+V_2}\right)$

~~NOTATION~~ Therefore,

$$\text{ARE}(\hat{g}_{opt}^*(\hat{\beta}_1, \hat{\beta}_2), \hat{\beta}) = \frac{\frac{2V_1 V_2}{V_1+V_2}}{\frac{2}{\hat{V}_1^{-1} + \hat{V}_2^{-1}}} \Bigg|_{\substack{\hat{V}_1, \hat{V}_2}} = \frac{V_1 V_2}{V_1+V_2} \left( \frac{1}{\frac{1}{V_1} + \frac{1}{V_2}} \right) \Bigg|_{\substack{\hat{V}_1, \hat{V}_2}} = 1$$

3c) Same approach as in (1), see Joe's answer