

$$1e) \quad Y = X_1 \beta_1 + X_2 \beta_2 + \epsilon. \quad \text{Let } X = (X_1, X_2) \text{ and } \beta = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}$$

$$\Rightarrow Y = X\beta + \epsilon. \quad M = X(X'X)^{-1}X' \text{ is an opo onto } X = (X_1, X_2)$$

$$\text{Then, } \hat{\sigma}_{\text{overfit}}^2 = \frac{Y'(I-M)Y}{n-p} \quad \text{where } p = \text{rank}(W) = \text{rank}(X_1, X_2).$$

$$\begin{aligned} \text{Then, } E[\hat{\sigma}_{\text{overfit}}^2] &= \frac{1}{n-p} E[Y'(I-M)Y] = \frac{1}{n-p} \{ E[Y]'(I-M)E[Y] + \text{tr}(\sigma^2(I-M)) \} \\ &= \frac{1}{n-p} \{ (X_1 \beta_1)' \overset{0}{(I-M)} (X_1 \beta_1) + \sigma^2(n-p) \} \\ &= \boxed{\sigma^2} \end{aligned}$$

$$\text{Since } (I-M)X_1 = X_1 - \frac{M X_1}{X_1} = 0. \quad \square$$

$$\begin{aligned} b) \quad \text{Take } Y &= X_1 \beta_1 + X_2 \beta_2 + \epsilon = X_1 \beta_1 + M_2 X_1 \beta_1 - M_2 X_1 \beta_1 + X_2 \beta_2 + \epsilon \\ &= \underline{X_1 \beta_1} + X_2 (X_2' X_2)^{-1} X_2' X_1 \beta_1 - \underline{X_2 (X_2' X_2)^{-1} X_2' X_1 \beta_1} + \underline{X_2 \beta_2} + \epsilon \\ &= (I + M_2) X_1 \beta_1 + X_2 [\beta_2 - (X_2' X_2)^{-1} X_2' X_1 \beta_1] + \epsilon \end{aligned}$$

$$\text{Then, } \hat{\beta}_1 = \left\{ [(I + M_2) X_1]' [(I + M_2) X_1] \right\}^{-1} [(I + M_2) X_1]' Y$$

under true model (2)

$$\text{Then, } E[\hat{\beta}_1] = \left\{ X_1' \underbrace{(I + M_2)' (I + M_2)}_{(I + M_2)} X_1 \right\}^{-1} X_1' (I + M_2) E[Y]$$

$$= \underbrace{\left\{ X_1' (I + M_2) X_1 \right\}^{-1} X_1' (I + M_2) X_1}_{I} \beta_1 = \boxed{\beta_1} \quad \square$$

1.c) Know that $\frac{Y'(I-M)Y}{n-p}$ is the UMVUE of σ^2 . If $Y \sim N(X\beta, \sigma^2 I)$,

and $(I-M)$ is an OPO of rank $n-p$, it follows by a thm given in the slides on "distribution of quadratic forms" that $\frac{1}{\sigma^2} Y'(I-M)Y \sim \chi^2(n-p, \gamma)$

where $\gamma = \frac{(X\beta)'(I-M)(X\beta)}{2\sigma^2} = \frac{\beta'X'(I-M)X\beta}{2\sigma^2} = 0$ since $(I-M)X = 0$.

Thus, $\frac{1}{\sigma^2} Y'(I-M)Y \sim \chi^2(n-p)$

So, for model (1): $\frac{1}{\sigma^2} Y'(I-M)Y \sim \chi^2(r(I-M)) \equiv \chi^2(n-p)$

$\Rightarrow (1-\alpha) \times 100\%$ CI (σ^2) based on (1) is:

$$C_1 = \left\{ \chi^2_{(n-p)}(\alpha/2) \leq \frac{Y'(I-M)Y}{\sigma^2} \leq \chi^2_{(n-p)}(1-\alpha/2) \right\} = \left\{ \frac{Y'(I-M)Y}{\chi^2_{(n-p)}(1-\alpha/2)} \leq \sigma^2 \leq \frac{Y'(I-M)Y}{\chi^2_{(n-p)}(\alpha/2)} \right\}$$

For model (2): $\frac{1}{\sigma^2} Y'(I-M_1)Y \sim \chi^2(r(I-M_1)) \equiv \chi^2(n-p_1)$

