

$$\beta_1^* = \beta_1 + (X_1^T X_1)^{-1} X_1^T X_3 \beta_3$$

3) a) * Reparametrize $Y = X_1 \beta_1 + X_3 \beta_3 + \varepsilon_2 \Leftrightarrow Y = X_1 \beta_1^* + (I - M_1) X_3 \beta_3 + \varepsilon_2$

now $\begin{pmatrix} X_1^T \\ X_3^T (I - M_1) \end{pmatrix} (X_1, (I - M_1) X_3) \begin{pmatrix} \hat{\beta}_1^* \\ \hat{\beta}_3 \end{pmatrix} = \begin{pmatrix} X_1^T Y \\ (I - M_1) X_3^T Y \end{pmatrix}$

thus $\boxed{X_3^T (I - M_1) X_3 \hat{\beta}_3 = X_3^T (I - M_1) X_3^T Y}$ ✓

* Reparametrize $Y = (X_1, X_2) \begin{pmatrix} \gamma_1^* \\ \gamma_2^* \end{pmatrix} + (I - M_{12}) X_3 \gamma_3 + \varepsilon_1$ ✓ and

by normal equations we obtain

$\boxed{X_3^T (I - M_{12}) X_3 \hat{\gamma}_3 = X_3^T (I - M_{12}) X_3^T Y}$ ✓

We prove the negation statement

* $\text{sign}(\hat{\beta}_3) = \text{sign}(\hat{\gamma}_3) \Leftrightarrow \text{sign}[X_3^T (I - M_1)] = \text{sign}[X_3^T (I - M_{12})]$

\Rightarrow : If $\text{sign}(\hat{\beta}_3) = \text{sign}(\hat{\gamma}_3)$ wlog assume are both positive
 then since $X_3^T (I - M_1) X_3$ and $X_3^T (I - M_{12}) X_3$ are both quadratic forms then

$\text{sign}(X_3^T (I - M_{12}) X_3) = \text{sign}(X_3^T (I - M_1) X_3) > 0$ ✓ thus we

must have $\text{sign}(X_3^T (I - M_1) Y) > 0$ and $\text{sign}(X_3^T (I - M_{12}) Y) > 0$.

\Leftarrow : If $\text{sign}(X_3^T (I - M_{12}) Y) = \text{sign}(X_3^T (I - M_1) Y) \xrightarrow{\text{wlog}} > 0$ and $\text{sign}(X_3^T (I - M_1) X_3) = \text{sign}(X_3^T (I - M_{12}) X_3) > 0$ then we must have that $\text{sign}(\hat{\beta}_3) > 0$ and $\text{sign}(\hat{\gamma}_3) > 0$ ✓ thus $\text{sign}(\hat{\beta}_3) = \text{sign}(\hat{\gamma}_3)$.

b) (i) M_{12} is the opo onto $C(X_1, X_2)$

$$\text{now } C(X_1, X_2) = C(X_1, X_2) \cap C(X_1) + C(X_1, X_2) \cap C(X_1)^\perp$$

$$\Leftrightarrow C(X_1, X_2) = C(X_1) + C(X_1, X_2) \cap C(X_1)^\perp$$

$$\text{thus } M_{12} = M_1 + M_3 \Rightarrow \boxed{M_3 = M_{12} - M_1} \checkmark$$

