

2010 Theory II #1

$$\begin{aligned}
 1a) P(Y_i = y_i | X_i) &= \frac{P(x_i, y_i)}{\sum_{y=0}^1 P(x_i, y_i)} \\
 &= \frac{\frac{1}{\sqrt{2\pi\sigma^2}} \left[\exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_1)^2 \right\} \right]^{y_i} \left[\exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_0)^2 \right\} \right]^{1-y_i} \pi_1^{y_i} (1-\pi_1)^{1-y_i}}{\frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_0)^2 \right\} + \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_1)^2 \right\}}
 \end{aligned}$$

Then

$$\begin{aligned}
 \text{logit}(P(Y_i = 1 | X_i)) &= \log \left\{ \frac{P(Y_i = 1 | X_i)}{P(Y_i = 0 | X_i)} \right\} \\
 &= \log \left\{ \frac{\pi_1 \exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_1)^2 \right\}}{(1-\pi_1) \exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_0)^2 \right\}} \right\} \\
 &= \text{logit}\pi_1 + \frac{1}{2\sigma^2} [x_i^2 - 2\mu_0 x_i + \mu_0^2 - (x_i^2 - 2\mu_1 x_i + \mu_1^2)] \\
 &= \underbrace{\text{logit}\pi_1}_{\alpha_0} + \underbrace{\frac{\mu_0^2 - \mu_1^2}{2\sigma^2}}_{\alpha_1} + \underbrace{2\frac{(\mu_1 - \mu_0)}{\sigma^2} x_i}_{\alpha_2}
 \end{aligned}$$

1b) See 4.9 in the GLM solutions for this.

No need to calculate θ yet

(2)

$$1c) p(x_i, y_i) = \frac{1}{\sqrt{2\pi\sigma^2}} \left[\exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_1)^2 \right\} \right]^{y_i} \left[\exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu_0)^2 \right\} \right]^{1-y_i} \frac{\pi_1^{y_i} (1-\pi_1)^{1-y_i}}{\pi_2^{y_i} (1-\pi_2)^{1-y_i}}$$

$$\ln(\theta) = \sum_{i=1}^n \left\{ -\frac{1}{2} \log(2\pi) - \frac{1}{2} \log(\sigma^2) - \frac{y_i}{2\sigma^2} (x_i - \mu_1)^2 - \frac{1-y_i}{2\sigma^2} (x_i - \mu_0)^2 + y_i \log\left(\frac{\pi_1}{1-\pi_1}\right) + \log(1-\pi_1) \right\}$$

Let $A = \{i: y_i = 0\}$, $B = \{i: y_i = 1\}$

$$\frac{\partial}{\partial \mu_0} \ln(\theta) = \frac{1}{\sigma^2} \sum_{i \in A} (x_i - \mu_0) \text{ set } 0 \Rightarrow \hat{\mu}_0 = \bar{x}_A$$

$$\frac{\partial}{\partial \mu_1} \ln(\theta) = \frac{1}{\sigma^2} \sum_{i \in B} (x_i - \mu_1) \text{ set } 0 \Rightarrow \hat{\mu}_1 = \bar{x}_B$$

$$\frac{\partial}{\partial \sigma^2} \ln(\theta) = -\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i \in A} (x_i - \mu_0)^2 + \frac{1}{2(\sigma^2)^2} \sum_{i \in B} (x_i - \mu_1)^2 \text{ set } 0$$

$$\Rightarrow \hat{\sigma}^2 = \frac{\sum_{i \in A} (x_i - \mu_0)^2 + \sum_{i \in B} (x_i - \mu_1)^2}{n}$$

$$\frac{\partial}{\partial \pi} \ln(\theta) = \sum_{i=1}^n \frac{\partial}{\partial \pi} \left\{ y_i \log \pi_1 + (1-y_i) \log (1-\pi) \right\}$$

$$= \sum_{i=1}^n \left\{ \frac{y_i}{\pi_1} - \frac{1-y_i}{1-\pi} \right\} = \frac{n\bar{y}_n}{\pi_1} - \frac{n(1-\bar{y}_n)}{1-\pi_1} \text{ set } 0$$