$\Rightarrow (1-\alpha) \times 100\%$ CI (σ^2) based on (2) is:

$$C_2 = \left\{ \chi^2_{(n-p_1)}(\alpha/2) \leq \frac{Y'(I-M_1)Y}{\sigma^2} \leq \chi^2_{(n-p_1)}(1-\alpha/2) \right\} = \left\{ \frac{Y'(I-M_1)Y}{\chi^2_{(n-p_1)}(1-\alpha/2)} \leq \sigma^2 \leq \frac{Y'(I-M_1)Y}{\chi^2_{(n-p_1)}(\alpha/2)} \right\}$$

$$\text{The } E[\text{length}(C_1)] = E\left[\frac{Y'(I-M)Y}{\chi^2_{(n-p)}(\alpha/2)} - \frac{Y'(I-M)Y}{\chi^2_{(n-p)}(1-\alpha/2)} \right] = \underbrace{\left[\frac{1}{\chi^2_{(n-p)}(\alpha/2)} - \frac{1}{\chi^2_{(n-p)}(1-\alpha/2)} \right]}_{a_1} E[Y'(I-M)Y]$$

$$= a_1 \{ E[Y'(I-M)Y] + E[Y'(I-M)Y] \}$$

$$= a_1 \{ 2(X\beta)'(I-M)X\beta + 2\text{tr}(\sigma^2(I-M)) \} = 2a_1(n-p)$$

$$a_2 = \frac{1}{\chi^2_{(n-p_1)}(\alpha/2)} - \frac{1}{\chi^2_{(n-p_1)}(1-\alpha/2)}$$

$$\text{Similarly, } E[\text{length}(C_2)] = 2a_2(n-p_1) \Rightarrow E[\text{length}(C_1)] > E[\text{length}(C_2)]$$

$$\text{when } 2a_1(n-p) > 2a_2(n-p_1) \Rightarrow (n-p) \left[\frac{1}{\chi^2_{(n-p)}(\alpha/2)} - \frac{1}{\chi^2_{(n-p)}(1-\alpha/2)} \right] > (n-p_1) \left[\frac{1}{\chi^2_{(n-p_1)}(\alpha/2)} - \frac{1}{\chi^2_{(n-p_1)}(1-\alpha/2)} \right]$$

1 d) (i) Constraint $y_1 + y_2 + y_3 = 180$

where $y_1 = \gamma_1 + \epsilon_1$

$y_2 = \gamma_2 + \epsilon_2$

$y_3 = \gamma_3 + \epsilon_3 = 180 - \gamma_1 - \gamma_2 + \epsilon_3$

$\Leftrightarrow y_1 - 60 = \gamma_1 - 60 + \epsilon_1$

$y_2 - 60 = \gamma_2 - 60 + \epsilon_2$

$y_3 - 60 = 120 - \gamma_1 - \gamma_2 + \epsilon_3$

Let $y = \begin{pmatrix} y_1 - 60 \\ y_2 - 60 \\ y_3 - 60 \end{pmatrix}$, $\gamma = \begin{pmatrix} \gamma_1 - 60 \\ \gamma_2 - 60 \end{pmatrix}$, $\epsilon = \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \end{pmatrix}$, $X = \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ -1 & -1 \end{pmatrix}$

Then, $y = X\beta + \epsilon$

$H_0: \gamma = 0$ vs. $H_1: \gamma \neq 0$

$F = \frac{y'(M - M_0)y / r(M - M_0)}{MSE} \sim F((M - M_0), (I - M), \gamma)$

where $\gamma = \frac{(xy)'(M - M_0)(xy)}{26^2}$, $M_0 = 0$, and $M = X(X'X)^{-1}X' = \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ -1 & -1 \end{pmatrix} \left[\begin{pmatrix} 1 & 0 & -1 \\ 0 & 1 & -1 \end{pmatrix} \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \right]^{-1} \begin{pmatrix} 1 & 0 & -1 \\ 0 & 1 & -1 \end{pmatrix}$

$\stackrel{\text{in R}}{=} \frac{1}{3} \begin{pmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{pmatrix}$

Numerator: Know $\frac{1}{3} \begin{pmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{pmatrix} \begin{pmatrix} y_1 - 60 \\ y_2 - 60 \\ y_3 - 60 \end{pmatrix} = \frac{1}{3} \begin{pmatrix} 2(y_1 - 60) - (y_2 - 60) - (y_3 - 60) \\ -(y_1 - 60) + 2(y_2 - 60) - (y_3 - 60) \\ -(y_1 - 60) - (y_2 - 60) + 2(y_3 - 60) \end{pmatrix} = \frac{1}{3} \begin{pmatrix} 2y_1 - y_2 - y_3 \\ -y_1 + 2y_2 - y_3 \\ -y_1 - y_2 + 2y_3 \end{pmatrix}$