(ii) We must show $\boxed{(I - M_1)M_3 = M_3}$ indeed

$$(I - M_1)(M_{12} - M_1) = M_{12} - M_1 - M_1 M_{12} + M_1^2 = \begin{matrix} \text{since } M_1(C(X_1) \subset C(X_1, X_2)) \\ M_{12} - M_1 - M_1 + M_1 = M_{12} - M_1 = M_3 \checkmark \text{ as required.} \end{matrix}$$

$$\begin{aligned} \text{c) } X_3^T (I - M_1)Y - X_3^T M_3 Y < 0 &\Leftrightarrow X_3^T (I - M_1)Y - X_3^T (M_{12} - M_1)Y < 0 \\ &\Leftrightarrow X_3^T (I - M_1 - M_{12} + M_1)Y < 0 \Leftrightarrow X_3^T (I - M_{12})Y < 0 \checkmark \end{aligned}$$

thus $\text{sign}(\hat{\gamma}_3) < 0 \checkmark$ by using part (a) using

part (a) Thus $\text{sign}(\hat{\gamma}_3) < 0$ and $\text{sign}(\hat{\beta}_3) > 0$ thus

$$\text{sign}(\hat{\gamma}_3) \neq \text{sign}(\hat{\beta}_3) \checkmark$$

□

$$M_0 = \frac{1}{n} J_n J_n^T$$

$$d) (i) R^2 = \frac{Y^T M Y}{Y^T (I - M_0) Y} = \frac{Y^T M Y}{Y^T M Y + Y^T (I - M_0 - M) Y}$$

$$\text{Now let } X = Y^T M Y \sim \chi^2(p) \checkmark \text{ and } Y^T (I - M_0 - M) Y \sim \chi^2(n-1-p) \checkmark$$

$$\text{Now } M(I - M_0 - M) = M - M M_0 - M^2 = M - M M_0 - M = -M M_0$$

$$M M_0 = \frac{1}{n} X (X^T X)^{-1} X^T J_n J_n^T = 0 \text{ thus } X \text{ is independent from } Y \checkmark$$

$$\text{thus } R^2 = \frac{X}{X+Y} \sim \text{Beta}\left(\frac{p}{2}, \frac{n-1-p}{2}\right) \checkmark$$

$$* E(R^2) = \frac{\frac{p}{2}}{\frac{n-1-p}{2} + \frac{p}{2}} = \frac{\frac{p}{2}}{\frac{n-1}{2}} = \boxed{\frac{p}{n-1}} \checkmark$$

$$(ii) \frac{n-1-p}{p} = \frac{1}{\lambda} \Rightarrow \frac{n-1}{p} - 1 = \frac{1}{\lambda} \Rightarrow \frac{n-1}{p} = \frac{1}{\lambda} + 1 \Rightarrow$$

$$\frac{p}{n-1} = \frac{1}{\frac{1+\lambda}{\lambda}} = \frac{\lambda}{1+\lambda}$$

$$\text{let } R = R^2$$

$$\boxed{E(R^2) = \frac{p}{n-1}}$$

$$\text{let } b = \frac{p}{2} \\ \text{let } c = \frac{n-1-p}{2}$$

$$* \text{ Now } \text{Var}(R) = E(R^2) - E(R)^2 =$$

$$E(R^2) = \int \frac{\Gamma(b+c)}{\Gamma(b)\Gamma(c)} x^{b-1} (1-x)^{c-1} dx \rightarrow \text{kernel Beta}(b, c)$$

$$\text{thus } E(R^2) = \frac{\Gamma(b+c) \Gamma(b+2)}{\Gamma(b) \Gamma(b+c+2)} = \frac{(b+1)(b)}{(b+c+1)(b+c)} = \frac{\left(\frac{p}{2}+1\right)\left(\frac{p}{2}\right)}{\left(\frac{n-1}{2}\right)\left(\frac{n-1}{2}+1\right)} = \left(\frac{\lambda}{1+\lambda}\right)^{\frac{-1}{2}} \checkmark$$

Thus $\text{Var}(R) = \left(\frac{\lambda}{1+\lambda}\right)^2 - \left(\frac{\lambda}{1+\lambda}\right)^2 = 0$

Thus the colimit of R^2 is the expectation

$$\frac{p}{n-1} = \boxed{\frac{\lambda}{1+\lambda}} \checkmark$$

In fact R^2 converges to a dirac delta function with point mass given at $\frac{\lambda}{1+\lambda}$.