$$\Rightarrow (1-\hat{\pi}_1)\bar{y}_n = \hat{\pi}_1(1-\bar{y}_n) \Rightarrow \bar{y}_n - \hat{\pi}_1\bar{y}_n = \hat{\pi}_1 - \hat{\pi}_1\bar{y}_n$$

$$\Rightarrow \hat{\pi}_1 = \bar{y}_n$$

(3)

$$= -\frac{1}{\sigma^2} \sum (1-y_i)$$

$$\frac{\partial^2}{\partial \mu_0^2} \ell_n(\theta) = -\frac{1}{\sigma^2} \sum_{i \in A} 1 \quad \frac{\partial^2}{\partial \mu_1^2} \ell_n(\theta) = -\frac{1}{\sigma^2} \sum_{i \in B} = -\frac{1}{\sigma^2} \sum y_i$$

$$\frac{\partial^2}{\partial (\sigma^2)^2} \ell_n(\theta) = \frac{n}{2(\sigma^2)^2} - \frac{1}{(\sigma^2)^3} \sum (1-y_i)(x_i - \mu_0)^2 - \frac{1}{(\sigma^2)^3} \sum y_i (x_i - \mu_1)^2$$

$$\frac{\partial^2}{\partial \sigma^2 \partial \mu_0} \ell_n(\theta) = -\frac{1}{(\sigma^2)^2} \sum_{i \in A} (x_i - \mu_0) = -\frac{1}{(\sigma^2)^2} \sum (1-y_i)(x_i - \mu_0)$$

$$\frac{\partial^2}{\partial \sigma^2 \partial \mu_1} \ell_n(\theta) = -\frac{1}{(\sigma^2)^2} \sum_{i \in B} (x_i - \mu_1) = -\frac{1}{(\sigma^2)^2} \sum y_i (x_i - \mu_1)$$

$$\frac{\partial^2}{\partial \pi_1^2} \ell_n(\theta) = -\frac{n \bar{y}_n}{\pi_1^2} - \frac{n(1-\bar{y}_n)}{(1-\pi_1)^2} = -n \left[\frac{\bar{y}_n(1-\pi_1)^2 + \pi_1^2(1-\bar{y}_n)}{\pi_1^2(1-\pi_1)^2} \right]$$

$$= -n \left[\frac{\bar{y}_n(1-2\pi_1 + \pi_1^2) + \pi_1^2 - \pi_1^2 \bar{y}_n}{\pi_1^2(1-\pi_1)^2} \right]$$

$$= -n \left[\frac{\bar{y}_n - 2\pi_1 \bar{y}_n + \pi_1^2}{\pi_1^2(1-\pi_1)^2} \right]$$

0 for all other partial derivatives

Notice that

$$\text{and similarly for } E[(1-y_i)(x_i - \mu_0)]$$

$$E[y_i(x_i - \mu_1)] = E(y_i E[x_i - \mu_1 | y_i])$$

$$\sum_{y_i=0}^1 y_i E[x_i - \mu_1 | y_i = y_i] \pi_1^{y_i} (1-\pi_1)^{1-y_i} = E[X_i - \mu_1 | Y_i = 1] \pi_1 = 0$$

Then

$$I_n(\mu_0, \mu_1, \sigma^2, \pi_1) = \left[\begin{array}{c} \frac{n(1-\pi_1)}{\sigma^2} \\ \frac{n\pi_1}{\sigma^2} \\ -\left(\frac{n}{2\sigma^2} - \frac{n\pi_1}{\sigma^2} - \frac{n(1-\pi_1)}{\sigma^2} \right) \\ \frac{n}{\pi_1(1-\pi_1)} \end{array} \right]$$

since

$$\begin{aligned} E[Y_i(X_i - \mu_1)^2] &= E(Y_i E[(X_i - \mu_1)^2 | Y_i]) \\ &= \sum_{y_i=0} y_i E[(X_i - \mu_1)^2 | Y_i = y_i] \pi_1^{y_i} (1 - \pi_1)^{1-y_i} \\ &= E[(X_i - \mu_1)^2 | Y_i = 1] \pi_1 = \sigma^2 \pi_1 \end{aligned}$$

and similarly

$$E[(1-Y_i)(X_i - \mu_0)^2] = \sigma^2(1-\pi_1)$$

Then

$$\left[\lim_{n \rightarrow \infty} \frac{1}{n} I_n(\theta) \right]^{-1} = \frac{\sigma^2}{\left[\begin{array}{c} (1-\pi_1)^{-1} \\ \pi_1^{-1} \\ 2\sigma^2 \\ \frac{\pi_1(1-\pi_1)}{\sigma^2} \end{array} \right]}$$

