

b) let $Z_i \sim \text{Gam}(\alpha, \frac{1}{\alpha})$ for $\alpha > 0$. calculate μ_{ij} & σ^2
 write the likelihood for (β, α) & show it can be expressed in closed form via the gamma function

from: $E(Y_{ij}|X_{ij}) = \mu_{ij} = e^{X_{ij}^T \beta} E(Z_i) = e^{X_{ij}^T \beta} \underbrace{\left(\frac{1}{\alpha}\right)}_{E(Z_i)=1} = e^{X_{ij}^T \beta} = \mu_{ij}$

$$\text{var}(Z_i) = \alpha \left(\frac{1}{\alpha}\right)^2 = \frac{1}{\alpha}$$

$$\Rightarrow \sigma^2 = \left(\frac{\sqrt{1/\alpha}}{1}\right)^2 = \left(\frac{1}{\sqrt{\alpha}}\right)^2 = \frac{1}{\alpha} = \sigma^2$$

$$f(y_{ij}, y_{ik}, z_i) = f(y_{ij}, y_{ik} | z_i) f(z_i) = f(y_{ij} | z_i) f(y_{ik} | z_i) f(z_i)$$

$$= \left(\frac{\lambda_{ij}^{y_{ij}} e^{-\lambda_{ij}}}{y_{ij}!} \right) \left(\frac{\lambda_{ik}^{y_{ik}} e^{-\lambda_{ik}}}{y_{ik}!} \right) \left(\frac{\alpha^\alpha}{\Gamma(\alpha)} z_i^{\alpha-1} e^{-\alpha z_i} \right)$$

$$= \left(\frac{e^{y_{ij} X_{ij}^T \beta} e^{-z_i} e^{X_{ij}^T \beta}}{y_{ij}!} \right) \left(\frac{e^{y_{ik} X_{ik}^T \beta} e^{-z_i} e^{X_{ik}^T \beta}}{y_{ik}!} \right) \left(\frac{\alpha^\alpha}{\Gamma(\alpha)} z_i^{\alpha-1} e^{-\alpha z_i} \right)$$

$$= \frac{\alpha^\alpha e^{(y_{ij} X_{ij} + y_{ik} X_{ik})^T \beta}}{y_{ij}! y_{ik}! \Gamma(\alpha)} z_i^{(\alpha + y_{ij} + y_{ik}) - 1} e^{-z_i(\alpha + e^{X_{ij}^T \beta} + e^{X_{ik}^T \beta})}$$

$$\Rightarrow f(y_{ij}, y_{ik}) = \int_0^\infty f(y_{ij}, y_{ik}, z_i) dz_i = \frac{\alpha^\alpha e^{(y_{ij} X_{ij} + y_{ik} X_{ik})^T \beta}}{y_{ij}! y_{ik}! \Gamma(\alpha)} \int_0^\infty z_i^{(\alpha + y_{ij} + y_{ik}) - 1} e^{-z_i(\alpha + e^{X_{ij}^T \beta} + e^{X_{ik}^T \beta})} dz_i$$

$$= \frac{\alpha^\alpha e^{(y_{ij} X_{ij} + y_{ik} X_{ik})^T \beta}}{y_{ij}! y_{ik}! \Gamma(\alpha)} \frac{\Gamma(\alpha + y_{ij} + y_{ik})}{(\alpha + e^{X_{ij}^T \beta} + e^{X_{ik}^T \beta})^{\alpha + y_{ij} + y_{ik}}} \int_0^\infty \frac{(\alpha + e^{X_{ij}^T \beta} + e^{X_{ik}^T \beta})^{\alpha + y_{ij} + y_{ik}}}{\Gamma(\alpha + y_{ij} + y_{ik})} z_i^{(\alpha + y_{ij} + y_{ik}) - 1} e^{-z_i(\alpha + e^{X_{ij}^T \beta} + e^{X_{ik}^T \beta})} dz_i$$

$$\Rightarrow L(\alpha, \beta) = \prod_{i=1}^m f(y_{ij}, y_{ik}) = \prod_{i=1}^m \frac{\alpha^\alpha \Gamma(\alpha + y_{ij} + y_{ik}) e^{(y_{ij} X_{ij} + y_{ik} X_{ik})^T \beta}}{y_{ij}! y_{ik}! \Gamma(\alpha) (\alpha + e^{X_{ij}^T \beta} + e^{X_{ik}^T \beta})^{\alpha + y_{ij} + y_{ik}}}$$

* note $j=1, k=2$
 (w/c $j=1, 2$)

$$b. \mu_{ij} = \exp\{x_{ij}^T \beta\} E[z_i] = \exp\{x_{ij}^T \beta\}$$

$$\gamma^2 = \frac{\text{Var}(z_i)}{E[z_i]^2} = \alpha \left(\frac{1}{\alpha}\right)^2 = \frac{1}{\alpha}$$

$$\begin{aligned} p(y_{i1}, y_{i2}, z_i) &= p(y_{i1}, y_{i2} | z_i) p(z_i) = \frac{\lambda_{i1}^{y_{i1}} e^{-\lambda_{i1}}}{y_{i1}!} \cdot \frac{\lambda_{i2}^{y_{i2}} e^{-\lambda_{i2}}}{y_{i2}!} \cdot \frac{\alpha^\alpha}{\Gamma(\alpha)} z_i^{\alpha-1} e^{-\alpha z_i} \\ &= \frac{(\exp\{x_{i1}^T \beta\} z_i)^{y_{i1}} e^{-\exp\{x_{i1}^T \beta\} z_i}}{y_{i1}!} \cdot \frac{(\exp\{x_{i2}^T \beta\} z_i)^{y_{i2}} e^{-\exp\{x_{i2}^T \beta\} z_i}}{y_{i2}!} \cdot \frac{\alpha^\alpha}{\Gamma(\alpha)} z_i^{\alpha-1} e^{-\alpha z_i} \\ &= \frac{e^{x_{i1}^T \beta y_{i1}}}{y_{i1}!} \cdot \frac{e^{x_{i2}^T \beta y_{i2}}}{y_{i2}!} \cdot \frac{\alpha^\alpha}{\Gamma(\alpha)} z_i^{y_{i1} + y_{i2} + \alpha - 1} e^{(-\exp\{x_{i1}^T \beta\} - \exp\{x_{i2}^T \beta\} - \alpha) z_i} \end{aligned}$$

$$So \quad p(y_{i1}, y_{i2}) = \int p(y_{i1}, y_{i2}, z_i) dz_i = \frac{e^{x_{i1}^T \beta y_{i1}}}{y_{i1}!} \cdot \frac{e^{x_{i2}^T \beta y_{i2}}}{y_{i2}!} \cdot \frac{\alpha^\alpha}{\Gamma(\alpha)} \cdot \frac{\Gamma(y_{i1} + y_{i2} + \alpha)}{(\exp\{x_{i1}^T \beta\} + \exp\{x_{i2}^T \beta\} + \alpha)^{y_{i1} + y_{i2} + \alpha}}$$

$$So \quad L(\alpha, \beta) = \prod_{i=1}^n \frac{\exp\{y_{i1} x_{i1}^T \beta + y_{i2} x_{i2}^T \beta\} \cdot \alpha^\alpha \cdot \Gamma(y_{i1} + y_{i2} + \alpha)}{y_{i1}! y_{i2}! \Gamma(\alpha) [\exp\{x_{i1}^T \beta\} + \exp\{x_{i2}^T \beta\} + \alpha]^{y_{i1} + y_{i2} + \alpha}}$$