$= \begin{pmatrix} y_1 - \bar{y} \\ y_2 - \bar{y} \\ y_3 - \bar{y} \end{pmatrix}$

Know $y'(M - M_0)y = y'My = (My)'(My)$

$= (y_1 - \bar{y})^2 + (y_2 - \bar{y})^2 + (y_3 - \bar{y})^2$ where $r(M - M_0) = r(M)$

Denominator: $I - M = \frac{1}{3} \begin{pmatrix} 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{pmatrix}$ $(I - M)y = y - My = \begin{pmatrix} y_1 - 60 \\ y_2 - 60 \\ y_3 - 60 \end{pmatrix} - \begin{pmatrix} y_1 - \bar{y} \\ y_2 - \bar{y} \\ y_3 - \bar{y} \end{pmatrix} = \begin{pmatrix} \bar{y} - 60 \\ \bar{y} - 60 \\ \bar{y} - 60 \end{pmatrix}$

Then, $y'(I - M)y = [(I - M)y]'[(I - M)y] = 3(\bar{y} - 60)^2$

Conclude: Thus, $F = \frac{y'My / r(M)}{y'(I - M)y / r(I - M)} = \frac{(y_1 - \bar{y})^2 + (y_2 - \bar{y})^2 + (y_3 - \bar{y})^2}{6(\bar{y} - 60)^2} \sim F(2, 1, \delta)$

$3 - 2 = 1$

Under H_0 : $F \sim F(2, 1)$

Under H_1 : $F \sim F(2, 1, \delta)$ where $\delta = \frac{(xy)'M(xy)}{26^2} = \frac{[M(xy)]'[M(xy)]}{26^2}$

$\downarrow \text{since } M_0 = 0$

where $M(xy) = \frac{1}{3} \begin{pmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{pmatrix} \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ -1 & -1 \end{pmatrix} \begin{pmatrix} y_1 - 60 \\ y_2 - 60 \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ -1 & -1 \end{pmatrix} \begin{pmatrix} y_1 - 60 \\ y_2 - 60 \end{pmatrix} = \begin{pmatrix} y_1 - 60 \\ y_2 - 60 \\ -(y_1 + y_2) + 120 \end{pmatrix}$

$= \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ -1 & -1 \end{pmatrix}$ since $MX = X$

$\Rightarrow \delta = \frac{(y_1 - 60)^2 + (y_2 - 60)^2 + (y_3 - 60)^2}{26^2}$

1d ii) Want to derive joint 95% CR(γ_1, γ_2)

Know $F = \frac{(\hat{\gamma} - x\gamma)'(M - M_0)(\hat{\gamma} - x\gamma)/r(M - M_0)}{MSE} \sim F\left(\frac{r(M - M_0)}{r(M)}, r(I - M)\right)$

Note $\delta = 0$
 ✓ b/c subtracted mean of γ , namely $E[\gamma] = x\gamma$

From i), Know $r(M) = 2$, $r(I - M) = 1$, $MSE = (\bar{y} - 60)^2 / 1 = (\bar{y} - 60)^2$

Need to find $(\hat{\gamma} - x\gamma)'M(\hat{\gamma} - x\gamma) = (\hat{\gamma} - x\gamma)'M(\hat{\gamma} - x\gamma) = (\hat{\gamma} - \gamma)' \underbrace{x'x(x'x)^{-1}x'x}_{(x'x)} (\hat{\gamma} - \gamma)$

$$= (\hat{\gamma} - \gamma)' \begin{pmatrix} 2 & 1 \\ 1 & 2 \end{pmatrix} (\hat{\gamma} - \gamma) = (\gamma - \hat{\gamma})' \begin{pmatrix} 2 & 1 \\ 1 & 2 \end{pmatrix} (\gamma - \hat{\gamma})$$

Then, $F = \frac{(\gamma - \hat{\gamma})' \begin{pmatrix} 2 & 1 \\ 1 & 2 \end{pmatrix} (\gamma - \hat{\gamma}) / 2}{(\bar{y} - 60)^2} \sim F(2, 1)$

$\Rightarrow (1 - \alpha) \times 100\% \text{ CR}(\gamma_1, \gamma_2)$

$= \left\{ \gamma : (\gamma - \hat{\gamma})' \begin{pmatrix} 2 & 1 \\ 1 & 2 \end{pmatrix} (\gamma - \hat{\gamma}) \leq 2(\bar{y} - 60)^2 F(0.95, 2, 1) \right\}$

The 95% percentile of a central $F(2, 1)$ distn.