1d) $g(\mu_0, \mu_1, \sigma^2, \pi_1) = \left(\text{logit} \pi_1 + \frac{\mu_0^2 - \mu_1^2}{2\sigma^2}, \frac{\mu_1 - \mu_0}{\sigma^2} \right)$

$$\frac{\partial}{\partial \mu_0} g = \left(-\frac{\mu_0}{\sigma^2}, -\frac{1}{\sigma^2} \right) \quad \frac{\partial}{\partial \mu_1} g = \left(-\frac{\mu_1}{\sigma^2}, \frac{1}{\sigma^2} \right)$$

$$\frac{\partial}{\partial \theta^2} g(\theta) = \left(-\frac{\mu_0^2 - \mu_1^2}{2(\theta^2)^2}, -\frac{\mu_1 - \mu_0}{(\theta^2)^2} \right)$$

$$\frac{\partial}{\partial \pi_2} g(\theta) = \left(\frac{d \log \frac{\pi_2}{1-\pi_2}}{d \pi_2}, 0 \right) = \left(\frac{(1-\pi_2)(\mu_1 - \mu_2)}{\pi_2(1-\pi_2)^2}, 0 \right) = \left(\frac{1}{\pi_2(1-\pi_2)}, 0 \right)$$

$$g(\theta) \in [g(\theta)]'$$

$$= \begin{bmatrix} \frac{1}{\theta^2} & \left[\theta^2 \pi (1-\pi) \right]^{-1} & \left[(1-\pi)^{-1} \right. \\ \begin{pmatrix} \mu_0^2 - \mu_1^2 & -\frac{\mu_0^2 - \mu_1^2}{2\theta^2} \\ -1 & 1 - \frac{\mu_1 - \mu_0}{\theta^2} \end{pmatrix} & 0 & \left. \frac{\pi^{-1}}{2\theta^2} \right] \\ \begin{pmatrix} \frac{\mu_0}{1-\pi_1} & -\frac{\mu_1}{\pi_1} & -\frac{\mu_0^2 - \mu_1^2}{2\theta^2} \\ -\frac{1}{1-\pi_1} & \frac{1}{\pi_1} - 2(\mu_1 - \mu_0) & 0 \end{pmatrix} & \begin{pmatrix} \frac{1}{\theta^2} & \left[\frac{\mu_0}{1-\pi_1} & -1 \\ -\frac{\mu_1}{\pi_1} & 1 & 1 \\ -\frac{\mu_0^2 - \mu_1^2}{2\theta^2} & -\frac{\mu_1 - \mu_0}{\theta^2} & 0 \end{pmatrix} \\ \frac{\pi_1(1-\pi_1)}{\theta^2} & 0 \end{pmatrix} \end{bmatrix} [g(\theta)]'$$

So we have

$$= \begin{bmatrix} \frac{\mu_0^2}{1-\pi_1} + \frac{\mu_1^2}{\pi_1} + \frac{(\mu_0^2 - \mu_1^2)^2}{2\theta^2} + \frac{\theta^2}{\pi_1(1-\pi_1)} & -\frac{\mu_0}{1-\pi_1} - \frac{\mu_1}{\pi_1} + \frac{(\mu_0^2 - \mu_1^2)(\mu_0 - \mu_1)}{\theta^2} \\ -\frac{\mu_0}{1-\pi_1} - \frac{\mu_1}{\pi_1} + \frac{(\mu_0^2 - \mu_1^2)(\mu_0 - \mu_1)}{\theta^2} & \frac{1}{\pi_1(1-\pi_1)} + \frac{2(\mu_1 - \mu_0)}{\theta^2} \end{bmatrix}$$

(6)

e) This implies that $x_i \perp\!\!\!\perp y_i$ since $p(x_i, y_i) = p(x_i|y_i)p(y_i)$
 $= p(x_i)p(y_i)$. so that $\text{logit}(p(y_i=1|x_i)) = \text{logit}(p(y_i=1))$
 $= \text{logit}\pi$

Then

$$\frac{\partial^2}{\partial \mu^2} \ln(\theta) = \frac{\partial}{\partial \mu} \sum_{i=1}^n -\frac{1}{2\sigma^2} (x_i - \mu)^2 = -\frac{n}{\sigma^2}$$

$$\frac{\partial^2}{\partial \sigma^2} \ln(\theta) = \frac{\partial}{\partial \sigma^2} \left\{ -\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (x_i - \mu)^2 \right\}$$

$$= \frac{n}{(\sigma^2)^2} - \frac{1}{(\sigma^2)^3} \sum_{i=1}^n (x_i - \mu)^2 \quad \frac{\partial^2}{\partial \mu \partial \sigma^2} = \frac{1}{(\sigma^2)^2} \sum_{i=1}^n (x_i - \mu)$$

$$\frac{\partial^2}{\partial \pi_1^2} \ln(\theta) = -n \left[\frac{\bar{y}_n - 2\pi_1 \bar{y}_n + \pi_1^2}{\pi_1^2 (1-\pi_1)^2} \right] \quad 0 \text{ otherwise}$$

so that

$$\left[\lim_{n \rightarrow \infty} \frac{1}{n} I_n(\theta) \right]^{-1} = \begin{bmatrix} \frac{1}{\sigma^2} & & \\ & \frac{1}{2\sigma^4} & \\ & & \frac{1}{\pi_1(1-\pi_1)} \end{bmatrix}^{-1} = \begin{bmatrix} \sigma^2 & & \\ & 2\sigma^4 & \\ & & \pi_1(1-\pi_1) \end{bmatrix}$$

Now we want

all of these details. Let $h^{(\theta)} = \text{logit}\pi$ then

$$\frac{\partial}{\partial \theta} h(\theta) = \frac{1-\pi_1}{\pi_1} \cdot \frac{1}{(1-\pi_1)^2} = \frac{1}{\pi_1(1-\pi_1)}$$

~~so that $\hat{h}(\theta) = (0, 0, [\pi, (1-\pi)]^{-1})$ and hence~~

$$\sqrt{n}(\text{logit } \hat{\pi} - \text{logit } \pi) \xrightarrow{D} N\left(0, \frac{1}{\pi(1-\pi)}\right)$$

The model is

$$\text{logit } (\pi_i) = \alpha_0 \Rightarrow \pi_i = \frac{e^{\alpha_0}}{1+e^{\alpha_0}}$$

Then

$$\begin{aligned} \frac{\partial}{\partial \alpha_0} \ln(\alpha_0) &= \frac{\partial}{\partial \alpha_0} \sum_{i=1}^n \left\{ y_i \text{logit } \pi_i + \log(1-\pi_i) \right\} \\ &= \frac{\partial}{\partial \alpha_0} \sum_{i=1}^n \left\{ y_i \alpha_0 - \log(1+e^{\alpha_0}) \right\} = \sum_{i=1}^n \left\{ y_i - \frac{e^{\alpha_0}}{1+e^{\alpha_0}} \right\} \end{aligned}$$

and

$$\frac{\partial^2}{\partial \alpha_0^2} \ln(\alpha_0) = \frac{\partial}{\partial \alpha_0} \sum_{i=1}^n \left\{ y_i - \frac{e^{\alpha_0}}{1+e^{\alpha_0}} \right\} = -n \frac{e^{\alpha_0}}{1+e^{\alpha_0}} \cdot \frac{1}{1+e^{\alpha_0}} = -n \pi_i(1-\pi_i)$$

so that

$$\left[\lim_{n \rightarrow \infty} \frac{1}{n} I_n(\alpha_0) \right]^{-1} = \frac{1}{\pi \pi_i(1-\pi_i)}$$